431 Class 25

Thomas E. Love

2017-11-30

Today's Agenda

- Ginzberg's Depression Data from the car package
 - Should we be modeling a transformed outcome?
 - Comparing Models with R², adjusted R², AIC, BIC
 - Comparing Model Predictions Out of Sample
 - Partitioning the Data Set
 - Building the Model in a Training Sample
 - Using a Test Sample, MAPE and MSPE for Validation
- Getting Better Calibrated on Residual Plots

Today's R Setup and Data Set

```
library(car); library(magrittr)
library(broom); library(tidyverse)

ginz0 <- tbl_df(car::Ginzberg)
ginz0$id <- 1:82
ginz <- select(ginz0, id, fatalism, simplicity, depression)

source("Love-boost.R")</pre>
```

Ginzberg's Depression Data

The Ginzberg data are part of the car package. The data describe psychiatric patients hospitalized for depression. We'll look at three variables, each of which is scaled to have mean 1 and standard deviation 0.5 in this sample. . .

- our outcome, fatalism, which measures the subject's fatalism (the belief that all events are inevitable)
- simplicity, which measures the need to see the world in black and white
- depression, which is the Beck self-report depression scale

Subjects with values exceeding 1 on these measures are reporting greater than average fatalism, simplicity or depression, respectively.

Standardized key variables in the ginz tibble

Remember that the values for each variable have been standardized to mean 1 and standard deviation 0.5

```
summary(select(ginz, -id))
```

```
fatalism
                  simplicity
                                  depression
Min. :-0.05837
                Min. :0.2507
                                Min. :0.4695
1st Qu.: 0.56301
                 1st Qu.:0.6563
                                1st Qu.:0.5664
Median: 0.97727
                 Median :0.8827
                                Median : 0.8247
                 Mean :1.0000
Mean : 1.00000
                                Mean :1.0000
3rd Qu.: 1.39152
                 3rd Qu.:1.2694
                                3rd Qu.:1.3737
Max. : 2.22003
                 Max. :2.8541
                                Max. :2.2456
```

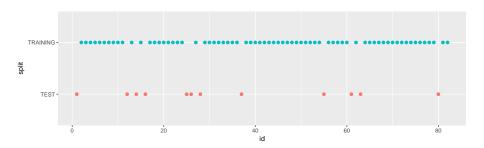
- What does a value of zero mean on these scales?
- A change of one unit on these scales is how large?

Partitioning into Training and Test Samples

We'll build a training sample (ginz.train) for building models with 70 patients, and hold back a test sample (ginz.test) of the remaining 12 patients for evaluating the model after it's been built.

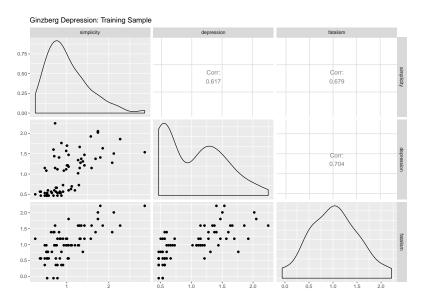
```
set.seed(43111)
ginz.train <- sample_n(ginz, 70, replace = FALSE)
ginz.test <- anti_join(ginz, ginz.train, by = "id")</pre>
```

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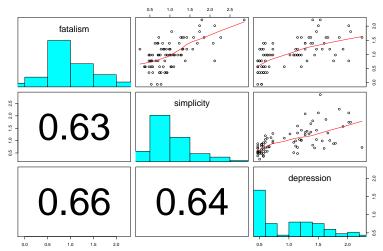


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Scatterplot Matrix for Ginzberg's Depression Data







Does Box-Cox suggest a transformation?

This throws an error message:

```
Error in bc1(out, lambda) :
   First argument must be strictly positive.
```

Oops, we have some non-positive values of our outcome

```
summary(ginz.train$fatalism)
```

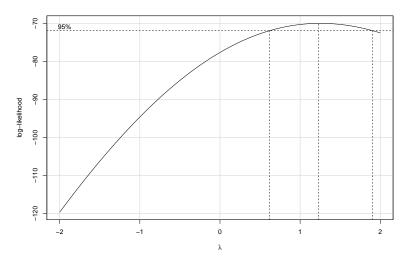
```
Min. 1st Qu. Median Mean 3rd Qu. Max. -0.05837 0.56301 0.97727 1.03349 1.39152 2.22003
```

Could just add 1 to every value for Box-Cox check...

```
ginz.train$fat <- ginz.train$fatalism + 1
m1a <- lm(fat ~ simplicity + depression, data = ginz.train)</pre>
```

```
boxCox(m1a)
```

Here's the new plot (on 1 + fatalism)



And, if we need backup for our eyes...

```
powerTransform(m1a)
```

```
Estimated transformation parameters
Y1
1.251195
```

Take advantage of the roundlam object contained within powerTransform.

```
powerTransform(m1a)$roundlam
```

```
Y1
1
```

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```
m1 <- lm(fatalism ~ simplicity + depression,
        data = ginz.train)
arm::display(m1)
lm(formula = fatalism ~ simplicity + depression, data = ginz.
           coef.est coef.se
(Intercept) 0.14 0.10
simplicity 0.40 0.10
depression 0.48 0.10
n = 70. k = 3
residual sd = 0.33, R-Squared = 0.59
```

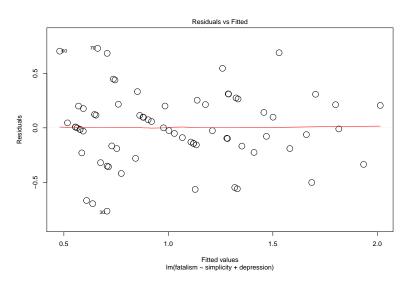
summary(m1) # edited output on next page

```
lm(fatalism ~ simplicity + depression, data = ginz.train)
Multiple R-squared: 0.593, Adjusted R-squared: 0.580
F-statistic: 48.71 on 2 and 67 DF, p-value: 8.699e-14
                        SE t
Coefficients: Estimate
(Intercept)
               0.140 0.099 1.42 0.161
simplicity
              0.400 0.100 3.98 0.0002 ***
             0.477 0.103 4.64 1.64e-05 ***
depression
Residuals: Min
                Q1 Med
                            Q3 Max
                                        SE
        -0.76 -0.17 -0.005 0.20 0.73 0.33
```

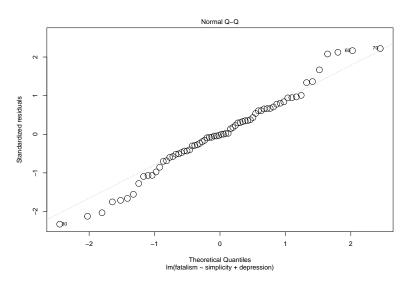
Is collinearity a big issue here?

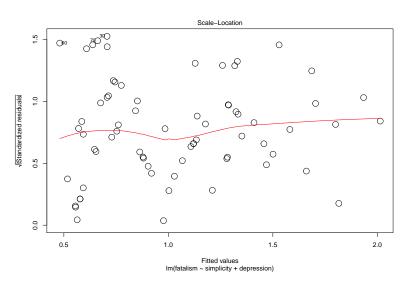
```
vif(m1)
```

```
simplicity depression 1.616002 1.616002
```

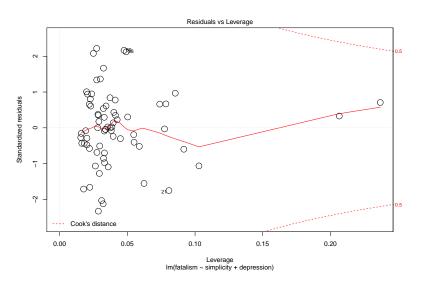


Residuals in a Normal Q-Q Plot

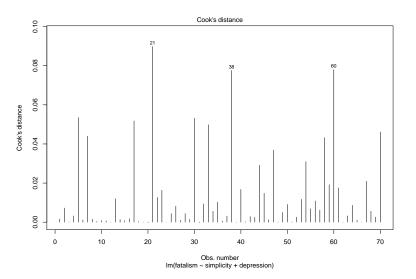




Plot 5: Residuals, Leverage, Influence?



Plot 4: Index plot of Cook's distance



- Model m1 included both depression and simplicity.
- Let's fit Model m2 which only includes depression.

```
m2 <- lm(fatalism ~ depression, data = ginz.train)
arm::display(m2)</pre>
```

```
Call:
lm(formula = fatalism ~ depression, data = ginz.train)
Multiple R-squared: 0.496, Adjusted R-squared: 0.489
F-statistic: 66.92 on 1 and 68 DF, p-value: 1.031e-11
Coefficients: Estimate SE t
                                     р
(Intercept) 0.289 0.101 2.86 0.006 **
depression 0.730 0.089 8.18 1.03e-11 ***
Residuals: Min Q1 Med Q3 Max
                                     SE
        -0.80 -0.19 -0.06 0.20 0.90 0.37
```

```
anova(m1, m2)
```

Analysis of Variance Table

```
Model 1: fatalism ~ simplicity + depression

Model 2: fatalism ~ depression

Res.Df RSS Df Sum of Sq F Pr(>F)

1 67 7.4448

2 68 9.2081 -1 -1.7632 15.868 0.0001699 ***
---

Signif. codes:
0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```

Does the order in which we list the models matter?

```
anova(m2, m1)
Analysis of Variance Table
Model 1: fatalism ~ depression
Model 2: fatalism ~ simplicity + depression
 Res.Df RSS Df Sum of Sq F Pr(>F)
1 68 9.2081
2 67 7.4448 1 1.7632 15.868 0.0001699 ***
Signif. codes:
0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

How does order matter here?

Which Model Looks Best in the training sample?

```
round(glance(m1),3) # depression and simplicity
```

```
round(glance(m2),3) # depression alone
```

```
r.squared adj.r.squared sigma statistic p.value df
1 0.496 0.489 0.368 66.916 0 2
logLik AIC BIC deviance df.residual
1 -28.331 62.662 69.408 9.208 68
```

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Making Predictions in the Test Sample with m1

Let's use model m1 to predict fatalism scores for our test sample group.

```
fatalism = 0.14 + 0.40 simplicity + 0.48 depression
```

```
head(ginz.test,1)
```

A tibble: 1 x 4

So, predicted fatalism for subject 1 here is...

$$fatalism = 0.14 + 0.40 (0.92983) + 0.48 (0.5987), or 0.80$$

Observed error is 0.36 - 0.80 = -0.44

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There must be an easier way

And there is...

```
predict(m1, newdata = ginz.test)
```

```
5
0.7977475 0.7377445 0.6013353 0.8820320 0.6698679
        6
                                                 10
0.6455903 0.6082102 0.5705040 1.2421200 1.0197707
       11
                 12
1.3530064 1.9511248
```

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Making Predictions in the Test Sample with m1

Let's use our model m1 to predict fatalism scores for the test sample group of 12 patients on the basis of their simplicity and depression scores.

```
m1.preds <- predict(m1, newdata = ginz.test)</pre>
# make predictions
m1.error <- ginz.test$fatalism - m1.preds
# calculate errors
m1.abserror <- abs(m1.error)
# absolute value of errors
m1.sqerror <- m1.error^2</pre>
# squared errors
```

Back to the first member of our Test Sample

```
head(ginz.test, 1)
# A tibble: 1 x 4
    id fatalism simplicity depression
 <int> <dbl> <dbl> <dbl>
 1 0.35589 0.92983 0.5987
m1.preds[1] # predicted fatalism from m1
0.7977475
m1.error[1] # error (observed - predicted)
```

-0.4418575

Making Predictions in the Test Sample with m2

Using model m2, we have:

```
m2.preds <- predict(m2, newdata = ginz.test) # predictions
m2.error <- ginz.test$fatalism - m2.preds # errors
m2.abserror <- abs(m2.error) # absolute value of errors
m2.sqerror <- m2.error^2 # squared errors</pre>
```

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Mean Absolute Prediction Error (MAPE) across the Models

```
summary(m1.abserror)
  Min. 1st Qu. Median Mean 3rd Qu. Max.
0.1245 0.2391 0.3758 0.3963 0.4456 1.1810
summary(m2.abserror)
```

Min. 1st Qu. Median Mean 3rd Qu. Max. 0.1146 0.2131 0.3350 0.3811 0.4420 0.9936

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Mean Squared Prediction Error (MSPE) across the Model

```
Min. 1st Qu. Median Mean 3rd Qu. Max.
0.01551 0.05771 0.14125 0.22686 0.19859 1.39473
summary(m2.sqerror)
```

Min. 1st Qu. Median Mean 3rd Qu. Max. 0.01313 0.04601 0.11351 0.20242 0.20093 0.98721

Which Model Looks Best in the test sample?

Model	MAPE	MSPE	Max Abs Err
m1 (depression + simplicity)	0.396	0.227	1.18
m2 (depression only)	0.381	0.202	0.99

What we see here in the 12(!) people in our test group doesn't entirely match what we saw in the training sample of 70 people.

- But should it?
- In 432, we'll learn some better ways to validate our models.

Calibrating Yourself on Residual Plots

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Multivariate Regression: Checking Assumptions

Assumptions (see Course Notes, Section 42)

- Linearity
- Normality
- Homoscedasticity
- Independence

Available Residual Plots

```
plot(model, which = c(1:3,5))
```

- Residuals vs. Fitted Values
- Normal Q-Q Plot of Standardized Residuals
- Scale-Location Plot
- Index Plot of Cook's Distance
- 5 Residuals, Leverage and Influence

```
set.seed(431122)
x1 \leftarrow rnorm(200, 20, 5)
x2 \leftarrow rnorm(200, 20, 12)
x3 \leftarrow rnorm(200, 20, 10)
er \leftarrow rnorm(200, 0, 1)
v < -.3*x1 - .2*x2 + .4*x3 + er
sim0 <- data.frame(y, x1, x2, x3) %>% tbl_df
mod0 < -lm(y ~ x1 + x2 + x3, data = sim0)
summary (mod0) # appears on next slide
```

An Idealized Model (by Simulation)

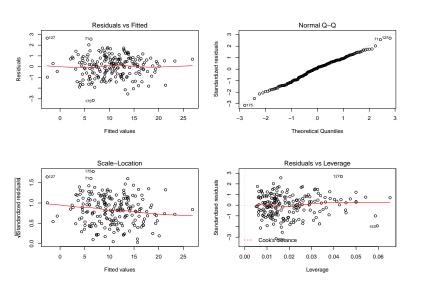
```
Call: lm(formula = y \sim x1 + x2 + x3, data = sim0)
Residuals: Min 1Q Median 3Q
                                         Max
        -3.14553 -0.68079 0.08096 0.69216 2.65265
Coefficients: Estimate Std. Error t value Pr(>|t|)
(Intercept) 0.122852 0.348584 0.352 0.725
x1 0.285539 0.014211 20.093 <2e-16 ***
x2 -0.204908 0.005828 -35.159 <2e-16 ***
x3 0.413308 0.007172 57.631 <2e-16 ***
Signif codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```

Residual standard error: 1.007 on 196 degrees of freedom Multiple R-squared: 0.9589, Adjusted R-squared: 0.9583 F-statistic: 1524 on 3 and 196 DF, p-value: < 2.2e-16

Building Residual Plots for Idealized Model

```
par(mfrow=c(2,2))
plot(mod0)
par(mfrow=c(1,1))
```

Residual Analysis (Idealized Model: n = 200)



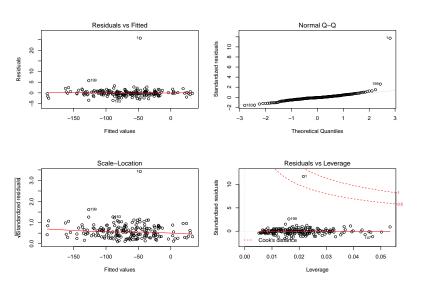
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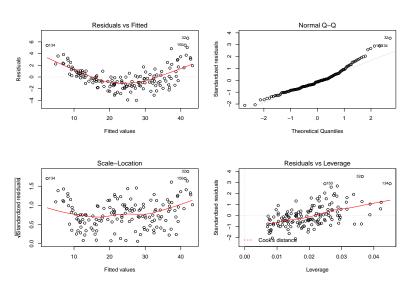
What's the Goal Here?

Develop an effective model. (?) (!)

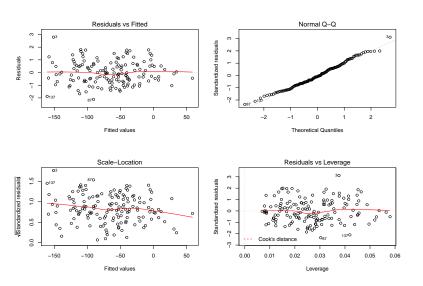
- Models can do many different things. What you're using the model for matters, a lot.
- Don't fall into the trap of making binary decisions (this model isn't perfect, no matter what you do, and so your assessment of residuals will also have shades of gray).
- The tools we have provided (scatterplots, mostly) are well designed for rather modest sample sizes. When you have truly large samples, they don't scale very well.
- Just because R chooses four plots for you to study doesn't mean they provide the only relevant information.
- Embrace the uncertainty. Look at it as an opportunity to study your data more effectively.

Simulation 1 (n = 200 subjects)

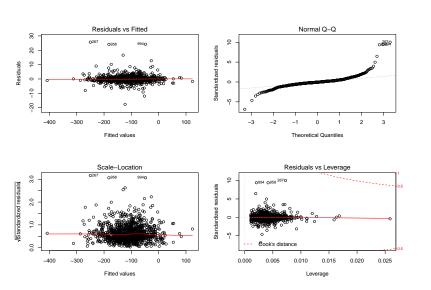


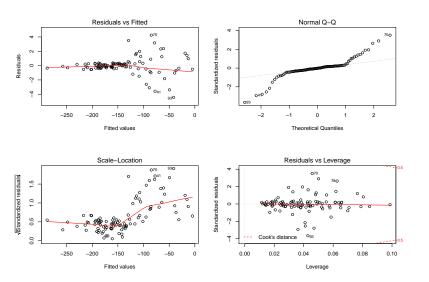


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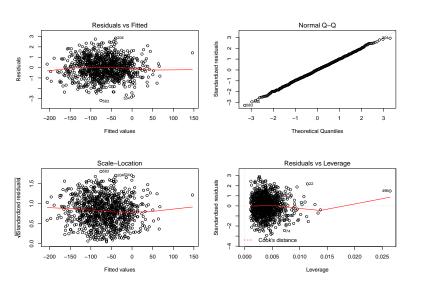


Simulation 4 (n = 1000)





Simulation 6 (n = 1000)



Some Reactions to the Six Simulations

For those of you playing along at home...

- Observation 1 has an impossibly large standardized residual, and some influence.
- Ourve in residuals vs. fitted values plot suggests potential non-linearity.
- No substantial problems, although there's a little bit of heteroscedasticity.
- Normality issues outlier-prone even with 1000 observations.
- Serious heteroscedasticity residuals much more varied for larger fitted values.
- No serious violations point 496 has very substantial leverage, though.