Geographically weighted correspondence matrices for local change analyses and error reporting: mapping the spatial distribution of errors and change

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Abstract

This letter describes and applies generic approaches for generating local measures from the correspondence table. These were developed by integrating the functionality of 2 R packages: gwxtab and diffeR in order to demonstrate the development of spatially explicit accuracy measures comparing comparison and reference data (predicted and observed) and change measures if the data represent classes at time t and t+1. The approach builds on on earlier work that considered the measures derived from the correspondence matrix in the context of generalized linear models and probability, and instead computes local geographically weighted correspondence matrices, from which local statistic can be calculated. In this case a selection of the overall and categorical difference measures proposed by Pontius and Milones (2011) and Pontius and Santacruz (2014) were examined spatially.

Key Words:

geographically weighted; accuracy and error; correspondence matrix; validation matrix; error matrix

1. Introduction

One of the under-explored areas in research of remote sensing land cover and land use is the investigation of local statistical models. Most remote sensing methods (classification, validation, change detection, etc) apply global approaches under the assumption that models and relationships between the variables or data under consideration remain constant over geographical space (Comber et al., 2012). Any spatial 'variation' in the results is driven by variation in data or variable values. Global models describe processes or patterns which are assumed to be location independent of stationary. However, this assumption of process spatial stationarity and invariance is contradicted by the many observations of spatial auto-correlation in landscape processes, and in coincident remotely sensed data, especially classification error (from Campbell 1981 to Comber in press 1) and more widely in geographic analyses under Tobler's 1st law of geography (Tobler, 1970). Spatial auto-correlation occurs when changes in properties of nearby features in geographic space are found to be correlated, contradicting the underlying assumption of independence in statistical analysis and inference. The result is process spatial non-stationarity when the statistical pattern or relationship observed in one region differs from that in another. An example of this global rather than local philosophy in remote sensing is the persistence of the use of the correspondence matrix. In error reporting this summarises the spatial coincidence of a classified dataset with a reference dataset, sometimes referred to as predicted and observed data, and in change analyses it summarises the class to class transitions between data collected at different times.

Recent research has started to address this fundamental statistical blind spot in remote sensing, and spatially sensitive approaches have been proposed for remote sensing classification (Comber et al, 2016 in press 2), data fusion (Lisev et al., 2016) and in applications where remote sensing data provides one of the input variables, for example mapping above ground biomass (Propastin, 2012), population segregation (Yu and Wu, 2004), net primary production (Wang et al., 2005) and in epidemiology (Khormi and Kumar, 2011).

In remote sensing error analysis Foody (2005) calculated geographically distributed correspondence matrices and interpolated between them to generate surfaces of error. Extensions by Comber et al. (2012) and Comber (2013) developed geographically weighted (GW) measures accuracy from a GW logistic regression which were further extended to examine the spatio-temporal characteristics of classification accuracy by Tsutsumida and Comber (2015). All of these approaches provide spatially distributed measures of error that can be easily quantified using a simple logistic regression of the data in the correspondence matrix as described in Comber (2013). However, many other measures may be calculated from the correspondence matrix including Kappa estimates (Congalton, 1991) and the quantity and allocation disagreements suggested by Pontius and Millones (2011). Indeed, within statistics there is a long literature describing statistical measures that can be derived from the family of contingency tables (e.g. Hartigan and Kleiner, 1981; Friendly, 1994), of which the correspondence matrix is member. These are not straightforward to describe or formalise in a logistic regression, geographically weighted or not.

This letter describes a generic method for calculating spatially distributed correspondence matrices that support a much wider set of geographically weighted analyses of error, accuracy and correspondence. It uses the gwxtab R package (Brunsdon et al, 2016) as a framework for calculating local correspondence matrices. Then it uses these to calculate *local* difference metrics as described in Pontius and Milones (2011) and Pontius and Santacruz (2014) and implemented in the differ R package (Pontius and Santacruz, 2015). For good measure we calculate a local Kappa estimate as well. This letter highlights how thinking locally rather than globally can result in more spatially nuanced reportings of accuracy and other comparative measures such as change. It illustrates how generic tools such as gwxtab can be used to calculate local versions of any correspondence table derived metric, which can in turn be mapped, to generate novel, spatially distributed measures of accuracy.

2. Data and Methods

2.1 Data

The dataset used in this analysis was collated by International Institute for Applied Systems Analysis (IIASA) in Austria and is included in the gwxtab package. It describes land cover at 2,439 locations in the British Isles (minus the islands!) from 4 sources collapsed to 10 classes as described in Comber et al (2013): volunteered land cover data collected by the Geo-Wiki initiative (Fritz et al., 2012), the GLC-2000 database (Fritz et al., 2003), the MODIS land cover product (Loveland et al., 2000) and GlobCover (Bicheron et al., 2008). The analyses in this letter compare the Geo-Wiki and Modis data and interested researchers can explore the links above to find out more about them. Here they are used simply to illustrate the methods being proposed. Table 1 shows the correspondence matrix of the Modis (rows) against the Geo-Wiki data (columns). The values in the table describe the counts of pixels recorded in each class, in each dataset. Off diagonal elements in the matrix summarise disagreemnts in the land cover classes allocated to pixels in each dataset.

Table 1. The Geo-Wiki	(columns)	and Modis	data (rows)	${\it correspondence\ matrix}$

	1	2	3	4	5	6	7	8	9	10
1. Forest	31	3	37	33	10	1	55	0	1	1
2. Shrub	0	0	0	1	0	0	31	0	0	0
3. Grass	11	4	26	62	15	1	19	0	5	1
4. Crop	13	1	117	355	10	1	78	0	2	3
5. Mosaic	53	1	293	135	1	0	89	0	2	1
6. Wetland	0	0	0	0	0	0	0	0	0	0
7. Urban	21	1	104	49	0	0	749	0	5	4
8. Snow	0	0	0	0	0	0	0	0	0	0
9. Barren	0	0	0	0	0	0	0	0	0	0
10. Water	0	0	1	1	0	0	1	0	0	0

[1] 0.4764248

2.2 Geographically weighted correspondence matrices

A correspondence matrix summarises the spatial intersection of 2 datasets and in a remote sensing error analysis, it compares the classified data with higher quality reference data at sample locations. In full a correspondence analysis, for example examining change over time, it summarises the spatial intersection of all data points or pixels. However, it provides no information about the spatial distribution of change or error, and the global measures derived the from correspondence matrix may mask local variations (McGwire and Fisher, 2001).

The basic idea of geographically weighted approach is that local correspondence matrices are computed from subsets of the full datasets at predefined locations, in order to quantify the spatial variation in the relationship between the 2 datasets, for example to generate spatially distributed measures of error. At each location, a subset of the data falling under a kernel are weighted by their distance to that location and then used to construct the correspondence matrix. This local correspondence matrix can be used to calculate the statistic of interest.

The kernel size or bandwidth, can be fixed (e.g. 20 km) or it can be adaptive to subset the nearest n data points (e.g. 15%) and different kernel functions (shapes) can be used for the distance weighting.

Generally larger bandwidths result in a greater degree of spatial smoothing. Gollini et al (2013) describe some of these and methods for determining bandwidth optimally. In this case, an adaptive bandwidth of 15% was specified and a bisquare kernel were applied. For a given bandwidth h, this is defined by:

$$f(d) = \begin{cases} \left(1 - \left(\frac{d}{h}\right)^2\right)^2 & \text{if } d < h; \\ 0 & \text{otherwise.} \end{cases}$$
 (1)

where d is the distance of the data point to the kernel centre.

Local, geographically weighted correspondence matrices were constructed to compare volunteered land cover data collected by the Geo-Wiki initiative (Frizt et al, 2012) with the MODIS global land cover products at 4304 locations on a hexagonal grid covering the study area. The study area, grid and data points are shown in Figure 1, with 2 example locations labeled. Tables 2 and 3 show the *geographically weighted* correspondence matrices at these locations. As in Table 1, he values in the table describe the counts of pixels recorded in each class, in each dataset, but now just for that location. The values are the sums of the *distance weighted* pixel counts.

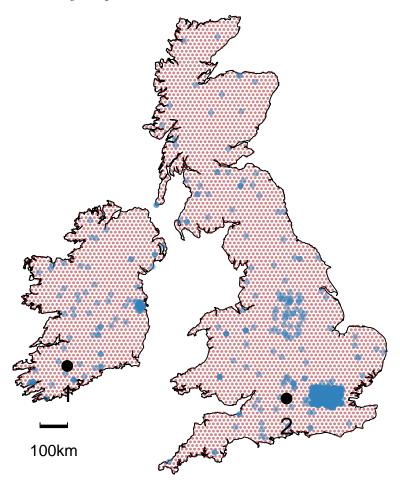


Figure 1: The study area, analysis grid in red and the data points,in blue with a transparency term, and 2 example locations.

Table 2. The local Geographically Weighted correspondence matrix at Location 1

	1	2	3	4	5	6	7	8	9	10
1. Forest	1.58	0	3.30	0.00	0.01	0	0.00	0	0	0
2. Shrub	0.05	0	2.81	0.74	0.00	0	0.76	0	0	0
3. Grass	0.60	0	5.01	0.16	0.01	0	0.00	0	0	0
4. Crop	3.94	0	32.24	7.44	0.58	0	0.23	0	0	0
5. Mosaic	2.25	0	5.51	1.55	0.03	0	0.00	0	0	0
6. Wetland	0.00	0	0.41	0.24	0.00	0	0.00	0	0	0
7. Urban	0.91	0	4.46	0.79	0.09	0	18.66	0	0	0
8. Snow	0.00	0	0.00	0.00	0.00	0	0.00	0	0	0
9. Barren	0.94	0	2.95	0.00	0.00	0	0.00	0	0	0
10. Water	0.00	0	0.00	0.24	0.00	0	0.00	0	0	0

Table 3. The local Geographically Weighted correspondence matrix at Location 2

	1	2	3	4	5	6	7	8	9	10
1. Forest	1.12	0	0.00	0.11	1.12	0	0.52	0	0	0
2. Shrub	0.18	0	0.00	0.00	0.07	0	0.00	0	0	0
3. Grass	0.83	0	0.11	2.08	2.10	0	0.62	0	0	0
4. Crop	1.15	0	0.11	11.27	3.73	0	0.94	0	0	0
5. Mosaic	0.88	0	0.04	1.57	0.24	0	0.00	0	0	0
6. Wetland	0.00	0	0.00	0.00	0.00	0	0.00	0	0	0
7. Urban	0.95	0	0.28	0.62	1.71	0	3.90	0	0	0
8. Snow	0.00	0	0.00	0.00	0.00	0	0.00	0	0	0
9. Barren	0.00	0	0.00	0.47	0.31	0	0.00	0	0	0
10. Water	0.00	0	0.00	0.00	0.15	0	0.00	0	0	0

Accuracy measures: Quantity and Allocation Disagreements

Pontius and Millones (2011) and Pontius and Santacruz (2014) describe a number of methods for calculating difference metrics and composite measures of accuracy from the correspondence matrix. These are based around map-to-map cross-tabulations or correspondence matrices and can be used to compare mapped land cover data for error or change. For example, quantity disagreement is defined as the amount of difference between the Observed reference data and the Predicted classified data relative to the proportions of the classes in the Observed and Predicted data. It is computed from the sum of the row totals (the Predicted data) minus the sum of the column totals (the Observed data) divided by 2. Similarly, the allocation disagreement is defined as the amount of difference between the Observed data and the Predicted data that are due mis-matches in the spatial allocation of classes, relative to the class proportions. It is computed from the total number of pixels minus the diagonal agreement, minus the quantity disagreement. In all cases the measures can be computed from correspondence tables of counts of coincident pixels or proportions. Some of the measures from the diffeR package that were applied in this analysis are summarised in Table 4.

Table 4: A summary of the difference and disagreement measures in the diffeR package

Measure	Function	Return	Descriptions
		value	
Overall allocation dif-	overallAllocD	Single	The amount of difference between observed and
ference		value	predicted data due to the *less than maximum
			match* in the spatial allocation of the cate-
			gories, given the proportions of the categories
Overall difference	overallDiff	Single	in both datasets. The overall difference between tabulated ob-
Overall difference	Overalibili	value	served and predicted data calculated from the
		varue	sum of the quantity and allocation components
			of difference.
Overall exchange differ-	overallExchangeD	Single	Exchange is the transition from category i to
ence		value	category j in some observations and a transi-
			tion from category j to category i in an identi-
			cal number of other observations.
Overall quantity differ-	overallQtyD	Single	This is the amount of difference between the
ence		value	observed variable and a predicted variable that is due to the less than maximum match in the
			proportions of the categories.
Overall shift difference	overallShiftD	Single	Shift describes to the difference remaining after
O vertair sinire dinorence	0 1 0 1 0 1 1 1 1 1 1 1 1	value	subtracting Quantity difference and Exchange
			from the Overall difference.
Category exchange dif-	exchangeDj	Value for	The exchange value as above at the class level.
ference		each class	
Category overall differ-	overallDiffCatj	Value for	The overall difference as above at the class
ence		each class	level.
Category quantity dif-	quantityDj	Value for	The quantity difference as above at the class
ference	ahif+Di	each class Value for	level. The shift difference as above at the class level.
Category shift difference	shiftDj	each class	The shift difference as above at the class level.
		Cacii Ciabb	

Addendum: Kappa Estimates

Although now widely discredited, many remote sensing analyses still use the Kappa estimate, $\hat{\kappa}$. Its derivation is below:

$$\hat{\kappa} = \frac{N \sum_{i=1}^{r} x_{ii} - \sum_{i=1}^{r} (x_{i+} * x_{+i})}{N^2 - \sum_{i=1}^{r} (x_{i+} * x_{+i})}$$
(2)

where N is the total number of observations in the matrix, r is the number of rows, x_{ii} is observations in row i and column i, x_{i+} and x_{+i} are the marginal totals of row i and column i, respectively. Essentially what this does is:

- 1. Multiply the sum of the diagonals by the table sum.
- 2. Then subtract from this the sum of the row and column marginal totals products.

3. Next, divide this by the sum of the table squared, minus the sum of the row and column marginal totals products, as above.

The top part of the equation gives a measure of chance agreement and the bottom part a measure of the expected disagreements.

Code

The Rmarkdown script used to produce this manuscript, including all the code used in the analysis and to produce the mapped figures, can be found at https://github.com/lexcomber/RSLcode

Results

A fixed bandwidth of 15% of the data points under a bisquare kernel was specified, and for each grid location the Kappa was estimated. A series of geographically weighted analyses were undertaken to compare the MODIS and Geo-Wiki data using the metrics described in Table 1. Maps of the spatial distribution of different accuracy measures are shown in Figure 2, 3 and 4. Figure 2 maps overall comparative measures and Figure 3 maps class specific measures from Pontius and Milones (2011) and Pontius and Santacruz (2014) along with the spatially distributed local Kappa estimate. Figure 4 maps User and Producer accuracies for the class or *Urban*. The definitions of the different measures can be found in the papers cited above or in the different package. Room precludes the full description here but the critical point is that they vary spatially For example the maps in Figure 4 show how the User and Producer accuracies from the local cross-tabulations vary from the global estimates derived from Table 1. Also, critical to note are that measures of agreement are *not* related to the density of the data points, *per se*. Rather they reflect local measures calculated from local correspondence matrices at the locations mapped in Figure 1.

Discussion

The correspondence matrix is the de facto method in remote sensing for reporting comparisons between classified data, habitats, land cover and land use, for accuracy and error reporting and for sumamrise the results of change detection analyses. Local, spatially distributed, geographically weighted cross-tabulations matrices provide a framework for examining how and where measures derived from the correspondence matrix vary. They are an advance on the logistical regression methods suggested in Comber et al (2012) and Comber (2013) because they support the generation of any local measure that the user wishes to specify. In this case the analyses integrated functions from 2 R packages: the gwxtab package to create local geographically weighted correspondence matrices and the overall and categorical difference measures in the diffeR package. The availability of open, free and transparent code provides a dynamic and rich research environment within which method extensions can be developed.

Whilst this paper advocates the application of local statistical models and spatially dependent methods, precisely because they reflect our understanding of nearly all processes and relationships we have encountered in natural and human sciences, (Tobler's 1st Law of Geography - Tobler, 1970), we recognise that for many policy makers global statistics simplify complex data and provide an overall summary of the data and are therefore widely used. Local measures may not be intuitively understood. However, maps provide an incredibly powerful and readily understood representation. So whilst policy makers may not immediately understand the various error measures in the DiffeR package, they will understand the mapped spatial distribution of a measure. It is therefore incumbent on the Remote Sensing community to more strongly engage with more advanced reporting techniques, such as local statistical models in their funded science. There is a previous example of this: land cover uncertainty reporting. In the 1980s and 1990s policy makers struggled to understand that maps might contain errors and variations in representation. Now they do not: 10% reported error rates are readily understood. Providing open code and transparent methods (e.g through sites like https://github.com/ is one way that the academic community could better support the up-skilling of policy makers (and maybe result in more interesting research being funded in Remote Sensing).

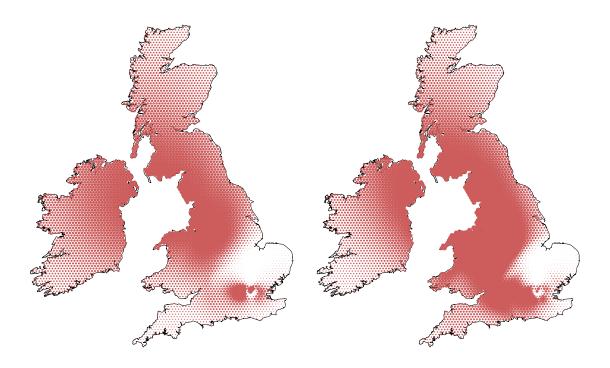
There are some important considerations related to the application of GW models, including bandwidth specification and kernel shape as described in Gollini et al (2013). Here a bandwidth of 15% of the data points was selected but exploration using a range of bandwidths is recommended as bandwidth affects the degree of smoothing and thus the sensitivity of the analyses to the data distribution. As yet the gwxtab package does not include a function for bandwidth optimisation (as do other GW packages in R such as GWmodel and spgwr). This is an area that the developers of gwxtab are working on and will provide greater confidence in the results of the spatially distributed models, without requiring a formal sampling strategy.

The GW framework for correspondence tables presented here supports greater understanding of the spatial process and statistical relationships under investigation. This is important as the number and diversity of remote sensing derived products and applications increases and reflects the original aims of geographically weighted regression (Brunsdon et al., 1996). For example, it could be used to dynamically visualize accuracy, to characterise error, to provide local distributions of Chi-squared statistics, to explore the implications of different to locally focus additional ground-truth sampling, to assess the value of the imagery itself locally (e.g. LANDSAT versus MODIS), to explore the utility of different class definitions (e.g. Cropland vs. Managed grassland, which can be particularly difficult to classify) and to identify locales with missing and misaligned data.

Finally, spatially explicit approaches such as the GW correspondence matrices allow some of the dominant assumptions of spatial non-staionarity of processes within remote sensing methods to be examined and tested. They accommodate the spatial auto-correlation found remote sensing data and analyses of many landscape processes.

a) Overall allocation difference

b) Overall exchange difference



c) Overall quantity difference

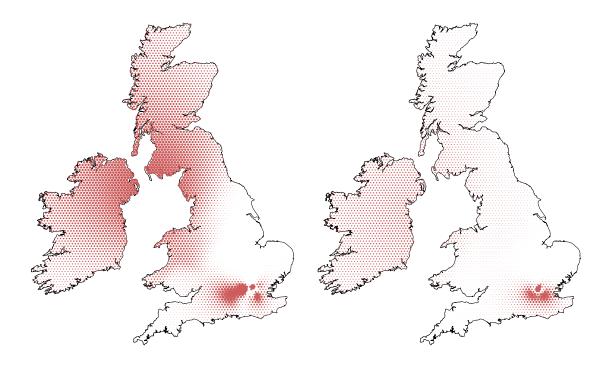
d) Kappa Estimate



Figure 2: The spatial distribution of the difference measures, scaled to $[0,\,1]$, comparing the Geo-Wiki crowdsourced data with MODIS data.

a) Exchange difference for Forest

b) Category Overall difference for Urban



c) Category quantity difference for Grass

d) Category shift difference

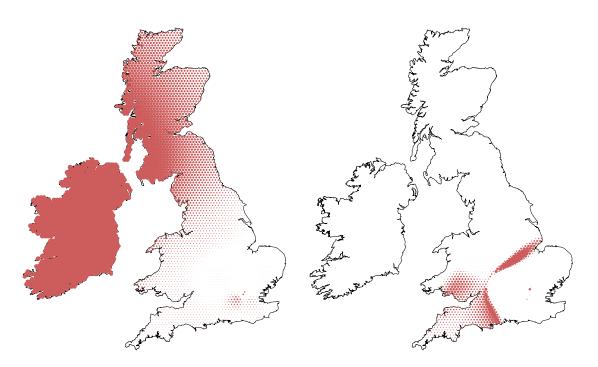


Figure 3: The spatial distribution of the class level, or categorical difference measures, scaled to [0, 1], comparing the Geo-Wiki crowdsourced data with MODIS data.

a) Urban User Accuracy, Globally 0.803 Locally 0.521 (1st Qu) to 0.723 (2nd Qu) a) Urban Producer Accuracy, Globally 0.733 Locally 0.752 (1st Qu) to 0.939 (2nd Qu)

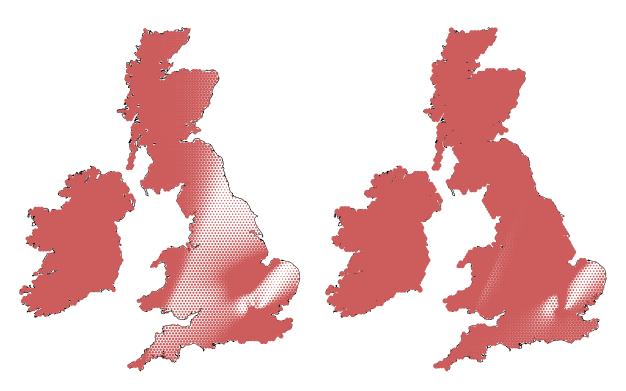


Figure 4: The spatial distribution of a) User, and b) Producer accuracy values for the class of Urban, scaled to [0, 1], using the Geo-Wiki data to validate MODIS data.

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