# Summary of SI

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# 1 Introduction

# 1.1 Comparison of two mathematical modeling methods

- i. Mechanism modeling is to write down a complete set of equations relating the different variables in the system based on physical laws.
- ii. System identification (SI) is to find out the relationship between input and output, as well as other signals, from a set of experimental data.

### 1.2 Definition of SI

System identification can be defined as the determination of a mathematic model from the observed input and output by minimizing some error criterion function.

### 1.3 Four entities of SI

Data, set of models, criterion, optimization approaches.

# 2 Identification models and LS estimation

#### 2.1 Identification models

Shift operator

- i. Forward z: zu(t) = u(t+1)
- ii. Backward  $z^{-1}: z^{-1}u(t) = u(t-1)$

### 2.1.1 Time series models

Denote

$$A(z^{-1}) = 1 + \sum_{i=1}^{n_a} a_i z^{-i}$$

$$B(z^{-1}) = \sum_{i=1}^{n_b} b_i z^{-i}$$

$$C(z^{-1}) = 1 + \sum_{i=1}^{n_c} c_i z^{-i}$$

$$D(z^{-1}) = 1 + \sum_{i=1}^{n_d} d_i z^{-i}$$

i. Autoregressive (AR) model

$$A(z^{-1})y(t) = v(t)$$

ii. Moving average (MA) model

$$y(t) = D(z^{-1})v(t)$$

iii. Autoregressive moving average (ARMA) model

$$A(z^{-1})y(t) = D(z^{-1})v(t)$$

#### 2.1.2 Equation error type models

$$A(z^{-1})y(t) = B(z^{-1})u(t) + w(t)$$

$$w(t) = \begin{cases} v(t) & \text{ARX} \\ D(z^{-1})v(t) & \text{ARMAX} \\ \frac{1}{C(z^{-1})}v(t) & \text{ARARX} \\ \frac{D(z^{-1})}{C(z^{-1})}v(t) & \text{ARARMAX} \end{cases}$$

where AR refers to the autoregressive part  $A(z^{-1})y(t)$  and X refers to the extra input  $B(z^{-1})u(t)$ .

#### 2.1.3 Output error type models

$$y(t) = \frac{B(z^{-1})}{A(z^{-1})}u(t) + w(t)$$

$$w(t) = \begin{cases} v(t) & \text{OE} \\ D(z^{-1})v(t) & \text{OEMA} \end{cases}$$
 
$$\frac{1}{C(z^{-1})}v(t) & \text{OEAR}$$
 
$$\frac{D(z^{-1})}{C(z^{-1})}v(t) & \text{OEARMA} (B-J)$$

where  $\frac{B(z^{-1})}{A(z^{-1})} := x(t)$  can be viewed as the true output.

### 2.2 LS principle

#### 2.2.1 One-dimensional case

**Question 1.** For a desk, we assume that n persons obtain n different values of the length of this desk  $x_1, x_2, ..., x_n$ . Now our question is, what is the most likely length of this desk?

**Answer 1.** Intuitively, if x is the length of this desk, it should minimize the following error function

$$f(x) = \sum_{i=1}^{n} (x_i - x)^2,$$

which is the sum of squares of the error. By setting the derivation to zero, one has

$$\frac{df(x)}{dx} = -2\sum_{i=1}^{n} (x_i - x) = 0,$$

which gives

$$x = \frac{x_1 + x_2 + \dots + x_n}{n}.$$

In addition, it is easily obtained that

$$\frac{df(x)}{dx} = 2n > 0.$$

Therefore,  $f(\frac{x_1+x_2+\cdots+x_n}{n})$  is the minimum value of the error function f(x). This is the basic principle of the method of least squares.

### 2.3 Statistical properties for LS estimation

**Theorem 2.1.** Consider the system (\*), and suppose that  $V_t$  is of mean zero, and  $V_t$  and  $H_t$  are statistically independent. Then the least squares estimate  $\hat{\theta}_{LS} = (H_t^T H_t)^{-1} H_t^T Y_t$  is an **unbiased** estimate, that is,  $\mathbb{E}[\hat{\theta}_{LS}] = \theta$ .

**Proof.** By using  $Y_t = H_t\theta + V_t$ , one has

$$\begin{split} \hat{\theta}_{\text{LS}} &= (H_t^T H_t)^{-1} H_t^T Y_t \\ &= (H_t^T H_t)^{-1} H_t^T (H_t \theta + V_t) \\ &= (H_t^T H_t)^{-1} H_t^T H_t \theta + (H_t^T H_t)^{-1} H_t^T V_t \\ &= \theta + (H_t^T H_t)^{-1} H_t^T V_t \end{split}$$

then

$$\mathbb{E}[\hat{\theta}_{LS}] = \theta + \mathbb{E}[(H_t^T H_t)^{-1} H_t^T V_t] = \theta$$

**Theorem 2.2.** Consider the system (\*), and suppose that  $V_t$  is of mean zero and has the covariance matrix  $\operatorname{cov}[V_t] = R_v$ . In addition, it is assumed that  $V_t$  and  $H_t$  are statistically independent. Then the covariance matrix of the estimation error  $\tilde{\theta}_{LS}(t) = \hat{\theta}_{LS}(t) - \theta$  is given by

$$\operatorname{cov} \left[ \tilde{\theta}_{\mathrm{LS}}(t) \right] = \mathbb{E} \left[ (H_t^T H_t)^{-1} H_t^T R_v H_t (H_t^T H_t)^{-1} \right]$$

**Proof.** It follows from the proof of the previous theorem that

$$\tilde{\theta}_{\mathrm{LS}} = \hat{\theta}_{\mathrm{LS}}(t) - \theta = (H_t^T H_t)^{-1} H_t^T V_t$$

and

$$\mathbb{E} \big[ \, \hat{\theta}_{\mathrm{LS}} \, \big] = \mathbb{E} \big[ \, \hat{\theta}_{\mathrm{LS}}(t) - \theta \, \big] = 0$$

then

$$\begin{aligned} \operatorname{cov} \left[ \tilde{\theta}_{\mathrm{LS}} \right] &= \mathbb{E} \left[ \tilde{\theta}_{\mathrm{LS}} \tilde{\theta}_{\mathrm{LS}}^{T} \right] \\ &= \mathbb{E} \left[ (H_{t}^{T} H_{t})^{-1} H_{t}^{T} V_{t} V_{t}^{T} H_{t} (H_{t}^{T} H_{t})^{-1} \right] \\ &= \mathbb{E} \left[ (H_{t}^{T} H_{t})^{-1} H_{t}^{T} R_{v} H_{t} (H_{t}^{T} H_{t})^{-1} \right] \end{aligned}$$

If  $\mathbb{E}[v(t)] = 0$  and  $E[v^2(t)] = \sigma^2$ , one has  $\text{cov}[V_t] = \sigma^2 I_t$ ,  $\text{cov}[\tilde{\theta}_{LS}(t)] = \sigma^2 \mathbb{E}[(H_t^T H_t)^{-1}]$ .

**Theorem 2.3.** An unbiased estimate of  $\sigma^2$  can be given by

$$\hat{\sigma}^2 = \frac{J(\hat{\theta}_{LS}(t))}{t - \dim \theta}$$

# 3 RLS

### 3.1 Basic RLS

# 3.1.1 P(0)

When the initial value P(0) takes  $p_0I > 0$ , it follows that

$$P(t) = \left[\frac{1}{p_0}I + \sum_{i=1}^t \varphi(i)\varphi^T(i)\right]^{-1}$$
$$= \left[\frac{1}{p_0}I + H_t^T H_t\right]^{-1}$$

where

$$H_t = [\varphi(1) \ \varphi(2) \ \dots \ \varphi(t)]^T$$

This implies that

$$\lim_{p_0 \to \infty} P(t) = \lim_{p_0 \to \infty} \left[ \frac{1}{p_0} I + H_t^T H_t \right]^{-1}$$
$$= [H_t^T H_t]^{-1}$$
$$= P(t)$$

This fact shows that the  $p_0$  should be chosen to be as large as possible.

# 3.2 Forgetting factor RLS algorithm

$$J(\theta) = \sum_{i=1}^{t} \lambda^{t-i} [y(i) - \varphi^{T}(i)\theta]^{2}$$
$$= (Y_{t} - H_{t}\theta)^{T} \Lambda_{t} (Y_{t} - H_{t}\theta)$$

$$\begin{array}{lll} \lambda & = & \rho^2 \\ H_t & = & \left[ \begin{array}{cccc} \varphi(1) & \varphi(2) & \dots & \varphi(t-1) & \varphi(t) \end{array} \right]^T \\ Y_t & = & \left[ \begin{array}{cccc} y(1) & y(2) & \dots & y(t-1) & y(t) \end{array} \right]^T \\ \Lambda_t & = & \operatorname{diag}\{\lambda^{t-1}, \lambda^{t-2}, \dots, \lambda, 1\} \end{array}$$

Take the derivation,

$$\frac{\partial J(\theta)}{\partial \theta} = \frac{\partial V_t}{\partial \theta} \frac{\partial J(\theta)}{\partial V_t}$$

$$= \frac{\partial (Y_t - H_t \theta)}{\partial \theta} \frac{\partial (V_t^T \Lambda_t V_t)}{\partial V_t}$$

$$= -2H_t^T \Lambda_t V_t$$

$$= -2H_t^T \Lambda_t (Y_t - H_t \theta)$$

$$= 0$$

which gives

$$\hat{\theta}(t) = (H_t^T \Lambda_t H_t)^{-1} H_t^T \Lambda_t Y_t$$

Denote  $P(t) = (H_t^T \Lambda_t H_t)^{-1}$ , then

$$\begin{split} P^{-1}(t) &= \sum_{i=1}^{t} \lambda^{t-i} \varphi(i) \varphi^T(i) \\ &= \lambda \sum_{i=1}^{t-1} \lambda^{t-i-1} \varphi(i) \varphi^T(i) + \varphi(t) \varphi^T(t) \\ &= \lambda P^{-1}(t-1) + \varphi(t) \varphi^T(t) \end{split}$$

By the matrix inversion lemma, one has

$$\begin{split} P(t) \; &= \; \frac{1}{\lambda} \bigg( P(t-1) - \frac{P(t-1)\varphi(t)\varphi^T(t)P(t-1)}{\lambda + \varphi^T(t)P(t-1)\varphi(t)} \bigg) \\ H_t^T \Lambda_t Y_t \; &= \; \sum_{i=1}^t \lambda^{t-i}\varphi(i)y(i) \\ &= \; \lambda \sum_{i=1}^{t-1} \lambda^{t-i-1}\varphi(i)y(i) + \varphi(t)y(t) \\ &= \; \lambda H_{t-1}^T \Lambda_{t-1} Y_{t-1} + \varphi(t)y(t) \\ \hat{\theta}(t) \; &= \; (H_t^T \Lambda_t H_t)^{-1} H_t^T \Lambda_t Y_t \\ &= \; P(t)P(t-1)^{-1} P(t-1)(\lambda H_{t-1}^T \Lambda_{t-1} Y_{t-1} + \varphi(t)y(t)) \\ &= \; \lambda P(t)P(t-1)^{-1} \hat{\theta}(t-1) + P(t)\varphi(t)y(t) \\ &= \; [I-P(t)\varphi(t)\varphi^T(t)] \hat{\theta}(t-1) + P(t)\varphi(t)y(t) \\ &= \; \hat{\theta}(t-1) + P(t)\varphi(t)[y(t)-\varphi^T(t)\hat{\theta}(t-1)] \end{split}$$

$$L(t) = P(t)\varphi(t) = \frac{P(t-1)\varphi(t)}{\lambda + \varphi^{T}(t)P(t-1)\varphi(t)}$$
  
$$P(t) = \frac{1}{\lambda}(I - L(t)\varphi^{T}(t))P(t-1)$$

# 3.3 Fixed memory identification

$$J(\theta) = \sum_{i=t-p+1}^{t} [y(i) - \varphi^{T}(i)\theta]^{2}$$
  
=  $(Y_{p,t} - H_{p,t}\theta)^{T} (Y_{p,t} - H_{p,t}\theta)$ 

then

$$\hat{\theta}(t) = (H_{p,t}^T H_{p,t})^{-1} H_{p,t}^T Y_{p,t}$$

where

$$H_{p,t} = [\varphi(t-p+1) \ \varphi(t-p+2) \ \dots \ \varphi(t)]^T$$
  
 $Y_{p,t} = [y(t-p+1) \ y(t-p+2) \ \dots \ y(t)]^T$ 

Denote  $P(t) = (H_{p,t}^T H_{p,t})^{-1}$ , thus

$$P^{-1}(t) = \sum_{i=t-p+1}^{t} \varphi(i)\varphi^{T}(i)$$

$$= \sum_{i=t-p}^{t-1} \varphi(i)\varphi^{T}(i) + \varphi(t)\varphi^{T}(t) - \varphi(t-p)\varphi^{T}(t-p)$$

$$= P^{-1}(t-1) + \varphi(t)\varphi^{T}(t) - \varphi(t-p)\varphi^{T}(t-p)$$

$$\begin{split} H_{p,t}^{T}Y_{p,t} &= \sum_{i=t-p+1}^{t} \varphi(i)y(i) \\ &= \sum_{i=t-p}^{t-1} \varphi(i)y(i) + \varphi(t)y(t) - \varphi(t-p)y(t-p) \\ &= H_{p,t-1}^{T}Y_{p,t-1} + \varphi(t)y(t) - \varphi(t-p)y(t-p) \end{split}$$

so

$$\begin{split} \hat{\theta}(t) &= P(t)P^{-1}(t-1)P(t-1)[H_{p,t-1}^{T}Y_{p,t-1} + \varphi(t)y(t) - \varphi(t-p)y(t-p)] \\ &= P(t)P^{-1}(t-1)\hat{\theta}(t-1) + P(t)[\varphi(t)y(t) - \varphi(t-p)y(t-p)] \\ &= [I - P(t)\varphi(t)\varphi^{T}(t) + P(t)\varphi(t-p)\varphi^{T}(t-p)]\hat{\theta}(t-1) + P(t)[\varphi(t)y(t) - \varphi(t-p)y(t-p)] \\ &= \hat{\theta}(t-1) + P(t)\varphi(t)[y(t) - \varphi^{T}(t)\hat{\theta}(t-1)] - P(t)\varphi(t-p)[y(t-p) - \varphi^{T}(t-p)\hat{\theta}(t-1)] \\ &= \hat{\theta}(t-1) + P(t)[\varphi(t) - \varphi(t-p)] \begin{bmatrix} y(t) - \varphi^{T}(t)\hat{\theta}(t-1) \\ y(t-p) - \varphi^{T}(t-p)\hat{\theta}(t-1) \end{bmatrix} \end{split}$$

# 3.4 Fixed memory identification with a forgetting factor

$$J(\theta) = \sum_{i=t-p+1}^{t} \lambda^{t-i} [y(i) - \varphi^{T}(i)\theta]^{2}$$
$$= (Y_{p,t} - H_{p,t}\theta)^{T} \Lambda (Y_{p,t} - H_{p,t}\theta)$$

then

$$\hat{\theta}(t) = (H_{p,t}^T \Lambda H_{p,t})^{-1} H_{p,t}^T \Lambda Y_{p,t}$$

$$\begin{array}{lll} H_{p,t} & = & [ \ \varphi(t-p+1) \ \ \varphi(t-p+2) \ \ \dots \ \ \varphi(t) \ ]^T \\ Y_{p,t} & = & [ \ y(t-p+1) \ \ y(t-p+2) \ \ \dots \ \ y(t) \ ]^T \\ \Lambda & = & \mathrm{diag}\{\lambda^{p-1},\lambda^{p-2},...,\lambda,1\} \end{array}$$

Denote  $P(t) = (H_{p,t}^T \Lambda H_{p,t})^{-1}$ , thus

$$\begin{split} P^{-1}(t) &= \sum_{i=t-p+1}^{t} \lambda^{t-i} \varphi(i) \varphi^T(i) \\ &= \sum_{i=t-p}^{t-1} \lambda^{t-i} \varphi(i) \varphi^T(i) + \varphi(t) \varphi^T(t) - \lambda^p \varphi(t-p) \varphi^T(t-p) \\ &= P^{-1}(t-1) + \varphi(t) \varphi^T(t) - \lambda^p \varphi(t-p) \varphi^T(t-p) \end{split}$$

$$\begin{split} H_{p,t}^T \Lambda Y_{p,t} &= \sum_{i=t-p+1}^t \lambda^{t-i} \varphi(i) y(i) \\ &= \sum_{i=t-p}^{t-1} \lambda^{t-i} \varphi(i) y(i) + \varphi(t) y(t) - \lambda^p \varphi(t-p) y(t-p) \\ &= H_{p,t-1}^T \Lambda Y_{p,t-1} + \varphi(t) y(t) - \lambda^p \varphi(t-p) y(t-p) \end{split}$$

so

$$\begin{split} \hat{\theta}(t) &= P(t)P^{-1}(t-1)P(t-1)[H_{p,t-1}^T\Lambda Y_{p,t-1} + \varphi(t)y(t) - \lambda^p \varphi(t-p)y(t-p)] \\ &= P(t)P^{-1}(t-1)\hat{\theta}(t-1) + P(t)[\varphi(t)y(t) - \lambda^p \varphi(t-p)y(t-p)] \\ &= [I-P(t)\varphi(t)\varphi^T(t) + \lambda^p P(t)\varphi(t-p)\varphi^T(t-p)]\hat{\theta}(t-1) + P(t)[\varphi(t)y(t) - \lambda^p \varphi(t-p)y(t-p)] \\ &= \hat{\theta}(t-1) + P(t)\{\varphi(t)[y(t) - \varphi^T(t)\hat{\theta}(t-1)] - \lambda^p \varphi(t-p)[y(t-p) - \varphi^T(t-p)\hat{\theta}(t-1)]\} \\ &= \hat{\theta}(t-1) + P(t)[\ \varphi(t) \ -\lambda^p \varphi(t-p)\ ] \begin{bmatrix} y(t) - \varphi^T(t)\hat{\theta}(t-1) \\ y(t-p) - \varphi^T(t-p)\hat{\theta}(t-1) \end{bmatrix} \end{split}$$

# 4 RLS with $\lambda(t)$

### 5 GERLS

### 5.1 Filtering based recursive generalized LS algorithm

Consider the following ARARX model

$$A(z^{-1})y(t) = B(z^{-1})u(t) + \frac{1}{C(z^{-1})}v(t)$$
(5.1)

Define the filtered input  $u_f(t) := C(z^{-1})u(t)$  and filtered output  $y_f(t) = C(z^{-1})y(t)$ . Multiplying the both sides of (5.1) by  $C(z^{-1})$ , yields

$$A(z^{-1})y_f(t) = B(z^{-1})u_f(t) + v(t)$$

For the above equation error type model, one has

$$y_f(t) = \varphi_f^T(t)\theta_s + v(t) \tag{5.2}$$

$$\varphi_f(t) = \begin{bmatrix} -y_f(t-1) & -y_f(t-2) & \dots & -y_f(t-n_a) & u_f(t-1) & u_f(t-2) & \dots & u_f(t-n_b) \end{bmatrix}^T$$
 
$$\theta_s = \begin{bmatrix} a_1 & a_2 & \dots & a_{n_a} & b_1 & b_2 & \dots & b_{n_b} \end{bmatrix}^T$$

Next, we will construct another algorithm to estimate  $C(z^{-1})$ . Define an intermediate variable

$$w(t) = \frac{1}{C(z^{-1})}v(t) \tag{5.3}$$

which gives

$$C(z^{-1})w(t) = v(t)$$

By this relation, it can be obtained

$$w(t) = \varphi_n^T(t)\theta_n + v(t) \tag{5.4}$$

where

$$\begin{aligned} \varphi_n(t) &= \begin{bmatrix} -\omega(t-1) & -\omega(t-2) & \dots & -\omega(t-n_c) \end{bmatrix}^T \\ \theta_n &= \begin{bmatrix} c_1 & c_2 & \dots & c_{n_c} \end{bmatrix}^T \end{aligned}$$

It follows from (5.1) and (5.3) that

$$w(t) = A(z^{-1})y(t) - B(z^{-1})u(t) = y(t) - \varphi_s^T \theta_s$$

where  $\varphi_s(t) = \begin{bmatrix} -y(t-1) & -y(t-2) & \dots & -y(t-n_a) & u(t-1) & u(t-2) & \dots & u(t-n_b) \end{bmatrix}^T$ .

Then we can obtain an estimate of w(t) as

$$\hat{w}(t) = y(t) - \hat{\varphi}_s^T(t)\hat{\theta}_s(t-1)$$

According to (5.2) and (5.4), two least square algorithms can be construct

$$\begin{cases}
\hat{\theta}_{s}(t) = \hat{\theta}_{s}(t-1) + L_{s}(t) \left[ \hat{y}_{f}(t) - \hat{\varphi}_{f}^{T}(t) \hat{\theta}_{s}(t) \right] \\
L_{s}(t) = \frac{P_{s}(t-1)\hat{\varphi}_{f}(t)}{1 + \hat{\varphi}_{f}^{T}(t)P_{s}(t-1)\hat{\varphi}_{f}(t)} \\
P_{s}(t) = \left[ I - L_{n}(t) \hat{\varphi}_{s}^{T}(t) \right] P_{s}(t-1), P_{s}(0) = p_{0}I
\end{cases}$$
(5.5)

$$\begin{cases}
\hat{\theta}_{n}(t) = \hat{\theta}_{n}(t-1) + L_{n}(t) [\hat{w}(t) - \hat{\varphi}_{n}^{T}(t)\hat{\theta}_{n}(t)] \\
L_{n}(t) = \frac{P_{n}(t-1)\hat{\varphi}_{n}(t)}{1 + \hat{\varphi}_{n}^{T}(t)P_{n}(t-1)\hat{\varphi}_{n}(t)} \\
P_{n}(t) = [I - L_{n}(t)\hat{\varphi}_{n}^{T}(t)]P_{n}(t-1), P_{n}(0) = p_{0}I
\end{cases} (5.6)$$

$$\begin{array}{lll} \hat{\theta}_n &=& [ \ \hat{c}_1(t) \ \ \hat{c}_2(t) \ \dots \ \hat{c}_{n_c}(t) \ ]^T \\ \hat{\theta}_s &=& [ \ \hat{a}_1(t) \ \ \hat{a}_2(t) \ \dots \ \hat{a}_{n_a}(t) \ \ \hat{b}_1(t) \ \ \hat{b}_2(t) \ \dots \ \hat{b}_{n_b}(t) \ ]^T \\ \hat{\varphi}_n(t) &=& [ \ -\hat{\omega}(t-1) \ \ -\hat{\omega}(t-2) \ \dots \ \ -\hat{\omega}(t-n_c) \ ]^T \\ \varphi_s(t) &=& [ \ -y(t-1) \ \ -y(t-2) \ \dots \ \ -y(t-n_a) \ \ u(t-1) \ \ u(t-2) \ \dots \ \ u(t-n_b) \ ]^T \\ \hat{\varphi}_f(t) &=& [ \ -\hat{y}_f(t-1) \ \ -\hat{y}_f(t-2) \ \dots \ \ -\hat{y}_f(t-n_a) \ \ \hat{u}_f(t-1) \ \ \hat{u}_f(t-2) \ \dots \ \ \hat{u}_f(t-n_b) \ ]^T \\ \hat{u}_f(t) &=& u(t) + \hat{c}_1(t)u(t-1) + \dots + \hat{c}_{n_c}(t)u(t-n_c) \\ \hat{y}_f(t) &=& y(t) - \hat{\varphi}_s^T(t)\hat{\theta}_s(t-1) \end{array}$$

# 5.2 Filtering based recursive generalized extended LS algorithm

Consider the following ARARX model

$$A(z^{-1})y(t) = B(z^{-1})u(t) + \frac{D(z^{-1})}{C(z^{-1})}v(t)$$
(5.7)

Define the filtered input  $u_f(t)$ :  $=\frac{D(z^{-1})}{C(z^{-1})}u(t)$  and filtered output  $y_f(t) = \frac{D(z^{-1})}{C(z^{-1})}y(t)$ . Multiplying the both sides of (5.7) by  $\frac{C(z^{-1})}{D(z^{-1})}$ , yields

$$A(z^{-1})y_f(t) = B(z^{-1})u_f(t) + v(t)$$

For the above equation error type model, one has

$$y_f(t) = \varphi_f^T(t)\theta_s + v(t) \tag{5.8}$$

where

Next, we will construct another algorithm to estimate  $\frac{D(z^{-1})}{C(z^{-1})}$ . Define an intermediate variable

$$w(t) = \frac{D(z^{-1})}{C(z^{-1})}v(t)$$
(5.9)

which gives

$$C(z^{-1})w(t) = D(z^{-1})v(t)$$

By this relation, it can be obtained

$$w(t) = \varphi_n^T(t)\theta_n + v(t) \tag{5.10}$$

where

It follows from (5.7) and (5.9) that

$$w(t) = A(z^{-1})y(t) - B(z^{-1})u(t) = y(t) - \varphi_s^T \theta_s$$

where  $\varphi_s(t) = [-y(t-1) \ -y(t-2) \ \dots \ -y(t-n_a) \ u(t-1) \ u(t-2) \ \dots \ u(t-n_b)]^T$ .

Then we can obtain an estimate of w(t) as

$$\hat{w}(t) = y(t) - \hat{\varphi}_s^T(t)\hat{\theta}_s(t-1)$$

According to (5.8) and (5.10), two least square algorithms can be construct

$$\begin{cases} \hat{\theta}_{s}(t) &= \hat{\theta}_{s}(t-1) + L_{s}(t) \left[ \hat{y}_{f}(t) - \hat{\varphi}_{f}^{T}(t) \hat{\theta}_{s}(t) \right] \\ L_{s}(t) &= \frac{P_{s}(t-1)\hat{\varphi}_{f}(t)}{1 + \hat{\varphi}_{f}^{T}(t)P_{s}(t-1)\hat{\varphi}_{f}(t)} \\ P_{s}(t) &= \left[ I - L_{n}(t) \hat{\varphi}_{s}^{T}(t) \right] P_{s}(t-1), P_{s}(0) = p_{0}I \end{cases}$$
(5.11)

$$\begin{cases} \hat{\theta}_{n}(t) &= \hat{\theta}_{n}(t-1) + L_{n}(t) \left[ \hat{w}(t) - \hat{\varphi}_{n}^{T}(t) \hat{\theta}_{n}(t) \right] \\ L_{n}(t) &= \frac{P_{n}(t-1)\hat{\varphi}_{n}(t)}{1 + \hat{\varphi}_{n}^{T}(t)P_{n}(t-1)\hat{\varphi}_{n}(t)} \\ P_{n}(t) &= \left[ I - L_{n}(t) \hat{\varphi}_{n}^{T}(t) \right] P_{n}(t-1), P_{n}(0) = p_{0}I \end{cases}$$
(5.12)