

Prior Distributions I

Bayesian Modeling for Socio-Environmental Data

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Outline

- Bayes Theorem
- Priors
- Conjugacy

Bayes Theorem

$$\underbrace{[\theta|y]}_{\text{Posterior}} = \frac{\overbrace{[y|\theta]}^{\text{Likelihood}} \overbrace{[\theta]}^{\text{Prior}}}{\underbrace{\int [y|\theta][\theta] d\theta}_{\text{Marginal}}}$$

- Likelihood: Links unobserved θ to observed y .

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- Likelihood: Links unobserved θ to observed y .
- Prior Distribution: What is already known about θ .
- Marginal Distribution: the area under the joint distribution curve.
Serves to normalize the curve with respect to θ .
- Posterior Distribution: a true PDF.

Prior distributions

Prior distributions can be informative, reflecting knowledge gained in previous research *or* they can be vague, reflecting a lack of information about θ before data are collected.

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- Conjugacy is important for two reasons:
 - 1 Conjugacy in simple cases minimizes computational work and, in more complicated cases, allows us to break down calculations into manageable chunks.
 - 2 Conjugacy plays an important role in Markov chain Monte Carlo procedures (more on this later).

Conjugate priors

Table A.3: Table of conjugate distributions

Likelihood	Prior distribution	Posterior distribution
$y_i \sim \text{binomial}(n, \phi)$	$\phi \sim \text{beta}(\alpha, \beta)$	$\phi \sim \text{beta}(\sum y_i + \alpha, n - \sum y_i + \beta)$
$y_i \sim \text{Bernoulli}(\phi)$	$\phi \sim \text{beta}(\alpha, \beta)$	$\phi \sim \text{beta}(\sum_{i=1}^n y_i + \alpha, \sum_{i=1}^n (1 - y_i) + \beta)$
$y_i \sim \text{Poisson}(\lambda)$	$\lambda \sim \text{gamma}(\alpha, \beta)$	$\lambda \sim \text{gamma}(\alpha + \sum_{i=1}^n y_i, \beta + n)$
$y_i \sim \text{normal}(\mu, \sigma^2)$ σ^2 is known.	$\mu \sim \text{normal}(\mu_0, \sigma_0^2)$	$\mu \sim \text{normal}\left(\frac{\left(\frac{\mu_0}{\sigma_0^2} + \frac{\sum_{i=1}^n y_i}{\sigma^2}\right)}{\left(\frac{1}{\sigma_0^2} + \frac{n}{\sigma^2}\right)}, \left(\frac{1}{\sigma_0^2} + \frac{n}{\sigma^2}\right)^{-1}\right)$
$y_i \sim \text{normal}(\mu, \sigma^2)$ μ is known.	$\sigma^2 \sim$ inverse gamma(α, β)	$\sigma^2 \sim$ inverse gamma($\alpha + \frac{n}{2}, \beta + \frac{\sum_{i=1}^n (y_i - \mu)^2}{2}$)
$y_i \sim \text{lognormal}(\mu, \sigma^2)$, μ is known	$\sigma^2 \sim$ inverse gamma(α, β)	$\sigma^2 \sim$ inverse gamma($n/2 + \alpha, \frac{(\log(y_i) - \mu)^2}{2} + \beta$)
$y_i \sim \text{lognormal}(\mu, \sigma^2)$ σ^2 is known	$\mu \sim \text{normal}(\mu_0, \sigma_0^2)$	$\mu \sim \text{normal}\left(\frac{\left(\frac{\mu_0}{\sigma_0^2} + \frac{\sum_{i=1}^n \log y_i}{\sigma^2}\right)}{\left(\frac{1}{\sigma_0^2} + \frac{n}{\sigma^2}\right)}, \left(\frac{1}{\sigma_0^2} + \frac{n}{\sigma^2}\right)^{-1}\right)$

and Hooten, 2015)

(Hol

Just a bit more on priors

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- There is no such thing as a truly non-informative prior—only those that influence the posterior more (or less) than others.
- Informative priors can be justified and useful.
- “Non-informative,” vague, or flat priors are provisional starting points for a Bayesian analysis.

Example: childhood asthma and PM

- You are studying the relationship between childhood asthma and industrial airborne PM.
 - In one school 17 of 80 students have been hospitalized for asthma-related issues.
- 1 What distributions would you choose for the likelihood and the prior?
 - 2 How would you draw the DAG?

Picking distributions

Prior: $([\phi])$

- continuous quantity ranging from 0 to 1

Likelihood: $([\mathbf{y}|\phi])$

- count data: $y = 17$ successes, given $n = 80$ trials

Posterior: $([\phi|\mathbf{y}])$

- Is there a conjugate to use?

Picking distributions

Prior: $([\phi])$

- continuous quantity ranging from 0 to 1.
- Uniform beta with parameters, $\alpha_{prior} = 1, \beta_{prior} = 1$
- $\phi \sim \text{beta}(1, 1)$

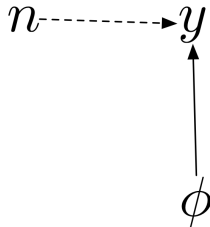
Likelihood: $([\mathbf{y}|\phi])$

- binomial distribution with $y = 17$ successes, given $n = 80$ trials

Posterior: $([\phi|\mathbf{y}])$

- Using the beta-binomial conjugate prior relationship
- $\phi \sim \text{beta}(\alpha_{post}, \beta_{post})$

Drawing the DAG



Writing out the full posterior

$$\text{beta}(\alpha_{post}, \beta_{post}) = \frac{\text{binomial}(y|\phi, n) \text{beta}(\alpha_{prior}, \beta_{prior})}{[y]}$$

Posterior distribution parameters

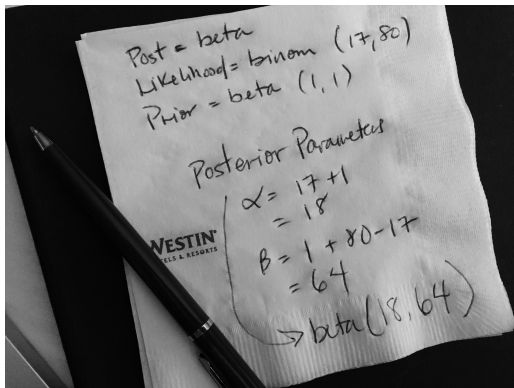
$$\alpha_{post} = \sum y_i + \alpha_{prior}$$

$$\beta_{post} = n - \sum y_i + \beta_{prior}$$

This means you can. . .

This means you can...

...literally calculate the parameters of the posterior distribution on the back of a hotel napkin.



Posterior distribution parameters

```
aPrior <- 1
bPrior <- 1
y <- 17
n <- 80
aPost <- aPrior + y
aPost
```

```
## [1] 18
```

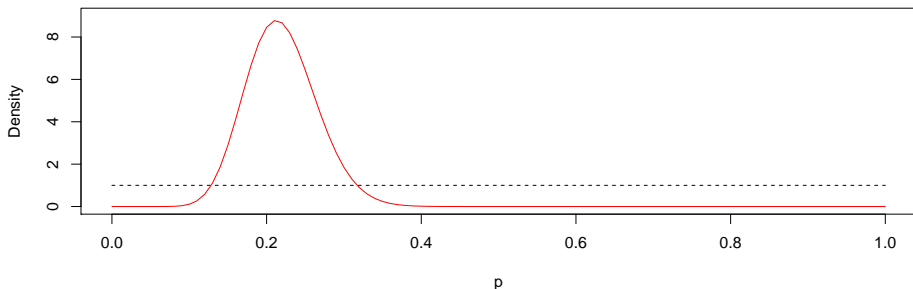
```
bPost <- bPrior + n - y
bPost
```

```
## [1] 64
```

$\phi \sim \text{beta}(18, 64)$

Plotting the prior and posterior

Beta Prior (dash) and Posterior (red)



Between which quantiles does ϕ lie with probability 0.95?

```
## [1] 0.1373419 0.3146275
```

Between which quantiles does ϕ lie with probability 0.95?

```
## [1] 0.1391926 0.2630073
```

What if, instead, the new study showed a much lower incidence of hospitalization, rather than closer finding (e.g. 3 of 75 hospitalized)?

```
aPrior <- 18
bPrior <- 64
y <- 3
n <- 75
aPost <- aPrior + y
aPost
```

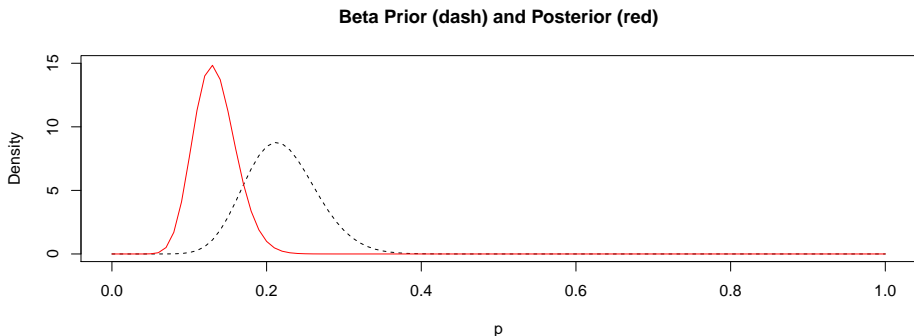
```
## [1] 21
```

```
bPost <- bPrior + n - y
bPost
```

```
## [1] 136
```

```
 $\phi \sim \text{beta}(21, 136)$ 
```

Plotting the prior and posterior



Between which quantiles does ϕ lie with probability 0.95?

```
## [1] 0.08529956 0.19103790
```

What if, instead, the new study showed a much higher incidence of hospitalization, rather than closer finding (e.g. 80 of 80 hospitalized)?

```
aPrior <- 18
bPrior <- 64
y <- 80
n <- 80
aPost <- aPrior + y
aPost
```

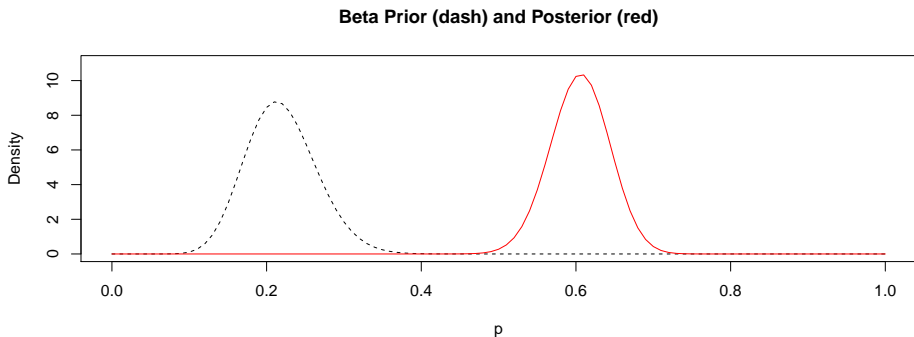
```
## [1] 98
```

```
bPost <- bPrior + n - y
bPost
```

```
## [1] 64
```

$\phi \sim \text{beta}(98, 64)$

Plotting the prior and posterior

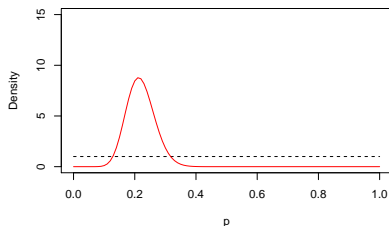


Between which quantiles does ϕ lie with probability 0.95?

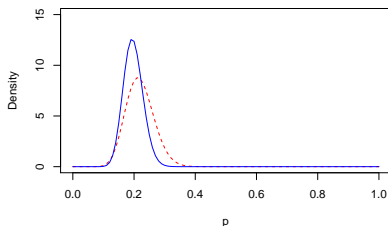
```
## [1] 0.5287686 0.6786532
```

Exploring the role of new data

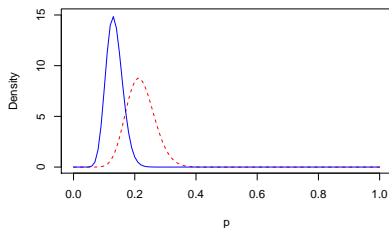
ORIGINAL STUDY:
Flat Beta Prior (black) and Posterior (red)



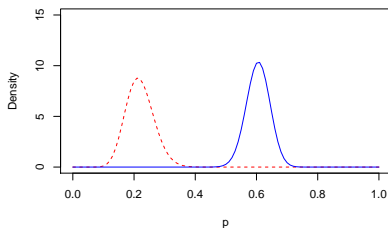
NEW DATA SIMILAR INCIDENCE:
Informed Beta Prior (red) and Posterior (blue)



NEW DATA SMALLER INCIDENCE:
Informed Beta Prior (dash) and Posterior (red)



NEW DATA LARGER INCIDENCE:
Informed Beta Prior (dash) and Posterior (red)

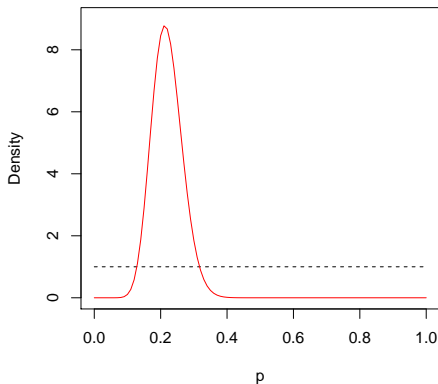


→ ##

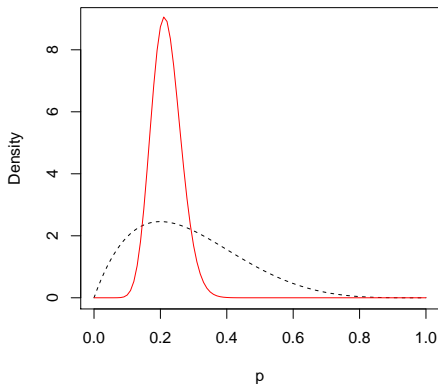
All priors and data influence on the posterior distribution

Comparing priors

Prior $[\phi]$: $\phi \sim \text{beta}(1, 1)$
Post $[\phi|y]$: $\phi \sim \text{beta}(18, 64)$
Beta Prior (dash) and Posterior (red)



Prior $[\phi]$: $\phi \sim \text{beta}(2, 5)$
Post $[\phi|y]$: $\phi \sim \text{beta}(19, 68)$
Beta Prior (dash) and Posterior (red)



Now consider the following data:

School: Hospitalizations/Total Students

School 1: 17/80

School 2: 17/75

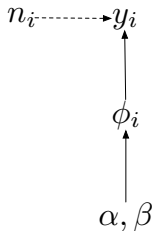
School 3: 19/100

School 4: 10/55

School 5: 33/111

For a childhood asthma hospitalization model that, uses the new data across schools, what is the DAG?

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Full posterior

$$y_i \sim \text{binomial}(n_i, \phi_i)$$

$$\phi_i \sim \text{beta}(\alpha, \beta)$$

$$\alpha \sim \text{uniform}(0, 500)$$

$$\beta \sim \text{uniform}(0, 500)$$

$$[\phi, \alpha, \beta | \mathbf{y}] \propto \prod_{i=1}^5 [y_i | \phi_i, n_i] [\phi_i | \alpha, \beta] [\alpha] [\beta]$$

This is a hierarchical model because parameter ϕ , is on both sides of the conditioning symbol.

→

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- Informative priors, when properly justified, can be useful.
- Strong data overwhelms a prior.

Role of priors in science

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- Priors represent our current knowledge (or lack of current knowledge), which is updated with data.
- We encourage you to think of vague priors as a provisional starting point.