MATH 534 HOMEWORK 6

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1.)

```
Time = c(0.25, 0.5, 0.75, 1, 1.5, 2, 3, 4, 6, 12, 24, 48)
Concentration = c(215.6, 189.2, 176, 162.8, 138.6, 121, 101.2, 88, 61.6, 22, 4.4, 0)
model = function(alpha_1, lambda_1, alpha_2, lambda_2, t){
 E_1 = \exp(\lambda_1 + t)
 E_2 = \exp(-lambda_2*t)
 f = alpha_1*E_1 + alpha_2*E_2
 grad = cbind(E_1, alpha_1*t*E_1, E_2, -alpha_2*t*E_2)
 attr(f,'gradient') = grad
 return(f)
}
Result = nls(Concentration ~ model(alpha_1, lambda_1, alpha_2, lambda_2, Time),
        start=list(alpha_1=100, lambda_1=0.05, alpha_2=100, lambda_2=0.05), trace = TRUE,
        nls.control(maxiter = 50, tol = 1e-5, minFactor = 1/1024, printEval=TRUE))
## 1467426 : 1e+02 5e-02 1e+02 5e-02
## 291337.6 : 117.70889622
                             0.02792691 78.34186855
                                                       0.36480855
## 14578.49 :
               15.72534624
                             0.02504509 201.31255034
                                                       0.14968154
## 1555.431 :
               19.699747699 -0.001236943 189.572837396
                                                          0.240197105
## 550.3467 :
              25.09567158 -0.03041565 190.82812021
                                                       0.29130563
## 416.1091 :
              56.00672585 -0.09272928 163.86795105
                                                       0.35424356
## 395.9537 : 133.3465269 -0.1686788 90.4777954
                                                   0.4934618
## 373.9764 : 156.4320010 -0.1784388 68.9979824
                                                    0.6033732
## 333.6786 : 166.3855802 -0.1812592 60.7703647
                                                    0.7203197
## 256.5963 : 173.6974436 -0.1810699 56.9454141
                                                    0.9306454
## 97.12377 : 171.743499 -0.173778 65.120558
                                                 1.207145
## 35.62475 : 162.5897018 -0.1618363 81.2745180
                                                   1.3302877
## 34.38149 : 162.5295380 -0.1617149 81.2564401
                                                    1.3032926
## 34.38027 : 162.6081936 -0.1618016 81.2367152
                                                   1.3063819
```

We can see that our estimates our $\alpha_1 = 162.59$, $\lambda_1 = -.1618$, $\alpha_2 = 81.24$, and $\lambda_2 = 1.306$.

34.38026 : 162.5968351 -0.1617894 81.2418058

34.38026 : 162.5981307 -0.1617908 81.2412402

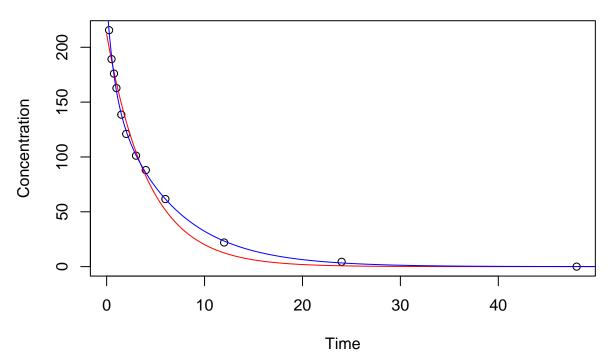
Now, we will plot our one component (RED line) and two component (BLUE line) models against the data provided.

1.3060291

1.3060701

```
plot(Time, Concentration, main='Data and the Fitted Curves')
t = seq(0,50, 0.1)
lines(t,211.9203*exp(-0.2357*t),col='red')
lines(t, 162.6*exp(-0.1618*t) + 81.24*exp(-1.306*t), col = 'blue')
```

Data and the Fitted Curves



By comparing the two models, we can observe that the two component model fits the data better than our one component model.

2.)

```
\mathbf{A}
```

```
dose=c(0,2,10,50,250)
animals=c(111,105,124,104,90)
tumor =c(4,4,11,13,60)

data=cbind(dose,animals,tumor)
data=as.matrix(data)
```

The IRLS algorithm is derived as follows:

```
# Start of the IRLS algorithm
X = data[,1:2]
y = as.matrix(data[,3])
JacfW_Binom = function(beta,X){
 x=X[,1]
 n=X[,2]
 p = (1-exp(-beta[1]-x*beta[2]-(x^2)*beta[3]))
 q=1-p
 f=n*p
  db0 = n*exp(-beta[1]-x*beta[2]-(x^2)*beta[3])
  db1 = n*x*exp(-beta[1]-x*beta[2]-(x^2)*beta[3])
  db2 = n*(x^2)*exp(-beta[1]-x*beta[2]-(x^2)*beta[3])
  J =cbind(db0,db1,db2)
 W = diag(1/(n*p*q))
 list(f=f, J=J, W=W) }
GN = function(y,X, beta0, Jac, Wt = 1, maxit,IRLS = TRUE){
                         b0", "
cat("Iteration", "
             b2". "
                            MRE", "
                                                Digits\n")
```

```
for (it in 1: maxit){
   a = do.call(Jac, list(beta0, X))
   J = a J
   f = a f
   if (IRLS==TRUE){
     Wt = a\$W 
   JW = t(J) %*% Wt
   dir = solve(JW %*% J) %*% JW %*% (y-f)
   beta1 = beta0 + dir
   relerr = norm(beta1-beta0)/max(1,norm(beta1))
   digits = -log10(norm(beta1-beta_star)/norm(beta_star))
                               %6.6f
cat(sprintf(' %2.0f
                                        %6.6f
                                                  %6.6f
                                                            %1.1e
                                                                          %1.1e\n'
                ,it,beta0[1],beta0[2],beta0[3],relerr, digits))
   beta0=beta1
  }
 MLE_beta<<-beta0
```

```
maxit=20
beta0 = matrix(c(.1,.001,.00001),3,1)
#found beta star by outputting more digits.
beta_star = matrix(c(0.045945, 0.001627, 0.00001019))
mle = GN(y,X, beta0, 'JacfW_Binom', Wt = 1, maxit, IRLS = TRUE)
```

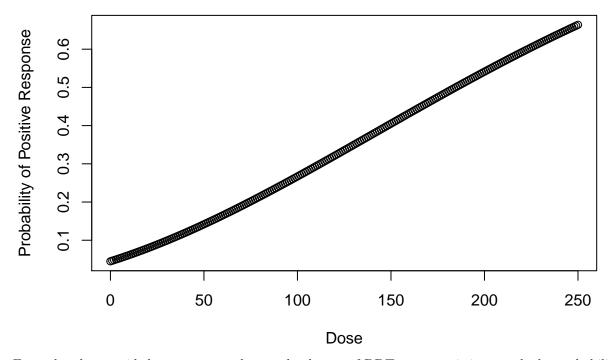
##	Iteration	b0	b1	b2	MRE	Digits
##	1	0.100000	0.001000	0.000010	5.4e-02	1.7e+00
##	2	0.046765	0.001434	0.000011	7.1e-04	2.2e+00
##	3	0.046228	0.001605	0.000010	2.6e-04	3.0e+00
##	4	0.045987	0.001624	0.000010	3.9e-05	3.9e+00
##	5	0.045951	0.001627	0.000010	6.0e-06	4.9e+00
##	6	0.045946	0.001627	0.000010	9.3e-07	5.2e+00
##	7	0.045945	0.001627	0.000010	1.4e-07	5.1e+00
##	8	0.045945	0.001627	0.000010	2.2e-08	5.0e+00
##	9	0.045945	0.001627	0.000010	3.5e-09	5.0e+00
##	10	0.045945	0.001627	0.000010	5.4e-10	5.0e+00
##	11	0.045945	0.001627	0.000010	8.3e-11	5.0e+00
##	12	0.045945	0.001627	0.000010	1.3e-11	5.0e+00
##	13	0.045945	0.001627	0.000010	2.0e-12	5.0e+00
##	14	0.045945	0.001627	0.000010	3.1e-13	5.0e+00
##	15	0.045945	0.001627	0.000010	4.8e-14	5.0e+00
##	16	0.045945	0.001627	0.000010	7.4e-15	5.0e+00
##	17	0.045945	0.001627	0.000010	1.2e-15	5.0e+00
##	18	0.045945	0.001627	0.000010	1.9e-16	5.0e+00
##	19	0.045945	0.001627	0.000010	3.6e-17	5.0e+00
##	20	0.045945	0.001627	0.000010	7.2e-18	5.0e+00

Thus, our MLE estimates are $\beta_0 = 0.0459$, $\beta_1 = 0.0016$, and $\beta_2 = 0.00001$. We can observe that the estimates are found before the 20 iterations are met.

\mathbf{B}

Now, we will plot the probability of positive response, using our MLE estimates.

```
x = seq(0,250, length=250)
y= 1 - exp(-MLE_beta[1] - MLE_beta[2]*x - MLE_beta[3]*x^2)
plot(x,y, xlab="Dose", ylab="Probability of Positive Response")
```



From the plot provided, we can note that as the dosage of DDT exposure is increased, the probability of positive response of an animal having a tumor also increases.

3.)

```
\mathbf{A}
```

```
TreeData = read.table('/Users/gustavo/Desktop/Math 534/8/trees.txt', header=TRUE)

D=TreeData$D
S=TreeData$S

X = as.matrix(cbind(1, log(D), S))
y=TreeData$y
```

The IRLS algorithm is derived as follows:

```
JacfW_Bern = function(beta,X){

xb = X%*%beta
exb = exp(xb)
p = exb/(1+exb)
q=1-p
f=p
frac = exb/((1+exb)^2)

J= (diag(as.vector(frac))) %*% X
W = diag(1/as.vector(p*q))
list(f=f, J=J, W=W)}
```

```
maxit = 20
beta_0=matrix(c(0,0,0))
#found beta star by outputting more digits.
beta_star = matrix(c(-9.562084874971, 3.197563503535, 4.508593181753))
MLE_beta = GN(y, X, beta_0, 'JacfW_Bern', Wt = 1, maxit, IRLS = TRUE)
```

##	Iteration	b0	b1	b2	MRE	Digits
##	1	0.000000	0.000000	0.000000	1.0e+00	4.4e-01
##	2	-6.124979	2.035718	2.857737	3.0e-01	1.0e+00
##	3	-8.673430	2.895867	4.076634	8.7e-02	2.1e+00
##	4	-9.492693	3.173968	4.474456	7.3e-03	4.3e+00
##	5	-9.561642	3.197413	4.508373	4.7e-05	8.7e+00
##	6	-9.562085	3.197563	4.508593	1.9e-09	1.3e+01
##	7	-9.562085	3.197564	4.508593	1.8e-16	1.3e+01
##	8	-9.562085	3.197564	4.508593	1.8e-16	1.3e+01
##	9	-9.562085	3.197564	4.508593	2.1e-16	1.3e+01
##	10	-9.562085	3.197564	4.508593	3.6e-16	1.3e+01
##	11	-9.562085	3.197564	4.508593	3.3e-16	1.3e+01
##	12	-9.562085	3.197564	4.508593	1.0e-16	1.3e+01
##	13	-9.562085	3.197564	4.508593	7.7e-17	1.3e+01
##	14	-9.562085	3.197564	4.508593	2.1e-16	1.3e+01
##	15	-9.562085	3.197564	4.508593	3.6e-16	1.3e+01
##	16	-9.562085	3.197564	4.508593	3.3e-16	1.3e+01
##	17	-9.562085	3.197564	4.508593	1.0e-16	1.3e+01
##	18	-9.562085	3.197564	4.508593	7.7e-17	1.3e+01
##	19	-9.562085	3.197564	4.508593	2.1e-16	1.3e+01
##	20	-9.562085	3.197564	4.508593	3.6e-16	1.3e+01

Thus, our MLE estimates are $\beta_0 = -9.5621$, $\beta_1 = 3.1976$, $\beta_2 = 4.508$. We can again observe that our estimates are obtained in very few interations.

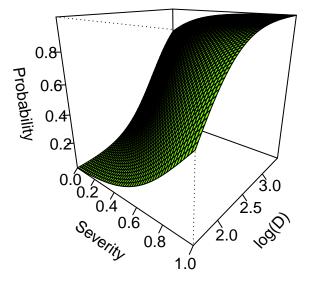
 \mathbf{B}

In order to plot our graph, we need to define two preliminary functions that compute our probability at a specific value and over a general interval.

```
#single value
pi_bern =function(s,logd){
    x= c(1,logd,s)
    xb = x%*%beta_star
    exb = exp(xb)
    p = exb/(1+exb)
    as.numeric(p)}
#interval
pi_bern2 = function(x1,x2){
    n=length(x1)
    pi=numeric(n)
    for(i in 1:n){
        pi[i] = pi_bern(x1[i],x2[i])}
        pi}
```

Now, we will establish the interval at which we wish to plot our function.

PI(B)



From the 3-D plot, we can observe a full perspective of how wind severity and tree diameter affect the

probability of a tree being blown down.

 ${\bf C}$ We will use our previously derived function for the prediction.

pi_bern(.3,log(10))

[1] 0.3000951

Thus, a tree with a diameter of 10 located in an area with wind severity of 0.3 has a probability of 0.30 of being blown down.

4.)

A.) If the response distribution is normal, then our Likelihood function for β is as follows:

$$L(\beta) = (2\pi\sigma^2)^{\frac{-n}{2}} * exp\left(\frac{-\sum_{i=1}^{n} (y_i - \mu_i)^2}{2\sigma^2}\right)$$

Since our exponent is being raised to a negative value, we can note that our Likelihood function is a decreasing function. Thus, our MLE occurs when the varying terms (those affected by the index i) are minimized. That is, when the following is minimized:

$$\sum_{i=1}^{n} (y_i - \mu_i)^2 = \sum_{i=1}^{n} (y_i - X_i^T \beta)^2$$

Which is equivalent to the L-2 norm.

B If the response distribution is double exponential, then our Likelihood function for β is as follows:

$$L(\beta) = (2\sigma)^{-n} * exp\left(\frac{-\sum_{i=1}^{n} |y_i - \mu_i|}{2\sigma^2}\right)$$

Again, since our exponent is being raised to a negative value, we can note that our Likelihood function is a decreasing function. Thus, our MLE occurs when the varying terms (those affected by the index i) are minimized. That is, when the following is minimized:

$$\sum_{i=1}^{n} |y_i - \mu_{i|} = \sum_{i=1}^{n} |y_i - X_i^T \beta|$$

Which is equivalent to the L-1 norm.

C If the response distribution is uniform, then our density, $f(y_i)$, can be expressed in the following manner:

$$f(y_i) \propto (1_{(\mu_i - \sigma_i, \mu_i + \sigma_i)})^{(y_i)}$$

$$\propto 1 \text{ if } \mu_i - \sigma_i < y_i < \mu_i + \sigma_i , 0 \text{ o.w.}$$

$$\propto 1 \text{ if } |y_i - \mu_i| < \sigma_i , 0 \text{ o.w.}$$

Thus, we can observe that the likelihood function of β is simply the product of the indicator function provided above over all y_i values. Then, the likelihood function is maximized when the likelihood function is equivalent to 1. This happens when all of our y_i values satisfy our indicator function, and more importantly, this can only occur when the maximum y_i value satisfies the following condition:

$$max_i|y_i - \mu_i| = max_i|y_i - X_i^T\beta| < \sigma_i$$

This is equivalent to minimizing the $L-\infty$ norm.