

431 Class 24

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Today's Agenda

- What Happened Over the Break
- Regression Modeling using the National Youth Fitness Survey data
 - Box-Cox to identify potential outcome transformations
 - Adjusted R^2 and when that indicates severe overfitting
 - Testing/Pruning a Kitchen Sink Model
 - Hypothesis testing approaches, t and F tests
 - Using step for variable selection using AIC
 - Identification of meaningful collinearity with `vif`
 - Checking Assumptions with Residual Plots
 - Summarizing a Model: Drawing Conclusions

Today's R Setup and Data Set

```
library(car); library(magrittr)
library(broom); library(tidyverse)
```

-- Attaching packages -----

```
v ggplot2 2.2.1      v purrr  0.2.4
v tibble  1.3.4      v dplyr  0.7.4
v tidyr   0.7.2      v stringr 1.2.0
v readr   1.1.1      v forcats 0.2.0
```

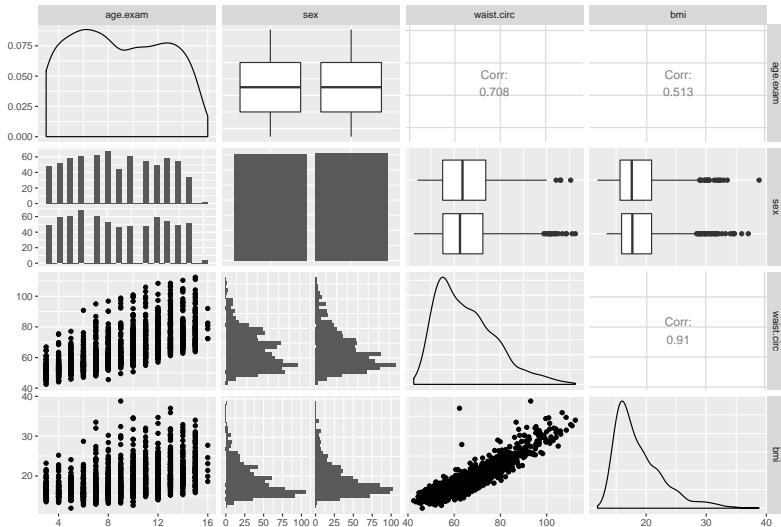
-- Conflicts -----

```
x tidyr::extract() masks magrittr::extract()
x dplyr::filter()  masks stats::filter()
x dplyr::lag()     masks stats::lag()
x dplyr::recode()  masks car::recode()
x purrr::set_names() masks magrittr::set_names()
x purrr::some()    masks car::some()
```

Model `m2` for the `nyfs1` data: Predicting `bmi`
using `waist.circ` as well as `age.exam` and `sex`

Scatterplot Matrix for some NYFS1 variables

Scatterplot Matrix for nyfs1 data



Building Model 2

```
m2 <- nyfs1 %$%  
  lm(bmi ~ waist.circ + age.exam + sex)  
  
m2
```

Call:

```
lm(formula = bmi ~ waist.circ + age.exam + sex)
```

Coefficients:

(Intercept)	waist.circ	age.exam	sexMale
-1.4452	0.3484	-0.2932	0.1881

Summary of Model m2 (rearranged a little)

Call: `lm(formula = bmi ~ waist.circ + age.exam + sex)`

Multiple R-squared: 0.8635, Adjusted R-squared: 0.8633
F-statistic: 2979 on 3 and 1412 DF, p-value: < 2.2e-16

Coefficients:	Estimate	Std. Error	t value	Pr(> t)	
(Intercept)	-1.445179	0.222018	-6.509	1.05e-10	***
waist.circ	0.348370	0.004421	78.801	< 2e-16	***
age.exam	-0.293247	0.015440	-18.992	< 2e-16	***
sexMale	0.188094	0.080208	2.345	0.0192	*

Sig. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residuals:	Min	1Q	Median	3Q	Max
	-4.0232	-0.8879	-0.0802	0.7836	20.3659

Residual standard error: 1.509 on 1412 degrees of freedom

Summary of m1 (for reference, rearranged)

Call: `lm(formula = bmi ~ waist.circ)`

Multiple R-squared: 0.8282, Adjusted R-squared: 0.828
F-statistic: 6814 on 1 and 1414 DF, p-value: < 2.2e-16

Coefficients:	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	-0.06646	0.23292	-0.285	0.775
waist.circ	0.28893	0.00350	82.548	<2e-16 ***

Sig. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residuals:	Min	1Q	Median	3Q	Max
	-4.2343	-1.0941	-0.0744	0.9254	19.0664

Residual standard error: 1.692 on 1414 degrees of freedom

95% Confidence Intervals for m_2 Coefficients

```
confint(m2, level = 0.95)
```

	2.5 %	97.5 %
(Intercept)	-1.88069873	-1.0096587
waist.circ	0.33969754	0.3570418
age.exam	-0.32353535	-0.2629586
sexMale	0.03075409	0.3454346

Tidying the Coefficients (with broom::tidy)

Places the coefficient summary into a tibble.

```
tidy(m2) ## from broom package
```

	term	estimate	std.error	statistic
1	(Intercept)	-1.4451787	0.222017690	-6.509295
2	waist.circ	0.3483697	0.004420855	78.801422
3	age.exam	-0.2932470	0.015440287	-18.992327
4	sexMale	0.1880943	0.080208291	2.345073

	p.value
1	1.046338e-10
2	0.000000e+00
3	8.257321e-72
4	1.916119e-02

Model Summaries, at a glance (with broom::glance)

```
glance(m2) ## also from broom
```

	r.squared	adj.r.squared	sigma	statistic	p.value
1	0.8635437	0.8632538	1.509103	2978.545	0

	df	logLik	AIC	BIC	deviance	df.residual
1	4	-2589.92	5189.84	5216.118	3215.678	1412

```
glance(m1) ## for comparison
```

	r.squared	adj.r.squared	sigma	statistic	p.value
1	0.8281523	0.8280307	1.692336	6814.214	0

	df	logLik	AIC	BIC	deviance	df.residual
1	2	-2753.188	5512.376	5528.143	4049.7	1414

Augmenting the Data with Model Results

(broom::augment) run with warning = FALSE

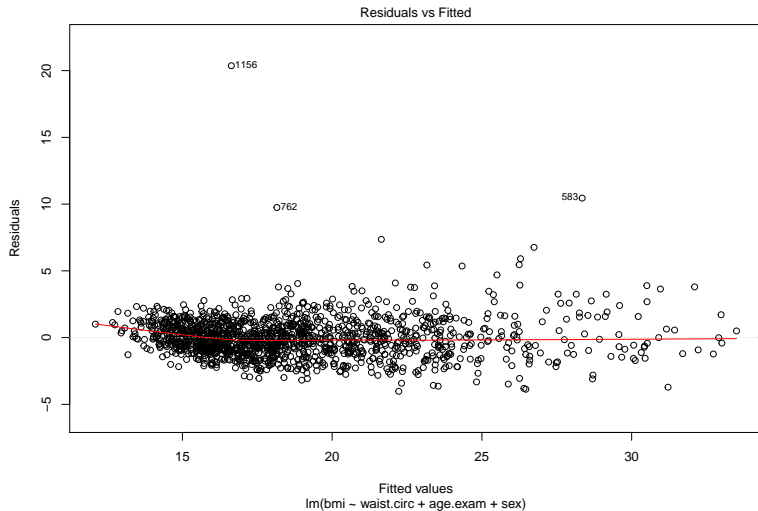
```
newdat2 <- augment(m2) ## again from broom  
head(newdat2,3)
```

	bmi	waist.circ	age.exam	sex	.fitted	.se.fit
1	22.3	71.9	8	Female	21.25663	0.06927752
2	19.8	79.4	14	Female	22.10992	0.08007933
3	15.2	46.8	3	Male	14.16688	0.08724867

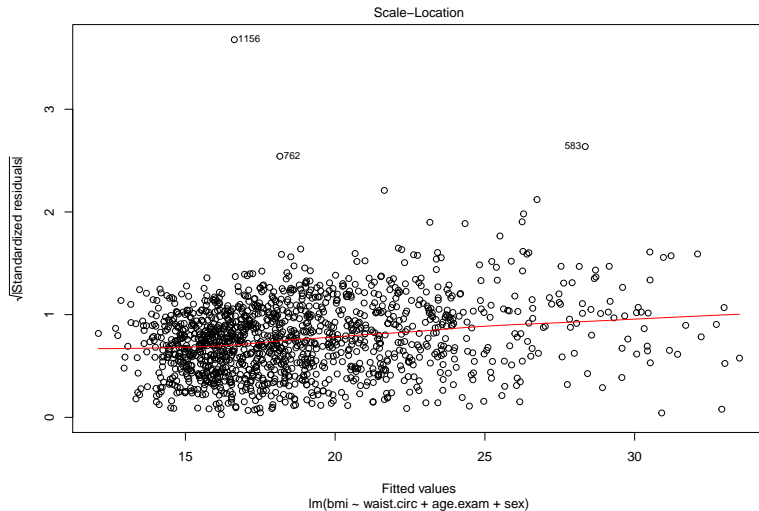
	.resid	.hat	.sigma	.cooksd
1	1.043374	0.002107399	1.509382	0.0002529070
2	-2.309917	0.002815807	1.508381	0.0016586209
3	1.033124	0.003342564	1.509387	0.0003942708

	.std.resid
1	0.6921161
2	-1.5328151
3	0.6857415

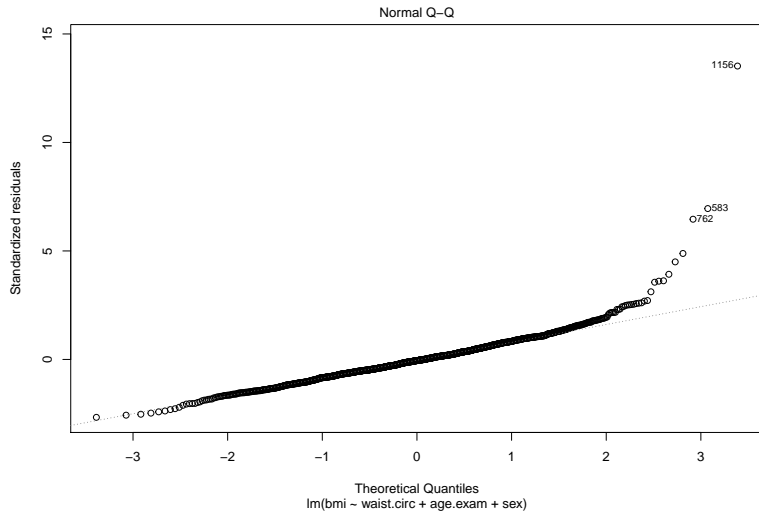
Residuals vs. Fitted Values plot(m2, which = 1)



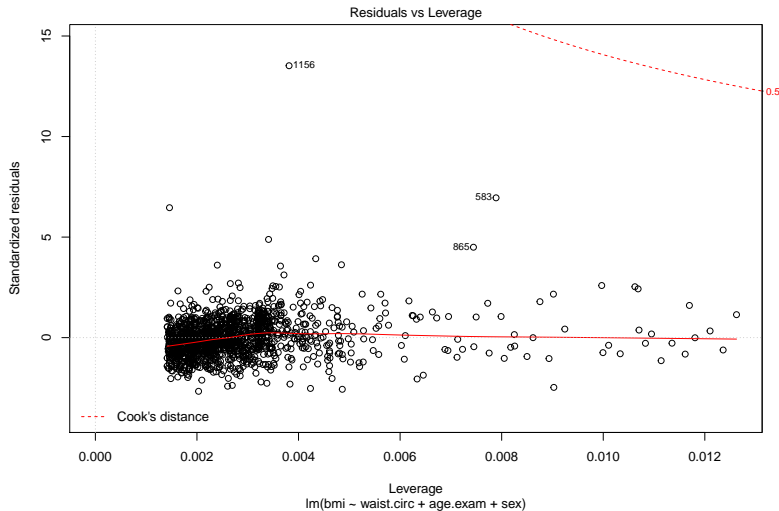
Scale-Location Plot `plot(m2, which = 3)`



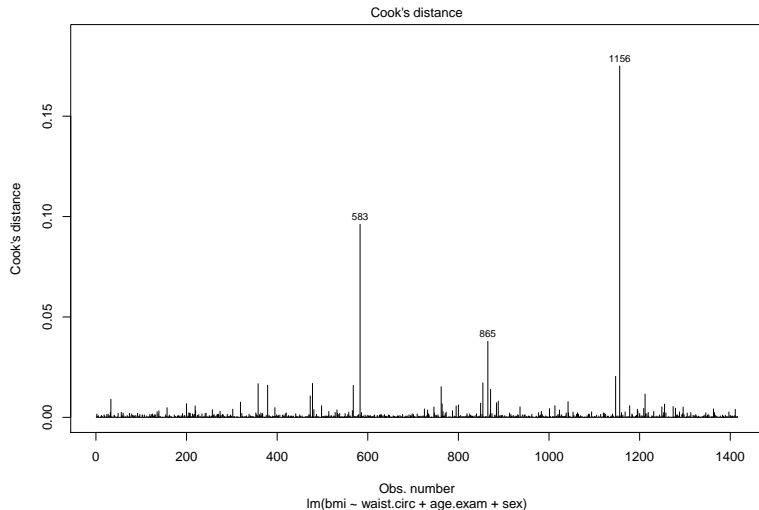
Standardized Residuals plot(m2, which = 2)



Residuals, Leverage, Influence plot(m2, which = 5)



Cook's Distance Index Plot `plot(m2, which = 4)`



Collinearity in Model m2

```
vif(m2)
```

waist.circ	age.exam	sex
2.006275	2.006268	1.000007

- If two predictors (A and B) are highly correlated (collinear) with one another, then the predictive value of the second one into the model (B) will be masked by its strong correlation with A if A is already in the model.
- If we see a variance inflation factor above 5 (certainly above 10) this will indicate that we have some highly correlated predictors and we might be better off including one or the other in our final model.
- In this case, while `waist.circ` and `age.exam` have some correlation with each other ($r = 0.7$ from our scatterplot matrix), it's not enough to worry us on this score.

Conclusions from m2?

Model m2 includes three inputs to predict bmi:

```
bmi = - 1.45  
      + 0.35 waist.circ  
      - 0.29 age.exam  
      + 0.19 if sex = Male
```

- $R^2 = 0.8635$ (adjusted $R^2 = 0.8633$)
- global F test is highly significant ($p < 0.0001$)
- some issues with our residual plots
- no sign of important collinearity

**Should We Transform our Outcome? And, if so,
how?**

Transforming / Re-expressing our Outcome?

We can use the Box-Cox family of transformations to isolate specific choices of the power transformation parameter λ for re-expressing our quantitative outcome which might lead to a more effective (yet still interpretable) model.

This approach is appropriate for strictly **positive** outcomes. If our minimum value is -14, we might add 15 to each observation before using Box-Cox.

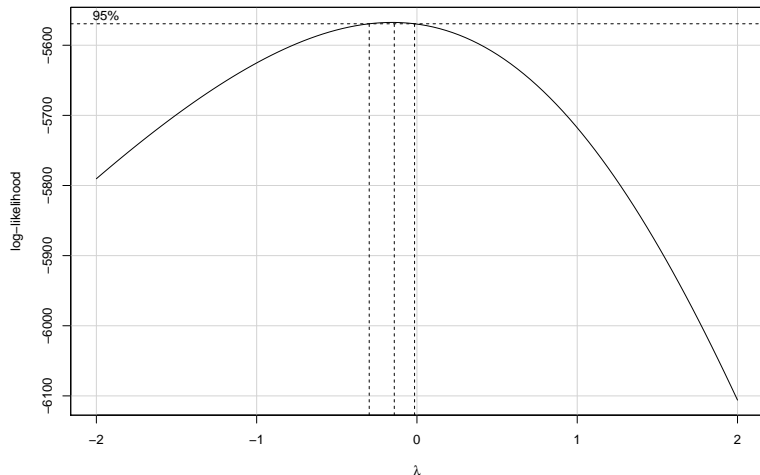
Ladder of Power Transformations

Power (λ)	Transformation
2	y^2
1	y (untransformed)
0.5	\sqrt{y}
0	$\log y$
-1	$\frac{1}{y}$

From the car package, we use `boxCox` and `powerTransform`.

The Box-Cox Transformation Plot for our model `m2`

```
boxCox(lm(bmi ~ waist.circ + age.exam + sex, data = nyfs1))
```



Power Transformation details for our model `m2`

```
powerTransform(lm(bmi ~ waist.circ + age.exam + sex,  
                  data = nyfs1))
```

Estimated transformation parameters

Y1

-0.1543853

What re-expression of the `bmi` data does the Box-Cox plot suggest?

Model m3: $\log(\text{bmi})$ as a function of waist circumference, age at exam and sex

Model m3

```
m2 <- nyfs1 %>% lm(bmi ~ waist.circ + age.exam + sex)
m3 <- nyfs1 %>% lm(log(bmi) ~ waist.circ + age.exam + sex)
```

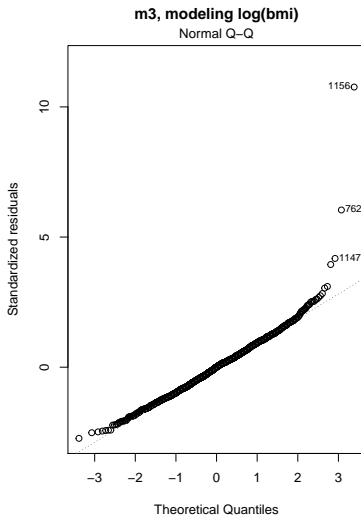
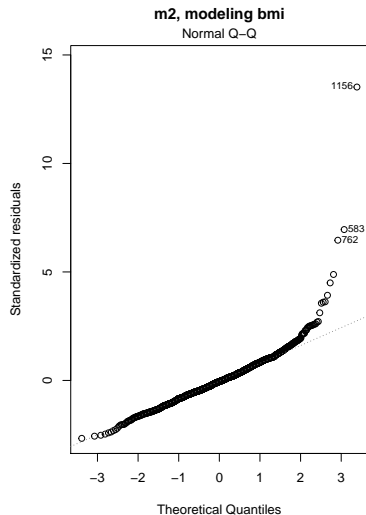
```
glance(m2)
```

	r.squared	adj.r.squared	sigma	statistic	p.value	
1	0.8635437	0.8632538	1.509103	2978.545	0	
	df	logLik	AIC	BIC	deviance	df.residual
1	4	-2589.92	5189.84	5216.118	3215.678	1412

```
glance(m3)
```

	r.squared	adj.r.squared	sigma	statistic	p.value	
1	0.8624454	0.8621531	0.07384071	2951.005	0	
	df	logLik	AIC	BIC	deviance	df.residual
1	4	1682.663	-3355.325	-3329.047	7.698859	1412

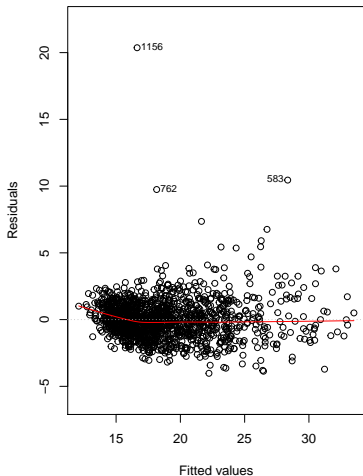
Does the Normality problem improve with $m3$?



More of a “fuzzy football” with $\log(\text{bmi})$: m3?

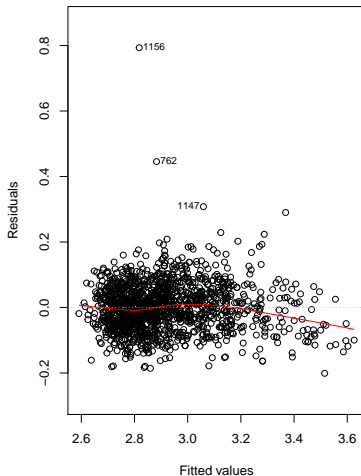
m2, modeling bmi

Residuals vs Fitted

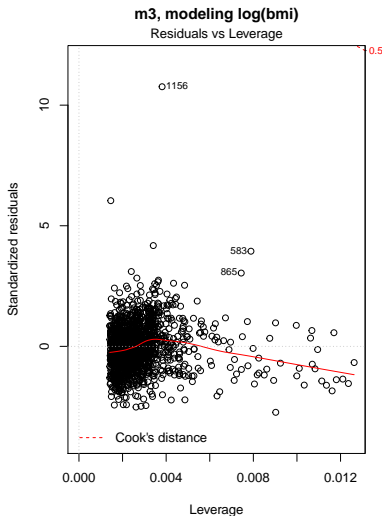
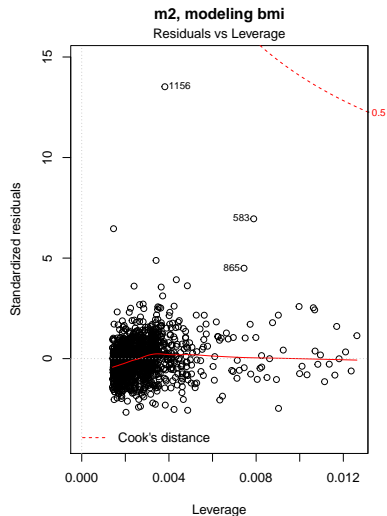


m3, modeling $\log(\text{bmi})$

Residuals vs Fitted



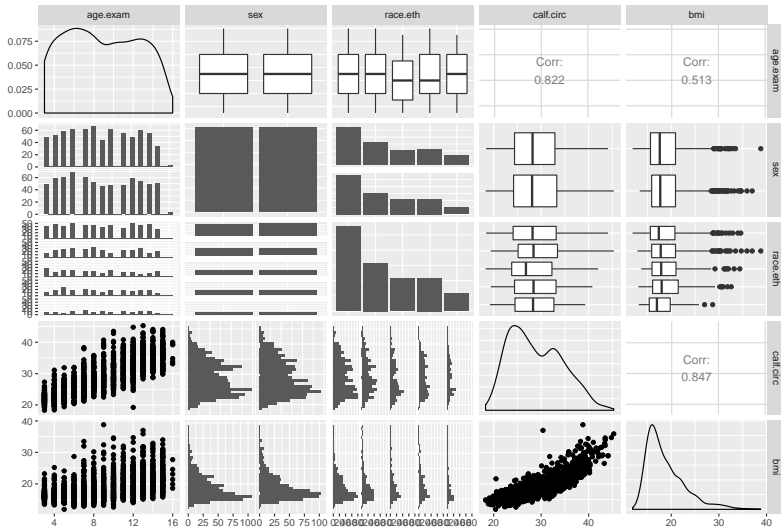
Residuals/Leverage/Influence comparing models



**Model `m4` for the expanded `nyfs2` data set:
Predicting `bmi` using `calf.circ`, `age.exam`, `sex`
and `race.eth`**

(Part of) the expanded nyfs2 data

Scatterplot Matrix for nyfs2 data



The Kitchen Sink Model (m4)

```
m4 <- lm(bmi ~ calf.circ + age.exam + race.eth + sex,  
         data = nyfs2)
```

```
glance(m4)
```

	r.squared	adj.r.squared	sigma	statistic	p.value	
1	0.8300404	0.8291954	1.686596	982.3316	0	
	df	logLik	AIC	BIC	deviance	df.residual
1	8	-2745.366	5508.732	5556.033	4005.205	1408

Summary of m4 (output lightly edited)

bmi ~ calf.circ + age.exam + race.eth + sex, data = nyfs2

Multiple R-squared: 0.83, Adjusted R-squared: 0.8292

F-statistic: 982.3 on 7 and 1408 DF, p-value: < 2.2e-16

Coefficients:	Estimate	SE	t	p	
(Intercept)	-4.005	0.291	-13.77	< 2e-16	***
calf.circ	0.972	0.014	67.83	< 2e-16	***
age.exam	-0.625	0.021	-29.10	< 2e-16	***
race.eth2 Non-Hispanic Black	-0.026	0.119	-0.22	0.825	
race.eth3 Mexican American	0.997	0.136	7.31	4.4e-13	***
race.eth4 Other Hispanic	0.375	0.135	2.78	0.006	**
race.eth5 Other or Multi-Race	-0.130	0.172	-0.76	0.450	
sexMale	0.021	0.090	0.24	0.812	

Residuals:	Min	1Q	Med	3Q	Max	SE
	-4.539	-0.994	-0.127	0.796	19.595	1.687

Can We/Should We Simplify the Model?

- ① t tests for individual predictors as “last predictor in”
- ② F tests for groups of predictors (order matters)
- ③ Stepwise Variable Selection with step using AIC

t test for sex looks NOT significant

Parsimony is an attractive feature in a model. We often want to reduce the number of regression inputs to our model, particularly if some of those inputs add no significant predictive value given the other regression inputs.

- Would the model be better without sex?
 - t test p value is 0.812
 - t test uses “Last Predictor In” approach
- Would the model be better without race.eth, too?
 - Here we have four t tests, some sig, some NS
 - What we need is a test of whether race.eth as an input is significant or not

anova(m4) Results (lightly edited)

Analysis of Variance Table

Response: bmi

	Df	Sum Sq	Mean Sq	F value	Pr(>F)	
calf.circ	1	16917.2	16917.2	5947.1300	< 2.2e-16	***
age.exam	1	2451.2	2451.2	861.7043	< 2.2e-16	***
race.eth	4	191.8	48.0	16.8576	1.653e-13	***
sex	1	0.2	0.2	0.0566	0.812	
Residuals	1408	4005.2	2.8			

ANOVA tests whether each regression input adds significant value given that the preceding inputs are already in the model. So, order matters here.

ANOVA for m4 built in a different order

```
anova(lm(bmi ~ calf.circ + sex + age.exam + race.eth,  
         data = nyfs2))
```

Analysis of Variance Table

Response: bmi

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
calf.circ	1	16917.2	16917.2	5947.1300	< 2.2e-16 ***
sex	1	2.8	2.8	0.9997	0.3176
age.exam	1	2448.7	2448.7	860.8399	< 2.2e-16 ***
race.eth	4	191.6	47.9	16.8379	1.715e-13 ***
Residuals	1408	4005.2	2.8		

Signif. codes:

0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

What Did We Discuss Today?

- Use `eda.ksam` instead of `eda.2sam`
- Box-Cox to identify potential outcome transformations
- Adjusted R^2 and when that indicates severe overfitting
- Testing/Pruning a Kitchen Sink Model
 - Hypothesis testing approaches, t and F tests
- Identification of meaningful collinearity with `vif`
- Checking Assumptions with Residual Plots
- Summarizing the Model: Drawing Conclusions