

10.

The labor supply of married women has been a subject of a great deal of economic research. Consider the following supply equation specification

$$HOURS = \beta_1 + \beta_2 WAGE + \beta_3 EDUC + \beta_4 AGE + \beta_5 KIDS6 + \beta_6 NWIFEINC + e$$

where $HOURS$ is the supply of labor, $WAGE$ is hourly wage, $EDUC$ is years of education, $KIDS6$ is the number of children in the household who are less than 6 years old, and $NWIFEINC$ is household income from sources other than the wife's employment.

- Discuss the signs you expect for each of the coefficients.
- Explain why this supply equation cannot be consistently estimated by OLS regression.
- Suppose we consider the woman's labor market experience $EXPER$ and its square, $EXPER^2$, to be instruments for $WAGE$. Explain how these variables satisfy the logic of instrumental variables.
- Is the supply equation identified? Explain.
- Describe the steps [not a computer command] you would take to obtain IV/2SLS estimates.

X: Endogeneity :

- 遗漏变数
- 互为因果
- 测量误差

a. $\beta_2 > 0$: 資薪越高, 工作意願上升

$\beta_3 > 0$: 教育程度高, 工作机会↑

$\beta_4 \text{ not sure}$: 年齡↑→累積經驗, 但也可能 : 高齡退休而 hours↓

$\beta_5 < 0$: 小孩多 → 謹慎 time↑ → 工作時間↓

$\beta_6 < 0$: 有其它額外收入 → 降低工作意願

b. 遺漏变数: wage 可能和誤差項有關, 有未被觀察到的變數
ex: 工作能力、健康狀況 → 產生內生性問題

c. 工具變數必須和 e_i 不相關, 但是不影響 y_i , 即 $E(e_i)$,
而 Experience 滿足和 Wage 高度正相關, 且為非線性關係, 所以用二次方捕捉關係
理論上 Experience 不會影響工作時數, 故 Exper 與 e_i 無關

d. identified 定義: 若工具變數 (IV) 相當於內生性變數, 則滿足 identified 條件
(在 IV or 2SLS 漢中, 必須滿足 identified 的條件!)

IV ($exper, exper^2$) 2個 > 內生變數 (wage), 故符合 identified 條件

e. $Wage = \gamma_0 + \gamma_1 EXPER + \gamma_2 EXPER^2 + \gamma_3 EDUC + \gamma_4 AGE + \gamma_5 KIDS6 + \gamma_6 NWIFEINC$
可以得到預測值 $\widehat{WAGE} \rightarrow$ 隔離掉 Wage 對誤差項的影響
再將預測值 \widehat{WAGE} 代入原迴歸式 $\rightarrow HOURS = \beta_1 + \beta_2 \widehat{WAGE} + \beta_3 EDUC + \dots$
得到的 β_i 估計量, 即 2SLS 的一致估計量。

- (10) In the regression model $y = \beta_1 + \beta_2 x + e$, assume x is endogenous and that z is a valid instrument. In Section 10.3.5, we saw that $\beta_2 = \text{cov}(z, y)/\text{cov}(z, x)$.
- Divide the denominator of $\beta_2 = \text{cov}(z, y)/\text{cov}(z, x)$ by $\text{var}(z)$. Show that $\text{cov}(z, x)/\text{var}(z)$ is the coefficient of the simple regression with dependent variable x and explanatory variable z .
 $\checkmark x = \gamma_1 + \theta_1 z + v$. [Hint: See Section 10.2.1.] Note that this is the first-stage equation in two-stage least squares.
 - Divide the numerator of $\beta_2 = \text{cov}(z, y)/\text{cov}(z, x)$ by $\text{var}(z)$. Show that $\text{cov}(z, y)/\text{var}(z)$ is the coefficient of a simple regression with dependent variable y and explanatory variable z .
 $\checkmark y = \pi_0 + \pi_1 z + u$. [Hint: See Section 10.2.1.]
 - In the model $y = \beta_1 + \beta_2 x + e$, substitute for x using $x = \gamma_1 + \theta_1 z + v$ and simplify to obtain $y = \pi_0 + \pi_1 z + u$. What are π_0 , π_1 , and u in terms of the regression model parameters and error and the first-stage parameters and error? The regression you have obtained is a **reduced-form** equation.
 - Show that $\beta_2 = \pi_1/\theta_1$.
 - If $\hat{\pi}_1$ and $\hat{\theta}_1$ are the OLS estimators of π_1 and θ_1 , show that $\hat{\beta}_2 = \hat{\pi}_1/\hat{\theta}_1$ is a consistent estimator of $\beta_2 = \pi_1/\theta_1$. The estimator $\hat{\beta}_2 = \hat{\pi}_1/\hat{\theta}_1$ is an **indirect least squares** estimator.

a. 一階殘差模型: $X = \gamma_1 + \theta_1 Z + V \quad \text{--- ①}$

期望值: $E(X) = \gamma_1 + \theta_1 E(Z) \quad \text{--- ②}$

① - ② $\Rightarrow X - E(X) = \theta_1 (Z - E(Z)) + V$

$$[X - E(X)][Z - E(Z)] = \theta_1 (Z - E(Z))^2 + (Z - E(Z))V$$

$$E\{[X - E(X)][Z - E(Z)]\} = \theta_1 E[(Z - E(Z))^2] + E[(Z - E(Z))V]$$

$\because Z - E(Z)$ 和 V 不相關 $\therefore E[(Z - E(Z))V] = 0$

$$\rightarrow E[(X - E(X))(Z - E(Z))] = \theta_1 E[(Z - E(Z))^2]$$

$$\Rightarrow \text{cov}(Z, X) = \theta_1 \text{var}(Z)$$

$$\Rightarrow \theta_1 = \frac{\text{cov}(Z, X)}{\text{var}(Z)}$$

b.

$$y = \pi_0 + \pi_1 Z + u \quad \text{--- ③}$$

$$E(y) = \pi_0 + \pi_1 E(Z) \quad \text{--- ④}$$

$$\text{③} - \text{④} \Rightarrow y - E(y) = \pi_1 (Z - E(Z)) + u \quad \nabla = 0.$$

$$E([y - E(y)][Z - E(Z)]) = E(\pi_1 (Z - E(Z))^2) + E(u(Z - E(Z)))$$

$$\Rightarrow E([y - E(y)][Z - E(Z)]) = E(\pi_1 (Z - E(Z))^2)$$

$$\Rightarrow \text{cov}(y, Z) = \pi_1 \text{var}(Z)$$

$$\Rightarrow \pi_1 = \frac{\text{cov}(y, Z)}{\text{var}(Z)}$$

$$C_1: Y_1 = \beta_1 + \theta_1 Z + V \quad \text{and} \quad Y = \beta_1 + \beta_2 X + e$$

$$\Rightarrow Y = \beta_1 + \beta_2(Y_1 + \theta_1 Z + V) + e$$

$$\Rightarrow Y = \frac{(\beta_1 + \beta_2 Y_1)}{\pi_0} + \frac{Z(\beta_2 \theta_1)}{\pi_1} + \frac{\beta_2 V}{\kappa} + e$$

$$d. \quad \because \pi_1 = \beta_2 \theta_1 \Rightarrow \beta_2 = \frac{\pi_1}{\theta_1}$$

$$e_1. \quad \hat{\theta}_1 = \frac{\text{cov}(z, Y)}{\text{var}(z)} = \frac{\frac{1}{n} \sum (z_i - \bar{z})(Y_i - \bar{Y})}{\frac{1}{n} \sum (z_i - \bar{z})^2} = \frac{\sum (z_i - \bar{z})(Y_i - \bar{Y})}{\sum (z_i - \bar{z})^2}$$

$$\hat{\pi}_1 = \frac{\text{cov}(z, y)}{\text{var}(z)} = \frac{\sum (z_i - \bar{z})(y_i - \bar{y})}{\sum (z_i - \bar{z})^2}$$

$$\hat{\beta}_2 = \frac{\pi_1}{\theta_1} = \frac{\hat{\text{cov}}(z, X)}{\hat{\text{cov}}(z, Y)}$$

$$\text{由大數法則可知 } \hat{\beta}_2 = \frac{\hat{\text{cov}}(z, X)}{\hat{\text{cov}}(z, Y)} \xrightarrow{P} \frac{\text{cov}(z, X)}{\text{cov}(z, Y)} = \beta_2$$