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### Nonlinear Regression <sup>1</sup>

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 $<sup>^{1}</sup>$ This section is based on Stock and Watson (2020), Chapter 8.4  $\equiv$   $^{1}$ 

### Nonlinear regression

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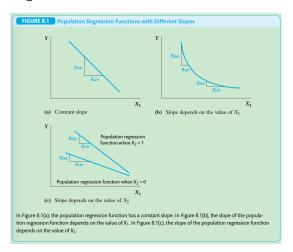
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### Motivation I

Goal: two groups of methods for detecting and modeling nonlinear population regression functions.



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### Motivation II

- 1. the relationship between Y and an independent variable,  $X_1$ , depends on the value of  $X_1$  itself.
  - E.g. reducing class sizes by one student per teacher might have a greater effect if class sizes are already manageably small.
  - If so, the test score (Y) is a nonlinear function of the student-teacher ratio (X<sub>1</sub>), where this function is steeper when X<sub>1</sub> is small.

#### Nonlinear regression

- 2. the effect on Y of a change in  $X_1$  depends on the value of another independent variable—say,  $X_2$ .
  - For example, students still learning English might especially benefit from having more one-on-one attention;
  - if so, the effect on test scores of reducing the student-teacher ratio will be greater in districts with many students still learning English than in districts with few English learners.
- ▶ If CEF is nonlinear function of the X's and of the parameters.
- ▶ If so, the parameters cannot be estimated by OLS.
- Can be estimated using nonlinear least squares.

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### General Strategy I

- be the nonlinear models are extensions of the multiple regression model and (can be estimated and tested using OLS).
- ▶ Consider the relationship between (*Test Scores* and *District Income*)
- ▶ Measure ratio of students with poor family background:
  - the percentage of students qualifying for a subsidized lunch
  - the percentage of students whose families qualify for income assistance.
  - Alternatively: the average annual per capita income in the school district ("district income").

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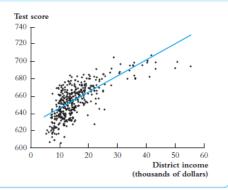
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### General Strategy II



There is a positive correlation between test scores and district income (correlation = 0.71), but the linear OLS regression line does not adequately describe the relationship between these variables.



- Data: income measured in thousands of 1998 dollars.
- ▶ median district income is 13.7(\$13,700 per person).
- ▶ ranges from 5.3 to 55.3.

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### General Strategy III

- ▶ Test scores and district income are strongly positively correlated, with a correlation coefficient of 0.71;
- There seems to be some curvature in the relationship between test scores and district income that is not captured by the linear regression.

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### General Formula for Nonlinear Regression

▶ The nonlinear population regression models considered in this chapter are of the form

$$Y_i = f(X_{1i}, X_{1i}, \dots, X_{ki}) + u_i, i = 1, \dots, n, (8.3)$$

where  $f(X_{1i}, X_{1i}, \ldots, X_{ki})$  is the population nonlinear regression function, a possibly nonlinear function of the independent variables  $X_{1i}, X_{1i}, \ldots, X_{ki}$  and  $u_i$  is the error term.

▶ E.g., in the quadratic regression model

$$f(Income_i) = \beta_0 + \beta_1 Income_i + \beta_2 Income_i^2$$
.

### Partial Effect

 $\triangleright$  Experiment: on individuals with the same values of  $X_2, \ldots, X_k$ , and participants are randomly assigned treatment levels  $X_1 = x_1$ or  $X_1 + \Delta X_1 = x_1 + \Delta x_1$ .

- ▶ The expected difference in Y is the causal effect of the treatment, holding constant  $X_2, \ldots, X_k$ .
- ▶ In the quadratic regression model, this effect on Y is  $\Delta Y = f(X_1 + \Delta X_1, \dots, X_k) - f(X_1, X_2, \dots, X_k).$
- $\triangleright$   $\triangle$  is the predicted difference in with difference  $X_1 + \triangle X_1$  and  $X_1$ .

### Partial Effect Application example

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- ▶ The regression function f is unknown, this population causal effect is also unknown.
  - $\diamond$  First estimate the regression function f.
  - $\diamond$  At a general level, denote this estimated function by  $f_n$ ;
  - $\diamond$  e.g. estimated quadratic regression function in Equation (8.2).
- $\triangleright \hat{Y} = \hat{f}(X_1 + \Delta X_1, \dots, X_k) \hat{f}(X_1, X_2, \dots, X_k).$

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### Application example I

$$\textit{TestScore} = \underset{(2.9)}{607.3} + \underset{(0.27)}{3.85} \textit{Income} - \underset{(0.0048)}{0.0423} \textit{Income}^2, R^2 = 0.554, (8.2)$$

- ▶ What is the predicted change in test cores associated with a change in district income of \$1000, based on the estimated quadratic regression function?
- b this effect depends on the initial district income. We therefore consider two cases: an increase in district income from 10 to 11 (i.e., from \$10,000 per capita to \$11,000 per capita) and an increase in district income from 40 to 41 (i.e., from \$40,000 per capita to \$41,000 per capita).

 $\triangleright$ 

$$\Delta \hat{Y} = (\hat{\beta}_0 + \hat{\beta}_1 \times 11 + \hat{\beta}_2 \times 11^2) - (\hat{\beta}_0 + \hat{\beta}_1 \times 10 + \hat{\beta}_2 \times 10^2).$$

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### Application example II

- when Income = 10, the predicted value of test scores is 607.3 + 3.85\*10 0.0423\*102 = 641.57. When Income = 11, the predicted value is 607.3 + 3.85\*11 0.0423\*112 = 644.53. The difference in these two predicted values is  $Y_n = 644.53 641.57 = 2.96$  points.
- when income changes from \$40,000 to \$41,000, the difference in the predicted values in Equation (8.6) is  $Y_n = (607.3 + 3.85 * 41 0.0423 * 41^2) (607.3 + 3.85 * 40 0.0423 * 40^2) = 694.04 693.62 = 0.42 points.$
- ▶ Thus a change of income of \$1000 is associated with a larger change in predicted test scores if the initial income is \$10,000 than if it is from \$40,000 to \$41,000.

Application example

### Standard errors of estimated effects I

- ▶ **Sampling error** The estimate of the effect on Y of changing X depends on  $\hat{f}$ , which varies from one sample to the next.
- Example:

$$SE(\Delta Y) = SE(\hat{\beta}_1 + 21\hat{\beta}_2)$$

 $\diamond$  if we can compute the standard error of  $\hat{\beta}_1 + 21\hat{\beta}_2$  then we have computed the standard error of Yn.

### Standard errors of estimated effects II

- ▶ If not directly computable, there are two other ways to compute it; these correspond to the two approaches for testing a single restriction on multiple coefficients.
  - 1. The first method is to use approach 1 of Section 7.3, which is to compute the F-statistic testing the hypothesis that  $\beta_1 + 21\beta_2 = 0$ . The standard error of  $Y_n$  is then given by

$$SE(\Delta \hat{Y}) = \frac{|\Delta \hat{Y}|}{\sqrt{F}}.(8.8)$$

the F-statistic testing the hypothesis that  $\beta_1 + 21\beta_2 = 0$  is F = 299.94. With  $\Delta \hat{Y} = 2.96$ ,  $SE(\Delta \hat{Y}) = 2.96/\sqrt{299.94} = 0.17$ .

2. The second approach is to compute the standard error using  $var(\hat{\beta})$ .

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### Comment on interpreting coefficients I

- ▶ In the multiple regression model, the regression coefficients had a natural interpretation.
- $\triangleright$  For example,  $\beta_1$  is the expected change in Y associated with a change in  $X_1$ , holding the other regressors constant.
- $\triangleright$  But as we have seen, this is not generally the case in a nonlinear model. e.g.  $\hat{\beta}_1$
- ▶ In nonlinear models, the regression function is best interpreted by graphing it and by calculating the predicted effect on *Y* of changing one or more of the independent variables.

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### General Approach I

The general approach to modeling nonlinear regression functions taken in this chapter has five elements:

- 1. Identify a possible nonlinear relationship.
  - Use economic theory and what you know about the application to suggest a possible nonlinear relationship.
  - Why might such nonlinear dependence exist? What nonlinear shapes does this suggest?
  - e.g. classroom dynamics with 11-year-olds suggests that cutting class size from 18 students to 17 could have a greater effect than cutting it from 30 to 29.
- 2. Specify a nonlinear function and estimate its parameters by OLS.
- Determine whether the nonlinear model improves upon a linear model.
  - Use t-statistics and F-statistics to test the null hypothesis that the population regression function is linear against the alternative that it is nonlinear.
- 4. Plot the estimated nonlinear regression function. Does
- 5. Estimate the effect on Y of a change in X.

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○ One way to specify a nonlinear regression function is to use a polynomial in X. In general, let r denote the highest power of X that is included in the regression. The polynomial regression model of degree r is

$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 X_i^2 + \dots \beta_r X_i^r + u_r.$$

- $\triangleright$  When r = 2, the model is called **quadratic regression** model.
- $\triangleright$  When r = 3, the model is called **cubic regression model**.

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# Testing the null hypothesis that the population regression function is linear I

- ▶ If the population regression function is linear, then the quadratic and higher-degree terms do not enter the population regression function.
- $\triangleright$  The null hypothesis  $H_0$  that the regression is linear and the alternative  $H_1$  that it is a polynomial of degree up to r

$$H0: \beta_2 = 0, \beta_3 = 0, \dots, \beta_r = 0 \text{ vs. } H1: \text{ at least one } \beta_j \neq 0, j = 2,$$

 $\triangleright$  Can be tested against the alternative that it is a polynomial of degree up to r by testing  $H_0$  against  $H_1$  using joint null hypothesis with q=r-1 restrictions.

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### Determine the order of the polynomial I

- ▶ how many powers of X should be included in a polynomial regression? trade-off between flexibility and statistical precision.
- ▶ Increasing the degree *r* introduces more flexibility.
- ▶ reduce the precision of the estimated coefficients.
- ▶ Include enough to model the nonlinear regression function adequately—but no more.

### Sequential Test:

- 1. Pick a maximum value of r, and estimate the polynomial regression for that r
- 2. Test  $H_0$ :  $\beta_r = 0$ , if reject, test for r + 1.
- 3. If fail to reject: test r-1.
- 4. continue this procedure until the coefficient on the highest power in your polynomial is statistically significant.

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### Application to district income and test scores I

The estimated cubic regression function relating district income to test scores is

$$\widehat{\textit{TestScore}} = \underset{(5.1)}{600.1} + \underset{(0.71)}{5.02} \\ \textit{Income} - \underset{(0.029)}{0.096} \\ \textit{Income}^2 + \underset{(0.00035)}{0.00036} \\ \bar{R}^2 = 0.555.$$

- ▶ The t-statistic on *Income*<sup>3</sup> is 1.97, so the null hypothesis that the regression function is a quadratic is rejected against the alternative that it is a cubic at the 5% level.
- b the F-statistic testing the joint null hypothesis that the coefficients on *Income*<sup>2</sup> and *Income*<sup>3</sup> are both 0 is 37.7, with a p-value less than 0.01%.
- > null hypothesis that the regression function is linear is rejected against the alternative that it is either a quadratic or a cubic.

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## Interpretation of coefficients in polynomial regression models I

- ▶ The coefficients in polynomial regressions do not have a simple interpretation.
- The best way to interpret polynomial regressions is to plot the estimated regression function and calculate the estimated effect on Y associated with a change in X for one or more values of X.

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### Logarithms I

- $\triangleright$  Another way to specify a nonlinear regression function is to use the natural logarithm of Y and/or X.
- ▶ Logarithms convert changes in vairables into percentage changes.

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### Example

- SW Ch 3, "Social Class or Education? Childhood Circumstances and Adult Earnings Revisited," examined the household earnings gap by socioeconomic classification.
  - Easier to compare wage gaps across professions and over time when they are expressed in percentage terms.
- ▶ In SW 8.1, we found that district income and test scores were nonlinearly related.
  - might it be that a change in district income of 1%—rather than \$1000—is associated with a change in test scores that is approximately constant.
- ▷ In the economic analysis of consumer demand, it is often assumed that a 1% increase in price leads to a certain percentage decrease in the quantity demanded(price elasticity).

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The logarithm function has the following useful properties

$$\begin{aligned} &\ln(1/x) = -\ln x \\ &\ln(ax) = \ln a + \ln x \\ &\ln(a/x) = \ln a - \ln x \\ &\ln(a^x) = x \ln a. \end{aligned}$$

Logarithms

### Logarithms and percentages I

The link between the logarithm and percentages relies on a key fact: When x is small, the difference between the logarithm of  $x + \Delta x$  and the logarithm of x is approximately  $\Delta x/x$ , the percentage change in x divided by 100. That is,

$$\ln(x + \Delta x) - \ln(x) \approx \frac{\Delta x}{x}$$
. (when  $\frac{\Delta x}{x}$  is small.)

Proof: using taylor expansion.

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Logarithms

### The three logarithmic regression models I

- 1. linear-log model.  $Y_i = \beta_0 + \beta_1 \ln(X_i)$ . In the linear-log model, a 1% change in X is associated with a change in Y of  $0.01\beta1$ .
- 2. log-linear model. In  $Y_i = \beta_0 + \beta_1 X_i$ . a one-unit change in  $\Delta X_1 = 1$  is associated with a  $(100 * \beta_1)\%$  change in Y.
- 3. log-log model. In  $Y_i = \beta_0 + \beta_1 \ln X_i$ . In the log-log model, a 1% change in X is associated with a  $\beta_1$  change in Y.

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# A difficulty with comparing logarithmic specifications I

- by the  $R^2$  can be used to compare the log-linear and log-log models; as it happened, the log-log model had the higher  $R^2$ .
- be the  $R^2$  can be used to compare the linearlog regression in and the linear regression of Y against X.

In the test score and district income regression, the linear-log regression has an  $R^2$  of 0.561, while the linear regression has an  $R^2$  of 0.508, so the linear-log model fits the data better.

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### A difficulty with comparing logarithmic specifications II

### How can we compare the linear-log model and the log-log model?

- $\triangleright$  Unfortunately, the  $R^2$  cannot be used to compare these two regressions because their dependent variables are different.
- $\triangleright$  Recall that the  $R^2$  measures the fraction of the variance of the dependent variable explained by the regressors.
- ▶ The dependent variables in the log-log and linear-log models are different, it does not make sense to compare their  $R^2$ 's.

Use economic theory to decide.

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# Computing predicted values of Y when Y is in logarithms I

Consider the log-linear regression model,

$$Y_i = \exp(\beta_0 + \beta_1 X_i + u_i).$$

Then 
$$\hat{Y}_i = \exp(\hat{\beta}_0 + \hat{\beta}_1 X_i)$$
.

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## Polynomial and Logarithmic Models have of Test Scores and District Income I

- economic theory or expert judgment might suggest a functional form to use,
- ▶ the true form of the population regression function is unknown.
- Need to decide which method or combination of methods works best.

**Polynomial specifications**. Because the coefficient on *Income*<sup>3</sup> was significant at the 5% level, select the cubic model as the preferred polynomial specification.

**Logarithmic specifications**. The logarithmic specification seemed to provide a good fit to these data.

**One way to test**: to augment it with higher powers of the logarithm of income. If not statistically different from 0, then we can conclude

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### Polynomial and Logarithmic Models have of Test Scores and District Income II

that the specification is adequate in the sense that it cannot be rejected against a polynomial function of the logarithm.

$$\widehat{TestScore} = \frac{(79.4)}{486.1} + \frac{(87.9)}{113.4} ln(Income) - \frac{(31.7)}{26.93} ln(Income)^{2} + \frac{(3.74)}{3.063} ln(Income)^{3}, R^{2} = 0.560.$$

he F-statistic testing the joint hypothesis that the true coefficients on the quadratic and cubic term are both 0 is 0.44, with a p-value of 0.64.

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### Interactions Between Two Binary Variables I

▶ Binary variable interaction regression model Introducing another regressor, the product of the two binary variables,  $D_{1i} * D_{2i}$ . The resulting regression is

$$Y_i = \beta_0 + \beta_1 D_{1i} + \beta_2 D_{2i} + \beta_3 (D_{1i} \times D_{2i}) + u_i$$

▶ The new regressor, the product  $D_{1i} * D_{2i}$ , is called an interaction term or an **interacted regressor**.

 $\triangleright$ 

$$E[Y_i|D_{1i}=1,D_{2i}=d_2]-E[Y_i|D_{1i}=0,D_{2i}=d_2]=\beta_1+\beta_3d_2.$$

- ▶ The effect of acquiring a college degree depends on the person's sex.
- ▶ The binary variable interaction regression allows the effect of changing one of the binary independent variables to depend on the value of the other binary variable.

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# Application to the student-teacher ratio and the percentage of English learners I

- $ightharpoonup HiSTR_i$  be a binary variable that equals 1 if the student-teacher ratio is 20 or more and that equals 0 otherwise.
- $ightharpoonup HiEL_i$  be a binary variable that equals 1 if the percentage of English learners is 10% or more and that equals 0 otherwise.

$$\widehat{TestScore} = 664.1 - 1.9 \underbrace{HiSTR}_{(1.9)} - 18.2 \underbrace{HiEL}_{(2.3)} - 3.51 \underbrace{HiSTR}_{(3.1)} * HiEL,$$

 $\triangleright$ 

- $\triangleright$  The predicted effect of moving from a district with a low student–teacher ratio to one with a high student–teacher ratio is -1.9-3.5 HiEL.
- $\triangleright$  if the fraction of English learners is low HiEL=0, then the effect on test scores of moving from HiSTR=0 to HiSTR=1 is for test scores to decline by 1.9 points.

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## Application to the student-teacher ratio and the percentage of English learners II

- $\triangleright$  If the fraction of English learners is high, then test scores are estimated to decline by 1.9 + 3.5 = 5.4 points.
- Can use to estimate the mean test scores for each of the four possible combinations of the binary variables.
- $\triangleright$  HiSTR<sub>i</sub> = 0 (low student-teacher ratios) and HiEL<sub>i</sub> = 0 (low fractions of English learners) is 664.1.
- $\triangleright$  HiSTR<sub>i</sub> = 1 (high student-teacher ratios) and HiEL<sub>i</sub> = 0 (low fractions of English learners), the sample average is 662.2
- $\triangleright$  When  $HiSTR_i = 0$  and  $HiEL_i = 1$ , the sample average is 645.9
- $\triangleright$  HiSTR<sub>i</sub> = 1 and HiEL<sub>i</sub> = 1, the sample average is 640.5.

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### Interactions Between a Continuous and a Binary Variable I

$$Y_{i} = \beta_{0} + \beta_{1}X_{i} + \beta_{2}D_{i} + \beta_{3}(X_{i} * D_{i2}) + u_{i},$$

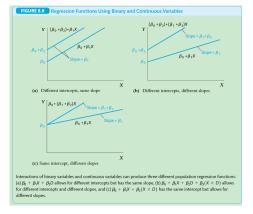
$$Y_{i} = \beta_{0} + \beta_{1}X_{i} + \beta_{2}D_{i} + u_{i},$$

$$Y_{i} = \beta_{0} + \beta_{1}X_{i} + \beta_{3}(X_{i} * D_{i2}) + u_{i}.$$

where  $X_i * D_i$  is a new variable, the product of  $X_i$  and  $D_i$ .

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### Interactions Between a Continuous and a Binary Variable II



All three specifications are versions of the multiple regression model, and once the new variable  $X_i * D_i$  is created, the coefficients of all three can be estimated by OLS.

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#### Interaction

# Application to the student-teacher ratio and the percentage of English learners I

**Research Question**: Does the effect on test scores of cutting the student—teacher ratio depend on whether the percentage of students still learning English is high or low?

Test if the slope/intercept is different.

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$$\widehat{TestScore} = \underset{(11.9)}{682.2} - \underset{(0.59)}{0.97}STR + \underset{(19.5)}{5.6}HiEL - \underset{(0.97)}{1.28}(STR * HiEL)$$

- ▶ reducing the student-teacher ratio by 1 is predicted to increase test scores by 0.97 points in districts with low fractions of English learners but by 2.25 points in districts with high fractions of English learners.
- ▶ The difference between these two effects, 1.28 points, is the coefficient on the interaction term.
- nuanced policy counterfactuals.

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## Application to the student-teacher ratio and the percentage of English learners II

#### ▶ Test

- 1. the two lines are the same. This *F*-statistic is 89.9, significant at the 1% level.
- 2. the hypothesis that two lines have the same slope can be tested. (Cannot reject at 10 % level)
- 3. Third, the hypothesis that the lines have the same intercept. (Cannot reject at 5 % level)
- ▶ Contradictory results.
- The reason is that the regressors, HiEL and STR ∗ HiEL, are highly correlated.
- ▶ results in large standard errors on the individual coefficients.
- be impossible to tell which of the coefficients is nonzero, there is strong evidence against the hypothesis that both are 0.

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# Application to the student-teacher ratio and the percentage of English learners III

▶ Finally, the hypothesis that the student-teacher ratio does not enter this specification can be tested by computing the F-statistic for the joint hypothesis that the coefficients on STR and on the interaction term are both 0. This F-statistic is 5.64(p-value: 0.004).

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## Application to the student-teacher ratio and the percentage of English learners I

$$\widehat{\textit{TestScore}} = \underset{(11.8)}{686.3} - \underset{(0.59)}{1.12} STR - \underset{(0.37)}{0.67} \textit{PctEL} + \underset{(0.019)}{0.0012} (\textit{STR} * \textit{PctEL}), \\ \bar{R}^2 = 0.422.(8.37)$$

- When the percentage of English learners is at the median (PctEL = 8.85), the slope of the line relating test scores and the student–teacher ratio is estimated to be -1.11 (= -1.12 + 0.0012 \* 8.85).
- ▶ at the 75th percentile (PctEL = 23.0), slope of -1.09 (= -1.12 + 0.0012 \* 23.0).
- ▶ The difference between these estimated effects is not statistically significant (t = 0.0012/0.019 = 0.06).

References

Stock, J. H. and Watson, M. W. (2020). Introduction to econometrics, volume 4. Pearson New York.