

10.2 The labor supply of married women has been a subject of a great deal of economic research. Consider the following supply equation specification

$$HOURS = \beta_1 + \beta_2 WAGE + \beta_3 EDUC + \beta_4 AGE + \beta_5 KIDSL6 + \beta_6 NWIFEINC + e$$

where $HOURS$ is the supply of labor, $WAGE$ is hourly wage, $EDUC$ is years of education, $KIDSL6$ is the number of children in the household who are less than 6 years old, and $NWIFEINC$ is household income from sources other than the wife's employment.

- Discuss the signs you expect for each of the coefficients.
- Explain why this supply equation cannot be consistently estimated by OLS regression.
- Suppose we consider the woman's labor market experience $EXPER$ and its square, $EXPER^2$, to be instruments for $WAGE$. Explain how these variables satisfy the logic of instrumental variables.
- Is the supply equation identified? Explain.
- Describe the steps [not a computer command] you would take to obtain IV/2SLS estimates.

a. $\beta_2 (WAGE) \rightarrow \oplus$, 工資提高, 激勵更多勞動

$\beta_3 (EDUC) \rightarrow$ 不確定, 教育程度愈高 \rightarrow 能力強, 有意願進入職場 \rightarrow 增加勞動供給

\rightarrow 效率高, 對時間彈性要求高, 可能投入其他地方 \rightarrow 減少勞動供給

$\beta_4 (AGE) \rightarrow$ 不確定, 年齡增加, 可能增加經驗 (提升供給), 也可能因為健康 (降低供給)

$\beta_5 (KIDSL6) \rightarrow \ominus$, 家中兒童越多, 負擔越重, 供給下降

$\beta_6 (NWIFEINC) \rightarrow \ominus$, 其他來源的收入高, 減少女性進入勞動市場的誘因

b. Endogeneity 內生性問題

Hours \longleftrightarrow WAGES 都由供需決定, 會造成估計失效

c. Instrumental variable

- ① Relevance: $EXPER$, $EXPER^2$ 和 $WAGE$ 通常顯著相關
- ② Exogeneity: $EXPER$, $EXPER^2$ 透過 $WAGE$ 影響 Hours, 所以和誤差不相關

d. Yes, 只有一個內生變數 ($WAGE$) 且至少有一個工具變數可使用

e. $WAGE = \gamma_1 + \gamma_2 EDUC + \gamma_3 AGE + \gamma_4 KIDSL6 + \gamma_5 KIDSL6^2 + \gamma_6 NWIFEINC + \alpha_1 EXPER + \alpha_2 EXPER^2 + u$

先對上面的式子回歸, 得到 \hat{WAGE} , 將原來的 $WAGE$ 用 \hat{WAGE} 替換, 再用 OLS 估計

10.3 In the regression model $y = \beta_1 + \beta_2 x + e$, assume x is endogenous and that z is a valid instrument. In Section 10.3.5, we saw that $\beta_2 = \text{cov}(z, y) / \text{cov}(z, x)$.

- Divide the denominator of $\beta_2 = \text{cov}(z, y) / \text{cov}(z, x)$ by $\text{var}(z)$. Show that $\text{cov}(z, x) / \text{var}(z)$ is the coefficient of the simple regression with dependent variable x and explanatory variable z , $x = \gamma_1 + \theta_1 z + v$. [Hint: See Section 10.2.1.] Note that this is the first-stage equation in two-stage least squares.
- Divide the numerator of $\beta_2 = \text{cov}(z, y) / \text{cov}(z, x)$ by $\text{var}(z)$. Show that $\text{cov}(z, y) / \text{var}(z)$ is the coefficient of a simple regression with dependent variable y and explanatory variable z , $y = \pi_0 + \pi_1 z + u$. [Hint: See Section 10.2.1.]
- In the model $y = \beta_1 + \beta_2 x + e$, substitute for x using $x = \gamma_1 + \theta_1 z + v$ and simplify to obtain $y = \pi_0 + \pi_1 z + u$. What are π_0 , π_1 , and u in terms of the regression model parameters and error and the first-stage parameters and error? The regression you have obtained is a **reduced-form** equation.
- Show that $\beta_2 = \pi_1 / \theta_1$.
- If $\hat{\pi}_1$ and $\hat{\theta}_1$ are the OLS estimators of π_1 and θ_1 , show that $\hat{\beta}_2 = \hat{\pi}_1 / \hat{\theta}_1$ is a consistent estimator of $\beta_2 = \pi_1 / \theta_1$. The estimator $\hat{\beta}_2 = \hat{\pi}_1 / \hat{\theta}_1$ is an **indirect least squares** estimator.

a. $x = \gamma_1 + \theta_1 z + v \Rightarrow$ 取期望值 $E(x) = \gamma_1 + \theta_1 E(z)$

$$\begin{aligned} x &= \gamma_1 + \theta_1 z + v \\ -) E(x) &= \gamma_1 + \theta_1 E(z) \\ \hline x - E(x) &= \theta_1 (z - E(z)) + v \end{aligned}$$

同乘 $z - E(z) \Rightarrow (z - E(z))(x - E(x)) = \theta_1 (z - E(z))^2 + v(z - E(z))$

同取期望值: $E[(z - E(z))(x - E(x))] = \theta_1 E[(z - E(z))^2]$

$$\Rightarrow \theta_1 = \frac{\text{cov}(z, x)}{\text{var}(z)}$$

b. $y = \pi_0 + \pi_1 z + u \Rightarrow$ 同 a. $E(y) = \pi_0 + \pi_1 E(z)$

$$\begin{aligned} y &= \pi_0 + \pi_1 z + u \\ -) E(y) &= \pi_0 + \pi_1 E(z) \\ \hline y - E(y) &= \pi_1 (z - E(z)) + u \end{aligned}$$

$v(z - E(z)) \Rightarrow E[(y - E(y))(z - E(z))] = \pi_1 E[(z - E(z))^2]$
取期望值

$$\Rightarrow \pi_1 = \frac{\text{cov}(z, y)}{\text{var}(z)}$$

c. $x = \gamma_1 + \theta_1 z + v$ 代入 $y = \beta_1 + \beta_2 x + e$

$$\begin{aligned} y &= \beta_1 + \beta_2 (\gamma_1 + \theta_1 z + v) + e \\ &= (\beta_1 + \beta_2 \gamma_1) + \beta_2 \theta_1 z + (\beta_2 v + e) \end{aligned}$$

对 π_1 reduced-form $y = \pi_0 + \pi_1 z + u$

$$\Rightarrow \pi_0 = \beta_1 + \beta_2 \gamma_1$$

$$\pi_1 = \beta_2 \theta_1$$

$$u = \beta_2 v + e$$

$$d. \pi_1 = \beta_2 \theta_1 \Rightarrow \beta_2 = \frac{\pi_1}{\theta_1}$$

e. by (a)

$$\begin{aligned} \hat{\theta}_1 &= \frac{\hat{\text{cov}}(z, x)}{\hat{\text{var}}(z)} = \frac{\sum (z_i - \bar{z})(x_i - \bar{x}) / N}{\sum (z_i - \bar{z})^2 / N} \\ &= \frac{\sum (z_i - \bar{z})(x_i - \bar{x})}{\sum (z_i - \bar{z})^2} \end{aligned}$$

$$\begin{aligned} \text{by (b)} \quad \hat{\pi}_1 &= \frac{\hat{\text{cov}}(z, y)}{\hat{\text{var}}(z)} = \frac{\sum (z_i - \bar{z})(y_i - \bar{y}) / N}{\sum (z_i - \bar{z})^2 / N} \\ &= \frac{\sum (z_i - \bar{z})(y_i - \bar{y})}{\sum (z_i - \bar{z})^2} \end{aligned}$$

$$\begin{aligned} \hat{\beta}_2 &= \frac{\hat{\pi}_1}{\hat{\theta}_1} = \frac{\sum (z_i - \bar{z})(y_i - \bar{y})}{\sum (z_i - \bar{z})(x_i - \bar{x})} = \frac{\sum (z_i - \bar{z})(y_i - \bar{y}) / N}{\sum (z_i - \bar{z})(x_i - \bar{x}) / N} \\ &= \frac{\hat{\text{cov}}(z, y)}{\hat{\text{cov}}(z, x)} \end{aligned}$$

$$\Rightarrow \hat{\beta}_2 \text{ is IV estimator}$$

consistent

$$\therefore \hat{\text{cov}}(z, y) \xrightarrow{P} \text{cov}(z, y)$$

$$\hat{\text{cov}}(z, x) \xrightarrow{P} \text{cov}(z, x)$$

$$\hat{\beta}_2 = \frac{\hat{\pi}_1}{\hat{\theta}_1} = \frac{\hat{\text{cov}}(z, y)}{\hat{\text{cov}}(z, x)} \xrightarrow{P} \frac{\text{cov}(z, y)}{\text{cov}(z, x)} = \beta_2$$