Final Project

- Option #1: Research Paper
 - Feb 12: Indicate choice
 - Feb 12: Choose your paper
 - 10 minute presentation (Feb 26)
 - Watch 2 other videos and comment (Mar 5)
 - max 2 page report (Mar 5)
- Option #2: Simulation Study
 - max 2 page report (Mar 5)
 - Simulation code (Mar 5)
 - Simulation can take a couple of hours to run (don't leave to last minute!)

Informative or Random dropout

- Once a subject's tumor diameter exceeds 1.2cm they are removed from the study
- In a breast feeding study, women who have stopped breast feeding are less likely to fill out follow-up surveys
- In a smoking cessation study, those that have started smoking are less likely to report their status

Recap of Day 1

Recap of Day 1

- Joint models are needed when the marker and survival processes are correlated
 - Interested in survival outcome but want to incorporate the effects of a marker measured with error
 - Interested in a longitudinal outcome but have to deal with informative dropout
- Components of a joint model
 - Longitudinal submodel
 - Survival submodel
- Both components are conditional on random effects ("shared random effects" models)

Longitudinal Data Analysis

Day 2

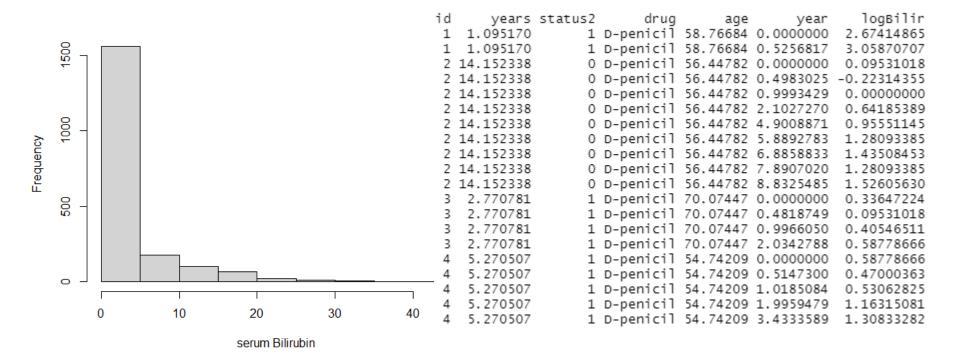
Review of linear mixed-effects models

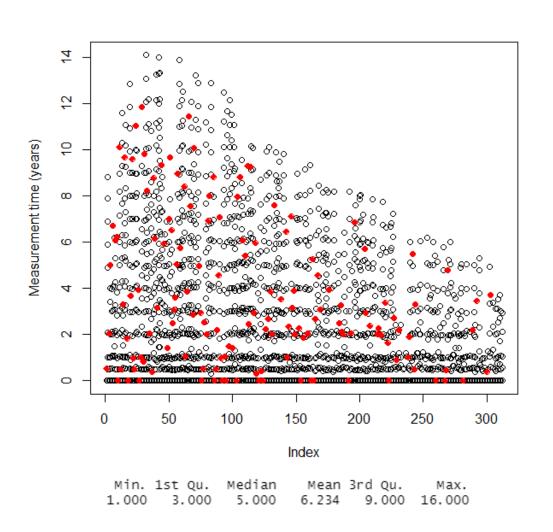
Breakout Session #2 (10 min)

- 1. What are two advantages of linear mixed effects models for modeling longitudinal data?
- 2. In what situations would you use REML (restricted maximum likelihood) to estimate parameters for a linear mixed model and why?
- 3. Why or why should we not consider an unstructured covariance matrix in our PBC data example?
- 4. What is your current TV show/movie recommendation?

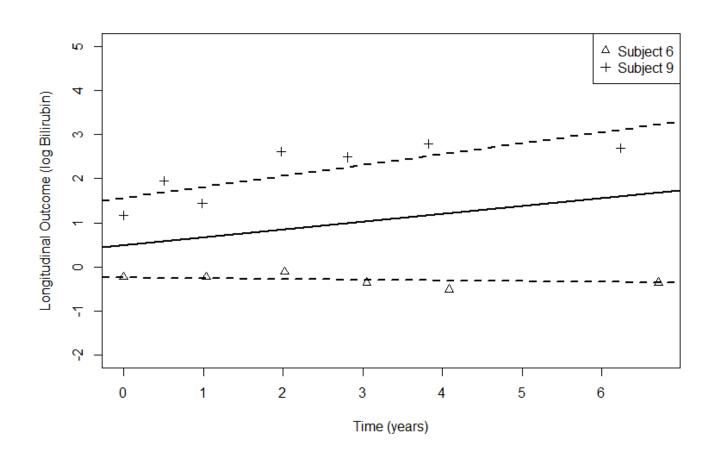
Features of Longitudinal Data

- Repeated evaluations of the same outcome in a subject over time
 - Serum bilirubin in PBC patients
- Interested in investigating:
 - How treatment means differ at specific time points (crosssectional effect)
 - How treatment means or differences between treatment means change over time (longitudinal effect)
- Measures on the same subject are expected to be (positively) correlated
 - Need to use appropriate statistical methods to account for this





• Each individual in the population has their own subjectspecific mean response profile over time



 The evolution of each subject's response in time can be described by a linear model

$$y_{ij} = \tilde{\beta}_{i0} + \tilde{\beta}_{i1}t_{ij} + \epsilon_{ij}, \quad \epsilon_{ij} \sim \mathcal{N}(0, \sigma^2)$$

- y_{ij} is the jth response for the ith subject
- $\tilde{\beta}_{i0}$ is the intercept and $\tilde{\beta}_{i1}$ is the slope for subject i
- Assume that since subjects are sampled from a population, subject-specific coefficients are sampled from a population of regression coefficients

$$\tilde{\beta}_i \sim \mathcal{N}(\beta, D)$$

where $\beta = (\beta_0, \beta_1)'$ and D is the variance-covariance matrix

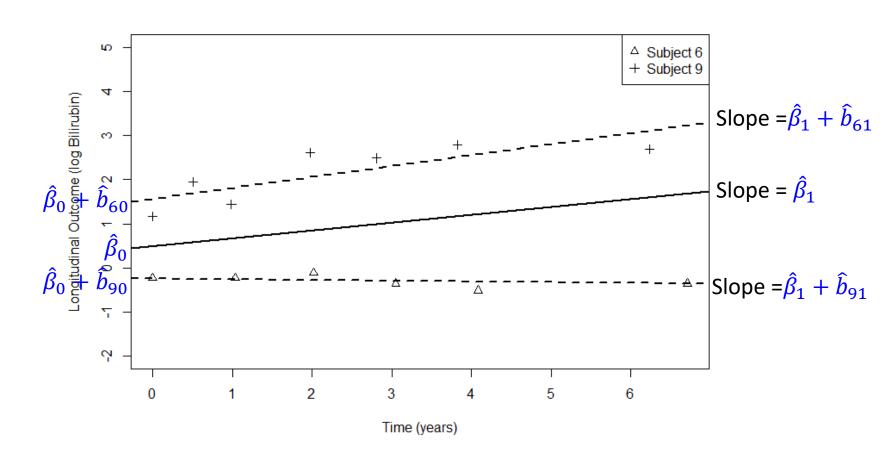
We can reformulate the model as

$$y_{ij} = (\beta_0 + b_{i0}) + (\beta_1 + b_{i1})t_{ij} + \epsilon_{ij}, \quad \epsilon_{ij} \sim N(0, \sigma^2)$$

- β s are the fixed effects
- b_i s are the random effects
- For the random effects, we assume

$$b_i = \begin{bmatrix} b_{i0} \\ b_{i1} \end{bmatrix} \sim \mathcal{N}(0, D) \qquad D = \begin{pmatrix} \sigma_0^2 & \sigma_{01} \\ \sigma_{01} & \sigma_1^2 \end{pmatrix}$$

• Each individual in the population has their own subjectspecific mean response profile over time



In general form

$$\begin{cases} y_i = X_i \beta + Z_i b_i + \epsilon_i, \\ b_i \sim \mathcal{N}(0, D), \\ \epsilon_i \sim \mathcal{N}(0, \sigma^2 I_{n_i}) \end{cases}$$

- X is the design matrix for the fixed effects β
- Z is the design matrix for the random effects b_i
- $b_i \perp \!\!\! \perp \epsilon_i$

Interpretation:

- β_j is the change in the average y_i when x_j is increased by one unit
- b_i represents how a subset of the regression parameters for the ith subject deviates from those in the population
- Advantage: Population + Subject-specific predictions
- β is the mean response changes in the population
- $\beta + b_i$ describes individual response trajectories

Benefits for Joint modeling:

- Can reconstruct the complete path of an individual's timedependent process
- Accommodates unbalanced data
- Accommodates correlation between repeated measurements (parsimonious)

Random effects

- How do the random effects capture correlation?
- (Conditional independence assumption) Given the random effects, the measurements of each subject are independent

$$p(y_i|b_i;\theta) = \prod_{j=1}^{n_i} p(y_{ij}|b_i;\theta)$$

 Marginally (integrating out the random effects), the measurements of each subject are correlated

$$p(y_i) = \int p(y_i|b_i)p(b_i)db_i$$

$$\Rightarrow y_i \sim \mathcal{N}(X_i\beta, Z_iDZ_i' + \sigma^2 I_{n_i})$$

Random effects

- Can consider more general covariances matrices for the subject-specific errors $\epsilon_i \sim N(0, \Sigma_i)$
- The corresponding marginal model is of the form

$$y_i \sim \mathcal{N}(X_i\beta, Z_iDZ_i' + \Sigma_i)$$

- This model does not assume conditional independence
- Both b_i and Σ_i try to capture the correlation in the observed responses y_i
- Philosophical choice: Random effects vs. Serial correlation

Random effects

- With random effects we model the correlations in the repeated measurements of each subject
- In using random effects for modeling the covariance matrix
 - The more random effects we include the more flexibly we capture the correlations
 - By using random effects (other than random intercept alone) we also directly allow for heteroscedasticity (i.e., non-constant variances over time)
 - We do assume a particular type of structure for the correlations and the variances (they are not allowed completely free)
 - Random effects work equally well with balanced or unbalanced data

Estimation

 Assuming independence across subjects, the log-likelihood of the linear mixed model is given by

$$l(\theta) = \sum_{i=1}^{n} \log p(y_i; \theta)$$

$$= \sum_{i=1}^{n} \log \int p(y_i|b_i; \beta, \sigma^2) p(b_i; \theta_b) db_i$$

where
$$\theta' = (\beta', \sigma^2, \theta'_b)$$
 and $\theta_b = \operatorname{vech}(D)$

Estimation

Estimation is based on Maximum Likelihood (ML)

$$p(y_i) = \int p(y_i|b_i)p(b_i)db_i \Rightarrow y_i \sim \mathcal{N}(X_i\beta, V_i)$$
$$V_i = Z_i DZ_i' + \sigma^2 I_{n_i}$$

Then

$$p(y_i; \theta) = (2\pi)^{-n_i/2} |V_i|^{-1/2} \exp\left\{-\frac{1}{2}(y_i - X_i\beta)' V_i^{-1}(y_i - X_i\beta)\right\}$$

where $\theta' = (\beta', \sigma^2, \theta_b')$ and $\theta_b = \text{vech}(D)$

Estimation

• Fixed effects: For known marginal covariance matrix $V_i = Z_i D Z_i' + \sigma^2 I_{n_i}$, the fixed effects are estimated using generalized least squares

$$\hat{\beta} = \left(\sum_{i=1}^{n} X_i' V_i^{-1} X_i\right)^{-1} \sum_{i=1}^{n} X_i' V_i^{-1} y_i$$

- Variance components: The unique parameters in V_i are estimated based on either
 - maximum likelihood (ML) profile likelihood
 - restricted maximum likelihood (REML)
 - REML provides unbiased estimates for the variance components in small samples

Hypothesis Tests

- Test of random effects
 - With REML, must the have same fixed effects to compare models
 - Likelihood ratio tests -> mixture of chi-square distributions
 - Compare with AIC
- Test of fixed effects
 - Must use ML to compare models with different fixed effects
 - REML depends on the fixedeffects design matrix

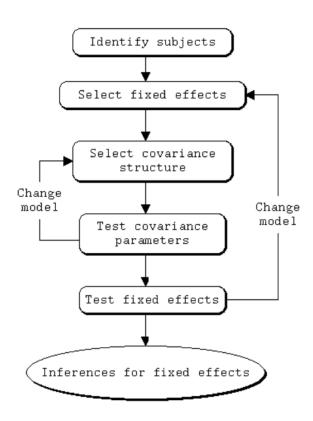


Figure 2. Repeated Measures Analysis in PROC MIXED

Mixed-Effects Models in R

nlme

- Fits linear and nonlinear mixed effects models, and marginal models for normal data
- Allows for both random effects and correlated error terms
- Several options for covariance matrices and variance functions

Ime4

- Fits linear, nonlinear and generalized mixed effects models
- Uses only random effects
- Allows for nested and crossed random-effects designs

Mixed-Effects Models in R

- We will only use package nlme because the joint modeling package JM accepts that as an argument
- The basic function to fit linear mixed models is Ime() and has three basic arguments
 - fixed: a formula specifying the response vector and the fixed-effects structure
 - random: a formal specifying the random-effects structure
 - data: data frame containing all the variables

- We fit a linear mixed model for the PBC dataset to describe the evolution of serum bilirubin
- We will fit the specific model assuming:
 - Different average longitudinal evolutions per treatment group (fixed part)
 - Random intercepts and random slopes (random part)

 Fit a linear mixed effects model for log serum bilirubin, assuming simple linear evolutions in time for each subject (i.e., random intercept, random slope) and different average evolutions per treatment group

$$y_i(t) = \beta_0 + \beta_1 t + \beta_2 \{ \text{D-penic}_i \times t \} + b_{i0} + b_{i1} t + \epsilon_i(t)$$

Note: We did not include a main effect for treatment due to randomization

fixed

random intercept + random slope

```
Linear mixed-effects model fit by REML
 Data: pbc2
                         logLik
       AIC
                  BIC
  3082.323 3121.323 -1534.161
Random effects:
 Formula: ~year | id
 Structure: General positive-definite, Log-Cholesky parametrization
        \sigma_0 StdDev Corr
(Intercept) 0.9990381 (Intr)
        ept) 0.9990381 (Intr) \sigma_{01} 0.1722826 0.417 \sigma_{00} 0.3489259 \sigma_{00}
Residual 0.3489259 \sigma
Fixed effects: log(serBilir) ~ year + drug:year
                eta_0 Value Std.Error DF t-value p-value 0.4956686 0.05807539 1631 8.534915 0.0000
                \beta_1 0.1761726 0.01754759 1631 10.039704 0.0000
year
year:drugD-penicil 0.0027708 0.02411083 1631 0.114920 0.9085
 Correlation: \beta_2
                     (Intr) year
                      0.177
year
year:drugD-penicil 0.002 -0.705
Standardized Within-Group Residuals:
         Min
                       01
                                   Med
                                                  Q3
                                                               Max
-4.31683669 -0.49673826 -0.01978638 0.45207679 5.28538333
Number of Observations: 1945
Number of Groups: 312
```

 Under the random-intercept, random-slope model, the implied marginal model is

$$y_i = X_i \beta + \epsilon_i, \quad \epsilon_i \sim \mathcal{N}(0, Z_i D Z_i' + \sigma^2 I_{n_i})$$

 And the marginal covariance function for any pair of responses on the same individual is of the form

$$Cov(y_{ij}, y_{ij'}) = \sigma_1^2 t_{ij} t_{ij'} + \sigma_{01} (t_{ij} + t_{ij'}) + \sigma_0^2$$

$$Var(y_{ij}) = \sigma_1^2 t_{ij}^2 + 2\sigma_{01} t_{ij} + \sigma_0^2 + \sigma^2$$

• The estimated marginal covariance matrix (from a random intercept and slope model) for Patient 6 in our data set is:

```
id drug year serBilir
6 placebo 0.000000 0.8
6 placebo 1.034936 0.8
6 placebo 2.017851 0.9
6 placebo 3.052787 0.7
6 placebo 4.084985 0.6
6 placebo 6.716132 0.7
```

```
margCov.p1 <- getVarCov(lmeFit.p1, individuals = 6, type = "marginal")</pre>
margCov.p1
## id 6
## Marginal variance covariance matrix
## 1 1.1198 1.0724 1.1430 1.2173 1.2914 1.4803
## 2 1.0724 1.3002 1.2792 1.3853 1.4912 1.7609
## 3 1.1430 1.2792 1.5304 1.5450 1.6809 2.0274
## 4 1.2173 1.3853 1.5450 1.8348 1.8807 2.3080
## 5 1.2914 1.4912 1.6809 1.8807 2.2017 2.5879
## 6 1.4803 1.7609 2.0274 2.3080 2.5879 3.4230
     Standard Deviations: 1.0582 1.1403 1.2371 1.3545 1.4838 1.8501
cov2cor(margCov.p1[[1]])
## 1 1.0000000 0.8887190 0.8730637 0.8492049 0.8224226 0.7560682
## 2 0.8887190 1.0000000 0.9068515 0.8969157 0.8813207 0.8346721
## 3 0.8730637 0.9068515 1.0000000 0.9219747 0.9157065 0.8857795
## 4 0.8492049 0.8969157 0.9219747 1.0000000 0.9357149 0.9209516
## 5 0.8224226 0.8813207 0.9157065 0.9357149 1.0000000 0.9426686
## 6 0.7560682 0.8346721 0.8857795 0.9209516 0.9426686 1.0000000
```

 The estimated marginal covariance matrix (from a random intercept model) for Patient 6 in our data set is:

```
margCov.p0 <- getVarCov(lmeFit.p0, individuals = 6, type = "marginal")
margCov.p0
## id 6
## Marginal variance covariance matrix
## 1 1.4375 1.1955 1.1955 1.1955 1.1955 1.1955
## 2 1.1955 1.4375 1.1955 1.1955 1.1955 1.1955
## 3 1.1955 1.1955 1.4375 1.1955 1.1955 1.1955
## 4 1.1955 1.1955 1.1955 1.4375 1.1955 1.1955
## 5 1.1955 1.1955 1.1955 1.1955 1.4375 1.1955
## 6 1.1955 1.1955 1.1955 1.1955 1.1955 1.4375
     Standard Deviations: 1.199 1.199 1.199 1.199 1.199
cov2cor(margCov.p0[[1]])
## 1 1.0000000 0.8316326 0.8316326 0.8316326 0.8316326 0.8316326
## 2 0.8316326 1.0000000 0.8316326 0.8316326 0.8316326 0.8316326
## 3 0.8316326 0.8316326 1.0000000 0.8316326 0.8316326 0.8316326
## 4 0.8316326 0.8316326 0.8316326 1.0000000 0.8316326 0.8316326
## 5 0.8316326 0.8316326 0.8316326 0.8316326 1.0000000 0.8316326
```

6 0.8316326 0.8316326 0.8316326 0.8316326 0.8316326 1.0000000

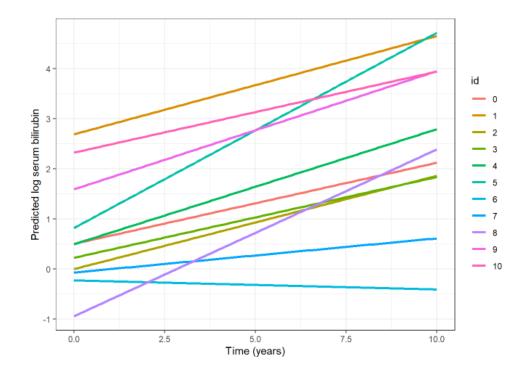
 We can extract and plot predicted longitudinal evolutions from a mixed model

```
> newdat <- expand.grid(id=1:2, year=c(0.5,1.5), drug=c("D-penicil", "placebo"))</pre>
> newdat$pred <- predict(lmeFit.p0, newdata=newdat)</pre>
> newdat
  id year drug
1 1 0.5 D-penicil 2.6815376
2 2 0.5 D-penicil 0.4094144
3 1 1.5 D-penicil 2.7815562
4 2 1.5 D-penicil 0.5094330
5 1 0.5 placebo 2.6764478
6 2 0.5 placebo 0.4043246
7 1 1.5 placebo 2.7662868
8 2 1.5 placebo 0.4941635
> newdat <- expand.grid(year=c(0.5,1.5), drug=c("D-penicil", "placebo"))</pre>
> newdat$pred <- predict(lmeFit.p0, newdata=newdat,level=0)</pre>
> newdat
       drug pred
  year
1 0.5 D-penicil 0.6207185
2 1.5 D-penicil 0.7207371
3 0.5 placebo 0.6156286
4 1.5
         placebo 0.7054676
```

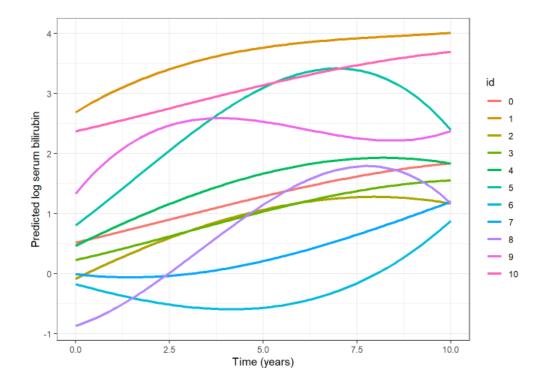
 Random intercepts and random slopes, with a diagonal covariance matrix (using the pdDiag() function)

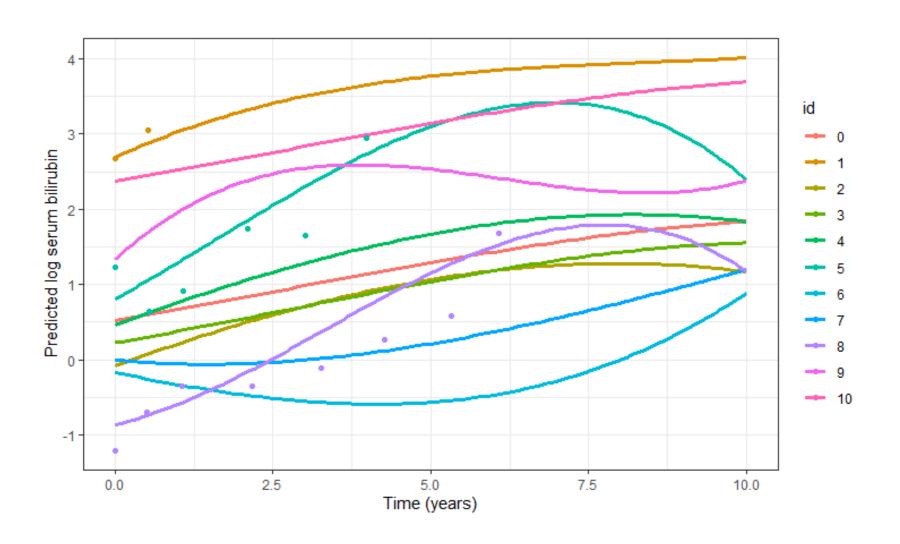
```
> lmeFit.p2 <- lme(log(serBilir) ~ year+drug:year, data = pbc2,</p>
                    random = list(id = pdDiag(form = \sim year))
> lmeFit.p2
Linear mixed-effects model fit by REML
  Data: pbc2
  Log-restricted-likelihood: -1545.702
  Fixed: log(serBilir) ~ year + drug:year
       (Intercept) year year:drugD-penicil
.5018477393 0.1615161046 0.0009132756
      0.5018477393
                                              0.0009132756
Random effects:
 Formula: ~year | id
 Structure: Diagonal
        (Intercept) year Residual
StdDev: 1.022807 0.1730064 0.3477198
Number of Observations: 1945
Number of Groups: 312
```

 Can consider different functions of time for describing the longitudinal trajectory

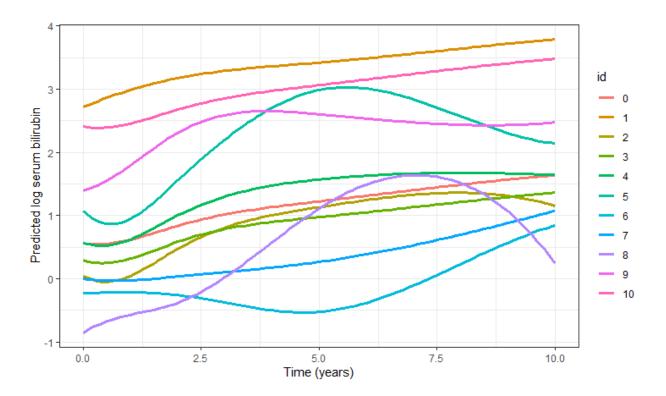


Model trajectory over time more flexibly using splines

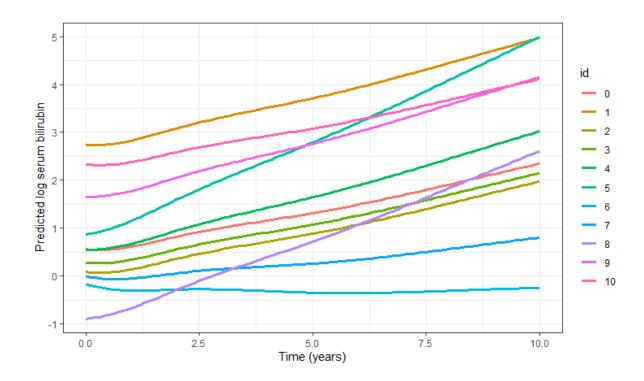




Model trajectory over time more flexibly using splines



 Model trajectory over time more flexibly using splines (basis, natural, etc.)



Breakout Session #2 (10 min)

- 1. What are two advantages of linear mixed effects models for modeling longitudinal data?
- 2. In what situations would you use REML (restricted maximum likelihood) to estimate parameters for a linear mixed model and why?
- 3. Why or why should we not consider an unstructured covariance matrix in our PBC data example to have increased flexibility?
- 4. What is your current TV show/movie recommendation?