

MA212 Probability and Statistics

Chapter Four: Expectation and Variance

Wenwu GONG & Qianqian WANG
Department of Statistics and Data Science

Email: 12031299@mail.sustech.edu.cn
12032005@mail.sustech.edu.cn

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1.1 Basic concepts

Expectation of R.V.

Discrete: Assuming the frequency function of r.v. X is $p(x)$, if

$$\sum_{k=1}^{\infty} |x_k| p_k < +\infty ,$$

Then

$$E(X) \triangleq \sum_{k=1}^{\infty} x_k p_k = \sum_{k=1}^{\infty} x_k P\{X = x_k\}$$

Continuous: Assuming the probability density function of r.v. X is $f(x)$, if

$$\int_{-\infty}^{\infty} |x| f(x) dx < +\infty$$

Then

$$E(X) \triangleq \int_{-\infty}^{\infty} x f(x) dx$$

1.1 Basic concepts (Cont'd)

Expectation of R.V.'s function $Y = g(X)$

- If X is **discrete** with frequency function $p(x)$,

when $\sum_{k=1}^{\infty} |g(x_k)| \cdot p_k < +\infty$, then

$$E(Y) = E[g(X)] = \sum_{k=1}^{\infty} g(x_k) p_k$$

- If X is **continuous** with density function $f(x)$,

when $\int_{-\infty}^{\infty} |g(x)| f(x) dx < \infty$, then

$$E(Y) = E[g(X)] = \int_{-\infty}^{\infty} g(x) f(x) dx$$

1.1 Basic concepts (Cont'd)

Expectation of R.V.s' function $Z = g(X, Y)$

- For the **joint frequency function** of X and Y is $p(i, j)$. If

$$\sum_{i=1}^{\infty} \sum_{j=1}^{\infty} |g(x_i, y_j)| p_{ij} < +\infty, \text{ then}$$

$$E(Z) = E[g(X, Y)] = \sum_{i=1}^{\infty} \sum_{j=1}^{\infty} g(x_i, y_j) p_{ij}$$

- For the **joint density** of X and Y is $f(x, y)$. If

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} |g(x, y)| f(x, y) dx dy < \infty, \text{ then}$$

$$E(Z) = E[g(X, Y)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x, y) f(x, y) dx dy$$

Note: the formula can be generalized to high-dimensional random variables.

1.2 The properties of expectation

Properties

If $a \leq X \leq b$ (a.e), then $a \leq E(X) \leq b$

If c is constant, then $E(cX) = cE(X)$

If X, Y are r.v., then $E(X+Y) = E(X) + E(Y)$

If X, Y are **independent**, then $E(XY) = E(X)E(Y)$

Inferences

If $X = c$ (a.e), then $E(X) = c$

If a_1, a_2, \dots, a_n are constants, X_1, X_2, \dots, X_n are r.v., then $E\left(\sum_{i=1}^n a_i X_i\right) = \sum_{i=1}^n a_i E(X_i)$

If X_1, X_2, \dots, X_n are independent, then

$$E(X_1 X_2 \cdots X_n) = E(X_1)E(X_2) \cdots E(X_n)$$

1.3 Relationship with probability: Inequality

Markov's Inequality

Theorem: If X is a R.V. with $P(X \geq 0) = 1$ and for which $E(X)$ exists, then

$$P\{X \geq t\} \leq \frac{E(X)}{t}.$$

Proof: We will prove this for the **discrete case**. the continuous case is entirely analogous.

$$E(X) = \sum_x xp(x) = \sum_{x < t} xp(x) + \sum_{x \geq t} xp(x)$$

All the terms in the sums are nonnegative because X takes on only nonnegative values. Thus

$$E(X) \geq \sum_{x \geq t} xp(x) \geq \sum_{x \geq t} tp(x) = tP\{X \geq t\}.$$

$$\rightarrow P\{X \geq t\} \leq \frac{E(X)}{t}.$$

1.4 Basic concepts (Cont'd)

Variance of R.V.

The calculation of variance is the expected value of $g(X) = (X - E(X))^2$.

$$\text{Var}(X) \triangleq D(X) \triangleq E(X - E(X))^2$$

- Assuming the **frequency function** of X is $p(x)$, then

$$D(X) = \sum_{k=1}^{\infty} (x_k - E(X))^2 \cdot p_k$$

- Assuming the **probability density** of X is, then

$$D(X) = \int_{-\infty}^{\infty} (x - E(X))^2 f(x) dx$$

More general, you can use the following equation to calculate variance

$$D(X) = E(X - E(X))^2 = E(X^2) - [E(X)]^2$$

1.5 The properties of variance

Properties

If $X = c(\text{constant})$, then $D(X) = 0$

If c is constant, then $D(cX) = c^2 D(X)$

For **independent** r.v X, Y $D(X + Y) = D(X) + D(Y)$

$$D(X - Y) = D(X) + D(Y)$$

If X and Y are not independent

$$D(X \pm Y) = D(X) + D(Y) \pm 2E[(X - E(X))(Y - E(Y))]$$

1.6 Summary of expectation and variances

Some common distributions

| Distribution of X | $E(X)$ | $D(X)$ |
|---------------------------|-------------------|------------------------|
| $X \sim p(\lambda)$ | λ | λ |
| $X \sim b(n, p)$ | np | $np(1 - p)$ |
| $X \sim U(a, b)$ | $\frac{a + b}{2}$ | $\frac{(b - a)^2}{12}$ |
| $X \sim EXP(1/\theta)$ | θ | θ^2 |
| $X \sim N(\mu, \sigma^2)$ | μ | σ^2 |

1.7 Basic concepts (Cont'd)

Covariance of R.V

Definition: If the variance of both X and Y exists, then

$$\text{Cov}(X, Y) \triangleq E[(X - E(X))(Y - E(Y))]$$

$\text{Cov}(X, Y)$ is the Covariance of X and Y .

X, Y are independent $\longrightarrow \text{Cov}(X, Y) = 0$

$\text{Cov}(X, Y) \neq 0 \longrightarrow X, Y$ are not independent
 $\longrightarrow X, Y$ must have a relation

1.8 The properties of Covariance

Properties

- If X, Y are independent, then $\text{Cov}(X, Y) = 0$
- $\text{Cov}(X, Y) = \text{Cov}(Y, X)$
- $\text{Cov}(X, Y) = E(XY) - E(X)E(Y)$
- $D(X + Y) = D(X) + D(Y) + 2E[(X - E(X))(Y - E(Y))]$
 $= D(X) + D(Y) + 2\text{Cov}(X, Y)$
- For any constant a, b
 $\text{Cov}(aX, bY) = E[(aX - E(aX))(bY - E(bY))]$
- $\text{Cov}(X_1 + X_2, Y) = \text{Cov}(X_1, Y) + \text{Cov}(X_2, Y)$

1.9 Basic concepts (Cont'd)

Conditional Expectation

- When $X = x$, the conditional expectation definition of Y is

$$E(Y | X = x) = \sum_y y p_{Y|X}(y | x) \quad (\text{Discrete})$$

$$E(Y | X = x) = \int y f_{Y|X}(y | x) dy \quad (\text{Continuous})$$

- Generally, the conditional expectation of function $h(Y)$ is

$$E[h(Y) | X = x] = \sum_y h(y) p_{Y|X}(y | x) \quad (\text{Discrete})$$

$$E[h(Y) | X = x] = \int h(y) f_{Y|X}(y | x) dy \quad (\text{Continuous})$$

- Theorem**

$$E(Y) = E[E(Y | X)].$$

$$D(Y) = D[E(Y | X)] + E[D(Y | X)].$$

2.1 Extra Exercises 1

Expectation calculation

- The CDF of R.V. X is:

$$F(x) = \begin{cases} \exp(x)/2, & x < 0 \\ 1/2, & 0 \leq x < 1 \\ 1 - 1/2 \exp(-\frac{1}{2}(x-1)), & x \geq 1 \end{cases}$$

Try to calculate $E(X)$.

- R.V. X has $E(X) = \text{Var}(X) = \lambda$. If $E(X-1)(X-2) = 1$, what is λ ?

Expectation and Probability

For non-negative R.V. X and $E(X) < +\infty$, try to prove: (Changeable of summation and integration)

- Discrete: $E(X) = \sum_{k=1}^{+\infty} P(X \geq k)$;
- Continuous: $E(X) = \int_0^{+\infty} P(X > x) dx$.

2.2 Homework1

6. Let X be a continuous random variable with probability density function $f(x) = 2x, 0 \leq x \leq 1$.
- Find $E(X)$.
 - Let $Y = X^2$. Find the probability mass function of Y and use it to find $E(Y)$.
 - Use Theorem A in Section 4.1.1 to find $E(X^2)$ and compare to your answer in part (b).
 - Find $\text{Var}(X)$ according to the definition of variance given in Section 4.2. Also find $\text{Var}(X)$ by using Theorem B of Section 4.2.
15. Suppose that two lotteries each have n possible numbers and the same payoff. In terms of expected gain, is it better to buy two tickets from one of the lotteries or one from each?
21. A random square has a side length that is a uniform $[0, 1]$ random variable. Find the expected area of the square.
30. Find $E[1/(X + 1)]$, where X is a Poisson random variable.

2.2 Homework1

31. Let X be uniformly distributed on the interval $[1, 2]$. Find $E(1/X)$. Is $E(1/X) = 1/E(X)$?

Supplementary Questions:

1. Suppose that the density function of a random variable is

$$f(x) = \begin{cases} 0, & x < 0 \\ e^{-x}, & x \geq 0 \end{cases}$$

Find the expectation of

(1) $Y = 2X$

(2) $Y = e^{-2x}$

2. Suppose that the joint density function of random variable (X, Y) is

$$f(x) = \begin{cases} 12y^2, & 0 < y < x < 1 \\ x, & \text{Otherwise} \end{cases}$$

Compute $E(X), E(Y), E(XY), E(X^2 + Y^2)$

2.3 Extra Exercises 2

Expexctation calculation for functions

- R.V. $X_i \stackrel{\text{iid}}{\sim} U(0, \theta)$, $i = 1, \dots, n$. Let $Y = \max\{X_1, X_2, \dots, X_n\}$, $Z = \min\{X_1, X_2, \dots, X_n\}$, try to calculate $E(Y), E(Z)$. (Distribution of order statistics)

Conditional expexctation

Let X, Y are $\stackrel{\text{iid}}{\sim} \text{Exp}(\lambda)$. We define

$$Z = \begin{cases} 3X + 1, & X \geq Y \\ 6Y, & X < Y \end{cases}$$

Try to calculate $E(Z)$. (Full expectation formula)

2.4 Homework2

49. Two independent measurements, X and Y , are taken of a quantity μ . $E(X) = E(Y) = \mu$, but σ_X and σ_Y are unequal. The two measurements are combined by means of a weighted average to give

$$Z = \alpha X + (1 - \alpha)Y$$

where α is a scalar and $0 \leq \alpha \leq 1$.

- Show that $E(Z) = \mu$.
 - Find α in terms of σ_X and σ_Y to minimize $\text{Var}(Z)$.
 - Under what circumstances is it better to use the average $(X + Y)/2$ than either X or Y alone?
50. Suppose that X_i , where $i = 1, \dots, n$, are independent random variables with $E(X_i) = \mu$ and $\text{Var}(X_i) = \sigma^2$. Let $\bar{X} = n^{-1} \sum_{i=1}^n X_i$. Show that $E(\bar{X}) = \mu$ and $\text{Var}(\bar{X}) = \sigma^2/n$.

2.4 Homework2

54. Let X , Y , and Z be uncorrelated random variables with variances σ_X^2 , σ_Y^2 , and σ_Z^2 , respectively. Let

$$U = Z + X$$

$$V = Z + Y$$

Find $\text{Cov}(U, V)$ and ρ_{UV} .

55. Let $T = \sum_{k=1}^n kX_k$, where the X_k are independent random variables with means μ and variances σ^2 . Find $E(T)$ and $\text{Var}(T)$.
60. Let Y have a density that is symmetric about zero, and let $X = SY$, where S is an independent random variable taking on the values $+1$ and -1 with probability $\frac{1}{2}$ each. Show that $\text{Cov}(X, Y) = 0$, but that X and Y are not independent.

Supplementary Questions:

1. Suppose that X and Y are independent random variables. $E(X) = 3$, $E(Y) = 1$, $D(X) = 4$, $D(Y) = 9$. If $Z = 5X - 2Y + 15$, compute $E(Z)$, $D(Z)$.
2. Suppose that $X_i (i = 1, 2, 3, 4)$ are mutually independent to each other. $(X_i) = 2i$, $D(X_i) = 5 - i$. If $Z = 2X_1 - X_2 + 3X_3 - 0.5X_4$, compute $E(Z)$ and $D(Z)$.

P171: 54, 60 and the extra question

Extra: Assume that the density function for X is

$$f(x) = \frac{1}{2} e^{-|x|}, -\infty < x < \infty$$

- (1) Compute $E(X)$ and $D(X)$.
- (2) Are X and $|X|$ independent or not? State your reason.
- (3) Are X and $|X|$ correlated or not? State your reason

Supplementary Questions

1. Suppose that the joint density function of (X, Y) is

$$f(x, y) = \begin{cases} \frac{x+y}{8}, & 0 \leq x \leq 2, 0 \leq y \leq 2 \\ 0, & \text{otherwise} \end{cases}$$

Compute $E(X)$, $E(Y)$, $Cov(X, Y)$, ρ_{XY} , $D(X + Y)$.

2. X and Y are independent random variables which both follow the normal distribution $N(\mu, \sigma^2)$. If $Z = \alpha X + \beta Y$, $W = \alpha X - \beta Y$, compute $Cov(Z, W)$ and ρ_{ZW} .