10.2 The labor supply of married women has been a subject of a great deal of economic research. Consider the following supply equation specification

$$HOURS = \beta_1 + \beta_2 WAGE + \beta_3 EDUC + \beta_4 AGE + \beta_5 KIDSL6 + \beta_6 NWIFEINC + e$$

where *HOURS* is the supply of labor, *WAGE* is hourly wage, *EDUC* is years of education, *KIDSL6* is the number of children in the household who are less than 6 years old, and *NWIFEINC* is household income from sources other than the wife's employment.

a. Discuss the signs you expect for each of the coefficients.

 $\beta_2 > 0$, 工資高, 工作越多回報增加多→勞動供給增加

 eta_3 不確定,教育程度高越能得到較佳的工作機會與薪資,但要考量具有較好薪資而有休閒與工作之間的 trade off,因此 eta_3 不確定

 eta_4 不確定,年齡增加(年輕時)在一開始可能增加勞動供給,而在老年時可能因為健康因素降低勞動供給

 β_5 < 0,家中兒童多,為照顧兒童可能使工作時間減少→降低勞動供給

 β_6 < 0,其他收入增加,減少工作誘因→降低勞動供給

b. Explain why this supply equation cannot be consistently estimated by OLS regression. 因為工資與 error 可能有內生問題,部分特質如工作能力並未加入,同時影響工時工資,且包含於 error 中

c. Suppose we consider the woman's labor market experience *EXPER* and its square, *EXPER*², to be instruments for *WAGE*. Explain how these variables satisfy the logic of instrumental variables.

EXPER, EXPER²與工資具有相關性 EXPER, EXPER²不直接影響Hours, 只锈過Wage影響

d. Is the supply equation identified? Explain.

Yes,有使用一個內生變數以及至少一個工具變數

- e. Describe the steps [not a computer command] you would take to obtain IV/2SLS estimates.
- 1.先以EXPER, EXPER²及其他外生變數對Wage做迴歸得 Wage
- 2.將原先迴歸中的Wage用Wage替代進行 OLS 迴歸

- 10.3 In the regression model $y = \beta_1 + \beta_2 x + e$, assume x is endogenous and that z is a valid instrument. In Section 10.3.5, we saw that $\beta_2 = \text{cov}(z, y)/\text{cov}(z, x)$.
 - a. Divide the denominator of $\beta_2 = \cos(z, y)/\cos(z, x)$ by $\sin(z)$. Show that $\cos(z, x)/\sin(z)$ is the coefficient of the simple regression with dependent variable x and explanatory variable z, $x = \gamma_1 + \theta_1 z + v$. [Hint: See Section 10.2.1.] Note that this is the first-stage equation in two-stage least squares.

$$\mathcal{A} = Y_1 + Q_1 + V$$

$$\Xi(x) = Y_1 + Q_1 \Xi(x)$$

$$\Rightarrow \Xi(x - \Xi(x))(z - \Xi(z)) = Q_1(z - \Xi(z) + V$$

$$\Rightarrow Z(x - \Xi(x))(z - \Xi(z)) = Q_1 \Xi(z - \Xi(z))$$

$$\Rightarrow Q_1 = \frac{COV(z \cdot x)}{Var(z)}$$

b. Divide the numerator of $\beta_2 = \text{cov}(z, y)/\text{cov}(z, x)$ by var(z). Show that cov(z, y)/var(z) is the coefficient of a simple regression with dependent variable y and explanatory variable z, $y = \pi_0 + \pi_1 z + u$. [Hint: See Section 10.2.1.]

$$y = \pi_0 + \pi_1 z + \pi_2$$

$$\Rightarrow y - E(y) = \pi_1 (z - \pm z) + \pi_2$$

$$\Rightarrow E[(y - E(y))(z - \pm z)] = \pi_1 E[(z - \pm z)]$$

$$\Rightarrow \pi_1 = \frac{Cov(z, y)}{Var(z)}$$

c. In the model $y = \beta_1 + \beta_2 x + e$, substitute for x using $x = \gamma_1 + \theta_1 z + v$ and simplify to obtain $y = \pi_0 + \pi_1 z + u$. What are π_0 , π_1 , and u in terms of the regression model parameters and error and the first-stage parameters and error? The regression you have obtained is a **reduced-form** equation.

$$y = \beta_1 + \beta_2 \cdot (Y_1 + Q_1 + V) + \ell = \beta_1 + \beta_2 Y_1 + \beta_2 Q_1 + \beta_2 Y_1 + \ell + \beta_2 Q_1 + \beta_2 Y_1 + \ell + \beta_2 Q_1 + \beta_2 Y_1 + \ell + \beta_2 Q_1 + \beta_2 Q_1$$

d. Show that $\beta_2 = \pi_1/\theta_1$.

e. If $\hat{\pi}_1$ and $\hat{\theta}_1$ are the OLS estimators of π_1 and θ_1 , show that $\hat{\beta}_2 = \hat{\pi}_1/\hat{\theta}_1$ is a consistent estimator of $\beta_2 = \pi_1/\theta_1$. The estimator $\hat{\beta}_2 = \hat{\pi}_1/\hat{\theta}_1$ is an **indirect least squares** estimator.

$$\widehat{\mathcal{U}}_{1} = \frac{\widehat{\mathcal{L}oV(Z,X)}}{\widehat{\mathcal{V}ar(Z)}} = \frac{\Sigma(\overline{Z_{\tilde{A}}} - \overline{Z})(\overline{A_{\tilde{b}}} - \overline{X})}{\Sigma(\overline{Z_{\tilde{b}}} - \overline{Z})^{*}} \xrightarrow{P} Q$$

$$\widehat{\beta}_{2} = \frac{\widehat{\mathcal{R}}_{1}}{\widehat{\mathcal{Q}}_{1}} = \frac{\Sigma(\overline{\mathcal{B}}_{1} - \overline{\mathcal{B}})(\overline{\mathcal{X}}_{1} - \overline{\mathcal{X}})}{\Sigma(\overline{\mathcal{B}}_{1} - \overline{\mathcal{B}})(\overline{\mathcal{Y}}_{1} - \overline{\mathcal{Y}})} = \frac{\widehat{cov}(\overline{\mathcal{B}}, x)}{\widehat{cov}(\overline{\mathcal{B}}, y)} \xrightarrow{P} \widehat{\mathcal{B}}_{2}$$