

The Ordinary Least Squares (OLS) Estimators

$$b_2 = \frac{\sum (x_i - \bar{x})(y_i - \bar{y})}{\sum (x_i - \bar{x})^2} \quad (2.7)$$

$$b_1 = \bar{y} - b_2 \bar{x} \quad (2.8)$$

where $\bar{y} = \sum y_i / N$ and $\bar{x} = \sum x_i / N$ are the sample means of the observations on y and x .

1. Let $K=2$, show that (b_1, b_2) in p.29 of slides in Ch5 reduces to the formula of (b_1, b_2) in (2.7)-(2.8)

$$\text{當 } K=2, \quad y_i = \beta_1 + \beta_2 x_i + e_i, \quad i=1, 2, \dots, N$$

$$\hat{X} = \begin{bmatrix} 1 & x_1 \\ 1 & x_2 \\ \vdots & \vdots \\ 1 & x_n \end{bmatrix} \Rightarrow \hat{X}'\hat{X} = \begin{bmatrix} 1 & 1 & 1 & \dots & 1 \\ x_1 & x_2 & \dots & x_n \end{bmatrix} \begin{bmatrix} 1 \\ x_1 \\ x_2 \\ \vdots \\ x_n \end{bmatrix} = \begin{bmatrix} n & \sum x_i \\ \sum x_i & \sum x_i^2 \end{bmatrix}$$

$$(\hat{X}'\hat{X})^{-1} = \frac{1}{n \sum x_i^2 - (\sum x_i)^2} \begin{bmatrix} \sum x_i^2 & -\sum x_i \\ -\sum x_i & n \end{bmatrix}$$

$$\hat{X}'\hat{Y} = \begin{bmatrix} 1 & \dots & 1 \\ x_1 & x_2 & \dots & x_n \end{bmatrix} \begin{bmatrix} y_1 \\ \vdots \\ y_n \end{bmatrix} = \begin{bmatrix} \sum y_i \\ \sum x_i y_i \end{bmatrix}$$

$$\begin{aligned} \text{則 } b = (\hat{X}'\hat{X})^{-1} \hat{X}'\hat{Y} &= \frac{1}{n \sum x_i^2 - (\sum x_i)^2} \begin{bmatrix} \sum x_i^2 & -\sum x_i \\ -\sum x_i & n \end{bmatrix} \begin{bmatrix} \sum y_i \\ \sum x_i y_i \end{bmatrix} \\ &= \frac{1}{n \sum x_i^2 - (\sum x_i)^2} \begin{bmatrix} \sum x_i^2 \sum y_i - \sum x_i \sum x_i y_i \\ -\sum x_i \sum y_i + n \sum x_i y_i \end{bmatrix} \end{aligned}$$

$$\text{而 } b_2 = \frac{-\sum x_i \sum y_i + n \sum x_i y_i}{n \sum x_i^2 - (\sum x_i)^2} = \frac{-n \bar{x} \bar{y} + n \sum x_i y_i}{n \sum x_i^2 - (n \bar{x})^2} = \frac{\sum (x_i - \bar{x})(y_i - \bar{y})}{\sum (x_i - \bar{x})^2}$$

$$b_1 = \frac{\sum x_i^2 \sum y_i - \sum x_i \sum x_i y_i}{n \sum x_i^2 - (\sum x_i)^2} = \frac{n \sum x_i^2 \bar{y} - n \bar{x} \sum x_i y_i}{n \sum x_i^2 - (n \bar{x})^2} = \frac{\bar{y} [\sum x_i^2 - n \bar{x}^2] - \bar{x} [\sum x_i y_i - n \bar{x} \bar{y}]}{\sum (x_i - \bar{x})^2} = \bar{y} - b_2 \bar{x}$$

$$\text{var}(b_1|\mathbf{x}) = \sigma^2 \left[\frac{\sum x_i^2}{N \sum (x_i - \bar{x})^2} \right] \quad (2.14)$$

$$\text{var}(b_2|\mathbf{x}) = \frac{\sigma^2}{\sum (x_i - \bar{x})^2} \quad (2.15)$$

$$\text{cov}(b_1, b_2|\mathbf{x}) = \sigma^2 \left[\frac{-\bar{x}}{\sum (x_i - \bar{x})^2} \right] \quad (2.16)$$

2. Let $K=2$, show that $\text{cov}(b_1, b_2)$ in p.30 of slides in Ch5 reduces to the formula of (b_1, b_2) in (2.14)-(2.16)

$$\begin{aligned} \text{Var}(b) &= \sigma^2 (X^T X)^{-1} = \frac{\sigma^2}{n \sum x_i^2 - (\sum x_i)^2} \begin{bmatrix} \sum x_i^2 & -\sum x_i \\ -\sum x_i & n \end{bmatrix} = \frac{\sigma^2}{n \sum x_i^2 - (n\bar{x})^2} \begin{bmatrix} \sum x_i^2 & -n\bar{x} \\ -n\bar{x} & n \end{bmatrix} \\ &= \begin{bmatrix} \frac{\sigma^2 \sum x_i^2}{n \sum x_i^2 - (n\bar{x})^2} & \frac{-\sigma^2 \bar{x}}{\sum x_i^2 - n\bar{x}^2} \\ \frac{-\sigma^2 \bar{x}}{\sum x_i^2 - n\bar{x}^2} & \frac{\sigma^2}{\sum x_i^2 - n\bar{x}^2} \end{bmatrix} \\ &= \begin{bmatrix} \frac{\sigma^2 \sum x_i^2}{n (\sum x_i - \bar{x})^2} & \frac{-\sigma^2 \bar{x}}{(\sum x_i - \bar{x})^2} \\ \frac{-\sigma^2 \bar{x}}{(\sum x_i - \bar{x})^2} & \frac{\sigma^2}{(\sum x_i - \bar{x})^2} \end{bmatrix} \\ &= \begin{bmatrix} \text{Var}(b_1|x) & \text{Cov}(b_1, b_2|x) \\ \text{Cov}(b_1, b_2|x) & \text{Var}(b_2|x) \end{bmatrix} \end{aligned}$$

5.3 Consider the following model that relates the percentage of a household's budget spent on alcohol $WALC$ to total expenditure $TOTEXP$, age of the household head AGE , and the number of children in the household NK .

$$WALC = \beta_1 + \beta_2 \ln(TOTEXP) + \beta_3 NK + \beta_4 AGE + e$$

This model was estimated using 1200 observations from London. An incomplete version of this output is provided in Table 5.6.

TABLE 5.6 Output for Exercise 5.3

Dependent Variable: WALC				
Included observations: 1200				
Variable	Coefficient	Std. Error	t-Statistic	Prob.
C	1.4515	2.2019		0.5099
$\ln(TOTEXP)$	2.7648		5.7103	0.0000
NK		0.3695	-3.9376	0.0001
AGE	-0.1503	0.0235	-6.4019	0.0000
R-squared		Mean dependent var		6.19434
S.E. of regression		S.D. dependent var		6.39547
Sum squared resid	46221.62			

- Fill in the following blank spaces that appear in this table.
 - The t -statistic for b_1 .
 - The standard error for b_2 .
 - The estimate b_3 .
 - R^2 .
 - $\hat{\sigma}$.
- Interpret each of the estimates b_2 , b_3 , and b_4 .
- Compute a 95% interval estimate for β_4 . What does this interval tell you?
- Are each of the coefficient estimates significant at a 5% level? Why?
- Test the hypothesis that the addition of an extra child decreases the mean budget share of alcohol by 2 percentage points against the alternative that the decrease is not equal to 2 percentage points. Use a 5% significance level.

a. b_1 的 t 值: $\frac{\hat{b}_1}{SE(\hat{b}_1)} = \frac{1.4515}{2.2019} \approx 0.6594$

a.ii $SE(b_2) = \frac{\hat{b}_2}{t} = \frac{2.7648}{5.7103} \approx 0.4841$

a.iii. $b_3 = t \cdot SE = -3.9376 \cdot 0.3695 \approx -1.4549$

a.iv. $R^2 = 1 - \frac{SSE}{SSR} \approx 1 - \left(\frac{46221.62}{(1200-1) \times (6.39547)^2} \right) \approx 0.0575$

a.v. $\hat{\sigma} \approx \sqrt{\frac{SSE}{n-k-1}} = \sqrt{\frac{46221.62}{1200-4}} \approx 0.6142$

- b. b_2 : 表示 $TOTEXP$ 对 $WALC$ 增加 1%, 酒精支出占比 增加约 2.76%
 b_3 : 表示每增一孩子, 酒精支出占比 平均 减少约 1.45%
 b_4 : 表示家庭主每增 1 岁, 酒精支出占比 平均 减少约 0.15%

c. $CI_{95\%} = \hat{b}_4 \pm t_{0.025, n-k-1} = -0.1503 \pm 1.96 \times 0.0235 = [-0.1964, -0.1042]$

d. $\begin{cases} b_2(0) < 0.05 \\ b_3(0.001) < 0.05 \\ b_4(0) < 0.05 \end{cases}$ 显著

e. $H_0: \beta_3 = -2, H_1: \beta_3 \neq -2$

$$t = \frac{\hat{b}_3 - (-2)}{SE(\hat{b}_3)} = \frac{-1.4549 + 2}{0.3695} \approx 1.4752$$

$$\text{而 } t^* = t_{0.025, 1196} \approx 1.96$$

$$\text{因 } 1.4752 < 1.96, \text{ 故 reject } H_0$$

5.23 The file *cocaine* contains 56 observations on variables related to sales of cocaine powder in northeastern California over the period 1984–1991. The data are a subset of those used in the study Caulkins, J. P. and R. Padman (1993), "Quantity Discounts and Quality Premia for Illicit Drugs," *Journal of the American Statistical Association*, 88, 748–757. The variables are

PRICE = price per gram in dollars for a cocaine sale
QUANT = number of grams of cocaine in a given sale
QUAL = quality of the cocaine expressed as percentage purity
TREND = a time variable with 1984 = 1 up to 1991 = 8
 Consider the regression model

$$PRICE = \beta_1 + \beta_2 QUANT + \beta_3 QUAL + \beta_4 TREND + e$$

- a. What signs would you expect on the coefficients β_2 , β_3 , and β_4 ?

0. β_2 預期負，量大压低購買價
 β_3 預期正，純度為成本荷
 β_4 不確定，無法推測

lm(formula = price ~ quant + qual + trend, data = cocaine)

Residuals:

Min	1Q	Median	3Q	Max
-43.479	-12.014	-3.743	13.969	43.753

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	90.84669	8.58025	10.588	1.39e-14 ***
quant	-0.05997	0.01018	-5.892	2.85e-07 ***
qual	0.11621	0.20326	0.572	0.5700
trend	-2.35458	1.38612	-1.699	0.0954 .

 Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 20.06 on 52 degrees of freedom
 Multiple R-squared: 0.5097, Adjusted R-squared: 0.4814
 F-statistic: 18.02 on 3 and 52 DF, p-value: 3.806e-08

$\beta_2 = -0.05997$ 符合預期
 $\beta_3 = 0.11621$ 符合預期
 $\beta_4 = -2.35458$ 不符合預期
 Multiple $R^2 = 0.5097$
 表約 51% 可被解釋

"t-value for $H_0: \beta_2 = 0$ is -5.892"
 "Critical t-value at 5% significance level: -1.675"

d $H_0: \beta_2 \geq 0$ / $H_1: \beta_2 < 0$

$$t = -5.892 < t_{0.95, 52} = -1.675$$

Reject H_0 .

e. t-value for testing $\beta_3 > 0$: 0.5717
 Critical t-value (5% level): 1.6462

$$H_0: \beta_3 = 0, H_1: \beta_3 > 0$$

$$t = 0.5717 < t_{0.95, 52} = 1.675, \text{ not reject } H_0$$

f. 平均年變化: $\beta_4 = -2.35458$

表示平均價格每年下降約 2.35 美元/每克

可能是因生產技術↑, S 過多 P 下降