

# Variable Selection In Additive Gene Environment Interactions with the Group Lasso

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## 1 Introduction

We consider a regression model for an outcome variable  $\mathbf{Y} = (Y_1, \dots, Y_n)$  where  $n$  is the number of subjects. Let  $E = (E_1, \dots, E_n)$  be a binary or continuous environment vector and  $\mathbf{X} = (X_1, \dots, X_n)^T$  be the  $n \times p$  matrix of high-dimensional data where  $X_i = (X_{i1}, \dots, X_{ij}, \dots, X_{ip}) \in [0, 1]^p$ , and  $\boldsymbol{\varepsilon} = (\varepsilon_1, \dots, \varepsilon_n)$  a vector of errors. Consider the regression model with main effects and their interactions with  $E$ :

$$Y_i = \beta_0^* + \sum_{j=1}^p \beta_j^* X_{ij} + \beta_E^* E_i + \sum_{j=1}^p \alpha_j^* E_i X_j + \varepsilon_i, \quad i = 1, \dots, n, \quad (1)$$

where  $\beta_0^*, \beta_j^*, \beta_E^*, \alpha_j^*$  are the true unknown model parameters for  $j = 1, \dots, p$ . This can be extended to the more general additive model:

$$Y_i = \beta_0^* + \sum_{j=1}^p f_j^*(X_{ij}) + f_E^*(E_i) + \sum_{j=1}^p f_{jE}^*(X_{ij}, E_i) + \varepsilon_i \quad i = 1, \dots, n \quad (2)$$

As in (Radchenko and James, 2010), we can express (2) as

$$\mathbf{Y} = \sum_{j=1}^p \mathbf{f}_j^* + \mathbf{f}_E^* + \sum_{j=1}^p \mathbf{f}_{jE}^* + \boldsymbol{\varepsilon} \quad (3)$$

where  $\mathbf{f}_j^* = (f_j^*(X_{1j}), \dots, f_j^*(X_{nj}))^T$ ,  $\mathbf{f}_{jE}^* = (f_{jE}^*(X_{1j}, X_{1E}), \dots, f_{jE}^*(X_{nj}, X_{nE}))^T$  and  $\mathbf{f}_E^* = f_E^*(E_i)$ . We consider the candidate vectors  $\{\mathbf{f}_j, \mathbf{f}_E, \mathbf{f}_{jE}\}$ . The general approach for fitting (3) is to minimize the following penalized regression criterion:

$$\frac{1}{2} \|\mathbf{Y} - \mathbf{f}\|^2 + P(\mathbf{f}) \quad (4)$$

where

$$\mathbf{f} = \sum_{j=1}^p \mathbf{f}_j + \mathbf{f}_E + \sum_{j=1}^p \mathbf{f}_{jE} \quad (5)$$

and  $P(\mathbf{f})$  is a penalty function on  $\mathbf{f}$

The smoothing method for variable  $X_j$  is a projection on to a set of basis functions. Consider

$$f_j(\cdot) = \sum_{\ell=1}^{p_j} \psi_{j\ell}(\cdot) \beta_{j\ell} \quad (6)$$

where the  $\{\psi_{j\ell}\}_1^{p_j}$  are a family of basis functions in  $X_j$  ([Hastie et al., 2015](#)). Let

$$f_{jE}(X_j, E) = \sum_{\ell=1}^{q_j} \phi_{j\ell}(X_j, E) \alpha_{j\ell} \quad (7)$$

where the  $\{\phi_{j\ell}\}_1^{q_j}$  are a family of basis functions in  $X_j \cdot E$ .

Following ([Choi et al., 2010](#)), we reparametrize the coefficients for the interaction terms as  $\alpha_{j\ell} = \gamma_{j\ell} \beta_{j\ell} \beta_E$ . Plugging this into (7):

$$f_{jE}(X_j, E) = \sum_{\ell=1}^{q_j} \phi_{j\ell}(X_j, E) \gamma_{j\ell} \beta_{j\ell} \beta_E \quad (8)$$

## Bibliography

- Peter Radchenko and Gareth M James. Variable selection using adaptive nonlinear interaction structures in high dimensions. *Journal of the American Statistical Association*, 105(492):1541–1553, 2010. [1](#)
- Trevor Hastie, Robert Tibshirani, and Martin Wainwright. *Statistical Learning with Sparsity: The Lasso and Generalizations*. CRC Press, 2015. [2](#)
- Nam Hee Choi, William Li, and Ji Zhu. Variable selection with the strong heredity constraint and its oracle property. *Journal of the American Statistical Association*, 105(489):354–364, 2010. [2](#)