- f. Column (5) contains the random effects estimates. Which coefficients, apart from the intercepts, show the most difference from the fixed effects estimates? Use the Hausman test statistic (15.36) to test whether there are significant differences between the random effects estimates and the fixed effects estimates in column (3) (Why that one?). Based on the test results, is random effects estimation in this model appropriate?
 - EXPER (株教美了 -0.0023 = 1.92倍

$$t_{\text{EXPER}} = \frac{0.0575 - 0.0984}{\sqrt{0.033^2 - 0.032^2}} = -1.67$$

$$t_{\text{EXPER}} = \frac{-0.0012 - (-0.0023)}{\sqrt{0.0011^2 - 0.0007^2}} = 1.296$$

tsouth =
$$\frac{-0.3261 - (-0.2326)}{\sqrt{0.1258^2 - 0.0317^2}} = -0.77$$

只有EXPER的係款在10% 顯著水準上有顯著差異

- ラ 因此內生性證據相當導弱
- 习敌在此例 採用隨机效果估計是合通的,

b. Estimate the model $LIQUOR_{it} = \beta_1 + \beta_2 INCOME_{it} + u_i + e_{it}$ using random effects. Construct a 95% interval estimate of the coefficient on *INCOME*. How does it compare to the interval in part (a)?

```
Call:
                                                       > confint(random mod)
plm(formula = liquor ~ income, data = pdat, model = "random")
                                                                           2.5 %
                                                                                     97.5 %
Balanced Panel: n = 40, T = 3, N = 120
                                                       (Intercept) -0.05211904 1.99018381
                                                                     0.01283111 0.04031983
                                                       income
Effects:
             var std.dev share
idiosyncratic 0.9640 0.9819 0.571
individual 0.7251 0.8515 0.429
                                                         INCOME(a) 係數的 95% 信賴區間非常寬,且包
theta: 0.4459
                                                         含 0,表示統計上並不顯著。
Residuals:
         1st Qu.
                  Median
                         3rd Ou.
-2.263634 -0.697383 0.078697 0.552680 2.225798
                                                         而在(b) 部分的隨機效果(Random Effects)
                                                         模型下,標準誤大幅下降至 0.00701,對應的
Coefficients:
          Estimate Std. Error z-value Pr(>|z|)
                                                         95% 信賴區間也明顯縮小為 [0.0128, 0.0403],
(Intercept) 0.9690324 0.5210052 1.8599 0.0628957 .
                                                         且不包含 0。
        0.0265755 0.0070126 3.7897 0.0001508 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
                                                         因此,隨機效果模型提供了強而有力的證據,我
                                                         們以 95% 的信心水準估計,每增加 1000 美元的
Total Sum of Squares:
                                                         收入,家庭在酒類上的支出將增加 12.83 至
Residual Sum of Squares: 112.88
R-Squared:
            0.1085
                                                         40.32 美元之間。
Adj. R-Squared: 0.10095
Chisq: 14.3618 on 1 DF, p-value: 0.00015083
```

LIQUOR; = 0.96903+ 0.02658 INCOME;

c. Test for the presence of random effects using the LM statistic in equation (15.35). Use the 5% level of significance.

在宏毅為水準下級界值為±1.96 把絕H。: cu²=0, 接受 H;: cu²>0 表示存在 統計上點書的 未觀察異質性。 **d.** For each individual, compute the time averages for the variable *INCOME*. Call this variable *INCOMEM*. Estimate the model $LIQUOR_{ii} = \beta_1 + \beta_2 INCOME_{ii} + \gamma INCOMEM_i + c_i + e_{ii}$ using the random effects estimator. Test the significance of the coefficient γ at the 5% level. Based on this test, what can we conclude about the correlation between the random effect u_i and *INCOME*? Is it OK to use the random effects estimator for the model in (b)?

```
Call:
plm(formula = liquor ~ income + INCOMEM, data = pdat1, model = "random")
Balanced Panel: n = 40, T = 3, N = 120
Effects:
                var std.dev share
idiosyncratic 0.9640 0.9819 0.571
           0.7251 0.8515 0.429
individual
theta: 0.4459
Residuals:
    Min.
          1st Qu.
                    Median 3rd Qu.
                                          Max.
-2.300955 -0.703840 0.054992 0.560255 2.257325
Coefficients:
            Estimate Std. Error z-value Pr(>|z|)
(Intercept) 0.9163337 0.5524439 1.6587 0.09718 .
          0.0207421 0.0209083 0.9921 0.32117
income
INCOMEM 0.0065792 0.0222048 0.2963 0.76700
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Total Sum of Squares: 126.61
Residual Sum of Squares: 112.79
             0.10917
R-Squared:
Adj. R-Squared: 0.093945
Chisq: 14.3386 on 2 DF, p-value: 0.00076987
```

INCOMEM 係款 Y 不 顯著,放根據 Mundlak 檢定結果 沒有證據 顯示收入與未觀察的異鹽性之間存在相関性。

i Land A Land A

b. Reestimate the model in part (a) with school fixed effects. Compare the results with those in part (a). Have any of your conclusions changed? [*Hint*: specify *SCHID* as the cross-section identifier and *ID* as the "time" identifier.]

> summary(FE model)

```
Oneway (individual) effect Within Model
 Call:
 plm(formula = readscore ~ small + aide + tchexper + boy + white asian +
    freelunch, data = pa data, effect = "individual", model = "within")
 Unbalanced Panel: n = 79, T = 34-137, N = 5766
 Residuals:
    Min. 1st Qu. Median 3rd Qu.
                                    Max.
 -102.6381 -16.7834 -2.8473 12.7591 198.4169
 Coefficients:
            Estimate Std. Error t-value Pr(>|t|)
 small
           6.490231 0.912962 7.1090 1.313e-12 ***
           0.996087 0.881693 1.1297 0.2586
          0.285567 0.070845 4.0309 5.629e-05 ***
 tchexper
          -5.455941 0.727589 -7.4987 7.440e-14 ***
 white_asian 8.028019 1.535656 5.2277 1.777e-07 ***
 freelunch -14.593572 0.880006 -16.5835 < 2.2e-16 ***
 Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 '' 1
 控制了學校特定的因素,小班教學的優勢實際上變得更明顯,且仍然具有高度顯著性。
 教學經驗的效果在數值上有所下降,表示 OLS 所觀察到的一部分效果可能來自於學校
 之間的差異,但其估計仍為正值且高度顯著。
 男孩的成績仍然比女孩低,這是在各校內部也持續存在的一個高度顯著差距。
 控制學校的異質性之後,種族 / 族群的優勢幾乎加倍,這顯示出在同一所學校內部的差
 異甚至比合併樣本所呈現的估計還要大。
c. Test for the significance of the school fixed effects. Under what conditions would we expect the
```

c. Test for the significance of the school fixed effects. Under what conditions would we expect the inclusion of significant fixed effects to have little influence on the coefficient estimates of the remaining variables? 如写情况下 我們預期 秋 服著的 固定效果料 其他发权估计影响小?

```
remaining variables? 可以有水子款价款明 納入 股者的 固定效果 有其他发权估计的向小
> pFtest(FE_model, pool_model)

F test for individual effects

data: readscore ~ small + aide + tchexper + boy + white_asian + freelunch
F = 16.698, dfl = 78, df2 = 5681, p-value < 2.2e-16
alternative hypothesis: significant effects
```

拒絕虛無假設,拒絕學校之間沒有顯著差異的說法。 如果學校指標變數與已包含的解釋變數之間沒有相關性,那麼將這些學校指標變數納入 或排除回歸模型,對回歸估計結果應該影響不大 d. Reestimate the model in part (a) with school random effects. Compare the results with those from parts (a) and (b). Are there any variables in the equation that might be correlated with the school effects? Use the LM test for the presence of random effects.

```
plm(formula = readscore ~ small + aide + tchexper + boy + white asian +
   freelunch, data = pa data, model = "random")
Unbalanced Panel: n = 79, T = 34-137, N = 5766
Effects:
               var std.dev share
idiosyncratic 751.43 27.41 0.829
            155.31 12.46 0.171
individual
theta:
  Min. 1st Qu. Median Mean 3rd Qu.
0.6470 0.7225 0.7523 0.7541 0.7831 0.8153
Residuals:
  Min. 1st Ou. Median Mean 3rd Ou.
-97.483 -17.236 -3.282 0.037 12.803 192.346
Coefficients:
            Estimate Std. Error z-value Pr(>|z|)
(Intercept) 436.126774 2.064782 211.2217 < 2.2e-16 ***
            6.458722 0.912548 7.0777 1.466e-12 ***
small
aide
            0.992146 0.881159 1.1260 0.2602
            0.302679 0.070292 4.3060 1.662e-05 ***
tchexper
           -5.512081 0.727639 -7.5753 3.583e-14 ***
white asian 7.350477 1.431376 5.1353 2.818e-07 ***
freelunch -14.584332 0.874676 -16.6740 < 2.2e-16 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 '' 1
```

整體而言,隨機效果 估計結果看起來與 OLS 估計結果大致相 同,但與與固定效果 估計結果更為相似

Total Sum of Squares: 6158000

```
+ readscore ~ small + aide + tchexper + boy + white_asian + freelunch,
+ data = pa_data,
+ type = "bp",
+ effect = "individual"
+)

Lagrange Multiplier Test - (Breusch-Pagan)

data: readscore ~ small + aide + tchexper + boy + white_asian + freelunch
chisq = 6677.4, df = 1, p-value < 2.2e-16
alternative hypothesis: significant effects
```

Breusch-Pagan 拉格朗日乘數檢 定顯示出顯著的統計結果,因此 我們拒絕「沒有隨機效果」的虛 無假設。換句話說,閱讀成績中 存在高度顯著的學校層級異質 性,這證實相較於單純的合併 OLS,應使用隨機效果模型。 e. Using the *t*-test statistic in equation (15.36) and a 5% significance level, test whether there are any significant differences between the fixed effects and random effects estimates of the coefficients on *SMALL*, *AIDE*, *TCHEXPER*, *WHITE_ASIAN*, and *FREELUNCH*. What are the implications of the test outcomes? What happens if we apply the test to the fixed and random effects estimates of the coefficient on *BOY*?

```
Hausman Test
```

```
data: readscore ~ small + aide + tchexper + boy + white_asian + freelunch
chisq = 13.809, df = 6, p-value = 0.03184
alternative hypothesis: one model is inconsistent
```

> print(t_values)

於,有時小於真實變異數。

Model 1: restricted model

Total Sum of Squares:

Residual Sum of Squares: 4281300

```
small aide tchexper white_asian freelunch
1.14600764 0.12843803 -1.93771666 1.21807432 -0.09555102
```

只有教師經驗係數的差異具有顯著性。 而這項檢定可能出現的問題可以從 BOY 變數的檢定中看出,固定效果的標準誤小於隨 機效果的標準誤,因此分母會包含負數的平方根,這種情況產生的原因是標準誤為估 計變異數的平方根,而像所有估計量一樣,變異數估計值會呈現抽樣變動,有時大

Hausman 對比統計量(Hausman contrast statistic),用來聯合檢定所有係數,其統計量大於這六個係數的比較臨界值為12.592 因此我們拒絕虛無假設—即「未觀察異質性與解釋變數之間沒有相關性」。根據這項

因此我們拒絕虛無假設—即「未觀察異質性與解釋變數之間沒有相關性」。根據這項 檢定結果,不建議採用隨機效果估計量。

f. Create school-averages of the variables and carry out the Mundlak test for correlation between them and the unobserved heterogeneity.

```
small_avg + aide_avg + tchexper_avg + boy_avg + white_asian_avg +
    freelunch avg
Note: Coefficient covariance matrix supplied.
  Res.Df Df
                 F Pr(>F)
 1 5695
 2 5689 6 2.2541 0.03557 *
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Coefficients:
                   Estimate Std. Error z-value Pr(>|z|)
(Intercept) 459.462989 20.529888 22.3802 < 2.2e-16 ***
              6.637460 0.922068 7.1985 6.090e-13 ***
small
aide
                   1.157620 0.889542 1.3014 0.1931
freelunch -14.699025 0.892109 -16.4767 < 2.2e-16 ***

        small m
        -18.410060
        22.273923
        -0.8265
        0.4085

        aide m
        16.811358
        20.793685
        0.8085
        0.4188

        tchexper m
        1.006007
        0.625690
        1.6078
        0.1079

        boy m
        -53.353521
        25.221654
        -2.1154
        0.0344

white_asian_m -6.648191 6.320012 -1.0519 0.2928
freelunch m -3.318853 8.779553 -0.3780 0.7054
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

6007200

Model 2: readscore ~ small + aide + tchexper + boy + white asian + freelunch +

BoYM的係該在5% 毀著 水準下是毀裝的。 對6個推鑽均變該的联合毀舊性檢定 PrOF) = 0.03557, 因此 把絕 H·, 結論收與 e 题 都分同。