

**10.2** The labor supply of married women has been a subject of a great deal of economic research. Consider the following supply equation specification

$$HOURS = \beta_1 + \beta_2 WAGE + \beta_3 EDUC + \beta_4 AGE + \beta_5 KIDS6 + \beta_6 NWIFEINC + e$$

where  $HOURS$  is the supply of labor,  $WAGE$  is hourly wage,  $EDUC$  is years of education,  $KIDS6$  is the number of children in the household who are less than 6 years old, and  $NWIFEINC$  is household income from sources other than the wife's employment.

- a. Discuss the signs you expect for each of the coefficients.
- b. Explain why this supply equation cannot be consistently estimated by OLS regression.
- c. Suppose we consider the woman's labor market experience  $EXPER$  and its square,  $EXPER^2$ , to be instruments for  $WAGE$ . Explain how these variables satisfy the logic of instrumental variables.
- d. Is the supply equation identified? Explain.
- e. Describe the steps [not a computer command] you would take to obtain IV/2SLS estimates.

a.  $\beta_2 (WAGE)$  → 正號，工資提高，激励提供更多勞動。

$\beta_3 (EDUC)$  → 不確定，教育程度越高 → 能力強有意願進入職場 → 增加勞動供給。

→ 效率高，對時間彈性要求高，可能移入其他地市 → 減少勞動供給。

$\beta_4 (AGE)$  → 不確定，年齡增加，可能增加經驗(提升供給)，也可能因為健康(降低供給)

$\beta_5 (KIDS6)$  → 負號，家中兒童越多，負擔重，供給下降。

$\beta_6 (NWIFEINC)$  → 負號，其他來源的收入高，減少女性進入勞動市場的誘因。

b. Endogeneity 內生性問題。

$HOURS \longleftrightarrow WAGE$  都由供需求決定，會造成估計失效

c. Instrumental variable

→ ① Relevance :  $EXPER$ 、 $EXPER^2$  和  $WAGE$  通常顯著相關。

② Exogeneity :  $EXPER$ 、 $EXPER^2$  透過  $WAGE$  影響  $HOURS$ 。

所以和誤差項不相關

d. Yes, 只有一個內生變數 (WAGE) 且至少有一個工具變數可使用

e.  $\text{WAGE} = \gamma_1 + \gamma_2 \text{EDUC} + \gamma_3 \text{AGE} + \gamma_4 \text{KIDS16} + \gamma_5 \text{KIDS18} + \gamma_6 \text{NWIFEINC} + \theta_1 \text{EXPER} + \theta_2 \text{EXPER}^2 + u$

先對上面的式子回歸，得到  $\hat{\text{WAGE}}$

將原本的 WAGE 用  $\hat{\text{WAGE}}$  替換，再用 OLS 估計。

10.3 In the regression model  $y = \beta_1 + \beta_2 x + e$ , assume  $x$  is endogenous and that  $z$  is a valid instrument.

In Section 10.3.5, we saw that  $\beta_2 = \text{cov}(z, y)/\text{cov}(z, x)$ .

- Divide the denominator of  $\beta_2 = \text{cov}(z, y)/\text{cov}(z, x)$  by  $\text{var}(z)$ . Show that  $\text{cov}(z, x)/\text{var}(z)$  is the coefficient of the simple regression with dependent variable  $x$  and explanatory variable  $z$ ,  $x = \gamma_1 + \theta_1 z + v$ . [Hint: See Section 10.2.1.] Note that this is the first-stage equation in two-stage least squares.
- Divide the numerator of  $\beta_2 = \text{cov}(z, y)/\text{cov}(z, x)$  by  $\text{var}(z)$ . Show that  $\text{cov}(z, y)/\text{var}(z)$  is the coefficient of a simple regression with dependent variable  $y$  and explanatory variable  $z$ ,  $y = \pi_0 + \pi_1 z + u$ . [Hint: See Section 10.2.1.]
- In the model  $y = \beta_1 + \beta_2 x + e$ , substitute for  $x$  using  $x = \gamma_1 + \theta_1 z + v$  and simplify to obtain  $y = \pi_0 + \pi_1 z + u$ . What are  $\pi_0$ ,  $\pi_1$ , and  $u$  in terms of the regression model parameters and error and the first-stage parameters and error? The regression you have obtained is a **reduced-form** equation.
- Show that  $\beta_2 = \pi_1/\theta_1$ .
- If  $\hat{\pi}_1$  and  $\hat{\theta}_1$  are the OLS estimators of  $\pi_1$  and  $\theta_1$ , show that  $\hat{\beta}_2 = \hat{\pi}_1/\hat{\theta}_1$  is a consistent estimator of  $\beta_2 = \pi_1/\theta_1$ . The estimator  $\hat{\beta}_2 = \hat{\pi}_1/\hat{\theta}_1$  is an **indirect least squares** estimator.

a.  $x = \gamma_1 + \theta_1 z + v \rightarrow \text{取期望值 } E(x) = \gamma_1 + \theta_1 E(z)$

$$\begin{aligned} x &= \gamma_1 + \theta_1 z + v \\ E(x) &= \gamma_1 + \theta_1 E(z) \end{aligned}$$

$$(x - E(x)) = \theta_1 (z - E(z)) + v$$

$$\text{同乘 } (z - E(z)) \Rightarrow (z - E(z))(x - E(x)) = \theta_1 (z - E(z))^2 + v(z - E(z))$$

$$\text{同取期望值: } E[(z - E(z))(x - E(x))] = \theta_1 E[(z - E(z))^2]$$

$$\Rightarrow \theta_1 = \frac{\text{cov}(z, x)}{\text{var}(z)}$$

$$b. y = \pi_0 + \pi_1 z + u \rightarrow \text{同 a. } E(y) = \pi_0 + \pi_1 E(z)$$

$$\begin{cases} y = \pi_0 + \pi_1 z + u \\ E(y) = \pi_0 + \pi_1 E(z) \end{cases}$$

$$y - E(y) = \pi_1(z - E(z)) + u$$

$$x(z - E(z)) \rightarrow E[(y - E(y))(z - E(z))] = \pi_1 E[(z - E(z))^2]$$

取期望值

$$\Rightarrow \pi_1 = \frac{\text{Cov}(z, y)}{\text{Var}(z)}$$

$$c. x = v_1 + \theta_1 z + v \text{ 代入 } y = \beta_1 + \beta_2 x + e$$

$$\begin{aligned} y &= \beta_1 + \beta_2(v_1 + \theta_1 z + v) + e \\ &= (\beta_1 + \beta_2 v_1) + \beta_2 \theta_1 z + (\beta_2 v + e) \end{aligned}$$

對照 reduced-form  $y = \pi_0 + \pi_1 z + u$

$$\Rightarrow \pi_0 = (\beta_1 + \beta_2 v_1)$$

$$\pi_1 = \beta_2 \theta_1$$

$$u = \beta_2 v + e$$

$$d. \pi_1 = \beta_2 \theta_1 \Rightarrow \beta_2 = \frac{\pi_1}{\theta_1}$$

$$e. \text{ by (a)} \quad \hat{\theta}_1 = \frac{\widehat{\text{Cov}}(z, y)}{\widehat{\text{Var}}(z)} = \frac{\sum (z_i - \bar{z})(y_i - \bar{y}) / N}{\sum (z_i - \bar{z})^2 / N} = \frac{\sum (z_i - \bar{z})(x_i - \bar{x})}{\sum (z_i - \bar{z})^2}$$

$$\text{by (b)} \quad \hat{\pi}_1 = \frac{\widehat{\text{Cov}}(z_i y)}{\widehat{\text{Var}}(z)} = \frac{\sum (z_i - \bar{z})(y_i - \bar{y}) / N}{\sum (z_i - \bar{z})^2 / N} = \frac{\sum (z_i - \bar{z})(y_i - \bar{y})}{\sum (z_i - \bar{z})^2}$$

$$\hat{\beta}_2 = \hat{\pi}_1 / \hat{\theta}_1 = \frac{\sum (z_i - \bar{z})(y_i - \bar{y})}{\sum (z_i - \bar{z})(x_i - \bar{x})} = \frac{\sum (z_i - \bar{z})(y_i - \bar{y}) / N}{\sum (z_i - \bar{z})(x_i - \bar{x}) / N} = \frac{\widehat{\text{Cov}}(z, y)}{\widehat{\text{Cov}}(z, x)}$$

$\rightarrow$  是 IV estimator.

consistent

$$\because \widehat{\text{Cov}}(z, y) \xrightarrow{P} \text{Cov}(z, y)$$

$$\widehat{\text{Cov}}(z, x) \xrightarrow{P} \text{Cov}(z, x)$$

$$\hat{\beta}_2 = \hat{\pi}_1 / \hat{\theta}_1 = \frac{\widehat{\text{Cov}}(z, y)}{\widehat{\text{Cov}}(z, x)} \xrightarrow{P} \frac{\text{Cov}(z, y)}{\text{Cov}(z, x)} = \beta_2$$