

Brownian Motion

Exercises

Exercise 9.1. Let Z be a standard normal random variable. For all $t \geq 0$, let $X_t = \sqrt{t}Z$. The stochastic process $X = \{X_t : t \geq 0\}$ has continuous paths and $\forall t \geq 0, X_t \sim N(0, t)$. Is X a Brownian motion? Justify. (ref. Baxter and Rennie, p. 49)

Exercise 9.2. Let W and \widetilde{W} be two independent Brownian motion and ρ is a constant contained in the unit interval. For all $t \geq 0$, let $X_t = \rho W_t + \sqrt{1 - \rho^2} \widetilde{W}_t$. The stochastic process $X = \{X_t : t \geq 0\}$ has continuous paths and $\forall t \geq 0, X_t \sim N(0, t)$. Is X a Brownian motion? Justify. (ref Baxter and Rennie, p. 49)

Exercise 9.3. Let W be a Brownian motion on the filtered probability space $(\Omega, \mathcal{F}, \mathbb{F} = \{\mathcal{F}_t : t \geq 0\}, P)$. Let $X_t = \exp\left[\sigma W_t - \frac{\sigma^2}{2}t\right]$. Show that $X = \{X_t : t \geq 0\}$ is a martingale.

Exercise 9.4. Let W be a Brownian motion on the filtered probability space $(\Omega, \mathcal{F}, \mathbb{F} = \{\mathcal{F}_t : t \geq 0\}, P)$. Show that $\{W_t^2 - t : t \geq 0\}$ is a martingale.

Exercise 9.5. Let W be a Brownian motion. Show that

$$\text{Cov}[W_t, W_s] = \min(s, t).$$

Exercise 9.6. Let W be a Brownian motion. Show that

- (i) For all $s > 0, \{W_{t+s} - W_s : t \geq 0\}$
- (ii) $\{-W_t : t \geq 0\}$
- (iii) $\left\{cW_{\frac{t}{c^2}} : t \geq 0\right\}$
- (iv) $\left\{V_0 = 0 \text{ and } V_t = tW_{\frac{1}{t}} \text{ if } t > 0 : t \geq 0\right\}$

are Brownian motions.

Exercise 9.7. Let \mathbf{B} be a four-dimensional Brownian motion with

$$\text{Corr}[\mathbf{B}_t] = \begin{pmatrix} 1 & 0.5 & 0.8 & 0.1 \\ 0.5 & 1 & 0.3 & 0.4 \\ 0.8 & 0.3 & 1 & 0.1 \\ 0.1 & 0.4 & 0.1 & 1 \end{pmatrix}.$$

Find the matrix \mathbf{A} such that $\mathbf{B} = \mathbf{A}\mathbf{W}$ and \mathbf{W} is a four-dimensional Brownian motion with independent components.

Solutions

1 Exercise 9.1

No since $0 \leq s \leq t < \infty$,

$$\begin{aligned}\text{Var}[X_t - X_s] &= \text{Var}[\sqrt{t}Z - \sqrt{s}Z] \\ &= (\sqrt{t} - \sqrt{s})^2 \text{Var}[Z] \\ &= t - 2\sqrt{t}\sqrt{s} + s \\ &\neq t - s.\end{aligned}$$

2 Exercise 9.2

Yes. Suffices to verify that (i) the time increments are independent and (ii) for all $0 \leq s \leq t < \infty$, $X_t - X_s \sim N(0, t - s)$.

(ii)

$$X_t - X_s = \underbrace{\rho(W_t - W_s)}_{N(0, \rho^2(t-s))} + \underbrace{\sqrt{1 - \rho^2}(\widetilde{W}_t - \widetilde{W}_s)}_{N(0, (1-\rho^2)(t-s))}$$

Since the two terms in the right hand side are two independent Gaussian random variables with an expectation of zero, their sum is also a zero-expectation Gaussian random variable. Finally

$$\text{Var}[X_t - X_s] = \rho^2(t - s) + (1 - \rho^2)(t - s) = t - s$$

which completes the first part.

(i) Let $0 \leq t_1 \leq t_2 \leq t_3 \leq t_4 < \infty$. Since $X_{t_2} - X_{t_1}$ and $X_{t_4} - X_{t_3}$ have a Gaussian distribution, it suffices to show that the covariance is null:

$$\begin{aligned}& \text{Cov}[X_{t_2} - X_{t_1}; X_{t_4} - X_{t_3}] \\ &= \text{Cov}\left[\rho(W_{t_2} - W_{t_1}) + \sqrt{1 - \rho^2}(\widetilde{W}_{t_2} - \widetilde{W}_{t_1}); \rho(W_{t_4} - W_{t_3}) + \sqrt{1 - \rho^2}(\widetilde{W}_{t_4} - \widetilde{W}_{t_3})\right] \\ &= \rho^2 \text{Cov}[W_{t_2} - W_{t_1}; W_{t_4} - W_{t_3}] + \rho\sqrt{1 - \rho^2} \text{Cov}[W_{t_2} - W_{t_1}; \widetilde{W}_{t_4} - \widetilde{W}_{t_3}] \\ &\quad + \rho\sqrt{1 - \rho^2} \text{Cov}[\widetilde{W}_{t_2} - \widetilde{W}_{t_1}; W_{t_4} - W_{t_3}] + (1 - \rho^2) \text{Cov}[\widetilde{W}_{t_2} - \widetilde{W}_{t_1}; \widetilde{W}_{t_4} - \widetilde{W}_{t_3}] \\ &= 0\end{aligned}$$

since the Brownian increment independence implies that

$$\text{Cov} [W_{t_2} - W_{t_1}; W_{t_4} - W_{t_3}] = 0$$

and

$$\text{Cov} [\widetilde{W}_{t_2} - \widetilde{W}_{t_1}; \widetilde{W}_{t_4} - \widetilde{W}_{t_3}] = 0.$$

The independence between the two Brownian motions implies that

$$\text{Cov} [W_{t_2} - W_{t_1}; \widetilde{W}_{t_4} - \widetilde{W}_{t_3}] = 0$$

and

$$\text{Cov} [\widetilde{W}_{t_2} - \widetilde{W}_{t_1}; W_{t_4} - W_{t_3}] = 0.$$

3 Exercise 9.3

(i) Integrability

$$\begin{aligned} \mathbb{E} [|X_t|] &= \mathbb{E} \left[\left| \exp \left[\sigma W_t - \frac{\sigma^2}{2} t \right] \right| \right] = \mathbb{E} \left[\exp \left[\sigma W_t - \frac{\sigma^2}{2} t \right] \right] \\ &= \int_{-\infty}^{\infty} \exp \left[\sigma w - \frac{\sigma^2}{2} t \right] \frac{1}{\sqrt{2\pi}} \frac{1}{\sqrt{t}} \exp \left[-\frac{w^2}{2t} \right] dw \\ &= \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} \frac{1}{\sqrt{t}} \exp \left[-\frac{w^2 - 2t\sigma w + \sigma^2 t^2}{2t} \right] dw \\ &= \underbrace{\int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} \frac{1}{\sqrt{t}} \exp \left[-\frac{(w - t\sigma)^2}{2t} \right] dw}_{N(t\sigma, t) \text{ density function}} = 1 < \infty \end{aligned}$$

(ii) Since X_t is a continuous function of \mathcal{F}_t -measurable random variables, X_t is itself \mathcal{F}_t -measurable.

(iii) For all $0 \leq s \leq t \leq \infty$,

$$\begin{aligned}
\mathbb{E}[X_t | \mathcal{F}_s] &= X_s \mathbb{E} \left[\frac{X_t}{X_s} \middle| \mathcal{F}_s \right] \text{ car } X_s > 0 \\
&= X_s \mathbb{E} \left[\frac{\exp \left[\sigma W_t - \frac{\sigma^2}{2} t \right]}{\exp \left[\sigma W_s - \frac{\sigma^2}{2} s \right]} \middle| \mathcal{F}_s \right] \\
&= X_s \mathbb{E} \left[\exp \left[\sigma (W_t - W_s) - \frac{\sigma^2}{2} (t - s) \right] \middle| \mathcal{F}_s \right] \\
&= X_s \mathbb{E} \left[\exp \left[\sigma (W_t - W_s) - \frac{\sigma^2}{2} (t - s) \right] \right] \text{ since } W_t - W_s \text{ is independent of } \mathcal{F}_s. \\
&= X_s \int_{-\infty}^{\infty} \exp \left[\sigma w - \frac{\sigma^2}{2} (t - s) \right] \frac{1}{\sqrt{2\pi}} \frac{1}{\sqrt{t - s}} \exp \left[-\frac{w^2}{2(t - s)} \right] dw \\
&= X_s \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} \frac{1}{\sqrt{t - s}} \exp \left[-\frac{w^2 - 2(t - s)\sigma w + \sigma^2(t - s)^2}{2(t - s)} \right] dw \\
&= X_s \underbrace{\int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} \frac{1}{\sqrt{t - s}} \exp \left[-\frac{(w - (t - s)\sigma)^2}{2(t - s)} \right] dw}_{N((t-s)\sigma, t-s) \text{ density function}} = X_s
\end{aligned}$$

4 Exercise 9.4

First, $W_t^2 - t$ is \mathcal{F}_t -measurable since it is a continuous function of W_t which is \mathcal{F}_t -measurable.

Second,

$$\mathbb{E}[|W_t^2 - t|] \leq \mathbb{E}[W_t^2] + t = t + t = 2t < \infty.$$

Third, $\forall 0 \leq s \leq t$,

$$\begin{aligned}
\mathbb{E}[W_t^2 - t | \mathcal{F}_s] &= \mathbb{E}[(W_t - W_s + W_s)^2 - t | \mathcal{F}_s] \\
&= \mathbb{E}[(W_t - W_s)^2 + 2W_s(W_t - W_s) + W_s^2 - t | \mathcal{F}_s] \\
&= \underbrace{\mathbb{E}[(W_t - W_s)^2 | \mathcal{F}_s]}_{=t-s} + 2W_s \underbrace{\mathbb{E}[W_t - W_s | \mathcal{F}_s]}_{=0} + W_s^2 - t \\
&= W_s^2 - s. \blacksquare
\end{aligned}$$

5 Exercise 9.5

Without loss of generality, let $0 < s < t$.

$$\begin{aligned}
 \text{Cov}[W_t, W_s] &= \text{Cov}[W_t - W_s + W_s, W_s] \\
 &= \text{Cov}[W_t - W_s, W_s] + \text{Cov}[W_s, W_s] \\
 &= \text{Cov}[W_t - W_s, W_s - W_0] + \text{Var}[W_s] \\
 &= 0 + s \text{ because the increments of } W \text{ are independent} \\
 &= \min(s, t) \text{ since } s < t. \blacksquare
 \end{aligned}$$

6 Exercise 9.6

Let

$$Z_t = W_{t+s} - W_s.$$

$$(MB1) \quad Z_0 = W_s - W_s = 0.$$

(MB2) Since $Z_{t_k} - Z_{t_{k-1}} = (W_{t_k+s} - W_s) - (W_{t_{k-1}+s} - W_s) = W_{t_k+s} - W_{t_{k-1}+s}$ and for $\forall 0 \leq t_0 < t_1 < \dots < t_k$, the random variables $W_{t_1+s} - W_{t_0+s}, W_{t_2+s} - W_{t_1+s}, \dots, W_{t_k+s} - W_{t_{k-1}+s}$ are independent, then for $\forall 0 \leq t_0 < t_1 < \dots < t_k$, the random variables $Z_{t_1} - Z_{t_0}, Z_{t_2} - Z_{t_1}, \dots, Z_{t_k} - Z_{t_{k-1}}$ are independent.

(MB3) $\forall u, t \geq 0$ such that $u < t$, $Z_t - Z_u = (W_{t+s} - W_s) - (W_{u+s} - W_s) = W_{t+s} - W_{u+s}$ has a zero-expectation Gaussian distribution 0 with variance $(t+s) - (u+s) = t - u$.

(MB4) $\forall \omega \in \Omega$, the path $t \rightarrow Z_t(\omega) = W_{t+s}(\omega) - W_s(\omega)$ is continuous since $t \rightarrow W_t(\omega)$ is continuous.

Let

$$Y_t = -W_t.$$

$$(MB1) \quad Y_0 = -W_0 = 0.$$

(MB2) Since $Y_{t_k} - Y_{t_{k-1}} = W_{t_{k-1}} - W_{t_k}$ and $\forall 0 \leq t_0 < t_1 < \dots < t_k$, the random variables $W_{t_1} - W_{t_0}, W_{t_2} - W_{t_1}, \dots, W_{t_k} - W_{t_{k-1}}$ are independent, then $\forall 0 \leq t_0 < t_1 < \dots < t_k$, the random variables $W_{t_0} - W_{t_1}, W_{t_1} - W_{t_2}, \dots, W_{t_{k-1}} - W_{t_k}$ are independent, which implies that $Y_{t_1} - Y_{t_0}, Y_{t_2} - Y_{t_1}, \dots, Y_{t_k} - Y_{t_{k-1}}$ are independent.

(MB3) $\forall s, t \geq 0$ such that $s < t$, $Y_t - Y_s = W_s - W_t$ has a Gaussian distribution with expectation 0 and variance $t - s$.

(MB4) $\forall \omega \in \Omega$, the path $t \rightarrow Y_t(\omega) = -W_t(\omega)$ is continuous since $t \rightarrow W_t(\omega)$ is continuous.

Let

$$X_t = cW_{\frac{t}{c^2}}.$$

$$(MB1) \quad X_0 = cW_0 = 0.$$

(MB2) Since $X_{t_k} - X_{t_{k-1}} = cW_{\frac{t_k}{c^2}} - cW_{\frac{t_{k-1}}{c^2}}$ and $\forall 0 \leq t_0 < t_1 < \dots < t_k$, the random variables $W_{\frac{t_1}{c^2}} - W_{\frac{t_0}{c^2}}, W_{\frac{t_2}{c^2}} - W_{\frac{t_1}{c^2}}, \dots, W_{\frac{t_k}{c^2}} - W_{\frac{t_{k-1}}{c^2}}$ are independent, then $\forall 0 \leq t_0 < t_1 < \dots < t_k$, the random variables $cW_{\frac{t_1}{c^2}} - cW_{\frac{t_0}{c^2}}, cW_{\frac{t_2}{c^2}} - cW_{\frac{t_1}{c^2}}, \dots, cW_{\frac{t_k}{c^2}} - cW_{\frac{t_{k-1}}{c^2}}$ are independent, which implies that $X_{t_1} - X_{t_0}, X_{t_2} - X_{t_1}, \dots, X_{t_k} - X_{t_{k-1}}$ are independent.

(MB3) Since cW is Gaussian if W is, $\forall s, t \geq 0$ such that $s < t$, $X_t - X_s = c \left(W_{\frac{t}{c^2}} - W_{\frac{s}{c^2}} \right)$ is Gaussian with $E[X_t - X_s] = cE \left[W_{\frac{t}{c^2}} - W_{\frac{s}{c^2}} \right] = 0$ and variance $\text{Var}[X_t - X_s] = c^2 \text{Var} \left[W_{\frac{t}{c^2}} - W_{\frac{s}{c^2}} \right] = c^2 \left(\frac{t}{c^2} - \frac{s}{c^2} \right) = t - s$.

(MB4) $\forall \omega \in \Omega$, the path $t \rightarrow X_t(\omega) = cW_{\frac{t}{c^2}}(\omega)$ is continuous since $t \rightarrow W_t(\omega)$ is.

Let

$$V_t = \begin{cases} 0 & \text{if } t = 0 \\ tW_{\frac{1}{t}} & \text{if } t > 0. \end{cases}$$

(MB1) $V_0 = 0$ from definition of V .

(MB3) $\forall s, t \geq 0$ such that $s < t$,

$$\begin{aligned} V_t - V_s &= tW_{\frac{1}{t}} - sW_{\frac{1}{s}} \\ &= -s \left(W_{\frac{1}{s}} - W_{\frac{1}{t}} \right) + (t - s) W_{\frac{1}{t}} \end{aligned}$$

is a linear combination of two independent Gaussian random variable. Therefore, $V_t - V_s$ is Gaussian with

$$E[V_t - V_s] = E[tW_{\frac{1}{t}} - sW_{\frac{1}{s}}] = tE[W_{\frac{1}{t}}] - sE[W_{\frac{1}{s}}] = 0$$

and

$$\begin{aligned} \text{Var}[V_t - V_s] &= \text{Var} \left[-s \left(W_{\frac{1}{s}} - W_{\frac{1}{t}} \right) + (t - s) W_{\frac{1}{t}} \right] \\ &= \text{Var} \left[-s \left(W_{\frac{1}{s}} - W_{\frac{1}{t}} \right) \right] + \text{Var} \left[(t - s) W_{\frac{1}{t}} \right] \\ &\quad \text{since } W_{\frac{1}{s}} - W_{\frac{1}{t}} \text{ is independent of } W_{\frac{1}{t}} \\ &= s^2 \text{Var} \left[W_{\frac{1}{s}} - W_{\frac{1}{t}} \right] + (t - s)^2 \text{Var} \left[W_{\frac{1}{t}} \right] \\ &= s^2 \left(\frac{1}{s} - \frac{1}{t} \right) + (t - s)^2 \frac{1}{t} \\ &= s - \frac{s^2}{t} + t - 2s + \frac{s^2}{t} \\ &= t - s. \end{aligned}$$

If $s = 0$, then $V_t = tW_{\frac{1}{t}}$ is Gaussian with

$$\mathbb{E}[V_t] = \mathbb{E}\left[tW_{\frac{1}{t}}\right] = t\mathbb{E}\left[W_{\frac{1}{t}}\right] = 0$$

and

$$\text{Var}[V_t] = \text{Var}\left[tW_{\frac{1}{t}}\right] = t^2\text{Var}\left[W_{\frac{1}{t}}\right] = t^2\frac{1}{t} = t.$$

(MB2) It suffices to show that $\forall 0 \leq t_1 < t_2 \leq t_3 < t_4$, the covariance between $V_{t_2} - V_{t_1}$ and $V_{t_4} - V_{t_3}$ is zero since these two random variables have a Gaussian distribution.

If $t_1 > 0$, then because $0 < \frac{1}{t_4} < \frac{1}{t_3} \leq \frac{1}{t_2} < \frac{1}{t_1}$,

$$\begin{aligned} \text{Cov}[V_{t_2} - V_{t_1}; V_{t_4} - V_{t_3}] &= \text{Cov}\left[t_2W_{\frac{1}{t_2}} - t_1W_{\frac{1}{t_1}}; t_4W_{\frac{1}{t_4}} - t_3W_{\frac{1}{t_3}}\right] \\ &= t_2t_4\text{Cov}\left[W_{\frac{1}{t_2}}; W_{\frac{1}{t_4}}\right] - t_2t_3\text{Cov}\left[W_{\frac{1}{t_2}}; W_{\frac{1}{t_3}}\right] \\ &\quad - t_1t_4\text{Cov}\left[W_{\frac{1}{t_1}}; W_{\frac{1}{t_4}}\right] + t_1t_3\text{Cov}\left[W_{\frac{1}{t_1}}; W_{\frac{1}{t_3}}\right] \\ &\quad \text{since } \text{Cov}(W_t, W_s) = \min(s, t). \\ &= t_2t_4\frac{1}{t_4} - t_2t_3\frac{1}{t_3} - t_1t_4\frac{1}{t_4} + t_1t_3\frac{1}{t_3} \\ &= t_2 - t_2 - t_1 + t_1 \\ &= 0 \end{aligned}$$

If $t_1 = 0$, then

$$\begin{aligned} \text{Cov}[V_{t_2} - V_{t_1}; V_{t_4} - V_{t_3}] &= \text{Cov}\left[t_2W_{\frac{1}{t_2}}; t_4W_{\frac{1}{t_4}} - t_3W_{\frac{1}{t_3}}\right] \\ &= t_2t_4\text{Cov}\left[W_{\frac{1}{t_2}}; W_{\frac{1}{t_4}}\right] - t_2t_3\text{Cov}\left[W_{\frac{1}{t_2}}; W_{\frac{1}{t_3}}\right] \\ &= t_2t_4\frac{1}{t_4} - t_2t_3\frac{1}{t_3} \\ &= t_2 - t_2 \\ &= 0 \end{aligned}$$

(MB4) $\forall \omega \in \Omega$, the path $t \rightarrow V_t(\omega) = tW_{\frac{1}{t}}(\omega)$ is continuous for all $t > 0$ since the functions $t \rightarrow W_t(\omega)$ and $t \rightarrow t$ are continuous and so is their product. Since $\lim_{t \rightarrow 0} tW_{\frac{1}{t}}(\omega) = 0$ almost-surely, the path $t \rightarrow V_t(\omega) = tW_{\frac{1}{t}}(\omega)$ is continuous for all t .