#### Statistics

Hao

Estimators

Hypothesis Testing

Hypothesis

Town Load To

Statistical

Significano

Interval

Hypothesis Testing

Test of Causa Effect

References

# Review of Statistics <sup>1</sup>

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 $<sup>^1</sup>This$  section is based on Stock and Watson (2020), Chapter 3.4  $\Xi$  > 3.4  $\Xi$  > 9.9  $^{\circ}$ 

#### Estimators

Hypothe

Hypothesis Type I and Ty

Statistica Significan

Interval

Hypothesis Testin

Test of Causa Effect

References

Suppose you want to understand the distribution of X in the population.

- ▶ When a statistic  $\hat{\theta} = \hat{\theta}(X_1, \dots, X_n)$  is a function of an i.i.d. sample, then the distribution is determined by the population distribution is F and the sample size is n.
- $\triangleright$  We call the distribution of  $\hat{\theta}$  the **sample distribution**.

The goal of an estimator  $\hat{\theta}$  is to learn about the parameter  $\theta$ , we evaluate the

- ▶ The exact bias and variance.
- ▶ The distribution under normality.
- $\triangleright$  The asymptotic distribution as  $n \to \infty$ .

Reference

### Goodness of Estimators

Let  $\hat{\theta}$  be an estimator of  $\theta$ . Then

- $\triangleright$  The bias of  $bias(\hat{\theta})$  is  $\hat{\theta} \theta$ .
  - ♦ We say an estimator is **unbiased** if the bias is 0.
- ightharpoonup The **mean squared error** of an estimator  $\hat{\theta}$  for  $\theta$  is

$$\mathit{mse}(\hat{\theta}) = \mathbb{E}[(\hat{\theta} - \theta)^2].$$

 $\diamond$  The mean squared error is  $mse(\hat{\theta}) = var(\hat{\theta}) + (bias(\hat{\theta}))^2$ .

Hypothes Testing

Hypothesi

Type I and Type

Statistical

Significand

Confide Interval

Example of Hypothesis Testin

Test of Cause

References

### Best Unbiased Estimator

Definition 1 (Best Linear Unbiased Estimator (BLUE))

If  $\sigma^2 < \infty$  the sample mean  $\bar{X}_n$  has the lowest variance among all linear unbiased estimators of  $\mu$ .

Hypothesis Testing

#### Hypothesis

Type I and Type error

Significa

Confiden Interval

Hypothesis Testin

Test of Causa Effect

Reference

- $\triangleright$  A point hypothesis is the statement that  $\theta$  equals a specific value  $\theta_0$ .
- ightharpoonup A common example is  $\theta$  measures the effect the proposed policy. A typical question is whether  $\theta=0$ , which can be written as  $\theta_0=0$ .
- ▶ The **null hypothesis**, written as  $H_0: \theta = \theta_0$ , is the restriction  $\theta = \theta_0$ .
- ▶ The **alternative hypothesis**, written as  $H_A$ :  $\theta \neq \theta_0$ , is the set  $\{\theta \in \Theta : \theta \neq \theta_0\}$ .
  - ♦ **One-sided** hypothesis:  $H_A$ :  $\theta > \theta_0$ .
  - ♦ **Two-sided** hypothesis:  $H_A$ :  $\theta \neq \theta_0$ .

Hao

Estimator

Hypothe

#### Hypothesis

error Statistical

Statistical Significanc

Interval

Hypothesis Testing

Test of Causa Effect

Reference

# Acceptance and Rejection

- ▶ A hypothesis test is a decision based on data. We can either fail to reject the null hypothesis or reject the alternative hypothesis.
- An alternative way to express a decision rule is to construct a real-valued function of the data called a **test statistics**

$$T = T(X_1, \ldots, X_n)$$

together with a **critical region** C.

- A hypothesis can be expressed as
  - ♦ Accept  $H_0$  if  $T \in C$ .
  - ⋄ Reject  $H_0$  if  $T \notin C$ .

Note: "Accept"  $H_0$  does not mean  $H_0$  is true.

#### Estimator

BLUE

Hypothes Testing

#### Hypothesis

Type I and Ty

Significan

Confider Interval

Example of Hypothesis Testi

Test of Caus Effect

Reference

# Example - Hypothesis Testing I

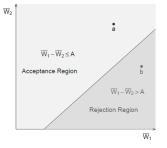
### Consider the following examples:

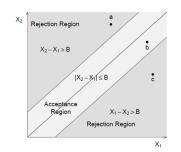
- ightharpoonup 2n adults who were raised in similar settings, n attended early childhood education. Let  $\bar{W}_1$  be the average wage in the early childhood education group, and let  $\bar{W}_2$  be the average wage in the remaining sample. Null hypothesis  $H_0: \bar{W}_1 > \bar{W}_2$ .
- $\triangleright$  You ride each bus once and record the time it takes to travel from home to the university. Let  $X_1$  and  $X_2$  be the two recorded travel times. You adopt the following decision rule: If the absolute difference in travel times is greater than B minutes you will reject the hypothesis that the average travel times are the same, otherwise you will accept the hypothesis.

Hao

Hypothesis

# Example - Hypothesis Testing II





(a) Early Childhood Education Example

(b) Bus Travel Example

Significar

Confider Interval

Example of Hypothesis Testir

Test of Caus Effect

Reference

# Type I and Type II error

- ▶ A false rejection of the null hypothesis is a **Type I error**.
- ▶ A false acceptance of the alternative hypothesis is a **Type II** error.

	Accept $H_0$	Reject $H_0$
$H_0$ true	Correct Decision	Type I Error
$H_1$ true	Type II Error	Correct Decision

stimator

Hypothesis Testing

Type I and Type II error

Significan

Interval Example of

Hypothesis Testin

Test of Caus Effect

Reference

The **power function** of a hypothesis test is the probability of rejection

$$\pi(F) = \mathbb{P}(\text{Reject } H_0|F) = \mathbb{P}(T \in C|F).$$

- ▶ The size of a hypothesis test is the probability of a Type I error.
- ▶ The power of a hypothesis test is the complement of the probability of the Type II error.

BLUE

Hypothesis Testing Hypothesis

Statistical Significance

Confiden

Example of Hypothesis Testin

Test of Causa Effect

Reference

Suppose we use a test which has the form: "Reject  $H_0$  when T > c", how to report the results? A simple choice is to report the "**p-value**", which is

$$p=1-G_0(T),$$

where  $G_0(\cdot)$  is the null sampling distribution. If  $G_0(c)=\alpha$ , the decision is identical to "Reject  $H_0$  if  $p<\alpha$ ". Reporting p-values is especially useful when T has complicated or unusual distribution.

# Computing p-value

- ▶ Suppose we are interested in testing the null hypothesis in  $H_0: \mathbb{E}(X) = \mu$  with the alternative hypothesis  $H_A: \mathbb{E}(X) \neq \mu$ .
  - Two-sided test.
- $\triangleright$  We observe the realization of  $X_1, \ldots, X_n$  as  $x_1, \ldots, x_n$ .
- $\triangleright$  Note that  $\bar{X}$  is a function of  $X_1, \ldots, X_n$ , which are i.i.d., therefore is a random variable.

$$\diamond \text{ Let } \bar{X} = \frac{1}{n} \sum_{i=1}^{n} X_{i}$$

$$\diamond \text{ and } \bar{x} = \frac{1}{n} \sum_{i=1}^{n} X_{i}.$$

▷ Under  $H_0$ , the distribution of  $\frac{\bar{X} - \mathbb{E}(X)}{\sigma_{\bar{X}}} \sim N(0, 1)(\mathsf{CLT})$ .

$$ho \; 
ho = 1 - \mathbb{P}\left( |rac{ar{X} - \mathbb{E}(X)}{\sigma_{ar{X}}}| < |rac{ar{x} - \mathbb{E}(X)}{\sigma_{ar{X}}}| 
ight).$$

Issue:  $\sigma_{\bar{x}}$  unknown.

Estimator

Hypothes Testing

Type I and Ty

Statistical Significance

Confide Interval

Hypothesis Testin

Test of Causa Effect

Reference

If the following assumptions hold:

- 1.  $X_1, ..., X_n$  are i.i.d.
- 2.  $\mathbb{E}(X_i) < \infty$ .

The sample variance is computed

$$\bar{s}^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2.$$

- $\triangleright \mu$  is unknown, need to be estimated.
- $\triangleright \mathbb{E}((X-\bar{X})^2) \to \tfrac{n-1}{n}\sigma.$
- ▶ The sample variance is a consistent estimator of the population variance.

Hypothesis Testing

Hypothesis

Statistical Significance

Confiden

Example of Hypothesis Testin

Test of Causa Effect

Reference

The standardized sample average can be constructed using

$$t = \frac{\bar{X} - \mu}{\sqrt{\bar{s}^2}}.$$

With the sample of  $x_1, \ldots, x_n$ , we can compute the sample t-statistic  $t^{sample}$ .

The p-value is given by

$$p$$
-value =  $2\Phi(-|t^{sample}|)$ .

Estimato

Hypothesis Testing

Type I and T error

Statistical Significance

Confider

Example of Hypothesis Testing

Test of Causa Effect

Reference

When construct hypothesis test, can fix a significance level.

- ho  $\alpha$ -significance test means the tolerance to make Type I error is  $\alpha$ .
- $\triangleright \alpha$  is referred to as the **size** of the test.

Suppose the two-sided test has the **significance level** of  $\alpha$ , the rule is "Reject  $H_0$  if  $|t^{sample}| > 1 - \Phi^{-1}(\alpha/2)$ ".

$$\alpha = 1\%, 1 - \Phi^{-1}(\alpha/2) = 2.58.$$

$$\alpha = 5\%, 1 - \Phi^{-1}(\alpha/2) = 1.96.$$

$$\alpha = 10\%, 1 - \Phi^{-1}(\alpha/2) = 1.64.$$

Statistical Significance

Interval Example of

Test of Causa

Reference

### Confidence Interval I

We are interested in learning a parameter of interest  $\theta$  from i.i.d. random sample of  $X_1, \ldots, X_n$ .

- ▶ With random sampling error, it's impossible to learn the exact value of the parameter of interest.
- ightharpoonup Construct a **confidence set**: the parameter of interested has  $1-\alpha$  probability to fall into the confidence set.
- ▶ The coverage probability of the interval estimator is the probability that the random interval contains the true parameter.
  - $\diamond$  An  $1-\alpha$  asymptotic confidence interval for a parameter has the asymptotic coverage probability  $1-\alpha$ .

Type I and Type

Statistical

Significano

Confidence Interval

Example of Hypothesis Testir

Test of Caus

Reference

### Confidence Interval II

A normal-based  $1-\alpha$  confidence interval is

$$CI = [\hat{\theta} - Z_{1-\alpha/2}s(\hat{\theta}), \hat{\theta} + Z_{1-\alpha/2}s(\hat{\theta})],$$

where  $\hat{\theta}$  is the estimator for  $\theta$  and  $se(\hat{\theta})$  is the estimated standard deviation.  $Z_{1-\alpha/2}$  is the  $1-\alpha/2$ -quantile of a normal distribution.

### Test for Difference Between Two Groups I

Suppose we observe the i.i.d sample  $W_1, \ldots, W_{n_1}, \ldots, W_n$ .

- ightharpoonup Sample  $W_1,\ldots,W_{n_1}$  are the monthly wage of graduates with master's degree, let  $\mu_1$  denote the population mean and  $\sigma_1^2$  the population variance of group 1.
- Sample  $W_{n_1+1}, \ldots, W_n$  are the monthly wage of graduates with bachelor's degree, let  $\mu_2$  denote the population mean and  $\sigma_2^2$  the population variance of group 2.
- ▶ Let  $n_2 = n n_1$ .
- $\vdash H_0: \mu_1 \mu_2 > d_0, H_1: \mu_1 \mu_2 \leq d_0$ , with significance level of  $\alpha$ .

Statistica

Confider

Interval Example of

Hypothesis Testing

Test of Causa Effect

Reference:

# Test for Difference Between Two Groups

- $\triangleright$  The parameter of interest is  $\theta = \mu_1 \mu_2$ .
- ▶ Let  $\overline{W}_1$  and  $\overline{W}_2$  be the estimated sample mean and  $s_1^2$  and  $s_2^2$  be the estimated sample variance for group 1 and group 2.
- ho The standard error of  $\hat{ heta}=ar{W}_1-ar{W}_2$  is  $se(\hat{ heta})=\sqrt{rac{s_1^2}{n_1}+rac{s_2^2}{n_2}}$ .
- $\triangleright$  We construct the t-statistic as  $t = \frac{\hat{\theta} d_0}{\sec(\hat{\theta})}$ .
- $\triangleright$  We reject  $H_0$  if  $t > Z_{1-\alpha}$ .

stimators

Estimators

Hypothesis

Hypothesi:

Type I and Type

Statistical

Significanc

Interval
Example of

Hypothesis Testin

Test of Causa Effect

References

Stock, J. H. and Watson, M. W. (2020). *Introduction to econometrics*, volume 4. Pearson New York.