Homework I: DATA620013

Advanced Statistical Learning

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1 Problem 1

Assume that X is not random, $y = X\beta + \varepsilon$

$$E(d^T y) = E(d^T(X\beta + \epsilon)) = d^T X\beta$$

As is given unbiased, $Ed^T y = Ec^T y = c^T\beta$, so we have $d^T X = c^T$

$$\begin{aligned}
Var(\hat{\beta}) &= E(\hat{\beta} - \beta)(\hat{\beta} - \beta)^T \\
&= E((X^T X)^{-1} X^T (X\beta + \varepsilon) - \beta)((X^T X)^{-1} X^T (X\beta + \varepsilon) - \beta)^T \\
&= E((X^T X)^{-1} X^T \varepsilon))((X^T X)^{-1} X^T \epsilon))^T \\
&= E((X^T X)^{-1} X^T \varepsilon\varepsilon^T X (X^T X)^{-1}) \\
&= (X^T X)^{-1}
\end{aligned}$$

so that, $Var(c^T\hat{\beta}) = c^T(X^T X)^{-1} c = d^T X (X^T X)^{-1} X^T d$
As the same, $Var(d^T y) = Var(d^T(X\beta + \varepsilon)) = E(d^T(X\beta + \varepsilon) - d^T X\beta)(d^T(X\beta + \varepsilon) - d^T X\beta)^T = d^T d$
We want to show that, $\forall d,\ s.t. Ed^T y = c^T\beta$

$$Var(c^T\hat{\beta}) \leq Var(d^T y) \iff d^T d - d^T X (X^T X)^{-1} X^T d \geq 0 \iff I - X(X^T X)^{-1} X^T \succeq 0$$

To prove that $I - X(X^T X)^{-1} X^T \succeq 0$, we set $I - X(X^T X)^{-1} X^T = M$

$$M^2 = (I - X(X^T X)^{-1} X^T)(I - X(X^T X)^{-1} X^T) = M$$

so that, $I - X(X^T X)^{-1} X^T = M = M^2 = M^T M$, $I - X(X^T X)^{-1} X^T$ is a positive semidefinite matrix. The proof is completed.

2 Problem 2

$$\theta = \{\mu_k, \Sigma\}, \quad k = 1, \cdots, K$$

$$p_k(x_i;\theta) = Pr(G = k|X = x_i;\theta) = \frac{f(x|G = k)Pr(G = k)}{Pr(X = x)} \propto f(x|G = k)Pr(G = k)$$

the likelihood of the data is

$$\mathcal{L}(\theta; x) = \prod_{i=1}^{N} p_k(x_i;\theta) = \prod_{i=1}^{N} \frac{f(x|G = k)Pr(G = k)}{Pr(X = x)}$$

the log likelihood of the data is

$$log(\mathcal{L}(\theta; x)) = \sum_{i=1}^{N} log(p_{k_i}(x_i;\theta)) \propto \sum_{i=1}^{N} log(f_{k_i}(x_i)\pi_{k_i}) = \sum_{i=1}^{N} log(f_{k_i}(x_i)) + log\pi_{k_i}$$

where x_i means the i_{th} sample, k_i means the class of the i_{th} sample and

$$f_k(x) = \frac{1}{(2\pi)^{\frac{P}{2}}|\Sigma|^{\frac{1}{2}}} e^{-\frac{1}{2}(x-\mu_k)^T\Sigma^{-1}(x-\mu_k)}$$

(1) $\hat{\mu_k}$

Suppose that the k_{th} class has n_k samples

$$\frac{\partial log(\mathcal{L}(\theta;x))}{\partial \mu_k} = \frac{\partial \sum_{i=1}^{N} log(f_{k_i}(x_i))}{\partial \mu_k} = \frac{\partial \sum_{j=1}^{n_k} log(f_k(x_j))}{\partial \mu_k} = \sum_{j=1}^{n_k} \frac{\frac{\partial f_k(x_j)}{\partial \mu_k}}{f_k(x_j)} \tag{1}$$

$$\frac{\partial f_k(x_j)}{\partial \mu_k} = f_k(x_j)(-\frac{1}{2}(\Sigma^{-1} + (\Sigma^{-1})^T)(\mu_k - x_j)) = f_k(x_j)(-\Sigma^{-1})(\mu_k - x_j)$$

so that (1) equals

$$\sum_{j=1}^{n_k}(-\Sigma^{-1})(\mu_k - x_j) \tag{2}$$

let (2) equals to 0, we have

$$\hat{\mu_k} = \frac{1}{n_k}\sum_{j=1}^{n_k} x_j$$

(2) $\hat{\Sigma}$

From the matrix derivative formula, we know $\mathrm{d}f = tr(\frac{\partial f^T}{\partial x}\mathrm{d}x)$ (*), and $d|X| = |X|\,\mathrm{tr}\left(X^{-1}dX\right)$ (**)

$$\arg\max_{\Sigma} log(\mathcal{L}(\theta;x)) = \arg\max_{\Sigma} \sum_{i=1}^{N} -\frac{1}{2}log|\Sigma| - \frac{1}{2}(x_i-\mu_{k_i})^T\Sigma^{-1}(x_i-\mu_{k_i}) = \arg\min_{\Sigma}\sum_{i=1}^{N} log|\Sigma| + (x_i-\mu_{k_i})^T\Sigma^{-1}(x_i-\mu_{k_i}) \tag{3}$$

the derivative of first item of (3) is

$$\sum_{i=1}^{N}\mathrm{d}log|\Sigma| = \sum_{i=1}^{N}|\Sigma|^{-1}\mathrm{d}|\Sigma| \overset{(**)}{=} \sum_{i=1}^{N} tr(\Sigma^{-1}\mathrm{d}\Sigma) = tr(N * \Sigma^{-1}\mathrm{d}\Sigma) \tag{4}$$

As the assumption above, all K classes and n_k samples about the k_{th} class, the second item of (3) also equals

$$\sum_{k=1}^{K}\sum_{j=1}^{n_k}(x_j - \mu_k)^T\Sigma^{-1}(x_j - \mu_k) \tag{5}$$

the derivative of second item of (3) is

$$\begin{aligned}
\sum_{k=1}^{K}\sum_{j=1}^{n_k}(x_j - \mu_k)^T\mathrm{d}\Sigma^{-1}(x_j - \mu_k) &= -\sum_{k=1}^{K}\sum_{j=1}^{n_k}(x_j - \mu_k)^T\Sigma^{-1}\mathrm{d}\Sigma\Sigma^{-1}(x_j - \mu_k) \\
&= tr(-\sum_{k=1}^{K}\sum_{j=1}^{n_k}(x_j - \mu_k)^T\Sigma^{-1}\mathrm{d}\Sigma\Sigma^{-1}(x_j - \mu_k)) \\
&= -\sum_{k=1}^{K}\sum_{j=1}^{n_k} tr((x_j - \mu_k)^T\Sigma^{-1}\mathrm{d}\Sigma\Sigma^{-1}(x_j - \mu_k)) \\
&= -\sum_{k=1}^{K}\sum_{j=1}^{n_k} tr(\Sigma^{-1}(x_j - \mu_k)(x_j - \mu_k)^T\Sigma^{-1}\mathrm{d}\Sigma) \\
&= -\sum_{k=1}^{K} tr(\Sigma^{-1}\sum_{j=1}^{n_k}(x_j - \mu_k)(x_j - \mu_k)^T\Sigma^{-1}\mathrm{d}\Sigma)
\end{aligned} \tag{6}$$

let $S_k = \sum_{j=1}^{n_k}(x_j - \mu_k)(x_j - \mu_k)^T$, the (6) equals

$$-\sum_{k=1}^{K} tr(\Sigma^{-1}S_k\Sigma^{-1}\mathrm{d}\Sigma) \tag{7}$$

from (3), (4) and (7), we have

$$\mathrm{d}log(\mathcal{L}) = tr(N * \Sigma^{-1}\mathrm{d}\Sigma - \Sigma^{-1}S_k\Sigma^{-1}\mathrm{d}\Sigma) = tr((N * \Sigma^{-1} - \Sigma^{-1}\sum_{k=1}^{K} S_k\Sigma^{-1})\mathrm{d}\Sigma)$$

moreover, from (*), we have

$$\frac{\partial log(\mathcal{L}(\theta;x))}{\partial\Sigma} = N * \Sigma^{-1} - \Sigma^{-1}\sum_{k=1}^{K} S_k\Sigma^{-1} \tag{8}$$

let (8) equals 0, we have

$$\hat{\Sigma} = \frac{1}{N}\sum_{k=1}^{K} S_k = \sum_{k=1}^{K}\frac{n_k - 1}{N}(\frac{1}{n_k - 1}S_k)$$

$$S_k = \sum_{j=1}^{n_k}(x_j - \hat{\mu_k})(x_j - \hat{\mu_k})^T$$

the $\frac{1}{n_k-1}S_k$ is the covariance matrix of the k_{th} class, so the weight of the pooled covariance estimate is $\frac{n_k-1}{N}$

3 Problem 3

4 Problem 4

$$p(\beta|x) = \frac{p(x|\beta)p(\beta)}{p(x)} \propto p(x|\beta)p(\beta)$$

$$\arg\max_{\beta} p(\beta|x) = \arg\max_{\beta} p(x|\beta)p(\beta)$$

let $\Pr(G = 1 \mid X = x, \beta) = \pi(x)$, so $\Pr(G = 0 \mid X = x, \beta) = 1 - \pi(x)$

$$\begin{aligned}
\arg\max_{\beta} p(x|\beta)p(\beta) &= \arg\max_{\beta}\prod_{i=1}^{n} p(x_i|\beta)p(\beta) \\
&= \arg\max_{\beta}\prod_{i=1}^{n} \pi(x_i)^{y_i}[1 - \pi(x_i)]^{(1-y_i)}p(\beta) \\
&= \arg\max_{\beta}\sum_{i=1}^{n} y_i log\pi(x_i) + (1 - y_i)log(1 - \pi(x_i)) + logp(\beta) \\
&= \arg\max_{\beta}\sum_{i=1}^{n} y_i log\pi(x_i) + (1 - y_i)log(1 - \pi(x_i))
\end{aligned}$$

so the loss function can be supposed to

$$\mathcal{L}(\beta) = -\sum_{i=1}^{n} y_i log\pi(x_i) + (1-y_i)log(1-\pi(x_i)) = -\sum_{i=1}^{n} y_i log\frac{\pi(x_i)}{1 - \pi(x_i)} + log(1-\pi(x_i)) = -\sum_{i=1}^{n}[y_i(\beta^T x_i) - log(1+e^{\beta^T x_i})]$$