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- 10.2 The labor supply of married women has been a subject of a great deal of economic research. Consider the following supply equation specification

$$HOURS = \beta_1 + \beta_2 WAGE + \beta_3 EDUC + \beta_4 AGE + \beta_5 KIDS6 + \beta_6 NWIFEINC + e$$

where $HOURS$ is the supply of labor, $WAGE$ is hourly wage, $EDUC$ is years of education, $KIDS6$ is the number of children in the household who are less than 6 years old, and $NWIFEINC$ is household income from sources other than the wife's employment.

- Discuss the signs you expect for each of the coefficients.
- Explain why this supply equation cannot be consistently estimated by OLS regression.
- Suppose we consider the woman's labor market experience $EXPER$ and its square, $EXPER^2$, to be instruments for $WAGE$. Explain how these variables satisfy the logic of instrumental variables.
- Is the supply equation identified? Explain.
- Describe the steps [not a computer command] you would take to obtain IV/2SLS estimates.

a. $\beta_2 (WAGE)$: 預期 +, 時薪越高, 勞動供給↑。

$\beta_3 (EDUC)$: 預期 +, 受教育程度↑, 工時↑。

$\beta_4 (AGE)$: 預期不確定, 年輕時可能工時為正, 但年紀大了可能負向影響工時。

$\beta_5 (KIDS6)$: 預期 -, 幼兒數量↑, 婦女外出工作↓。

$\beta_6 (NWIFEINC)$: 預期 -, 其他收入↑, 經濟壓力較低, 勞動供給↓。

b. : $WAGE$ 是內生變數, 可能和誤差項 (e) 有關。

- $WAGE$ 並非隨機決定, 勞動供給 ($HOURS$) 同時影響其工資 ($WAGE$).
- e 也可能包含工作意願、個人特質、健康等, 這些因素會和 $WAGE$ 有關。

c. IV 必需符合以下兩條件：

① IV 和內生變數高度相關

: $EXPER$ 累積越多會影響個人的 $WAGE$ 水準；

$EXPER^2$ 隨著年資增加, 可能邊際效益下降。

② IV 和 e 無關

: $EXPER$ 和 $EXPER^2$ 主要影響 $WAGE$

而不會直接影響 $HOURS$ 。

d. 該方程式 identified. : 工具變數 (2個) 數量大於 內生變數 (1個)
且 IV 和 內生變數強相關, 和 e 無關

e. First Stage: 用 $EXPER$ 和 $EXPER^2$ 迴歸 $WAGE$, 得到 $WAGE$ 的預測值 \widehat{WAGE}

Second Stage: 將 \widehat{WAGE} 代入原方程式, 用 \widehat{WAGE} 、 $EDUC$ 、 AGE 、 $KIDS6$ 、 $NWIFEINC$ 迴歸 $HOURS$

10.3 In the regression model $y = \beta_1 + \beta_2 x + e$, assume x is endogenous and that z is a valid instrument. In Section 10.3.5, we saw that $\beta_2 = \text{cov}(z, y)/\text{cov}(z, x)$.

- Divide the denominator of $\beta_2 = \text{cov}(z, y)/\text{cov}(z, x)$ by $\text{var}(z)$. Show that $\text{cov}(z, x)/\text{var}(z)$ is the coefficient of the simple regression with dependent variable x and explanatory variable z : $x = \gamma_1 + \theta_1 z + v$. [Hint: See Section 10.2.1.] Note that this is the first-stage equation in two-stage least squares.
- Divide the numerator of $\beta_2 = \text{cov}(z, y)/\text{cov}(z, x)$ by $\text{var}(z)$. Show that $\text{cov}(z, y)/\text{var}(z)$ is the coefficient of a simple regression with dependent variable y and explanatory variable z : $y = \pi_0 + \pi_1 z + u$. [Hint: See Section 10.2.1.]
- In the model $y = \beta_1 + \beta_2 x + e$, substitute for x using $x = \gamma_1 + \theta_1 z + v$ and simplify to obtain $y = \pi_0 + \pi_1 z + u$. What are π_0 , π_1 , and u in terms of the regression model parameters and error and the first-stage parameters and error? The regression you have obtained is a **reduced-form** equation.
- Show that $\beta_2 = \pi_1/\theta_1$.
- If $\hat{\pi}_1$ and $\hat{\theta}_1$ are the OLS estimators of π_1 and θ_1 , show that $\hat{\beta}_2 = \hat{\pi}_1/\hat{\theta}_1$ is a consistent estimator of $\beta_2 = \pi_1/\theta_1$. The estimator $\hat{\beta}_2 = \hat{\pi}_1/\hat{\theta}_1$ is an **indirect least squares** estimator.

a. $X = r_1 + \theta_1 Z + V$

$$E(X) = r_1 + \theta_1 E(Z)$$

$$X - E(X) = \theta_1 (Z - E(Z)) + V$$

$$(Z - E(Z))(X - E(X)) = \theta_1 (Z - E(Z))^2 + V(Z - E(Z))$$

$$E[(Z - E(Z))(X - E(X))] = \theta_1 E[(Z - E(Z))^2] + V E(Z - E(Z))$$

Assuming $V E(Z - E(Z)) = 0$, we get $E[(Z - E(Z))(X - E(X))] = \theta_1 E[(Z - E(Z))^2]$

$$\text{Solve } \theta_1 = \frac{E[(Z - E(Z))(X - E(X))]}{E[(Z - E(Z))^2]} = \frac{\text{cov}(Z, X)}{\text{var}(Z)} \#$$

This is the OLS estimator of θ_1 in the 1st regression $X = r_1 + \theta_1 Z + V$.

b.

$$y = \pi_0 + \pi_1 Z + U$$

$$E(y) = \pi_0 + \pi_1 E(Z)$$

$$Y - E(Y) = \pi_0 + \pi_1 (Z - E(Z)) + U$$

$$(Z - E(Z))(Y - E(Y)) = \pi_1 (Z - E(Z))^2 + U (Z - E(Z))$$

$$E[(Z - E(Z))(Y - E(Y))] = \pi_1 E[(Z - E(Z))^2] + U E(Z - E(Z))$$

Assuming $U E(Z - E(Z)) = 0$, we get $E[(Z - E(Z))(Y - E(Y))] = \pi_1 E[(Z - E(Z))^2]$

$$\text{Solve } \pi_1 = \frac{E[(Z - E(Z))(Y - E(Y))]}{E[(Z - E(Z))^2]} = \frac{\text{cov}(Z, Y)}{\text{var}(Z)}$$

This is the OLS estimator of π_1 in the regression $y = \pi_0 + \pi_1 Z + U$.

$$C. \quad y = \beta_1 + \beta_2 x + e$$

x 用 1st regression : $x = r_1 + \theta_1 z + v$. 代入原模型

$$y = \beta_1 + \beta_2(r_1 + \theta_1 z + v) + e$$

$$= (\beta_1 + \beta_2 r_1) + \beta_2 \theta_1 z + (\beta_2 v + e)$$

對應 $y = \pi_0 + \pi_1 z + u$, 得 $\begin{cases} \pi_0 = \beta_1 + \beta_2 r_1 \\ \pi_1 = \beta_2 \theta_1 \\ u = \beta_2 v + e \end{cases}$

$$d. \quad \pi_1 = \beta_2 \theta_1 \Rightarrow \beta_2 = \frac{\pi_1}{\theta_1}$$

$$e. \quad \theta_1 = \frac{E[(z - E(z))(x - E(x))]}{E[(z - E(z))^2]} = \frac{\text{cov}(z, x)}{\text{var}(z)}$$

$$\hat{\theta}_1 = \frac{\widehat{\text{cov}}(z, x)}{\widehat{\text{var}}(z)} = \frac{\frac{\sum (z_i - \bar{z})(x_i - \bar{x})}{N}}{\frac{\sum (z_i - \bar{z})^2}{N}} = \frac{\sum (z_i - \bar{z})(x_i - \bar{x})}{\sum (z_i - \bar{z})^2}$$

This estimator is consistent if z is uncorrelated with v .

$$\pi_1 = \frac{E[(z - E(z))(y - E(y))]}{E(z - E(z))^2} = \frac{\text{cov}(z, y)}{\text{var}(z)}$$

$$\hat{\pi}_1 = \frac{\widehat{\text{cov}}(z, y)}{\widehat{\text{var}}(z)} = \frac{\frac{\sum (z_i - \bar{z})(y_i - \bar{y})}{N}}{\frac{\sum (z_i - \bar{z})^2}{N}} = \frac{\sum (z_i - \bar{z})(y_i - \bar{y})}{\sum (z_i - \bar{z})^2}$$

This estimator is consistent if z is uncorrelated with u .

Next.

$$\beta_2 = \frac{x_i}{\theta_1}$$

$$\hat{\beta}_2 = \frac{\hat{x}_i}{\hat{\theta}_1}$$

$$\frac{\left[\frac{\sum(z_i - \bar{z})(y_i - \bar{y})}{\sum(z_i - \bar{z})^2} \right]}{\left[\frac{\sum(z_i - \bar{z})(x_i - \bar{x})}{\sum(z_i - \bar{z})^2} \right]} = \frac{\sum(z_i - \bar{z})(y_i - \bar{y})}{\sum(z_i - \bar{z})(x_i - \bar{x})}$$

$$= - \frac{\frac{\sum(z_i - \bar{z})(y_i - \bar{y})}{N}}{\frac{\sum(z_i - \bar{z})(x_i - \bar{x})}{N}} = \frac{\hat{\text{cov}}(z, y)}{\hat{\text{cov}}(z, x)}$$

當 $N \rightarrow \infty$, $\hat{\text{cov}}(z, y) \xrightarrow{P} \text{cov}(z, y)$; $\hat{\text{cov}}(z, x) \xrightarrow{P} \text{cov}(z, x)$

$$\therefore \hat{\beta}_2 = \frac{\hat{\text{cov}}(z, y)}{\hat{\text{cov}}(z, z)} \xrightarrow{P} \beta_2 = \frac{\text{cov}(z, y)}{\text{cov}(z, z)}$$