

ESS101- Modeling and Simulation

Lecture 10-12

Paolo Falcone

*Department of Signals and Systems
Chalmers University of Technology
Göteborg, Sweden*

Today (Chapter 9)

- ➡ Parametric identification
 - ➡ Tailor-made models
 - ➡ Ready-made models
 - ➡ Model structures (ARMAX, ARX, OE)
- ➡ Prediction
- ➡ Prediction Error Method
 - ➡ Least squares method
 - ➡ Convergence analysis
 - ➡ Excitation

Parametric system identification

Problem statement. Given a system in the form

$$\dot{x} = f(x, u, \theta)$$

$$y = g(x, u, \theta)$$

find the vector θ based on measurements of \underline{u} and y .

Linear parametric identification. The function f is linear. Possible forms

$$\dot{x} = A(\theta)x + B(\theta)u$$



$$y = C(\theta)x + D(\theta)u$$

or

$$Y(s) = G(s, \theta)U(s)$$

Parametric system identification. Models types

Tailor-made models

-  obtained as the results of physical modeling
-  the parameters have a physical meaning

$$\dot{x} = f(x, u, \theta) \quad \dot{x} = A(\theta)x + B(\theta)u$$

$$y = g(x, u, \theta) \quad y = C(\theta)x + D(\theta)u$$

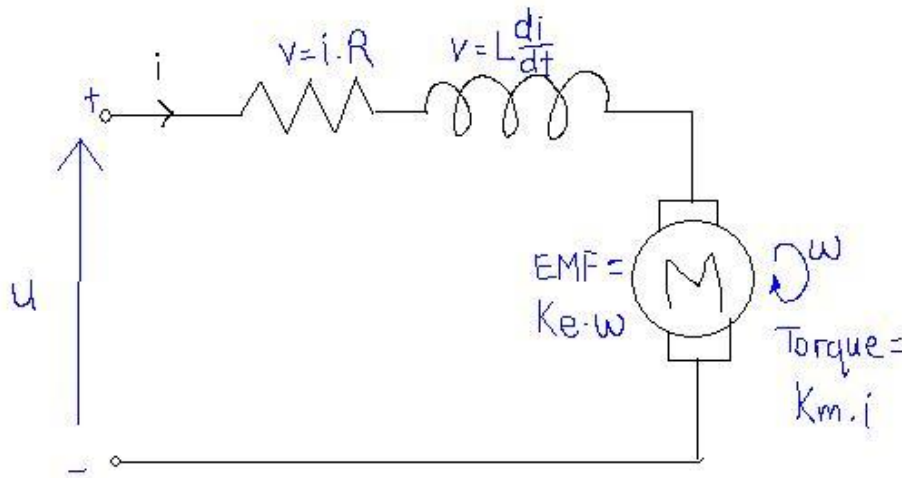
$$Y(s) = G(s, \theta)U(s)$$

Ready-made models (or “black-box” models)

-  describe the input-output relationships
-  no physical interpretation

$$Y(z) = G(z, \theta)U(z)$$

Tailor-made models. Example



$$\frac{di(t)}{dt} = -\frac{R}{L}i(t) - \frac{k}{L}\omega(t) + \frac{1}{L}V(t)$$

$$\frac{d\omega(t)}{dt} = \frac{k}{J}i(t) - \frac{b}{J}\omega(t)$$

$$\dot{x} = A(\theta)x + B(\theta)u$$

$$y = Cx$$

$$x = \begin{bmatrix} i \\ \omega \end{bmatrix}, \quad u = V$$

$$y = \omega$$

$$\theta = [R \quad L \quad k \quad b \quad J]^T$$

$$A(\theta) = \begin{bmatrix} -\frac{\theta_1}{\theta_2} & -\frac{\theta_3}{\theta_2} \\ \frac{\theta_3}{\theta_5} & -\frac{\theta_4}{\theta_5} \end{bmatrix} \quad B(\theta) = \begin{bmatrix} \frac{1}{\theta_2} \\ 0 \end{bmatrix}$$

Linear ready-made models

Consider the *parametrized, linear discrete time* model

$$y(t) = G(q, \theta)u(t) + w(t)$$

where

⊗ $G(q, \theta)$ is a parametrized transfer function

$$G(q, \theta) = \frac{B(q)}{F(q)} = \frac{b_1 q^{-nk} + b_2 q^{-nk-1} + \dots + b_{n_b} q^{-nk-n_b+1}}{1 + f_1 q^{-1} + \dots + f_{n_f} q^{-n_f}}$$

$$\theta = \begin{bmatrix} b_1 & \dots & b_{n_b} & f_1 & \dots & f_{n_f} & c_1 & \dots & c_{n_c} & d_1 & \dots & d_{n_d} \end{bmatrix}^T$$

⊗ w is filtered white noise

$$w(t) = H(q, \theta)e(t) \quad H(q, \theta) = \frac{C(q)}{D(q)} = \frac{1 + c_1 q^{-1} + \dots + c_{n_c} q^{-n_c}}{1 + d_1 q^{-1} + \dots + d_{n_d} q^{-n_d}}$$

Model structures

The most general model structure is

$$y(t) = \frac{B(q,\theta)}{F(q,\theta)} u(t) + \frac{C(q,\theta)}{D(q,\theta)} e(t)$$

❖ **A**uto **R**egressive **M**oving **A**verage with e**X**ogenous input (**ARMAX**)

$$F(q,\theta) = D(q,\theta) = A(q,\theta), \quad \underbrace{A(q,\theta)y(t)}_{\text{Auto Regression}} = \underbrace{B(q,\theta)u(t)}_{\text{eXogenous input}} + \underbrace{C(q,\theta)e(t)}_{\text{Moving Average}}$$

❖ **A**uto **R**egressive with e**X**ogenous input (**ARX**)

$$F(q,\theta) = D(q,\theta) = A(q,\theta), \quad \underbrace{A(q,\theta)y(t)}_{\text{Auto Regression}} = \underbrace{B(q,\theta)u(t)}_{\text{eXogenous input}} + e(t)$$

$$C(q,\theta) = 1$$

Model structures

The most general model structure is

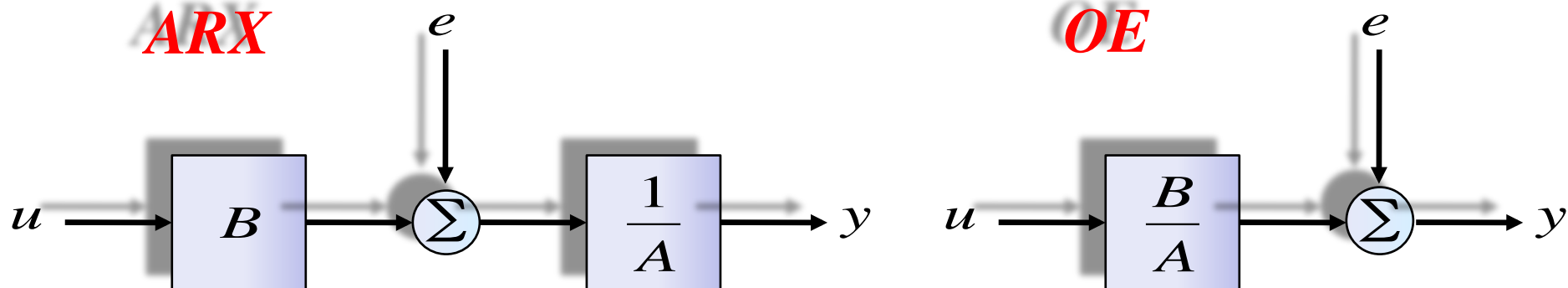
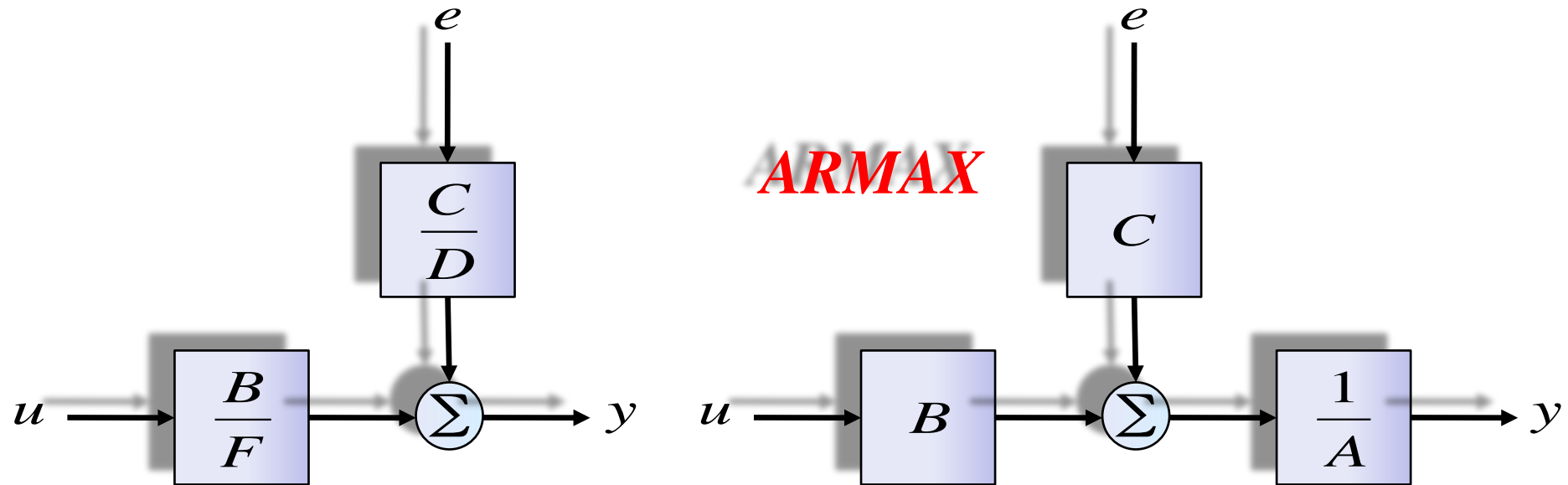
$$y(t) = \frac{B(q,\theta)}{F(q,\theta)} u(t) + \frac{C(q,\theta)}{D(q,\theta)} e(t)$$

➤ **Output Error (OE)**

$$\frac{C(q,\theta)}{D(q,\theta)} = 1, \quad y(t) = \frac{B(q,\theta)}{A(q,\theta)} u(t) + e(t)$$

Observe that in the *ARX* models the noise passes through the model dynamics while in *OE* model it does not

Model structures



Notation

ARMAX(n_a, n_b, n_c)

$$\underbrace{A(q, \theta)}_{n_a \text{ \# coeff}} y(t) = \underbrace{B(q, \theta)}_{n_b \text{ \# coeff}} u(t) + \underbrace{C(q, \theta)}_{n_c \text{ \# coeff}} e(t)$$

ARX(n_a, n_b)

$$\underbrace{A(q, \theta)}_{n_a \text{ \# coeff}} y(t) = \underbrace{B(q, \theta)}_{n_b \text{ \# coeff}} u(t) + e(t)$$

Prediction

Definition. The *predictor* is a mathematical object we use to calculate $y(t)$, based on measurements $y(s)$ and $u(s)$ with $s < t$.

Obtain the predictor by *removing the noise*. Example the OE model

OE model

$$y(t) = \frac{B(q, \theta)}{A(q, \theta)} u(t) + e(t)$$

OE predictor

$$\hat{y}(t | \theta) = \frac{B(q, \theta)}{A(q, \theta)} u(t)$$

Prediction

In general

$$y(t) = \underbrace{\frac{B(q,\theta)}{F(q,\theta)}}_{G(q,\theta)} u(t) + \underbrace{\frac{C(q,\theta)}{D(q,\theta)}}_{H(q,\theta)} e(t)$$



$$H^{-1}(q,\theta)y(t) = H^{-1}(q,\theta)G(q,\theta)u(t) + e(t)$$



$$y(t) + H^{-1}(q,\theta)y(t) = y(t) + H^{-1}(q,\theta)G(q,\theta)u(t) + e(t)$$



$$y(t) = [1 - H^{-1}(q,\theta)]y(t) + H^{-1}(q,\theta)G(q,\theta)u(t) + e(t)$$



$$\hat{y}(t | \theta) = [1 - H^{-1}(q,\theta)]y(t) + H^{-1}(q,\theta)G(q,\theta)u(t)$$

Prediction

In general

$$y(t) = \underbrace{\frac{B(q,\theta)}{F(q,\theta)}}_{G(q,\theta)} u(t) + \underbrace{\frac{C(q,\theta)}{D(q,\theta)}}_{H(q,\theta)} e(t)$$



$$\hat{y}(t | \theta) = [1 - H^{-1}(q,\theta)] y(t) + H^{-1}(q,\theta) G(q,\theta) u(t)$$

Observe that

$$1 - H^{-1}(q,\theta) = 1 - \frac{D(q,\theta)}{C(q,\theta)}$$

$$= \frac{(c_1 - d_1)q^{-1} + \cdots + (c_{n_c} - d_{n_c})q^{-n_c}}{1 + c_1 q^{-1} + \cdots + c_{n_c} q^{-n_c}}$$

$\hat{y}(t | \theta)$ only
depends on $y(s)$,
 $s < t$

Prediction

Consider the **ARX** model

$$y(t) = \frac{B(q,\theta)}{A(q,\theta)} u(t) + \frac{1}{A(q,\theta)} e(t)$$

The corresponding predictor is obtained by observing that

$$A(q,\theta)y(t) = B(q,\theta)u(t) + e(t)$$



$$\left(1 + a_1 q^{-1} + \cdots + a_{n_a} q^{-n_a}\right) y(t) = B(q,\theta)u(t) + e(t)$$



$$\hat{y}(t | \theta) = -a_1 y(t-1) - \cdots - a_{n_a} y(t-n_a) + B(q,\theta)u(t)$$

Prediction error method (PEM)

1. Calculate the *predictor* of the system output

$$\hat{y}(t | \theta)$$

2. Calculate the *prediction error*

$$\varepsilon(t | \theta) = y(t) - \hat{y}(t | \theta)$$

3. Calculate the *criterion (loss or cost) function*

$$V_N(\theta) = \frac{1}{N} \sum_{t=1}^N \varepsilon^2(t | \theta)$$

4. Calculate the *“best” vector of parameters*

$$\theta_N^* = \arg \min_{\theta} V_N(\theta)$$

Iterative search for minimum

In general the solution of the minimization problem is found numerically, through iterative methods

Solve the *necessary* optimality condition

$$\frac{d}{d\theta} V_N(\theta) = 0$$

$$\hat{\theta}^{i+1} = \hat{\theta}^i - \mu_i M_i V'_N(\hat{\theta}^i)$$

$$M_i = \left(V''_N(\hat{\theta}^i) \right)^{-1}$$

Assuming V_N
twice
differentiable

PEM for ARX models

Consider the **ARX** model

$$y(t) = \frac{B(q,\theta)}{A(q,\theta)} u(t) + \frac{1}{A(q,\theta)} e(t)$$

and its predictor

$$\hat{y}(t | \theta) = -a_1 y(t-1) - \dots - a_{n_a} y(t-n_a) + B(q,\theta)u(t)$$

It can be rewritten as

$$\hat{y}(t | \theta) = \theta^T \varphi(t), \quad \theta = \begin{bmatrix} a_1 \\ \vdots \\ a_{n_a} \\ b_1 \\ \vdots \\ b_{n_b} \end{bmatrix}, \quad \varphi(t) = \begin{bmatrix} -y(t-1) \\ \vdots \\ -y(t-n_a) \\ u(t) \\ \vdots \\ u(t-n_b) \end{bmatrix},$$

PEM for ARX models

The predictor is linear in the parameters vector

Apply the PEM method. The prediction error is

$$\varepsilon(t | \theta) = y(t) - \theta^T \varphi(t)$$

The criterion function becomes

$$\begin{aligned} V_N(\theta) &= \frac{1}{N} \sum_{t=1}^N \left(y(t) - \theta^T \varphi(t) \right)^2 \\ &= \frac{1}{N} \sum_{t=1}^N y^2(t) - 2\theta^T f_N + \theta^T R_N \theta \end{aligned} \quad \text{with} \quad \begin{cases} f_N = \frac{1}{N} \sum_{t=1}^N \varphi(t) y(t) \\ R_N = \frac{1}{N} \sum_{t=1}^N \varphi(t) \varphi^T(t) \end{cases}$$

$$\theta_N^* = \arg \min_{\theta} V_N(\theta) = R_N^{-1} f_N$$

PEM for ARX models

For **ARX** models the PEM gives

$$V_N(\theta) = \frac{1}{N} \sum_{t=1}^N \left(y(t) - \theta^T \varphi(t) \right)^2$$
$$= \frac{1}{N} \sum_{t=1}^N y^2(t) - 2\theta^T f_N + \theta^T R_N \theta$$

with $\begin{cases} f_N = \frac{1}{N} \sum_{t=1}^N \varphi(t) y(t) \\ R_N = \frac{1}{N} \sum_{t=1}^N \varphi(t) \varphi^T(t) \end{cases}$

$$\theta_N^* = \arg \min_{\theta} V_N(\theta) = R_N^{-1} f_N$$


Observe that in f_N and R_N are estimates of the covariance and cross covariance of y and u

$$R_u^N(\tau) = \frac{1}{N} \sum_{t=1}^N u(t-i)u(t-j)$$
$$R_{yu}^N(\tau) = \frac{1}{N} \sum_{t=1}^N y(t-i)u(t-j)$$

Model properties

Bias error.

 The chosen model structure is not capable of describing the system

 The parameters change when identification is made in different operating conditions

Variance errors.

 Parameters change when identification is repeated

 Use longer measurement sequences

Convergence of the estimate

1. Select the family of models
2. Minimize the criterion function

$$V_N(\theta) = \frac{1}{N} \sum_{t=1}^N \varepsilon(t | \theta)^2$$

θ_N^* is the optimum

Convergence of the estimate

Observations

- θ_N^* is the minimum of $V_N(\theta)$
- $V_N(\theta)$ depends on data (stochastic processes)

Hence

- θ_N^* is a stochastic process

Q: what happens to θ_N^* when $N \rightarrow \infty$?

Convergence of the estimate

Observe that

$\hat{y}(t | \theta)$ is a stationary process, then $\varepsilon(\cdot | \theta) = y(\cdot) - \hat{y}(\cdot | \theta)$
and $\varepsilon(\cdot, \theta)^2$ are a stationary process as well

$V_N(\theta)$ tends to the variance of the prediction error as $N \rightarrow \infty$

$$V_N(\theta) \rightarrow E[\varepsilon(\cdot | \theta)^2] = \hat{V}(\theta)$$

NOTE. $E[\varepsilon(\cdot | \theta)^2]$ depends on θ

Convergence of the estimate

Denote by

$$\Delta = \left\{ \theta \mid \hat{V}(\hat{\theta}) \leq \hat{V}(\theta), \forall \theta \in \Theta \right\}$$

the set of minimum points (in general multiple minima)

Since $V_N(\theta) \rightarrow E[\varepsilon(\cdot \mid \theta)^2] = \hat{V}(\theta)$ we expect the convergence to uniformly hold in the parameters space as well

$$\theta_N^* \rightarrow \Delta$$

Conclusion. The estimated vector of parameters converges to a value *minimizing the variance of the prediction error*

Convergence of the estimate

Assume the data has been generated by the model $M(\theta^0)$

Q: does θ_N^* converge to θ^0 ?

Consider the prediction error $\varepsilon(t | \theta) = y(t) - \hat{y}(t | \theta)$

$$\varepsilon(t | \theta) = e(t) + \left(\hat{y}(t | \theta^0) - \hat{y}(t | \theta) \right)$$

with $e(t) = y(t) - \hat{y}(t | \theta^0)$

Convergence of the estimate

Observe that

$\hat{y}(t | \theta^0) - \hat{y}(t | \theta)$ depends on the past values of $u(\cdot), y(\cdot)$ while $e(t)$ does not.

$\hat{y}(t | \theta^0) - \hat{y}(t | \theta)$ and $e(t)$ are uncorrelated. Hence:

$$\underbrace{\text{Var}[\varepsilon(t | \theta)]}_{\hat{V}_N(\theta)} = \underbrace{\text{Var}[e(t)]}_{\lambda(e(t) \text{ white noise})} + \text{Var}[\hat{y}(t | \theta^0) - \hat{y}(t | \theta)]$$

Conclusion. $\hat{V}_N(\theta)$ (the rhs) is minimized by setting $\theta = \theta^0$

Observe that the prediction error is a white noise

Variance error

Assume the *bias is zero*. I.e.:

$$\theta_N^* \rightarrow \theta^0, \text{ as } N \rightarrow \infty$$

then, it can be shown that

$$E\left[(\theta_N^* - \theta^0)(\theta_N^* - \theta^0)^T\right] \approx \frac{1}{N} \underbrace{\lambda}_{\text{noise variance}} R^{-1}$$

$$R = E\left[\left(\frac{d}{d\theta} \hat{y}(t | \theta^0)\right)\left(\frac{d}{d\theta} \hat{y}(t | \theta^0)\right)^T\right]$$

Conclusion. The variance of the estimated parameters can be decreased by increasing the size of the data set.

Excitation

Consider an **ARX(1,1)** model

$$\boxed{\theta_N^* = R_N^{-1} f_N} \quad \text{with} \quad \begin{cases} f_N = \frac{1}{N} \sum_{t=1}^N \varphi(t) y(t) \\ R_N = \frac{1}{N} \sum_{t=1}^N \varphi(t) \varphi^T(t) \end{cases} \quad \text{Least Squares (LS) formula}$$

Requirement. R_N invertible

$$R_N = \begin{bmatrix} \frac{1}{N} \sum_{t=1}^N y^2(t-1) & \frac{1}{N} \sum_{t=1}^N y(t-1)u(t-1) \\ \frac{1}{N} \sum_{t=1}^N u(t-1)y(t-1) & \frac{1}{N} \sum_{t=1}^N u^2(t-1) \end{bmatrix} = \begin{bmatrix} R_y(0) & R_{yu}(0) \\ R_{uy}(0) & R_u(0) \end{bmatrix}$$

Excitation

In general

$$R_N = \begin{bmatrix} R_y^{n_a-1} & R_{yu} \\ R_{uy} & R_u^{n_b-1} \end{bmatrix} \quad \text{where}$$

$$R_y^{n_a-1} = E \begin{bmatrix} \begin{bmatrix} y(t-1) \\ y(t-2) \\ \vdots \\ y(t-n_a) \end{bmatrix} \\ \begin{bmatrix} y(t-1) & y(t-2) & \cdots & y(t-n_a) \end{bmatrix} \end{bmatrix}$$
$$R_u^{n_b-1} = E \begin{bmatrix} \begin{bmatrix} u(t-1) \\ u(t-2) \\ \vdots \\ u(t-n_b) \end{bmatrix} \\ \begin{bmatrix} u(t-1) & u(t-2) & \cdots & u(t-n_b) \end{bmatrix} \end{bmatrix}$$

Excitation

A *necessary condition** on the regularity of R_N is that

$R_u^{n_b-1}$ *must* be invertible

**deriving from the Sylvester criterion of the positive definiteness of a matrix*

$$R_u^{n_b-1} = \begin{bmatrix} R_u(0) & R_u(1) & \cdots & R_u(n_b-1) \\ R_u(1) & R_u(0) & \cdots & R_u(n_b-2) \\ \vdots & \vdots & \ddots & \vdots \\ R_u(n_b-1) & R_u(n_b-2) & \cdots & R_u(0) \end{bmatrix}$$

In this case u is said to be *persistently exciting*

Excitation

Definition. A signal w is persistently exciting of order n if R_w^n is invertible.

Q: Which signal is persistently exciting of every order?

Q: Which signal would you **NOT** use in the LS formula?

Note that singularity of R_N may occur for other reasons as well (system structure)