

Q6

TABLE 15.10 Estimation Results for Exercise 15.6

	(1) OLS 1987	(2) OLS 1988	(3) FE	(4) FE Robust	(5) RE
<i>C</i>	0.9348 (0.2010)	0.8993 (0.2407)	1.5468 (0.2522)	1.5468 (0.2688)	1.1497 (0.1597)
<i>EXPER</i>	0.1270 (0.0295)	0.1265 (0.0323)	0.0575 (0.0330)	0.0575 (0.0328)	0.0986 (0.0220)
<i>EXPER</i> ²	-0.0033 (0.0011)	-0.0031 (0.0011)	-0.0012 (0.0011)	-0.0012 (0.0011)	-0.0023 (0.0007)
<i>SOUTH</i>	-0.2128 (0.0338)	-0.2384 (0.0344)	-0.3261 (0.1258)	-0.3261 (0.2495)	-0.2326 (0.0317)
<i>UNION</i>	0.1445 (0.0382)	0.1102 (0.0387)	0.0822 (0.0312)	0.0822 (0.0367)	0.1027 (0.0245)
<i>N</i>	716	716	1432	1432	1432

f. The most of difference \Rightarrow South

$$\text{EXPER}^2 \quad \frac{0.023}{0.012} \approx 1.92. \quad \text{Hausman test: } t = \frac{\hat{\beta}_{FE} - \hat{\beta}_{RE}}{\sqrt{\text{Var}(\hat{\beta}_{FE}) - \text{Var}(\hat{\beta}_{RE})}}$$

$$t_{\text{EXPER}} = \frac{0.0575 - 0.0986}{\sqrt{0.033^2 - 0.022^2}} \approx -1.6 \quad , \quad t_{\text{EXPER}^2} = \frac{-0.0012 - (-0.0021)}{\sqrt{0.0011^2 - 0.0007^2}} \approx 1.29$$

$$t_{\text{SOUTH}} = \frac{-0.3261 - (-0.2326)}{\sqrt{0.1258^2 - 0.0317^2}} \approx -0.77 \quad t_{\text{UNION}} = \frac{0.0822 - 0.1027}{\sqrt{0.0312^2 - 0.0245^2}} \approx -1.06$$

Random effect is appropriate *

Q17

15.17 The data file *liquor* contains observations on annual expenditure on liquor (*LIQUOR*) and annual income (*INCOME*) (both in thousands of dollars) for 40 randomly selected households for three consecutive years.

- b. Estimate the model $LIQUOR_{it} = \beta_1 + \beta_2 INCOME_{it} + u_i + e_{it}$ using random effects. Construct a 95% interval estimate of the coefficient on *INCOME*. How does it compare to the interval in part (a)?

```
2.5 %.income 97.5 %.income
0.01283111 0.04031983
```

與(a)相比，不包含 0，表示 β_2 顯著不為 0，支持收入變化會影響酒類支出

- c. Test for the presence of random effects using the LM statistic in equation (15.35). Use the 5% level of significance.

Lagrange Multiplier Test - (Breusch-Pagan)

```
data: liquor ~ income
chisq = 20.68, df = 1, p-value = 5.429e-06
alternative hypothesis: significant effects
```

```
> qchisq(p=0.95,df=1)
[1] 3.841459
```

Reject H_0 (H_0 : No random effect)

- d. For each individual, compute the time averages for the variable *INCOME*. Call this variable *INCOMEM*. Estimate the model $LIQUOR_{it} = \beta_1 + \beta_2 INCOME_{it} + \gamma INCOMEM_i + c_i + e_{it}$ using the random effects estimator. Test the significance of the coefficient γ at the 5% level. Based on this test, what can we conclude about the correlation between the random effect u_i and *INCOME*? Is it OK to use the random effects estimator for the model in (b)?

```
Call:
plm(formula = liquor ~ income + INCOMEM, data = pdat2, model = "random")
```

Balanced Panel: n = 40, T = 3, N = 120

Effects:

	var	std.dev	share
idiosyncratic	0.9640	0.9819	0.571
individual	0.7251	0.8515	0.429
theta:	0.4459		

Residuals:

Min.	1st Qu.	Median	3rd Qu.	Max.
-2.300955	-0.703840	0.054992	0.560255	2.257325

Coefficients:

	Estimate	Std. Error	z-value	Pr(> z)
(Intercept)	0.9163337	0.5524439	1.6587	0.09718 .
income	0.0207421	0.0209083	0.9921	0.32117
INCOMEM	0.0065792	0.0222048	0.2963	0.76700

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Total Sum of Squares: 126.61

Residual Sum of Squares: 112.79

R-Squared: 0.10917

Adj. R-Squared: 0.093945

Chisq: 14.3386 on 2 DF, p-value: 0.00076987

$H_0: r = 0$, $H_1: r \neq 0$

p-value = 0.767 > 0.05

無法拒絕 H_0

INCOMEM 和個體隨機效果 u_i 沒有相關性，所以可以使用 random effect model

Q20

15.20 This exercise uses data from the STAR experiment introduced to illustrate fixed and random effects for grouped data. In the STAR experiment, children were randomly assigned within schools into three types of classes: small classes with 13–17 students, regular-sized classes with 22–25 students, and regular-sized classes with a full-time teacher aide to assist the teacher. Student scores on achievement tests were recorded as well as some information about the students, teachers, and schools. Data for the kindergarten classes are contained in the data file *star*.

- d. Reestimate the model in part (a) with school random effects. Compare the results with those from parts (a) and (b). Are there any variables in the equation that might be correlated with the school effects? Use the LM test for the presence of random effects.

Oneway (individual) effect Random Effect Model
(Swamy-Arora's transformation)

```
call:
plm(formula = readscore ~ small + aide + tchexper + boy + white_asian +
      freelunch, data = pdata, model = "random")
```

Unbalanced Panel: n = 79, T = 34-137, N = 5766

Effects:

	var	std.dev	share
idiosyncratic	751.43	27.41	0.829
individual	155.31	12.46	0.171

theta:

	Min.	1st Qu.	Median	Mean	3rd Qu.	Max.
	0.6470	0.7225	0.7523	0.7541	0.7831	0.8153

Residuals:

	Min.	1st Qu.	Median	Mean	3rd Qu.	Max.
	-97.483	-17.236	-3.282	0.037	12.803	192.346

Coefficients:

	Estimate	Std. Error	z-value	Pr(> z)
(Intercept)	436.126774	2.064782	211.2217	< 2.2e-16 ***
small	6.458722	0.912548	7.0777	1.466e-12 ***
aide	0.992146	0.881159	1.1260	0.2602
tchexper	0.302679	0.070292	4.3060	1.662e-05 ***
boy	-5.512081	0.727639	-7.5753	3.583e-14 ***
white_asian	7.350477	1.431376	5.1353	2.818e-07 ***
freelunch	-14.584332	0.874676	-16.6740	< 2.2e-16 ***

估計相近

Lagrange Multiplier Test - (Breusch-Pagan)

```
data: readscore ~ small + aide + tchexper + boy + white_asian + freelunch
chisq = 6677.4, df = 1, p-value < 2.2e-16
alternative hypothesis: significant effects
```

p-value < 0.05, reject H0, 存在 RE, school level unobserved heterogeneity

- e. Using the *t*-test statistic in equation (15.36) and a 5% significance level, test whether there are any significant differences between the fixed effects and random effects estimates of the coefficients on *SMALL*, *AIDE*, *TCHEXPER*, *WHITE_ASIAN*, and *FREELUNCH*. What are the implications of the test outcomes? What happens if we apply the test to the fixed and random effects estimates of the coefficient on *BOY*?

Hausman Test

```
data: readscore ~ small + aide + tchexper + boy + white_asian + freelunch
chisq = 13.809, df = 6, p-value = 0.03184
alternative hypothesis: one model is inconsistent
```

H0: $\beta_{FE} = \beta_{RE}$, H1: $\beta_{FE} \neq \beta_{RE}$

```
small      : t = 1.15, p = 0.252
aide       : t = 0.13, p = 0.898
tchexper   : t = -1.94, p = 0.053
white_asian : t = 1.22, p = 0.223
freelunch  : t = -0.10, p = 0.924
```

- f. Create school-averages of the variables and carry out the Mundlak test for correlation between them and the unobserved heterogeneity.

Unbalanced Panel: n = 78, T = 34-136, N = 5681

Effects:

	var	std.dev	share
idiosyncratic	756.11	27.50	0.817
individual	169.40	13.02	0.183

theta:

Min.	1st Qu.	Median	Mean	3rd Qu.	Max.
0.6593	0.7327	0.7615	0.7630	0.7892	0.8217

Residuals:

Min.	1st Qu.	Median	Mean	3rd Qu.	Max.
-98.886	-17.051	-3.166	0.039	12.846	193.321

Coefficients:

	Estimate	Std. Error	z-value	Pr(> z)
(Intercept)	459.462989	20.529888	22.3802	< 2.2e-16 ***
small	6.637460	0.922068	7.1985	6.090e-13 ***
aide	1.157620	0.889542	1.3014	0.1931
tchexper	0.289286	0.071754	4.0316	5.539e-05 ***
boy	-5.386109	0.735063	-7.3274	2.346e-13 ***
white_asian	8.081423	1.550155	5.2133	1.855e-07 ***
freelunch	-14.699025	0.892109	-16.4767	< 2.2e-16 ***
small_m	-18.410060	22.273923	-0.8265	0.4085
aide_m	16.811358	20.793685	0.8085	0.4188
tchexper_m	1.006007	0.625690	1.6078	0.1079
boy_m	-53.353521	25.221654	-2.1154	0.0344 *
white_asian_m	-6.648191	6.320012	-1.0519	0.2928
freelunch_m	-3.318853	8.779553	-0.3780	0.7054

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Boy_m's p-value < 0.05, 應採用 FE

其他使用 RE