

# 1 Quantities

- $J \in \mathbb{Z}^+ =$  number of scales decomposition
- $T = 2^J =$  number of time periods
- $N \in \mathbb{Z}^+ =$  number of cross-section elements
- $K(\leq N) =$  number of common factors

## 2 Multivariate Locally Stationary Wavelet process (Park et al. (2014))

The vector  $(N \times 1)$  of stochastic processes  $\mathbf{X}_{t;T}$  follows the given decomposition :

$$\mathbf{X}_{t;T} = \sum_{j=-J}^{-1} \sum_{k=0}^T \mathbf{W}_j(k/T) \boldsymbol{\xi}_{j,k} \psi_{j,k}(t) \quad (2.1)$$

where

- $\mathbf{W}_j(\mathbf{z})$  is a lower-triangular  $(N \times N)$  matrix.  
For each  $(m, n)$ -element,

$$W_j^{(m,n)}(z) \text{ is a Lipschitz continuous function on } z \in (0, 1) \quad (2.2)$$

$$\sum_{j=-\infty}^{-1} \left| W_j^{(m,n)}(z) \right|^2 < \infty, \quad \forall z \in (0, 1) \quad (\text{finite energy}) \quad (2.3)$$

$$\sum_{j=-\infty}^{-1} 2^{-j} L_j^{(m,n)} < \infty \quad (\text{uniformly bounded Lipschitz constants } L_j) \quad (2.4)$$

- $\boldsymbol{\xi}_{j,k}$  is the vector  $(N \times 1)$  of random orthonormal increments.

$$\mathbb{E} \left[ \boldsymbol{\xi}_{j,k}^{(u)} \right] = 0, \quad \forall j, k, u \quad (2.5)$$

$$\text{Cov} \left[ \boldsymbol{\xi}_{j,k}^{(u)}, \boldsymbol{\xi}_{j',k'}^{(u')} \right] = \delta_{j,j'} \delta_{k,k'} \delta_{u,u'}, \quad \forall j, j', k, k', u, u' \quad (2.6)$$

- $\psi_{j,k}(t) = \psi_{j,k-t}$  is a scalar representing a non-decimated wavelet.

We can define the *Cross-Evolutionary Wavelet Spectrum*  $(N \times N)$  matrix :  $\mathbf{S}_j(z) = \mathbf{W}_j(z) \mathbf{W}_j(z)'$ . This gives us the ability to express the *local autocovariance* :  $c^{(u,u')}(z, \tau) = \sum_{j=-\infty}^{-1} S_j^{(u,u')}(z) \boldsymbol{\Psi}_j(\tau)$  where  $\boldsymbol{\Psi}_j(\tau) = \sum_k \psi_{j,k}(0) \psi_{j,k}(\tau)$ , the *autocorrelation wavelet*. The latter also define the *inner product matrix of discrete autocorrelation wavelets* :  $A_{jl} = \sum_{\tau} \boldsymbol{\Psi}_j(\tau) \boldsymbol{\Psi}_l(\tau)$ ,  $A = \{A_{jl}\}_{j,l \in \mathbb{N}}$  and its inverse :  $\bar{A} = A^{-1}$ . A rather simple extension of the autocorrelation wavelet is the *cross-correlation wavelet* which characterizes the dependence between two wavelets at different scales. The latter wavelet is thus defined as  $\boldsymbol{\Psi}_{j,j'}(\tau) = \sum_k \psi_{j,k}(0) \psi_{j',k}(\tau)$ .

Each  $(m, n)$ -element of the Cross-Evolutionary Wavelet Spectrum can be expressed as

$$S_j^{(m,n)}(k/T) = \sum_{u=1}^N W_j^{(m,u)}(k/T) W_j^{(u,n)}(k/T), \quad \forall j, k$$

From this definition it is not difficult to extend the CEWS to take into account the dependence structure between different scales and through time :

$$S_{j,j'}^{m,n}(k/T, k'/T) = \sum_{u=1}^N W_j^{(m,u)}(k/T) W_{j'}^{(u,n)}(k'/T), \quad \forall j, j', k, k' \quad (2.7)$$

We make the following assumption regarding the latter object,

$$S_{j,j'}^{(m,n)}(k/T, k'/T) = \begin{cases} S_j^{(m,n)}(k/T) & \text{if } j = j' \\ 0 & \text{otherwise} \end{cases} \quad (2.8)$$

This assumption (possible improvement : condition similar to Chamberlain ?) imposes no dependence between different scales of decomposition. Notice that we don't restrict the serial dependence.

## 2.1 Estimation of MvLSW

- $E[\mathbf{I}_{j,k}] = \sum_{l=-J}^{-1} A_{jl} \mathbf{S}_l(k/T) + O(T^{-1})$  (biased estimator)

## 2.2 Estimation of MvLSW

$$\mathbf{d}_{j,k} = \sum_{t=0}^{T-1} \mathbf{X}_t \psi_{j,k}(t) \quad (\text{empirical wavelet coefficients}) \quad (2.9)$$

$$\mathbf{I}_{j,k} = \mathbf{d}_{j,k} \mathbf{d}_{j,k}' \quad (\text{raw wavelet periodogram}) \quad (2.10)$$

$$\bar{\mathbf{I}}_{j,k} = \sum_{l=-J}^{-1} \bar{A}_{jl} \mathbf{I}_{l,k} \quad (\text{corrected periodogram}) \quad (2.11)$$

$$\tilde{\mathbf{I}}_{j,k} = \frac{1}{2M+1} \sum_{m=-M}^M \mathbf{I}_{j,k+m} \quad (\text{smooth periodogram}) \quad (2.12)$$

$$\begin{aligned} \hat{\mathbf{S}}_j(k/T) &= \sum_{l=-J}^{-1} \bar{A}_{jl} \tilde{\mathbf{I}}_{l,k} \\ &= \frac{1}{2M+1} \sum_{m=-M}^M \bar{\mathbf{I}}_{j,k+m} \\ &= \frac{1}{2M+1} \sum_{m=-M}^M \sum_{l=-J}^{-1} \bar{A}_{jl} \mathbf{I}_{l,k+m} \end{aligned} \quad (\text{final estimator of CEWS}) \quad (2.13)$$

## 2.3 Notes

- The dependence structure is entirely in  $\mathbf{W}_j(z)$ , not in  $\boldsymbol{\xi}_{j,k}$ .
- The lower-triangular form of  $\mathbf{W}_j(z)$  allows us to use the Cholesky decomposition on  $\mathbf{S}_j(z)$ .

## 3 Factor Model

- The factor structure is imposed on the following :

$$\mathbf{W}_j(k/T) \boldsymbol{\xi}_{j,k} = \boldsymbol{\Lambda}_{j,k} \mathbf{F}_k + \boldsymbol{\epsilon}_{j,k} \quad (3.1)$$

, not only on  $\boldsymbol{\xi}_{j,k}$  since they are assumed orthonormal.

- Assumptions :

1.  $\mathbf{F}_k \sim (\mathbf{0}, \boldsymbol{\Sigma}_F)$ , where  $\boldsymbol{\Sigma}_F$  is a diagonal positive definite  $(K \times K)$  matrix.
2.  $E[F_k^{(u)4}] < \infty, \forall k, u$
3.  $\mathbf{F}_k \perp \boldsymbol{\epsilon}_{j,k'}, \forall j, k, k'$
4.  $\boldsymbol{\epsilon}_{j,k} \sim (\mathbf{0}, \boldsymbol{\Sigma}_\epsilon)$ , where  $\boldsymbol{\Sigma}_\epsilon$  has bounded eigenvalues. *Note : make  $\boldsymbol{\Sigma}_\epsilon$  dependent on time ? what about serial dependence ?*
5.  $\boldsymbol{\Lambda}_{j,k}' \boldsymbol{\Lambda}_{l,m} = \mathbf{0}, \forall j \neq l, \forall k \neq m$ .

- We can then represent the CEWS with the factor structure :

$$\begin{aligned} \text{Var}[\mathbf{W}_j(k/T) \boldsymbol{\xi}_{j,k}] &= \text{Var}[\boldsymbol{\Lambda}_{j,k} \mathbf{F}_k] + \text{Var}[\boldsymbol{\epsilon}_{j,k}] \\ \mathbf{W}_j(k/T) \text{Var}[\boldsymbol{\xi}_{j,k}] \mathbf{W}_j(k/T)' &= \boldsymbol{\Lambda}_{j,k} \text{Var}[\mathbf{F}_k] \boldsymbol{\Lambda}_{j,k}' + \text{Var}[\boldsymbol{\epsilon}_{j,k}] \\ \mathbf{W}_j(k/T) \mathbf{W}_j(k/T)' &= \boldsymbol{\Lambda}_{j,k} \boldsymbol{\Sigma}_F \boldsymbol{\Lambda}_{j,k}' + \boldsymbol{\Sigma}_\epsilon \\ \mathbf{S}_j(k/T) &= \boldsymbol{\Lambda}_{j,k} \boldsymbol{\Sigma}_F \boldsymbol{\Lambda}_{j,k}' + \boldsymbol{\Sigma}_\epsilon \end{aligned} \quad \text{from (2.6)}$$