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- 4.4** The general manager of a large engineering firm wants to know whether the experience of technical artists influences their work quality. A random sample of 50 artists is selected. Using years of work experience (*EXPER*) and a performance rating (*RATING*, on a 100-point scale), two models are estimated by least squares. The estimates and standard errors are as follows:

Model 1:

$$\widehat{RATING} = 64.289 + 0.990EXPER \quad N = 50 \quad R^2 = 0.3793$$

(se) (2.422) (0.183)

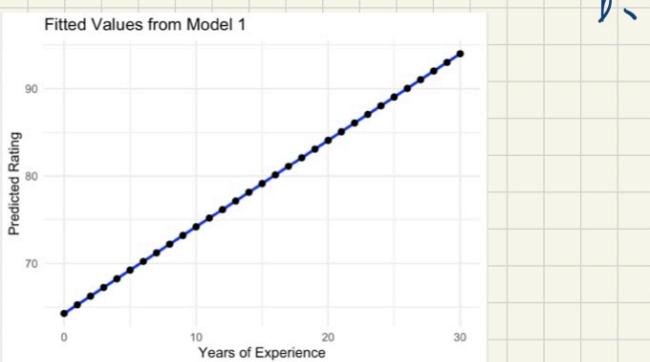
Model 2:

$$\widehat{RATING} = 39.464 + 15.312 \ln(EXPER) \quad N = 46 \quad R^2 = 0.6414$$

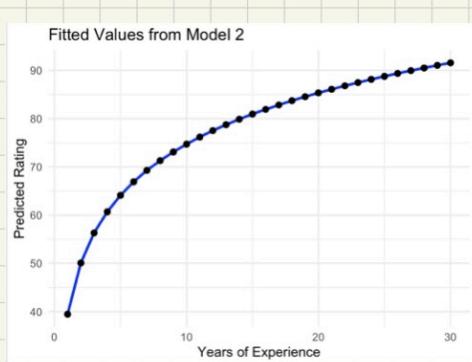
(se) (4.198) (1.727)

- a. Sketch the fitted values from Model 1 for *EXPER* = 0 to 30 years.
- b. Sketch the fitted values from Model 2 against *EXPER* = 1 to 30 years. Explain why the four artists with no experience are not used in the estimation of Model 2.
- c. Using Model 1, compute the marginal effect on *RATING* of another year of experience for (i) an artist with 10 years of experience and (ii) an artist with 20 years of experience.
- d. Using Model 2, compute the marginal effect on *RATING* of another year of experience for (i) an artist with 10 years of experience and (ii) an artist with 20 years of experience.
- e. Which of the two models fits the data better? Estimation of Model 1 using just the technical artists with some experience yields $R^2 = 0.4858$.
- f. Do you find Model 1 or Model 2 more reasonable, or plausible, based on economic reasoning? Explain.

a.



b.



c.

Model 1 : Margin effect: $\frac{d \widehat{RATING}}{d EXPER} = 0.99$

(i) 10 年經驗，每增加 1 年，邊際效應增加 0.99

(ii) 20 年經驗，每增加 1 年，邊際效應增加 0.99

代表不管目前的經驗多少。

邊際效應都是一樣的。

d. Model 2. Margin effect: $\frac{d \hat{RATING}}{d EXPER} = 15.312 \cdot \frac{1}{EXPER}$

(i) 10 年經驗
邊際效應: $15.312 \times \frac{1}{10} = 1.5312 \rightarrow$ 代表經驗帶來的效益是遞減的.

(ii) 20 年經驗
邊際效應: $15.312 \times \frac{1}{20} = 0.7656$

e. Model 1: $R^2 = 0.3793$

R^2 越高代表模型擬合度好.

Model 1 (just technical artists): $R^2 = 0.4858$

\rightarrow Model 2 的擬合度高.

Model 2: $R^2 = 0.6414$

f. Model 1 (線性模型): 隨著經驗上升，但邊際效應會不變，在現實中不合理.

Model 2 (對數模型): 隨著經驗上升，邊際效應遞減，符合現實，因為會在後期成長放緩.

4.28

A.

回歸

```
> summary(mod1)
Call:
lm(formula = northampton ~ time, data = wa_wheat)

Residuals:
    Min      1Q  Median      3Q     Max 
-0.62394 -0.17302  0.03342  0.12996  0.72050 

Coefficients:
            Estimate Std. Error t value Pr(>|t|)    
(Intercept) 0.603245  0.081858  7.369 2.55e-09 ***  
time        0.023078  0.002908  7.935 3.69e-10 ***  
---
Signif. codes:  0 ‘***’ 0.001 ‘**’ 0.01 ‘*’ 0.05 ‘.’ 0.1 ‘ ’ 1 

Residual standard error: 0.2791 on 46 degrees of freedom
Multiple R-squared:  0.5778,   Adjusted R-squared:  0.5687 
F-statistic: 62.96 on 1 and 46 DF,  p-value: 3.689e-10

> summary(mod2)
Call:
lm(formula = northampton ~ log(time), data = wa_wheat)

Residuals:
    Min      1Q  Median      3Q     Max 
-0.78488 -0.20711 -0.06382  0.15447  0.91573 

Coefficients:
            Estimate Std. Error t value Pr(>|t|)    
(Intercept) 0.3510    0.1759   1.995   0.052 .  
log(time)   0.2790    0.0575   4.852 1.44e-05 ***  
---
Signif. codes:  0 ‘***’ 0.001 ‘**’ 0.01 ‘*’ 0.05 ‘.’ 0.1 ‘ ’ 1 

Residual standard error: 0.3494 on 46 degrees of freedom
Multiple R-squared:  0.3386,   Adjusted R-squared:  0.3242 
F-statistic: 23.55 on 1 and 46 DF,  p-value: 1.44e-05

> summary(mod3)
Call:
lm(formula = northampton ~ I(time^2), data = wa_wheat)

Residuals:
    Min      1Q  Median      3Q     Max 
-0.56899 -0.14970  0.03119  0.12176  0.62049 

Coefficients:
            Estimate Std. Error t value Pr(>|t|)    
(Intercept) 7.737e-01  5.222e-02 14.82 < 2e-16 ***  
I(time^2)  4.986e-04  4.939e-05 10.10 3.01e-13 ***  
---
Signif. codes:  0 ‘***’ 0.001 ‘**’ 0.01 ‘*’ 0.05 ‘.’ 0.1 ‘ ’ 1 

Residual standard error: 0.2396 on 46 degrees of freedom
Multiple R-squared:  0.689,   Adjusted R-squared:  0.6822 
F-statistic: 101.9 on 1 and 46 DF,  p-value: 3.008e-13

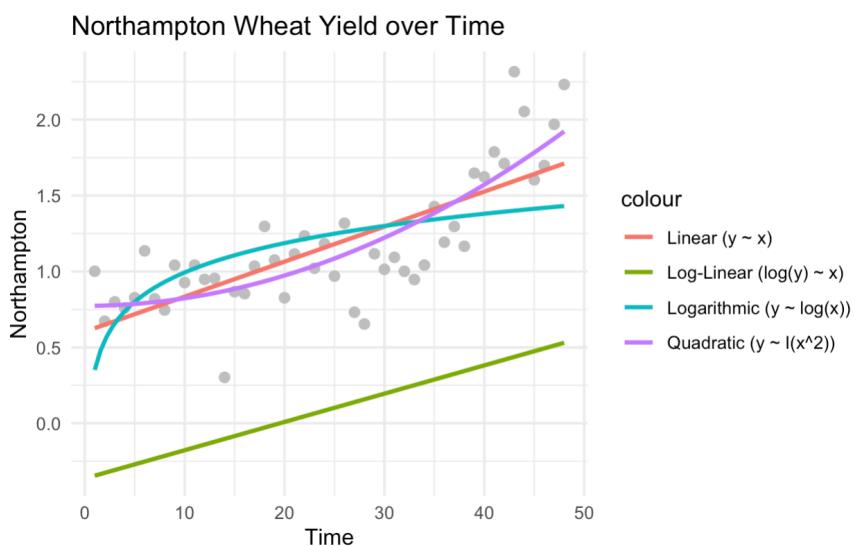
> summary(mod4)
Call:
lm(formula = log(northampton) ~ time, data = wa_wheat)

Residuals:
    Min      1Q  Median      3Q     Max 
-1.09292 -0.10049  0.07125  0.14140  0.40263 

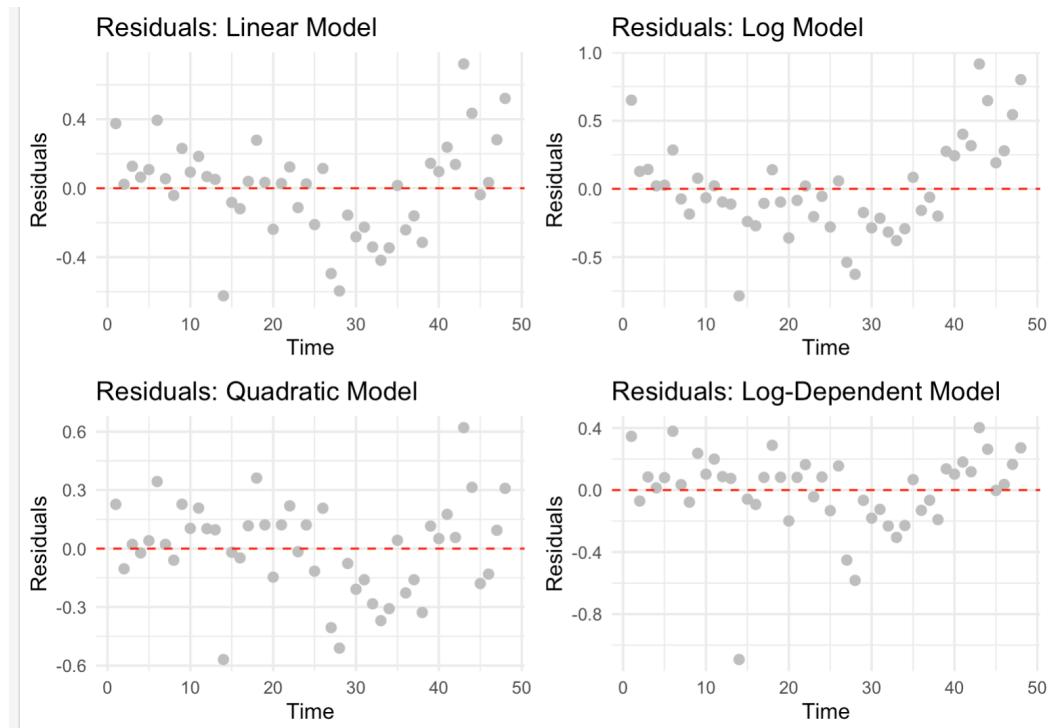
Coefficients:
            Estimate Std. Error t value Pr(>|t|)    
(Intercept) -0.363938  0.076192 -4.777 1.85e-05 ***  
time        0.018632  0.002707  6.883 1.37e-08 ***  
---
Signif. codes:  0 ‘***’ 0.001 ‘**’ 0.01 ‘*’ 0.05 ‘.’ 0.1 ‘ ’ 1 

Residual standard error: 0.2598 on 46 degrees of freedom
Multiple R-squared:  0.5074,   Adjusted R-squared:  0.4966 
F-statistic: 47.37 on 1 and 46 DF,  p-value: 1.366e-08
```

回歸圖



殘差圖



殘差常態分佈檢定

Shapiro-Wilk normality test

```
data: residuals(mod1)
W = 0.98236, p-value = 0.6792
```

Shapiro-Wilk normality test

```
data: residuals(mod2)
W = 0.96657, p-value = 0.1856
```

Shapiro-Wilk normality test

```
data: residuals(mod3)
W = 0.98589, p-value = 0.8266
```

Shapiro-Wilk normality test

```
data: residuals(mod4)
W = 0.86894, p-value = 7.205e-05
```

p-value:

Mod1 > 0.05 , 殘差呈現常態分佈

Mod2 > 0.05 , 殘差呈現常態分佈

Mod3 > 0.05 , 殘差呈現常態分佈

Mod4 < 0.05 , 殘差非常態分佈

R^2 :

Mod1 R^2 : 0.5778

Mod2 R^2 : 0.3494

Mod3 R^2 : 0.689

Mod4 R^2 : 0.2598

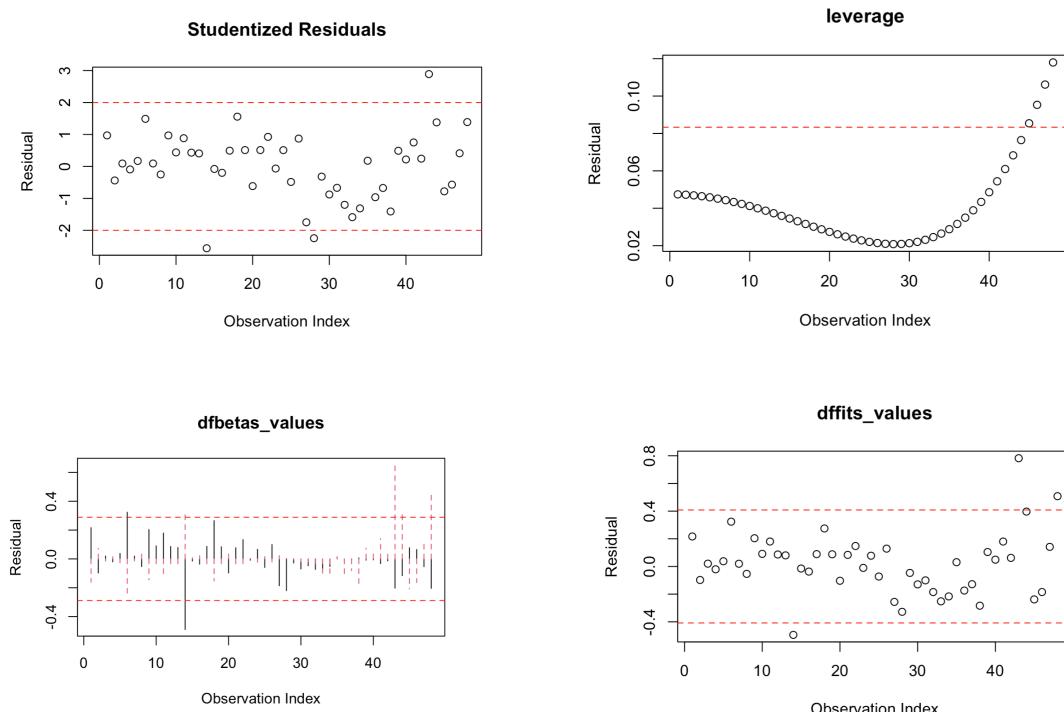
因此我覺得模型 3 的表現較好，因為他的 R^2 較接近 1

B. 係數解釋

當 $\text{time} = 0$ 時，值為 0.7737

The coefficient of time^2 is 0.0004986，代表時間越大，產出會隨之增加，也就是每增加 1 單位的 time^2 ，yield 會增加約 0.0004986 單位

C.



Studentized residuals :

大部分觀測值的殘差皆落在 ± 2 範圍內，顯示模型預測結果整體穩定。然而，仍有少數觀測值超出此範圍，可能為潛在的離群點或異常觀測值

Leverage:

大部分觀測值的槓桿值均低於參考線，顯示其對模型的影響有限。然而，在資料序列後段，有數個觀測值的槓桿值明顯偏高，可能為潛在的高影響力資料點

dfbetas_values:

這張 DFBETAs 圖顯示大多數觀測值的 DFBETA 值均落在 ± 0.3 的門檻範圍內，顯示這些資料對模型係數的影響有限。然而，在觀測值編號接近 43 至 47 的區間中，有數筆資料的 DFBETA 超出紅線，表示這些觀測值可能對某一或多個迴歸係數產生顯著影響，可能為潛在的高影響觀測點

dffits_values:

這張 DFFITS 值圖用於檢查每筆觀測值對模型預測的影響程度。結果顯示，大多數觀測值的 DFFITS 值均落在 ± 0.4 的警戒範圍內，顯示其對模型預測的影響有限。然而，在觀測值編號約 45 至 47 處，DFFITS 值超出警戒門檻，顯示這些資料對模型的預測結果影響較大，，可能為潛在的高影響觀測點

D.

```
> pred
    fit      lwr      upr
1 1.881111 1.372403 2.389819
> origin
[1] 2.2318
```

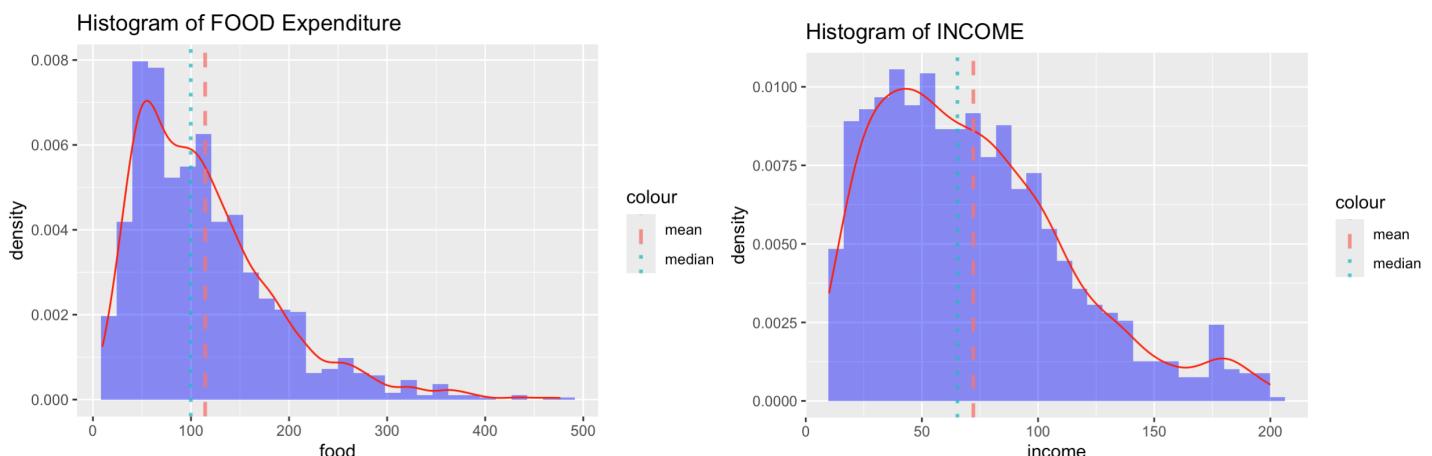
2.2318有落在區間之中

4.29

A.

```
> describe(cx5_small$food)
   vars   n   mean     sd median trimmed   mad min   max range skew kurtosis se
X1     1 1200 114.44 72.66    99.8 105.03 66.18 9.63 476.67 467.04 1.35     2.36 2.1
> describe(cx5_small$income)
   vars   n   mean     sd median trimmed   mad min   max range skew kurtosis se
X1     1 1200 72.14 41.65    65.29  67.94 41.65 10  200  190 0.84     0.32 1.2
```

	mean	median	Minimum	maximum	Standard Deviation
food	114.44	99.8	9.63	476.67	72.66
income	72.14	65.29	10	200	41.65



The histogram of food expenditure are not symmetrical and bell-shaped curve.
food_mean > food_median : right-skewed.

The histogram of income are not symmetrical and bell-shaped curve.
income_mean > income_median : right-skewed

```
> print(jb_food)
```

```
Jarque-Bera Normality Test  
  
data: cex5_small$food  
JB = 648.65, p-value < 2.2e-16  
alternative hypothesis: greater
```

```
> print(jb_income)
```

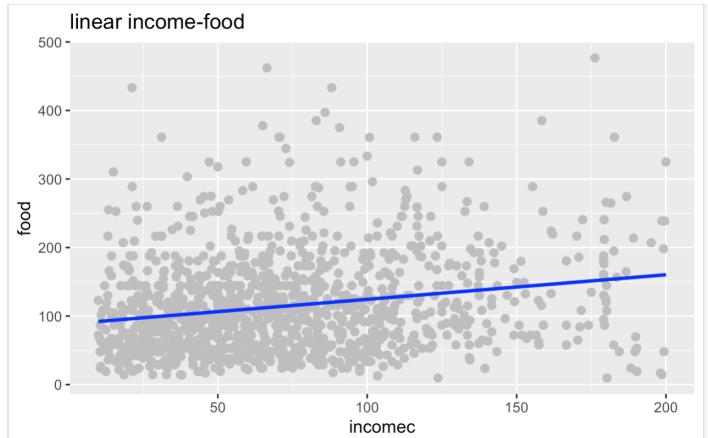
```
Jarque-Bera Normality Test  
  
data: cex5_small$income  
JB = 148.21, p-value < 2.2e-16  
alternative hypothesis: greater
```

Food of JB = 648.65
P-value < 0.05, 所以拒絕虛無假設
也就是 food 並非常態分佈

Income of JB = 648.65
P-value < 0.05, 所以拒絕虛無假設
也就是 income 並非常態分佈

B.

```
Call:  
lm(formula = food ~ income, data = cex5_small)  
  
Residuals:  
    Min      1Q  Median      3Q     Max  
-145.37 -51.48 -13.52  35.50 349.81  
  
Coefficients:  
            Estimate Std. Error t value Pr(>|t|)  
(Intercept) 88.56650   4.10819 21.559 < 2e-16 ***  
income       0.35869   0.04932  7.272 6.36e-13 ***  
---  
Signif. codes:  0 ‘***’ 0.001 ‘**’ 0.01 ‘*’ 0.05 ‘.’ 0.1 ‘ ’ 1  
  
Residual standard error: 71.13 on 1198 degrees of freedom  
Multiple R-squared:  0.04228, Adjusted R-squared:  0.04148  
F-statistic: 52.89 on 1 and 1198 DF,  p-value: 6.357e-13
```



```
> confint(mod1, level = 0.95)  
    2.5 %    97.5 %  
(Intercept) 80.5064570 96.626543  
income       0.2619215  0.455452  
>
```

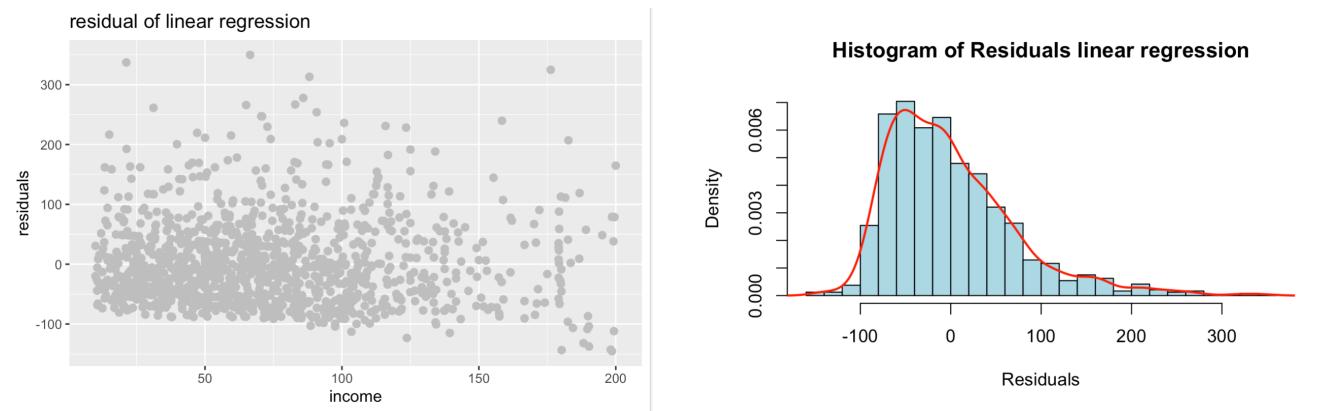
Estimate linear regression:
Food = 88.5665 + 0.3587*Income

95% interval estimate for β_2 is
[0.2619, 0.4555]

從係數的角度來看，t值為7.272 是具有顯著性的，且p值極小且顯著，再加上95%信賴區間為 [0.2619, 0.4555]，相對窄，因此可以說收入對於食物支出的影響是統計上精確的。

但是從 $R^2 = 0.04228$ 的角度來看，此模型的解釋力很低，因此可能還有其他因素會影像食物支出。

C.



從散佈圖中來看並沒有明顯的線性趨勢，但是可以發現隨著收入增加殘差的變異增大
從直方圖來看可以發現圖形有右偏，且非完全對稱

```
> jarque.bera.test(residuals(mod1)) #測試殘差的常態性
```

Jarque Bera Test

```
data: residuals(mod1)
X-squared = 624.19, df = 2, p-value < 2.2e-16
```

從 Jarque Bera Test 來看

虛無假設 H_0 ：殘差為常態分配

替代假設 H_1 ：殘差非常態分配

p-value < 0.05, 所以拒絕虛無假設，接受替代假設，也就是殘差非常態分佈

殘差為常態分佈比較重要，因為殘差為常態分佈，才可以確保後續的統計推論是正確的

D.

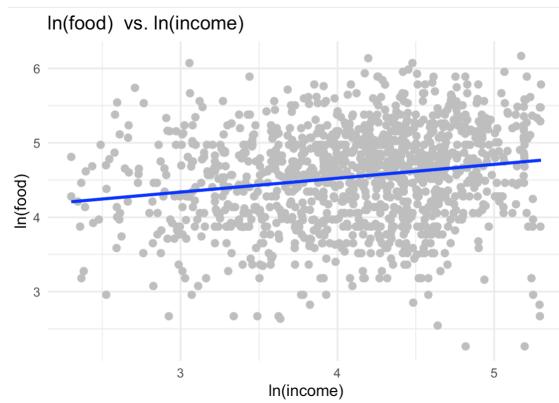
```
income = 19 expected food expenditure 95.38155 , elasticity = 0.07145038 ,se =  0.00982475 ,elasticity_interval = 0.05217475 0.09072601
income = 65 expected food expenditure 111.8811 , elasticity = 0.2083876 ,se =  0.02865423 ,elasticity_interval = 0.1521695 0.2646056
income = 160 expected food expenditure 145.9564 , elasticity = 0.3931988 ,se =  0.05406661 ,elasticity_interval = 0.287123 0.4992746
> |
```

彈性隨著收入增加而增加，代表收入增加時會增加食物支出

信賴區間並沒有重疊，代表在統計上有顯著差異

根據經濟學原則，隨著收入上升，食物支出的彈性會降低，但在此模型中看到食物彈性隨著收入增加而增加，代表在模型設定的收入範圍內，食物仍是具有收入敏感性的商品

E.



```
Call:
lm(formula = ln_food ~ ln_income, data = df_log)

Residuals:
    Min      1Q  Median      3Q     Max 
-2.48175 -0.45497  0.06151  0.46063  1.72315 

Coefficients:
            Estimate Std. Error t value Pr(>|t|)    
(Intercept) 3.77893   0.12035 31.400 <2e-16 ***
ln_income   0.18631   0.02903  6.417  2e-10 ***
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.6418 on 1198 degrees of freedom
Multiple R-squared:  0.03323, Adjusted R-squared:  0.03242 
F-statistic: 41.18 on 1 and 1198 DF,  p-value: 1.999e-10
```

log-log model 的 R^2 為 0.0332， linear model 的 R^2 為 0.0423

因此 log-log model 的模型適配度較差，也就是 linear model seems fit data better

F.

```
> cat("the point of the elasticity", b2_log)
the point of the elasticity 0.1863054
> cat ("95% interval estimate of the elasticity",interval_log)
95% interval estimate of the elasticity 0.1293432 0.2432675
```

```
income = 19 ,elasticity_interval = 0.05217475 0.09072601 elasticity = 0.07145038
income = 65 ,elasticity_interval = 0.1521695 0.2646056 elasticity = 0.2083876
income = 160 ,elasticity_interval = 0.287123 0.4992746 elasticity = 0.3931988
>
```

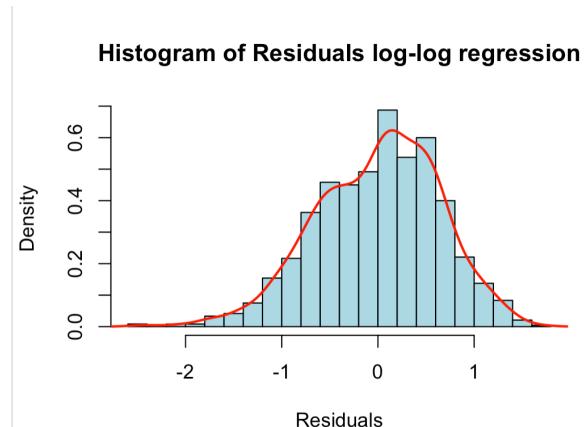
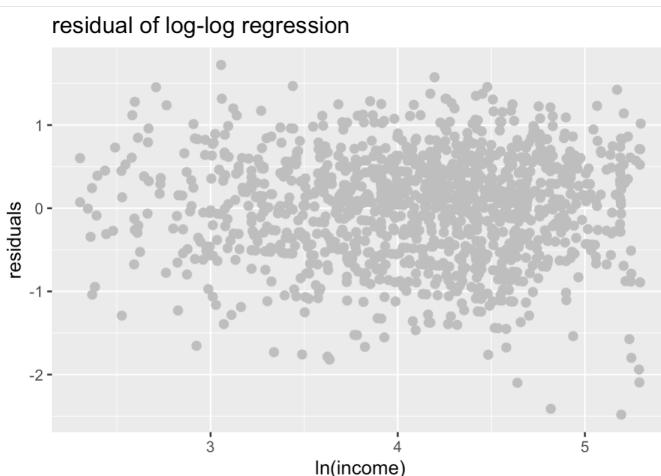
Linear–linear 資料

log–log model 的彈性區間 和 linear–linear model 在 Income = 19 與 160 時的彈性區間並沒有重疊，因此在統計上有顯著差異

log–log model 的彈性區間和 linear–linear model 在 Income = 65 時的彈性區間有重疊，因此在統計上沒有顯著差異

因此可以說 log–log model 所估計的彈性與 linear–linear model 在 income = 65 時的彈性相近，但與低收入與高收入時的彈性明顯不同。

G.

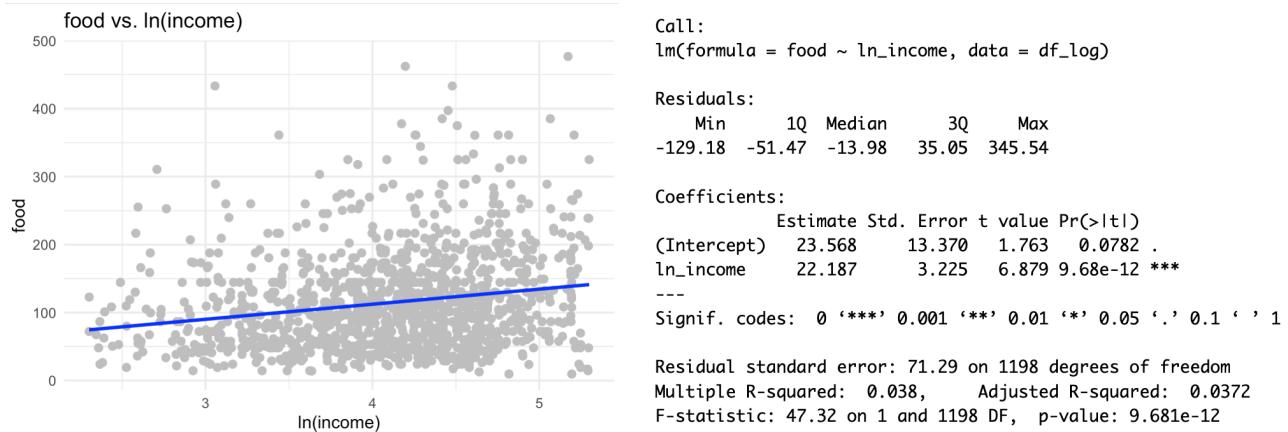


Jarque Bera Test

```
data: residuals(mod2)
X-squared = 25.85, df = 2, p-value = 2.436e-06
```

從殘差圖來看，並沒有明顯的趨勢，從直方圖來看，大致是鐘狀，但有偏態與峰態異常。
從JB檢定來看，p–value小於 0.05拒絕虛無假設，代表殘差非常態分佈

H.



Linear-log model 的 R^2 為 0.038

log-log model 的 R^2 為 0.0332 , linear-linear model 的 R^2 為 0.0423

可以得出 linear-linear model R^2 最大，模型配適度最高

I.

```
income = 19 expected food expenditure 88.89788 , elasticity = 0.2495828 ,se = 0.03628131 ,elasticity_interval = 0.1784009 0.3207648
income = 65 expected food expenditure 116.1872 , elasticity = 0.1909624 ,se = 0.02775978 ,elasticity_interval = 0.1364992 0.2454256
income = 160 expected food expenditure 136.1733 , elasticity = 0.1629349 ,se = 0.02368549 ,elasticity_interval = 0.1164652 0.2094046
```

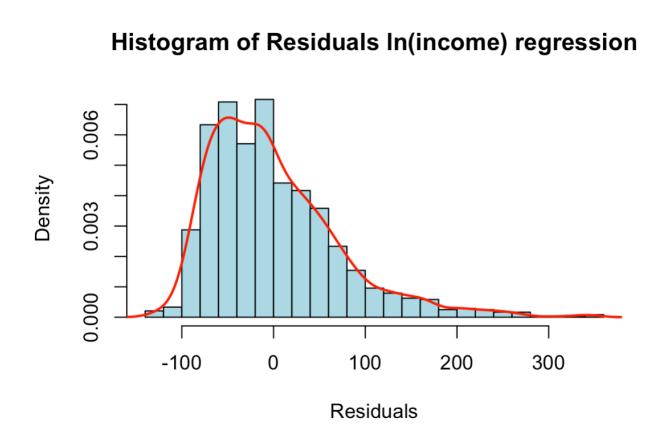
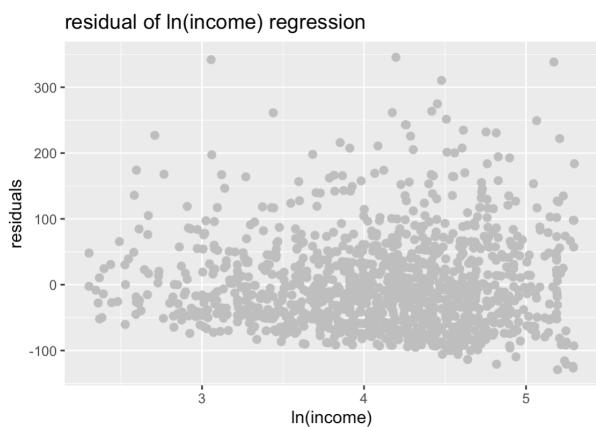
Linear-log model vs. log-log model : 在不同的收入水準之下，彈性區間幾乎重疊，代表沒有顯著差異

Linear-log model vs. log-log model :

linear-log model 的彈性區間 和 linear-linear model 在 Income = 19 與 160 時的彈性區間並沒有重疊，因此在統計上有顯著差異

linear-log model 的彈性區間和 linear-linear model 在 Income = 65 時的彈性區間有重疊，因此在統計上沒有顯著差異

J.



Jarque Bera Test

```
data: residuals(mod3)
X-squared = 628.07, df = 2, p-value < 2.2e-16
```

從散佈圖來看是呈現隨機分佈，但是隨著收入越大有殘差有變大的趨勢，從直方圖可以看出有左偏的趨勢，從JB檢定來看，p-value 小於 0.05 ，殘差非常態分佈

K. 我會選擇log-log模型，因為他在三個模型之中殘差是最接近常態分佈的