431 Class 10

Thomas E. Love

2017-09-28

Today's Agenda

- Forming Project Task B Groups
- The Western Collaborative Group Study
 - Dealing with Factors
 - Building Tables Effectively
 - Dealing with Missingness
 - Scatterplot and Correlation Matrices

15 Questions Dr. Love plans to include in the Survey

The following items will be included in the survey. As a result, you will not want to ask these questions in your Task B, although you should consider these groupings as candidates for application in your research questions.

These 8 items will be provided in groups after the application of cutpoints we will identify together after the survey is complete.

- 1. In what year were you born?
- 2. How would you rate your current health overall (Excellent, Very Good, Good, Fair, Poor)
- 3. For how long, in months, have you lived in Northeast Ohio?
- 4. What is your height in inches? (If you are five feet, eight inches tall, please write 68 inches. To convert from centimeters to inches, multiply your height in centimeters by 0.3937, and then round the result to the nearest inch.)
- 5. What is your weight in pounds? (To convert from kilograms to pounds, multiply your weight in kilograms by 2.2046, and then round the result to the nearest pound.)
- What is your pulse rate, in beats per minute? (Please either use a tracking device, or count your pulse for 15 seconds then
 multiply by 4)
- 7. Last week, on how many days did you exercise? (0 7)
- 8. Last night, how many hours of sleep did you get?

The following 7 items will have yes/no responses, and thus produce binary groups for analysis.

- 1. Were you born in the United States?
- 2. Is English the language you speak better than any other?
- 3. Do you identify as female?
- 4. Do you wear prescription glasses or contact lenses?
- 5. Before taking 431, had you ever used R before?
- 6. Are you currently married or in a stable domestic relationship?
- 7. Have you smoked 100 cigarettes or more in your entire life?

Project Task B groups

We need ten such groups, each with about 5 people involved.

Google Form is available at https://goo.gl/forms/WaQOdCEAW0wxdjJh2 and needs to be done by 5 PM today.

Task B meetings in class will be held next Tuesday, and also on 2017-10-12.

Details on Task B specified at https://github.com/thomaselove/431project

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The Form's Questions...

Fall 2017 Project Task B Groups

This is the form to specify the group name and membership for Task B. Only one person from your group should fill out this form. The Task B groups will be formed in class on 2017-09-26. This form needs to be submitted by noon on 2017-09-27. If you have questions, contact Dr. Love directly. Your email address (tel3@case.edu) will be recorded when you submit this form. Not you? Switch account * Required What is the name of your Task B group? * 100 characters or less, please. Please select the names of the members of your group from the list below. Your group must include 4-6 people, in total. Be sure to check the box for each group member, including yourself. In Our Group Albar, Zainab Asagba, Oghenerukema

Today's R Setup

```
library(Epi); library(viridis)
library(GGally); library(mice)
library(forcats); library(tidyverse)
source("Love-boost.R")
```

Cleaning up Loose Ends in the VHL Study

VHL <- read.csv("vonHippel-Lindau.csv") %>% tbl_df

Von Hippel - Lindau study Codebook

- p.ne = plasma norepinephrine (pg/ml)
- tumorvol = tumor volume (ml)
- ullet disease =1 for patients with multiple endocrine neoplasia type 2
- disease = 0 for patients with von Hippel-Lindau disease

We want to add a new variable (factor) called diagnosis, which takes the values von H-L or neoplasia.

Creating a Factor to represent disease diagnosis

We want to add a new variable, specifically a factor, called diagnosis, which will take the values von H-L or neoplasia.

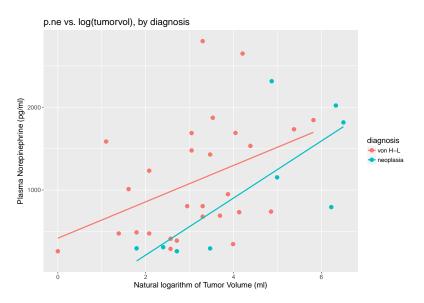
- Recall disease is a numeric 1/0 variable (0 = von H-L, 1 = neoplasia)
- Use fct_recode from the forcats package...

Now, what does VHL look like?

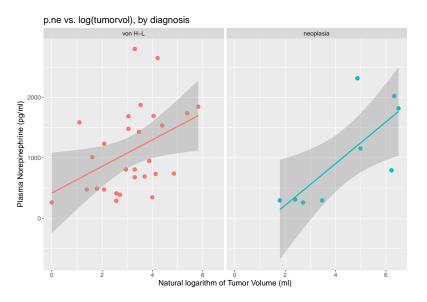
VHL

```
A tibble: 37 \times 5
      id disease p.ne tumorvol diagnosis
   <int>
            <int> <int>
                            <int>
                                      <fctr>
     101
                0
                    289
                                13
                                     von H-L
     102
                1
                    294
                               32 neoplasia
 3
     103
                0
                  2799
                               27
                                     von H-L
4
     104
                   2649
                               67
                0
                                     von H-L
 5
     105
                0
                    346
                                54
                                    von H-L
6
     106
                0
                   1690
                               57
                                     von H-L
     107
                0
                    805
                                19
                                     von H-L
8
     108
                1
                   1153
                               147 neoplasia
9
     109
                    678
                0
                                27
                                     von H-I.
10
     110
                   1817
                              665 neoplasia
  ... with 27 more rows
```

Compare the patients by diagnosis



Facetted Scatterplots by diagnosis



```
model2 <- lm(p.ne ~ log(tumorvol) * diagnosis, data = VHL)
model2</pre>
```

```
Call:
lm(formula = p.ne ~ log(tumorvol) * diagnosis, data = VHL)
Coefficients:
                      (Intercept)
                            417.2
                   log(tumorvol)
                            220.0
              diagnosisneoplasia
                           -893.3
log(tumorvol):diagnosisneoplasia
                            124.8
```

```
 \texttt{p.ne} = 417 + 220 \, \log(\texttt{tumorvol}) - 893 \, (\texttt{diagnosis} = \texttt{neoplasia}) + 125 \, (\texttt{diagnosis} = \texttt{neoplasia}) * \log(\texttt{tumorvol})
```

where the indicator variable (diagnosis = neoplasia) = 1 for neoplasia subjects, and 0 for other subjects...

- Model for p.ne in von H-L patients:
 - 417 + 220 log(tumorvol)
- Model for p.ne in neoplasia patients:
 - $(417 893) + (220 + 125) \log(tumorvol)$
 - -476 + 345 log(tumorvol)

What is the predicted p.ne for a single new subject with tumorvol = 55 ml (so log(tumorvol) = 4.01) in each diagnosis category?

```
fit lwr upr
1 905.7322 -456.1596 2267.624
```

```
fit lwr upr
1 1299.003 -23.21001 2621.215
```

The Western Collaborative Group Study

Original description: 3524 men aged 39-59 and employed in the San Francisco Bay or Los Angeles areas were enrolled in 1960 and 1961. In addition to determinations of behavior pattern, the initial examination included medical and parental history, socioeconomic factors, exercise, diet, smoking, alcohol consumption, diet, serum lipid and lipoprotein studies, blood coagulation studies, and cardiovascular examination.

 http://www.epi.umn.edu/cvdepi/study-synopsis/ western-collaborative-group-study/

The WCGS data describe 3,154 subjects and 22 variables. For now, let's examine a few interesting variables, and a sample of 500 of the observations.

```
wcgs.full <- read.csv("wcgs.csv")
wcgs.full <- tbl_df(wcgs.full)</pre>
```

Select the variables of interest, and sample 500 subjects

```
# A tibble: 500 x 10
         age chol arcus dibpat bmi wghtcat
  <int> <int> <int> <int> <fctr> <dbl> <fctr>
   6039 42 218
                      0 Type A 25.05680 140-170
  3713 44 207
                      0 Type B 22.87093 170-200
3 3465 50 249
                      0 Type A 24.38980 140-170
4 3923 42 236
                      0 Type B 25.10714 170-200
5 3503 46 247
                      1 Type A 21.47629 140-170
6 13167 40 206
                      0 Type B 26.45372 170-200
7 11325 39 267
                      1 Type B 23.56510 140-170
  3800 51 214
                      0 Type A 23.67245 140-170
9 12582 43 235
                      1 Type A 19.25676 140-170
10 19096 48 234
                      0 Type B 28.12608 170-200
# ... with 490 more rows, and 3 more variables:
#
   smoke <fctr>, ncigs <int>, chd69 <fctr>
```

Name	Stored As	Туре	Details (units, levels, etc.)	
id	integer	(nominal)	ID $\#$, nominal and uninteresting	
age	integer	quantitative	age, in years - no decimal places	
chol	integer	quantitative	total cholesterol, mg/dL	
arcus	integer	(nominal)	arcus senilis present (1) or absent (0)	
dibpat	factor (2)	(binary)	behavioral pattern: A or B	
bmi	number	quantitative	body-mass index	
wghtcat	factor (4)	(ordinal)	wt: < 140 , 140-170, 170-200, > 200	
smoke	factor (2)	(binary)	cigarette smoker: Yes or No	
ncigs	integer	quantitative	number of cigarettes smoked per day	
chd69	factor (2)	(binary)	CHD event: Yes or No	

Summary of this sample (without id)

```
cho1
    age
                                arcus
Min.
      :39.00
              Min.
                    :110.0
                            Min.
                                   :0.0000
1st Qu.:42.00
              1st Qu.:201.0
                            1st Qu.:0.0000
Median :45.00
              Median :224.0
                            Median :0.0000
      :46.19 Mean :228.2 Mean
                                  :0.2966
Mean
3rd Qu.:50.00 3rd Qu.:255.0 3rd Qu.:1.0000
Max.
      : 59.00
              Max. :400.0
                            Max. :1.0000
              NA's :3
                            NA's :1
  dibpat
                bmi
                            wahtcat smoke
Type A:251 Min.
                 :17.22 < 140 : 34 No :255
           1st Qu.:22.87 > 200 : 44 Yes:245
Type B:249
           Median: 24.39 140-170:252
           Mean :24.49 170-200:170
           3rd Ou.:25.84
           Max. :38.95
           chd69
   ncigs
Min.
      : 0.00 No :462
1st Qu.: 0.00 Yes: 38
Median: 0.00
      :11.62
Mean
3rd Ou.: 20.00
Max. :70.00
```

A Key Research Question

Were the men with Type A behavioral patterns more likely to suffer a CHD event?

```
table(wcgs1$dibpat, wcgs1$chd69)
```

```
No Yes
Type A 222 29
Type B 240 9
```

What's not so great about this table?

Re-specify the factor to re-order of the levels

```
wcgs1$chdevent <-
  factor(wcgs1$chd69, levels = c("Yes", "No"))
tab1 <- table(wcgs1$dibpat, wcgs1$chdevent)
tab1</pre>
```

```
Yes No
Type A 29 222
Type B 9 240
```

and this is standard epidemiological format.

Aha! A Two-by-Two Table!

twoby2(tab1) ## twoby2 is part of the Epi package

```
2 by 2 table analysis:
```

Outcome : Yes

Comparing : Type A vs. Type B

```
Yes No P(Yes) 95% conf. interval
Type A 29 222 0.1155 0.0815 0.1613
Type B 9 240 0.0361 0.0189 0.0680
```

```
95% conf. interval
Relative Risk: 3.1965 1.5450 6.6134
Sample Odds Ratio: 3.4835 1.6132 7.5221
Conditional MLE Odds Ratio: 3.4754 1.5589 8.5398
Probability difference: 0.0794 0.0334 0.1279
```

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A 4 by 2 table

```
table(wcgs1$wghtcat, wcgs1$chdevent)
```

Why is this a poor table?

An Improved 4 x 2 table

	Yes	No
< 140	1	33
140-170	20	232
170-200	14	156
> 200	3	41

```
table(wcgs1$dibpat, wcgs1$wghtcat, wcgs1$chdevent)
  = Yes
       < 140 140-170 170-200 > 200
 Type A
                 15 13
 Type B
   = No
       < 140 140-170 170-200 > 200
 Type A 15 104 82 21
```

Type B 18

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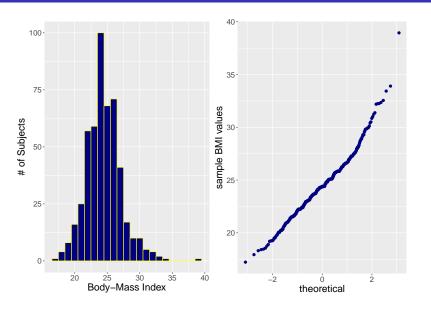
128 74

A Three-Way Table (Flattened)

```
ftable(wcgs1$dibpat, wcgs1$wghtcat, wcgs1$chdevent)
```

```
Type A < 140 0 15
140-170 15 104
170-200 13 82
> 200 1 21
Type B < 140 1 18
140-170 5 128
170-200 1 74
> 200 2 20
```

How can we describe the distribution of BMI?



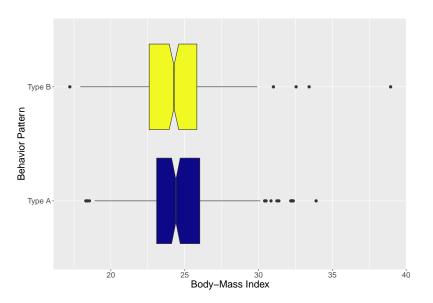
```
mosaic::favstats(wcgs1$bmi)

min Q1 median Q3 max mean
17.22242 22.87091 24.3898 25.84016 38.94737 24.48819
sd n missing
2.680243 500 0

psych::describe(wcgs1$bmi)
```

```
vars n mean sd median trimmed mad min max
X1     1 500 24.49 2.68 24.39 24.39 2.15 17.22 38.95
    range skew kurtosis se
X1 21.72 0.65     2.11 0.12
```

Comparing BMI by Behavior Pattern

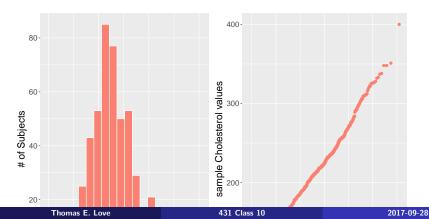


```
by(wcgs1$bmi, wcgs1$dibpat, mosaic::favstats)
wcgs1$dibpat: Type A
     min Q1 median Q3
                                      max
                                             mean
 18.30729 23.09703 24.40514 26.02431 33.90239 24.64689
      sd n missing
2.700341 251
wcgs1$dibpat: Type B
           Q1 median
                               Q3
     min
                                      max
                                             mean
 17.22242 22.59412 24.27378 25.82449 38.94737 24.32823
     sd n missing
2.65565 249
```

How about Total Cholesterol, instead?

Warning: Removed 3 rows containing non-finite values (stat_bin).

Warning: Removed 3 rows containing non-finite values (stat_qq).



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```
mosaic::favstats(wcgs1$chol)

min Q1 median Q3 max mean sd n missing
110 201 224 255 400 228.167 44.20081 497 3

psych::describe(wcgs1$chol)
```

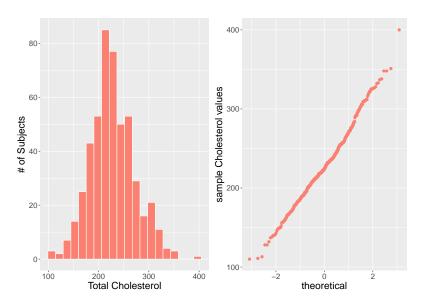
```
vars n mean sd median trimmed mad min max
X1   1 497 228.17 44.2   224 226.64 41.51 110 400
   range skew kurtosis se
X1   290 0.33   0.36 1.98
```

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What should we do about missing values?

- ggplot2 doesn't include missing values in the plot, but it does warn that they've been removed.
 - We could suppress that warning by setting na.rm = TRUE in the call to a geom like geom_histogram or geom_qq.

Plot with na.rm = TRUE



Classification of Missing Data

There are three classifications we will think about in 431. Subtleties abound, but these three will suffice for most practical work.

- MCAR: Missing Completely at Random. This is the desirable scenario
 for us. MCAR means that there is no relationship between the
 probability of a data point being missing, and any values in the data
 set, either missing or observed.
 - The missing data are just a random subset of the data.
 - This is one kind of "ignorable" missingness.

Classification of Missing Data

- MAR: Missing at Random, which is definitely an unfortunate name. Less desirable, but there is still some hope. MAR means that the probability of a data point being missing has nothing to do with the missing value that would have been observed, but does have something to do with the values of some other variable that you did observe.
 - The idea is that if we can control for this other variable in our analysis, then we can treat this missingness as just a random subset of the data after that adjustment, which will eventually be pretty straightforward.
 - This is another kind of "ignorable" missingness.

Classification of Missing Data

- MNAR: Missing not at Random is a more serious issue. Now, additional thought and some special methods may well be required. Here, there is a relationship between the probability that a value is missing and what the actual (missing) value is.
 - This is what we mean by "non-ignorable" missingness.
 - Multiple Imputation methods (and Maximum Likelihood approaches)
 assume the data are at least MAR, so that's the important distinction to
 make, generally.
- Some of these explanations come from this link

What should we do about missing values?

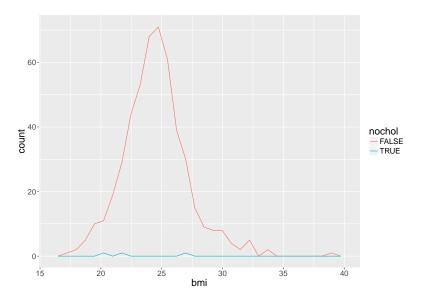
- At times you will want to try to understand what makes observations with missing values different from observations with meaningful recorded values, especially if we're thinking that the missing mechanism is MCAR or MAR.
 - We might, for instance, compare the BMI values or perhaps the smoking status for those with and without missing cholesterol values, using the is.na() function to make a new variable to indicate those subjects without a cholesterol level.
 - Some of this material is drawn from R for Data Science

Do those missing cholesterol look unusual in terms of BMI?

First, we'll build a new (logical) variable (TRUE/FALSE) to indicate a missing cholesterol level, and then we'll plot the BMI distributions for each level of the new variable.

```
wcgs1 %>%
  mutate(
    nochol = is.na(chol)
) %>%
  ggplot(aes(x = bmi)) +
  geom_freqpoly(aes(col = nochol), bins = 30) +
  theme(text = element_text(size = 18))
```

Do the BMIs of people without chol look different?



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Are the people without a cholesterol value unusual in terms of their smoking status?

```
temp1 <- table(is.na(wcgs1$chol), wcgs1$smoke)
knitr::kable(temp1)</pre>
```

	No	Yes
FALSE	253	244
TRUE	2	1

What should we do about missing values?

- 4 How many missing values are there?
 - If the missing values are more than, say, 5% of a variable, we're going to need some strong, almost heroic assumptions in order to feel confident about using such a variable in building a model or making an inference.
 - If the amount of missing data is very small relative to the size of the data as a whole, then leaving out a few samples and just running models or comparisons ignoring those observations may not be too damaging.
 - Depending on the situation, you may want to look for other fixes besides just dropping these cases and wiping out potentially useful data.
 - Some of this material comes from this R-bloggers post

What should we do about missing values?

Could we **impute** missing values?

- One approach is simple imputation, where a single value is created to "fill in" the missing observation. This is pretty easy to do, but very rarely a good idea.
 - Rarely, substituting the mean is a reasonable thing to do, as it reduces variance in your estimate of the distribution, among other problems.
 - Sometimes, but still pretty rarely, substituting in a random value observed in the rest of the data set is a reasonable thing to do.
 - Better, although still problematic, imputation approaches use other variables in the data set to predict the missing value, and contain a random component. Using other variables preserves the relationships among variables in the imputations. The random component is important so that all missing values of a single variable are not all exactly equal. One example would be to use a regression equation to predict missing values, then add a random error term.
- See http://www.theanalysisfactor.com/multiple-imputation-in-a-nutshell/

What's so bad about simple imputation?

Although there are several simple imputation approaches that solve many of the problems inherent in mean imputation, one problem remains. Because the imputed value is an estimate - a predicted value - there is uncertainty about its true value. Every statistic has uncertainty, measured by its standard error. Statistics computed using imputed data have even more uncertainty than its standard error measures. Your statistical package cannot distinguish between an imputed value and a real value.

Since the standard errors of statistics based on imputed values, such as sample means or regression coefficients, are too small, corresponding reported p-values are also too small. P-values that are reported as smaller than they, in reality, are, lead to Type I errors.

- It turns out that *multiple* imputation is a much better approach.
- Various types of "hot deck" procedures can help, too. See the HotDeckImputation package in R, or this link

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So what is multiple imputation?

Multiple imputation has solved this problem by incorporating the uncertainty inherent in imputation. It has four steps:

- Create *m* sets of imputations for the missing values using an imputation process with a random component.
- ② The result is m full data sets. Each data set will have slightly different values for the imputed data because of the random component.
- Analyze each completed data set. Each set of parameter estimates will differ slightly because the data differs slightly.
- Ombine results, calculating the variation in parameter estimates.

Multiple Imputation is amazing

Remarkably, m, the number of sufficient imputations, can be only 5 to 10 imputations, although it depends on the percentage of data that are missing. The result is unbiased parameter estimates and a full sample size, when done well.

Doing multiple imputation well, however, is not always quick or easy. First, it requires that the missing data be ignorable. Second, it requires a very good imputation model. Creating a good imputation model requires knowing your data very well and having variables that will predict missing values.

Source:

http://www.theanalysis factor.com/multiple-imputation-in-a-nutshell/

What will we do in 431?

- Often, we'll be willing to simply exclude the data with missing values from our graphs or other analyses.
- Sometimes, we'll be willing to assume (heroically) that the data are
 missing at random and we'll use a simple imputation approach, via the
 mice package.
- Later in the term (and definitely in 432) we'll move on up to multiple imputation, using mice sometimes and Hmisc at other times.

Using mice to build imputations for chol

```
mice::md.pattern(wcgs1)
    id age dibpat bmi wghtcat smoke ncigs chd69
496
    chdevent arcus chol
496
  3
```

Build 5 actual imputations using the "predictive mean matching" (pmm) approach

```
wcgs.temp <- mice(wcgs1,m=5,maxit=50,meth='pmm',seed=431)</pre>
```

```
iter imp variable
 1
       chol
             arcus
    2 chol arcus
 1
    3
      chol arcus
      chol arcus
    4
     5
      chol
             arcus
2
       chol
             arcus
2
       chol
             arcus
2
    3
       chol
             arcus
2
    4
       chol
             arcus
2
     5
       chol
             arcus
3
        chol
             arcus
```

View imputation results, summarized

```
> summary(wcqs.temp)
Multiply imputed data set
call:
mice(data = wcgs1. m = 5. method = "pmm". maxit = 50. seed = 431)
Number of multiple imputations: 5
Missing cells per column:
       id
               age
                        cho1
                                arcus
                                         dibpat
                                                      bmi
                                                           wahtcat
                                                                       smoke
                                                                                 ncias
                                                                                          chd69 chdevent
       0
                 0
Imputation methods:
                       cho1
                                         dibpat
                                                           wghtcat
                                                                       smoke
                                                                                 ncigs
                                                                                           chd69 chdevent
      id
               age
                                arcus
                                                      bmi
   "pmm"
             "pmm"
                       "mmq"
                                 'mmm"
                                           'mmm"
                                                    "mmq"
                                                              "mmq"
                                                                       "mmd"
                                                                                 "pmm"
                                                                                           "pmm"
VisitSequence:
 chol arcus
PredictorMatrix:
             age chol arcus dibpat bmi wahtcat smoke ncigs chd69 chdevent
id
                    0
                                                            0
age
                           0
                                       0
                                                                   0
                                                                             0
cho1
               1
                                       1
                                                                   1
arcus
                           0
                                                            1
dibpat
hmi
                                       0
                                                                   0
wahtcat
smoke
                                       0
                                                      0
                                                            0
                                                                   0
                                                            0
ncigs
                           0
                                       0
                                                      0
                                                                   0
chd69
                                       0
                                                      0
                                                            0
                                                                   0
chdevent
                                                                   0
Random generator
                 seed value:
```

"mmd

Inspect the imputed values, if you like

```
wcgs.temp$imp$chol
```

```
1 2 3 4 5
62 137 268 211 269 224
237 212 244 295 258 282
472 255 205 252 205 178
```

wcgs.temp\$imp\$arcus

```
1 2 3 4 5
480 0 1 0 0 1
```

Simple Imputation: Complete data with, let's say, the fourth of the five imputations we built

```
completedwcgs1 <- mice::complete(wcgs.temp,4)</pre>
```

```
favstats with and without imputation
```

```
mosaic::favstats(wcgs1$chol)
```

```
min Q1 median Q3 max mean sd n missing 110 201 224 255 400 228.167 44.20081 497 3
```

```
mosaic::favstats(completedwcgs1$chol)
```

```
min Q1 median Q3 max mean sd n missing 110 201 224 255 400 228.262 44.13794 500 0
```

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```
> modelFit0 <- with(wcqs1,lm(chol ~ bmi * dibpat))</pre>
> summary(modelFit0)
call:
lm(formula = chol ~ bmi * dibpat)
Residuals:
             1Q Median
    Min
                             30
                                      Max
-117.432 -28.537 -4.019 25.859 175.580
Coefficients:
               Estimate Std. Error t value Pr(>|t|)
(Intercept) 185.096 25.576 7.237 1.77e-12 ***
bmi
                1.974 1.031 1.914 0.0562 .
dibpatType B 70.345 36.237 1.941 0.0528 .
bmi:dibpatTvpe B -3.325 1.471 -2.261 0.0242 *
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' '1
Residual standard error: 43.74 on 493 degrees of freedom
  (3 observations deleted due to missingness)
Multiple R-squared: 0.02664. Adjusted R-squared: 0.02072
F-statistic: 4.498 on 3 and 493 DF. p-value: 0.003984
> confint(modelFit0)
                      2.5 % 97.5 %
(Intercept) 134.84532681 235.3476621
bmi
               -0.05226551 4.0000876
dibpatType B -0.85177283 141.5423274
bmi:dibpatTvpe B -6.21478658 -0.4358619
```

Multiple imputation and pooling

Suppose that the next step in our analysis is to fit a linear model to the data. You may ask what imputed data set to choose. The mice package makes it again very easy to fit a a model to each of the imputed data sets and then pool the results together

```
> modelFit1 <- with(wcgs.temp.lm(chol ~ bmi * dibpat))</pre>
> round(summary(pool(modelFit1)),3)
                      se
                                   df Pr(>|t|)
                                                lo 95 hi 95 nmis
                                                                    fmi lambda
(Intercept) 185.409 25.776 7.193 434.927
                                         0.000 134.749 236.069
                                                               NA 0.034
                                                                         0.029
bmi
     1.958 1.039 1.884 438.572
                                         0.060 -0.084
                                                        4.000
                                                                0 0.032
                                                                         0.028
dibpat2 71.446 36.278 1.969 479.010
                                         0.049 0.163 142.730
                                                               NA 0.016 0.012
bmi:dibpat2 -3.362 1.472 -2.283 480.131
                                         0.023 -6.255 -0.469
                                                               NA 0.016 0.012
```

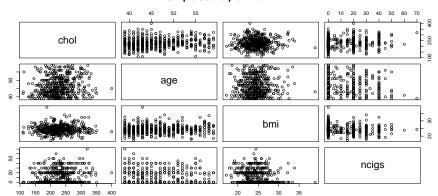
Details of linear model after pooling

modelFit1 contains the results of the fitting performed over the imputed data sets, while the pool() function pools them all together.

- fmi = fraction of missing information
- lambda = proportion of total variance attributable to the missing data
- Note that if we were looking at a strict alpha of 0.05, we'd have a significant dibpat2 main effect now, when we didn't before.

Multivariable Descriptions: A Scatterplot Matrix

Simple Scatterplot Matrix



Correlation Matrix (after imputation)

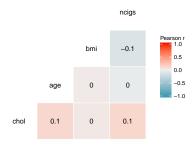
```
completedwcgs1 <- mice::complete(wcgs.temp,1)
round(cor(completedwcgs1[c("chol", "age", "bmi")]),3)</pre>
```

```
chol age bmi
chol 1.000 0.133 0.036
age 0.133 1.000 0.029
bmi 0.036 0.029 1.000
```

Using GGally for a Correlation Matrix

```
tempdat <- completedwcgs1 %>%
  select(chol, age, bmi, ncigs)

ggcorr(tempdat, name = "Pearson r", label = TRUE)
```



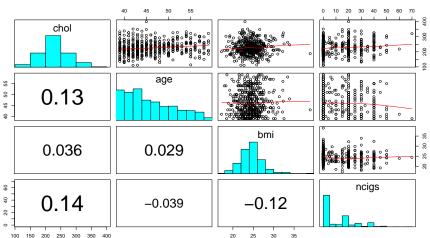
My Favorite Scatterplot Matrix

My favorite way to augment this plot adds loess smooths to the upper panel, and correlations in the lower panel, with histograms down the diagonal. To do this, we first create two functions (these modifications come from Chang's R Graphics Cookbook), called panel.hist and panel.cor.

These functions are in the Love-boost.R script.

Augmented Scatterplot Matrix

Augmented Scatterplot Matrix



Code for Augmented Scatterplot Matrix

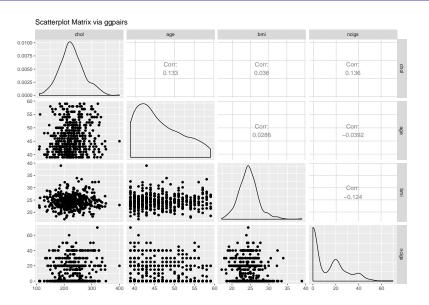
Using GGally for a Scatterplot Matrix (Code)

```
tempdat <- completedwcgs1 %>%
  select(chol, age, bmi, ncigs)

ggpairs(tempdat, title = "Scatterplot Matrix via ggpairs")
```

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Using GGally for a Scatterplot Matrix (Result)



For Tuesday: Working with Tables

For Class 11, I'd like you to read a 1981 paper by A.S.C. Ehrenberg that you'll find linked on the Class 10 README.

The paper is called *The Problem of Numeracy* and it provides some very helpful tips for working with tables, in particular.

There are three key tips related to the development of tables, in practice, as described by Ehrenberg, and also by Howard Wainer¹ who concisely states them as:

- ① Order the rows and columns in a way that makes sense.
- 2 Round a lot!
- ALL is different and important.

¹Visual Revelations (1997), Chapter 10.