Lecture 7 - Thompson sampling

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Recap: Bayesian statistics and Bernoulli-Beta conjugate

Recall that the reward r(i) of arm i follows some distribution. Assume that the reward distributions of arms belong to the same family with respective parameters, which writes

$$r(i) \sim p(x \mid \theta_i)$$
.

Recall that when estimating θ , the posterior is

$$p(\theta \mid x) = \frac{p(x \mid \theta)p(\theta)}{\int_{\theta'} p(x \mid \theta')p(\theta')d\theta'}.$$

Conjugate distributions: The posterior distributions $p(\theta \mid x)$ are in the same probability distribution family as the prior probability distribution $p(\theta)$.

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Recap: Bayesian statistics and Bernoulli-Beta conjugate

The Bernoulli-Beta is important for Thompson sampling for Bernoulli bandits. Recall that the Beta distribution Beta (α, β) with parameter $\theta = \{\alpha, \beta\}$ follows the probability density function of

$$p(x) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha - 1} (1 - x)^{\beta - 1},$$

where $\Gamma(z) = \int_0^\infty x^{z-1} \exp(-x) dx$, $z \in \mathbb{C}$ is the Gamma function.

Recap: Bayesian statistics and Bernoulli-Beta conjugate

When $p(\theta) \sim \text{Beta}(\alpha_0, \beta_0)$ and we observe $x_1, \dots, x_{\alpha' + \beta'} \sim x$ i.i.d. with α' ones and β' zeros (observe a new data $x \sim \text{Ber}(\theta)$), then the posterior should follow:

$$\begin{split} p(\theta \mid x_{1}, \dots, x_{\alpha' + \beta'}) &= \frac{p(x_{1}, \dots, x_{\alpha' + \beta'} \mid \theta) p(\theta)}{\int_{\theta'} p(x_{1}, \dots, x_{\alpha' + \beta'} \mid \theta') p(\theta') d\theta'} \\ &= \frac{\binom{\alpha' + \beta'}{\alpha'} \theta^{\alpha'} (1 - \theta)^{\beta'} \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha) \Gamma(\beta)} \theta^{\alpha_{0} - 1} (1 - \theta)^{\beta_{0} - 1}}{\int_{\theta'} p(x_{1}, \dots, x_{\alpha' + \beta'} \mid \theta') p(\theta') d\theta'} \\ &= \frac{\binom{\alpha' + \beta'}{\alpha'} \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha) \Gamma(\beta)}}{\int_{\theta'} p(x_{1}, \dots, x_{\alpha' + \beta'} \mid \theta') p(\theta') d\theta'} \theta^{\alpha_{0} + \alpha' - 1} (1 - \theta)^{\beta_{0} + \alpha' - 1}} \\ &\sim \text{Beta}(\alpha_{0} + \alpha', \beta_{0} + \beta'). \end{split}$$
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Thompson sampling algorithms

- Before the game starts, the learner sets a prior over possible bandit environments.
- In each round, the learner samples an environment from the posterior and acts according to the optimal action in that environment.
- The exploration in Thompson sampling comes from the randomization, i.e., whether the posterior concentrates or not.

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Algorithm 1: Thompson sampling (Bernoulli bandits)

Input: Prior \alpha_0, \beta_0

Output: a_t, t \in [T]

Initialize \alpha_i = \alpha_0, \beta_i = \beta_0, for i \in [m]

while t \leq T - 1 do

Sample \theta_i(t) \sim \text{Beta}(\alpha_i, \beta_i) independently for i \in [m]

a_t = \arg\max_{i \in [m]} \theta_i(t) with arbitrary tiebreaker

If r_t = 1 then \alpha_{a_t} + = 1; If r_t = 0 then \beta_{a_t} + = 1;

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Thompson sampling algorithms

When the family of the underlying reward distribution is unknown, a Gaussian-Gaussian conjugate (the non-informative prior) can be useful.

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Algorithm 2: Thompson sampling
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Input: Prior θ_0

Output: $a_t, t \in [T]$

Initialize $\theta_i = \theta_0$, for $i \in [m]$

while $t \leq T - 1$ do

Sample independently for $i \in [m]$, $\theta_i(t) \sim p(\theta \mid \{r_{t'}\}_{1 \{a_{t'}=i,t' \leq t-1\}})$ $a_t = \arg\max_{i \in [m]} \theta_i(t)$ with arbitrary tiebreaker

Update the posterior probability distribution of $\theta_{a_t}(t+1)$ by

$$p(\theta_{a_t}(t+1) \mid \{r_{t'}\}_{\mathbb{1}\{a_{t'}=i\}}) = \frac{p(\{r_{t'}\}_{\mathbb{1}\{a_{t'}=i\}} \mid \theta)p(\theta)}{\int_{\theta'} p(\{r_{t'}\}_{\mathbb{1}\{a_{t'}=i\}} \mid \theta')p(\theta')d\theta'}$$

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The Regret of Thompson sampling Algorithms

Theorem

Assume the rewards of arms are μ_i -Bernoulli. The regret under TS (Bernoulli bandits) is at most:

$$\overline{R}_T \leq \sum_{i:\Delta_i > 0} rac{\mu_1 - \mu_i}{d_{\mathsf{KL}}(\mu_1 \| \mu_i)} \log T + o(\log T),$$

where the Kullback-Leibler divergence:

$$d_{\mathit{KL}}(\mu_1 \| \mu_i) = \mu_1 \log(\frac{\mu_1}{\mu_i}) + (1 - \mu_1) \log(\frac{1 - \mu_1}{1 - \mu_i}).$$



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Question and Answering (Q&A)



