

# FDR control via Data Splitting

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## FDR control via Data Splitting

- Review of Multiple testing
- Motivation of Data Splitting
- Data Splitting

# Review of Multiple testing

- In High dimensional linear regression model, we explain variable selection procedure using the “FDR control via Data splitting” (C Dai, 2020).
- Before explaining the paper, we review the several multiple testing procedures

# Multiple Testing

- The simultaneous testing of more than one hypothesis
- Having observed a large number  $N$  of test statistics, how should we decide which if any of the null hypotheses to reject
- Assume that there are  $N$  hypotheses testing

$$H_{0i} \text{ vs } H_{1i}, \quad i = 1, 2, \dots, N$$

# Multiple Testing

		Decision		
		Null	Non-Null	
Actual	Null	$N_0 - a$	$a$	$N_0$
	Non-Null	$N_1 - b$	$b$	$N_1$
		$N - R$	$R$	$N$

- Decision rule  $\mathcal{D}$  has rejected  $R$  out of  $N$  null hypotheses. ( $N$ ,  $R$  is known)
- $a$  of these decisions were incorrect. (false discoveries)
- $N_0$ ,  $N_1$ ,  $a$ ,  $b$  are unknown random variables.

# Family-Wise Error Rate(FWER)

- The FWER criterion aims to control the probability of making even one false rejection among  $N$  simultaneous hypothesis tests.
- The FWER is

$$FWER = Pr\{\text{reject any true } H_{0i}\} = P(a \geq 1)$$

- There are several methods for control FWER at level  $\alpha$ . (Bonferroni's procedure, Holm's procedure)
- This means we control  $FWER = P(a \geq 1) \leq \alpha$ .
- But FWER usually proved too conservative when  $N$  is large.

# False Discovery Rates

- Define the false discovery proportion(Fdp)  $Fdp(\mathcal{D}) = a/R$ .
- Since Fdp is unobservable, we control the false discovery rate(fdr) at level  $q$  ( $0 < q < 1$ ) defined

$$FDR(\mathcal{D}) = E(Fdp(\mathcal{D})) \leq q$$

- How can we choose the decision rule  $\mathcal{D}$ ?
- The idea of finding the decision rule is to order the observed p-values from smallest to largest.
- Let  $p_i$  be the p-value corresponding the  $H_{0i}$  for all  $i = 1, 2, \dots, N$  and  $p_{(1)} \leq p_{(2)} \leq \dots \leq p_{(N)}$  be the ordering p-values.

# Benjamini–Hochberg FDR Control

## Theorem(Benjamini–Hochberg FDR Control, 1995)

Let  $H_{0(i)}$  be the null hypothesis corresponding the  $i$ th ordering p-value  $p_{(i)}$ . Define  $i_{\max} = \max\{i : p_{(i)} \leq \frac{i}{N}q\}$  and let  $\mathcal{D}_q$  be the rule that rejects  $H_{0(i)}$  for  $i \leq i_{\max}$ . If the p-values corresponding to valid null hypotheses are independent of each other, then

$$FDR(\mathcal{D}_q) = \pi_0 q \leq q; \quad \text{where } \pi_0 = N_0/N$$



# False Discovery Rates

- The Benjamin Hochberg (BHq) procedure is the most commonly used basic method of FDR control
- But the BHq procedure requires the assumption of independence for all the p-values.
- Benjamini and Yekutieli(2001) generalized BHq to handle dependent p-values by using a shrinkage of the control level  $\tilde{q} = \frac{q}{\sum_{j=1}^N 1/j}$ .
- There are many methods of the FDR control.
  - Sarkar(2002), Storey (2004) and so on.

## FDR control via Data Splitting

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# Notation and Definition

- Let  $\mathbf{X}_{n \times p} = (X_1, X_2, \dots, X_p)$  be the explanatory features with  $p$  being large.
- For each row of the design matrix  $\mathbf{X}$  independently follows a  $p$ -dimensional distribution with a covariance matrix  $\Sigma$ .
- For each feature has been normalized with zero mean and unit variance.
- Let  $Y = (y_1, \dots, y_n)$  be the vector of  $n$  independent response variable. Consider the linear regression model

$$Y = X_{n \times p} \beta + \epsilon$$

# Notation and Definition

- $S_0 = \{i : \beta_i = 0\}$  : the index set of the null features
- $S_1 = \{i : \beta_i \neq 0\}$  : the index set of the relevant features
- $p_0 = |S_0|$  and  $p_1 = |S_1|$
- $\hat{S}$  : the index set of the selected features (estimator of  $S_1$ )
- In this case, the hypotheses are

$$H_{0i} : \beta_i = 0 \text{ vs } H_{1i} : \beta_i \neq 0$$

# Motivation of Data Splitting

- In High-Dimensional linear regression model, we can select the significant variable by applying the FDR control procedures.

$$FDR = \mathbb{E}[FDP], \quad FDP = \frac{\#\{j : j \in S_0, j \in \hat{S}\}}{\#\{j : j \in \hat{S}\} \vee 1}$$

- But it is difficult to calculate p-values and estimate the joint distribution of features.
- For these reasons, there is a limit to applying the BHq procedure.
- To solve these problems, we use the data-splitting

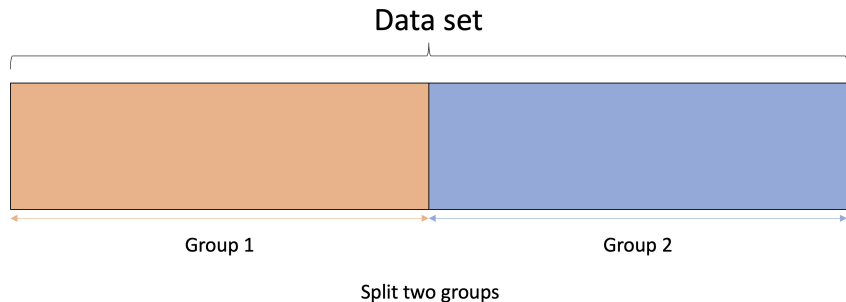
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# Single Data Splitting(DS)

- Features selection depend on  $\hat{\beta}$ . For example,  $\hat{\beta}$  can be estimated via OLS or some shrinkage methods.
- In contrast to those commonly methods, we split the data into two groups of equal size denoted as  $(\mathbf{y}^{(1)}, \mathbf{X}^{(1)})$ , and  $(\mathbf{y}^{(2)}, \mathbf{X}^{(2)})$ .
- So we can estimate two independent regression coefficients denoted  $\hat{\beta}^{(1)}, \hat{\beta}^{(2)}$  for each groups.
- To achieve FDR control under our data-splitting, two independent coefficients should satisfy the following assumption.

# Single Data Splitting(DS)





## Assumption 1 (Symmetry)

For each null feature index  $j \in S_0$ , the sampling distribution of either  $\hat{\beta}_j^{(1)}$  or  $\hat{\beta}_j^{(2)}$  is symmetric about 0

- For  $j \in S_0$ , only one of  $\hat{\beta}_j^{(1)}$  and  $\hat{\beta}_j^{(2)}$  is symmetric about 0.
- Define the mirror statistics  $M_j$  by

$$M_j = \text{sign}(\hat{\beta}_j^{(1)} \hat{\beta}_j^{(2)}) f(|\hat{\beta}_j^{(1)}|, |\hat{\beta}_j^{(2)}|)$$

where the function  $f(u, v)$  is non-negative, symmetric about  $u$  and  $v$ , and monotone increasing in both  $u$  and  $v$ .

- For a relevant feature, the corresponding mirror statistic tends to be positive and relatively large.

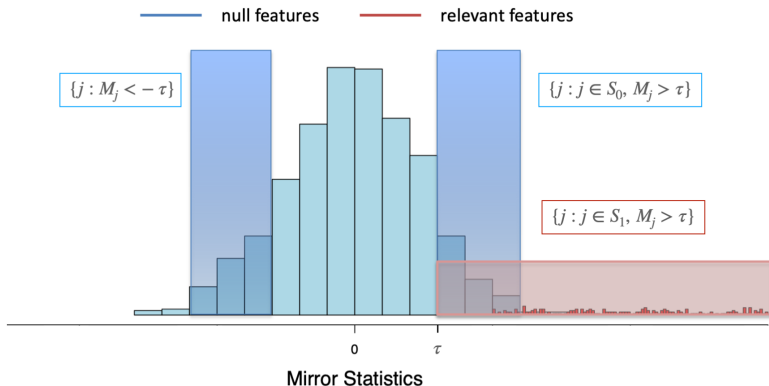
## Lemma 1

Under Assumption 1, regardless of the data-splitting procedures, the sampling distribution of  $M_j$  is symmetric about 0 for  $j \in S_0$

- We use mirror statistics as test statistics.
- The mirror statistics satisfy the following two properties.
  - (A1) A feature with a larger mirror statistic is more likely to be a relevant feature.
  - (A2) The sampling distribution of the mirror statistic of any null feature is symmetric about 0.

# Mirror Statistics

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# Mirror Statistics

- There are several choice of  $f(u, v)$ . The optimal choice that obtains the highest power is  $f(u, v) = u + v$ .
- By using the Assumption 1, we can make an upper bound of the number of false positives

$$\#\{j \in S_0 : M_j > t\} \approx \#\{j \in S_0 : M_j < -t\} \leq \#\{j : M_j < -t\}, \forall t > 0.$$

- For given  $t > 0$ , we can define the FDP(t) of the selection  $\hat{S}_t = \{j : M_j > t\}$  and the estimate of the FDP(t) as  $\widehat{FDP}(t)$ .

$$FDP(t) = \frac{\#\{j : M_j > t, j \in S_0\}}{\#\{j : M_j > t\} \vee 1}, \quad \widehat{FDP}(t) = \frac{\#\{j : M_j < -t\}}{\#\{j : M_j > t\} \vee 1}$$

- Let  $\forall q \in (0, 1)$  be given FDR control level.
- Then, we can find the cutoff value  $\tau_q$  as follows

$$\tau_q = \min\{t > 0 : \widehat{FDP}(t) \leq q\}$$

- Thus, we finally select  $\hat{S}_{\tau_q} = \{j : M_j > \tau_q\}$  as a set of the index of relevant features.

# Algorithm of FDR control via Single DS

## Algorithm 1

1. Split the data into two groups, independent to the response vector  $y$ .
2. Estimate the "impact" coefficient  $\hat{\beta}^{(1)}$  and  $\hat{\beta}^{(2)}$  on each part of the data. The two estimation procedures can be potentially different.
3. Calculate the mirror statistics following

$$M_j = \text{sign}(\hat{\beta}_j^{(1)} \hat{\beta}_j^{(2)}) (|\hat{\beta}_j^{(1)}| + |\hat{\beta}_j^{(2)}|)$$

4. Given a designated FDR level  $q \in (0, 1)$ , calculate the cutoff  $\tau_q$  as :

$$\tau_q = \min\{t > 0 : \widehat{FDP}(t) = \frac{\#\{j : M_j < -t\}}{\#\{j : M_j > t\} \vee 1} \leq q\}$$

5. Select the features  $\{j : M_j > \tau_q\}$

# Single Data Splitting

- To obtain a good estimate of the number of false positives, the mirror statistics of the null features cannot be too correlated. So we require the following weak dependence assumption.

## Assumption 2 (Weak dependence among the null features)

The mirror statistics  $M_j$ 's are continuous random variables, and there exist constants  $c > 0$  and  $\alpha \in (0, 2)$  such that

$$\text{Var} \left( \sum_{j \in S_0} 1(M_j > t) \right) \leq c p_0^\alpha, \quad \forall t \in R, \text{ where } p_0 = |S_0|$$

- Additionally, we assume the variances of the mirror statistics are bounded.

## Proposition

For any designated FDR control level  $q \in (0, 1)$ , assume that there exists a constant  $t_q > 0$  such that  $\mathbb{P}(FDP(t_q) \leq q) \rightarrow 1$  as  $p \rightarrow \infty$ . Then, under Assumption 1 and 2, the procedure in Algorithm 1 satisfies

$$FDP(\tau_q) \leq q + o_p(1) \text{ and } \limsup_{p \rightarrow \infty} FDR(\tau_q) \leq q$$

- We note that the existence of  $t_q > 0$  such that  $\mathbb{P}(FDP(t_q) \leq q) \rightarrow 1$  essentially guarantees the asymptotic feasibility of FDR control based upon the rankings of features by their mirror statistics.



# Multiple Data Splitting(MDS)

- There are two problems about DS.
- First, splitting the data into two halves inflates the variance of the estimated regression coefficient. So, DS can potentially suffer a power loss.
- Second, the selection result of DS may not be stable and can vary substantially across different sample splits.
- To solve this problem, we use a multiple data splitting procedure to aggregate the selection results obtained from independent replications of DS.

# Multiple Data Splitting(MDS)

- Suppose we independently repeat DS  $m$  times with random sample splits.
- Each time the set of selected features is denoted as  $\hat{S}^{(k)}$  for  $k = 1, 2, \dots, m$ .
- For each feature  $X_j$ , we define the associated inclusion rate  $I_j$  and its estimate  $\hat{I}_j$  as

$$I_j = \mathbb{E} \left[ \frac{1(j \in \hat{S})}{|\hat{S}| \vee 1} | X, y \right], \quad \hat{I}_j = \frac{1}{m} \sum_{k=1}^m \frac{1(j \in \hat{S}^{(k)})}{|\hat{S}^{(k)}| \vee 1}$$

# Multiple Data Splitting(MDS)

- This rate is an importance measurement of each feature relative to the DS selection procedure.
- MDS is most useful if the following statement is approximately true.
  - If a feature is selected less frequently in the repeated sample splitting, it is less likely to be a relevant feature.
- If this holds, we can choose a proper inclusion rate cutoff to control the FDR, and select those features with inclusion rates larger than cutoff.

# Algorithm of aggregating selection results from multiple data splits

## Algorithm 2

1. Sort the estimated inclusion rates :  $0 \leq \hat{l}_{(1)} \leq \hat{l}_{(2)} \leq \dots \leq \hat{l}_{(p)}$
2. Find the largest  $l \in \{1, 2, \dots, p\}$  such that  $\hat{l}_{(1)} + \hat{l}_{(2)} + \dots + \hat{l}_{(l)} \leq q$
3. Select the features  $\hat{S} = \{j : \hat{l}_j > \hat{l}_{(l)}\}$ 
  - The following proposition points out a key factor for MDS to achieve FDR control

# Multiple Data Splitting(MDS)

## Proposition

Suppose we can asymptotically control the FDP of DS for any designated level  $q \in (0, 1)$ . Furthermore, we assume that with probability approaching 1, the power of DS is bounded below by some  $\kappa > 0$ . We consider the following two regimes with  $n, p \rightarrow \infty$  at a proper rate.

- (a) In the non-sparse regime where  $\liminf p_1/p > 0$ , we assume that the mirror statistics are consistent at ranking features, i.e.,  
$$\sup_{i \in S_1, j \in S_0} P(I_i < I_j) \rightarrow 0.$$
- (b) In the sparse regime where  $\limsup p_1/p = 0$ , we assume that the mirror statistics are strongly consistent at ranking features, i.e.,  
$$\sup_{i \in S_1} P(I_i < \max_{j \in S_0} I_j) \rightarrow 0.$$

Then, for MDS (see Algorithm 2) in both the non-sparse and the sparse regimes, we have

$$FDP \leq q + o_p(1) \quad \text{and} \quad \limsup_{n, p \rightarrow \infty} FDR \leq q$$

# Application for Linear models

- Consider the process of above method applying to linear models.
- We proceed the data splitting by using a Lasso + OLS procedure.
- In detail, on the first half of the data  $(\mathbf{y}^{(1)}, \mathbf{X}^{(1)})$ , we apply Lasso for dimension reduction. Let  $\hat{\beta}^{(1)}$  be the estimated regression coefficients and denotes  $\hat{S}^{(1)} = \{j : \hat{\beta}_j^{(1)} \neq 0\}$ .
- Next, we restrict the features to  $\hat{S}^{(1)}$  obtained above. Then, we run OLS using the second half of the data  $(\mathbf{y}^{(2)}, \mathbf{X}^{(2)})$  to obtain the estimated coefficients  $\hat{\beta}^{(2)}$ .
- So, we can make the mirror statistics by using  $\hat{\beta}^{(1)}$  and  $\hat{\beta}^{(2)}$ .

- The main advantage of DS and MDS is that they do not require the p-values and the joint distribution of features.
- MDS stabilizes the result and improves the power of a single DS.
- DS and MDS control the FDR at the designated level in linear models.
- Both DS and MDS are conceptually simple and easy to implement based on existing softwares for high-dimensional regression methods.

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