

# **A tutorial of Hidden Markov Model (HMM)**

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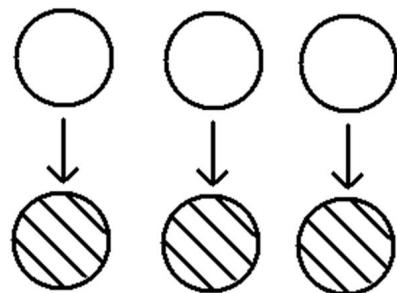
## **Application of HMM**

- Stock price analysis
- Auto speech recognition
- Character recognition
- Sequence alignment
- etc.

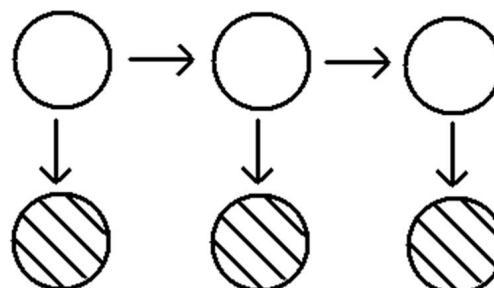
# Probabilistic Graphical Model (PGM)

PGM is a probabilistic model. The graph expresses the conditional dependence structure between random variables.

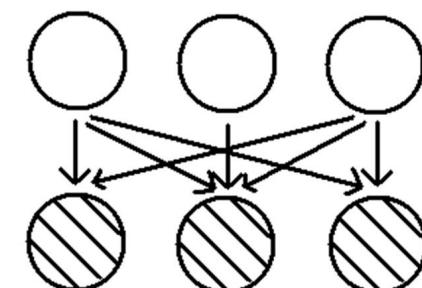
Example:



GMM



HMM



Others

White circle represents latent variable.

Solid circle represents observed variable.

All of GMM, HMM, Others belongs to PGM.

# EM algorithm for PGM

EM algorithm is used to estimate parameters in probabilistic graphic model (PGM) with latent variables\*.

## EM algorithm for PGM with latent variables

1. Init parameters
2. E step:  $q(\mathbf{Z}) = p(\mathbf{Z} | \mathbf{X}\theta^{\text{old}})$
3. M step:  $\hat{\theta} = \arg \max_{\theta} Q(\theta, \theta^{\text{old}})$   
where  $Q(\theta, \theta^{\text{old}}) = \sum_{\mathbf{Z}} p(\mathbf{Z} | \mathbf{X}\theta^{\text{old}}) \ln p(\mathbf{XZ} | \theta)$
4. If converge then stop, otherwise goto 2

\* More details can be seen in GMM Tutorial.

Generally, estimation of  $p(\mathbf{Z}|\mathbf{X}\theta^{\text{old}})$  is difficult.

However, the independence of PGM will simplify the model.

### D-separation property<sup>[1]</sup>

Let A, B, C be set of nodes. We can check whether A and B are conditional independent on C in belowing way.

Check all the paths from node in A to node C. The path is said to be blocked if

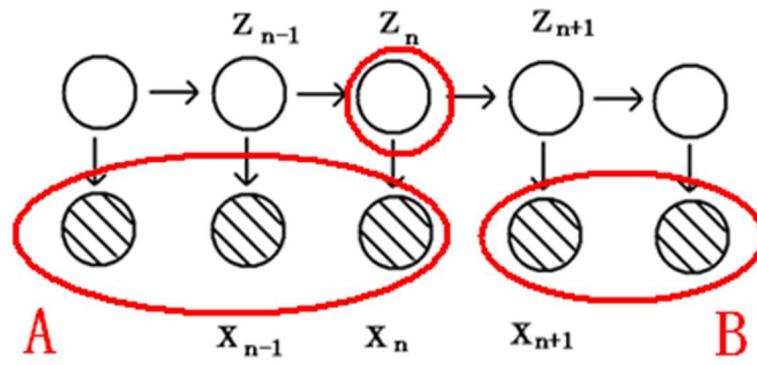
- (a)  $\exists$  head-to-tail or tail-to-tail nodes on the path, and the node is in C.
- (b)  $\forall$  head-to-head nodes, neither the node, nor any of its descendants is in C

If all paths are blocked, then A is said to be d-separated from B by C.

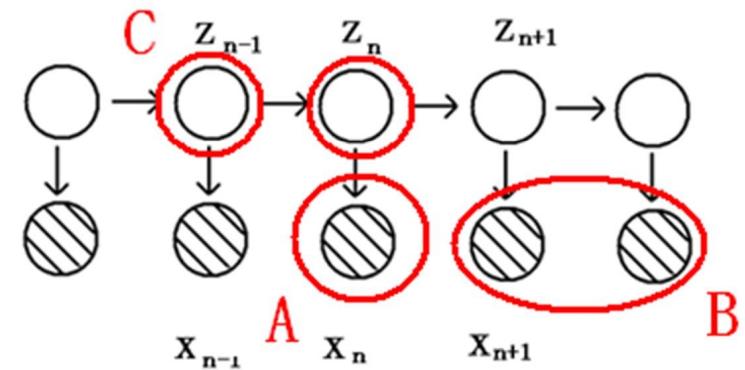
That is  $P(AB|C) = P(A|C)P(B|C)$ .

[1] Bishop, Christopher M. *Pattern recognition and machine learning*. Springer, 2006. Chap. 8.

## Exercises



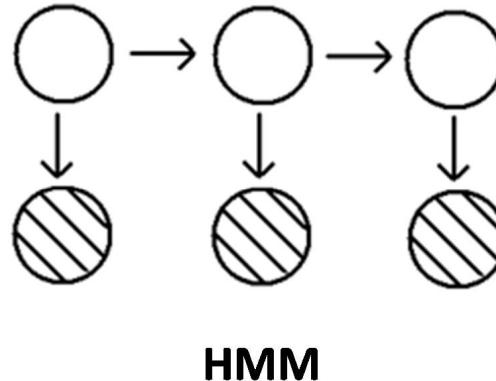
$$p(AB|\mathbf{z}_n) = p(A|\mathbf{z}_n)p(B|\mathbf{z}_n)$$



$$p(ABC|\mathbf{z}_n) = p(A|\mathbf{z}_n)p(B|\mathbf{z}_n)p(C|\mathbf{z}_n)$$

## HMM model

Hidden Markov Model (HMM) is a kind of PGM with latent variables.



HMM's **joint probability distribution** over both latent and observed variables is

$$p(\mathbf{XZ}|\theta) = p(\mathbf{z}_1|\boldsymbol{\pi}) \prod_{n=2}^N p(\mathbf{z}_n | \mathbf{z}_{n-1}, \mathbf{A}) \prod_{n=1}^N p(\mathbf{x}_n | \mathbf{z}_n, \phi) \quad (1)$$

where the parameters are  $\theta = \{\boldsymbol{\pi}, \mathbf{A}, \phi\}$

$\boldsymbol{\pi}$  is the start probability of each state.

$\mathbf{A}$  is the transition matrix.

$\phi$  is the parameter associated with emission distribution (can be multinomial, Gauss, GMM, etc.)

## **Three basic problem of HMM**

- **How to train a HMM model?**
- **Given a sequence X, how to get the likelihood  $p(X)$  from HMM model?**
- **How to find the best decoding path of HMM model?**

# How to train HMM model?

HMM is a kind of probabilistic graphic model (PGM) with latent variables.

Apply EM algorithm to HMM,  $Q(\theta, \theta^{\text{old}})$  is

$$\begin{aligned}
 Q(\theta, \theta^{\text{old}}) &= \sum_{\mathbf{Z}} p(\mathbf{Z} | \mathbf{X}\theta^{\text{old}}) \ln p(\mathbf{XZ} | \theta) \\
 &= \sum_{\mathbf{Z}} p(\mathbf{Z} | \mathbf{X}\theta^{\text{old}}) \left[ \ln p(\mathbf{z}_1) + \sum_{n=2}^N \ln p(\mathbf{z}_n | \mathbf{z}_{n-1}) + \sum_{n=1}^N \ln p(\mathbf{x}_n | \mathbf{z}_n) \right] \\
 &= \sum_{\mathbf{z}_1} p(\mathbf{z}_1 | \mathbf{X}\theta^{\text{old}}) \ln p(\mathbf{z}_1) + \sum_{n=2}^N \sum_{\mathbf{z}_{n-1}, \mathbf{z}_n} p(\mathbf{z}_{n-1}, \mathbf{z}_n | \mathbf{X}\theta^{\text{old}}) \ln p(\mathbf{z}_n | \mathbf{z}_{n-1}) + \sum_{n=1}^N \sum_{\mathbf{z}_n} p(\mathbf{z}_n | \mathbf{X}\theta^{\text{old}}) \ln p(\mathbf{x}_n | \mathbf{z}_n) \\
 &= \underbrace{\sum_{k=1}^K \gamma(z_{nk}) \ln \pi_k}_{\textcircled{1}} + \underbrace{\sum_{n=2}^N \sum_{j=1}^K \sum_{k=1}^K \xi(z_{n-1,j} z_{n,k}) \ln A_{jk}}_{\textcircled{2}} + \underbrace{\sum_{n=1}^N \sum_{k=1}^K \gamma(z_{nk}) \ln p(\mathbf{x}_n | z_{nk} \phi)}_{\textcircled{3}} \quad (2)
 \end{aligned}$$

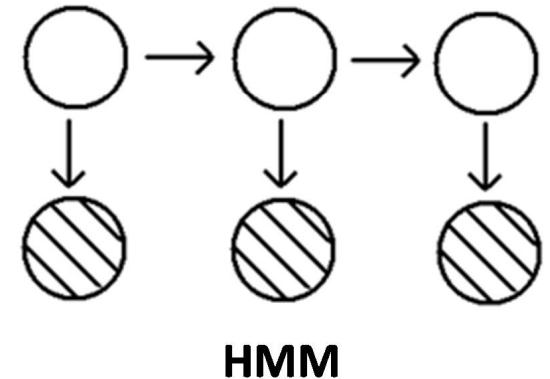
$$\text{Where } \gamma(z_{nk}) = p(z_{nk} | \mathbf{X}\theta^{\text{old}}) \quad (3)$$

$$\xi(z_{n-1,j} z_{n,k}) = p(z_{n-1,j} z_{n,k} | \mathbf{X}\theta^{\text{old}}) \quad (4)$$

$\pi$  is only associated with  $\textcircled{1}$

$A$  is only associated with  $\textcircled{2}$

$\phi$  is only associated with  $\textcircled{3}$



## E step

Evaluate  $p(\mathbf{Z} | \mathbf{X}\theta^{\text{old}})$ .

From the form of  $Q(\theta, \theta^{\text{old}})$  of HMM, there is no need to calculate  $p(\mathbf{Z} | \mathbf{X}\theta^{\text{old}})$  for all  $\mathbf{Z}$ . Many terms vanish. So just need to calculate

$$\gamma(z_{nk}) = p(z_{nk} | \mathbf{X}\theta^{\text{old}})$$

$$\xi(z_{n-1,j} z_{n,k}) = p(z_{n-1,j} z_{n,k} | \mathbf{X}\theta^{\text{old}})$$

Generally,  $p(\mathbf{Z} | \mathbf{X}\theta^{\text{old}})$  is difficult to estimate in PGM. While using dependency of HMM, we can simplify the computation.

We will show  $\gamma(z_{nk})$  can be decomposed to forward and backward term.

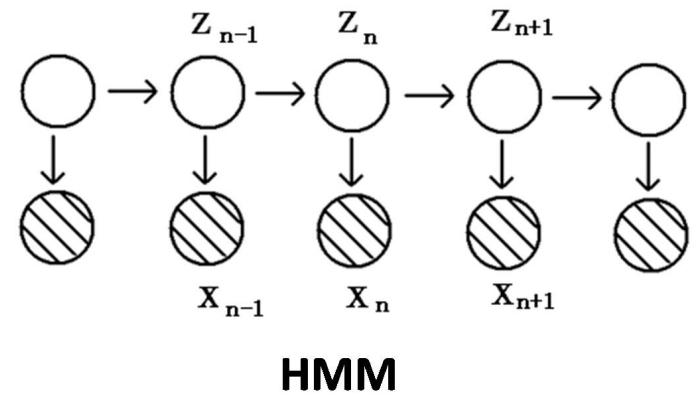
$$\gamma(\mathbf{z}_n) = \frac{\hat{\alpha}(\mathbf{z}_n)}{\text{forward}} \frac{\hat{\beta}(\mathbf{z}_n)}{\text{backward}} \quad (5)$$

Both of forward and backward term can be computed efficiently.

## Forward Backward algorithm

Using independence of PGM, we can decompose  $\gamma(\mathbf{z}_n | \mathbf{X})$  into forward and backward term.

$$\begin{aligned}
 \gamma(\mathbf{z}_n) &= p(\mathbf{z}_n | \mathbf{X}) \\
 &= \frac{p(\mathbf{X} | \mathbf{z}_n) p(\mathbf{z}_n)}{p(\mathbf{X})} \\
 &= \frac{p(\mathbf{x}_1, \dots, \mathbf{x}_n | \mathbf{z}_n) p(\mathbf{x}_{n+1}, \dots, \mathbf{x}_N | \mathbf{z}_n) p(\mathbf{z}_n)}{p(\mathbf{x}_1, \dots, \mathbf{x}_n) p(\mathbf{x}_{n+1}, \dots, \mathbf{x}_N | \mathbf{x}_1, \dots, \mathbf{x}_n)} \\
 &= \frac{p(\mathbf{x}_1, \dots, \mathbf{x}_n, \mathbf{z}_n) p(\mathbf{x}_{n+1}, \dots, \mathbf{x}_N | \mathbf{z}_n)}{p(\mathbf{x}_1, \dots, \mathbf{x}_n) p(\mathbf{x}_{n+1}, \dots, \mathbf{x}_N | \mathbf{x}_1, \dots, \mathbf{x}_n)} \\
 &= \hat{\alpha}(\mathbf{z}_n) \hat{\beta}(\mathbf{z}_n)
 \end{aligned} \tag{6}$$



where  $\hat{\alpha}(\mathbf{z}_n) = \frac{p(\mathbf{x}_1, \dots, \mathbf{x}_n, \mathbf{z}_n)}{p(\mathbf{x}_1, \dots, \mathbf{x}_n)}$  (Forward)  $\tag{7}$

$$\hat{\beta}(\mathbf{z}_n) = \frac{p(\mathbf{x}_{n+1}, \dots, \mathbf{x}_N | \mathbf{z}_n)}{p(\mathbf{x}_{n+1}, \dots, \mathbf{x}_N | \mathbf{x}_1, \dots, \mathbf{x}_n)} \tag{Backward} \tag{8}$$

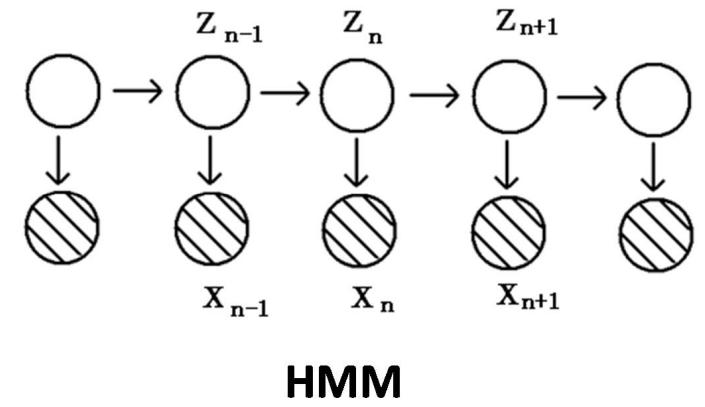
**Compute**  $\hat{\alpha}(\mathbf{z}_n)$

$\hat{\alpha}(\mathbf{z}_n)$  Can be computed efficiently using dependency of HMM

$$\text{denote } c_n = p(\mathbf{x}_n | \mathbf{x}_1, \dots, \mathbf{x}_{n-1}) \quad (9)$$

$$\hat{\alpha}(\mathbf{z}_1) = p(\mathbf{z}_1 | \mathbf{x}_1) = \frac{p(\mathbf{x}_1 | \mathbf{z}_1)p(\mathbf{z}_1)}{c_1} \quad (10)$$

$$\begin{aligned} \hat{\alpha}(\mathbf{z}_n) &= p(\mathbf{x}_1, \dots, \mathbf{x}_n, \mathbf{z}_n) / p(\mathbf{x}_1, \dots, \mathbf{x}_n) \\ &= p(\mathbf{x}_1, \dots, \mathbf{x}_n | \mathbf{z}_n)p(\mathbf{z}_n) / p(\mathbf{x}_1, \dots, \mathbf{x}_n) \\ &= p(\mathbf{x}_1, \dots, \mathbf{x}_{n-1}, \mathbf{z}_n)p(\mathbf{x}_n | \mathbf{z}_n) / p(\mathbf{x}_1, \dots, \mathbf{x}_n) \\ &= \sum_{\mathbf{z}_{n-1}} p(\mathbf{x}_1, \dots, \mathbf{x}_{n-1}, \mathbf{z}_{n-1}, \mathbf{z}_n)p(\mathbf{x}_n | \mathbf{z}_n) / p(\mathbf{x}_1, \dots, \mathbf{x}_n) \\ &= \sum_{\mathbf{z}_{n-1}} p(\mathbf{x}_1, \dots, \mathbf{x}_{n-1}, \mathbf{z}_n | \mathbf{z}_{n-1})p(\mathbf{z}_{n-1})p(\mathbf{x}_n | \mathbf{z}_n) / p(\mathbf{x}_1, \dots, \mathbf{x}_n) \\ &= \sum_{\mathbf{z}_{n-1}} p(\mathbf{x}_1, \dots, \mathbf{x}_{n-1} | \mathbf{z}_{n-1})p(\mathbf{z}_n | \mathbf{z}_{n-1})p(\mathbf{z}_{n-1})p(\mathbf{x}_n | \mathbf{z}_n) / (p(\mathbf{x}_1, \dots, \mathbf{x}_{n-1})p(\mathbf{x}_n | \mathbf{x}_1, \dots, \mathbf{x}_{n-1})) \\ &= p(\mathbf{x}_n | \mathbf{z}_n) \sum_{\mathbf{z}_{n-1}} p(\mathbf{x}_1, \dots, \mathbf{x}_{n-1} | \mathbf{z}_{n-1})p(\mathbf{z}_n | \mathbf{z}_{n-1}) / (p(\mathbf{x}_1, \dots, \mathbf{x}_{n-1})c_n) \\ &= \frac{1}{c_n} p(\mathbf{x}_n | \mathbf{z}_n) \sum_{\mathbf{z}_{n-1}} \hat{\alpha}(\mathbf{z}_{n-1})p(\mathbf{z}_n | \mathbf{z}_{n-1}) \end{aligned} \quad (11)$$



where  $c_1 = \sum_{\mathbf{z}_1} p(\mathbf{x}_1 | \mathbf{z}_1) p(\mathbf{z}_1)$  (12)

$$c_n = \sum_{\mathbf{z}_n} \left[ p(\mathbf{x}_n | \mathbf{z}_n) \sum_{\mathbf{z}_{n-1}} \hat{\alpha}(\mathbf{z}_{n-1}) p(\mathbf{z}_n | \mathbf{z}_{n-1}) \right] \quad (\text{integrate both side of (11)}) \quad (13)$$

**\*Byproduct: the likelihood of a sequence  $\mathbf{X}$ :**  $p(\mathbf{X}) = \prod_{n=1}^N c_n$  (14)

**Compute  $\hat{\beta}(\mathbf{z}_n)$**

$$\begin{aligned} \hat{\beta}(\mathbf{z}_n) &= p(\mathbf{x}_{n+1}, \dots, \mathbf{x}_N | \mathbf{z}_n) / p(\mathbf{x}_{n+1}, \dots, \mathbf{x}_N | \mathbf{x}_1, \dots, \mathbf{x}_n) \\ &= \sum_{\mathbf{z}_{n+1}} p(\mathbf{x}_{n+1}, \dots, \mathbf{x}_N | \mathbf{z}_{n+1}) / \left[ p(\mathbf{x}_{n+2}, \dots, \mathbf{x}_N | \mathbf{x}_1, \dots, \mathbf{x}_{n+1}) p(\mathbf{x}_{n+1} | \mathbf{x}_1, \dots, \mathbf{x}_n) \right] \\ &= \sum_{\mathbf{z}_{n+1}} p(\mathbf{x}_{n+1}, \dots, \mathbf{x}_N | \mathbf{z}_{n+1}) p(\mathbf{z}_{n+1}) / \left[ p(\mathbf{z}_n) / \left[ p(\mathbf{x}_{n+2}, \dots, \mathbf{x}_N | \mathbf{x}_1, \dots, \mathbf{x}_{n+1}) c_{n+1} \right] \right] \\ &= \sum_{\mathbf{z}_{n+1}} p(\mathbf{x}_{n+2}, \dots, \mathbf{x}_N | \mathbf{z}_{n+1}) p(\mathbf{x}_{n+1} | \mathbf{z}_{n+1}) p(\mathbf{z}_n | \mathbf{z}_{n+1}) p(\mathbf{z}_{n+1}) / \left[ p(\mathbf{z}_n) / \left[ p(\mathbf{x}_{n+2}, \dots, \mathbf{x}_N | \mathbf{x}_1, \dots, \mathbf{x}_{n+1}) c_{n+1} \right] \right] \\ &= \frac{1}{c_{n+1}} \sum_{\mathbf{z}_{n+1}} \hat{\beta}(\mathbf{z}_{n+1}) p(\mathbf{x}_{n+1} | \mathbf{z}_{n+1}) p(\mathbf{z}_{n+1} | \mathbf{z}_n) \end{aligned} \quad (15)$$

where  $\hat{\beta}(\mathbf{z}_n) = \frac{\gamma(\mathbf{z}_n)}{\hat{\alpha}(\mathbf{z}_n)} = \frac{p(\mathbf{z}_n | \mathbf{X})}{p(\mathbf{z}_n | \mathbf{X})} = 1$  (16)

**Compute**  $\xi(\mathbf{z}_{n-1}\mathbf{z}_n)$

Using conditional dependency of HMM

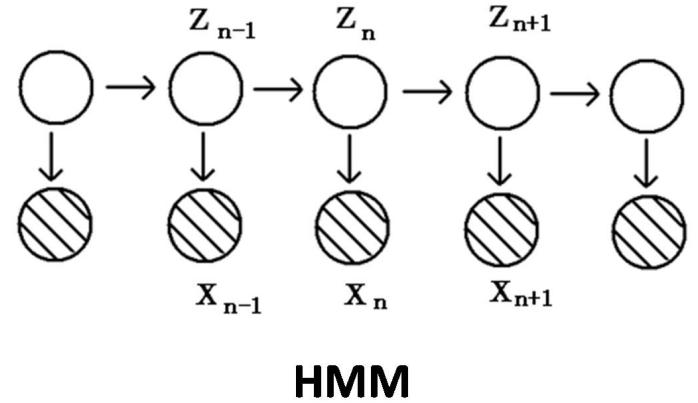
$$\xi(\mathbf{z}_{n-1}\mathbf{z}_n) = p(\mathbf{z}_{n-1}\mathbf{z}_n | \mathbf{X})$$

$$= \frac{p(\mathbf{X} | \mathbf{z}_{n-1}\mathbf{z}_n) p(\mathbf{z}_{n-1}\mathbf{z}_n)}{p(\mathbf{X})}$$

$$= \frac{p(\mathbf{x}_1, \dots, \mathbf{x}_{n-1} | \mathbf{z}_n) p(\mathbf{x}_n | \mathbf{z}_n) p(\mathbf{x}_{n+1}, \dots, \mathbf{x}_N | \mathbf{z}_n) p(\mathbf{z}_{n-1}\mathbf{z}_n)}{p(\mathbf{X})}$$

$$= \frac{p(\mathbf{x}_1, \dots, \mathbf{x}_{n-1}, \mathbf{z}_n) / p(\mathbf{z}_n) p(\mathbf{x}_n | \mathbf{z}_n) p(\mathbf{x}_{n+1}, \dots, \mathbf{x}_N | \mathbf{z}_n) p(\mathbf{z}_{n-1}\mathbf{z}_n)}{p(\mathbf{X})}$$

$$= \hat{\alpha}(\mathbf{z}_{n-1}) p(\mathbf{x}_n | \mathbf{z}_n) p(\mathbf{z}_n | \mathbf{z}_{n-1}) \hat{\beta}(\mathbf{z}_n) \quad (17)$$



## M step

In M step, need to optimize  $Q(\theta, \theta^{\text{old}})$  with respect to  $\theta = \{\pi, \mathbf{A}, \phi\}$

$$Q(\theta, \theta^{\text{old}}) = \sum_{k=1}^K \gamma(z_{nk}) \ln \pi_k + \sum_{n=2}^N \sum_{j=1}^K \sum_{k=1}^K \xi(z_{n-1,j} z_{n,k}) \ln A_{jk} + \sum_{n=1}^N \sum_{k=1}^K \gamma(z_{nk}) \ln p(\mathbf{x}_n | z_{nk} \phi) \quad (18)$$

  
①                    ②                    ③

$\pi$  is only associated with ①

$\mathbf{A}$  is only associated with ②

$\phi$  is only associated with ③

$\pi, \mathbf{A}, \phi$  are independent, so they can be optimized respectively.

① maximize  $\pi$

Take constraint  $\sum_{j=1}^K \pi_j = 1$  into consideration, introduce Lagrange function

$$R_1(\boldsymbol{\pi}, \lambda) = \sum_{k=1}^K \gamma(z_{nk}) \ln \pi_k + \lambda \left( \sum_{j=1}^K \pi_j - 1 \right) \quad (19)$$

Let  $\frac{\partial R_1(\boldsymbol{\pi}, \lambda)}{\partial \pi_k} = 0 \quad k = 1, \dots, K$   $\rightarrow \pi_k = \frac{\gamma_{1k}}{\sum_{j=1}^K \gamma_{1j}}, \quad k = 1, \dots, K$  (20)

② maximize  $\mathbf{A}$

Take constraint  $\sum_{l=1}^K A_{jl} = 1 \quad l = 1, \dots, K$  into consideration, introduce Lagrange function

$$R_2(\mathbf{A}, \lambda_1, \dots, \lambda_K) = \sum_{n=2}^N \sum_{j=1}^K \sum_{k=1}^K \xi(z_{n-1,j} z_{n,k}) \ln A_{jk} + \lambda_1 \left( \sum_{l=1}^K A_{1l} - 1 \right) + \dots + \lambda_K \left( \sum_{l=1}^K A_{Kl} - 1 \right) \quad (21)$$

Let  $\frac{\partial R_2(\mathbf{A}, \lambda_1, \dots, \lambda_K)}{\partial A_{jk}} = 0 \quad k = 1, \dots, K$   $\rightarrow A_{jk} = \frac{\sum_{n=2}^N \xi(z_{n-1,j} z_{n,k})}{\sum_{l=1}^K \sum_{n=2}^N \xi(z_{n-1,j} z_{n,l})} \quad j, k = 1, \dots, K$  (22)

This shows the elements which are zero in  $A_{jk}$  will keep zero all the time.  
So if you want to get left-right HMM just initialize  $\mathbf{A}$  as upper diagonal matrix.

### ③ Maximize $\phi$

the emission distribution can be multinomial, Gaussian, etc. We will discuss separately.

#### I . Emission distribution is discrete multinomial distribution

probability density function

$$p(\mathbf{x}_n | z_{nk} \phi) = \prod_{m=1}^M B_{km}^{x_{nm}} \quad s.t. \sum_{l=1}^M B_{kl} = 1, \quad k = 1, \dots, K \quad (23)$$

where  $\phi = \{B_{km}\}$   $k = 1, \dots, K, m = 1, \dots, M$

$B_{km}$  represents the probability of m-th event at k-th state.

To estimate  $B_{km}$ , introduce Lagrange function

$$R_3(\phi, \phi^{\text{old}}, \lambda_1, \dots, \lambda_K) = \sum_{n=1}^N \sum_{k=1}^K \gamma(z_{nk}) \ln p(\mathbf{x}_n | z_{nk} \phi) + \lambda_1 \left( \sum_{l=1}^M B_{1l} - 1 \right) + \lambda_K \left( \sum_{l=1}^M B_{Kl} - 1 \right) \quad (24)$$

$$\frac{\partial R_3(\phi, \phi^{\text{old}}, \lambda_1, \dots, \lambda_K)}{\partial B_{km}} = 0 \quad \rightarrow \quad B_{km} = \frac{\sum_{n=1}^N \gamma_{nk} x_{nm}}{\sum_{n=1}^N \gamma_{nk} \sum_{m=1}^M x_{nm}} \quad k = 1, \dots, K; m = 1, \dots, M \quad (25)$$

## EM algorithm for multinomial-HMM

1. Init parameters  $\theta = \{\pi, \mathbf{A}, \phi\}$

where  $\phi = \{B_{km}\} \quad k = 1, \dots, K; m = 1, \dots, M$

2. E step

Calculate  $\gamma(\mathbf{z}_n)$  using (6)

Calculate  $\xi(\mathbf{z}_{n-1} \mathbf{z}_n)$  using (17)

3. M step

$$\pi_k = \frac{\gamma_{1k}}{\sum_{j=1}^K \gamma_{1j}}, \quad k = 1, \dots, K \quad A_{jk} = \frac{\sum_{n=2}^N \xi(z_{n-1,j} z_{n,k})}{\sum_{l=1}^K \sum_{n=2}^N \xi(z_{n-1,j} z_{n,l})} \quad j, k = 1, \dots, K$$

$$B_{km} = \frac{\sum_{n=1}^N \gamma_{nk} x_{nm}}{\sum_{n=1}^N \gamma_{nk} \sum_{m=1}^M x_{nm}} \quad k = 1, \dots, K; m = 1, \dots, M$$

4. If converge then stop, otherwise goto 2.

## **Example**

Generate Multinomial-HMM for belowing sequences

Data{1} = [1 1 1 4 1 1 1 2 2 2 2 2 1 2 2 2 2 3 3 3 3 3 1 3 3 3]

Data{2} = [1 1 2 1 1 1 1 2 2 2 3 2 2 2 2 2 3 3 3 3 3 4 3 3 3]

State num: 3, multinomial num: 4

## **Output**

$$\pi = [1, 0, 0]$$

$$\mathbf{A} = \begin{bmatrix} 0.87 & 0.13 & 0.00 \\ 0.00 & 0.90 & 0.10 \\ 0.00 & 0.00 & 0.10 \end{bmatrix}$$

$$\mathbf{B} = \begin{bmatrix} 0.85 & 0.06 & 0.05 \\ 0.08 & 0.89 & 0.00 \\ 0.00 & 0.05 & 0.89 \\ 0.07 & 0.00 & 0.06 \end{bmatrix}$$

## II . Emission distribution is Gaussian distribution

$$\text{probability density function } p(\mathbf{x}_n | z_{nk}, \phi) = N(\mathbf{x}_n | \boldsymbol{\mu}_k \Sigma_k) \quad (26)$$

where  $\phi = \{\boldsymbol{\mu}_k, \Sigma_k\} \quad k = 1, \dots, K$

$$\begin{aligned} \text{Define } R_3(\phi, \phi^{old}) &= \sum_{n=1}^N \sum_{k=1}^K \gamma(z_{nk}) \ln p(\mathbf{x}_n | z_{nk}, \phi) \\ &= \sum_{n=1}^N \sum_{k=1}^K \gamma_{nk} \left[ -\frac{D}{2} \ln(2\pi) - \frac{1}{2} \ln |\Sigma_k| - \frac{1}{2} (\mathbf{x}_n - \boldsymbol{\mu}_k)^T \Sigma_k^{-1} (\mathbf{x}_n - \boldsymbol{\mu}_k) \right] \end{aligned} \quad (27)$$

To estimate parameters, let derivative of (27) with respect to

$\phi = \{\boldsymbol{\mu}_k, \Sigma_k\} \quad k = 1, \dots, K$  be zero

$$\begin{aligned} \frac{\partial R_3(\phi, \phi^{old})}{\partial \boldsymbol{\mu}_k} = 0 \quad k = 1, \dots, K &\quad \rightarrow \quad \boldsymbol{\mu}_k = \frac{\sum_{n=1}^N \gamma_{nk} \mathbf{x}_n}{\sum_{n=1}^N \gamma_{nk}} \quad k = 1, \dots, K \\ \frac{\partial R_3(\phi, \phi^{old})}{\partial \Sigma_k} = 0 \quad k = 1, \dots, K &\quad \rightarrow \quad \Sigma_k = \frac{\sum_{n=1}^N \gamma_{nk} (\mathbf{x}_n - \boldsymbol{\mu}_k)(\mathbf{x}_n - \boldsymbol{\mu}_k)^T}{\sum_{n=1}^N \gamma_{nk}} \quad k = 1, \dots, K \end{aligned} \quad (28)$$

## EM algorithm for Gaussian-HMM

1. Init parameters  $\theta = \{\pi, \mathbf{A}, \phi\}$

where  $\phi = \{\boldsymbol{\mu}_1, \dots, \boldsymbol{\mu}_K, \boldsymbol{\Sigma}_1, \dots, \boldsymbol{\Sigma}_K\} \quad k = 1, \dots, K$

2. E step

Calculate  $\gamma(\mathbf{z}_n)$  using (6)

Calculate  $\xi(\mathbf{z}_{n-1}\mathbf{z}_n)$  using (17)

3. M step

$$\pi_k = \frac{\gamma_{1k}}{\sum_{j=1}^K \gamma_{1j}}, \quad k = 1, \dots, K$$

$$A_{jk} = \frac{\sum_{n=2}^N \xi(z_{n-1,j} z_{n,k})}{\sum_{l=1}^K \sum_{n=2}^N \xi(z_{n-1,j} z_{n,l})} \quad j, k = 1, \dots, K$$

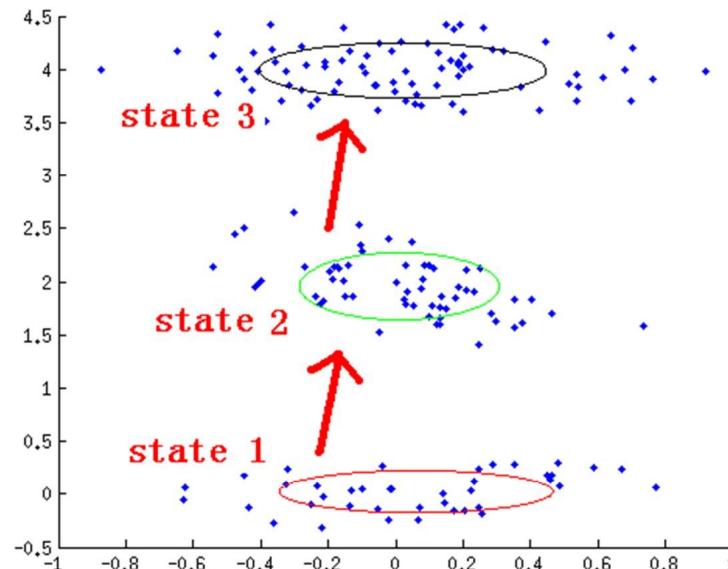
$$\boldsymbol{\mu}_k = \frac{\sum_{n=1}^N \gamma_{nk} \mathbf{x}_n}{\sum_{n=1}^N \gamma_{nk}} \quad k = 1, \dots, K$$

$$\boldsymbol{\Sigma}_k = \frac{\sum_{n=1}^N \gamma_{nk} (\mathbf{x}_n - \boldsymbol{\mu}_k)(\mathbf{x}_n - \boldsymbol{\mu}_k)^T}{\sum_{n=1}^N \gamma_{nk}} \quad k = 1, \dots, K$$

4. If converge then stop, otherwise goto 2.

## Example

The graph below shows the trained Gaussian-HMM model using created data.



State num: 3

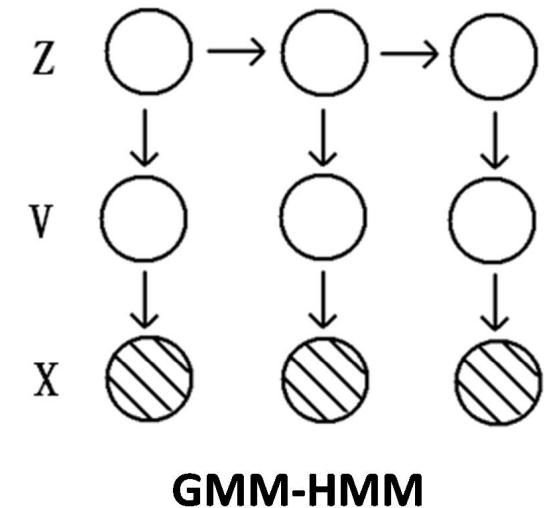
$$\pi = [1, 0, 0] \quad A = [0.95, 0.05, 0.00 \\ 0.00, 0.97, 0.03 \\ 0.00, 0.00, 1.00 ]$$

$\phi = \{\mu_k, \Sigma_k\} \quad k = 1, \dots, K$  is shown in graph

## GMM-HMM

Gaussian distribution is not able to capture distributions with many centers. Especially in ASR, where the timbre differentiates from person to person.

In GMM-HMM, every single state is a GMM model. Define  $v$  as 1-of- $K^*M$  random variable, where  $K$  is the number of state.  $M$  is the mixture number of GMM.  $v_{km}=1$  represents the  $k$ -th state,  $m$ -th mixture occurs.



Parameters are  $\theta = \{\pi, \mathbf{A}, \phi\}$

where  $\phi = \{B_{km}, \boldsymbol{\mu}_{km}, \boldsymbol{\Sigma}_{km}\} \quad k=1, \dots, K; m=1, \dots, M$

$B_{km} = p(v_{km} | z_k)$  is the probability of  $m$ -th mixture under  $k$ -th state

$\boldsymbol{\mu}_{km}, \boldsymbol{\Sigma}_{km}$  is mean and covariance of the  $k$ -th state,  $m$ -mixture, respectively.

Probability density function

$$p(\mathbf{x} | z_k \phi) = \sum_{m=1}^M p(v_{km} | z_k \phi) p(\mathbf{x} | v_{km} \phi) = \sum_{m=1}^M B_{km} N(\mathbf{x} | \boldsymbol{\mu}_{km}, \boldsymbol{\Sigma}_{km}) \quad (29)$$

## Apply EM algorithm to GMM-HMM

Compared with GMM-HMM, there are other latent variables  $\mathbf{v}_n$ , which dominates the emission GMM

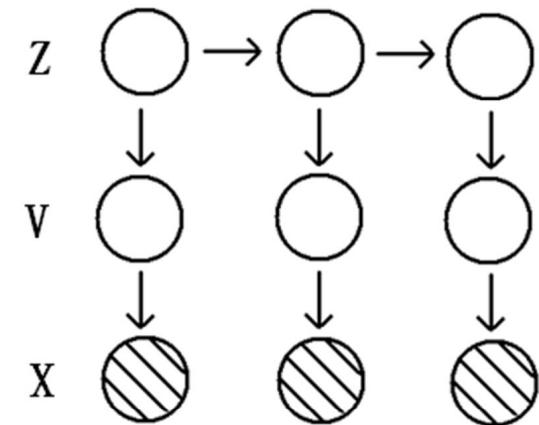
denote  $\mathbf{X} = \{\mathbf{x}_1, \dots, \mathbf{x}_N\}$   
 $\mathbf{V} = \{\mathbf{v}_1, \dots, \mathbf{v}_N\}$   
 $\mathbf{Z} = \{\mathbf{z}_1, \dots, \mathbf{z}_N\}$

(30)

Using independence of GMM-HMM, calculate  $Q(\theta, \theta^{\text{old}})$

**GMM-HMM**

$$\begin{aligned}
 Q(\theta, \theta^{\text{old}}) &= \sum_{\mathbf{VZ}} p(\mathbf{VZ} | \mathbf{X}\theta^{\text{old}}) \ln p(\mathbf{VXZ} | \theta) \\
 &= \sum_{\mathbf{VZ}} p(\mathbf{VZ} | \mathbf{X}\theta^{\text{old}}) \left[ \ln p(\mathbf{z}_1) + \sum_{n=2}^N \ln p(\mathbf{z}_n | \mathbf{z}_{n-1}) + \sum_{n=1}^N \ln p(\mathbf{v}_n | \mathbf{z}_n) + \sum_{n=1}^N \ln p(\mathbf{x}_n | \mathbf{v}_n) \right] \\
 &= \sum_{\mathbf{z}_1} p(\mathbf{z}_1 | \mathbf{X}\theta^{\text{old}}) \ln p(\mathbf{z}_1) + \sum_{n=2}^N \sum_{\mathbf{z}_{n-1}\mathbf{z}_n} p(\mathbf{z}_{n-1}\mathbf{z}_n | \mathbf{X}\theta^{\text{old}}) \ln p(\mathbf{z}_n | \mathbf{z}_{n-1}) + \sum_{n=1}^N \sum_{\mathbf{v}_n\mathbf{z}_n} p(\mathbf{v}_n\mathbf{z}_n | \mathbf{X}\theta^{\text{old}}) [\ln p(\mathbf{v}_n | \mathbf{z}_n) + \ln p(\mathbf{x}_n | \mathbf{v}_n)] \\
 &= \sum_{k=1}^K \gamma(z_{1k}) \ln \pi_k + \underbrace{\sum_{n=1}^N \sum_{j=1}^K \sum_{k=1}^K \xi(z_{n-1,j} z_{n,k}) \ln A_{jk}}_{\textcircled{1}} + \underbrace{\sum_{n=1}^N \sum_{k=1}^K \sum_{m=1}^M \eta_{nkm} \left[ \ln B_{km} - \frac{D}{2} \ln(2\pi) - \frac{1}{2} \ln |\Sigma_{km}| - \frac{1}{2} (\mathbf{x}_n - \boldsymbol{\mu}_k)^T \Sigma_{km}^{-1} (\mathbf{x}_n - \boldsymbol{\mu}_k) \right]}_{\textcircled{2}} \quad (31)
 \end{aligned}$$



## E step

Calculate  $p(\mathbf{VZ} | \mathbf{X}\theta^{\text{old}})$

From (31), no need to calculate all  $p(\mathbf{VZ} | \mathbf{X}\theta^{\text{old}})$ , only the terms below is needed

$$\gamma(z_{nk}) = p(z_{nk} | \mathbf{X}\theta^{\text{old}}) \quad (32)$$

$$\xi(z_{n-1,j} z_{n,k}) = p(z_{n-1,j} z_{n,k} | \mathbf{X}\theta^{\text{old}}) \quad (33)$$

$$\eta_{nkm} = p(v_{nkm} z_{nk} | \mathbf{X}\theta^{\text{old}}) \quad (34)$$

(32), (33) can be calculated in same way as (6), (17)

To calculate (34), just use forward-backward algorithm similar to (6)

$$\eta(\mathbf{v}_n \mathbf{z}_n) = p(\mathbf{v}_n \mathbf{z}_n | \mathbf{X}) = \underbrace{\hat{\alpha}(\mathbf{v}_n \mathbf{z}_n)}_{\text{forward}} \underbrace{\hat{\beta}(\mathbf{v}_n \mathbf{z}_n)}_{\text{backward}} \quad (35)$$

$$\text{where } \hat{\alpha}(\mathbf{v}_n \mathbf{z}_n) = \frac{p(\mathbf{x}_1, \dots, \mathbf{x}_n, \mathbf{v}_n \mathbf{z}_n)}{p(\mathbf{x}_1, \dots, \mathbf{x}_1)} \quad (36)$$

$$\hat{\beta}(\mathbf{v}_n \mathbf{z}_n) = \frac{p(\mathbf{x}_{n+1}, \dots, \mathbf{x}_N | \mathbf{v}_n \mathbf{z}_n)}{p(\mathbf{x}_{n+1}, \dots, \mathbf{x}_N | \mathbf{x}_1, \dots, \mathbf{x}_n)} \quad (37)$$

denote  $c_n = p(\mathbf{x}_n | \mathbf{x}_1, \dots, \mathbf{x}_{n-1})$

$\hat{\alpha}(\mathbf{v}_n \mathbf{z}_n)$ ,  $\hat{\beta}(\mathbf{v}_n \mathbf{z}_n)$ ,  $c_n$  can be computed similar to (10) – (16)

Results are given without detailed deduction

$$\hat{\alpha}(\mathbf{v}_1 \mathbf{z}_1) = p(\mathbf{v}_1 \mathbf{z}_1 | \mathbf{x}_1) = \frac{p(\mathbf{x}_1 | \mathbf{v}_1 \mathbf{z}_1) p(\mathbf{v}_1 \mathbf{z}_1)}{p(\mathbf{x}_1)} = \frac{p(\mathbf{x}_1 | \mathbf{v}_1) p(\mathbf{v}_1 | \mathbf{z}_1) p(\mathbf{z}_1)}{c_1} \quad (38)$$

$$\hat{\alpha}(\mathbf{v}_n \mathbf{z}_n) = \frac{1}{c_n} p(\mathbf{x}_n | \mathbf{v}_n) p(\mathbf{v}_n | \mathbf{z}_n) \sum_{\mathbf{v}_{n-1} \mathbf{z}_{n-1}} \hat{\alpha}(\mathbf{v}_{n-1} \mathbf{z}_{n-1}) p(\mathbf{z}_n | \mathbf{z}_{n-1}) \quad (39)$$

$$c_1 = p(\mathbf{x}_1) = \sum_{\mathbf{v}_1 \mathbf{z}_1} p(\mathbf{v}_1 \mathbf{z}_1 | \mathbf{x}_1) = \sum_{\mathbf{v}_1 \mathbf{z}_1} p(\mathbf{z}_1) p(\mathbf{v}_1 | \mathbf{z}_1) p(\mathbf{x}_1 | \mathbf{v}_1) \quad (40)$$

$$c_n = \sum_{\mathbf{v}_n \mathbf{z}_n} \left[ p(\mathbf{x}_n | \mathbf{v}_n) p(\mathbf{v}_n | \mathbf{z}_n) \sum_{\mathbf{v}_{n-1} \mathbf{z}_{n-1}} \hat{\alpha}(\mathbf{v}_{n-1} \mathbf{z}_{n-1}) p(\mathbf{z}_n | \mathbf{z}_{n-1}) \right] \quad (41)$$

$$\hat{\beta}(\mathbf{v}_N \mathbf{z}_N) = 1 \quad (42)$$

$$\hat{\beta}(\mathbf{v}_n \mathbf{z}_n) = \frac{1}{c_{n+1}} \sum_{\mathbf{v}_{n+1} \mathbf{z}_{n+1}} \hat{\beta}(\mathbf{v}_{n+1} \mathbf{z}_{n+1}) p(\mathbf{x}_{n+1} | \mathbf{v}_{n+1}) p(\mathbf{z}_{n+1} | \mathbf{z}_n) p(\mathbf{v}_{n+1} | \mathbf{z}_{n+1}) \quad (43)$$

## M step

**Maximize  $\pi, A$**

The maximization of  $\pi, A$  is same as (20), (22).

$$\pi_k = \frac{\gamma_{1k}}{\sum_{j=1}^K \gamma_{1j}}, \quad k = 1, \dots, K \quad (44)$$

$$A_{jk} = \frac{\sum_{n=2}^N \xi(z_{n-1,j} z_{n,k})}{\sum_{l=1}^K \sum_{n=2}^N \xi(z_{n-1,j} z_{n,l})} \quad j, k = 1, \dots, K \quad (45)$$

**Maximize  $\phi$**

To maximize  $\phi = \{B_{km}, \mu_{km}, \Sigma_{km}\}$   $k = 1, \dots, K; m = 1, \dots, M$

Taking constraint  $\sum_{l=1}^M B_{kl} = 1, \quad k = 1, \dots, K$  into account

Introduce Lagrange function

$$R_3(\phi, \phi^{old}, \lambda_1, \dots, \lambda_K) = \sum_{n=1}^N \sum_{k=1}^K \sum_{m=1}^M \eta_{nkm} \left[ \ln B_{km} - \frac{D}{2} \ln(2\pi) - \frac{1}{2} \ln |\Sigma_{km}| - \frac{1}{2} (\mathbf{x}_n - \boldsymbol{\mu}_k)^T \boldsymbol{\Sigma}_{km}^{-1} (\mathbf{x}_n - \boldsymbol{\mu}_k) \right] + \lambda_1 \left( \sum_{l=1}^M B_{1l} - 1 \right) + \lambda_K \left( \sum_{l=1}^M B_{Kl} - 1 \right) \quad (46)$$

Let the derivative of  $R_3(\phi, \phi^{\text{old}}, \lambda_1, \dots, \lambda_K)$  With respect to  
 $\phi = \{B_{km}, \mu_{km}, \Sigma_{km}\} \quad k = 1, \dots, K; m = 1, \dots, M$  be zero

$$\frac{\partial R_3(\phi, \phi^{\text{old}}, \lambda_1, \dots, \lambda_K)}{\partial B_{km}} = 0$$

$$B_{km} = \frac{\sum_{n=1}^N \eta_{nkm}}{\sum_{n=1}^N \sum_{j=1}^M \eta_{nj}}$$

$$\frac{\partial R_3(\phi, \phi^{\text{old}}, \lambda_1, \dots, \lambda_K)}{\partial \mu_{km}} = 0$$



$$\mu_{km} = \frac{\sum_{n=1}^N \eta_{nkm} \mathbf{x}_n}{\sum_{n=1}^N \eta_{nkm}} \quad k = 1, \dots, K; m = 1, \dots, M \quad (47)$$

$$\frac{\partial R_3(\phi, \phi^{\text{old}}, \lambda_1, \dots, \lambda_K)}{\partial \Sigma_{km}} = 0$$

$$\Sigma_{km} = \frac{\sum_{n=1}^N \eta_{nkm} (\mathbf{x}_n - \mu_{km})(\mathbf{x}_n - \mu_{km})^T}{\sum_{n=1}^N \eta_{nkm}}$$

## EM algorithm for GMM-HMM

1. Init parameters  $\theta = \{\pi, \mathbf{A}, \phi\}$

where  $\phi = \{B_{km}, \boldsymbol{\mu}_{km}, \boldsymbol{\Sigma}_{km}\} \quad k = 1, \dots, K; m = 1, \dots, M$

2. E step

Calculate  $\gamma(\mathbf{z}_n)$ ,  $\xi(\mathbf{z}_{n-1}\mathbf{z}_n)$ ,  $\eta(\mathbf{v}_n \mathbf{z}_n)$  from (32), (33), (34)

3. M step

$$\pi_k = \frac{\gamma_{1k}}{\sum_{j=1}^K \gamma_{1j}}, \quad k = 1, \dots, K$$

$$A_{jk} = \frac{\sum_{n=2}^N \xi(z_{n-1,j} z_{n,k})}{\sum_{l=1}^K \sum_{n=2}^N \xi(z_{n-1,j} z_{n,l})} \quad j, k = 1, \dots, K$$

$$B_{km} = \frac{\sum_{n=1}^N \eta_{nkm}}{\sum_{n=1}^N \sum_{j=1}^M \eta_{nj}}$$

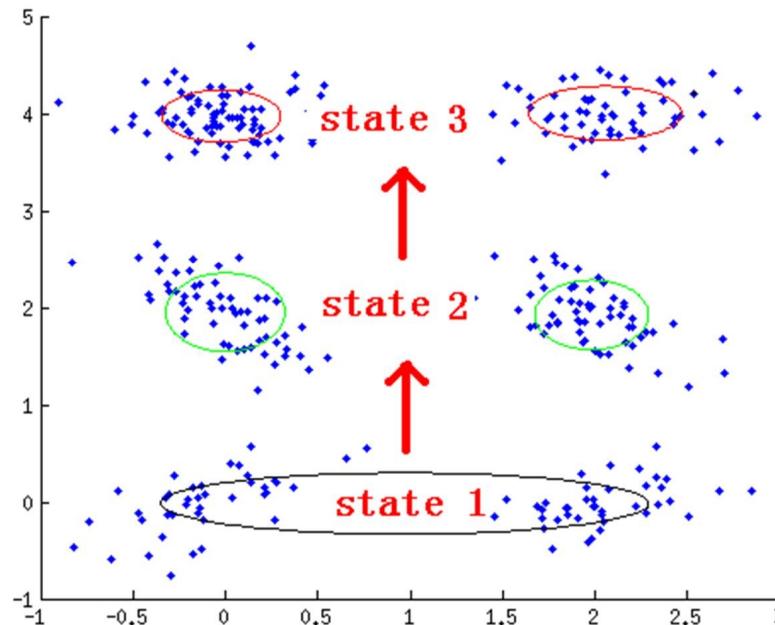
$$\boldsymbol{\mu}_{km} = \frac{\sum_{n=1}^N \eta_{nkm} \mathbf{x}_n}{\sum_{n=1}^N \eta_{nkm}}$$

$$\boldsymbol{\Sigma}_{km} = \frac{\sum_{n=1}^N \eta_{nkm} (\mathbf{x}_n - \boldsymbol{\mu}_{km})(\mathbf{x}_n - \boldsymbol{\mu}_{km})^T}{\sum_{n=1}^N \eta_{nkm}}$$

4. If converge then stop, otherwise goto 2.

## Example

A GMM-HMM with state num:3, mix num:2



$$\pi: [0, 0, 1]$$

$$A: [1.00, 0.00, 0.00, \\ 0.03, 0.97, 0.00, \\ 0.00, 0.05, 0.95]$$

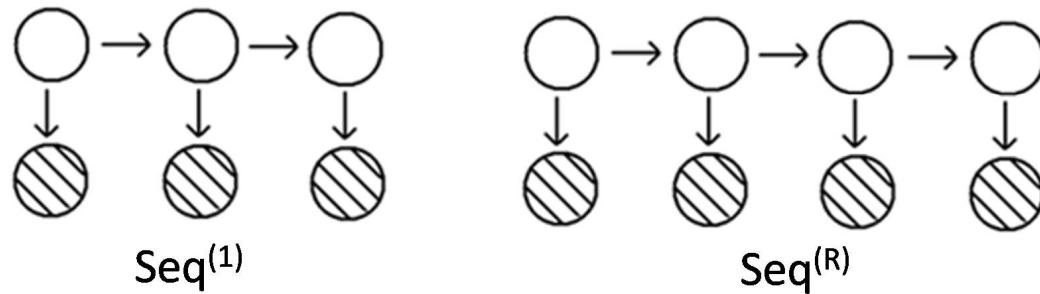
$$\phi = \{B_{km}, \mu_{km}, \Sigma_{km}\} \quad k=1, \dots, K; m=1, \dots, M$$

is shown on the graph

## Multi Sequence Training

Till now, we train HMM only use one sequence.

In ASR, we have many utterance to train one HMM model for each phoneme / word.  
This part will show how to train HMM using multi sequences.



Using independency of PGM,  $Q(\theta, \theta^{\text{old}})$

$$\begin{aligned}
 Q(\theta, \theta^{\text{old}}) &= \sum_{\mathbf{Z}} p(\mathbf{Z} | \mathbf{X}\theta^{\text{old}}) \ln p(\mathbf{XZ} | \theta) \\
 &= \sum_{r=1}^R \left( \sum_{\mathbf{Z}^{(r)}} p(\mathbf{Z}^{(r)} | \mathbf{X}^{(r)}\theta^{\text{old}}) \ln p(\mathbf{X}^{(r)}\mathbf{Z}^{(r)} | \theta) \right) \\
 &= \sum_{r=1}^R \sum_{k=1}^K \gamma_{1k}^{(r)} \ln \pi_k + \sum_{r=1}^R \sum_{n=2}^N \sum_{j=1}^K \sum_{k=1}^K \xi^{(r)}(z_{n-1,j} z_{n,k}) \ln A_{jk} + \sum_{l=r}^R \sum_{n=1}^N \sum_{k=1}^K \gamma_{nk}^{(r)} \ln p(\mathbf{x}_n | z_{nk}) \quad (48)
 \end{aligned}$$

Where  $\mathbf{X}^{(r)}$  is the observation variables of r-th seq.  
 $\mathbf{Z}^{(r)}$  is the latent variables of r-th seq.

## E STEP

Estimate  $\gamma$ ,  $\xi$ ,  $\eta$  for each sequences separately.

$$\gamma_{nk}^{(r)} = p(z_{nk}^{(r)} \mid \mathbf{X}^{(r)} \theta^{old}) \quad (49)$$

$$\xi^{(r)}(z_{n-1,j} z_{n,k}) = p(z_{n-1,j}^{(r)} z_{n,k}^{(r)} \mid \mathbf{X}^{(r)} \theta^{old}) \quad (50)$$

$$\eta_{nkm}^{(r)} = p(v_{nkm}^{(r)} z_{nk}^{(r)} \mid \mathbf{X} \theta^{old}) \quad (\text{for GMM only}) \quad (51)$$

## M STEP

Optimize (48) with respect to  $\theta = \{\pi, \mathbf{A}, \phi\}$

$$\pi_k = \frac{\sum_{r=1}^R \gamma_{1k}^{(r)}}{\sum_{r=1}^R \sum_{j=1}^K \gamma_{1j}^{(r)}} \quad (52)$$

$$A_{jk} = \frac{\sum_{r=1}^R \sum_{n=2}^N \xi^{(r)}(z_{n-1,j} z_{n,k})}{\sum_{r=1}^R \sum_{l=1}^K \sum_{n=2}^N \xi^{(r)}(z_{n-1,j} z_{n,l})} \quad (53)$$

## I . Multinomial Distribution

$$B_{km} = \frac{\sum_{r=1}^R \sum_{n=1}^{N^{(r)}} \gamma_{nk}^{(r)} x_{nm}^{(r)}}{\sum_{r=1}^R \sum_{n=1}^{N^{(r)}} \gamma_{nk}^{(r)} \sum_{m=1}^M x_{nm}^{(r)}} \quad (54)$$

## II . Gauss Distribution

$$\mu_k = \frac{\sum_{r=1}^R \sum_{n=1}^{N^{(r)}} \gamma_{nk}^{(r)} \mathbf{x}_n^{(r)}}{\sum_{r=1}^R \sum_{n=1}^{N^{(r)}} \gamma_{nk}^{(r)}} \quad \Sigma_k = \frac{\sum_{r=1}^R \sum_{n=1}^{N^{(r)}} \gamma_{nk}^{(r)} (\mathbf{x}_n^{(r)} - \mu_k)(\mathbf{x}_n^{(r)} - \mu_k)^T}{\sum_{r=1}^R \sum_{n=1}^{N^{(r)}} \gamma_{nk}^{(r)}} \quad (55)$$

## III. GMM

$$B_{km} = \frac{\sum_{r=1}^R \sum_{n=1}^{N^{(r)}} \eta_{nkm}^{(r)}}{\sum_{r=1}^R \sum_{n=1}^{N^{(r)}} \sum_{j=1}^M \eta_{nj}^{(r)}} \quad \mu_{km} = \frac{\sum_{r=1}^R \sum_{n=1}^{N^{(r)}} \eta_{nkm}^{(r)} \mathbf{x}_n}{\sum_{r=1}^R \sum_{n=1}^{N^{(r)}} \eta_{nkm}^{(r)}} \quad \Sigma_{km} = \frac{\sum_{r=1}^R \sum_{n=1}^{N^{(r)}} \eta_{nkm}^{(r)} (\mathbf{x}_n - \mu_{km})(\mathbf{x}_n - \mu_{km})^T}{\sum_{r=1}^R \sum_{n=1}^{N^{(r)}} \eta_{nkm}^{(r)}} \quad (56)$$

## Decoding of HMM

**Q:** How to find the best decoding path?

$$\begin{aligned}\mathbf{A:} \quad \mathbf{Z}_{opt} &= \max_{\mathbf{Z}} p(\mathbf{Z} | \mathbf{X}) = \max_{\mathbf{Z}} p(\mathbf{ZX}) \\ &= \max_{\mathbf{Z}} p(\mathbf{z}_1) \prod_{n=2}^N p(\mathbf{z}_n | \mathbf{z}_{n-1}) \prod_{n=1}^N p(\mathbf{x}_n | \mathbf{z}_n)\end{aligned}\tag{57}$$

However, we need to evaluate all possible  $\mathbf{Z}$  which is  $K^N$  times to get accurate solution. This infeasible.

## Viterbi Algorithm

Use greedy algorithm to estimate optimized path step by step.

By discarding paths with low probability and storing previous step,  
the computation complexity decreased to K\*N

### Viterbi Algorithm

$$V_{1k} = p(\mathbf{x}_1 | z_{1k}) \times p(z_{1k})$$

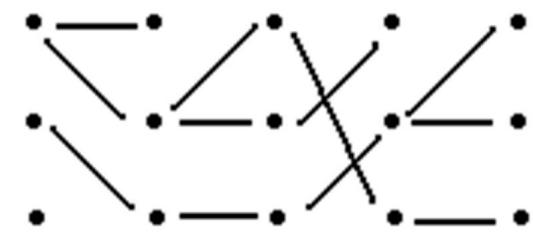
$$V_{1k} = p(\mathbf{x}_1 | z_{1k}) \times p(z_{1k})$$

for  $n = 2 : N$

$$V_{nk} = \max_j (V_{n-1,j} \times p(z_{nk} | z_{n-1,j}) \times p(\mathbf{x}_n | z_{nj}))$$

$$\text{path}(n-1) = j$$

$$\text{path}(n) = \arg \max_j V_{nj}$$

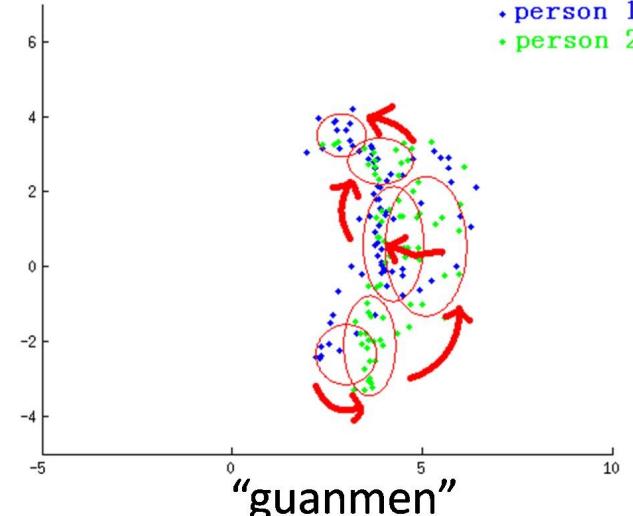
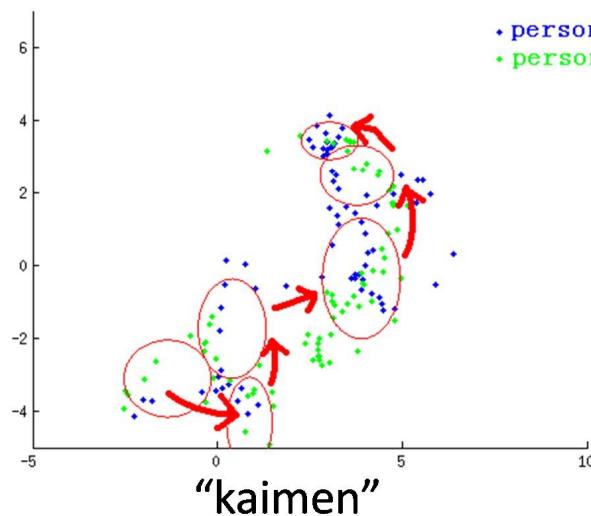


Path stored

# Experiments in small vocabulary ASR

12 Mfcc feature (only choose the 1<sup>st</sup> and 2<sup>nd</sup> dimension to plot)

## 1. Gaussian-HMM, state num: 6



$$\pi = [1, 0, 0, 0, 0, 0]$$

$$A = \begin{matrix} 0.86 & 0.14 & 0 & 0 & 0 & 0 \\ 0 & 0.90 & 0.10 & 0 & 0 & 0 \\ 0 & 0 & 0.87 & 0.13 & 0 & 0 \\ 0 & 0 & 0 & 0.97 & 0.03 & 0 \\ 0 & 0 & 0 & 0 & 0.91 & 0.09 \\ 0 & 0 & 0 & 0 & 0 & 1.00 \end{matrix}$$

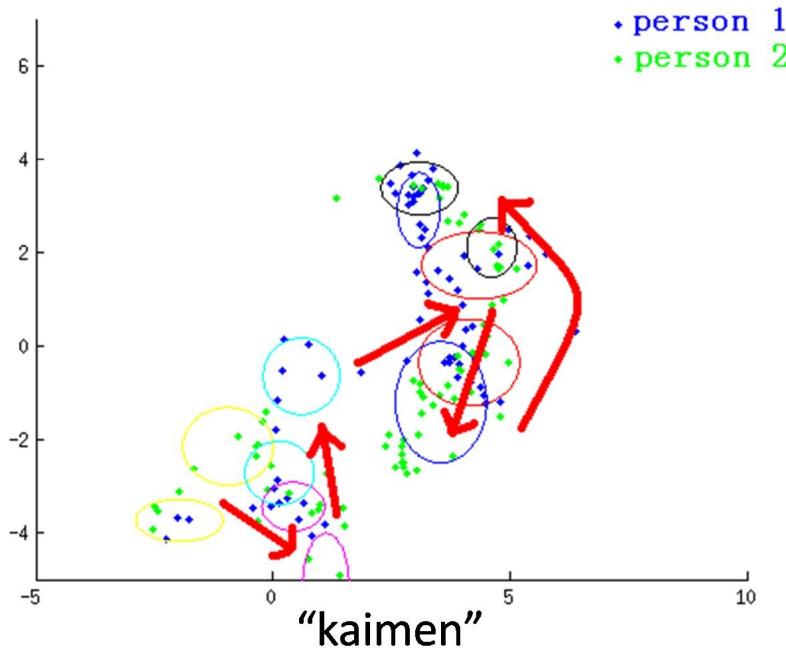
$$\pi = [1, 0, 0, 0, 0, 0]$$

$$A = \begin{matrix} 0.89 & 0.11 & 0 & 0 & 0 & 0 \\ 0 & 0.87 & 0.13 & 0 & 0 & 0 \\ 0 & 0 & 0.77 & 0.23 & 0 & 0 \\ 0 & 0 & 0 & 0.97 & 0.03 & 0 \\ 0 & 0 & 0 & 0 & 0.95 & 0.05 \\ 0 & 0 & 0 & 0 & 0 & 1.00 \end{matrix}$$

## 2. GMM-HMM (6 states, 2 mixutre)

With the increase of dataset, Gaussian-HMM is not able to capture the timbre of different people, gender, age.

We can use GMM-HMM model instead.



However, GMM-HMM is sensitive to the initial parameters. The tricks we are using in this example is:

1. Use Gaussian-HMM to train different people separately.
2. Combine the data point which are in the same state. Use GMM to initialize the parameters.
3. Run GMM-HMM to fine-tune the model.

# Results

**Dataset:** 20 Isolated Chinese words. 11 male + 9 female. Altogether 800 pronunciations.

10 male and 9 female for training. 10 male and 10 female for testing.

**Feature:** 12 dimension MFCC

**Model:** Gaussian-HMM, GMM-HMM

Model	Accuracy
Gaussian-HMM	82.25%
GMM-HMM	84.00%

## Weakness of HMM

### 1. Markov assumption

The next state is only dependent upon the current state. So is poor at capturing long-range correlations between the observed variables.

$$p(\mathbf{z}_{n+1} | \mathbf{z}_1, \dots, \mathbf{z}_n) = p(\mathbf{z}_{n+1} | \mathbf{z}_n) \quad (58)$$

### 2. Stationary assumption

$$p(\mathbf{z}_{n+1} | \mathbf{z}_n) = p(\mathbf{z}_n | \mathbf{z}_{n-1}) \quad (59)$$

### 3. Output independence assumption

The current output is conditionally independent of the previous output.

$$p(\mathbf{X} | \mathbf{Z}) = \prod_{n=1}^N p(\mathbf{x}_n | \mathbf{Z}) \quad (60)$$

## Tricks of HMM

1. HMM is more sensitive to the initial parameters than GMM. So it is easy to get into local minimum.

**Solve:** Use GMM or other methods to initialize parameters.

Initialize parameters randomly and run HMM separately for several times.

2. For Gaussian-HMM & GMM-HMM, if  $\text{eig}(\Sigma)$  is too small. Then  $\Sigma^{-1}$  will be unstable.

**Solve:** if  $\text{eig}(\Sigma) < \epsilon$  then  $\Sigma = \Sigma + \sigma I$

3. If  $p(\mathbf{x}|\mathbf{z})$  is Gaussian or GMM pdf, underflow may occurs.

$$p(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{(2\pi)^{D/2} |\boldsymbol{\Sigma}|^{1/2}} \exp\left(-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu})\right)$$

if this term is too small,  
after exp it will underflow

**Solve:** use  $\ln p(\mathbf{x}|\mathbf{z})$  to instead  $p(\mathbf{x}|\mathbf{z})$  in code implementation.

For Gaussian

$$\ln p(\mathbf{x}|\mathbf{z}) = \ln \pi - \frac{D}{2} \ln(2\pi) - \frac{1}{2} \ln |\boldsymbol{\Sigma}| - \frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu})$$

For GMM

$$\begin{aligned} \ln p(\mathbf{x}|\mathbf{z}) &= \ln \sum_{m=1}^m \pi_m N(\mathbf{x}|\boldsymbol{\mu}_m, \boldsymbol{\Sigma}_m) \\ &= \ln \sum_{m=1}^m \exp\left(\ln \pi_m - \frac{D}{2} \ln(2\pi) - \frac{1}{2} \ln |\boldsymbol{\Sigma}_m| - \frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_m)^T \boldsymbol{\Sigma}_m^{-1} (\mathbf{x} - \boldsymbol{\mu}_m)\right) \\ &= \left[ \ln \sum_{m=1}^m \exp\left(\ln \pi_m - \frac{D}{2} \ln(2\pi) - \frac{1}{2} \ln |\boldsymbol{\Sigma}_m| - \frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_m)^T \boldsymbol{\Sigma}_m^{-1} (\mathbf{x} - \boldsymbol{\mu}_m) - U\right) \right] + U \end{aligned}$$

if this term is too small,  
after exp it will underflow

Normalization factor,  
To avoid underflow

4. According to (11), (15),  $\hat{\alpha}(\mathbf{z}_n), \hat{\beta}(\mathbf{z}_n)$  may underflow, and  $\gamma(\mathbf{z}_n)$  will be unstable

**Solve:** Use  $\ln \gamma(\mathbf{z}_n), \ln c_n, \ln \hat{\alpha}(\mathbf{z}_n), \ln \hat{\beta}(\mathbf{z}_n)$  replace  $\gamma(\mathbf{z}_n), c_n, \hat{\alpha}(\mathbf{z}_n), \hat{\beta}(\mathbf{z}_n)$

$$\begin{aligned}\ln \hat{\alpha}(\mathbf{z}_n) &= -\ln c_n + \ln p(\mathbf{x}_n | \mathbf{z}_n) + \ln \sum_{\mathbf{z}_{n-1}} \exp \left( \ln \hat{\alpha}(\mathbf{z}_{n-1}) + \ln p(\mathbf{z}_n | \mathbf{z}_{n-1}) \right) \\ &= -\ln c_n + \ln p(\mathbf{x}_n | \mathbf{z}_n) + \left[ \ln \sum_{\mathbf{z}_{n-1}} \exp \left( \ln \hat{\alpha}(\mathbf{z}_{n-1}) + \ln p(\mathbf{z}_n | \mathbf{z}_{n-1}) \right) - U \right] + U\end{aligned}$$

if this term is too small,  
after exp it will underflow

Normalization factor,  
To avoid underflow

The same strategy can be applied to  $\ln c_n, \ln \hat{\beta}(\mathbf{z}_n)$

$\gamma(\mathbf{z}_n) = \hat{\alpha}(\mathbf{z}_n) \hat{\beta}(\mathbf{z}_n)$  will turn to

$$\ln \gamma(\mathbf{z}_n) = \ln \hat{\alpha}(\mathbf{z}_n) + \ln \hat{\beta}(\mathbf{z}_n)$$

Furthermore, for parameter estimation, such as (28) will turn to

$$\boldsymbol{\mu}_k = \frac{\sum_{n=1}^N \gamma_{nk} \mathbf{x}_n}{\sum_{n=1}^N \gamma_{nk}} = \frac{\sum_{n=1}^N \exp[\ln \gamma_{nk}] \mathbf{x}_n}{\sum_{n=1}^N \exp[\ln \gamma_{nk}]} = \frac{\sum_{n=1}^N \exp[(\ln \gamma_{nk}) - U] \mathbf{x}_n}{\sum_{n=1}^N \exp[(\ln \gamma_{nk}) - U]}$$

**if this term is too small,  
after exp it will underflow**

**Normalization factor,  
To avoid underflow**

## **Matlab Code**

<https://github.com/qiuqiangkong/matlab-hmm>

**THANK YOU!**