Relationship between SVD and PCA. How to use SVD to perform PCA?

Principal component analysis (PCA) is usually explained via an eigen-decomposition of the covariance matrix. However, it can also be performed via singular value decomposition (SVD) of the data matrix X. How does it work? What is the connection between these two approaches? What is the relationship between SVD and PCA?

Or in other words, how to use SVD of the data matrix to perform dimensionality reduction?

pca dimensionality-reduction matrix svd





- 6 I wrote this FAQ-style question together with my own answer, because it is frequently being asked in various forms, but there is no canonical thread and so closing duplicates is difficult. Please provide meta comments in this accompanying meta thread. amoeba Jan 22 '15 at 11:25
- 2 stats.stackexchange.com/questions/177102/... kjetil b halvorsen Feb 3 '16 at 10:45
- 2 In addition to an excellent and detailed amoeba's answer with its further links I might recommend to check this, where PCA is considered side by side some other SVD-based techniques. The discussion there presents algebra almost identical to amoeba's with just minor difference that the speech there, in describing PCA, goes about svd decomposition of \mathbf{X}' , \sqrt{n} [or \mathbf{X}' , $\sqrt{n-1}$] instead of \mathbf{X}' which is simply convenient as it relates to the PCA done via the eigendecomposition of the covariance matrix. trophns Feb 3 '16 at 12:18

PCA is a special case of SVD, PCA needs the data normalized, ideally same unit. The matrix is nxn in PCA, - Oryar Koryar Oct 17 '17 at 9:12

@OrvarKorvar: What n x n matrix are you talking about ? - Chhihe Mar 29 at 15:16

3 Answers

Let the data matrix \mathbf{X} be of $n \times p$ size, where n is the number of samples and p is the number of variables. Let us assume that it is *centered*, i.e. column means have been subtracted and are now equal to zero.

Then the $p \times p$ covariance matrix C is given by $C = \mathbf{X}^T \mathbf{X} / (n-1)$. It is a symmetric matrix and so it can be diagonalized:

$$C = VLV^{\mathsf{T}}$$
,

where ${\bf V}$ is a matrix of eigenvectors (each column is an eigenvector) and ${\bf L}$ is a diagonal matrix with eigenvalues λ_i in the decreasing order on the diagonal. The eigenvectors are called *principal axes* or *principal directions* of the data. Projections of the data on the principal axes are called *principal components*, also known as *PC scores*; these can be seen as new, transformed, variables. The j-th principal component is given by j-th column of ${\bf XV}$. The coordinates of the i-th data point in the new PC space are given by the i-th row of ${\bf XV}$.

If we now perform singular value decomposition of X, we obtain a decomposition

$$X = USV^{T}$$
.

where ${\bf U}$ is a unitary matrix and ${\bf S}$ is the diagonal matrix of singular values s_i . From here one can easily

$$\mathbf{C} = \mathbf{V}\mathbf{S}\mathbf{U}^{\mathsf{T}}\mathbf{U}\mathbf{S}\mathbf{V}^{\mathsf{T}}/(n-1) = \mathbf{V}\frac{\mathbf{S}^{2}}{n-1}\mathbf{V}^{\mathsf{T}},$$

meaning that right singular vectors ${\bf V}$ are principal directions and that singular values are related to the eigenvalues of covariance matrix via $\lambda_i=s_i^2/(n-1)$. Principal components are given by

$$XV = USV^TV = US$$

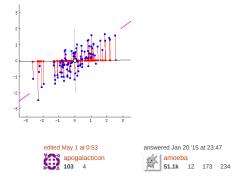
To summarize:

- 1. If $X = USV^{\top}$, then columns of V are principal directions/axes.
- 2. Columns of US are principal components ("scores").
- 3. Singular values are related to the eigenvalues of covariance matrix via $\lambda_i=s_i^2/(n-1)$. Eigenvalues λ_i show variances of the respective PCs.
- 4. Standardized scores are given by columns of $\sqrt{n-1}\mathbf{U}$ and loadings are given by columns of $\mathbf{VS}/\sqrt{n-1}$. See e.g. here and here for why "loadings" should not be confused with principal directions.
- 5. The above is correct only if X is centered. Only then is covariance matrix equal to $X^TX/(n-1)$.
- 6. The above is correct only for X having samples in rows and variables in columns. If variables are in rows and samples in columns, then U and V exchange interpretations.
- 7. If one wants to perform PCA on a correlation matrix (instead of a covariance matrix), then columns of X should not only be centered, but standardized as well, i.e. divided by their standard deviations.
- 8. To reduce the dimensionality of the data from p to k < p, select k first columns of \mathbf{U} , and $k \times k$ upper-left part of \mathbf{S} . Their product $\mathbf{U}_k\mathbf{S}_k$ is the required $n \times k$ matrix containing first k PCs.
- 9. Further multiplying the first k PCs by the corresponding principal axes $\mathbf{V}_k^{\mathsf{T}}$ yields $\mathbf{X}_k = \mathbf{U}_k \mathbf{S}_k \mathbf{V}_k^{\mathsf{T}}$ matrix that has the original $n \times p$ size but is of lower rank (of rank k). This matrix \mathbf{X}_k provides a reconstruction of the original data from the first k PCs. It has the lowest possible reconstruction error, see my answer here.
- 10. Strictly speaking, \mathbf{U} is of $n \times n$ size and \mathbf{V} is of $p \times p$ size. However, if n > p then the last n p columns of \mathbf{U} will be zeros; one should therefore use an "economy size" (or "thin") SVD that returns \mathbf{U} of $n \times p$ size, dropping the zero columns. For large $n \gg p$ the matrix \mathbf{U} would otherwise be unnecessarily huge. The same applies for an opposite situation of $n \ll p$.

Further links

• What is the intuitive relationship between SVD and PCA -- a very nonular and very similar thread

- PCA and Correspondence analysis in their relation to Biplot -- PCA in the context of some congeneric techniques, all based on SVD.
- Is there any advantage of SVD over PCA? -- a question asking if there any benefits in using SVD instead of PCA [short answer: ill-posed question].
- Making sense of principal component analysis, eigenvectors & eigenvalues -- my answer giving a non-technical explanation of PCA. To draw attention, I reproduce one figure here:



 ± 1 for both Q&A. Thanks for sharing. I have one question: why do you have to assume that the data matrix is centered initially? \pm Antoine Aug 6 '15 at 8:39

after reading your summary point #5, I restate my question as: why is the covariance matrix equal to this quantity only if the data matrix is centered initially? – Antoine Aug 6 '15 at 8:58

- 4 @Antoine, covariance matrix is by definition equal to $\langle (\mathbf{x}_i \bar{\mathbf{x}})(\mathbf{x}_i \bar{\mathbf{x}})^{\top} \rangle$, where angle brackets denote average value. If all \mathbf{x}_i are stacked as rows in one matrix \mathbf{X} , then this expression is equal to $(\mathbf{X} \bar{\mathbf{X}})(\mathbf{X} \bar{\mathbf{X}})^{\top}/(n-1)$. If \mathbf{X} is centered then it simplifies to $\mathbf{X}^{\top}/(n-1)$. Think of variance, it's equal to $\langle (\mathbf{x}_i \bar{\mathbf{x}})^2 \rangle$. But if $\bar{\mathbf{x}} = \mathbf{0}$ (i.e. data are centered), then it's simply the average value of \mathbf{x}_i^2 . amoeba Aug 6 '15 at 9:43
- 1 A code sample for PCA by SVD: stackoverflow.com/questions/3181593/... optimist Mar 23 '16 at 11:51
- 2 @amoeba +1 nice and clear math and logic flow, exactly what I needed. hxd1011 Jun 16 '16 at 14:04

I wrote a Python & Numpy snippet that accompanies @amoeba's answer and I leave it here in case it is useful for someone. The comments are mostly taken from @amoeba's answer.

```
import numpy as np
from numpy import linalg as la
np.random.seed(42)
def flip_signs(A, B):
        utility function for resolving the sign ambiguity in SVD http://stats.stackexchange.com/q/34396/115202
        signs = np.sign(A) * np.sign(B)
return A, B * signs
# Let the data matrix X be of n x p size, 
# where n is the number of samples and p is the number of variables n, p = 5, 3 X = np.random.rand(n, p) 
# Let us assume that it is centered 
X -= np.mean(X, axis=0)
 # the n x n covariance matrix
# the p x p covariance matrix
C = np.cov(x, rowvar=False)
print "C = \n", c
# C is a symmetric matrix and so it can be diagonalized:
1, principal_axes = la.eig(C)
# sort results wrt. eigenvalues
idx = l.argort([::-1]
1, principal_axes = l[idx], principal_axes[:, idx]
# the discovarialues in decreasing order.
l, principal_axes = l[idx], principal_axes[:, idx]
# the eigenvalues in decreasing order
print "l = \n", 1
# a matrix of eigenvectors (each column is an eigenvector)
print "V = \n", principal_axes
# projections of X on the principal axes are called principal components
principal_components = X.dot(principal_axes)
print "Y = \n", principal_components
# we now perform singular value decomposition of X
# "economy size" (or "thin") SVD
U, s, Vt = la.svd(X, full_matrices=False)
V = Vt.T
S = np.diag(s)
# 1) then columns of V are principal directions/axes
assert np.allclose(*flip_signs(V, principal_axes))
# 2) columns of US are principal components
assert np.allclose(*flip_signs(U.dot(S), principal_components))
\# 3) singular values are related to the eigenvalues of covariance matrix assert np.allclose((s ** 2) / (n - 1), l)
# 8) dimensionality reduction
# 10) we used "economy size" (or "thin") SVD
assert U.shape == (n, p)
assert S.shape == (p, p)
assert V.shape == (p, p)
                                                                                edited Oct 16 '16 at 9:52
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user115202
```

6/5/2018

the data to make a matrix

$$X = \begin{pmatrix} & & x_1^T - \mu^T \\ & & x_2^T - \mu^T \\ & & \vdots \\ & & x_n^T - \mu^T \end{pmatrix}$$

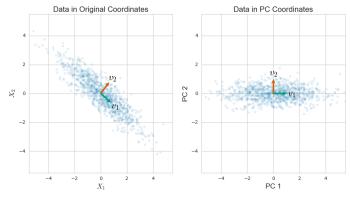
The covariance matrix

$$S = \frac{1}{n-1} \sum_{i=1}^{n} (x_i - \mu)(x_i - \mu)^T = \frac{1}{n-1} X^T X$$

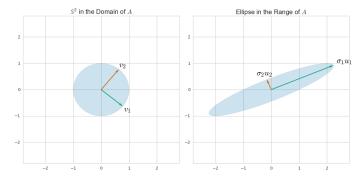
measures to which degree the different coordinates in which your data is given vary together. So, it's maybe not surprising that PCA -- which is designed to capture the variation of your data -- can be given in terms of the covariance matrix. In particular, the eigenvalue decomposition of S turns out to be

$$S = V\Lambda V^T = \sum_{i=1}^r \lambda_i \nu_i \nu_i^T,$$

where v_i is the i-th Principal Component, or PC, and λ_i is the i-th eigenvalue of S and is also equal to the variance of the data along the i-th PC. This decomposition comes from a general theorem in linear algebra, and some work does have to be done to motivate the relatino to PCA.



SVD is a general way to understand a matrix in terms of its column-space and row-space. (It's a way to rewrite any matrix in terms of other matrices with an intuitive relation to the row and column space.) For example, for the matrix $A = \begin{pmatrix} 1 & 2 \\ 0 & 1 \end{pmatrix}$ we can find directions u_i and v_i in the domain and range so that



You can find these by considering how A as a linear transformation morphs a unit sphere $\mathbb S$ in its domain to an ellipse: the principal semi-axes of the ellipse align with the u_i and the v_i are their preimages.

In any case, for the data matrix X above (really, just set A = X), SVD lets us write

$$X = \sum_{i=1}^{r} \sigma_i u_i v_j^T ,$$

where $\{u_i\}$ and $\{v_i\}$ are orthonormal sets of vectors. A comparison with the eigenvalue decomposition of S reveals that the "right singular vectors" v_i are equal to the PCs, the "right singular vectors" are

$$u_i = \frac{1}{\sqrt{(n-1)\lambda_i}} X v_i \,,$$

and the "singular values" σ_i are related to the data matrix via

$$\sigma_i^2 = (n-1)\lambda_i.$$

It's a general fact that the right singular vectors u_i span the column space of X. In this specific case, u_i give us a scaled projection of the data X onto the direction of the i-th principal component. The left singular vectors v_i in general span the row space of X, which gives us a set of orthonormal vectors that spans the data much like PCs.

I go into some more details and benefits of the relationship between PCA and SVD in this longer article.

protected by amoeba Oct 17 '17 at 9:15

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