

Homework 3

MSSC 6010- Computational Probability

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Question 1. (3.6.39) From book. Let the random variable X be the sum of the numbers on two fair dice. Find an upper bound on $\mathbb{P}(|X - 7| \geq 4)$ using Chebyshev's Inequality as well as the exact probability for $\mathbb{P}(|X - 7| \geq 4)$.

```
X = c(2,3,4,5,6,7,8,9,10,11,12)
X = rbind(X,c(1/36,2/36,3/36,4/36,5/36,6/36,5/36,4/36,3/36,2/36,1/36))
# Round for the output
X = signif(X,digits=3)
X
```

```
      [,1] [,2] [,3] [,4] [,5] [,6] [,7] [,8] [,9] [,10] [,11]
X 2.0000 3.0000 4.0000 5.000 6.000 7.000 8.000 9.000 10.0000 11.0000 12.0000
    0.0278 0.0556 0.0833 0.111 0.139 0.167 0.139 0.111 0.0833 0.0556 0.0278
```

```
# Go back to the original
X = c(2,3,4,5,6,7,8,9,10,11,12)
X = rbind(X,c(1/36,2/36,3/36,4/36,5/36,6/36,5/36,4/36,3/36,2/36,1/36))
```

Then we need μ and σ^2 in order to be able to use Chebyshev's Inequality. We use $\mu = \sum x\mathbb{P}(x)$ and $\sigma^2 = E[X^2] - \mu^2$

```
mu = sum(X[1,]*X[2,])
sigma_2 = sum(X[1,]**2*X[2,])-mu2
print(paste('mu = ', mu, ', sigma2 = ', sigma_2))
```

```
[1] "mu = 7 , sigma^2 = 5.83333333333333"
```

```
# Calculate Chebychev's RHS
my_prob = sigma_2/(4**2)
```

Plugging to Chebychev's Inequality

$$\begin{aligned}\mathbb{P}(|X - \mu| \geq k) &\leq \frac{\sigma^2}{k^2} \\ \mathbb{P}(|X - 7| \geq 4) &\leq \frac{5.83333^2}{4^2} \\ \mathbb{P}(|X - 7| \geq 4) &\leq 0.36458333\end{aligned}$$

Now, the exact probability is $\mathbb{P}(X = 2) + \mathbb{P}(X = 3) + \mathbb{P}(X = 11) + \mathbb{P}(X = 12) = 0.166666$

```
(exact_prob = X[2,1]+X[2,2]+X[2,10]+X[2,11])
```

0.1666667

```
# Clear environment  
rm(list = ls(all=TRUE))
```

Question 2. (3.6.48) From book. Consider the random variable X , which takes the values 1, 2, 3, and 4 with probabilities 0.2, 0.3, 0.1, and 0.4, respectively. Calculate $E[X]$, $\frac{1}{E[X]}$, $E\left[\frac{1}{X}\right]$, $E[X^2]$, and $E[X]^2$, and check empirically that $E[X]^2 \neq E[X^2]$ and $\frac{1}{E[X]} \neq E\left[\frac{1}{X}\right]$

```
X = c(1, 2, 3, 4)
X = rbind(X, c(0.2, 0.3, 0.1, 0.4))
X
```

```
      [,1] [,2] [,3] [,4]
X 1.0  2.0  3.0  4.0
    0.2  0.3  0.1  0.4
```

Now we can compute all the expectations

```
E_X = sum(X[,1]*X[,2])
one_div_E_X = 1/E_X
E_one_div_X = sum(1/X[,1]*X[,2])
E_X_sqr = sum(X[,1]**2*X[,2])
E_X_qnty_sqr = E_X2
cat("E[X]**2 = ", E_X_qnty_sqr, ", E[X**2] = ",
    E_X_sqr, ", E[X]**2 == E[X**2] is ", E_X_qnty_sqr == E_X_sqr, "\n")
```

```
E[X]**2 = 7.29 , E[X**2] = 8.7 , E[X]**2 == E[X**2] is FALSE
```

```
cat("And E[1/X] = ", E_one_div_X, ", 1/E[X] = ",
    one_div_E_X, ", E[1/X] == 1/E[X] is ", E_one_div_X==one_div_E_X)
```

```
And E[1/X] = 0.4833333 , 1/E[X] = 0.3703704 , E[1/X] == 1/E[X] is FALSE
```

Therefore $E[X]^2 \neq E[X^2]$ and $\frac{1}{E[X]} \neq E\left[\frac{1}{X}\right]$.

Question 3. (3.6.50) From book. Find the values of k such that the following functions are probability density functions.

- (a) $f(x) = \frac{kx^4}{5}, 0 < x < 1$

$$\begin{aligned} 1 &= \int_{-\infty}^{\infty} f(x) dx \\ &= \int_0^1 \frac{kx^4}{5} dx \\ &= \frac{kx^5}{25} \Big|_0^1 \\ &= k \left(\frac{x^5}{25} \right) \Big|_0^1 \\ &= k \left(\frac{1}{25} - 0 \right) \\ 1 &= \frac{k}{25} \\ k &= 25 \end{aligned}$$

Question 4. (3.6.52) From book. Consider an experiment where two dice are rolled. Let the random variable X equal the sum of the two dice and the random variable Y be the difference of the two dice.

```
# Clear environment
rm(list = ls(all=TRUE))
```

First let's define X and Y

```
X = c(2,3,4,5,6,7,8,9,10,11,12)
X = rbind(X,c(1/36,2/36,3/36,4/36,5/36,6/36,5/36,4/36,3/36,2/36,1/36))
Y = c(-5,-4,-3,-2,-1,0,1,2,3,4,5)
Y = rbind(Y,c(1/36,2/36,3/36,4/36,5/36,6/36,5/36,4/36,3/36,2/36,1/36))

# Round for the output
X = signif(X,digits=3)
X
```

```
      [,1] [,2] [,3] [,4] [,5] [,6] [,7] [,8] [,9] [,10] [,11]
X 2.0000 3.0000 4.0000 5.000 6.000 7.000 8.000 9.000 10.0000 11.0000 12.0000
    0.0278 0.0556 0.0833 0.111 0.139 0.167 0.139 0.111 0.0833 0.0556 0.0278
```

```
# Go back to the original
X = c(2,3,4,5,6,7,8,9,10,11,12)
X = rbind(X,c(1/36,2/36,3/36,4/36,5/36,6/36,5/36,4/36,3/36,2/36,1/36))

Y = signif(Y,digits=3)
Y
```

```
      [,1] [,2] [,3] [,4] [,5] [,6] [,7] [,8] [,9] [,10] [,11]
Y -5.0000 -4.0000 -3.0000 -2.000 -1.000 0.000 1.000 2.000 3.0000 4.0000 5.0000
    0.0278 0.0556 0.0833 0.111 0.139 0.167 0.139 0.111 0.0833 0.0556 0.0278
```

```
# Go back to the original
Y = c(-5,-4,-3,-2,-1,0,1,2,3,4,5)
Y = rbind(Y,c(1/36,2/36,3/36,4/36,5/36,6/36,5/36,4/36,3/36,2/36,1/36))
```

- (a) Find the mean of X

```
mu = sum(X[,1]*X[,2])
cat("mean = ", mu)
```

mean = 7

- (b) Find the variance of X

```
sigma_sqr = sum(X[1,]**2*X[2,])-mu2  
cat("variance = ", sigma_sqr)
```

variance = 5.833333

- (c) Find the skewness of X

```
skewness = sum((X[1,]-mu)**3*X[2,])/(sqrt(sigma_sqr))**3  
cat("skewness = ", skewness)
```

skewness = 0

- (d) Find the mean of Y

```
mu_y = sum(Y[1,]*Y[2,])  
cat("mean_y = ", mu_y)
```

mean_y = 0

- (e) Find the variance of Y

```
sigma_sqr_y = sum(Y[1,]**2*Y[2,])-mu_y2  
cat("variance_y = ", sigma_sqr_y)
```

variance_y = 5.833333

- (f) Find the skewness of Y

```
skewness_y = sum((Y[1,]-mu_y)**3*Y[2,])/(sqrt(sigma_sqr_y))**3  
cat("skewness_y = ", skewness_y)
```

skewness_y = 0

Question 5. (3.6.58) From book. Consider the probability density function

$$f(x) = \frac{1}{36} x e^{-\frac{x}{6}}, x > 0$$

Derive the moment generating function, and calculate the mean and the variance.

$$\begin{aligned}
M_X(t) &= E[e^{tX}] \\
&= \int_{-\infty}^{\infty} e^{tX} \left(\frac{1}{36} x e^{-\frac{x}{6}} \right) dx, x > 0 \\
&= \frac{1}{36} \int_0^{\infty} \left(x e^{-x(\frac{1}{6}-t)} \right) dx \\
&= \frac{1}{36} \left[x \left(\frac{1}{t - \frac{1}{6}} e^{-x(\frac{1}{6}-t)} \right) - \int_0^{\infty} \left(\frac{1}{t - \frac{1}{6}} e^{-x(\frac{1}{6}-t)} \right) dx \right] \\
&= \frac{1}{36} \left[x \left(\frac{1}{t - \frac{1}{6}} e^{-x(\frac{1}{6}-t)} \right) - \left(\frac{1}{(t - \frac{1}{6})^2} e^{-x(\frac{1}{6}-t)} \right) \right] \Bigg|_0^{\infty} \\
&= \lim_{b \rightarrow \infty} \frac{1}{36} \left[x \left(\frac{1}{t - \frac{1}{6}} e^{-x(\frac{1}{6}-t)} \right) - \left(\frac{1}{(t - \frac{1}{6})^2} e^{-x(\frac{1}{6}-t)} \right) \right] \Bigg|_0^b \\
&= \frac{1}{36} \lim_{b \rightarrow \infty} \left[b \left(\frac{1}{t - \frac{1}{6}} e^{-b(\frac{1}{6}-t)} \right) - \left(\frac{1}{(t - \frac{1}{6})^2} e^{-b(\frac{1}{6}-t)} \right) - \left(0 - \left(\frac{1}{(t - \frac{1}{6})^2} e^0 \right) \right) \right] \\
&= \frac{1}{36} \lim_{b \rightarrow \infty} \left[b \left(\frac{1}{t - \frac{1}{6}} e^{-b(\frac{1}{6}-t)} \right) - \left(\frac{1}{(t - \frac{1}{6})^2} e^{-b(\frac{1}{6}-t)} \right) - \left(-\frac{1}{(t - \frac{1}{6})^2} \right) \right] \\
&= \frac{1}{36} \left(\frac{1}{(t - \frac{1}{6})^2} \right) + \lim_{b \rightarrow \infty} \left[b \left(\frac{1}{t - \frac{1}{6}} e^{-b(\frac{1}{6}-t)} \right) - \left(\frac{1}{(t - \frac{1}{6})^2} e^{-b(\frac{1}{6}-t)} \right) \right] \\
M_X(t) &= \frac{1}{36} \left(\frac{1}{(t - \frac{1}{6})^2} \right)
\end{aligned}$$

Since

$$\begin{aligned}
& \lim_{b \rightarrow \infty} \left[b \left(\frac{1}{t - \frac{1}{6}} e^{-b(\frac{1}{6} - t)} \right) - \left(\frac{1}{(t - \frac{1}{6})^2} e^{-b(\frac{1}{6} - t)} \right) \right] = 0 \\
& \lim_{b \rightarrow \infty} \left[b \left(\frac{1}{t - \frac{1}{6}} e^{-b(\frac{1}{6} - t)} \right) \right] - \lim_{b \rightarrow \infty} \left[\left(\frac{1}{(t - \frac{1}{6})^2} e^{-b(\frac{1}{6} - t)} \right) \right] = \\
& \lim_{b \rightarrow \infty} \left[b \left(\frac{1}{t - \frac{1}{6}} e^{-b(\frac{1}{6} - t)} \right) \right] - \left[\frac{1}{\infty} \right] = \\
& \lim_{b \rightarrow \infty} \left[b \left(\frac{1}{t - \frac{1}{6}} e^{-b(\frac{1}{6} - t)} \right) \right] = \\
& \frac{1}{t - \frac{1}{6}} \lim_{b \rightarrow \infty} \left[b \left(e^{-b(\frac{1}{6} - t)} \right) \right] = \\
& \frac{1}{t - \frac{1}{6}} \lim_{b \rightarrow \infty} \left[\frac{b}{e^{-b(\frac{1}{6} - t)}} \right] = \\
& \frac{1}{t - \frac{1}{6}} \lim_{b \rightarrow \infty} \left[\frac{1}{(-(\frac{1}{6} - t)e^{-b(\frac{1}{6} - t)})} \right] = 0 \text{ by L'Hopital} \\
& \frac{1}{\infty} = 0
\end{aligned}$$

Finally, we can compute the mean and variance.

The mean is the first derivative evaluated at $t = 0$

$$\begin{aligned}
E[X] &= \frac{d}{dt} \frac{1}{36(t - 1/6)^2} \\
&= \frac{-2}{36} (t - \frac{1}{6})^{-3} \\
&= \frac{-2}{36} (-\frac{1}{6})^{-3} \\
&= 12
\end{aligned}$$

And the variance is the second moment, which is the second derivative evaluated at 0, minus $E[X]^2$

$$\begin{aligned}
var(X) &= E[X^2] - E[X]^2 \\
&= \frac{6}{36} (t - \frac{1}{6})^{-4} - 12^2 \\
&= \frac{6}{36} (\frac{1}{6})^{-4} - 12^2 \\
&= 72
\end{aligned}$$

Question 6. (3.6.60) From book. Prove that if a and b are real-valued constants, then

- (1) $M_{X+a}(t) = E[e^{(X+a)t}] = e^{at} M_X(t).$

$$\begin{aligned}
 M_{X+a}(t) &= E[e^{(X+a)t}] \\
 &= \int_{-\infty}^{\infty} e^{(X+a)t} f(x) dx \\
 &= \int_{-\infty}^{\infty} e^{Xt+at} f(x) dx \\
 &= \int_{-\infty}^{\infty} e^{Xt} e^{at} f(x) dx, \text{ and } e^{at} \text{ is a constant} \\
 &= e^{at} \int_{-\infty}^{\infty} e^{Xt} f(x) dx \\
 &= e^{at} M_X(t)
 \end{aligned}$$

- (2) $M_{bX}(t) = E[e^{bXt}] = M_X(bt).$

$$\begin{aligned}
 M_{bX}(t) &= E[e^{bXt}] \\
 &= \int_{-\infty}^{\infty} e^{bXt} f(x) dx \\
 &= \int_{-\infty}^{\infty} e^{(bt)X} f(x) dx \\
 &= M_X(bt)
 \end{aligned}$$

- (3) $M_{\frac{X+a}{b}}(t) = E[e^{(\frac{X+a}{b})t}] = e^{\frac{a}{b}t} M_X(\frac{t}{b}).$

$$\begin{aligned}
 M_{\frac{X+a}{b}}(t) &= E[e^{(\frac{X+a}{b})t}] \\
 &= \int_{-\infty}^{\infty} e^{(\frac{X+a}{b})t} f(x) dx \\
 &= \int_{-\infty}^{\infty} e^{\frac{Xt}{b}} e^{\frac{at}{b}} f(x) dx \\
 &= e^{\frac{at}{b}} \int_{-\infty}^{\infty} e^{X\frac{t}{b}} f(x) dx \\
 &= e^{\frac{at}{b}} M_X(\frac{t}{b})
 \end{aligned}$$