

# Meta-analysis and synthesis

## Bayesian Modeling for Socio-Environmental Data

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# Why this course?

SESYNC is dedicated to fostering synthetic, actionable science related to the structure, functioning, and sustainability of socio-environmental systems.



Synthesis: gaining new insight by combining information from multiple sources in novel ways.

- ▶ Multiple sources of data can enter through likelihoods.
- ▶ Results of previous studies can enter through priors.
- ▶ Hierarchical models allow “borrowing strength.”

## Multiple sources of data

$$[\boldsymbol{\theta} \mid \mathbf{y}_1, \mathbf{y}_2, \mathbf{y}_3] \propto \prod_{i=1}^{n_1} [y_{1,i} \mid \boldsymbol{\theta}] \prod_{i=1}^{n_2} [y_{2,i} \mid \boldsymbol{\theta}] \prod_{i=1}^{n_3} [y_{3,i} \mid \boldsymbol{\theta}] [\boldsymbol{\theta}] \quad (1)$$

- ▶ Can have different  $\theta'$ s in the different likelihoods.
- ▶ Likelihoods can have different distributions.

## Example

- ▶ Deterministic model of process:  $g(\boldsymbol{\theta}, \mathbf{z}_{t-1}) = \mathbf{A}\mathbf{z}_{t-1}$
- ▶ Stochastic model of process:

$$\log(\mathbf{z}_t) \sim \text{multivariate normal}(\log(g(\boldsymbol{\theta}, \mathbf{z}_{t-1})), \sigma_p^2 \mathbf{I})$$

- ▶ Posterior and joint distribution:

$$\begin{aligned} [\boldsymbol{\theta}, \mathbf{z}, \sigma_p^2 | \mathbf{y}_1, \mathbf{Y}_2] &\propto \underbrace{\prod_{t=2}^T \text{Poisson}\left(y_{1,t} \mid \sum_{i=1}^m z_{it}\right)}_{\text{likelihood for census data}} \times \\ &\times \underbrace{\text{multinomial}\left(\mathbf{y}_{2,t} \mid \sum y_{2,it}, \frac{\mathbf{z}_t}{\sum_{i=1}^m z_{it}}\right)}_{\text{likelihood for classification data}} \\ &\times \underbrace{\text{multivariate normal}(\log(\mathbf{z}_t) | \log(g(\boldsymbol{\theta}, \mathbf{z}_{t-1})), \sigma_p^2)}_{\text{process model}} \\ &\times [\boldsymbol{\theta}] [\mathbf{z}_1] [\sigma_p^2]. \end{aligned}$$

# Combining priors from multiple studies I

As derived quantity:

$$\begin{aligned} [\theta_{meta}] &= \sum_{i=1}^n w_i [\theta] \\ \sum_{i=1}^n w_i &= 1 \end{aligned}$$

The  $w_i$  are chosen by the investigator based on quality of study, sample size, etc.

## Combining priors from multiple studies II

As hyper-parameter:

$$[\mathbf{z}, \theta_{\text{meta}}, \varsigma^2 \mid \boldsymbol{\theta}] \propto \prod_{j=1}^J [\theta_i \mid z_i, \sigma_{\theta_i}^2] [z_i \mid \theta_{\text{meta}}, \varsigma^2] [\theta_{\text{meta}}] [\varsigma^2] \quad (2)$$

## Example: multiple priors on an intercept $\beta_0$

Step 1: Do MCMC to approximate mean  $\tilde{\beta}_0$  and variance of the mean  $\sigma_{\tilde{\beta}_0}^2$ .

$$\beta_{0,1} \sim \text{normal}(3.2, .35)$$

$$\beta_{0,2} \sim \text{normal}(5.6, 1.2)$$

$$\beta_{0,3} \sim \text{normal}(1.2, .07)$$

$$w_i = \frac{1}{3}$$

$$\tilde{\beta}_0 = \sum_{i=1}^3 w_i \beta_{0,i}$$

You will have a numeric value for the mean of  $\tilde{\beta}$  and a standard deviation of the distribution of  $\tilde{\beta}_0$ .



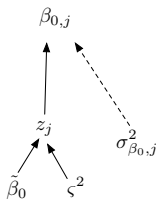
## Example: multiple priors on intercept $\beta_0$

Step 2: Using the numeric values for  $\tilde{\beta}_0$  and  $\sigma_{\tilde{\beta}_0}^2$ , find the marginal posterior distribution of the intercept based on new data and prior information ( $\beta_0^{\text{new}}$ ) in a separate MCMC algorithm:

$$\begin{aligned} [\beta_0^{\text{new}}, \beta_0, \beta_1, \sigma_{\text{reg}}^2 \mid \mathbf{y}] &\propto \prod_{i=1}^n \text{normal}(y_i \mid \beta_0^{\text{new}} + \beta_1 x_i, \sigma_{\text{reg}}^2) \\ &\quad \times \text{normal}(\beta_0^{\text{new}} \mid \underbrace{\tilde{\beta}_0, \tilde{\sigma}_{\beta_0}^2}_{\text{numeric}}) \\ &\quad \times \text{normal}(\beta_1 \mid 0, 10000) \\ &\quad \times \text{uniform}(\sigma^2 \mid 0, 2000) \end{aligned}$$

## Another approach: a hierarchical prior

Step1: Do MCMC to approximate mean  $\tilde{\beta}_0$  and variance of the mean  $\tilde{\sigma}_{\tilde{\beta}_0}^2$ .



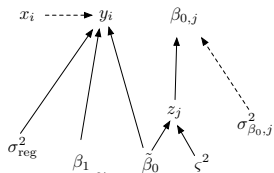
$$[\tilde{\beta}_0, \mathbf{z}, \zeta^2 | \boldsymbol{\beta}_0] \propto \prod_{j=1}^J [\beta_{0,j} | z_i, \sigma_{\beta_{0,j}}^2] [z_i | \tilde{\beta}_0, \zeta^2] [\mathbf{z}] [\tilde{\beta}_0] [\zeta^2]$$

## Another approach: a hierarchical prior

Step 2: As before, use the numeric values for  $\tilde{\beta}_0$  and  $\sigma_{\tilde{\beta}_0}^2$ , find the marginal posterior distribution of the intercept based on new data and prior information ( $\beta_0^{\text{new}}$ ) in a separate MCMC algorithm:

$$\begin{aligned} [\beta_0^{\text{new}}, \beta_1, \sigma_{\text{reg}}^2 \mid \mathbf{y}] &\propto \prod_{i=1}^n \text{normal}(y_i \mid \beta_0^{\text{new}} + \beta_1 x_i, \sigma_{\text{reg}}^2) \\ &\quad \times \underbrace{\text{normal}(\beta_0^{\text{new}} \mid \tilde{\beta}_0, \sigma_{\tilde{\beta}_0}^2)}_{\text{numeric}} \\ &\quad \times \text{normal}(\beta_1 \mid 0, 10000) \\ &\quad \times \text{uniform}(\zeta^2 \mid 0, 2000) \end{aligned}$$

## Including a hierarchical prior in a single model



$$[\tilde{\beta}_0, \beta_1, \sigma_{\text{reg}}, \zeta \mid \mathbf{y}] \propto \prod_{i=1}^n \text{normal}(y_i \mid \tilde{\beta}_0 + \beta_1 x_i, \sigma_{\text{reg}}^2) \quad (3)$$

$$\times \prod_{j=1}^J \text{normal}(\beta_{0,j} \mid z_j, \sigma_{\beta_{0,j}}^2) \quad (4)$$

$$\times \text{normal}(z_j \mid \tilde{\beta}_0, \zeta^2) \quad (5)$$

$$\times \text{normal}(\tilde{\beta}_0 \mid 0, 1000) \quad (6)$$

$$\times \text{normal}(\beta_1 \mid 0, 10000) \quad (7)$$

$$\times \text{uniform}(\zeta \mid 0, 200) \quad (8)$$

$$\times \text{uniform}(\sigma_{\text{reg}} \mid 0, 200) \quad (9)$$

These two approaches give results that are very similar, but not exactly the same. Why?

- ▶ Use the first approach when you want to weight studies with specific values for the  $w_i$ .
- ▶ Otherwise, the second approach which is more Bayesian and perhaps more easily explained.