

# **Homework 02**

## **Vision-Aided Navigation**

**086761**

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<b>1.</b>	.....	
<b>2.</b>	.....	<b>1</b>
(a)	.....	1
(b)	.....	2
(c)	.....	2
(d)	.....	3
<b>3.</b>	.....	<b>4</b>
(a)	.....	4
(b)	.....	4
(c)	.....	4
(d)	.....	5
<b>4.</b>	.....	<b>6</b>
(a)	.....	
(b)	.....	6
(c)	.....	7
<b>5.</b>	.....	<b>7</b>
(a)	.....	
(b)	.....	7
(c)	.....	7
(d)	.....	10
<b>6.</b>	.....	<b>10</b>
(a)	.....	
(b)	.....	11
(c)	.....	11

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## 1.

To show the correspondence between the Gaussian distribution in its standard form and its information form, we start with the Gaussian distribution given by

$$x \sim \mathcal{N}(\mu, \Sigma), \quad (1.1)$$

which has the probability density function (PDF)

$$p(x) = \frac{1}{\sqrt{(2\pi)^n \det(\Sigma)}} \exp\left(-\frac{1}{2}(x - \mu)^\top \Sigma^{-1}(x - \mu)\right) \quad (1.2)$$

and we aim to express it in terms of the information form

$$x \sim N^{-1}(\eta, \Lambda). \quad (1.3)$$

where  $\Lambda = \Sigma^{-1}$  is the information, or precision matrix (inverse of the covariance matrix) and  $\eta = \Lambda\mu$  is the information vector. The PDF in the information form can be written as

$$N^{-1}(x; \eta, \Lambda) = \frac{e^{-\frac{1}{2}\eta^\top \Lambda^{-1} \eta}}{\sqrt{\det(2\pi \Lambda^{-1})}} e^{-\frac{1}{2}x^\top \Lambda x + \eta^\top x}. \quad (1.4)$$

To show this, we start by expressing the exponent of the Gaussian PDF in terms of  $\eta$  and  $\Lambda$ :

$$\begin{aligned} -\frac{1}{2}(x - \mu)^\top \Sigma^{-1}(x - \mu) &= -\frac{1}{2}x^\top \Lambda x + x^\top \Lambda \mu - \frac{1}{2}\mu^\top \Lambda \mu \\ &\stackrel{\eta = \Lambda \mu}{=} -\frac{1}{2}x^\top \Lambda x + \eta^\top x - \frac{1}{2}\eta^\top \Lambda^{-1} \eta \end{aligned}$$

The normalization factor also needs to be expressed in terms of  $\Lambda$ . Since  $\Lambda = \Sigma^{-1}$ , we have:

$$\sqrt{(2\pi)^n \det(\Sigma)} = \sqrt{\det(2\pi \Lambda^{-1})}.$$

Putting it all together, the PDF in the information form is:

$$N^{-1}(x; \eta, \Lambda) = \frac{e^{-\frac{1}{2}\eta^\top \Lambda^{-1} \eta}}{\sqrt{\det(2\pi \Lambda^{-1})}} e^{-\frac{1}{2}x^\top \Lambda x + \eta^\top x}$$

## 2.

Given the standard observation model involving a random variable  $x \in \mathbb{R}^n$ :

$$z = h(x) + v \quad (2.1)$$

where  $v \sim N(0, \Sigma_v)$ , and assuming the initial belief regarding the state  $x$  is Gaussian with mean  $\hat{x}_0$  and covariance  $\Sigma_0$ .

### (a)

We can write the expressions for the prior  $p(x)$  and the measurement likelihood  $p(z|x)$  as follows:

The prior distribution of  $x$ ,  $p(x)$ , is given by the initial belief about the state, which is a Gaussian distribution

$$\begin{aligned}
p(x) &= N(x; \hat{x}_0, \Sigma_0) \\
&= \frac{1}{\sqrt{(2\pi)^n \det(\Sigma_0)}} \exp\left(-\frac{1}{2}(x - \hat{x}_0)^\top \Sigma_0^{-1}(x - \hat{x}_0)\right)
\end{aligned}$$

The measurement likelihood  $p(z|x)$  describes the probability of observing  $z$  given the state  $x$ , taking into account the noise  $v$ . Since  $v$  is Gaussian distributed, and  $z = h(x) + v$ , the likelihood is also Gaussian, centered around  $h(x)$  with covariance  $\Sigma_v$ :

$$\begin{aligned}
p(z|x) &= N(z; h(x), \Sigma_v) \\
&= \frac{1}{\sqrt{(2\pi)^m \det(\Sigma_v)}} \exp\left(-\frac{1}{2}(z - h(x))^\top \Sigma_v^{-1}(z - h(x))\right)
\end{aligned}$$

Here,  $m$  is the dimension of  $z$ , and  $h(x)$  represents the nonlinear transformation applied to  $x$  before adding the measurement noise  $v$ . This formulation allows for incorporating both the prior knowledge about the state  $x$  and the information gained from observing  $z$  in a Bayesian framework.

### (b)

Given a measurement  $z_1$  acquired through the measurement model  $z = h(x) + v$ , where  $v$  is the measurement noise, and assuming  $v \sim N(0, \Sigma_v)$ , the posterior probability  $p(x|z_1)$  can be derived using Bayes' theorem. Bayes' theorem allows us to update our prior belief  $p(x)$  about the state  $x$  in light of the new evidence provided by the measurement  $z_1$ . The posterior probability  $p(x|z_1)$  is given by:

$$\begin{aligned}
p(x|z_1) &= \frac{p(z_1|x) \cdot p(x)}{p(z_1)} \\
&= \frac{p(z_1|x) \cdot p(x)}{\int_{-\infty}^{\infty} p(z_1|\tilde{x})p(\tilde{x})d\tilde{x}}
\end{aligned} \tag{2.2}$$

Here,  $p(z_1|x)$  is the measurement likelihood, which is the probability of observing  $z_1$  given the state  $x$ , and  $p(x)$  is the prior probability of the state  $x$  before observing  $z_1$ . The denominator  $p(z_1)$  serves as a normalization factor, ensuring that the posterior probabilities sum to 1.

### (c)

By incorporating the explicit expressions for  $p(x)$  and  $p(z_1|x)$  as defined within Gaussian distribution frameworks, we arrive at:

- The prior  $p(x) = N(x; \hat{x}_0, \Sigma_0)$
- The likelihood  $p(z_1|x) = N(z_1; h(x), \Sigma_v)$

Under the assumptions of minor errors and a linear correlation, it can be deduced that:

$$\begin{aligned} x^\star &= \hat{x} + \tilde{x} \\ h(x^\star) &= h(\hat{x}_0) + H\tilde{x} \end{aligned}$$

where  $H = \left. \frac{dh}{dx} \right|_{x=\hat{x}_0}$ . When applied to  $p(x|z)$ :

$$p(x|z) = \exp \left[ -\frac{1}{2} \left( \|x - \hat{x}_0\|_{\Sigma_0}^2 + \|z - h(\hat{x}_0) - H\tilde{x}\|_{\Sigma_v}^2 \right) \right]. \quad (2.3)$$

To minimize and derive the optimal solution, we utilize the relation  $\|a\|_{\Sigma}^2 = \left\| \Sigma^{-1/2} a \right\|^2$

$$\|x - \hat{x}_0\|_{\Sigma_0}^2 + \|z - h(\hat{x}_0) - H\tilde{x}\|_{\Sigma_v}^2 = \left\| \Sigma_0^{-1/2} x \right\|^2 + \left\| \Sigma_v^{-1/2} \left( z - h(\hat{x}_0) - H\tilde{x} \right) \right\|^2 \quad (2.4)$$

This leads to a linear least squares problem, with a well-established solution:

$$\tilde{x} = (A^\top A)^{-1} A^\top B$$

Here,

$$A = \begin{bmatrix} \Sigma_0^{-1/2} \\ \Sigma_v^{-1/2} H \end{bmatrix}, \quad B = \begin{bmatrix} 0 \\ \Sigma_v^{-1/2} h(\hat{x}_0 - z) \end{bmatrix}$$

Yielding:

$$\tilde{x} = (\Sigma_0^{-1} + H^\top \Sigma_v^{-1} H)^{-1} H^\top \Sigma_v^{-1} h(\hat{x}_0 - z) \quad (2.5)$$

Therefore, we have:

$$\begin{aligned} \hat{x}_1 &= \hat{x}_0 + (\Sigma_0^{-1} + H^\top \Sigma_v^{-1} H)^{-1} H^\top \Sigma_v^{-1} h(\hat{x}_0 - z) \\ \Sigma_1 &= (\Sigma_0^{-1} + H^\top \Sigma_v^{-1} H)^{-1} \end{aligned}$$

**(d)**

Under markov assumption

$$p(x|z_1, z_2) = \frac{p(z_2|x)p(x|z_1)}{p(z_2)}$$

Hence

$$\begin{aligned} \hat{x}_2 &= \hat{x}_1 + (\Sigma_1^{-1} + H^\top \Sigma_v^{-1} H)^{-1} H^\top \Sigma_v^{-1} (h(\hat{x}_1) - z) \\ \Sigma_2 &= (\Sigma_1^{-1} + H^\top \Sigma_v^{-1} H)^{-1} \end{aligned}$$

**3.**

**(a)**

$$p(x_k|x_{k-1}, u_{k-1}) = \frac{1}{\sqrt{|2\pi\Sigma_w|}} \exp\left(-\frac{1}{2}\|x_k - f(x_{k-1}, u_{k-1})\|_{\Sigma_w}^2\right)$$

**(b)**

using Bayes rule and Markov:

$$\begin{aligned} \underbrace{p(x_1|z_1, u_0)}_{post} &= \frac{p(z_1|x_1, u_0)p(x_1, u_0)}{p(z_1, u_0)} \propto \eta p(z_1|x_1, u_0)p(x_1, u_0) = \eta p(z_1|x_1)p(x_1, u_0) \\ &= \eta p(z_1|x_1) \int p(x_1, x_0|u_0) dx_0 = \eta p(z_1|x_1) \int p(x_1|x_0, u_0) p(x_0, u_0) dx_0 \\ &= \eta p(z_1|x_1) \int \underbrace{p(x_1|x_0, u_0)}_{prior} p(x_0) dx_0 \end{aligned}$$

also used motion and observation models.

**(c)**

also using the Bayes Theorem:

$$\begin{aligned} p(x_0, x_1|z_1, u_0) &= \frac{p(z_1|x_1, x_0, u_0)p(x_1, x_0, u_0)}{p(z_1, u_0)} \propto \eta p(z_1|x_1, x_0, u_0)p(x_1, x_0, u_0) \\ &= \eta p(z_1|x_1)p(x_1, x_0|u_0) \\ &= \eta p(z_1|x_1)p(x_1|x_0, u_0)p(x_0|u_0) \\ &= \eta p(z_1|x_1)p(x_1|x_0, u_0)p(x_0) \end{aligned}$$

That is,

$$\begin{aligned} p(x_0, x_1|z_1, u_0) &= \eta \frac{1}{\sqrt{|2\pi\Sigma_v|}} \exp\left(-\frac{1}{2}\|z_1 - h(x_1)\|_{\Sigma_v}^2\right) \\ &\quad \cdot \eta \frac{1}{\sqrt{|2\pi\Sigma_w|}} \exp\left(-\frac{1}{2}\|x_1 - f(x_0, u_0)\|_{\Sigma_w}^2\right) \\ &\quad \cdot \eta \frac{1}{\sqrt{|2\pi\Sigma_0|}} \exp\left(-\frac{1}{2}\|x_0 - \hat{x}_0\|_{\Sigma_w}^2\right) \end{aligned}$$

MAP for  $x$  is:

$$x^* = \arg_x \max \left\{ \eta \frac{1}{\sqrt{|2\pi\Sigma_v|}} \exp\left(-\frac{1}{2}\|z_1 - h(x_1)\|_{\Sigma_v}^2\right) \right\}$$

$$\cdot \eta \frac{1}{\sqrt{|2\pi\Sigma_w|}} \exp\left(-\frac{1}{2}\|x_1 - f(x_0, u_0)\|_{\Sigma_w}^2\right) \\ \cdot \eta \frac{1}{\sqrt{|2\pi\Sigma_0|}} \exp\left(-\frac{1}{2}\|x_0 - \hat{x}_0\|_{\Sigma_w}^2\right) \}$$

log func of this term, won't change the solution.

$$x^* = \arg_x \max \left\{ \log \left[ \eta \frac{1}{\sqrt{|2\pi\Sigma_v|}} \cdot \eta \frac{1}{\sqrt{|2\pi\Sigma_w|}} \cdot \eta \frac{1}{\sqrt{|2\pi\Sigma_0|}} \right] \right. \\ \left. \cdot \left(-\frac{1}{2}\|z_1 - h(x_1)\|_{\Sigma_v}^2\right) \cdot \left(-\frac{1}{2}\|x_1 - f(x_0, u_0)\|_{\Sigma_w}^2\right) \cdot \left(-\frac{1}{2}\|x_0 - \hat{x}_0\|_{\Sigma_w}^2\right) \right\}$$

extracting the  $-\frac{1}{2} \Rightarrow$  will change the argument to min:

$$x^* = \arg_x \min \left\{ \log \left[ \eta \frac{1}{\sqrt{|2\pi\Sigma_v|}} \cdot \eta \frac{1}{\sqrt{|2\pi\Sigma_w|}} \cdot \eta \frac{1}{\sqrt{|2\pi\Sigma_0|}} \right] \right. \\ \left. \cdot \left(\|z_1 - h(x_1)\|_{\Sigma_v}^2\right) \cdot \left(\|x_1 - f(x_0, u_0)\|_{\Sigma_w}^2\right) \cdot \left(\|x_0 - \hat{x}_0\|_{\Sigma_w}^2\right) \right\}$$

taking out all the determinants in the log function since they are independent of  $x$

$$x^* = \arg_x \min \left\{ \left(\|z_1 - h(x_1)\|_{\Sigma_v}^2\right) \cdot \left(\|x_1 - f(x_0, u_0)\|_{\Sigma_w}^2\right) \cdot \left(\|x_0 - \hat{x}_0\|_{\Sigma_w}^2\right) \right\}$$

now using this relation  $\|a\|_{\Sigma}^2 = \|\Sigma^{-\frac{1}{2}}a\|$  :

$$x^* = \arg_x \min \left\{ \left(\|\Sigma_v^{-\frac{1}{2}}(z_1 - h(x_1))\|\right) \cdot \left(\|\Sigma_w^{-\frac{1}{2}}(x_1 - f(x_0, u_0))\|\right) \cdot \left(\|\Sigma_w^{-\frac{1}{2}}(x_0 - \hat{x}_0)\|\right) \right\}$$

This problem, which involves minimizing the squared differences where the function  $h(x)$  is non-linear, can be effectively tackled using optimization techniques introduced in our classes, such as the Gauss-Newton or Levenberg-Marquardt algorithms.

**(d)**

seems to be like:

The marginal covariance  $\Sigma_1$  for  $x_1$  can be extracted directly from the appropriate block of the joint covariance matrix since the off-diagonal blocks do not influence it after marginalization.

$$\Sigma_1 = \Sigma_{11}$$

The marginal information matrix  $I_1$  for  $x_1$  requires subtracting the influence of  $x_0$  from the total information about  $x_1$  which is done by the right term.

$$I_1 = I_{11} - I_{01}^T I_{00}^{-1} I_{10}$$

## Hands-on Exercises

4.

\*this Q in matlab code.

(a)

Camera projection matrix is:

$$P = \begin{bmatrix} f_x & s & u_0 \\ 0 & f_y & v_0 \\ 0 & 0 & 1 \end{bmatrix} \cdot [R \mid t]$$

$[R \mid t]$  is the camera pose,  $s$  is the skew.

will write down 3D transformation matrix:

$$T_G^C = \begin{bmatrix} 0.5363 & -0.8440 & 0 & -451.2459 \\ 0.8440 & 0.5363 & 0 & 257.0322 \\ 0 & 0 & 1 & 400 \\ 0 & 0 & 0 & 1 \end{bmatrix}$$

given  $(u_0, v_0)$  and also  $f_x = f_y = 400$ .

assume  $s = 0$ .

$$P = \begin{bmatrix} 480 & 0 & 320 \\ 0 & 480 & 270 \\ 0 & 0 & 1 \end{bmatrix} \begin{bmatrix} 0.5363 & -0.8440 & 0 & -451.2459 \\ 0.8440 & 0.5363 & 0 & 257.0322 \\ 0 & 0 & 1 & 400 \end{bmatrix}$$

$$P = \begin{bmatrix} 257.424 & -405.12 & 320 & -8.8598e4 \\ 405.12 & 257.424 & 270 & 2.318e5 \\ 0 & 0 & 1 & 400 \end{bmatrix}$$

(b)

need to calculate:

$$\begin{pmatrix} \tilde{u} \\ \tilde{v} \\ \tilde{w} \end{pmatrix} = P \cdot X^G$$



$$\begin{pmatrix} \tilde{u} \\ \tilde{v} \\ \tilde{w} \end{pmatrix} = \begin{pmatrix} 9.158e4 \\ 2.994e5 \\ 365 \end{pmatrix}$$

$$\begin{pmatrix} u \\ v \end{pmatrix} = \begin{pmatrix} \tilde{u} \\ \tilde{v} \\ \tilde{w} \end{pmatrix} \cdot \frac{1}{\tilde{w}} = \begin{pmatrix} \tilde{u} \\ \tilde{v} \\ \tilde{w} \end{pmatrix} = \begin{pmatrix} 250.905 \\ 820.168 \end{pmatrix}$$

**(c)**

Re-projection error will be calculated by:

$$Err = z - \pi(x, l) = \begin{pmatrix} u \\ v \end{pmatrix}_{pix} - \begin{pmatrix} 250.905 \\ 820.168 \end{pmatrix} = \begin{pmatrix} -9.45 \\ -651.17 \end{pmatrix}$$

the total difference in pixels is: 651.236

\*seems quite big

**5.**

\*this Q in matlab code.

**(a)**

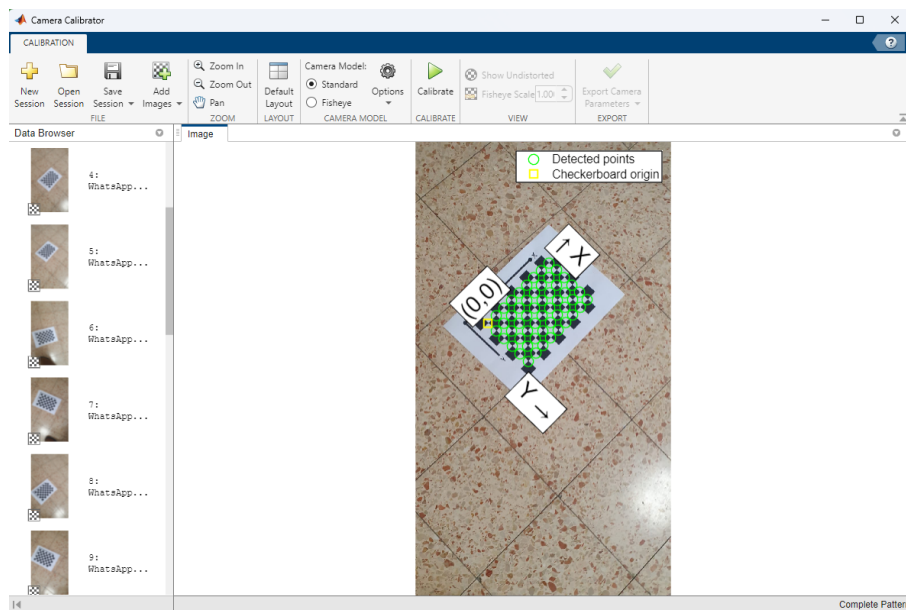
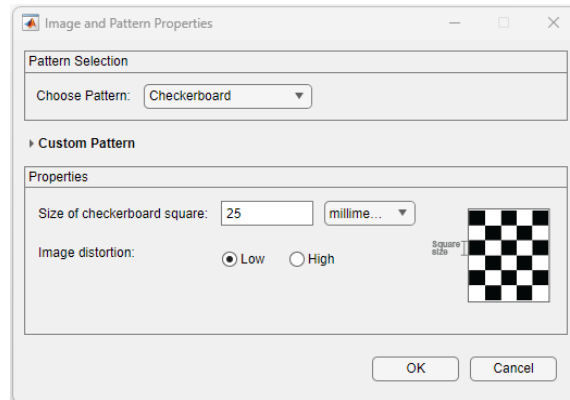
done.

**(b)**

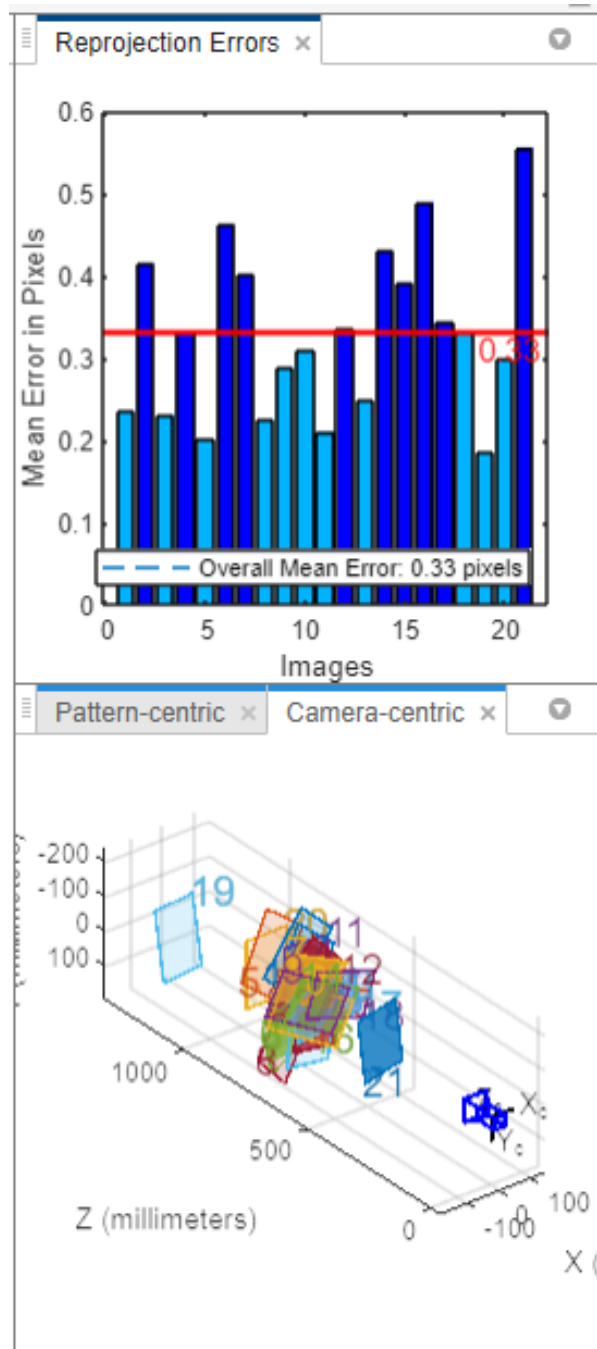
In total, 21 images of a checkerboard pattern  
using Samsung Galaxy s10+ camera.

**(c)**

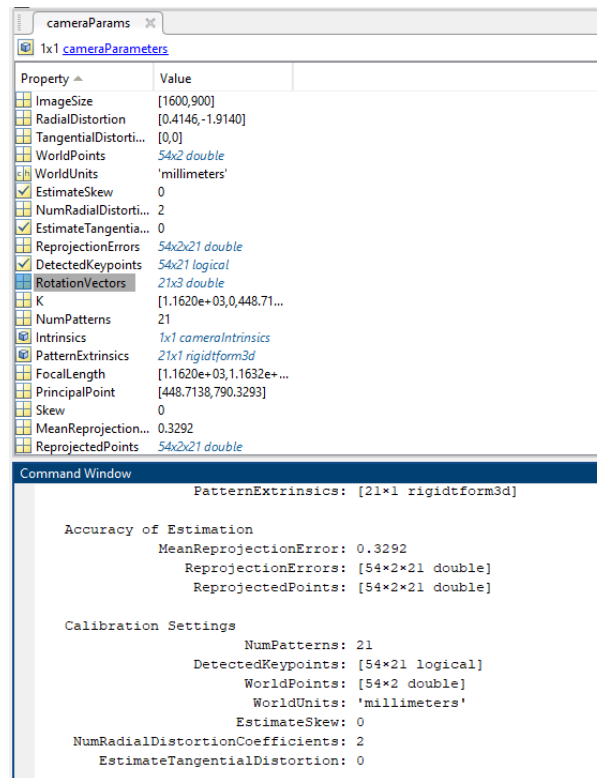
First uploading all the 21 images into Matlab Calibration GUI, verified that each rib is 25mm.:



Pressing Calibration button and derived all data:



the data derived:



(d)

As we saw before, the definition for K matrix:

$$K = \begin{bmatrix} f_x & s & u_0 \\ 0 & f_y & v_0 \\ 0 & 0 & 1 \end{bmatrix} = \begin{bmatrix} 1.162e3 & 0 & 448.7138 \\ 0 & 1.163e3 & 790.3293 \\ 0 & 0 & 1 \end{bmatrix}$$

The following is the principal point, in each axis:

$$\begin{bmatrix} 448.7138 \\ 790.3293 \end{bmatrix}$$

The following is the focal length, in each axis:

$$\begin{bmatrix} 1.162e3 \\ 1.163e3 \end{bmatrix}$$

6.

\*this Q in Python code.

(a)

done

(b)

done

**(c)**

SIFT features were extracted in both images

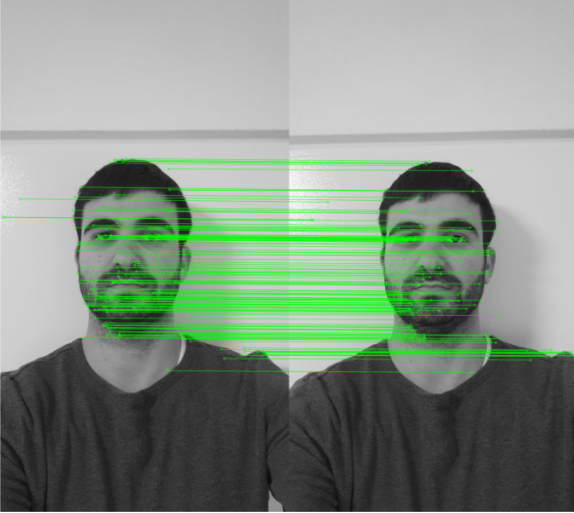


$$Scale = 1.94134, \text{ Orientation} = 3.718 \text{ rad}$$

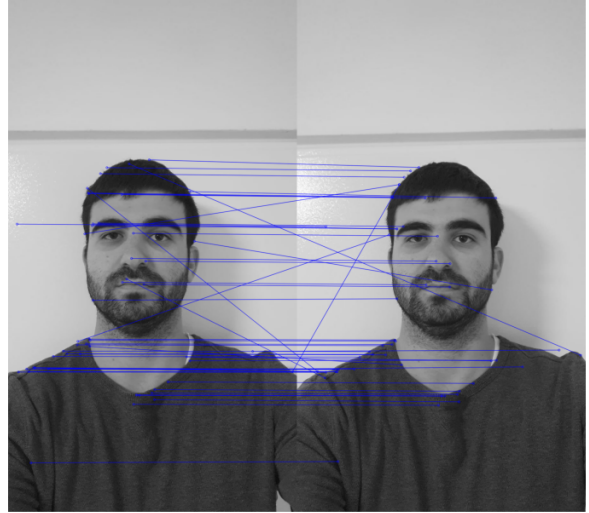
\*for the first keypoint that measured.

**(d)**

Inliers (Good Matches)



Outliers (Bad Matches)



\* with Threshold of 5.

