

# Renewal Theory

(1) Recurrent Times:

① Forward:

Def: For  $\varphi = (t_n)$ , renewal process.  $t_{\mu(n)} \leq t \leq t_{\mu(n)+1}$

the forward recurrent time is  $\Delta t_f =$

$$t_{N(t)+1} - t \quad ( \text{Timeline from } t \text{ to } t_{N(t)+1} \text{ with } A(t) \text{ events} )$$

Exk: If  $\phi$  is Poisson( $\lambda$ ) process. Then by

memoryless property:  $A(t) \sim \text{Exp}(\lambda)$ .

Denote:  $X = (X_n)$  i.i.d interarrival times.  $X \sim F$

prop. i)  $\lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t A(s) ds = E(X^2) / 2E(X) \text{ a.s.}$

$$ii) \lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t I_{\{A(s) > a\}} ds = \lambda E (X - a)^+, \text{ a.s.}$$

$$\text{iii) } \lim_{t \rightarrow \infty} \frac{1}{t} \int_1^t P(A(s) > a) ds = \lambda E(X - a)^+.$$

Pf: 1) Intuitively,  $\int_0^t A(s) ds \approx \sum_1^{N(t)} \frac{X_k^2}{2}$ .

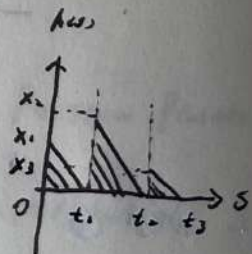
Actually, we have:

$$\frac{1}{t} \sum_{i=1}^{N(t)} \frac{X_i^2}{2} \approx \frac{1}{t} \int_0^t A(s) ds \approx \frac{1}{t} \sum_{i=1}^{N(t)+1} \frac{X_i^2}{2}$$

Then apply ERT. as before

$$2) \int_0^t I_{\{A(s) > n\}} ds \approx \sum_i^{N(t)} (X_k - n)^+$$

3') By DCT. in the expectation.

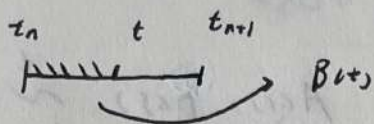




## ② Backward:

Def: For  $\varphi = \{t_n\}$  renewal process.  $t_{n(s)} \leq t \leq t_{n(s)+1}$ .

$B(t) = t - t_{n(s)}$  is backward recurrence time.

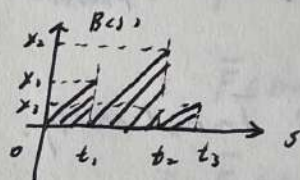


prop, i)  $\lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t B(s) \lambda ds = E(X) / 2 E(X)$ . a.s.

ii)  $\lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t I_{\{B(s) > a\}} \lambda ds = \lambda E(X - a)^+$ . a.s.

iii)  $\lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t p(B(s) > a) \lambda ds = \lambda E(X - a)^+$ .

Pf: Similar as before:



## (2) Distribution:

### ① Equilibrium List:

Denote:  $\bar{F}(x) = 1 - F(x) = p(X > x)$ . tail prob.

rmk:  $\lambda E(X - a)^+ = \lambda \int_a^\infty p(X > \eta) \lambda \eta$   
 $= \lambda \int_a^\infty \bar{F}(\eta) \lambda \eta.$

Def: As  $n \geq 0$  varies.  $\rho_{mk}$  defines a tail prob. of a list. We denote cdf of it by  $F_e$ , which is called equilibrium list. of  $F$ .

$F_e(x) =: \lambda \int_0^x \bar{F}(\eta) \lambda \eta.$  Set  $X_e \sim F_e$ . r.v.

rmk:  $X_e$  is conti. since density  $= F_e' = \lambda \bar{F}$  exists.

Note that in a long term:

$$\lim_{t \rightarrow \infty} \int_0^t \frac{p(A(s) < x)}{t} ds = \lim_{t \rightarrow \infty} \int_0^t \frac{p(B(s) < x)}{t} ds = 1 - \lambda \overline{E}(X - x)^+ \\ = \lambda \int_0^x \overline{F}(\eta) d\eta = F_e(x).$$

$\Rightarrow$  The stationary dist of  $A(s), B(s) \stackrel{d}{\sim} F_e$

② Spread:

Def: Spread as length of interarrival time is

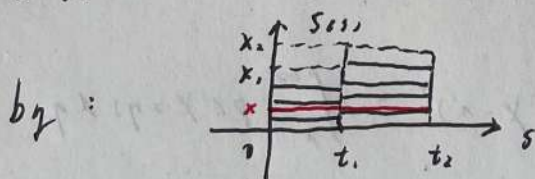
$$S(t) =: t_{\mu_{n+1}} - t_{\mu_n} = A(t) + B(t) = X_{\mu_{n+1}}$$

Remark: Immediately,  $\lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t S(s) ds = \overline{E}(X) / \overline{E}(X)$ .

Prop. i)  $\lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t I_{\{S(s) > x\}} ds = \lambda \overline{E}(X I_{\{X > x\}})$  a.s.

$$ii) \lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t p(S(s) > x) ds = \lambda \overline{E}(X I_{\{X > x\}})$$

Pfc:  $\frac{1}{t} \int_0^t I_{\{S(s) > x\}} ds \approx \frac{1}{t} \sum_{i=1}^{N(t)} X_i I_{\{X_i > x\}}$



Remark: In a long term:

$$\lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t p(S(s) < x) ds = \lambda \overline{E}(X I_{\{X < x\}})$$

$$\text{Denote: } \overline{F}_s(x) = \lambda \overline{E}(X I_{\{X > x\}})$$

$$= \lambda \int_x^\infty (t-x)^+ + x dF$$

$$= \lambda \overline{E}(X - x)^+ + \lambda x \overline{F}(x).$$

$$= \overline{F}_e(x) + \lambda x \overline{F}(x).$$



by prop above. the stationary dist. of  $S(x) \sim 1 - \bar{F}_S(x) = F_S(x)$ . Assume r.v.  $X_s$ .

$\Rightarrow X_s$  may not be conti. generally.  $\{F' \text{ doesn't exist}\}$ .  $\bar{E}(X_s) = \bar{E}(X) / \bar{E}(x)$ .

### (3) Inspection Paradox:

prop.  $S(t) \geq_{\text{stoch}} X$ . i.e.  $p(X > x) \leq p(S(t) > x)$ .  $\forall t, x \geq 0$ .

Moreover.  $p(X_s \geq x) \geq p(X \geq x)$ .  $X_s \geq_{\text{stoch}} X$ .

Pf: 1')  $p(S(t) > x \mid N(t) = n, t_n = s) = p(X_{n+1} > x \mid X_{n+1} > t-s)$   
 $= \bar{F}(\max\{x, t-s\}) / \bar{F}(t-s)$   
 $\geq \bar{F}(x)$ .

$\Rightarrow \bar{E}_{N, t_N}(p(S(t) > x \mid \square)) = p(S(t) > x) \geq \bar{F}(x)$

2') By  $p(X_s \geq x) = \lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t p(S(s) > x) ds \geq p(X \geq x)$

Rmk. Note that:  $\bar{E}(S(t)) \geq \bar{E}(X)$

$\bar{E}(X_s) \geq \bar{E}(X)$

$\Rightarrow$  The prop. above implies a paradox: Observing makes the expectation of lifetime be longer than usual.

That's because if we observe at time  $t$ , then the part of lifetime  $< t$  will be ignored.

e.g. A extreme case:

$X$  is r.v. of lifetime of bulbs.  $\begin{cases} p(X=0) = 0.9 \\ p(X=1) = 0.1 \end{cases}$

Then. all bulbs we observed has lifetime 1.

The right way to estimate lifetime of them

is SLLN:  $\frac{1}{n} \sum_{i=1}^n X_k \approx E(X)$ . for large  $n$ .

#### (4) Renewal Reward Thm:

Suppose  $R_i$  is r.v. of reward between  $t_i, t_{i+1}$ .

$\Rightarrow R(t) = \sum_{i=1}^{N(t)} R_j$  Assume  $(X_i, R_i)_{i \in \mathbb{Z}^+}$  i.i.d. but  $R_i$

can depend on  $X_i$ .  $(X, R)$  denote a cycle.

#### Thm. (RRT)

For a positive recurrent renewal process

where a reward  $R_i$  is earned during a

cycle with length  $X_i$ . If  $E(|R_i|) < \infty$ .

Then:  $\lim_{t \rightarrow \infty} \frac{R(t)}{t} = E(R)/E(X)$  a.s.

$$\lim_{t \rightarrow \infty} \frac{E(R(t))}{t} = E(R)/E(X).$$

p.f. 1°) Suppose  $R_i \geq 0$ . or set  $R_i = R_i^+ - R_i^-$ .

$$N(t): \frac{1}{t} \sum_{i=1}^{N(t)} R_i \leq \frac{R(t)}{t} \leq \frac{1}{t} \sum_{i=1}^{N(t)} R_i.$$

Apply ERT.

$$2^\circ) |R(t)|/t \leq Y(t) = \frac{1}{t} \sum_{i=1}^{N(t)} |R_i| \xrightarrow{\text{a.s.}} E(|R|)/E(X)$$

$\Rightarrow |R(t)|/t$  is bdd a.s. By DCT.



Application: We can see some byproduct as "reward" to calculate its long time rate by RRT.

i) Let  $R(t) = \int_0^t A(s) ds$ . Conti reward.

$$R_j = \int_{t_{j-1}}^{t_j} A(s) ds = \int_{t_{j-1}}^{t_j} (t_j - s) \lambda ds$$

$$\stackrel{s=t_{j-1}+u}{=} \int_0^{X_j} (X_j - u) \lambda du = \frac{X_j^2}{2}$$

$$\Rightarrow \text{By RRT: } \frac{1}{t} \int_0^t A(s) ds \xrightarrow{a.s.} E\left(\frac{X^2}{2}\right) / E(X).$$

ii) Similarly,  $R(t) = \int_0^t B(s) ds$ .

iii)  $R(t) = \int_0^t I_{\{A(s) > x\}} \lambda ds$ .  $R_i = \int_0^{X_i} I_{\{X_i - s > x\}} \lambda ds$

### (5) Central Limit Thm:

Consider renewal process  $\{t_n\}$  with interarrival times

$$X_n = t_n - t_{n-1}, \text{ s.t. } E(X) = \frac{1}{\lambda}, \text{ Var}(X) = \sigma^2.$$

Thm. CCLT for counting process

$$Z(t) = \frac{N(t) - \lambda t}{\sigma \sqrt{\lambda^3 t}} \xrightarrow[t \rightarrow \infty]{d} Z \sim N(0,1).$$

RMk: It implies:  $E(N(t)) \sim \lambda t$ .  $\text{Var}(N(t)) \sim \sigma^2 \lambda^3 t$

Pf: The key is:  $P(N(t) < n) = P(t_n > t)$ .

$$\text{set } n(t, x) = \lfloor \lambda t + x \sqrt{\sigma^2 \lambda^3 t} \rfloor.$$

$$\begin{aligned} \Rightarrow P(Z(t) < x) &= P(N(t) < n(t, x)) \\ &= P(t_{n(t, x)} > t) \end{aligned}$$



$$= P\left( \frac{t_{n(t,x)} - \mu(t,x)/\lambda}{\sigma \sqrt{\mu(t,x)}} > \frac{t - \mu(t,x)/\lambda}{\sigma \sqrt{\mu(t,x)}} \right)$$

Note by CLT,  $\frac{t_n - \mu/\lambda}{\sigma \sqrt{\mu}} \xrightarrow{d} Z \sim N(0,1)$ .

$$\text{So: } P(Z(t) < x) \xrightarrow{t \rightarrow \infty} P(Z > -x) = P(Z < x)$$

Remark: We don't need "Renewal" actually,  
but require:  $\varphi = \{t_n\}$  satisfies CLT.

### (6) Delayed Renewal Process:

Note that the counting process  $N(t)$  of renewal process  $\varphi = \{t_n\}$  generally doesn't have stationary increment like poisson process.

Def:  $\varphi_s$  is a shift by time  $s$  version of point process  $\varphi = \{t_n\}$ , which is  $\{t_n + s\}$ , moving the origin to be  $t=s$ . with counting process;  
 $N_s(t) = N(s+t) - N(s)$ .

Remark:  $\varphi_s \stackrel{d}{\sim} \varphi$ .  $\forall s \geq 0 \Leftrightarrow N(t)$  has stationary increment.

Def: A delayed renewal process is a renewal process where the first arrival time  $t_1 = X_1$  indeptly has a different dist.  $\sim F$ .  
For  $(X_n)_{n \geq 2}$  interarrival times. are i.i.d.  $\sim F$

Remark: 1) ERT remains valid.  $F$  doesn't need



to have finite first moment.

ii)  $\varphi$  is stationary renewal process  $\Leftrightarrow t, (s) = A(s)$   
of  $\varphi_s$  has the same list.  $\forall s \geq 0$ .

As  $s \rightarrow \infty$ .  $\varphi_s$  has a limiting list. of the delayed  
version  $\varphi$ . denoted by  $\varphi^* = \{t_n^*\}_{n \geq 1}$ .  $t_i^* \sim F_e$ .  
With its counting process  $N^*(t)$ . forward recurrent  
time process  $\{A^*(t)\}_{t \geq 0}$ .

Def.  $\varphi^*$  is stationary version of  $\varphi$ .

Prop.  $\varphi^*$  is stationary renewal process:

$$p(A^*(u) \leq x) = \lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t p(A(s+u) \leq x) ds \\ = \lim_{t \rightarrow \infty} \frac{1}{t} \int_u^{t+u} p(A(s) \leq x) ds = F_e(x)$$

$$\Rightarrow A^* \sim F_e. \text{ i.e. } t_i^* \sim F_e. \varphi_s^* = \varphi^*.$$

prop. Stationary version  $\varphi^*$  of renewal process with  
rate  $\lambda = E(X)^{-1}$  satisfies:  $E(N^*(t)) = \lambda t$ . and  
 $\lambda = E(N^*(1))$ .

Pf. Denote  $\mu(t) = E(N^*(t)) \Rightarrow \mu(n) = n\mu(1). \forall n$ .

So:  $\mu(t) = t\mu(1)$ . With  $\mu(t)/t \rightarrow \lambda$ . So  $\mu(1) = \lambda$ .

follows from ERT.  $\lim_{t \rightarrow \infty} \lim_{s \rightarrow \infty} \square = \lim_{s \rightarrow \infty} \lim_{t \rightarrow \infty} \square$

## (7) Renewal Equations:

For renewal process  $\varphi = \{t_n\}$  with  $p(X \leq y) = F_e(y)$ .



$$\begin{aligned}
 \text{Note: } p(A(t) > x) &= p(A(t) > x | X_1 > t) p(X_1 > t) + \int_0^t p(\square | X_1 = s) p(X_1 = s) ds \\
 &= p(X_1 - t > x) + \int_0^t p(A(t-s) > x) dF(s) \\
 &= \bar{F}(t+x) + \int_0^t p(A(t-s) > x) dF(s)
 \end{aligned}$$

Denote:  $\bar{F}(x+t) = Q(t)$ .  $H(t) = p(A(t) > x)$ .

$\Rightarrow H(t) = Q(t) + H * F(t)$ . renewal equation.

By Iteration:  $H(t) = Q + (Q + H * F) * F$

$$= \dots = Q + \sum_{n=1}^{\infty} Q * F^{*n}. \quad H \text{ is unknown}$$

prop.  $m(t) = E(N(t))$  satisfies:  $H(t) = Q + Q * m(t)$ . (\*)

Pf:  $E(N(t)) = E\left(\sum_i I_{[t_i, t_{i+1})}\right) = \sum F^{*n}$ .

since  $F^{*n}(t) = p\left(\sum_{i=1}^n X_i \leq t\right) = p(t_n \leq t) \xrightarrow{n \rightarrow \infty} 0$

Thm. (key renewal Thm)

If the renewal equation holds for given non-lattice  $F$  with mean  $1/\lambda$  and  $Q$  is DRI (directly Riemann integrable, i.e.  $\int_0^\infty Q(t) dt$  exists)

Then solution for (\*) holds.  $\lim_{t \rightarrow \infty} H(t) = \lambda \int_0^\infty Q(t) dt$ .

Rmk: These remain valid for delayed case:

in sense,  $F_0 = F$ .  $F_1$  is delay  $X_1$ 's dist. then.

$$H_1 = Q_1 + H_0 * F_1, \quad H_0 = Q_0 + Q_0 * m_0, \quad m_0 = E(N(t))$$

$$\text{With } H_1 = Q_1 + Q_0 * m_1, \quad m_1 = E(N(t))$$

$$\text{And } \lim_{t \rightarrow \infty} Q_0 * m_1 = \lim_{t \rightarrow \infty} Q_0 * m_0 = \lambda \int_0^\infty Q_0(t) dt.$$

Pf: It's equi. with Blackwell's Thm