

10.2 The labor supply of married women has been a subject of a great deal of economic research. Consider the following supply equation specification

$$HOURS = \beta_1 + \beta_2 WAGE + \beta_3 EDUC + \beta_4 AGE + \beta_5 KIDSL6 + \beta_6 NWIFEINC + e$$

where  $HOURS$  is the supply of labor,  $WAGE$  is hourly wage,  $EDUC$  is years of education,  $KIDSL6$  is the number of children in the household who are less than 6 years old, and  $NWIFEINC$  is household income from sources other than the wife's employment.

a. Discuss the signs you expect for each of the coefficients.

expect  $\beta_2 > 0$ , 工資 ↑, 更願意投入工資時間, 與勞動供給正相關

expect  $\beta_3 > 0$ , 教育程度 ↑, 可能有更高薪資的工作, 更可能參與勞動市場, 工時可能 ↑

expect  $\beta_4$  not sure, 年齡較大可能工時更長, 但可能因家庭或健康問題減少工作時間

expect  $\beta_5 < 0$ , 6歲以下幼兒 ↑, 要負擔更多家務, 導致工時 ↓

expect  $\beta_6 < 0$ , 家庭其他來源收入 ↑, 經濟壓力 ↓, 勞動供給 ↓

b. Explain why this supply equation cannot be consistently estimated by OLS regression.

∵  $WAGE$  變數具有内生性, 可能受到未觀察到的個人特質影響, 這些因素同時也會影響被解釋變數 ( $HOURS$ )。

( $WAGE$  與  $e$  相關)

c. Suppose we consider the woman's labor market experience  $EXPER$  and its square,  $EXPER^2$ , to be instruments for  $WAGE$ . Explain how these variables satisfy the logic of instrumental variables.

工具變數須滿足兩個條件!

1. 相關性: 工具變數與內生變數  $WAGE$  有關
2. 外生性: ..... 不能與誤差項  $\epsilon$  相關

∴  $EXPER, EXPER^2$  與  $WAGE$  高度相關, 且可合理假設只影響工資, 不直接影響勞動供給決策

d. Is the supply equation identified? Explain.

∴ 工具變數數量: 2 > 1 : 內生變數數量  
∴ is identified

e. Describe the steps [not a computer command] you would take to obtain IV/2SLS estimates.

1. 用所有外生變數和工具變數去迴歸內生變數, 取得預測值  $\widehat{WAGE}$
2. 在原本的勞動供給方程式中以  $\widehat{WAGE}$  代替  $WAGE$ , 然後用 OLS 估計這個新方程式
3. 所得估計值即為 2SLS 的一致估計結果

**10.3** In the regression model  $y = \beta_1 + \beta_2 x + e$ , assume  $x$  is endogenous and that  $z$  is a valid instrument. In Section 10.3.5, we saw that  $\beta_2 = \text{cov}(z, y) / \text{cov}(z, x)$ .

- a. Divide the denominator of  $\beta_2 = \text{cov}(z, y) / \text{cov}(z, x)$  by  $\text{var}(z)$ . Show that  $\text{cov}(z, x) / \text{var}(z)$  is the coefficient of the simple regression with dependent variable  $x$  and explanatory variable  $z$ ,  $x = \gamma_1 + \theta_1 z + v$ . [Hint: See Section 10.2.1.] Note that this is the first-stage equation in two-stage least squares.

$$X = \gamma_1 + \theta_1 Z + V$$

$$\rightarrow E(X) = \gamma_1 + \theta_1 E(Z)$$

$$\Rightarrow X - E(X) = \theta_1 (Z - E(Z)) + V$$

同乘  $(Z - E(Z))$

$$\Rightarrow (Z - E(Z))(X - E(X)) = \theta_1 (Z - E(Z))^2 + (Z - E(Z))V$$

取期望值

$$E[(Z - E(Z))(X - E(X))] = \theta_1 E[(Z - E(Z))^2] + E[(Z - E(Z))V]$$

$\therefore$  By assumption,  $\text{cov}(Z, V) = 0$  ( $\because E(V) = 0$ )

$$\text{cov}(Z, V) = E[(Z - E(Z))(V - E(V))] = E[(Z - E(Z))V] = 0$$

$$\therefore \theta_1 = \frac{E[(Z - E(Z))(X - E(X))]}{E[(Z - E(Z))^2]} = \frac{\text{cov}(Z, X)}{\text{Var}(Z)}$$

與 OLS 估計迴歸係數公式一致

- b. Divide the numerator of  $\beta_2 = \text{cov}(z, y) / \text{cov}(z, x)$  by  $\text{var}(z)$ . Show that  $\text{cov}(z, y) / \text{var}(z)$  is the coefficient of a simple regression with dependent variable  $y$  and explanatory variable  $z$ ,  $y = \pi_0 + \pi_1 z + u$ . [Hint: See Section 10.2.1.]

similar to (a)

$$Y = \pi_0 + \pi_1 Z + u$$

$$\rightarrow E(Y) = \pi_0 + \pi_1 E(Z)$$

$$\Rightarrow Y - E(Y) = \pi_1 (Z - E(Z)) + u$$

同乘  $(Z - E(Z))$

$$(Z - E(Z))(Y - E(Y)) = \pi_1 (Z - E(Z))^2 + (Z - E(Z))u$$

取期望值

$$E[(Z - E(Z))(Y - E(Y))] = \pi_1 E[(Z - E(Z))^2] + \underbrace{E[(Z - E(Z))u]}_{=0}$$

$$\Rightarrow \pi_1 = \frac{E[(Z - E(Z))(Y - E(Y))]}{E[(Z - E(Z))^2]} = \frac{\text{Cov}(Z, Y)}{\text{Var}(Z)}$$

與 OLS 估計迴歸係數公式一致

- c. In the model  $y = \beta_1 + \beta_2 x + e$ , substitute for  $x$  using  $x = \gamma_1 + \theta_1 z + v$  and simplify to obtain  $y = \pi_0 + \pi_1 z + u$ . What are  $\pi_0$ ,  $\pi_1$ , and  $u$  in terms of the regression model parameters and error and the first-stage parameters and error? The regression you have obtained is a **reduced-form** equation.

$$Y = \beta_1 + \beta_2 X + e \quad (X \text{ 用 } \gamma_1 + \theta_1 Z + v \text{ 代入})$$

$$\begin{aligned} \Rightarrow Y &= \beta_1 + \beta_2 (\gamma_1 + \theta_1 Z + v) + e \\ &= (\beta_1 + \beta_2 \gamma_1) + \beta_2 \theta_1 Z + (\beta_2 v + e) \\ &= \pi_0 + \pi_1 Z + u \end{aligned}$$

$$\pi_0 = \beta_1 + \beta_2 \gamma_1, \quad \pi_1 = \beta_2 \theta_1, \quad u = \beta_2 v + e$$

得到  $y$  只跟  $z$  有關  $fn$ , 内生變數  $x$  被解掉

d. Show that  $\beta_2 = \pi_1/\theta_1$ .

by (c),  $\because \pi_1 = \beta_2 \theta_1, \beta_2 = \frac{\pi_1}{\theta_1}$

e. If  $\hat{\pi}_1$  and  $\hat{\theta}_1$  are the OLS estimators of  $\pi_1$  and  $\theta_1$ , show that  $\hat{\beta}_2 = \hat{\pi}_1/\hat{\theta}_1$  is a consistent estimator of  $\beta_2 = \pi_1/\theta_1$ . The estimator  $\hat{\beta}_2 = \hat{\pi}_1/\hat{\theta}_1$  is an **indirect least squares** estimator.

by (a), (b)

$\hat{\theta}_1$  是  $X = r_1 + \theta_1 z + v$  的迴歸係數估計量

$$\hat{\theta}_1 = \frac{\text{cov}(z, X)}{\text{var}(z)} \quad \text{且} \quad \hat{\theta}_1 \xrightarrow{P} \theta_1 \quad (\text{若 } z \text{ 與 } v \text{ 不相關})$$

$\hat{\pi}_1$  是  $y = \pi_0 + \pi_1 z + u$  的迴歸係數估計量

$$\hat{\pi}_1 = \frac{\text{cov}(z, y)}{\text{var}(z)} \quad \text{且} \quad \hat{\pi}_1 \xrightarrow{P} \pi_1 \quad (\text{若 } z \text{ 與 } u \text{ 不相關})$$

consider  $\hat{\beta}_2 = \frac{\hat{\pi}_1}{\hat{\theta}_1}$  (Assume  $\theta_1 \neq 0$ )

$\because \hat{\pi}_1 \xrightarrow{P} \pi_1, \hat{\theta}_1 \xrightarrow{P} \theta_1, g(a, b) = \frac{a}{b}$  is continuous

By continuous mapping Thm

$\therefore \hat{\beta}_2 = \frac{\hat{\pi}_1}{\hat{\theta}_1} \xrightarrow{P} \frac{\pi_1}{\theta_1} = \beta_2$ , this is called ILS

estimator (間接最小平方估計)

\* prior condition  $\left\{ \begin{array}{l} \theta_1 \neq 0 \quad (\text{工具變數與因變數有關}) \\ \text{cov}(z, v) = 0 \\ \text{cov}(z, u) = 0 \end{array} \right.$