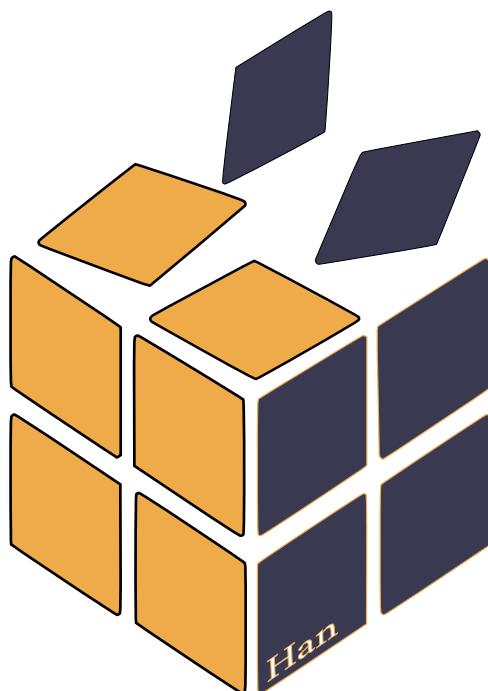


Deep Statistical Learning



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The logo depicts a cube puzzle gradually coming together, reflecting the journey of learning. Each piece represents a fragment of knowledge, and as they fall into place, they reveal the larger structure of understanding. It conveys the idea that growth is a process — knowledge is completed bit by bit.

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Part I

Introduction

Chapter 1

Basics

1.1 Naive Bayes

Let's say we have a prediction model and it gives us a prediction. We want to measure the precision of the prediction. For example, the prediction has 90% chance of being a correct prediction. This can be modeled probabilistically. Given data \mathbf{x} , we want to know the probability of it being y . Similarly, if we have N data, then with an i.i.d., assumption,

$$\prod_{i=1}^N p(y_i|\mathbf{x}_i).$$

What we want to is to build such a probabilistic model parameterized by some parameters.

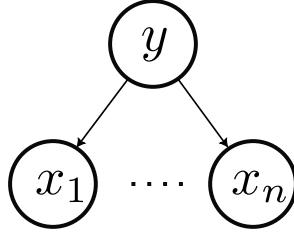
In this section, we discuss how to classify vectors of discrete-valued features \mathbf{x} . Recall that we discussed how to classify a feature vector \mathbf{x} by applying Bayes rule to a generative classifier of the form

$$p(y = c|\mathbf{x}, \boldsymbol{\theta}) \propto p(\mathbf{x}|y = c, \boldsymbol{\theta})p(y = c|\boldsymbol{\theta})$$

The key to using such models is specifying a suitable form for the class-conditional density $p(\mathbf{x}|y = c, \boldsymbol{\theta})$, which defines what kind of data we expect to see in each class.

- $\mathbf{x} \in \{1, \dots, K\}^D$,
 - K : the number of values for each feature.
 - D : the number of features.
- We will use a generative approach.
- Need to specify the class conditional distribution, $p(\mathbf{x}|y = c)$.
- A simple approach is to assume the features are **conditionally independence** given the class label.
- This allows us to write the class conditional density as a product of one dimensional densities:

$$p(\mathbf{x}|y = c, \boldsymbol{\theta}) = \prod_{j=1}^D p(x_j|y = c, \boldsymbol{\theta}_{jc})$$



The resulting model is called a **naive Bayes classifier (NBC)**. The model is called “naive” since we assume the independence between the features, which is not true in practice. However, it often results in classifiers that work well.

The form of the class-conditional density depends on the type of each feature. We give some possibilities below:

- In the case of real-valued features, we can use the Gaussian distribution: $p(\mathbf{x}|y = c, \boldsymbol{\theta}) = \prod_{j=1}^D \mathcal{N}(x_j|\mu_{jc}^2, \sigma_{jc}^2)$, where μ_{jc} is the mean of feature j in objects of class c , and σ_{jc}^2 is its variance.
- In the case of binary features, we can use the Bernoulli distribution: $p(\mathbf{x}|y = c, \boldsymbol{\theta}) = \prod_{j=1}^D \text{Ber}(x_j|\mu_{jc})$, where μ_{jc} is the probability that feature j occurs in class c . This is sometimes called the **multivariate Bernoulli naive Bayes** model.
- In the case of categorical features, $x_j \in \{1, \dots, K\}$, we can model the multinomial distribution: $p(\mathbf{x}|y = c, \boldsymbol{\theta}) = \prod_{j=1}^D \text{Cat}(x_j|\boldsymbol{\mu}_{jc})$, where $\boldsymbol{\mu}_{jc}$ is a histogram over the K possible values for x_j in class c .

The probability for a single data case is given by

$$p(\mathbf{x}_i, y_i | \boldsymbol{\theta}) = p(y_i | \boldsymbol{\pi}) \prod_j p(x_{ij} | \boldsymbol{\theta}_j) = \prod_c \pi_c^{\mathbb{I}(y_i=c)} \prod_j \prod_c p(x_{ij} | \boldsymbol{\theta}_{jc})^{\mathbb{I}(y_i=c)},$$

where $\boldsymbol{\pi}$ is a vector of class probability. Hence the log-likelihood is given by

$$\log p(\mathcal{D} | \boldsymbol{\theta}) = \sum_{c=1}^C N_c \log \pi_c + \sum_{j=1}^D \sum_{c=1}^C \sum_{i:y_i=c} \log p(x_{ij} | \boldsymbol{\theta}_{jc})$$

Algorithm 1: Fitting a naive Bayes classifier to binary features

```

Initialize  $N_c = 0, N_{jc} = 0$  ;
for  $i = 1 : N$  do
   $c = y_i$  //Class label of  $i$ -th example;
   $N_c := N_c + 1$ ;
  for  $j = 1 : D$  do
    if  $x_{ij} = 1$  then
       $| N_{jc} := N_{jc} + 1$ 
    end
  end
   $\hat{\pi} = \frac{N_c}{N}, \hat{\theta}_{jc} = \frac{N_{jc}}{N_c}$ 

```

Chapter 2

Evaluation Metrics

2.1 Basic Metrics

2.1.1 Confusion matrix

		Predicted	
		T	N
Actual	T	TP	FN
	N	FP	TN

Some special metrics:

- **Recall (or TPR):** Find all relevant cases within a dataset.

$$\frac{TP}{TP + FN}$$

- **Precision:** While recall expresses the ability to find all relevant instances in a dataset, precision expresses the proportion of the data points our model says was relevant actually were relevant.

$$\frac{TP}{TP + FP}$$

- **FPR:** Of all the actual negative cases, what fraction did my model incorrectly classify as positive?

$$\frac{FP}{FP + TN}$$

2.1.2 Receiver Operating Characteristic (ROC)

ROC curves are helpful when we need to identify how well each threshold performed in terms of the TPR and the FPR of binary classifier. *ROC* stands for *Receiver Operating Characteristic*, and the name comes from the graphs drawn during World War II that summarized how well radar operators correctly identified airplanes in radar signals. The illustrative ROC curve is given as follows:

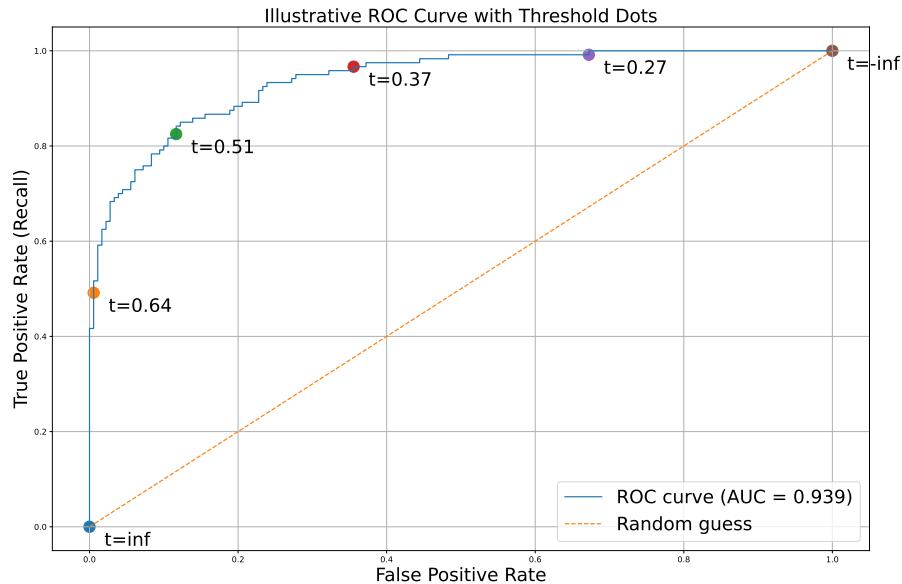


Figure 2.1: Each dot on the ROC curve represents the TPR and the FPR at a specific threshold.

- The diagonal line shows where the TRP = FPR.
- A point near $(0, 1)$ is ideal: low false alarms, high detection.
- Higher curve = better. The *AUC* (area under the ROC) summarizes performance across all thresholds (1.0 is perfect). AUC is especially helpful when we have to compare multiple models, since it provides a simple comparison.
- Choose the point that best matches your costs/requirements.

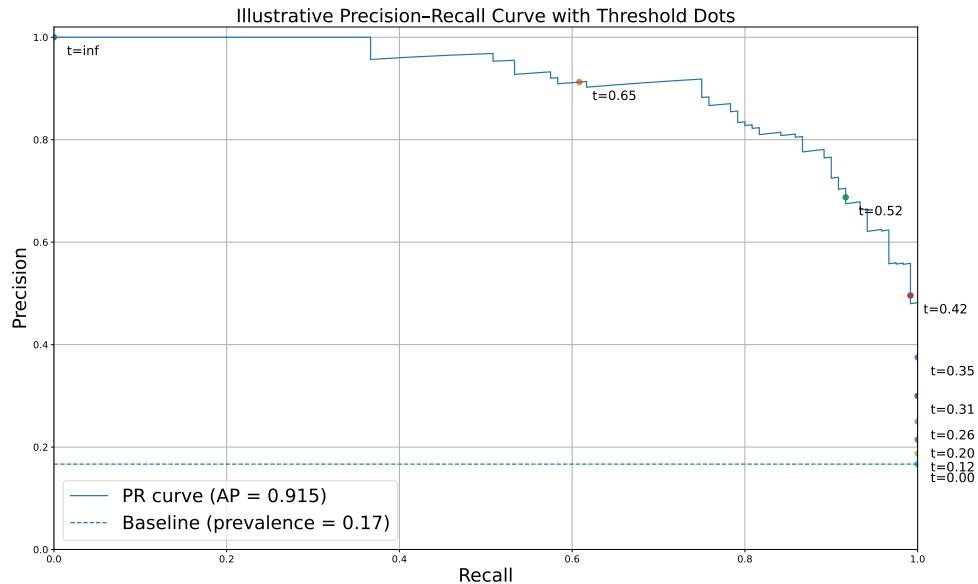
2.1.3 Precision Recall Graphs

When classes are highly imbalanced, ROC curves can be misleading. **With very few positives, the false positive rate (FPR) can stay low even for weak models, giving an overly optimistic picture.** In these cases, **precision–recall (PR) curves are more informative because they focus on the positive class:** they plot precision (y-axis) versus recall (x-axis), where recall equals TPR. PR curves highlight how many of the predicted positives are actually correct (precision) at each level of coverage (recall), making them better suited for rare-event detection.

This is because precision is more informative than FPR under heavy class imbalance: precision ignores true negatives (TN), whereas FPR uses them. The illustrative plot below shows the difference.

In sum, for imbalanced data, use PR curves (precision vs recall) instead of ROC: FPR can look small simply because true negatives dominate, while PR directly measures the quality of positive predictions.

Rule of thumbs:



- For rare positives or alerting workflows: use PR (and report precision@recall or recall@precision).
- For balanced data or TN-sensitive views: also show ROC/AUC.

Chapter 3

Dimensionality Reduction

3.1 Principal Component Analysis

3.1.1 Covariance and the weight vector

When deriving PCA, we seek a vector \mathbf{w} (the weight vector or loading vector) such that the projection of the data onto this vector maximizes the variance. Maximizing the variance is equivalent to projecting the data onto a lower-dimensional linear subspace in such a way that the distance between a vector and its projection is not too large. For a given data matrix \mathbf{X} with mean zero (mean-centered data), the projection of the data onto \mathbf{w} is given by $\mathbf{X}\mathbf{w}$.

In other words, PCA aims to find the most accurate data representation in a lower dimensional space spanned by the maximum-variance directions.

The variance of the projected data can be expressed as:

$$\text{Var}(\mathbf{X}\mathbf{w}) = \frac{1}{n}(\mathbf{X}\mathbf{w})^T(\mathbf{X}\mathbf{w}) = \frac{1}{n}\mathbf{w}^T\mathbf{X}^T\mathbf{X}\mathbf{w}$$

Where n is the number of data points. The matrix $\mathbf{X}^T\mathbf{X}$ is the covariance matrix of the data (up to a scaling factor).

The goal is to maximize the variance $\mathbf{w}^T\mathbf{X}^T\mathbf{X}\mathbf{w}$ with respect to the weight vector \mathbf{w} , subject to the constraint that $\mathbf{w}^T\mathbf{w} = 1$ (to prevent the trivial solution where the variance could be made arbitrarily large just by scaling \mathbf{w}).

$$\mathcal{L} = \frac{1}{n}\mathbf{w}^T\mathbf{X}^T\mathbf{X}\mathbf{w} - \lambda(\mathbf{w}^T\mathbf{w} - 1)$$

$$\frac{\partial \mathcal{L}}{\partial \mathbf{w}} = \frac{2}{n}\mathbf{X}^T\mathbf{X}\mathbf{w} - 2\lambda\mathbf{w} = \mathbf{0}$$

$$\underbrace{\frac{1}{n}\mathbf{X}^T\mathbf{X}\mathbf{w}}_{:=\mathbf{S}} = \lambda\mathbf{w} \quad \Rightarrow \quad \mathbf{S}\mathbf{w} = \lambda\mathbf{w}$$

This is exactly the eigenvalue equation. The eigenvectors \mathbf{w} are the directions that maximize the variance, and the eigenvalues λ represent the magnitude of the variance along those directions.

- Eigenvectors: Each eigenvector of the covariance matrix represents a direction in the feature space. These directions are the principal components.
- Eigenvalues: The corresponding eigenvalue tells us how much variance is captured along that direction. The larger the eigenvalue, the more variance is captured by the corresponding eigenvector.

$$\text{Var}(\mathbf{X}\mathbf{w}) = \frac{1}{n}(\mathbf{X}\mathbf{w})^T(\mathbf{X}\mathbf{w}) = \frac{1}{n}\mathbf{w}^T\mathbf{X}^T\mathbf{X}\mathbf{w} = \frac{1}{n}\mathbf{w}^T\mathbf{S}\mathbf{w} = \frac{1}{n}\mathbf{w}^T\lambda\mathbf{w} = \frac{1}{n}\lambda$$

Capturing the Most Information:

- Variance is a measure of how spread out the data is along a particular direction. By maximizing the variance, we ensure that the principal components capture the most significant patterns in the data.
- If we reduce the dimensionality by selecting components with the highest variance, we retain the most information about the data, effectively compressing the data without losing critical details.
- High variance indicates that the data points are spread out and less likely to be redundant. Conversely, low variance implies that data points are clustered close together, often making the information less significant.

Sort eigenvectors v_1, v_2, \dots, v_d in descending order of their corresponding eigenvalues $\lambda_1 \geq \lambda_2 \geq \dots \geq \lambda_d$.

Then, select the first k eigenvectors to form a projection matrix:

$$W_k = [v_1, v_2, \dots, v_k] \in \mathbb{R}^{d \times k}$$

Now project the centered data onto this new basis:

$$Z = X_{\text{centered}}W_k \in \mathbb{R}^{n \times k}$$

This gives the data in the lower-dimensional space.

3.2 Linear Discriminant Analysis

We seek to obtain a transformation of X to Y through projecting the samples in X onto a hyperplane.

- Data matrix $\mathbf{X} = [\mathbf{x}_1, \dots, \mathbf{x}_N]^T \in \mathbb{R}^{N \times m}$ with row-vectors \mathbf{x}_i .

- Class labels $i \in \{1, \dots, C\}$, and each class has N_i samples.
- Each sample is belong to class ω_i .

Of all the possible lines we would like to select the one that maximizes the separability. In order to find a good projection vector, we need to define a measure of separation between the projections.

The mean vector of each class in \mathbf{x} and \mathbf{y} is given by

$$\begin{aligned}\mu_i &= \frac{1}{N} \sum_{\mathbf{x} \in \omega_i} \mathbf{x} \\ \tilde{\mu}_i &= \frac{1}{N} \sum_{\mathbf{y} \in \omega_i} \mathbf{y} = \frac{1}{N} \sum_{\mathbf{y} \in \omega_i} \mathbf{w}^T \mathbf{x} = \mathbf{w}^T \frac{1}{N} \sum_{\mathbf{y} \in \omega_i} \mathbf{x} = \mathbf{w}^T \mu_i\end{aligned}$$

This indicates that projecting leads to projecting the mean of \mathbf{x} to the mean of \mathbf{y} . Then, we could set the distance between the projected means as our objective function as follows:

$$J(\mathbf{w}) = |\tilde{\mu}_1 - \tilde{\mu}_2| = |\mathbf{w}^T(\mu_1 - \mu_2)|.$$

However, maximizing the distance between the projected means is not enough to yield better class separability. Thus, we will introduce the within-class variability, so-called *scatter matrix*:

$$\begin{aligned}S_i &= \sum_{\mathbf{x} \in \omega_i} (\mathbf{x} - \mu_i)(\mathbf{x} - \mu_i)^T \\ S_W &= S_1 + S_2\end{aligned}$$

Where S_i is the covariance matrix of class ω_i , and S_W is called the *within-class scatter matrix*.

Now the scatter of the projection \mathbf{y} can then be expressed as a function of the scatter matrix in feature space \mathbf{x} .

$$\begin{aligned}\tilde{s}_i^2 &= \sum_{y \in \omega_i} (y - \mu_i)^2 = \sum_{\mathbf{x} \in \omega_i} (w^T \mathbf{x} - w^T \mu_i)^2 \\ &= \sum_{\mathbf{x} \in \omega_i} (w^T(\mathbf{x} - \mu_i))^2 \\ &= \sum_{\mathbf{x} \in \omega_i} \underbrace{(w^T(\mathbf{x} - \mu_i))}_{\text{scalar}} \underbrace{(w^T(\mathbf{x} - \mu_i))}_{\text{scalar}^T=\text{scalar}} \\ &= \sum_{\mathbf{x} \in \omega_i} w^T(\mathbf{x} - \mu_i)(\mathbf{x} - \mu_i)^T w \\ &= w^T S_i w.\end{aligned}$$

Thus,

$$\tilde{s}_1^2 + \tilde{s}_2^2 = w^T S_1 w + w^T S_2 w = w^T (S_1 + S_2) w = w^T S_W w = \tilde{S}_W.$$

Similarly, the difference between the projected means can be expressed as follows:

$$\begin{aligned}(\tilde{\mu}_1 - \tilde{\mu}_2)^2 &= (w^T \mu_1 - w^T \mu_2) \\ &= w^T(\mu_1 - \mu_2)(\mu_1 - \mu_2)^T w \\ &= w^T S_B w = \tilde{S}_B\end{aligned}$$

This matrix is called the *between-class scatter*. Since S_B is the outer product of two vectors, its rank is at most one.

For a projection vector $\mathbf{w} \in \mathbb{R}^d$ (with $\mathbf{w} \neq \mathbf{0}$) define the *Fisher ratio*

$$J(\mathbf{w}) = \frac{|\tilde{\mu}_1 - \tilde{\mu}_2|^2}{\tilde{s}_1^2 + \tilde{s}_2^2} = \frac{\mathbf{w}^\top S_B \mathbf{w}}{\mathbf{w}^\top S_W \mathbf{w}}.$$

The optimal w satisfies that $|\tilde{\mu}_1 - \tilde{\mu}_2|$ (inter-class spread) is big and $\tilde{s}_1^2 + \tilde{s}_2^2$ (intra-class spread) is small. Note that this objective function is called *Rayleigh quotient*.

Maximising the objective J with respect to \mathbf{w} is equivalent to maximising the numerator while holding the denominator constant as both the numerator and denominator have the same order of $\|\mathbf{w}\|$ and thus J is invariant to scale changes in $\|\mathbf{w}\|$. Further, as we are only interested in the direction of \mathbf{w} , we can hold the denominator constant to 1. This results in a Lagrangian:

$$\mathcal{L}(\mathbf{w}, \lambda) = \mathbf{w}^\top S_B \mathbf{w} - \lambda(\mathbf{w}^\top S_W \mathbf{w} - 1).$$

- If you scale \mathbf{w} by any nonzero scalar c , then numerator: $(c\mathbf{w})^\top S_B(c\mathbf{w}) = c^2 \mathbf{w}^\top S_B \mathbf{w}$ denominator: $(c\mathbf{w})^\top S_W(c\mathbf{w}) = c^2 \mathbf{w}^\top S_W \mathbf{w}$. The factor c^2 cancels in the ratio, so $J(c\mathbf{w}) = J(\mathbf{w})$. Therefore, only the direction of \mathbf{w} matters, not its length. That's what invariant to scale changes in $\|\mathbf{w}\|$ means.
- Because the ratio is scale-invariant, we can arbitrarily fix the scale of \mathbf{w} . The common choice is: $\mathbf{w}^\top S_W \mathbf{w} - 1$.

Then, $\partial \mathcal{L} / \partial \mathbf{w} = \mathbf{0}$ gives

$$\begin{aligned} S_B \mathbf{w} &= \lambda S_W \mathbf{w}, \\ S_W^{-1} S_B \mathbf{w} &= \lambda \mathbf{w}. \end{aligned}$$

By using the definition of S_B ,

$$S_W^{-1} S_B \mathbf{w} = S_W^{-1} (\tilde{\mu}_1 - \tilde{\mu}_2) (\tilde{\mu}_1 - \tilde{\mu}_2)^T \mathbf{w} = \lambda \mathbf{w}.$$

Let's set $(\tilde{\mu}_1 - \tilde{\mu}_2)^T \mathbf{w} = \alpha$ and α is a scalar. Then, we get

$$S_W^{-1} (\tilde{\mu}_1 - \tilde{\mu}_2) = \frac{\lambda \mathbf{w}}{\alpha}$$

Since we are only interested in the direction of \mathbf{w} , we can discard the scalar term, then

$$\mathbf{w}^* = S_W^{-1} (\tilde{\mu}_1 - \tilde{\mu}_2).$$

Chapter 4

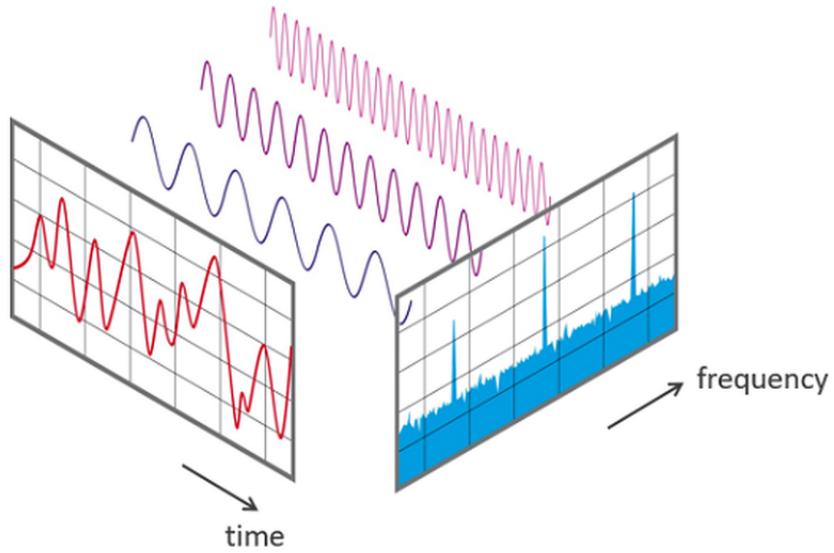
Fourier Analysis

Fourier analysis is a mathematical tool used to understand and analyze periodic phenomena. It is named after the French mathematician Joseph Fourier, it's based on the idea that any periodic function can be represented as a sum of sine and cosine functions with different frequencies, amplitudes, and phases.

The Fourier transform is one of the most important mathematical tools used for analyzing functions. Given an arbitrary function $f(x)$, with a real domain ($x \in \mathbb{R}$), we can express it as a linear combination of complex waves (\approx change of basis to a frequency domain).

- Periodic Functions: Fourier analysis deals with functions that repeat over regular intervals. This repetition could be over time (like sound waves or electrical signals) or space (like patterns on a surface).
- Fourier Series: It states that any periodic function can be represented as an infinite sum of sine and cosine functions. The Fourier series expansion expresses the original function as a weighted sum of sinusoids.
- Frequency Domain: While the original function is in the time or spatial domain, Fourier analysis allows us to examine it in the frequency domain. This means we can understand the different frequency components present in the signal and their respective strengths.
- Fast Fourier Transform (FFT): This is an efficient algorithm for calculating the Fourier transform of a discrete signal. It's widely used in digital signal processing due to its speed and accuracy.
- Continuous vs. Discrete Fourier Transform: While Fourier series deals with periodic continuous functions, Fourier transforms extend the concept to non-periodic functions or signals. The discrete Fourier transform (DFT) is used for discrete, sampled signals, as in digital signal processing.
- Inverse Fourier Transform: Just as you can transform a function into the frequency domain using Fourier analysis, you can also transform it back using the inverse Fourier transform. This allows you to reconstruct the original function from its frequency components.

Very intuitive explanation is [here!](#)



4.1 Fourier Series

We begin by discussing the Fourier series, which is used to analyze functions that are periodic in their inputs. A periodic function $f(x)$ is a function of a real variable x that repeats itself every time x changes by T . The constant T is called the *period*. We can write the periodicity condition as

$$f(x + T) = f(x), \forall x \in \mathbb{R}.$$

The value of $f(x)$ can be real or complex, but x should be real.

Let's consider what it means to specify a periodic function $f(x)$. One way to specify the function is to give an explicit mathematical formula for it. Another approach might be to specify the function values in $-T/2 \leq x < T/2$. Since there's an uncountably infinite number of points in this domain, we can generally only achieve an approximate specification of f this way, by giving the values of f at a large but finite set x points.

There is another interesting approach to specifying f . We can express it as a linear combination of simpler periodic functions, consisting of sines and cosines:

$$f(x) = \frac{a_0}{2} + \sum_{n=1}^{\infty} a_n \cos\left(\frac{2\pi nx}{T}\right) + \sum_{m=1}^{\infty} b_m \sin\left(\frac{2\pi mx}{T}\right), \quad (4.1)$$

where a_0 , a_n , and b_m are coefficients determined by integrating $f(x)$ over one period. The coefficients are given by

$$\begin{aligned} a_0 &= \frac{1}{T} \int_0^T f(x) dx \\ a_n &= \frac{2}{T} \int_0^T f(x) \cos\left(\frac{2\pi nx}{T}\right) dx \\ b_m &= \frac{2}{T} \int_0^T f(x) \sin\left(\frac{2\pi mx}{T}\right) dx \end{aligned}$$

The above formula is called a *Fourier series*. Given the numbers $\{a_n, b_m\}$, which are called the *Fourier coefficients*, $f(x)$ can be calculated for any x . The Fourier coefficients are real if $f(x)$ is a real function, or complex if $f(x)$ is complex.

Example: Consider a periodic function $f(x)$ defined on the interval $-\pi \leq x \leq \pi$ as follows:

$$f(x) = \begin{cases} 0 & \text{if } -\pi \leq x < 0 \\ x & \text{if } 0 \leq x \leq \pi \end{cases}$$

We want to find the Fourier series representation of $f(x)$.

Solution: Step 1: Determine the coefficients a_0 , a_n , and b_n .

1. Calculate a_0 :

$$\begin{aligned} a_0 &= \frac{1}{\pi} \int_{-\pi}^{\pi} f(x) dx \\ &= \frac{1}{\pi} \left(\int_{-\pi}^0 0 dx + \int_0^{\pi} x dx \right) \\ &= \frac{1}{\pi} \left(\int_0^{\pi} x dx \right) \\ &= \frac{1}{\pi} \left(\frac{1}{2}\pi^2 \right) \\ &= \frac{\pi}{2} \end{aligned}$$

2. Calculate a_n :

$$a_n = \frac{1}{\pi} \int_{-\pi}^{\pi} f(x) \cos(nx) dx$$

Since $f(x)$ is odd and $\cos(nx)$ is even, a_n will be zero for all n .

3. Calculate b_n :

$$\begin{aligned} b_n &= \frac{1}{\pi} \int_{-\pi}^{\pi} f(x) \sin(nx) dx \\ &= \frac{1}{\pi} \left(\int_0^{\pi} x \sin(nx) dx \right) \end{aligned}$$

This integral can be evaluated using integration by parts:

$$\begin{aligned} u &= x, \quad dv = \sin(nx) dx \\ du &= dx, \quad v = -\frac{1}{n} \cos(nx) \\ b_n &= \frac{1}{\pi} \left(\left[-\frac{x}{n} \cos(nx) \right]_0^\pi - \frac{1}{n} \int_0^\pi -\cos(nx) dx \right) \\ &= \frac{1}{\pi} \left(\frac{\pi}{n} - \frac{1}{n^2} [\sin(nx)]_0^\pi \right) \\ &= \frac{1}{\pi} \left(\frac{\pi}{n} \right) \\ &= \frac{1}{n} \end{aligned}$$

4. Write the Fourier series:

$$f(x) = \frac{\pi}{2} + \sum_{n=1}^{\infty} \frac{1}{n} \sin(nx)$$

4.1.1 Square-integrable functions

Can arbitrary periodic functions always be expressed as a Fourier series? It turns out that a certain class of periodic functions, commonly encountered in physical contexts, are guaranteed to always be expressible as Fourier series. These are called square-integrable functions such that the integral of the square of the absolute value is finite:

$$\int_{-a/2}^{a/2} |f(x)|^2 dx < \infty.$$

Unless otherwise stated, we will always assume that the functions we're dealing with are square-integrable.

4.1.2 Complex Fourier series and inverse relations

We have written the Fourier series as a sum of sine and cosine functions. However, sines and cosines can be expressed by exponential functions by using *Euler's formula*.

$$e^{ix} = \cos x + i \sin x \quad (4.2)$$

- $\cos x = \frac{e^{ix} + e^{-ix}}{2}$
- $\sin x = \frac{e^{ix} - e^{-ix}}{2i}$

Thus, Fourier series can be expressed as follows:

$$f(x) = \sum_{n=-\infty}^{\infty} c_n e^{inx},$$

where the complex Fourier coefficients c_n are given by:

$$c_n = \frac{1}{2\pi} \int_{-\pi}^{\pi} f(x) e^{-inx} dx$$

- i : complex number
- n : integer

Example: Consider the periodic function $f(x)$ defined on the interval $-\pi \leq x \leq \pi$:

$$f(x) = \begin{cases} 0 & \text{if } -\pi \leq x < 0 \\ x & \text{if } 0 \leq x \leq \pi \end{cases}$$

We will find the complex form of the Fourier series for this function.

Step 1: Calculate the complex Fourier coefficients c_n .

$$c_n = \frac{1}{2\pi} \int_{-\pi}^{\pi} f(x) e^{-inx} dx$$

Since $f(x)$ is piecewise, we split the integral into two parts:

$$c_n = \frac{1}{2\pi} \left(\int_{-\pi}^0 0 \cdot e^{-inx} dx + \int_0^\pi xe^{-inx} dx \right)$$

The first integral is zero, so we focus on the second part:

$$c_n = \frac{1}{2\pi} \int_0^\pi xe^{-inx} dx$$

This integral can be solved using integration by parts. Let:

$$u = x \quad \text{and} \quad dv = e^{-inx} dx$$

Then:

$$du = dx \quad \text{and} \quad v = \frac{e^{-inx}}{-in} = -\frac{1}{in}e^{-inx}$$

Applying integration by parts:

$$\int u dv = uv - \int v du$$

So,

$$\int_0^\pi xe^{-inx} dx = \left[-\frac{x}{in}e^{-inx} \right]_0^\pi + \frac{1}{in} \int_0^\pi e^{-inx} dx$$

Evaluating the boundary terms:

$$\left[-\frac{x}{in}e^{-inx} \right]_0^\pi = -\frac{\pi}{in}e^{-in\pi} - 0 = -\frac{\pi}{in}e^{-in\pi}$$

And for the integral:

$$\frac{1}{in} \int_0^\pi e^{-inx} dx = \frac{1}{in} \left[\frac{e^{-inx}}{-in} \right]_0^\pi = -\frac{1}{n^2} (e^{-in\pi} - 1)$$

Combining these results:

$$\int_0^\pi xe^{-inx} dx = -\frac{\pi}{in}e^{-in\pi} + \frac{1}{n^2}(1 - e^{-in\pi})$$

Thus, the Fourier coefficient c_n is:

$$c_n = \frac{1}{2\pi} \left(-\frac{\pi}{in} e^{-in\pi} + \frac{1}{n^2} (1 - e^{-in\pi}) \right)$$

$$c_n = \frac{1}{2\pi} \left(-\frac{\pi}{in} (-1)^n + \frac{1}{n^2} (1 - (-1)^n) \right)$$

This expression simplifies to:

$$c_n = \frac{i(-1)^n}{2n} + \frac{1}{2\pi n^2} (1 - (-1)^n)$$

Step 2: Write the Complex Fourier Series:

$$f(x) = \sum_{n=-\infty}^{\infty} c_n e^{inx}$$

With c_n now calculated, you can substitute these coefficients back into the Fourier series expression to represent $f(x)$.

4.2 Fourier Transform

4.2.1 Fourier Series v.s. Fourier Transform

The Fourier series is used to represent periodic functions as a sum of sine and cosine functions. On the other hand, the Fourier transform extends the idea to non-periodic functions or signals. The Fourier transform of a function $f(t)$ is given by:

$$F(\omega) = \int_{-\infty}^{\infty} f(t) e^{-i\omega t} dt,$$

where $F(\omega)$ is the frequency domain representation of the signal $f(t)$ and $\omega = 2\pi f$, where f is a frequency.

4.2.2 Equations of Fourier Transform

The Fourier transform and its inverse are defined as follows:

- Fourier Transform:

$$F(\omega) = \int_{-\infty}^{\infty} f(t) e^{-i\omega t} dt$$

- Inverse Fourier Transform:

$$f(t) = \frac{1}{2\pi} \int_{-\infty}^{\infty} F(\omega) e^{i\omega t} d\omega$$

These equations show how a time-domain signal $f(t)$ can be transformed into its frequency-domain representation $F(\omega)$, and vice versa.

Example: The Fourier transform of a rectangular pulse. Define a rectangular pulse function $f(t)$ as:

$$f(t) = \begin{cases} 1 & \text{if } |t| \leq \frac{1}{2} \\ 0 & \text{otherwise} \end{cases}$$

To find its Fourier transform $F(\omega)$:

$$F(\omega) = \int_{-\infty}^{\infty} f(t)e^{-i\omega t} dt$$

Since $f(t)$ is non-zero only between $-\frac{1}{2}$ and $\frac{1}{2}$, the integral becomes:

$$F(\omega) = \int_{-\frac{1}{2}}^{\frac{1}{2}} e^{-i\omega t} dt$$

This integral can be evaluated as:

$$F(\omega) = \left[\frac{e^{-i\omega t}}{-i\omega} \right]_{-\frac{1}{2}}^{\frac{1}{2}} = \frac{e^{-i\omega/2} - e^{i\omega/2}}{-i\omega} = \frac{2 \sin(\omega/2)}{\omega}$$

Thus, the Fourier transform of the rectangular pulse is:

$$F(\omega) = \frac{\sin(\omega/2)}{\omega/2} = \text{sinc}\left(\frac{\omega}{2}\right)$$

where $\text{sinc}(x) = \frac{\sin(x)}{x}$.

4.3 Fourier Transform as a Change of Basis

A change of basis in linear algebra is a fundamental concept that involves transitioning from one coordinate system to another within a vector space. This can simplify calculations, provide deeper insights, or align with different geometric interpretations. Let's break down the concept with definitions and steps:

Fourier analysis is closely related to the concept of change of basis in linear algebra. In Fourier analysis, we are essentially changing the basis from the standard basis of functions (e.g., the time domain for signals) to a basis of sinusoidal functions (e.g., the frequency domain). Here's how this relationship works:

The Fourier transform (both continuous and discrete versions) can be seen as a change of basis from the time domain to the frequency domain. When we apply the Fourier transform to a signal $x(t)x(t)$, we express it as a linear combination of sinusoidal basis functions.

- Standard Basis in Time Domain: In the time domain, a signal or function can be represented as a sum of basis functions, which are typically δ functions (impulse functions) or other

simple functions like polynomial terms. For example, a discrete-time signal $x[n]$ can be considered in terms of standard basis vectors.

- Fourier Basis in Frequency Domain: In the frequency domain, the basis functions are complex exponentials or sinusoidal functions $e^{i\omega t}$ (where i is the imaginary unit and ω is the angular frequency). The Fourier transform decomposes a signal into a sum of these sinusoidal basis functions.

The Fourier transform (both continuous and discrete versions) can be seen as a change of basis from the time domain to the frequency domain. When we apply the Fourier transform to a signal $x(t)$, we express it as a linear combination of sinusoidal basis functions.

Continuous Fourier Transform: The continuous Fourier transform of a function $x(t)$ is given by:

$$X(f) = \int_{-\infty}^{\infty} x(t)e^{-i2\pi ft} dt$$

Here, $X(f)$ represents the coordinates of the function $x(t)$ in the new basis (frequency domain). The inverse Fourier transform reconstructs the original signal from its frequency domain representation:

$$x(t) = \int_{-\infty}^{\infty} X(f)e^{i2\pi ft} df$$

This process is analogous to converting coordinates from the new basis back to the original basis.

4.3.1 Change of Basis

- Vector Space: A collection of vectors where vector addition and scalar multiplication are defined and satisfy specific axioms.
- Basis: A set of linearly independent vectors that span the entire vector space. Every vector in the space can be uniquely represented as a linear combination of the basis vectors.
- Coordinate Vector: The representation of a vector in terms of the basis vectors, typically given as a column of coefficients.

Steps for Change of Basis:

1. Representing Vectors in Original Basis: Given a vector space V with a basis $B = \{\mathbf{b}_1, \mathbf{b}_2, \dots, \mathbf{b}_n\}$, any vector \mathbf{v} in V can be written as:

$$\mathbf{v} = c_1 \mathbf{b}_1 + c_2 \mathbf{b}_2 + \cdots + c_n \mathbf{b}_n$$

Here, c_1, c_2, \dots, c_n are the coordinates of \mathbf{v} relative to basis B , often written as $[\mathbf{v}]_B$.

2. Defining the New Basis: Suppose we want to change to a new basis $B' = \{\mathbf{b}'_1, \mathbf{b}'_2, \dots, \mathbf{b}'_n\}$. We need to express the original basis vectors in terms of the new basis vectors. Let:

$$\mathbf{b}_i = a_{i1} \mathbf{b}'_1 + a_{i2} \mathbf{b}'_2 + \cdots + a_{in} \mathbf{b}'_n$$

for $i = 1, 2, \dots, n$.

3. Constructing the Change of Basis Matrix: The coefficients a_{ij} form the columns of the change of basis matrix P from B to B' :

$$P = \begin{bmatrix} a_{11} & a_{12} & \cdots & a_{1n} \\ a_{21} & a_{22} & \cdots & a_{2n} \\ \vdots & \vdots & \ddots & \vdots \\ a_{n1} & a_{n2} & \cdots & a_{nn} \end{bmatrix}$$

4. Converting Coordinates: To convert the coordinates of a vector \mathbf{v} from the original basis B to the new basis B' , we use the change of basis matrix P :

$$[\mathbf{v}]_{B'} = P^{-1}[\mathbf{v}]_B$$

where P^{-1} is the inverse of the change of basis matrix and $[\mathbf{v}]_B = (c_1, \dots, c_n)$ is a coordinate vector.

Conversely, to convert coordinates from the new basis B' back to the original basis B , we use:

$$[\mathbf{v}]_B = P[\mathbf{v}]_{B'}$$

Example: Consider a vector space \mathbb{R}^2 with an original basis $B = \{\mathbf{b}_1, \mathbf{b}_2\}$ and a new basis $B' = \{\mathbf{b}'_1, \mathbf{b}'_2\}$. Let:

$$\mathbf{b}_1 = \begin{bmatrix} 1 \\ 0 \end{bmatrix}, \quad \mathbf{b}_2 = \begin{bmatrix} 0 \\ 1 \end{bmatrix}$$

and

$$\mathbf{b}'_1 = \begin{bmatrix} 1 \\ 1 \end{bmatrix}, \quad \mathbf{b}'_2 = \begin{bmatrix} 1 \\ -1 \end{bmatrix}$$

To find the change of basis matrix P from B to B' , express the original basis vectors in terms of B' :

$$\begin{aligned} \mathbf{b}_1 &= \frac{1}{2}\mathbf{b}'_1 + \frac{1}{2}\mathbf{b}'_2 \\ \mathbf{b}_2 &= \frac{1}{2}\mathbf{b}'_1 - \frac{1}{2}\mathbf{b}'_2 \end{aligned}$$

So, the change of basis matrix P is:

$$P = \begin{bmatrix} \frac{1}{2} & \frac{1}{2} \\ \frac{1}{2} & -\frac{1}{2} \end{bmatrix}$$

To convert a vector \mathbf{v} with coordinates $[\mathbf{v}]_B = \begin{bmatrix} x \\ y \end{bmatrix}$ to the new basis B' , we compute:

$$[\mathbf{v}]_{B'} = P^{-1}[\mathbf{v}]_B$$

And to revert to the original basis:

$$[\mathbf{v}]_B = P[\mathbf{v}]_{B'}$$

Chapter 5

Training, Testing, and Regularization

5.1 Sources of Error in ML

$$E_{out} \leq E_{ml} + \Omega$$

- E_{out} : estimation of error.
- E_{ml} : error from a learning algorithm
- Ω : error caused by the variance from observations.

We also define

- f : target function
- g : learning function
- $g^{(D)}$: learned function based on D , or simply *hypothesis*.
- D : dataset drawn from the real world.
- \bar{g} : the average hypothesis of a given infinite number of D s.

$$\bar{g}(x) = \mathbb{E}_D[g^{(D)}(x)].$$

Error of a single instance x from g learnt from D is given by

$$Err_{out}(g^{(D)}(x)) = \mathbb{E}_X[(g^{(D)}(x) - f(x))^2],$$

where X can be considered as test sets. Then, the expected error over the infinite number of datasets D sampled from a true data distribution is

$$\begin{aligned}\mathbb{E}_D[Err_{out}(g^{(D)}(x))] &= \mathbb{E}_D[\mathbb{E}_X[(g^{(D)}(x) - f(x))^2]] \\ &= \mathbb{E}_X[\mathbb{E}_D[(g^{(D)}(x) - f(x))^2]]\end{aligned}$$

Let's simplify the term inside with an average of hypothesis $\bar{g}(x)$:

$$\begin{aligned}\mathbb{E}_D[(g^{(D)}(x) - f(x))^2] &= \mathbb{E}_D[(g^{(D)}(x) - \bar{g}(x) + \bar{g}(x) - f(x))^2] \\ &= \mathbb{E}_D[(g^{(D)}(x) - \bar{g}(x))^2 + (\bar{g}(x) - f(x))^2 \\ &\quad + 2(g^{(D)}(x) - \bar{g}(x))(\bar{g}(x) - f(x))] \\ &= \mathbb{E}_D[(g^{(D)}(x) - \bar{g}(x))^2] + (\bar{g}(x) - f(x))^2 \\ &\quad + \mathbb{E}_D[2(g^{(D)}(x) - \bar{g}(x))(\bar{g}(x) - f(x))]\end{aligned}$$

Since, $\mathbb{E}_D[2(g^{(D)}(x) - \bar{g}(x))(\bar{g}(x) - f(x))]$ is 0, the expectation of the error becomes

$$\mathbb{E}_D[\text{Err}_{\text{out}}(g^{(D)}(x))] = \mathbb{E}_X[\mathbb{E}_D[(g^{(D)}(x) - \bar{g}(x))^2] + (\bar{g}(x) - f(x))^2].$$

Let's closely look at this formula. The errors are from two sources:

- **Variance:** $\mathbb{E}_D[(g^{(D)}(x) - \bar{g}(x))^2]$. Variance captures how much your classifier changes if you train on a different training set. We need to collect more data to reduce the variance.
- **Bias:** $(\bar{g}(x) - f(x))^2$. Bias is the inherent error that you obtain from your classifier even with infinite training data. We need to build a more complex model to reduce the bias.

However, if we reduce the bias, then the variance tends to increase.

5.1.1 Alternative Derivation

The derivation of the bias–variance decomposition for squared error proceeds as follows.[6][7] For notational convenience, abbreviate $f = f(x)$ and $\hat{f} = \hat{f}(x)$. First, recall that, by definition, for any random variable \mathbf{X} , we have

$$\text{Var}[\hat{f}(x)] = \mathbb{E}[X^2] - \mathbb{E}[X]^2.$$

By rearranging, we get

$$\mathbb{E}[X^2] = \text{Var}[\hat{f}(x)] + \mathbb{E}[X]^2.$$

Since f is deterministic

$$\mathbb{E}[f] = f$$

Thus, given $y = f + \varepsilon$ and $\mathbb{E}[\varepsilon] = 0$, implies $\mathbb{E}[y] = \mathbb{E}[f + \varepsilon] = \mathbb{E}[f] = f$

Also, since $\text{Var}[\varepsilon] = \sigma^2$

$$\text{Var}[y] = \mathbb{E}[(y - \mathbb{E}[y])^2] = \mathbb{E}[(y - f)^2] = \mathbb{E}[(f + \varepsilon - f)^2] = \mathbb{E}[\varepsilon^2] = \text{Var}[\varepsilon] + (\mathbb{E}[\varepsilon])^2 = \sigma^2$$

Thus, since ε and \hat{f} are independent, we can write:

$$\begin{aligned}
 E[(y - \hat{f})^2] &= E[(f + \varepsilon - \hat{f})^2] \\
 &= E[(f + \varepsilon - \hat{f} + E[\hat{f}] - E[\hat{f}])^2] \\
 &= E[(f - E[\hat{f}])^2] + E[\varepsilon^2] + E[(E[\hat{f}] - \hat{f})^2] + 2E[(f - E[\hat{f}])\varepsilon] + \\
 &\quad 2E[\varepsilon(E[\hat{f}] - \hat{f})] + 2E[(E[\hat{f}] - \hat{f})(f - E[\hat{f}])] \\
 &= (f - E[\hat{f}])^2 + E[\varepsilon^2] + E[(E[\hat{f}] - \hat{f})^2] + \\
 &\quad 2(f - E[\hat{f}])E[\varepsilon] + 2E[\varepsilon]E[(E[\hat{f}] - \hat{f})] + 2E[(E[\hat{f}] - \hat{f})(f - E[\hat{f}])] \\
 &= (f - E[\hat{f}])^2 + E[\varepsilon^2] + E[(E[\hat{f}] - \hat{f})^2] \\
 &= (f - E[\hat{f}])^2 + \text{Var}[y] + \text{Var}[\hat{f}] \\
 &= \text{Bias}[\hat{f}]^2 + \text{Var}[y] + \text{Var}[\hat{f}] \\
 &= \text{Bias}[\hat{f}]^2 + \sigma^2 + \text{Var}[\hat{f}]
 \end{aligned}$$

Chapter 6

Optimization

6.1 Intuition of Gradient

Gradient descent is an optimization algorithm used to minimize a function by iteratively moving towards the function's minimum value. It is a fundamental concept in machine learning, particularly in training models such as neural networks. The gradient is a vector that represents the direction of the steepest increase of the function at a given point. For example, for a convex function $z = ax^2 + by^2$, the gradient is $[2ax, 2by]$, which points in the direction of the steepest ascent.

In gradient descent, the goal is to minimize the function, so the algorithm moves in the opposite direction of the gradient, which is $[-2ax, -2by]$. This opposite direction is chosen because it is the direction of the steepest decrease in the function value. But how do we know that moving in this direction will strictly decrease the function value?

6.1.1 Direction of Gradient Descent

Let's investigate the direction of gradient descent.

- The derivative of the objective function $f(\mathbf{x})$ provides the slope of $f(\mathbf{x})$ at the point $f(\mathbf{x})$.
- It tells us how to change \mathbf{x} in order to make a small improvement in our goal.

A function $f(\mathbf{x})$ can be approximated by its first-order Taylor expansion at $\bar{\mathbf{x}}$:

$$f(\mathbf{x}) \approx f(\bar{\mathbf{x}}) + \nabla f(\bar{\mathbf{x}})^T (\mathbf{x} - \bar{\mathbf{x}})$$

Now let $\mathbf{d} \neq 0$, $\|\mathbf{d}\| = 1$ be a direction, and in consideration of a new point $\mathbf{x} := \bar{\mathbf{x}} + \mathbf{d}$, we define:

$$f(\bar{\mathbf{x}} + \mathbf{d}) \approx f(\bar{\mathbf{x}}) + \nabla f(\bar{\mathbf{x}})^T \mathbf{d}$$

We would like to choose \mathbf{d} that minimizes the function f . From the Cauchy-Schwarz inequality ¹, we know that

$$|\nabla f(\bar{\mathbf{x}})^T \mathbf{d}| \leq \|\nabla f(\bar{\mathbf{x}})\| \|\mathbf{d}\|.$$

¹Cauchy-Schwarz Inequality: $|\mathbf{a} \cdot \mathbf{b}| \leq \|\mathbf{a}\| \|\mathbf{b}\|$. Equality holds if and only if either \mathbf{a} or \mathbf{b} is a multiple of the other.

The equality holds if and only if $\mathbf{d} = \lambda \nabla f(\bar{\mathbf{x}})$, where $\lambda \in \mathbb{R}$. Since we want to minimize the function f , we negate the steepest direction \mathbf{d}^* , then

$$f(\bar{\mathbf{x}} + \mathbf{d}) \approx f(\bar{\mathbf{x}}) - \lambda \nabla f(\bar{\mathbf{x}})^T \nabla f(\bar{\mathbf{x}}).$$

Since $\nabla f(\bar{\mathbf{x}})^T \nabla f(\bar{\mathbf{x}})$ is **always positive**, the term $-\lambda \nabla f(\bar{\mathbf{x}})^T \nabla f(\bar{\mathbf{x}})$ is always negative. Therefore by updating \mathbf{x}

$$\mathbf{x}^{(k+1)} = \mathbf{x}^{(k)} - \eta \nabla f(\mathbf{x}^{(k)}),$$

we get

$$f(\mathbf{x}^{(k+1)}) < f(\mathbf{x}^{(k)}).$$

6.2 Normalized Gradient Descent

The underlying issue of the vanila gradient descent is the presence of saddle points in nonconvex functions; the gradient $\nabla f(x)$ vanishes near saddle points, which causes GD to “stall” in neighboring regions. This both slows the overall convergence rate and makes detection of local minima difficult. The detrimental effects of this issue become particularly severe in high-dimensional problems where the number of saddle points may proliferate.

However, in the normalized gradient descent

$$\frac{\nabla f(x)}{\|\nabla f(x)\|}$$

The normalized gradient preserves the direction of the gradient but ignores magnitude, because the normalization does not vanish near saddle points, the intuitive expectation is that NGD should not slow down in the neighborhood of saddle points and should therefore escape quickly.

6.3 Projected Gradient Descent

Gradient Descent (GD) is a standard way to solve unconstrained optimization problem. Starting from an initial point $x \in \mathbb{R}^n$, GD itereates until a stopping criterion is met. Projected Gradient Descent (PGD) is a way to solve constrained optimization problem. Consider a constraint set \mathcal{Q} , starting from a initial point $x_0 \in \mathcal{Q}$, PGD iterates the following equation until a stopping condition is met:

$$x_{k+1} = P_{\mathcal{Q}}(x_k - t_k \nabla f(x_k)),$$

where $P_{\mathcal{Q}}$ is the projection operator

$$P_{\mathcal{Q}}(x_0) = \operatorname{argmin}_{x \in \mathcal{Q}} \frac{1}{2} \|x - x_0\|_2^2$$

In other words, given a point x_0 , $P_{\mathcal{Q}}$ tries to to find a point $x \in \mathcal{Q}$ which is “closest”to x_0 .

Note that a vector projection can be expressed as follows:

$$a_1 = \|\mathbf{a}\| \cos \theta = \mathbf{a} \cdot \hat{\mathbf{b}} = \mathbf{a} \cdot \frac{\mathbf{b}}{\|\mathbf{b}\|}$$

Thus, a projection for unit L_2 ball is given by the solution of the equation as follows:

$$\mathbf{x} = \mathcal{P}_{\|\mathbf{x}\|_2 \leq 1}(\mathbf{y})$$

The solution is

$$\mathbf{x} = \frac{\mathbf{y}}{\max\{1, \|\mathbf{y}\|_2\}}$$

The “geometric” proof is given as follows: Let $\mathcal{S} = \{\mathbf{x} \in \mathbb{R}^n : \|\mathbf{x}\|_2 \leq 1\}$.

- If $\mathbf{y} \in \mathcal{S}$, then $\|\mathbf{y}\|_2 \leq 1$ and \mathbf{y} itself is the closest point to \mathbf{y} .
- If $\mathbf{y} \notin \mathcal{S}$, then $\|\mathbf{y}\|_2 > 1$ and the closest point $\mathbf{x} \in \mathcal{S}$ to \mathbf{y} will be simply $\frac{\mathbf{y}}{\|\mathbf{y}\|_2}$ as the norm of $\frac{\mathbf{y}}{\|\mathbf{y}\|_2} = 1$.

By combining the best cases, we have

$$\mathbf{x} = \frac{\mathbf{y}}{\max\{1, \|\mathbf{y}\|_2\}}$$

6.4 Exponentially Weighted Average

$$v_t = \beta v_{t-1} + (1 - \beta) \theta_t$$

Larger β value covers more longer history. EMA is exponentially weighted average the previous result.

6.5 Bias Correction

The initial values of v_t will be very low which need to be compensated, since the curve starts from 0, there are not many values to average on in the initial points. Thus, the curve is lower than the correct value initially and then moves in line with expected values. The β is the same as the averaging coefficient. As t becomes large, the impact of the bias correction will be decreased.

$$v_t = \frac{v_t}{1 - \beta^t}$$

6.6 Momentum

Momentum can reduce the oscillation in the gradients. Let’s say w has a small value and b is in charge of oscillation. Then momentum can cancel out db by averaging them.

$$v_{dw} = \beta v_{dw} + (1 - \beta) dw$$

$$v_{db} = \beta v_{db} + (1 - \beta) db$$

$$w = w - \alpha v_{dw}$$

$$b = b - \alpha v_{db}$$

6.7 Adagrad: Adaptive Gradient

$$\begin{aligned} v_{dw} &= v_{dw} + dw \cdot dw \\ w &= w - \frac{\alpha}{\sqrt{v_{dw}} + \epsilon} v_{dw} \end{aligned}$$

A con of Adagrad is learning rate will become very small

6.8 RMS Prop

$$\begin{aligned} s_{dw} &= \beta s_{dw} + (1 - \beta) dw^2 \\ s_{db} &= \beta s_{db} + (1 - \beta) db^2 \\ w &= w - \alpha \frac{dw}{\sqrt{s_{dw}}} \\ b &= b - \alpha \frac{db}{\sqrt{s_{db}}} \end{aligned}$$

6.9 ADAM

Its name is derived from adaptive moment estimation, and the reason it's called that is because Adam uses estimations of first and second moments of gradient to adapt the learning rate for each weight of the neural network. N -th moment of a random variable is defined as the expected value of that variable to the power of n . More formally:

$$m_n = \mathbb{E}[X^n]$$

To estimates the moments, Adam utilizes exponentially moving averages, computed on the gradient evaluated on a current mini-batch:

Since m and v are estimates of first and second moments, we want to have the following property:

$$\mathbb{E}[m_t] = \mathbb{E}[g_t] \tag{6.1}$$

$$\mathbb{E}[v_t] = \mathbb{E}[g_t^2] \tag{6.2}$$

Unbiased estimators

ADAM uses both momentum style and RMS prop style averaging.

- $v_{dw} = \beta v_{dw} + (1 - \beta) dw$
- $v_{db} = \beta v_{db} + (1 - \beta) db$
- $s_{dw} = \beta s_{dw} + (1 - \beta) dw^2$

- $s_{db} = \beta s_{db} + (1 - \beta)db^2$

Using them,

$$\bullet \quad v_{dw}^{\text{corr}} = \frac{v_{dw}}{1 - \beta_1^t}$$

$$\bullet \quad v_{db}^{\text{corr}} = \frac{v_{db}}{1 - \beta_1^t}$$

$$\bullet \quad s_{dw}^{\text{corr}} = \frac{s_{dw}}{1 - \beta_2^t}$$

$$\bullet \quad s_{db}^{\text{corr}} = \frac{s_{db}}{1 - \beta_2^t}$$

Finally,

$$w = w - \alpha \frac{v_{dw}^{\text{corr}}}{\sqrt{s_{dw}^{\text{corr}}} + \varepsilon}$$

$$b = b - \alpha \frac{v_{db}^{\text{corr}}}{\sqrt{s_{db}^{\text{corr}}} + \varepsilon}$$

Chapter 7

Trees

A classification tree is a flowchart that helps a computer make decisions. It asks a series of yes/no questions about your data like “Is age ≤ 30 ?” and follows the answers down branches until it reaches a final box (a *leaf*). That leaf says which class the item belongs to, for example spam or not spam.

Trees are popular because they’re easy to read and explain. You can show the rules to anyone on your team and they’ll understand why a prediction was made. Trees also work well with mixed data (numbers and categories) and need very little data cleaning.

At each step, the tree tries many possible questions and picks the one that best separates the classes. We measure *how mixed* a node is using a number called *impurity* (common choices are *Gini* or *entropy*). A good question makes child nodes less mixed (ideally, each child is mostly one class).

Once a new example flows through the questions to a leaf, the tree predicts the majority class in that leaf. It can also report simple probabilities, like Class A: 80%, Class B: 20%, based on how the training examples in that leaf were labeled.

Deep trees can memorize the training data and make mistakes on new data (*overfitting*). You can control this by limiting the depth, requiring a minimum number of samples per split/leaf, or pruning back weak branches. When you need more accuracy and stability, ensembles like *Random Forests* or *Gradient Boosted Trees* build on the same idea but combine many trees.

Decision trees are more prone to overfitting than many of the models, as their learning algorithms can produce large, complicated decision trees that perfectly model every training instance but fail to generalize the real relationship. Several techniques can mitigate over-fitting in decision trees. *Pruning* is a common strategy that removes some of the tallest nodes and leaves of a decision tree.

7.1 Classification Tree

Let (p_k) be the fraction of samples of class (k) in the node. **Gini impurity**:

$$\text{Gini} = 1 - \sum_k p_k^2$$

- Range: $[0, 0.5]$ (0) (pure) to 0.5 ($p = 0.5$) for a binary class problem.

Entropy (information entropy)

$$H = - \sum_k p_k \log_2 p_k$$

- Range: (0) (pure) to $(\log_2 K)$ (maximally mixed; for binary, max = 1 bit at ($p = 0.5$)).

In both cases, the node is pure when it has 0 impurity and maximally impure when classes are evenly mixed.

Example

A node with 80% class A, 20% class B: ($p = [0.8, 0.2]$)

- Gini: $1 - (0.8^2 + 0.2^2) = 1 - (0.64 + 0.04) = 0.32$
- Entropy: $-[0.8 \log_2 0.8 + 0.2 \log_2 0.2] \approx 0.72$ bits

A node with 60% / 40%: ($p = [0.6, 0.4]$)

- Gini: $1 - (0.36 + 0.16) = 0.48$
- Entropy ≈ 0.97 bits

How a split is scored (impurity decrease):

Trees choose the split that **reduces impurity** the most:

$$\text{Gain} = \text{Impurity}(\text{parent}) - \left(\frac{n_L}{n} \text{Impurity}(L) + \frac{n_R}{n} \text{Impurity}(R) \right)$$

Example

Let's say a node splits samples as follows:

- 100 samples, classes A/B = 60/40 (parent ($p = [0.6, 0.4]$)))
- Candidate split: Left: 50 samples (40/10), Right: 50 samples (20/30)

The impurity of the node is given by

- Gini: 0.48,
- Entropy: ≈ 0.97

Then, we can compute the total Gini impurity, which is a weighted average of Gini impurities for the leaves,

- Left (40/10, ($p = [0.8, 0.2]$)):
 - Gini: 0.32,
 - Entropy: ≈ 0.72
- Right (20/30, ($p = [0.4, 0.6]$)):

- Gini: 0.48
- Entropy ≈ 0.97

Then,

- Gini: $0.5 \times 0.32 + 0.5 \times 0.48 = 0.40.$
 - Gain: $0.48 - 0.40 = 0.08$
- Entropy: $0.5 \times 0.72 + 0.5 \times 0.97 = 0.8465$
 - Gain $\approx 0.97 - 0.8465 = 0.1245$ bits

This split improves purity; you'd compare it with other candidate splits and pick the best gain.

Chapter 8

Ensemble Learning

8.1 Bagging

Bagging (bootstrap aggregating) is a simple way to make a noisy model more stable. Imagine you have a wobbly predictor-like a deep decision tree—that can change its mind a lot if the training data changes a little. Instead of trusting a single tree, bagging builds many versions of that model on slightly different views of the data, and then averages their answers (or takes a majority vote for classification). Averaging smooths out the noise, so the final prediction is steadier and usually more accurate.

How do we get those slightly different views? That’s where bootstrapping comes in. From the original training set, we repeatedly draw new datasets of the same size with replacement (so the same example can appear multiple times, and some examples are left out). Each bootstrapped dataset trains its own model. Because each model sees a different sample, their mistakes won’t line up perfectly—averaging then cancels out a lot of the randomness.

Why does this help? Many models, especially high-variance ones like decision trees, can overfit the quirks of a single dataset. **Bagging reduces variance by blending many perspectives, without increasing bias too much.** In practice, this often yields a solid accuracy boost and better reliability. A nice side effect is you get out-of-bag estimates: the examples left out of each bootstrap can be used as a built-in validation set to measure performance without a separate hold-out.

Bagging isn’t magic—if the base model is already very stable (low variance) or consistently biased in the wrong direction, averaging won’t fix the core problem. It also costs more compute and can feel less interpretable since you’re dealing with many models instead of one. But as a foundational ensemble idea, bagging is both elegant and powerful—and it’s the backbone of methods like Random Forests, which add an extra twist by randomly selecting features, too.

8.1.1 Intuition

We start with a single training set ($D = (x_i, y_i)_{i=1}^n$). Instead of hoping for many independent datasets, we **simulate** them by drawing **bootstrap samples** (D_1, \dots, D_M): each (D_m) is formed by sampling (n) examples **with replacement** from (D). For each (D_m), we fit a base learner (h_m) (e.g., a decision tree, ridge regression, logistic regression). We then **aggregate** the

models:

$$\bar{h}(x) = \frac{1}{M} \sum_{m=1}^M h_m(x).$$

If we could draw truly independent datasets $(D^{(1)}, \dots, D^{(M)})$ from the population and train $(h^{D(m)})$ on each, then the **empirical average hypothesis** would converge to the **theoretical average hypothesis** ($\mathbb{E}_D[h^D(x)]$) as $(M \rightarrow \infty)$. With bagging, our bootstraps come from the **empirical distribution** of (D) , not the population, so they're not independent.

- Each bootstrap sample is just a reshuffled, reweighted version of the same original points.
- So two bootstrap samples will overlap a lot (they reuse the same observations), unlike two brand-new datasets collected from the real world.
- In short: they're new mixes of the same ingredients, not new ingredients.

Still, two facts make bagging effective:

1. **Variance reduction by averaging:** Training many models on these different mixes and averaging their predictions smooths out noise (reduces variance), especially for models that change a lot when the data changes (like deep trees).
2. **Bootstrap as a population proxy:** Drawing with replacement from (D) mimics sampling from the underlying data-generating process via the empirical distribution. This is not perfect independence, but for many learners it's good enough to bring (\bar{h}) close to $(\mathbb{E}_D[h^D(x)])$ in practice, especially for unstable learners (small data perturbations \rightarrow large model changes), like deep decision trees.

8.1.2 Bias–variance intuition

- Variance: Bagging typically decreases variance because averaging stabilizes predictions.
- Bias: Bagging usually does not increase bias, and may slightly reduce it for some learners. But its main win is variance.
- Who benefits most? Unstable learners (*e.g.*, trees, k-NN with small (k)) benefit a lot. Stable learners (*e.g.*, linear/ridge regression) benefit less, though averaging can still help in noisy regimes.

There are some costs:

- Interpretability
- Computational costs

When and why it works (and when it doesn't)

- Correlation matters. If base learners are too correlated, variance reduction is limited. This is why Random Forests add feature subsampling at each split to further decorrelate trees, improving over plain bagging of trees.
- Computational cost. Training (M) models costs $(M \times)$ as much compute (though it parallelizes nicely).
- Interpretability. An averaged ensemble is harder to interpret than a single model (especially vs. a small tree or linear model). You can mitigate with feature importance, partial dependence, SHAP, and so on., but it's still less transparent.
- Diminishing returns in (M) . Gains flatten as (M) grows because you approach the correlation floor. In practice, tens to a few hundreds of models are often sufficient.

8.2 Random Forest

Now, what exactly makes a random forest different from a bagged forest? The answer to this question has to do with the way in which single trees are grown in a random forests. We are going to introduce a twist in the tree-building algorithm. More specifically, we are going to modify the node splitting procedure. Assuming we have p features, we predetermine a number $k \ll p$ such that at each node, k features are randomly selected (from the set of p predictors), to find the best split.

The fundamental reason to introduce another random sampling operation, which is repeated for every node partition, is to attempt reducing the number of correlated predictors that could potentially be present in the construction of a single tree. This seemingly awkward operation turns out to return high dividends when a forest is deployed in practice.

8.3 Boosting

Boosting builds a strong model by adding many small, simple models (weak learners) sequentially. **Each new model focuses on the mistakes of the models before it.** In the end, we add their predictions to get one strong predictor.

- Bagging = parallel averaging to reduce **variance**.
- Boosting = sequential fixing to reduce **bias** (and often variance too).

Types of Boosting Algorithms:

- AdaBoost (Adaptive Boosting)
- Gradient Boosting
- XGBoost

8.3.1 AdaBoost

Boosting is a general strategy for learning classifiers by combining simpler ones. The idea of boosting is to take a “weak classifier” - that is, any classifier that will do at least slightly better than chance — and use it to build a much better classifier, thereby boosting the performance of the weak classification algorithm. This boosting is done by averaging the outputs of a collection of weak classifiers.

The most popular boosting algorithm is *AdaBoost*, so-called because it is “adaptive.” AdaBoost is extremely simple to use and implement (far simpler than SVMs), and often gives very effective results. There is tremendous flexibility in the choice of weak classifier as well. Boosting is a specific example of a general class of learning algorithms called ensemble methods, which attempt to build better learning algorithms by combining multiple simpler algorithms.

Suppose we are given training data $\{(\mathbf{x}_i, y_i)\}_{i=1}^N$, where $\mathbf{x}_i \in \mathbb{R}^K$ and $y_i \in \{-1, 1\}$. And suppose we are given a (potentially large) number of weak classifiers, denoted $f_m(\mathbf{x}) \in \{-1, 1\}$, and a **0-1** loss function I , defined as

$$I(f_m(\mathbf{x}), y) = \begin{cases} 0 & \text{if } f_m(\mathbf{x}_i) = y_i \\ 1 & \text{if } f_m(\mathbf{x}_i) \neq y_i \end{cases}$$

- $I = 0$ if prediction equals the truth ($\hat{y} = y$)
- $I = 1$ if prediction is wrong ($\hat{y} \neq y$)

AdaBoost aims to build a strong classifier by combining many weak ones. It does this in rounds $m = 1, 2, \dots, M$:

1. Give every training example the same weight. Every sample is equally important at the beginning.
2. Fit a *weak learner* f_m on the weighted data
3. Compute how good it is: measure its *weighted error* of weak learner, e_m .

$$e_m = \sum_{i=1}^N w_i^m \cdot I(f_m(\mathbf{x}_i), y_i),$$

where w_i^m is the current example weight.

- This measures how many (weighted) mistakes that stump/tree makes this round.
4. Compute α_m , which is the *vote weight* (strength) assigned to that weak learner when forming the final ensemble:

$$\alpha_m = \frac{1}{2} \ln \left(\frac{1 - e_m}{e_m} \right).$$

- minimizing the weighted exponential loss for fixed f_m .
- bigger when the error e_m is smaller
- As $e_m \rightarrow 0$, $\alpha_m \rightarrow \infty$.

- As $e_m \rightarrow 1$, $\alpha_m \rightarrow -\infty$.
5. Repeat: fit the next weak learner on the reweighted data, and so on.

$$w_i^{(m+1)} \propto w_i^{(m)} \cdot \exp(-\alpha_m y_i f_m(\mathbf{x}_i)).$$

After learning, the final classifier is based on a linear combination of the weak classifiers:

$$F(\mathbf{x}) = \sum_{m=1}^M \alpha_m f_m(\mathbf{x}).$$

If more of the strong (high- α) weak learners say $+1$ than -1 , the final answer is $+1$, etc.

This is why it's called **adaptive**: each new weak learner adapts to the mistakes of the previous ones by focusing more weight on hard examples.

8.3.2 Gradient Boosting

Similar to AdaBoost, *gradient boosting* is an ensemble method that combines multiple weak learners (such as decision trees) into a strong learner. Also, similar to AdaBoost (and in contrast to Bagging), gradient boosting is a sequential (rather than parallel) algorithm – it is powerful but rather expensive to train.

In contrast to AdaBoost, we do not adjust the weights for the training examples that have been either misclassified or correctly classified. Also, we do not compute a weight for each model in the sequential gradient boosting ensemble. Instead, in gradient boosting, we *optimize a differentiable loss function* (*e.g.*, mean-squared-error for regression or negative log-likelihood for classification) of a weak learner via consecutive rounds of boosting. The output of gradient boosting is an additive models of multiple weak learners (we do not apply majority voting to the ensemble of models like AdaBoost).

Note that in gradient boosting, we may use decision trees (the most common base learning algorithm for gradient boosting is in fact a decision tree algorithm), but we do not necessarily use decision tree stumps. The first model in gradient boosting is basically just a decision tree's root node (on which we do majority voting in classification or averaging in regression). The consequent models in gradient boosting are deeper decision trees. The depth is usually determined by the practitioner. Values like 8 or 32 are common, depending on the complexity of the dataset/difficulty of the task.

Conceptually, the main idea behind gradient boosting can be summarized via the following three steps:

1. Construct a base tree (just the root node)
2. Build next tree based on errors of the previous tree
3. Combine tree from step 1 with trees from step 2. Go to step 2.

Build a strong predictor by stacking many tiny corrections: start with a simple guess, then repeatedly add small models that fix what's still wrong.

- Start simple: Make an initial prediction for every data point (*e.g.*, the average target value).
- Look at what you got wrong: Compute the residuals = (true predicted).
- For classification, this error is the negative gradient of the loss (that's the "gradient" in gradient boosting).
- Learn to fix the errors: Train a small, weak model (usually a shallow decision tree) to predict those residuals.
- Add a small step: Add a fraction (the learning rate) of that weak model to your running prediction.
- Repeat: Recompute the residuals with the improved predictions, fit another small tree to those, add it in, and so on.

After (M) rounds you have:

$$\hat{y}(x) = f_0(x) + \nu \cdot h_1(x) + \nu \cdot h_2(x) + \cdots + \nu \cdot h_M(x)$$

where (f_0) is the initial guess, each (h_m) is a small tree trained to fix current mistakes, and (ν) is the learning rate (*e.g.*, 0.1).

- Targets: [3, 5, 7].
- Initial guess: mean = 5 → predictions [5, 5, 5].
- Residuals: [-2, 0, +2].
- Fit a tiny tree to map features → residuals (learns left points are low, right points are high).
- Add a small step (say $0.5 \times$ tree output) to your predictions.
- Recompute residuals and repeat. Each round chips away at what's still wrong.

Part II

Regression

Chapter 9

Introduction to Regression Methods

9.1 Introduction

Regression is a process for finding the relationship between the inputs and the outputs. In a regression problem, we consider a set of noisy measurement (or noisy output data) $\mathbf{y} = [y_1, \dots, y_d]^T$ with measurement noise $\boldsymbol{\eta} = [\eta_1, \dots, \eta_d]^T$. We also consider a set of input data $\mathbf{x} = [x_1, \dots, x_d]$. We call the set of these input-output pairs $\mathcal{D} = \{(\mathbf{x}_1, \mathbf{y}_1), \dots, (\mathbf{x}_m, \mathbf{y}_m)\}$ the training data. The true relationship between the input and the output data is unknown. We denote this unknown relationship as a mapping $f(\cdot)$ that takes \mathbf{x}_n and maps it to y_n ,

$$\mathbf{y} = f(\mathbf{x}).$$

Finding the true $f(\cdot)$ from a finite number of data points D is infeasible. There are infinitely many ways to design $f(\cdot)$ for every \mathbf{x}_i . The idea of regression is to add a structure to the problem. Instead of looking for the true $f(\cdot)$, we find a proxy $g_{\theta}(\cdot)$ that takes a certain parameters $\boldsymbol{\theta} = [\theta_1, \dots, \theta_d]^T$. For instance, we can postulate that $(\mathbf{x}_n, \mathbf{y}_n)$ has a linear relationship:

$$g_{\boldsymbol{\theta}}(\mathbf{y}) = \mathbf{X}\boldsymbol{\theta} + \boldsymbol{\eta},$$

where \mathbf{X} is a $m \times d$ input matrix (or our observations). Since we do not know the true relationship, any choice of model is our conjecture. However, we can model the error of our choice. Given a parameter $\boldsymbol{\theta}$, we consider the difference between the noisy measurements and estimated value as follows:

$$\boldsymbol{\epsilon} = \mathbf{y} - \mathbf{X}\boldsymbol{\theta}$$

The purpose of regression is to find the best $\boldsymbol{\theta}$ such that the error is minimized. Therefore, we can consider a following objective function:

$$J(\boldsymbol{\theta}) = \boldsymbol{\epsilon}^T \boldsymbol{\epsilon}$$

Note that this is equivalent to minimizing the mean squared error:

$$MSE = \frac{1}{n} \sum_{i=1}^n (y_i - \mathbf{x}_i \boldsymbol{\theta})^2.$$

We can optimize this in a closed-form as follows:

$$\begin{aligned} J(\boldsymbol{\theta}) &= \|\mathbf{y} - \mathbf{X}\boldsymbol{\theta}\|_2^2 \\ &= (\mathbf{y} - \mathbf{X}\boldsymbol{\theta})^T (\mathbf{y} - \mathbf{X}\boldsymbol{\theta}) \\ &= (\mathbf{y}^T - \boldsymbol{\theta}^T \mathbf{X}^T)(\mathbf{y} - \mathbf{X}\boldsymbol{\theta}) \\ &= \mathbf{y}^T \mathbf{y} - \boldsymbol{\theta}^T \mathbf{X}^T \mathbf{y} - \mathbf{y}^T \mathbf{X}\boldsymbol{\theta} + \boldsymbol{\theta}^T \mathbf{X}^T \mathbf{X}\boldsymbol{\theta} \end{aligned}$$

To find the $\boldsymbol{\theta}$ that minimizes the objective function, we will compute a derivative of the function while setting it equal to zero:

$$\begin{aligned}\frac{\partial J}{\partial \boldsymbol{\theta}} &= -\mathbf{X}^T \mathbf{y} - \mathbf{X}^T \mathbf{y} + \mathbf{X}^T \mathbf{X} \boldsymbol{\theta} + \mathbf{X}^T \mathbf{X} \boldsymbol{\theta} = 0 \\ &\Rightarrow \mathbf{X}^T (\mathbf{X} \boldsymbol{\theta} - \mathbf{y}) = 0 \\ \boldsymbol{\theta} &= (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}\end{aligned}$$

Note that the $\mathbf{X}^T (\mathbf{X} \boldsymbol{\theta} - \mathbf{y})$ is called the *normal equation*.

```

1 import numpy as np
2 import matplotlib.pyplot as plt
3
4 N = 50
5 x = np.random.randn(N)
6 w_1 = 3.4 # True Parameter
7 w_0 = 0.9 # True Parameter
8 y = w_1*x + w_0 + 0.3*np.random.randn(N) # Synthesize training data
9
10 X = np.column_stack((x, np.ones(N)))
11 W = np.array([w_1, w_0])
12
13 # From Scratch
14 XtX = np.dot(X.T, X)
15 XtXinvX = np.dot(np.linalg.inv(XtX), X.T) # d x m
16 W_best = np.dot(XtXinvX, y.T)
17 print(f"W_best: {W_best}")
18
19 # Pythonic Approach
20 theta = np.linalg.lstsq(X, y, rcond=None)[0]
21 print(f"Theta: {theta}")
22
23 t = np.linspace(0, 1, 200)
24 y_pred = W_best[0]*t+W_best[1]
25 yhat = theta[0]*t+theta[1]
26 plt.plot(x, y, 'o')
27 plt.plot(t, y_pred, 'r', linewidth=4)
28 plt.show()
```

9.1.1 MLE Interpretation

If we assume $\epsilon_i \sim \mathcal{N}(0, \sigma^2)$, then

$$y_i | \mathbf{x}_i \sim \mathcal{N}(\mathbf{X} \boldsymbol{\theta}, \sigma^2).$$

Thus, the likelihood (*i.e.*, the probability density of the observations given the parameters) function is given by

$$\begin{aligned}\mathcal{L}(\boldsymbol{\theta}; \mathcal{D}) &= p(\mathbf{y}|X, \boldsymbol{\theta}) = \prod_i^n p(y_i|\mathbf{x}_i, \boldsymbol{\theta}) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(y_i - \boldsymbol{\theta} \mathbf{x}_i)^2}{2\sigma^2}\right) \\ \log \mathcal{L}(\boldsymbol{\theta}; \mathcal{D}) &= \text{const} - \frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \boldsymbol{\theta} \mathbf{x}_i)^2\end{aligned}$$

For a fixed σ , it is equivalent to minimizing the MSE.

9.1.2 Time Complexity

The time complexity of the training process involves the following steps:

- Computing $X^T X$: The product $X^T X$ involves multiplying a $d \times m$ matrix with an $m \times d$ matrix. The time complexity of this matrix multiplication is $O(md^2)$.
- Computing $X^T \mathbf{y}$: $X^T \mathbf{y}$ involves multiplying a $d \times m$ matrix with an $m \times 1$ vector. - The time complexity is $O(md)$.
- Computing $(X^T X)^{-1}$: - The inversion of a $d \times d$ matrix $X^T X$ has a time complexity of $O(d^3)$.
- Multiplying $(X^T X)^{-1}$ with $X^T \mathbf{y}$: This is a matrix-vector multiplication involving a $d \times d$ matrix and a $d \times 1$ vector. Thus, the time complexity is $O(d^2)$.
- Total Training Time Complexity: The dominant terms in the training process are $O(md^2)$ (for computing $X^T X$) and $O(d^3)$ (for inverting $X^T X$). - Therefore, the total time complexity for training a linear regression model is $O(md^2 + d^3)$.

Inference time complexity:

- The inference step requires a matrix-vector multiplication between an $m' \times d$ inference data matrix and a $d \times 1$ vector. - The time complexity of this operation is $O(m'd)$.

9.2 Overdetermined and Underdetermined Systems

Recall that the linear regression problem is an optimization problem of finding the optimal parameter as follows:

$$\boldsymbol{\theta}_{opt} = \underset{\boldsymbol{\theta} \in \mathbb{R}^d}{\operatorname{argmin}} \|y - \mathbf{X}\boldsymbol{\theta}\|^2.$$

We say the optimization problem is *overdetermined* if $\mathbf{X} \in \mathbb{R}^{m \times d}$ is tall and skinny, *i.e.*, $m > d$. This problem has a unique solution $\boldsymbol{\theta} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$ if and only if $\mathbf{X}^T \mathbf{X}$ is invertible. Equivalently, \mathbf{X} should be linearly independent (*i.e.*, full rank).

If \mathbf{X} is fat and short (*i.e.*, $m < d$), a problem is called *underdetermined*. **This problem will have infinitely many solutions.** Among all the feasible solutions, we will pick the one that minimizes the squared norm. The solution is called the *minimum-norm* least squares. Consider an underdetermined linear regression problem:

$$\boldsymbol{\theta} = \underset{\boldsymbol{\theta} \in \mathbb{R}^d}{\operatorname{argmin}} \|\boldsymbol{\theta}\|^2, \text{ subject to } \mathbf{y} = \mathbf{X}\boldsymbol{\theta},$$

where $\mathbf{X} \in \mathbb{R}^{m \times d}$, $\boldsymbol{\theta} \in \mathbb{R}^d$, and $\mathbf{y} \in \mathbb{R}^m$. If the matrix has $\operatorname{rank}(\mathbf{X}) = m$, then the linear regression problem will have a unique global minimum

$$\boldsymbol{\theta} = \mathbf{X}^T (\mathbf{X} \mathbf{X}^T)^{-1} \mathbf{y}.$$

This solution is called the minimum-norm least-squares solution. The proof of this solution is given by:

$$\mathcal{L}(\boldsymbol{\theta}, \boldsymbol{\lambda}) = \|\boldsymbol{\theta}\|^2 + \boldsymbol{\lambda}^T (\mathbf{X}\boldsymbol{\theta} - \mathbf{y}),$$

where λ is a Lagrange multiplier. The solution of the constrained optimization is the stationary point of the Lagrangian. To find it, we take the derivatives w.r.t., λ and θ as follows:

$$\begin{aligned}\nabla_{\theta} &= 2\theta + \mathbf{X}^T \lambda = 0 \\ \nabla_{\lambda} &= \mathbf{X}\theta - \mathbf{y} = 0\end{aligned}$$

The first equation gives us $\theta = -\mathbf{X}^T \lambda / 2$. Substituting it into the second equation, and assuming that $\text{rank}(\mathbf{X}) = m$ so that $\mathbf{X}^T \mathbf{X}$ is invertible, we have $\lambda = -2(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{y}$. Thus, we have

$$\theta = \mathbf{X}^T (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{y}.$$

Note that $\mathbf{X}^T \mathbf{X}$ is often called a *Gram matrix*, \mathbf{G} .

9.3 Overfitting

We examine the relationship between the number of training samples and the complexity of the model.

9.4 Ridge Regression

Regularization means that instead of seeking the model parameters by minimizing the training loss alone, we add a penalty term to force the parameters to “behave better”.

With the ridge regression principle, we can optimize it as follows:

$$J(\theta) = \|\mathbf{y} - \mathbf{X}\theta\|_2^2 + \lambda \|\theta\|_2^2 \quad (9.1)$$

$$= (\mathbf{y} - \mathbf{X}\theta)^T (\mathbf{y} - \mathbf{X}\theta) + \lambda \theta^T \theta \quad (9.2)$$

$$= (\mathbf{y}^T - \theta^T \mathbf{X}^T)(\mathbf{y} - \mathbf{X}\theta) + \lambda \theta^T \theta \quad (9.3)$$

$$= \mathbf{y}^T \mathbf{y} - \theta^T \mathbf{X}^T \mathbf{y} - \mathbf{y}^T \mathbf{X}\theta + \theta^T \mathbf{X}^T \mathbf{X}\theta + \theta^T \lambda \mathbf{I}\theta \quad (9.4)$$

$$\frac{\partial J}{\partial \theta} = -\mathbf{X}^T \mathbf{y} - \mathbf{X}^T \mathbf{y} + \mathbf{X}^T \mathbf{X}\theta + \mathbf{X}^T \mathbf{X}\theta + 2\lambda \mathbf{I}\theta = 0 \quad (9.5)$$

$$\theta = (\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I})^{-1} \mathbf{X}^T \mathbf{y} \quad (9.6)$$

- If $\lambda \rightarrow 0$, then $\|\mathbf{y} - \mathbf{X}\theta\|_2^2 + \underbrace{\lambda \|\theta\|_2^2}_{=0}$
- $\lambda \rightarrow \infty$, then $\underbrace{\frac{1}{\lambda} \|\mathbf{y} - \mathbf{X}\theta\|_2^2 + \|\theta\|_2^2}_{=0}$, since what we want to do is to minimize the objective function, we can divide it by λ . Therefore, the solution will be $\theta = 0$, because it is the smallest value the squared function can achieve.

Note that $\mathbf{X}^T \mathbf{X}$ is always symmetric ¹. Thus, it can be decomposed as $Q\Lambda Q^T$ by the Spectral theorem. The Q and Λ are eigenvector and eigenvalue matrices, respectively. Then, the inverse

¹ $(\mathbf{X}^T \mathbf{X})^T = \mathbf{X}^T \mathbf{X}$.

operation in the ridge regression can be expressed as follows:

$$\begin{aligned}\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I} &= \mathbf{Q} \boldsymbol{\Lambda} \mathbf{Q}^T + \lambda \mathbf{I} \\ &= \mathbf{Q} \boldsymbol{\Lambda} \mathbf{Q}^T + \lambda \mathbf{Q} \mathbf{Q}^T \\ &= \mathbf{Q} (\boldsymbol{\Lambda} + \lambda \mathbf{I}) \mathbf{Q}^T.\end{aligned}$$

Even if the symmetric matrix is not invertible or close to not invertible, the regularization constant λ makes it invertible (by making it to be a full-rank).

Note that we can change the ridge regression into a dual form:

$$(\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I}) \boldsymbol{\theta} = (\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I})(\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I})^{-1} \mathbf{X}^T \mathbf{y} \quad (9.7)$$

$$(\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I}) \boldsymbol{\theta} = \mathbf{X}^T \mathbf{y} \quad (9.8)$$

$$\boldsymbol{\theta} = \lambda^{-1} \mathbf{I} (\mathbf{X}^T \mathbf{y} - \mathbf{X}^T \mathbf{X} \boldsymbol{\theta}) \quad (9.9)$$

$$\begin{aligned}&= \mathbf{X}^T \underbrace{\lambda^{-1} (\mathbf{y} - \mathbf{X} \boldsymbol{\theta})}_{=\alpha} \\ &= \mathbf{X}^T \alpha \quad (9.10)\end{aligned}$$

$$= \mathbf{X}^T \alpha \quad (9.11)$$

This tells us that $\boldsymbol{\theta}$ is a linear combination of \mathbf{X} (*i.e.*, training data points) with α . Then,

$$\lambda \alpha = (\mathbf{y} - \mathbf{X} \boldsymbol{\theta}) \quad (9.12)$$

$$= (\mathbf{y} - \mathbf{X} \mathbf{X}^T \alpha) \quad (9.13)$$

$$\mathbf{y} = (\mathbf{X} \mathbf{X}^T \alpha + \lambda \alpha) \quad (9.14)$$

$$\alpha = (\mathbf{X} \mathbf{X}^T + \lambda)^{-1} \mathbf{y} \quad (9.15)$$

$$= (\mathbf{G} + \lambda)^{-1} \mathbf{y}. \quad (9.16)$$

This gives us the solution of the *underdetermined* problems. For a new data point \mathbf{x}_{new} , we can make an inference by computing dot products between the new data point and each training data point:

$$\begin{aligned}y &= \boldsymbol{\theta}^T \mathbf{x}_{new} \\ &= \alpha^T \mathbf{X} \mathbf{x}_{new}\end{aligned}$$

9.4.1 Time Complexity

- Training time: $O(md^2 + m^3)$
- Inference time: $O(md)$

9.5 Weighted LSE

The OLEs assume an equal confidence on all the measurements. Now we look at varying confidence in the measurements. We assume that the noise for each measurement has zero mean and is independent, then the covariance matrix for all measurement noise is given by

$$\begin{aligned}R &= \mathbb{E}(\eta \eta^T) \\ &= \begin{bmatrix} \sigma_1^2 & \dots & 0 \\ \vdots & \ddots & \vdots \\ 0 & \dots & \sigma_l^2 \end{bmatrix}\end{aligned}$$

By denoting the error vector $\mathbf{y} - \mathbf{X}\boldsymbol{\theta}$ as $\boldsymbol{\epsilon} = (\epsilon_1, \dots, \epsilon_l)^T$, we will minimize the sum of squared differences weighted over the variations of the measurements:

$$\begin{aligned} J(\tilde{\mathbf{x}}) &= \boldsymbol{\epsilon}^T R^{-1} \boldsymbol{\epsilon} = \frac{\epsilon_1^2}{\sigma_1^2} + \dots + \frac{\epsilon_l^2}{\sigma_l^2} \\ &= (\mathbf{y} - \mathbf{X}\boldsymbol{\theta})^T R^{-1} (\mathbf{y} - \mathbf{X}\boldsymbol{\theta}) \end{aligned}$$

Note that by dividing each residual by its variance, we effectively equalize the influence of each data point on the overall fitting process. Subsequently, by taking the partial derivative of J with respect to $\boldsymbol{\theta}$, we get the best estimate of the parameter, which is given by

$$\boldsymbol{\theta} = (\mathbf{X}^T R^{-1} \mathbf{X})^{-1} \mathbf{X}^T R^{-1} \mathbf{y}.$$

Note that the measurement noise matrix R must be non-singular for a solution to exist.

9.6 Robust Linear Regression

The linear regression is based on the squared error criterion. This criterion often suffers from a serious drawback caused by outliers. By the definition of a squared error, our training loss is given by

$$\mathcal{L}(\boldsymbol{\theta}) = \sum_{n=1}^N \left(y_n - g_{\boldsymbol{\theta}}(x_n) \right)^2.$$

Let's assume that one of these error terms is large due to an outlier.

$$\mathcal{L}(\boldsymbol{\theta}) = \left(y_1 - g_{\boldsymbol{\theta}}(x_1) \right)^2 + \underbrace{\left(y_2 - g_{\boldsymbol{\theta}}(x_2) \right)^2}_{\text{Large}} + \dots$$

If one or a few of these individual error terms are large, the square operation will amplify them. Consequently, the outliers suddenly have a very large contribution to the error. Since the goal of linear regression is to minimize the total loss, the presence of the outliers will drive the optimization solution to compensate for the large error.

Chapter 10

Recursive Least Squares

10.1 Recursive Least Squares

The ordinary least-squares assumes that all measurements are available at a certain time. However, this often might not be the case in practice. **More often, we obtain measurements sequentially and want to update our estimate with each new measurement.** In this case, the matrix \mathbf{X} needs to be augmented. This update can be very expensive especially if the number of measurements is extremely large, the solutions of the least squares problem become prohibitive to compute.

This motivates the Recursive Least Squares (RLS). Suppose we have an estimate $\boldsymbol{\theta}_{k-1}$ after $(k-1)$ measurements and obtain a new measurement \mathbf{y}_k . How can we update our estimate without completely reworking on the pseudo-inverse problem?

A linear recursive estimator can be expressed in the following form:

$$\begin{aligned}\mathbf{y}_k &= \mathbf{X}_k \boldsymbol{\theta} + \boldsymbol{\eta}_k \\ \boldsymbol{\theta}_k &= \boldsymbol{\theta}_{k-1} + K_k (\mathbf{y}_k - \mathbf{X}_k \boldsymbol{\theta}_{k-1})\end{aligned}$$

Here, \mathbf{X}_k is an $m \times d$ matrix (observations) and K_k is $d \times m$ and referred to as the *estimator gain matrix*. We refer to $(\mathbf{y}_k - \mathbf{X}_k \boldsymbol{\theta}_{k-1})$ as the *correction term*. Also, $\boldsymbol{\eta}_k$ is the measurement error. The new estimate is modified from the previous estimate $\boldsymbol{\theta}_{k-1}$ with a correction via the gain matrix.

Intuitively, we can notice that we have to compute the optimal gain matrix to update our estimate. To this end, we have to set an *estimation error*, which is our learning objective. The error can be expressed as follows:

$$\begin{aligned}\boldsymbol{\epsilon}_k &= \boldsymbol{\theta} - \boldsymbol{\theta}_k \\ &= \boldsymbol{\theta} - \boldsymbol{\theta}_{k-1} - K_k (\mathbf{y}_k - \mathbf{X}_k \boldsymbol{\theta}_{k-1}) \\ &= \boldsymbol{\epsilon}_{k-1} - K_k (\mathbf{X}_k \boldsymbol{\theta} + \boldsymbol{\eta}_k - \mathbf{X}_k \boldsymbol{\theta}_{k-1}) \\ &= \boldsymbol{\epsilon}_{k-1} - K_k \mathbf{X}_k (\boldsymbol{\theta} - \boldsymbol{\theta}_{k-1}) - K_k \boldsymbol{\eta}_k \\ &= (I - K_k \mathbf{X}_k) \boldsymbol{\epsilon}_{k-1} - K_k \boldsymbol{\eta}_k,\end{aligned}$$

where I is the $d \times d$ identity matrix. The mean of this error is then

$$\mathbb{E}[\boldsymbol{\epsilon}_k] = (I - K_k \mathbf{X}_k) \mathbb{E}[\boldsymbol{\epsilon}_{k-1}] - K_k \mathbb{E}[\boldsymbol{\eta}_k]$$

If $\mathbb{E}[\boldsymbol{\eta}_k] = 0$ and $\mathbb{E}[\boldsymbol{\epsilon}_{k-1}] = 0$, then $\mathbb{E}[\boldsymbol{\epsilon}_k] = 0$. So if the measurement noise has zero mean for all k , and the initial estimate of $\boldsymbol{\theta}$ is set equal to its expected value, then $\boldsymbol{\theta}_k = \boldsymbol{\theta}_k, \forall k$. This property tells us that the estimator $\boldsymbol{\theta}_k = \boldsymbol{\theta}_{k-1} + K_k(\mathbf{y}_k - \mathbf{X}_k\boldsymbol{\theta}_{k-1})$ is *unbiased*. This property holds regardless of the value of the gain vector K_k . This means the estimate will be equal to the true value $\boldsymbol{\theta}$ on average.

The key is to determine the optimal value of the gain vector K_k . The optimality criterion is to minimize the aggregated variance of the estimation errors at time k :

$$\begin{aligned} J_k &= \mathbb{E}[\|\boldsymbol{\theta} - \boldsymbol{\theta}_k\|^2] \\ &= \mathbb{E}[\boldsymbol{\epsilon}_k^T \boldsymbol{\epsilon}_k] \\ &= \mathbb{E}[tr(\boldsymbol{\epsilon}_k \boldsymbol{\epsilon}_k^T)] \\ &= tr(P_k), \end{aligned}$$

where $P_k = \mathbb{E}[\boldsymbol{\epsilon}_k \boldsymbol{\epsilon}_k^T]$ is the *estimation-error covariance* (i.e., **covariance matrix**). Note that the third line holds by the trace of a product (i.e., *cyclic property*) and the expectation in the third line can go into the trace operator by its linearity. Next, we can obtain P_k by

$$P_k = \mathbb{E}\left[((I - K_k \mathbf{X}_k)\boldsymbol{\epsilon}_{k-1} - K_k \boldsymbol{\eta}_k)((I - K_k \mathbf{X}_k)\boldsymbol{\epsilon}_{k-1} - K_k \boldsymbol{\eta}_k)^T\right]$$

By rearranging the above equation with the property that the mean of noise is zero, we can get

$$P_k = (I - K_k \mathbf{X}_k)P_{k-1}(I - K_k \mathbf{X}_k)^T + K_k R_k K_k^T, \quad (10.1)$$

where $R_k = \mathbb{E}[\boldsymbol{\eta}_k \boldsymbol{\eta}_k^T]$ as covariance of $\boldsymbol{\eta}_k$. This equation is the recurrence for the covariance of the least squares estimation error. It is consistent with the intuition that as the measurement noise R_k increases, the uncertainty in our estimate also increases (i.e., P_k increases). Note that P_k should be positive definite since it is a covariance matrix.

Next, let's compute K_k that minimizes the cost function given by the error equation. We are going to utilize the following property:

$$\begin{aligned} \frac{\partial tr(AC^T)}{\partial A} &= C \\ \frac{\partial tr(ACA^T)}{\partial A} &= AC + AC^T \end{aligned}$$

Next, we are going to take a derivative to the objective function:

$$\begin{aligned} \frac{\partial J_k}{\partial K_k} &= \frac{\partial tr(P_k)}{\partial K_k} = \frac{\partial tr}{\partial K_k} \underbrace{(I - K_k \mathbf{X}_k)P_{k-1}(I - K_k \mathbf{X}_k)^T}_{=ACA^T} + \frac{\partial}{\partial K_k} tr(K_k R_k K_k^T) \\ &= \frac{\partial tr(ACA^T)}{\partial(I - K_k \mathbf{X}_k)} \frac{\partial(I - K_k \mathbf{X}_k)}{\partial K_k} + \frac{\partial}{\partial K_k} tr(K_k R_k K_k^T) \quad \text{by Chain Rule} \\ &= ((I - K_k \mathbf{X}_k)P_{k-1} + (I - K_k \mathbf{X}_k)P_{k-1}^T)(-\mathbf{X}_k^T) + \frac{\partial}{\partial K_k} tr(K_k R_k K_k^T) \\ &= 2(I - K_k \mathbf{X}_k)P_{k-1}(-\mathbf{X}_k^T) + \frac{\partial}{\partial K_k} tr(K_k R_k K_k^T) \quad , \text{since } P_{k-1} \text{ is symmetric.} \\ &= -2(I - K_k \mathbf{X}_k)P_{k-1}\mathbf{X}_k^T + 2K_k R_k \end{aligned}$$

By setting the partial derivative to zero, we get

$$K_k = P_{k-1}\mathbf{X}_k^T(\mathbf{X}_k P_{k-1}\mathbf{X}_k^T + R_k)^{-1}.$$

10.1.1 Alternative Form

Sometimes it is useful to write the equations for P_k and K_k in alternate forms. Although these alternate forms are mathematically identical, they can be beneficial from a computational point of view. Let's first set $\mathbf{X}_k P_{k-1} \mathbf{X}_k^T + R_k = S_k$, then we get

$$K_k = P_{k-1} \mathbf{X}_k^T S_k^{-1}.$$

By putting this into Eq. (10.1),

$$\begin{aligned} P_k &= (I - P_{k-1} \mathbf{X}_k^T S_k^{-1} \mathbf{X}_k) P_{k-1} (I - P_{k-1} \mathbf{X}_k^T S_k^{-1} \mathbf{X}_k)^T + P_{k-1} \mathbf{X}_k^T S_k^{-1} R_k S_k^{-1} \mathbf{X}_k P_{k-1} \\ &\quad \vdots \\ &= P_{k-1} - P_{k-1} \mathbf{X}_k^T S_k^{-1} \mathbf{X}_k^T P_{k-1} \\ &= (I - K_k \mathbf{X}_k) P_{k-1}. \end{aligned}$$

Note that P_k is symmetric (*c.f.*, $P_k = \epsilon_k \epsilon_k^T$), since it is a covariance matrix, and so is S_k .

We take the inverse of both sides of

$$P_{k-1}^{-1} = \left(\underbrace{P_{k-1}}_A - \underbrace{P_{k-1} \mathbf{X}_k^T}_{B} \underbrace{(\mathbf{X}_k P_{k-1} \mathbf{X}_k^T)}_D^{-1} \underbrace{\mathbf{X}_k P_{k-1}}_C \right)^{-1}.$$

Next, we apply the matrix inversion lemma which is known as *Sherman-Morrison-Woodbury matrix identity* (or *matrix inversion lemma*) identity:

$$(A - BD^{-1}C)^{-1} = A^{-1} + A^{-1}B(D - CA^{-1}B)^{-1}CA^{-1}.$$

Then, rewrite P_k^{-1} as follows:

$$\begin{aligned} P_k^{-1} &= P_{k-1}^{-1} + P_{k-1}^{-1} P_{k-1} \mathbf{X}_k^T ((\mathbf{X}_k P_{k-1} \mathbf{X}_k^T + R_k) - \mathbf{X}_k P_{k-1} P_{k-1}^{-1} (P_{k-1} \mathbf{X}_k^T))^{-1} \mathbf{X}_k P_{k-1} P_{k-1}^{-1} \\ &= P_{k-1}^{-1} + \mathbf{X}_k^T R_k^{-1} \mathbf{X}_k \end{aligned}$$

This yields an alternative expression for the covariance matrix:

$$P_k = (P_{k-1}^{-1} + \mathbf{X}_k^T R_k^{-1} \mathbf{X}_k)^{-1}$$

We can also obtain

$$K_k = P_k \mathbf{X}_k^T R_k^{-1}$$

By

$$\begin{aligned} P_k &= (I - K_k \mathbf{X}_k) P_{k-1} \\ P_k P_{k-1}^{-1} &= (I - K_k \mathbf{X}_k) \\ P_k P_k^{-1} &= P_k P_{k-1}^{-1} + P_k \mathbf{X}_k^T R_k^{-1} \mathbf{X}_k = I \\ I &= (I - K_k \mathbf{X}_k) + P_k \mathbf{X}_k^T R_k^{-1} \mathbf{X}_k \\ K_k &= P_k \mathbf{X}_k^T R_k^{-1}. \end{aligned}$$

10.1.2 Summary of RLS

In sum, RLS can be updated as follows:

- Update the gain matrix:

$$\begin{aligned} - K_k &= P_{k-1} \mathbf{X}_k^T (\mathbf{X}_k P_{k-1} \mathbf{X}_k^T + R_k)^{-1} \text{ or} \\ - K_k &= P_k \mathbf{X}_k^T R_k^{-1} \end{aligned}$$

- Update estimate: $\boldsymbol{\theta}_k = \boldsymbol{\theta}_{k-1} + K_k (\mathbf{y}_k - \mathbf{X}_k \boldsymbol{\theta}_{k-1})$

- Update error covariance matrix by either:

$$\begin{aligned} - P_k &= (I - K_k \mathbf{X}_k) P_{k-1}. \\ - P_k &= (I - K_k \mathbf{X}_k) P_{k-1} (I - K_k \mathbf{X}_k)^T + K_k R_k K_k^T, \end{aligned}$$

Example: At sample time k , our measurement is

- $y_k = X_k \theta + \eta_k$
- $X_k = 1$
- $R_k = \mathbb{E}[\eta_k^2]$

For this scalar problem, the measurement matrix X_k is a scalar too, and the measurement noise covariance R_k is also a scalar. We will suppose that each measurement has the same covariance so the measurement covariance R_k is not a function of k , and can be written as R . Initially, before we have any measurements, we have some idea about the value of the θ , and this forms our initial estimate. We also have some uncertainty about our initial estimate, and this forms our initial covariance:

$$\begin{aligned} \hat{\theta}_0 &= \mathbb{E}[\theta] \\ P_0 &= \mathbb{E}[(\theta - \hat{\theta}_0)(\theta - \hat{\theta}_0)^T] \\ &= \mathbb{E}[(\theta - \hat{\theta}_0)^2] \end{aligned}$$

If we have absolutely no idea about θ , then $P(0) = \infty I$. If we are 100% certain about the θ before taking any measurements, then $P(0) = 0$. Let's compute the gain matrix at $k = 1$ by using the following equation:

$$K_k = P_{k-1} \mathbf{X}_k^T (\mathbf{X}_k P_{k-1} \mathbf{X}_k^T + R_k)^{-1}.$$

Then, we get

$$K_1 = P_0 (P_0 + R)^{-1}.$$

Similarly, by

$$\boldsymbol{\theta}_k = \boldsymbol{\theta}_{k-1} + K_k (\mathbf{y}_k - \mathbf{X}_k \boldsymbol{\theta}_{k-1}),$$

we obtain

$$\hat{\theta}_1 = \hat{\theta}_0 + \frac{P_0}{P_0 + R} (y_1 - \hat{\theta}_0).$$

Finally, let's update our covariance matrix P_k by

$$P_k = (I - K_k \mathbf{X}_k) P_{k-1} (I - K_k \mathbf{X}_k)^T + K_k R_k K_k^T.$$

Then,

$$\begin{aligned} P_1 &= \left(I - \frac{P_0}{P_0 + R} \right) P_0 I - \frac{P_0}{P_0 + R} + \frac{P_0}{P_0 + R} R \frac{P_0}{P_0 + R} \\ &= \left(\frac{P_0 R^2}{(P_0 + R)^2} \right) + \frac{P_0^2 R}{(P_0 + R)^2} \\ &= \frac{P_0 R (P_0 + R)}{(P_0 + R)^2} \\ &= \frac{P_0 R}{P_0 + R} \end{aligned}$$

By repeating these calculations, we can update the above parameters and find general expressions:

$$\begin{aligned} P_{k-1} &= \frac{P_0 R}{(k-1)P_0 + R} \\ K_k &= \frac{P_0}{kP_0 + R} \\ \hat{\theta}_k &= \frac{(k-1)P_0 + R}{kP_0 + R} \hat{\theta}_{k-1} + \frac{P_0}{kP_0 + R} y_k \end{aligned}$$

Note that if θ is known perfectly *a priori* (i.e., θ is known perfectly before any measurements are obtained) then $P_0 = 0$ and the above equation show that $K_k = 0$ and $\hat{\theta} = \hat{\theta}_0$. That is, the optimal estimate of θ is independent of any measurements that are obtained. In sum, this indicates that no update from measurements is needed, as the estimate is already perfect.

On the other hand, if x is completely unknown *a priori*, then $P_0 \rightarrow \infty$, and the above equations simplify to

$$\begin{aligned} \hat{\theta}_k &= \frac{(k-1)P_0}{kP_0} \hat{\theta}_{k-1} + \frac{P_0}{kP_0} y_k \\ &= \frac{k-1}{k} \hat{\theta}_{k-1} + \frac{1}{k} y_k \\ &= \frac{1}{k} [(k-1)\hat{\theta}_{k-1} + y_k] \end{aligned}$$

In other words, the optimal estimate of θ is equal to the running average of the measurements y_k , which can be written as

$$\begin{aligned} \bar{y}_k &= \frac{1}{k} \sum_{j=1}^k y_j \\ &= \frac{1}{k} \left(\sum_{j=1}^{k-1} y_j + y_k \right) \\ &= \frac{1}{k} \left[(k-1) \frac{1}{k-1} \sum_{j=1}^{k-1} y_j + y_k \right] \\ &= \frac{1}{k} [(k-1)\bar{y}_{k-1} + y_k] \end{aligned}$$

10.1.3 Curve Fitting

In the recursive curve fitting problem, we measure data one sample at a time ($y_1, y_2 \dots$,) and want to find the best fit of a curve to the data. The curve that we want to fit to the data could be constrained to be linear or quadratic and so on.

Example: Suppose that we want to fit a straight line to a set of data points. The linear data fitting problem can be written as

$$\mathbb{E}[\eta_k] = R_k$$

t_k is the independent variable, y_k is the noisy data, and we want to find the linear relationship between y_k and t_k . In sum, we want to estimate the constants θ_1 and θ_2 . The measurement matrix can be written as

$$\mathbf{X}_k = \begin{bmatrix} 1 & t_k \end{bmatrix}.$$

Then,

$$\mathbf{y}_k = \mathbf{X}_k \boldsymbol{\theta} + \boldsymbol{\eta}_k.$$

10.1.4 Python Implementation

```
1 class RecursiveLeastSquares(object):
2
3     # theta0 - initial estimate used to initialize the estimator
4     # P0 - initial estimation error covariance matrix
5     # R - covariance matrix of the measurement noise
6     def __init__(self,theta0,P0,R)
7
8         # initialize the values
9         self.theta0=theta0
10        self.P0=P0
11        self.R=R
12
13        # this variable is used to track the current time step k of the
14        estimator
15        # after every time step arrives, this variables increases for one
16        # in this way, we can track the number of variblaes
17        self.curr_step=0
18
19        # this list is used to store the estimates xk starting from the initial
20        estimate
21        self.estimates=[]
22        self.estimates.append(theta0)
23
24        # this list is used to store the estimation error covariance matrices
25        Pk
26        self.est_error_cov=[]
27        self.est_error_cov.append(P0)
28
29        # this list is used to store the gain matrices Kk
30        self.gainMatrices=[]
31
32        # this list is used to store estimation error vectors
33        self.errors=[]
```

```

32
33     # this function takes the current measurement and the current measurement
34     # matrix X
35     # and computes the estimation error covariance matrix, updates the estimate
36     ,
37     # computes the gain matrix, and the estimation error
38     # it fills the lists self.estimates, self.est_error_cov, self.gainMatrices,
39     # and self.errors
40     # it also increments the variable curr_step for 1
41
42     def predict(self,measurementValue,X):
43         import numpy as np
44
45         # Compute the L matrix and its inverse
46         #  $K_k = P_{k-1}X_k^T(R_k+X_kP_{k-1}X_k^T)^{-1}$ 
47         Lmatrix=self.R+np.matmul(X,np.matmul(self.est_error_cov[self.curr_step],X.T))
48         LmatrixInv=np.linalg.inv(Lmatrix)
49         # Compute the gain matrix
50         gainMatrix=np.matmul(self.est_error_cov[self.curr_step],np.matmul(X.T,LmatrixInv))
51
52         # Compute the estimation error
53         #  $\theta_k = \theta_{k-1} + K_k (y_k - X_k \theta_{k-1})$ 
54         error=measurementValue-np.matmul(X,self.estimates[self.curr_step])
55         # Compute the estimate
56         estimate=self.estimates[self.curr_step]+np.matmul(gainMatrix,error)
57
58         # Propagate the estimation error covariance matrix
59         #  $P_k = (I - K_k X_k)P_{k-1}(I - K_k X_k)^T + K_k R_k K_k^T$ 
60         ImKc=np.eye(np.size(self.theta0),np.size(self.theta0))-np.matmul(
gainMatrix,X)
61         error_cov=np.matmul(ImKc,self.est_error_cov[self.curr_step])
62
63         # add computed elements to the list
64         self.estimates.append(estimate)
65         self.est_error_cov.append(error_cov)
66         self.gainMatrices.append(gainMatrix)
67         self.errors.append(error)
68
69         # increment the current time step
70         self.curr_step=self.curr_step+1

```

10.2 Alternate Derivation of RLS

Suppose the training samples arrive one by one in the following sequence $\mathbf{x}_1, \dots, \mathbf{x}_m, \mathbf{x}_{m+1}$, where \mathbf{x}_{m+1} denotes the newly arrived sample vector. These samples can be projected onto the feature space by linear projection and expressed into a matrix $\mathbf{P}^T \in \mathbb{R}^{(d+1) \times (m+1)}$ as follows:

$$\mathbf{P}^T = [\mathbf{p}(\mathbf{x}_1), \dots, \mathbf{p}(\mathbf{x}_{m+1})],$$

where $\mathbf{p}(\cdot) \in \mathbb{R}^{d+1}$. Subsequently, let

$$\begin{aligned}\mathbf{R}_{m+1} &= \mathbf{P}^T \mathbf{P} + b\mathbf{I} \\ \mathbf{Q}_{m+1} &= \mathbf{P}^T \mathbf{y}.\end{aligned}$$

By separating the covariance of the newly arrived sample $p(\mathbf{x}_{m+1})$ from the remaining stack, we can write:

$$\begin{aligned}\mathbf{P}^T \mathbf{P} &= \sum_{i=1}^{m+1} \mathbf{p}(\mathbf{x}_i) \mathbf{p}(\mathbf{x}_i)^T \\ &= \sum_{i=1}^m \mathbf{p}(\mathbf{x}_i) \mathbf{p}(\mathbf{x}_i)^T + \mathbf{p}(\mathbf{x}_{m+1}) \mathbf{p}(\mathbf{x}_{m+1})^T \\ &= \mathbf{P}_m^T \mathbf{P}_m + \mathbf{p}(\mathbf{x}_{m+1}) \mathbf{p}(\mathbf{x}_{m+1})^T.\end{aligned}$$

Hence,

$$\begin{aligned}\mathbf{R}_{m+1} &= \mathbf{P}^T \mathbf{P} + b\mathbf{I} \\ &= (\mathbf{P}_m^T \mathbf{P}_m + \mathbf{p}(\mathbf{x}_{m+1}) \mathbf{p}(\mathbf{x}_{m+1})^T) + b\mathbf{I} \\ &= \underbrace{\mathbf{P}_m^T \mathbf{P}_m + b\mathbf{I}}_{=\mathbf{R}_m} + \mathbf{p}(\mathbf{x}_{m+1}) \mathbf{p}(\mathbf{x}_{m+1})^T \\ &= \mathbf{R}_m + \mathbf{p}(\mathbf{x}_{m+1}) \mathbf{p}(\mathbf{x}_{m+1})^T\end{aligned}$$

Similarly,

$$\mathbf{Q}_{m+1} = \mathbf{Q}_m + \mathbf{p}(\mathbf{x}_{m+1}) y_{m+1}$$

If the system is designed to forget the old training samples (*i.e.*, weighted averaging),

$$\begin{aligned}\mathbf{R}_{m+1} &= (1 - \lambda) \mathbf{R}_m + \lambda \mathbf{p}(\mathbf{x}_{m+1}) \mathbf{p}(\mathbf{x}_{m+1})^T, \\ \mathbf{Q}_{m+1} &= (1 - \lambda) \mathbf{Q}_m + \lambda \mathbf{p}(\mathbf{x}_{m+1}) y_{m+1},\end{aligned}$$

where $\lambda \in (0, 1)$ is often called a *forgetting factor*.

Let $\mathbf{A} = \mathbf{R}_m$, $\mathbf{B} = p(\mathbf{x}_{m+1})$, $\mathbf{C} = 1$ (scalar), $\mathbf{D} = p(\mathbf{x}_{m+1})^T = \mathbf{p}^T$, then based on the matrix inversion lemma (Woodbury, 1950; Sherman and Morrison, 1950),

$$(\mathbf{A} + \mathbf{BCD})^{-1} = \mathbf{A}^{-1} - \mathbf{A}^{-1} \mathbf{B} (\mathbf{C}^{-1} + \mathbf{D} \mathbf{A}^{-1} \mathbf{B})^{-1} \mathbf{D} \mathbf{A}^{-1},$$

we have

$$\begin{aligned}\mathbf{R}_{m+1}^{-1} &= [(1 - \lambda) \mathbf{R}_m + \lambda \mathbf{p} \mathbf{p}^T]^{-1} \\ &= \frac{1}{1 - \lambda} \mathbf{R}_m^{-1} - \frac{1}{1 - \lambda} \mathbf{R}_m^{-1} \lambda \mathbf{p} \left(\mathbf{I} + \mathbf{p}^T \frac{\lambda}{1 - \lambda} \mathbf{R}_m^{-1} \mathbf{p} \right)^{-1} \mathbf{p}^T \frac{1}{1 - \lambda} \mathbf{R}_m^{-1} \\ &= \frac{1}{1 - \lambda} \mathbf{R}_m^{-1} - \frac{1}{(1 - \lambda)^2} \mathbf{R}_m^{-1} \mathbf{p} \mathbf{p}^T \mathbf{R}_m^{-1} \left(\frac{1}{\lambda} + \frac{1}{1 - \lambda} \mathbf{p}^T \mathbf{R}_m^{-1} \mathbf{p} \right)^{-1}.\end{aligned}$$

We can obtain

$$\mathbf{w}_{m+1} = (\mathbf{P}^T \mathbf{P} + b\mathbf{I})^{-1} \mathbf{P}^T \mathbf{y} = \mathbf{R}_{m+1}^{-1} \mathbf{Q}_{m+1},$$

Substitute \mathbf{R}_{m+1}^{-1} and $\mathbf{Q}_{m+1} = (1 - \lambda) \mathbf{Q}_m + \lambda \mathbf{p}(\mathbf{x}_{m+1}) y_{m+1}$:

$$\begin{aligned}\mathbf{w}_{m+1} &= \left[\frac{1}{1 - \lambda} \mathbf{R}_m^{-1} - \frac{1}{(1 - \lambda)^2} \mathbf{R}_m^{-1} \mathbf{p} \mathbf{p}^T \mathbf{R}_m^{-1} \left(\frac{1}{\lambda} + \frac{1}{1 - \lambda} \mathbf{p}^T \mathbf{R}_m^{-1} \mathbf{p} \right)^{-1} \right] [(1 - \lambda) \mathbf{Q}_m + \lambda y_{m+1} \mathbf{p}] \\ &= \underbrace{\mathbf{R}_m^{-1} \mathbf{Q}_m}_{=w_m} + \frac{1}{1 - \lambda} \mathbf{R}_m^{-1} \mathbf{p} \mathbf{p}^T \mathbf{R}_m^{-1} \left(\frac{1}{\lambda} + \frac{1}{1 - \lambda} \mathbf{p}^T \mathbf{R}_m^{-1} \mathbf{p} \right)^{-1} \mathbf{Q}_m + \frac{\lambda}{1 - \lambda} \mathbf{R}_m^{-1} \mathbf{p} y_{m+1} \\ &\quad - \frac{\lambda}{(1 - \lambda)^2} \mathbf{R}_m^{-1} \mathbf{p} \mathbf{p}^T \mathbf{R}_m^{-1} \left(\frac{1}{\lambda} + \frac{1}{1 - \lambda} \mathbf{p}^T \mathbf{R}_m^{-1} \mathbf{p} \right)^{-1} \mathbf{p} y_{m+1}\end{aligned}$$

Let

$$\begin{aligned} A &= \left(\frac{1}{\lambda} + \frac{1}{1-\lambda} \mathbf{p}^T \mathbf{R}_m^{-1} \mathbf{p} \right)^{-1} \\ &= \frac{\lambda(1-\lambda)}{\lambda \mathbf{p}^T \mathbf{R}_m^{-1} \mathbf{p} + (1-\lambda)}, \end{aligned}$$

which is a constant. Then

$$\begin{aligned} w_{m+1} &= w_m - \frac{\mathbf{R}_m^{-1} \mathbf{p}}{(1-\lambda)^2} \cdot A \cdot ((1-\lambda) \mathbf{p}^T w_m + \lambda \mathbf{p}^T \mathbf{R}_m^{-1} \mathbf{p} y_{m+1}) + \frac{\lambda}{1-\lambda} \mathbf{R}_m^{-1} \mathbf{p} y_{m+1} \\ &= w_m - \frac{\lambda \mathbf{R}_m^{-1} \mathbf{p}}{(1-\lambda)} \cdot \frac{1}{\lambda \mathbf{p}^T \mathbf{R}_m^{-1} \mathbf{p} + (1-\lambda)} ((1-\lambda) \mathbf{p}^T w_m + \lambda \mathbf{p}^T \mathbf{R}_m^{-1} \mathbf{p} y_{m+1}) + \frac{\lambda}{1-\lambda} \mathbf{R}_m^{-1} \mathbf{p} y_{m+1} \\ &= w_m - \frac{\lambda \mathbf{R}_m^{-1} \mathbf{p}}{(1-\lambda)} \cdot \frac{1}{\lambda \mathbf{p}^T \mathbf{R}_m^{-1} \mathbf{p} + (1-\lambda)} ((1-\lambda) \mathbf{p}^T w_m + \lambda \mathbf{p}^T \mathbf{R}_m^{-1} \mathbf{p} y_{m+1}) \\ &\quad + \frac{\lambda}{(1-\lambda)} \cdot \frac{\lambda \mathbf{p}^T \mathbf{R}_m^{-1} \mathbf{p} + (1-\lambda)}{\lambda \mathbf{p}^T \mathbf{R}_m^{-1} \mathbf{p} + (1-\lambda)} \mathbf{R}_m^{-1} \mathbf{p} y_{m+1}, \text{ since } \mathbf{p}^T \mathbf{R}_m^{-1} \mathbf{p} \text{ is a scalar, we can put } \mathbf{R}_m^{-1} \mathbf{p} \text{ to the right} \\ &= w_m + \frac{\lambda}{\lambda \mathbf{p}^T \mathbf{R}_m^{-1} \mathbf{p} + (1-\lambda)} \cdot \frac{1}{(1-\lambda)} (- (1-\lambda) \mathbf{R}_m^{-1} \mathbf{p} \mathbf{p}^T w_m - \lambda \mathbf{R}_m^{-1} \mathbf{p} \mathbf{p}^T \mathbf{R}_m^{-1} \mathbf{p} y_{m+1} \\ &\quad + \lambda \mathbf{R}_m^{-1} \mathbf{p} \mathbf{p}^T \mathbf{R}_m^{-1} \mathbf{p} y_{m+1} + (1-\lambda) \mathbf{R}_m^{-1} \mathbf{p} y_{m+1}) \\ &= w_m + \frac{\lambda}{\lambda \mathbf{p}^T \mathbf{R}_m^{-1} \mathbf{p} + (1-\lambda)} \cdot \frac{1}{(1-\lambda)} (- (1-\lambda) \mathbf{R}_m^{-1} \mathbf{p} \mathbf{p}^T w_m + (1-\lambda) \mathbf{R}_m^{-1} \mathbf{p} y_{m+1}) \end{aligned}$$

Thus, the final recursive solution for the weight vector \mathbf{w}_{m+1} is given by

$$\mathbf{w}_{m+1} = \mathbf{w}_m + \frac{\lambda \mathbf{R}_m^{-1} \mathbf{p} (y_{m+1} - \mathbf{p}^T \mathbf{w}_m)}{\lambda \mathbf{p}^T \mathbf{R}_m^{-1} \mathbf{p} + (1-\lambda)}$$

Here, we note that λ controls the strength of update with respect to the accumulated solution with $\lambda \rightarrow 1$ having the strongest weight for newly arrived sample. When $\lambda = 0.5$ in , we have the following regularized recursive least squares solution:

$$w_{m+1} = w_m + \frac{\mathbf{R}_m^{-1} \mathbf{p} (y_{m+1} - \mathbf{p}^T w_m)}{\mathbf{p}^T \mathbf{R}_m^{-1} \mathbf{p} + 1}.$$

Chapter 11

Logistic Regression

11.1 Logistic Regression

Whereas linear regression maps inputs to \mathbb{R} , logistic regression predicts values in the bounded interval $[0, 1]$, letting us interpret the outputs as probabilities. We therefore pass the linear score through the *sigmoid* (logistic) function

$$\sigma(z) = \frac{1}{1 + \exp(-z)}, \quad z = \beta^\top \mathbf{x}.$$

The class probabilities are

$$P\{Y = 1 | X\} = \sigma(z), \quad P\{Y = 0 | X\} = 1 - \sigma(z).$$

We are going to use the Maximum Likelihood Estimation (MLE) to find the best parameter. The MLE is a method used in statistics to estimate the parameters of a probability distribution by maximizing a likelihood function. Essentially, MLE finds the parameter values that make the observed data most probable.

Let's first define the Likelihood Function: Based on the probability distribution of the data, write down the likelihood function. For a set of observations $X = \{x_1, x_2, \dots, x_n\}$, and a parameter β , the likelihood function $L(\beta; X)$ represents the probability of observing the given data under the parameter β .

Since the likelihood can be a product of probabilities, it might be easier to work with the natural logarithm of the likelihood function, called the log-likelihood. This transforms the product into a sum, simplifying the computation:

$$\ell(\beta; X) = \log L(\beta; X)$$

Finally, to find the optimal parameter, find the parameter value $\hat{\beta}$ that maximizes the log-likelihood function.

For a dataset with n observations, the likelihood function is the product of the probabilities of observing the given outcomes. For logistic regression, the likelihood function $L(\beta; X, Y)$ is given by:

$$L(\beta; X, Y) = \prod_{i=1}^n P(Y_i | X_i, \beta)$$

Given the binary nature of Y , this can be written as:

$$L(\beta; X, Y) = \prod_{i=1}^n [P(Y_i = 1|X_i, \beta)]^{Y_i} [P(Y_i = 0|X_i, \beta)]^{1-Y_i}$$

Substituting the logistic function, we get:

$$L(\beta; X, Y) = \prod_{i=1}^n \left[\frac{1}{1 + \exp(-X_i\beta)} \right]^{Y_i} \left[1 - \frac{1}{1 + \exp(-X_i\beta)} \right]^{1-Y_i}$$

The negative log-likelihood for logistic regression is then given by

$$\text{NLL}(\beta) = - \sum_{i=1}^N y_i \ln \sigma(\beta^T \mathbf{x}) + (1 - y_i) \ln(1 - \sigma(\beta^T \mathbf{x})).$$

This is also called **cross-entropy** error function.

To compute the derivative of NLL, we first need to know the following tricks:

- The derivative of $\ln(x)$:

$$\frac{\partial}{\partial x} \ln(x) = \frac{1}{x}.$$

- The derivative of the sigmoid is given by:

$$\frac{\partial \sigma(z)}{\partial z} = \sigma(z)(1 - \sigma(z)).$$

- Finally, the chain rule of derivative. Suppose we are computing the derivative of a composite function $f(x) = u(v(x))$. The derivative of $f(x)$ is the derivative of $u(x)$ with respect to $v(x)$ times the derivative of $v(x)$ with respect to x .

$$\frac{\partial f}{\partial x} = \frac{\partial u}{\partial v} \frac{\partial v}{\partial x}$$

The derivative of the loss function w.r.t., a single weight w_j is given by

$$\begin{aligned} \frac{\partial \mathcal{L}}{\partial w_j} &= \frac{\partial}{\partial w_j} - [\mathbf{y} \ln \sigma(\mathbf{w}\mathbf{x} + \mathbf{b}) + (1 - \mathbf{y}) \ln(1 - \sigma(\mathbf{w}\mathbf{x} + \mathbf{b}))] \\ &= -[\frac{\partial}{\partial w_j} y \ln \sigma(wx + b) + \frac{\partial}{\partial w_j} (1 - y) \ln(1 - \sigma(wx + b))] \\ &= -\frac{y}{\sigma(wx + b)} \frac{\partial}{\partial w_j} \sigma(wx + b) - \frac{1 - y}{1 - \sigma(wx + b)} \frac{\partial}{\partial w_j} [1 - \sigma(wx + b)] \\ &= -\left[\frac{y}{\sigma(wx + b)} - \frac{1 - y}{1 - \sigma(wx + b)} \right] \frac{\partial}{\partial w_j} \sigma(wx + b) \\ &= -\left[\frac{y - \sigma(wx + b)}{\sigma(wx + b)[1 - \sigma(wx + b)]} \right] \sigma(wx + b)[1 - \sigma(wx + b)] \frac{\partial \sigma(wx + b)}{\partial w_j} \\ &= -\left[\frac{y - \sigma(wx + b)}{\sigma(wx + b)[1 - \sigma(wx + b)]} \right] \sigma(wx + b)[1 - \sigma(wx + b)] x_j \\ &= -(y - \sigma(wx + b)) x_j \\ &= (\sigma(wx + b) - y) x_j. \end{aligned}$$

⁰ $\mathbb{I}(y_i = 1) = y_i$, because $y_i \in \{0, 1\}$ is a binary variable

Thus, the gradient of the cost function $J(\boldsymbol{\theta})$ with respect to the weights $\boldsymbol{\theta}$ is:

$$\nabla_{\boldsymbol{\theta}} J(\boldsymbol{\theta}) = \frac{1}{m} \mathbf{X}^T (\sigma(\mathbf{X}\boldsymbol{\theta}) - \mathbf{y})$$

This gradient is used in gradient descent to update the weights $\boldsymbol{\theta}$:

$$\boldsymbol{\theta} := \boldsymbol{\theta} - \alpha \nabla_{\boldsymbol{\theta}} J(\boldsymbol{\theta}),$$

where α is the learning rate.

11.1.1 Another Interpretation

Logistic regression models *logits* (log odds) through a linear model. For binary data, the goal is to model the probability p that one of two outcomes occurs. Recall that an ordinary linear regression model is not bounded. Thus, we will pass a linear model through a sigmoid function, which is also known as logistic function.

$$\sigma(z) = \frac{1}{1 + \exp^{-z}},$$

where $z = wx + b$. The sigmoid function has the property

$$\sigma(-x) = 1 - \sigma(x).$$

The z is often called the logit. Note that the inverse of the sigmoid is the log of the odds ratio $\frac{p}{1-p}$.

The logit function is $\log \frac{p}{1-p}$, which varies between $-\infty$ and $+\infty$ as p varies between 0 and 1.

$$\log \frac{p}{1-p} = w_0 x_0 + w_1 x_1 + \cdots + w_n x_n.$$

Note that the logistic regression model assumes that the log-odds (also called the *logit*) of an observation y can be expressed as a linear function of the predictor variables. In this context, the logit function is referred to as the *link* function because it “links” the probability of an event to a linear combination of the predictors. The logit provides information about the ratio of two outcomes. For example, suppose an event has a probability of occurring of 60%. If a certain factor makes the event twice as likely to occur, the logit can be used to quantify this by comparing the probability of the event occurring to the probability of it not occurring.

Chapter 12

Bayesian Regression

12.1 Bayesian Regression

Chain rule: $P(A, B|C) = P(A|B, C)P(B|C)$

$$\begin{aligned} P(\text{heads} | D) &= \int_{\theta} P(\text{heads}, \theta | D) d\theta \\ &= \int_{\theta} P(\text{heads} | \theta, D) P(\theta | D) d\theta \\ &= \int_{\theta} \theta P(\theta | D) d\theta \\ &= E[\theta | D] \\ &= \frac{n_H + \alpha}{n_H + \alpha + n_T + \beta} \end{aligned}$$

12.2 Bayesian Regression

Bayesian regression is derived by combining the concepts of linear regression with Bayesian probability. In this note, we will derive Bayesian linear regression step-by-step, starting from a basic linear model and incorporating Bayesian inference.

12.3 Linear Regression Model

We assume a simple linear model:

$$y = X\beta + \epsilon,$$

where:

- y is the vector of observed responses (dependent variable),
- X is the matrix of observed features (independent variables),

- β is the vector of regression coefficients,
- $\epsilon \sim \mathcal{N}(0, \sigma^2 I)$ is the noise term (normally distributed with mean 0 and variance σ^2).

This gives the likelihood function:

$$p(y|X, \beta, \sigma^2) = \mathcal{N}(y|X\beta, \sigma^2 I)$$

In the Bayesian framework, we treat the regression coefficients β as random variables with a prior distribution. Bayesian regression involves the following key elements:

- **Prior:** $p(\beta)$, which represents our belief about β before observing the data.
- **Likelihood:** $p(y|X, \beta)$, the likelihood of observing the data given the regression coefficients.
- **Posterior:** $p(\beta|y, X)$, which updates the prior belief based on the observed data.

According to Bayes' theorem, the posterior is:

$$p(\beta|y, X) = \frac{p(y|X, \beta)p(\beta)}{p(y|X)}$$

The term $p(y|X)$ is the marginal likelihood (evidence), which acts as a normalizing constant to ensure that the posterior is a valid probability distribution.

12.3.1 Prior on β

A common choice for the prior distribution on β is a Gaussian distribution:

$$p(\beta) = \mathcal{N}(\beta|\mu_0, \Sigma_0),$$

where μ_0 is the prior mean (our prior belief about β), and Σ_0 is the prior covariance matrix (representing the uncertainty in our prior belief).

12.3.2 Posterior Distribution

The posterior distribution of β is obtained by combining the likelihood with the prior using Bayes' theorem. The likelihood and prior are given by

$$\begin{aligned} p(y|X, \beta) &= \mathcal{N}(y|X\beta, \sigma^2 I) \\ p(\beta) &= \mathcal{N}(\beta|\mu_0, \Sigma_0). \end{aligned}$$

Since the posterior is proportional to the product of the prior and the likelihood (and applying properties of Gaussian distributions), we obtain the posterior distribution and it is also Gaussian:

$$p(\beta|y, X) = \mathcal{N}(\beta|\mu_n, \Sigma_n),$$

where:

- The **posterior mean** μ_n is given by:

$$\mu_n = \Sigma_n \left(\Sigma_0^{-1} \mu_0 + \frac{1}{\sigma^2} X^T y \right)$$

- The **posterior covariance** Σ_n is given by:

$$\Sigma_n = \left(\Sigma_0^{-1} + \frac{1}{\sigma^2} X^T X \right)^{-1}$$

We can find the best β by computing the highest posterior probability and it is called *maximum a posterior* (MAP) approach.

12.3.3 Prediction

Once we have the posterior distribution $p(\beta|y, X)$, we can make predictions on new data X_* .

The predictive distribution for the new output y_* given new input X_* is obtained by integrating over the posterior distribution of β :

$$p(y_*|X_*, y, X) = \int p(y_*|X_*, \beta)p(\beta|y, X)d\beta$$

Since both the likelihood and posterior are Gaussian, the predictive distribution is also Gaussian:

$$p(y_*|X_*, y, X) = \mathcal{N}(y_*|X_*\mu_n, X_*\Sigma_n X_*^T + \sigma^2 I)$$

In Bayesian regression, we incorporate prior beliefs about the regression coefficients through the prior distribution $p(\beta)$, and update this belief after observing data to obtain the posterior distribution $p(\beta|y, X)$. This method provides not only point estimates of the regression coefficients but also quantifies the uncertainty in these estimates through the posterior distribution.

Part III

Kernel Methods

Chapter 13

Introduction to Kernel Methods

13.1 Introduction

In machine learning a *kernel* is a function that measures similarity between two objects and, at the same time, implicitly maps those objects into (possibly very-high-dimensional) feature space without ever computing the coordinates of that space directly.

A basic idea of kernel machine learning algorithms is that we can make ML models non-linear by applying basis function (feature) transformations on the input space. For an input vector $\mathbf{x} \in \mathbb{R}^d$, do transformation $\mathbf{x} \rightarrow \phi(\mathbf{x}) \in \mathbb{R}^D$, where $D \gg d$.

13.2 Kernel Trick

Kernel trick aims to avoid computing explicit feature transformations $\phi(\mathbf{x})$ by changing the entire ML algorithm to use only *inner products* of feature transformations $\langle \phi(\mathbf{x}), \phi(\mathbf{x}') \rangle$ which are computed directly using a *kernel* (rather than the feature transformations themselves.). Note that even the RBF kernel also can be expressed as an inner product.

This is contrast to many other ML algorithms that require explicit transformation of input data into feature vectors using a feature map, kernel methods directly compute similarity using a user-specified kernel function.

In other words, the terminology *kernel trick* simply refers to taking any algorithm that depends on the data only through inner products and replacing each inner product $\langle \mathbf{x}, \mathbf{x}' \rangle$ by a kernel value $k(\mathbf{x}, \mathbf{x}')$.

13.2.1 From Feature Transformations to Kernels

Instead of computing $\phi(\mathbf{x}_i), \forall i$, we pre-compute $\langle \phi(\mathbf{x}_i), \phi(\mathbf{x}_j) \rangle$ for all \mathbf{x}_i and \mathbf{x}_j . Then we store the values in an $n \times n$ matrix K where $K_{ij} = \langle \phi(\mathbf{x}_i), \phi(\mathbf{x}_j) \rangle$. K is called the *kernel matrix* or *Gram matrix*.

Pre-computing the inner product can be more efficient:

Approach	What we actually compute	Cost per evaluation
Explicit map	build $\varphi(x) \in \mathbb{R}^D$ then take $\langle \varphi(x), \varphi(z) \rangle$	$\mathcal{O}(D)$ (often gigantic)
Kernel trick	evaluate $K(x, z) = \langle \varphi(x), \varphi(z) \rangle$ directly	$\mathcal{O}(d)$, usually $d \ll D$

For instance, one might map $x \rightarrow (1, x, x^2)$ for quadratic features, $x \rightarrow (1, x, x^2, x^3)$ for cubic features, and so on.

$$y_t = \begin{cases} +1 & x_1^2 + x_2^2 \leq 1 \\ -1 & \text{otherwise} \end{cases}$$

For $\mathbf{x}[x_1, x_2]^T$, let $\phi(\mathbf{x}) = (x_1^2, \sqrt{2}x_1x_2, x_2^2) \in \mathbb{R}^3$:

$$\begin{aligned} K(\mathbf{x}, \mathbf{x}') &= x_1^2 x_1'^2 + \sqrt{2}x_1x_2\sqrt{2}x_1'x_2' + x_2^2 x_2'^2 \\ &= (x_1x_1' + x_2x_2')^2 \\ &= \langle \mathbf{x}, \mathbf{x}' \rangle^2 \end{aligned}$$

13.3 Kernels

Formally, we can define that the kernel is the inner product of a fixed feature space mapping $\phi(\mathbf{x})$

$$k(\mathbf{x}, \mathbf{x}') = \langle \phi(\mathbf{x}), \phi(\mathbf{x}') \rangle.$$

Hence, a kernel is a function $k(\cdot, \cdot) : \mathcal{X} \times \mathcal{X} \rightarrow \mathbb{R}$. By the definition, a kernel function measures the similarity between data points. The kernel matrix is always an $n \times n$ matrix, whatever the nature of data. Also, it has a *modularity* between the choice of kernel function and the choice of learning algorithm. However, it tends to suffer from a poor scalability with respect to the dataset size.

Then, we can ask can any function be a kernel? The answer is no. To be a valid or well-defined kernel, it has to be *symmetric* (*i.e.*, $k(\mathbf{x}, \mathbf{x}') = k(\mathbf{x}', \mathbf{x})$ for all \mathbf{x} and \mathbf{x}') and *positive semi-definite*. Also the feature map ϕ is a mapping from the input space to a feature space (*i.e.*, Hilbert space).

13.4 Hilbert Spaces

In functional analysis, $f(\cdot)$ denotes a function itself, while $f(x)$ represents the specific value that function takes at the input x .

Consider a function f that takes an input $\mathbf{x} = (x_1, x_2)$:

$$f(\mathbf{x}) = f_1x_1 + f_2x_2 + f_3x_1x_2.$$

First we recall what is meant by a linear function. Given a vector space X over the reals, a function

$$f : X \longrightarrow \mathbb{R}$$

is linear if $f(\alpha\mathbf{x}) = \alpha f(\mathbf{x})$ and $f(\mathbf{x} + \mathbf{z}) = f(\mathbf{x}) + f(\mathbf{z})$, $\forall \mathbf{x}, \mathbf{z} \in X, \alpha \in \mathbb{R}$.

Inner product space A vector space X over the reals \mathbb{R} is an *inner product space* if there exists a real-valued symmetric bilinear map $\langle \cdot, \cdot \rangle$, that satisfies

$$\langle \mathbf{x}, \mathbf{x} \rangle \geq 0.$$

The bilinear map is known as the inner, dot or scalar product. Furthermore we will say the inner product is strict if

$$\langle \mathbf{x}, \mathbf{x} \rangle = 0 \iff \mathbf{x} = \mathbf{0}.$$

Given a strict inner product space, we can define a norm on the space X by

$$|\mathbf{x}|_2 = \sqrt{\langle \mathbf{x}, \mathbf{x} \rangle}.$$

The associated metric or distance between two vectors is defined as $d(\mathbf{x}, \mathbf{z}) = |\mathbf{x} - \mathbf{z}|_2$. For the vector space \mathbb{R}^n the standard inner product is given by

$$\langle \mathbf{x}, \mathbf{z} \rangle = \sum_{i=1}^n x_i z_i.$$

An inner product space is sometimes referred to as a *Hilbert space*, though most researchers require the additional properties of completeness and separability, as well as sometimes requiring that the dimension be infinite. We give a formal definition.

Hilbert Space A Hilbert Space \mathcal{F} is an inner product space with the additional properties that it is *separable* and *complete*. Completeness refers to the property that every Cauchy sequence $\{h_n\}_{n \geq 1}$ of elements of \mathcal{F} converges to an element $h \in \mathcal{F}$, where a Cauchy sequence is one satisfying the property that

$$\sup_{m > n} |h_n - h_m| \rightarrow 0, \text{ as } n \rightarrow \infty.$$

A space \mathcal{F} is separable if for any $\epsilon > 0$ there is a finite set of elements h_1, \dots, h_N of \mathcal{F} such taht for all $h \in \mathcal{F}$

$$\min_i |h_i - h| < \epsilon.$$

For instance, Let X be the set of all countable sequences of real numbers $\mathbf{x} = (x_1, \dots, x_n, \dots)$, such that the sum

$$\sum_{i=1}^{\infty} x_i^2 < \infty,$$

with the inner product between two sequences \mathbf{x} and \mathbf{y} defined by

$$\langle \mathbf{x}, \mathbf{y} \rangle = \sum_{i=1}^{\infty} x_i y_i.$$

This is the space known as L_2 .

- A vector space is just a collection of objects you can add and scale (e.g., $\mathbf{v} + \mathbf{w}, 3\mathbf{v}$) and all the usual algebra rules hold.

- An inner product $\langle \cdot, \cdot \rangle$ on a vector space is a generalisation of the dot product. It must satisfy
 - Bilinearity $\langle a\mathbf{v} + b\mathbf{w}, \mathbf{z} \rangle = a\langle \mathbf{v}, \mathbf{z} \rangle + b\langle \mathbf{w}, \mathbf{z} \rangle$ (and similarly in the second slot).
 - Symmetry $\langle \mathbf{v}, \mathbf{w} \rangle = \langle \mathbf{w}, \mathbf{v} \rangle$.
 - Positive-definiteness $\langle \mathbf{v}, \mathbf{v} \rangle \geq 0$ with equality only when $\mathbf{v} = 0$.
- A metric space is *complete* when every Cauchy sequence actually converges to a limit inside the space.
 - Analogy: rational numbers \mathbb{Q} are not complete as Cauchy sequences like $3, 3.1, 3.14, 3.141, \dots$ hover toward π , which is missing from \mathbb{Q} . The reals \mathbb{R} are complete.
 - Completeness is crucial for analysis: it guarantees limits, projections, series expansions (*e.g.*, Fourier), and many theorems work internally without you stepping outside the space.

13.5 Properties

Given a finite subset $S = \{\mathbf{x}_1, \dots, \mathbf{x}_l\}$ of an input space \mathbf{X} , a kernel $\kappa(\mathbf{x}, \mathbf{z})$ and a feature map ϕ into a feature space F satisfying

$$\kappa \langle \phi(\mathbf{x}), \phi(\mathbf{z}) \rangle,$$

let $\phi(S) = \{\phi(\mathbf{x}_1), \dots, \phi(\mathbf{x}_l)\}$ be the image of S under the map ϕ . Hence $\phi(S)$ is a subset of the inner product of space F . Then the kernel matrix \mathbf{K} of kernel evaluations between all pairs of elements of S is given by

$$\mathbf{K}_{ij} = \kappa \langle \phi(\mathbf{x}_i), \phi(\mathbf{x}_j) \rangle, \quad i, j = 1, \dots, l$$

Working in a kernel-defined feature space means that we are not able to explicitly represent points. For example the image of an input point \mathbf{x} is $\phi(\mathbf{x})$, but we do not have access to the components of this vector, only to the evaluation of inner products between this point and the images of other points. Despite this handicap there is a surprising amount of useful information that can be gleaned about $\phi(\mathbf{S})$.

For two data points the transformed kernel $\hat{\kappa}$ is given by

$$\begin{aligned} \kappa(\mathbf{x}, \mathbf{z}) &= \left\langle \hat{\phi}(\mathbf{x}), \hat{\phi}(\mathbf{z}) \right\rangle = \left\langle \frac{\phi(\mathbf{x})}{\|\phi(\mathbf{x})\|}, \frac{\phi(\mathbf{z})}{\|\phi(\mathbf{z})\|} \right\rangle = \frac{\phi(\mathbf{x})\phi(\mathbf{z})}{\|\phi(\mathbf{x})\|\|\phi(\mathbf{z})\|} \\ &= \frac{\kappa(\mathbf{x}, \mathbf{z})}{\sqrt{\kappa(\mathbf{x}, \mathbf{x})\kappa(\mathbf{z}, \mathbf{z})}} \end{aligned}$$

```

1 % original kernel matrix stored in variable K
2 % output uses the same variable K
3 % D is a diagonal matrix storing the inverse of the norms
4 import numpy as np
5
6 # Example kernel matrix (must be symmetric with non-negative diagonal entries
7 )
```

$$7 \mathbf{K} = \text{np.array}([[4.0, 2.0, 2.0],$$

```

8      [2.0, 3.0, 0.5],
9      [2.0, 0.5, 1.0]])
10
11 D = np.diag(1.0 / np.sqrt(np.diag(K)))
12
13 # Normalize the kernel
14 K = D @ K @ D           # same variable name as in MATLAB code

```

We can also evaluate the norms of linear combinations of images in the feature space. For example, we have

$$\begin{aligned}
\left\| \sum_{i=1}^l \alpha_i \phi(\mathbf{x}_i) \right\|^2 &= \left\langle \sum_{i=1}^l \alpha_i \phi(\mathbf{x}_i), \sum_{i=j}^l \alpha_j \phi(\mathbf{x}_j) \right\rangle \\
&= \sum_{i=1}^l \alpha_i \sum_{i=j}^l \alpha_j \langle \phi(\mathbf{x}), \phi(\mathbf{z}) \rangle \\
&= \sum_{i,j=1}^l \alpha_i \alpha_j \kappa(\mathbf{x}, \mathbf{z}).
\end{aligned}$$

A special case of the norm is the length of the line joining two images $\phi(\mathbf{x})$ and $\phi(\mathbf{z})$, which can be computed as

13.6 Kernel Regression

Recall that

$$\begin{aligned}
(\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I}) \boldsymbol{\theta} &= (\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I})(\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I})^{-1} \mathbf{X}^T \mathbf{y} \\
(\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I}) \boldsymbol{\theta} &= \mathbf{X}^T \mathbf{y} \\
\boldsymbol{\theta} &= \lambda^{-1} \mathbf{I}(\mathbf{X}^T \mathbf{y} - \mathbf{X}^T \mathbf{X} \boldsymbol{\theta}) \\
&= \mathbf{X}^T \underbrace{\lambda^{-1}(\mathbf{y} - \mathbf{X} \boldsymbol{\theta})}_{=\alpha} \\
&= \mathbf{X}^T \alpha
\end{aligned}$$

This tells us that $\boldsymbol{\theta}$ is a linear combination of \mathbf{X} with α . Then,

$$\begin{aligned}
\lambda \alpha &= (\mathbf{y} - \mathbf{X} \boldsymbol{\theta}) \\
&= (\mathbf{y} - \mathbf{X} \mathbf{X}^T \alpha) \\
\mathbf{y} &= (\mathbf{X} \mathbf{X}^T \alpha + \lambda \alpha) \\
\alpha &= (\mathbf{X} \mathbf{X}^T + \lambda)^{-1} \mathbf{y} \\
&= (\mathbf{G} + \lambda)^{-1} \mathbf{y}.
\end{aligned}$$

$$\hat{\boldsymbol{\theta}} = \mathbf{X}^T (\mathbf{X} \mathbf{X}^T + \lambda \mathbf{I})^{-1} \mathbf{y}.$$

By substituting into $\hat{y}(\mathbf{x}') = \hat{\boldsymbol{\theta}}^T \mathbf{x}' = (\mathbf{x}')^T \hat{\boldsymbol{\theta}}$ gives

$$\hat{y}(\mathbf{x}') = (\mathbf{x}')^T \hat{\boldsymbol{\theta}} = (\mathbf{x}')^T \mathbf{X}^T (\mathbf{X} \mathbf{X}^T + \lambda \mathbf{I})^{-1} \mathbf{y}.$$

The prediction depends on the data only through inner products, since

$$(\mathbf{x}')^T \mathbf{X}^T = \begin{bmatrix} \langle \mathbf{x}', \mathbf{x}_1 \rangle \\ \vdots \\ \langle \mathbf{x}', \mathbf{x}_n \rangle \end{bmatrix}, \quad \mathbf{X} \mathbf{X}^T = \begin{bmatrix} \langle \mathbf{x}_1, \mathbf{x}_1 \rangle & \cdots & \langle \mathbf{x}_1, \mathbf{x}_n \rangle \\ \vdots & \ddots & \vdots \\ \langle \mathbf{x}_n, \mathbf{x}_1 \rangle & \cdots & \langle \mathbf{x}_n, \mathbf{x}_n \rangle \end{bmatrix},$$

The kernel trick replaces the inner products by kernel evaluations, which can be represented as follows:

$$\hat{y}(\mathbf{x}') = k(\mathbf{x}') (K + \lambda \mathbf{I})^{-1} \mathbf{y},$$

where

$$k(\mathbf{x}') = \begin{bmatrix} k(\mathbf{x}', \mathbf{x}_1) \\ \vdots \\ k(\mathbf{x}', \mathbf{x}_n) \end{bmatrix} \quad K = \begin{bmatrix} k(\mathbf{x}_1, \mathbf{x}_1) & \cdots & k(\mathbf{x}_1, \mathbf{x}_n) \\ \vdots & \ddots & \vdots \\ k(\mathbf{x}_n, \mathbf{x}_1) & \cdots & k(\mathbf{x}_n, \mathbf{x}_n) \end{bmatrix},$$

- The estimate \hat{y} is a weighted sum of the previously observed outputs, \mathbf{y} . The more similar \mathbf{x}' is to the corresponding \mathbf{x}_t (*i.e.*, the higher value of $k(\mathbf{x}, \mathbf{x}')$), the more weight is given to that \mathbf{y} .
- Note that we have turned the original learning problem that had d parameters into one that has n parameters (*i.e.*, the number of data samples).

13.6.1 Non-Parametric Regression

Kernel methods corresponding to infinite-dimensional $\phi(\mathbf{x})$ can usually be considered as **non-parametric**. For instance, 1-D case can be considered as a joining the dots by interpolation.

13.7 From Feature Transformations to Kernels

Instead of computing $\phi(\mathbf{x}_i), \forall i$, we pre-compute $\langle \phi(\mathbf{x}_i), \phi(\mathbf{x}_j) \rangle$ for all \mathbf{x}_i and \mathbf{x}_j . Then we store the values in an $n \times n$ matrix K where $K_{ij} = \langle \phi(\mathbf{x}_i), \phi(\mathbf{x}_j) \rangle$. K is called the kernel matrix or Gram matrix.

Computing $\phi(\mathbf{x}_i)$ is $\mathcal{O}(2^d) \times n = \mathcal{O}(2^d n)$. On the other hand, pre-computing the inner products can be more efficient as follows:

$$\begin{aligned} \langle \phi(\mathbf{x}), \phi(\mathbf{x}') \rangle &= \phi(\mathbf{x})^T \phi(\mathbf{x}') \\ &= \prod_{k=1}^d (1 + x_k x'_k). \end{aligned}$$

This takes $\mathcal{O}(d)$ and we can compute the entire kernel matrix K in $\mathcal{O}(dn^2)$.

Note that all kernel methods are non-parametric models as we need to keep training data to be able to compute the kernel values between new test inputs and the training inputs.

13.8 Kernels

$k(\cdot, \cdot)$ is a valid or well-defined kernel, if the function $k(\mathbf{x}, \mathbf{x}')$ is both

- Symmetric: $k(\mathbf{x}, \mathbf{x}') = k(\mathbf{x}', \mathbf{x})$ for all \mathbf{x}, \mathbf{x}' .
- Positive semi-definite: $k(\cdot, \cdot)$ is PSD if for all finite subsets $\{\mathbf{x}_1, \dots, \mathbf{x}_m\}$, $\mathbf{x}_i \in \mathcal{X}$, K is a PSD matrix.

Note that

- A matrix $A \in \mathbb{R}^{m \times m}$ is PSD iff $\forall \mathbf{q} \in \mathbb{R}^m$ the following holds: $\mathbf{q}^T A \mathbf{q} \geq 0$.
- A symmetric matrix $A \in \mathbb{R}^{m \times m}$ is PSD iff all eigenvalues λ of A are non-negative.
- A symmetric matrix A is PSD if all its upper left sub-matrices have non-negative determinants.

For symmetric matrices all of the above are equivalent.

Some of the most popular kernels are

- Linear kernel: $k(\mathbf{x}, \mathbf{x}') = \mathbf{x}^T \mathbf{x}'$
- Polynomial kernel: $k(\mathbf{x}, \mathbf{x}') = (1 + \mathbf{x}^T \mathbf{x}')^{\tilde{d}}$, where \tilde{d} is the degree of the polynomial

Note that some kernels such as the RBF kernel have an infinite-dimensional feature space transformation. Hence, it is impossible to compute the feature space transformation explicitly. The RBF kernel is given by

$$k(\mathbf{x}, \mathbf{x}') = \exp(-\gamma \|\mathbf{x} - \mathbf{x}'\|^2)$$

Let, $\gamma = \frac{1}{2}$,

$$\begin{aligned} k(\mathbf{x}, \mathbf{x}') &= \exp\left(-\frac{1}{2}\|\mathbf{x} - \mathbf{x}'\|^2\right) \\ &= \exp\left[-\frac{1}{2}\langle \mathbf{x} - \mathbf{x}', \mathbf{x} - \mathbf{x}' \rangle\right] \\ &= \exp\left[-\frac{1}{2}\langle \mathbf{x}, \mathbf{x} - \mathbf{x}' \rangle - \langle \mathbf{x}', \mathbf{x} - \mathbf{x}' \rangle\right] \\ &= \exp\left[-\frac{1}{2}(\|\mathbf{x}\|^2 + \|\mathbf{x}'\|^2 - 2\langle \mathbf{x}, \mathbf{x}' \rangle)\right] \\ &= C \exp\langle \mathbf{x}, \mathbf{x}' \rangle \\ &= C \sum_{n=0}^{\infty} \frac{\langle \mathbf{x}, \mathbf{x}' \rangle^n}{n!} \\ &= C \sum_{n=0}^{\infty} \frac{k_{\text{poly}}(\mathbf{x}, \mathbf{x}')}{n!} \end{aligned}$$

The second to last step is by Taylor expansion of e^x .

13.8.1 Some Intuitions

Let $y = f(\mathbf{x}) + \epsilon$. Then f does not vary a lot if \mathbf{x} , \mathbf{x}' are close enough. This can be modeled by the co-variance of the y 's.

Let's represent the co-variance of the outputs in terms of the co-variance of the inputs.

$$\text{cov}(y, y') = k(\mathbf{x}, \mathbf{x}') + \underbrace{\sigma_n^2 \delta_{\mathbf{x}, \mathbf{x}'}}_{\text{noise term}} .$$

As y is a function values of $f(\mathbf{x})$, the kernel function can be viewed as an extension of the covariance matrix Σ of random vectors to the covariance of (random) functions.

Chapter 14

Gaussian Process

14.1 Introduction

A *Gaussian process* (GP) is a probability distribution over possible functions that fit a set of points fully specified by a mean and covariance function (posterior over f , function space view). It is a generalization of the Gaussian probability distribution. Whereas a probability distribution describes random variables which are scalars or vectors (for multivariate distributions), a stochastic process governs the properties of functions.

It means that GPs don't give you a single function to fit data. They will give you an entire distribution of possible functions. More formally, GPs define **distributions over functions** $f(x)$ of which the distribution is defined by a mean function $m(\mathbf{x}) = \mathbb{E}[f(\mathbf{x})]$ and positive definite covariance function $k(\mathbf{x}, \mathbf{x}') = \mathbb{E}[(f(\mathbf{x}) - m(\mathbf{x}))(f(\mathbf{x}') - m(\mathbf{x}'))]$, which tells us how points are related to each other (*i.e.*, *kernel*). Thus, it is a distribution over functions whose shape is defined by the kernel:

$$f(\mathbf{x}) \sim \mathcal{GP}(m(\mathbf{x}), k(\mathbf{x}, \mathbf{x}')).$$

Let's say we have points then GPs assume that the function values of the points follow a multivariate Gaussian distribution. For example, we have observations $\mathbf{X} = \{x_1, \dots, x_5\}$, then we have

$$\begin{bmatrix} f(x_1) \\ f(x_2) \\ \vdots \\ f(x_5) \end{bmatrix} \sim \mathcal{N} \left(\begin{bmatrix} m(x_1) \\ m(x_2) \\ \vdots \\ m(x_5) \end{bmatrix}, \begin{bmatrix} k(x_1, x_1) & \cdots & k(x_1, x_5) \\ \vdots & \ddots & \vdots \\ k(x_5, x_1) & \cdots & k(x_5, x_5) \end{bmatrix} \right),$$

More formally, let's say we have estimated functions $\mathbf{f} = [f(x_1), \dots, f(x_n)]$ with these observations. Now say we have some new points \mathbf{X}_* where we want to predict $f(\mathbf{X}_*)$.

The joint distribution of \mathbf{f} and \mathbf{f}_* can be modeled as:

$$\begin{pmatrix} \mathbf{f} \\ \mathbf{f}_* \end{pmatrix} \sim \mathcal{N} \left(\begin{pmatrix} m(\mathbf{X}) \\ m(\mathbf{X}_*) \end{pmatrix}, \begin{pmatrix} \mathbf{K} & \mathbf{K}_* \\ \mathbf{K}_*^T & \mathbf{K}_{**} \end{pmatrix} \right),$$

where $m(X) = [m(x_1), \dots, m(x_n)]$ and

$$\begin{aligned}\mathbf{K} &= \kappa(\mathbf{X}, \mathbf{X}) \\ \mathbf{K}_* &= \kappa(\mathbf{X}, \mathbf{X}_*) \\ \mathbf{K}_{**} &= \kappa(\mathbf{X}_*, \mathbf{X}_*).\end{aligned}$$

The mean is assumed to be $(m(\mathbf{X}), m(\mathbf{X}_*)) = 0$.

While this equation describes the joint probability distribution $P(\mathbf{f}, \mathbf{f}_* | \mathbf{X}, \mathbf{X}_*)$ over \mathbf{f} and \mathbf{f}_* in regressions, we need the conditional distribution $P(\mathbf{f}_* | \mathbf{f}, \mathbf{X}, \mathbf{X}_*)$ for a prediction. The conditional distribution can be achieved by the Marginal and conditional distributions of MVN theorem:

Theorem 1 *Marginals and conditionals of an MVN: A random vector \mathbf{X} follows a multivariate normal distribution:*

$$\mathbf{X} \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$$

where \mathbf{X} can be partitioned as:

$$\mathbf{X} = \begin{bmatrix} \mathbf{X}_1 \\ \mathbf{X}_2 \end{bmatrix}$$

with corresponding partitions of the mean vector $\boldsymbol{\mu}$ and the covariance matrix $\boldsymbol{\Sigma}$:

$$\boldsymbol{\mu} = \begin{bmatrix} \boldsymbol{\mu}_1 \\ \boldsymbol{\mu}_2 \end{bmatrix}, \quad \boldsymbol{\Sigma} = \begin{bmatrix} \boldsymbol{\Sigma}_{11} & \boldsymbol{\Sigma}_{12} \\ \boldsymbol{\Sigma}_{21} & \boldsymbol{\Sigma}_{22} \end{bmatrix}, \quad \boldsymbol{\Lambda} = \boldsymbol{\Sigma}^{-1} = \begin{bmatrix} \boldsymbol{\Lambda}_{11} & \boldsymbol{\Lambda}_{12} \\ \boldsymbol{\Lambda}_{21} & \boldsymbol{\Lambda}_{22} \end{bmatrix},$$

Then, the marginal distribution of \mathbf{X}_1 and \mathbf{X}_2 are given by

$$\begin{aligned}\mathbf{X}_1 &\sim \mathcal{N}(\boldsymbol{\mu}_1, \boldsymbol{\Sigma}_{11}) \\ \mathbf{X}_2 &\sim \mathcal{N}(\boldsymbol{\mu}_2, \boldsymbol{\Sigma}_{22})\end{aligned}$$

Subsequently, the (posterior) conditional distribution of \mathbf{X}_1 given \mathbf{X}_2 is:

$$\mathbf{X}_1 | \mathbf{X}_2 = \mathbf{x}_2 \sim \mathcal{N}(\boldsymbol{\mu}', \boldsymbol{\Sigma}'),$$

where the conditional mean $\boldsymbol{\mu}_{1|2}$ and the conditional covariance $\boldsymbol{\Sigma}_{1|2}$ are given by:

$$\begin{aligned}\boldsymbol{\mu}' &= \boldsymbol{\mu}_1 + \boldsymbol{\Sigma}_{12} \boldsymbol{\Sigma}_{22}^{-1} (\mathbf{x}_2 - \boldsymbol{\mu}_2) \\ \boldsymbol{\Sigma}' &= \boldsymbol{\Sigma}_{11} - \boldsymbol{\Sigma}_{12} \boldsymbol{\Sigma}_{22}^{-1} \boldsymbol{\Sigma}_{21}\end{aligned}$$

Note that $\boldsymbol{\Sigma}'$ does not depend on \mathbf{x}_2 . It doesn't hold for general (non-Gaussian) random variables. By using the theorem, we get

$$\mathbf{f}_* | \mathbf{f}, \mathbf{X}, \mathbf{X}_* \sim \mathcal{N}(\mathbf{K}_*^T \mathbf{K}^{-1} \mathbf{f}, \mathbf{K}_{**} - \mathbf{K}_*^T \mathbf{K}^{-1} \mathbf{K}_*)$$

In practice, we typically have noisy estimations of target functions, $y = f(x) + \epsilon$. By assuming the i.i.d., Gaussian noise with variance σ_n^2 , the prior on these noisy observations then becomes $cov(y) = \mathbf{K} + \sigma_n^2 \mathbf{I}$. The joint distribution of the observed target values and the function values at the test locations under the prior as

$$\begin{bmatrix} \mathbf{y} \\ \mathbf{f}_* \end{bmatrix} \sim \mathcal{N} \left[\mathbf{0}, \begin{bmatrix} \mathbf{K} + \sigma_n^2 \mathbf{I} & \mathbf{K}_* \\ \mathbf{K}_*^T & \mathbf{K}_{**} \end{bmatrix} \right]$$

14.2 Regression using Gaussian Process

- A set of input points $X = \{x_1, x_2, \dots, x_n\}$ (*i.e.*, training data).
- A set of corresponding output values $y = \{y_1, y_2, \dots, y_n\}$.

We assume that the output values are generated by some unknown function $f(x)$ with some Gaussian noise ϵ :

$$y = f(X) + \epsilon, \quad \epsilon \sim \mathcal{N}(0, \sigma_n^2)$$

The goal is to predict the value of the function $f(x_*)$ at a new input point x_* .

We assume that the function $f(x)$ is drawn from a **Gaussian Process**:

$$f(x) \sim \mathcal{GP}(m(x), k(x, x')),$$

where:

- $m(x)$ is the mean function (usually assumed to be 0 for simplicity).
- $k(x, x')$ is the covariance function (or kernel), which encodes assumptions about the smoothness of the function.

The actual observations $y = f(x) + \epsilon$ include Gaussian noise ϵ , which is typically assumed to have zero mean and variance σ_n^2 .

To predict the function value $f_* = f(x_*)$ at a new point x_* , we use the **posterior distribution** over the function, conditioned on the observed data X and y . This leads to two main equations:

- Posterior Mean (Prediction):

$$\mathbb{E}[f_*] = K(X, x_*)^\top [K(X, X) + \sigma_n^2 I]^{-1} y$$

- Posterior Variance (Uncertainty):

$$\text{Var}[f_*] = K(x_*, x_*) - K(X, x_*)^\top [K(X, X) + \sigma_n^2 I]^{-1} K(X, x_*)$$

These equations give us the mean and variance of the predictive distribution for the function value at the new point x_* .

Example: First, choose a covariance function (*i.e.*, Kernel). The kernel function $k(x, x')$ is crucial in defining the relationship between different input points. Some commonly used kernels are:

Squared Exponential (RBF) Kernel:

$$k(x, x') = \sigma_f^2 \exp\left(-\frac{(x - x')^2}{2l^2}\right),$$

where l is the length scale and σ_f^2 is the signal variance. The kernel function governs the smoothness and behavior of the GP, so selecting an appropriate kernel is important.

Subsequently, compute the covariance matrices:

- Covariance Matrix for Training Points: $K(X, X)$

$$K(X, X) = \begin{bmatrix} k(x_1, x_1) & k(x_1, x_2) & \dots & k(x_1, x_n) \\ k(x_2, x_1) & k(x_2, x_2) & \dots & k(x_2, x_n) \\ \vdots & \vdots & \ddots & \vdots \\ k(x_n, x_1) & k(x_n, x_2) & \dots & k(x_n, x_n) \end{bmatrix}$$

- Covariance Between Training Points and Test Point: $K(X, x_*)$

$$K(X, x_*) = \begin{bmatrix} k(x_1, x_*) \\ k(x_2, x_*) \\ \vdots \\ k(x_n, x_*) \end{bmatrix}$$

- Covariance at the Test Point: $K(x_*, x_*) = k(x_*, x_*)$

Then, compute the posterior mean and variance by Using the posterior mean and posterior variance equations from earlier, compute:

- The mean prediction $\mathbb{E}[f_*]$ at the test point x_* .
- The uncertainty $\text{Var}[f_*]$ at x_* .

This gives the complete predictive distribution for the function value at x_* .

14.3 Time Complexity

One drawback of Gaussian processes is that it scales very badly with the number of observations N . It takes $O(N^3)$ time, since we need to compute the inverse of the kernel matrix.

14.4 Example

We observe two training points in \mathbb{R}^2 :

$$X = \{(0, 0), (1, 1)\}, \quad y = \{1, 2\}.$$

We want to predict the function value at a new test point:

$$x_* = (0, 1).$$

We use the squared exponential (RBF) kernel for 2D inputs:

$$k(x, x') = \exp\left(-\frac{\|x - x'\|^2}{2\ell^2}\right), \quad \ell = 1.$$

Assume zero mean and no noise ($\sigma_n^2 = 0$).

Compute covariance matrices Training covariance:

$$K(X, X) = \begin{bmatrix} k((0, 0), (0, 0)) & k((0, 0), (1, 1)) \\ k((1, 1), (0, 0)) & k((1, 1), (1, 1)) \end{bmatrix}.$$

Compute distances:

- $\|(0, 0) - (0, 0)\|^2 = 0 \Rightarrow k = 1.$
- $\|(0, 0) - (1, 1)\|^2 = 1^2 + 1^2 = 2 \Rightarrow k = e^{-1} = 0.3679.$
- $\|(1, 1) - (1, 1)\|^2 = 0 \Rightarrow k = 1.$

So:

$$K = \begin{bmatrix} 1 & 0.3679 \\ 0.3679 & 1 \end{bmatrix}.$$

Cross-covariance $K(X, x_*)$:

Distances:

- $\|(0, 0) - (0, 1)\|^2 = 1 \Rightarrow k = e^{-0.5} = 0.6065.$
- $\|(1, 1) - (0, 1)\|^2 = 1 \Rightarrow k = e^{-0.5} = 0.6065.$

So:

$$K(X, x_*) = \begin{bmatrix} 0.6065 \\ 0.6065 \end{bmatrix}.$$

Test covariance:

$$K(x_*, x_*) = 1.$$

Posterior mean

$$\mathbb{E}[f_*] = K(X, x_*)^\top K(X, X)^{-1}y.$$

Compute K^{-1} . For a 2×2 symmetric matrix:

$$K^{-1} = \frac{1}{1 - (0.3679)^2} \begin{bmatrix} 1 & -0.3679 \\ -0.3679 & 1 \end{bmatrix}.$$

Denominator: $1 - 0.1353 = 0.8647$. So:

$$K^{-1} \approx \begin{bmatrix} 1.156 & -0.425 \\ -0.425 & 1.156 \end{bmatrix}.$$

Now multiply:

$$K^{-1}y = \begin{bmatrix} 1.156 & -0.425 \\ -0.425 & 1.156 \end{bmatrix} \begin{bmatrix} 1 \\ 2 \end{bmatrix}.$$

$$= \begin{bmatrix} 1.156(1) - 0.425(2) \\ -0.425(1) + 1.156(2) \end{bmatrix} = \begin{bmatrix} 0.306 \\ 1.887 \end{bmatrix}.$$

Now compute mean:

$$\mathbb{E}[f_*] = [0.6065 \ 0.6065] \begin{bmatrix} 0.306 \\ 1.887 \end{bmatrix}.$$

$$= 0.6065(0.306 + 1.887) = 1.33.$$

Posterior mean:

$$\mathbb{E}[f(0, 1)] \approx 1.33.$$

Posterior variance

$$\text{Var}[f_*] = K(x_*, x_*) - K(X, x_*)^\top K(X, X)^{-1} K(X, x_*).$$

First compute:

$$\begin{aligned} v &= K^{-1}K(X, x_*) = \begin{bmatrix} 1.156 & -0.425 \\ -0.425 & 1.156 \end{bmatrix} \begin{bmatrix} 0.6065 \\ 0.6065 \end{bmatrix}. \\ &= \begin{bmatrix} (1.156 - 0.425)(0.6065) \\ (-0.425 + 1.156)(0.6065) \end{bmatrix} = \begin{bmatrix} 0.443 \\ 0.443 \end{bmatrix}. \end{aligned}$$

Now:

$$K(X, x_*)^\top v = [0.6065 \ 0.6065] \begin{bmatrix} 0.443 \\ 0.443 \end{bmatrix}.$$

$$= 0.6065(0.443 + 0.443) = 0.537.$$

So:

$$\text{Var}[f_*] = 1 - 0.537 = 0.463.$$

Posterior variance

$$\text{Std}[f(0, 1)] \approx \sqrt{0.463} \approx 0.68.$$

For training data $((0, 0), 1), ((1, 1), 2)$ and test point $(0, 1)$:

$$f(0, 1) \sim \mathcal{N}(1.33, 0.68^2).$$

Chapter 15

Support Vector Machine

Support Vector Machines (SVMs) are among the most effective and versatile tools in machine learning, widely used for various tasks. SVMs work by finding the optimal boundary, or hyperplane, that separates different classes of data with the maximum margin, making them highly reliable for classification, especially with complex datasets.

What truly sets SVMs apart is their ability to handle both linear and non-linear data through the *kernel trick*, allowing them to adapt to a wide range of problems with impressive accuracy. In this blog post, we'll delve into how SVMs work and gently explore the mathematical foundations behind their powerful performance.

15.1 Decision Boundary with Margin

A hyperplane(or decision surface) is used to separate data points belonging to different classes. The goal of SVM is to find the optimal separating hyperplane. However, what is the optimal separating hyperplanes? **The optimal hyperplane is the one which maximizes the distance from the hyperplane to the nearest data point of any class. Support vectors are the data points that lie closest to the hyperplane.** The distance is referred to as the *margin*. SVMs maximize the margin around the separating hyperplane.

The equation of a hyperplane in \mathbb{R}^p can be expressed as:

$$\mathbf{w} \cdot \mathbf{x} + b = 0.$$

Here, \mathbf{w} is the normal vector to the hyperplane. It is clear when we set $b = \mathbf{w} \cdot \mathbf{x}_0$

$$\mathbf{w}(\mathbf{x} - \mathbf{x}_0) = 0.$$

Let's consider a simple scenario, where training data is linearly separable:

$$\mathcal{D} = \{(\mathbf{x}_i, y_i) \mid \mathbf{x}_i \in \mathbb{R}^p, y_i \in \{-1, 1\}\}_{i=1}^N.$$

Then, we can build two supporting hyperplanes, which separate the data with no points between them:

- $H_1 : \mathbf{w} \cdot \mathbf{x} + b = 1$

- $H_2 : \mathbf{w} \cdot \mathbf{x} + b = -1$

All samples have to satisfy one of two constraints:

1. $\mathbf{w} \cdot \mathbf{x} + b \geq 1$

2. $\mathbf{w} \cdot \mathbf{x} + b \leq -1$

For instance, if we set $\mathbf{w} = (0.5, 0.5)$, $b = -3$, then $\mathbf{x} = (4, 5)$ will be at H_1 and $\mathbf{x} = (2, 1)$ at H_2 . Collectively, these constraints can be combined into a single expression:

$$y \cdot (\mathbf{w} \cdot \mathbf{x} + b) \geq 1.$$

To maximize the margin, we can consider a unit vector $\mathbf{u} = \frac{\mathbf{w}}{\|\mathbf{w}\|}$, which is perpendicular to the hyperplanes and a point \mathbf{x}_0 on H_2 . If we scale \mathbf{u} from \mathbf{x}_0 , we get $\mathbf{z} = \mathbf{x}_0 + \gamma\mathbf{u}$ and we assume that \mathbf{z} reaches at H_1 . Then, $\mathbf{w} \cdot \mathbf{z} + b = 1$. This is equivalent to

$$\begin{aligned} \mathbf{w} \cdot (\mathbf{x}_0 + \gamma\mathbf{u}) + b &= 1 \\ \mathbf{w}\mathbf{x}_0 + \mathbf{w}\gamma\frac{\mathbf{w}}{\|\mathbf{w}\|} + b &= 1 \\ \mathbf{w}\mathbf{x}_0 + \gamma\|\mathbf{w}\| + b &= 1 \\ \mathbf{w}\mathbf{x}_0 + b &= 1 - \gamma\|\mathbf{w}\| \end{aligned}$$

As \mathbf{x}_0 is on H_2 , we get $\mathbf{w}\mathbf{x}_0 + b = -1$. Finally, we obtain

$$\begin{aligned} -1 &= 1 - \gamma\|\mathbf{w}\| \\ \gamma &= \frac{2}{\|\mathbf{w}\|}. \end{aligned}$$

Note that the scaled unit vector $\gamma\mathbf{u}$'s magnitude is γ . Thus, the maximization of margin is equivalent to maximize γ (*i.e.*, margin). To maximize γ (*i.e.*, margin), we have to minimize $\|\mathbf{w}\|$. Thus, finding the optimal hyperplane reduces to solving the following optimization problem:

$$\begin{aligned} \min \|\mathbf{w}\|, \text{ subject to} \\ y_i(\mathbf{w} \cdot \mathbf{x}_i + b) \geq 1 \quad \forall i. \end{aligned}$$

Equivalently,

$$\begin{aligned} \min \frac{1}{2}\|\mathbf{w}\|^2, \text{ subject to} \\ y_i(\mathbf{w} \cdot \mathbf{x}_i + b) \geq 1 \quad \forall i. \end{aligned}$$

Now, we have *convex quadratic optimization problem*. The solution of this problem gives us the optimal hyperplane that maximizes the margin (*c.f.*, §15.3). However, in practice, the data may not be perfectly separable. To address this, we introduce a *soft margin* that allows for some misclassification. This is done by admitting small errors in classification and potentially using a more complex, *nonlinear decision boundary*, improving the generalization of the model.

15.1.1 Alternative Derivation

Consider a point \mathbf{x} . Let \mathbf{d} be the vector from a hyperplane (*i.e.*, $\mathbf{w}\mathbf{x} + b = 0$) to \mathbf{x} of minimum length. Let \mathbf{x}_0 be the projection of \mathbf{x} onto the hyperplane. Then,

$$\mathbf{x}_0 = \mathbf{x} - \mathbf{d}.$$

As \mathbf{d} is parallel to \mathbf{w} , so $\mathbf{d} = \alpha\mathbf{w}$ for some $\alpha \in \mathbb{R}$. Since \mathbf{x}_0 is on the hyperplane, $\mathbf{w}\mathbf{x}_0 + b = 0$. Thus,

$$\mathbf{w}\mathbf{x}_0 + b = \mathbf{w}(\mathbf{x} - \mathbf{d}) + b = \mathbf{w}(\mathbf{x} - \alpha\mathbf{w}) + b = 0.$$

Then, we get

$$\alpha = \frac{\mathbf{w}\mathbf{x} + b}{\mathbf{w}^T\mathbf{w}}.$$

The length of \mathbf{d} is given by

$$\|\mathbf{d}\|_2 = \sqrt{\alpha^2\mathbf{w}^T\mathbf{w}} = \frac{|\mathbf{w}\mathbf{x} + b|}{\sqrt{\mathbf{w}^T\mathbf{w}}} = \frac{|\mathbf{w}\mathbf{x} + b|}{\|\mathbf{w}\|_2}.$$

We can obtain the margin by choosing the support vector, which is the closest point to the hyperplane by

$$\gamma(\mathbf{w}, b) = \min_{\mathbf{x} \in \mathcal{D}} \frac{|\mathbf{w}\mathbf{x} + b|}{\|\mathbf{w}\|_2}.$$

Note that the margin and hyperplane are scale invariant:

$$\gamma(\beta\mathbf{w}, \beta b) = \gamma(\mathbf{w}, b).$$

15.2 Error Handling in SVM

In practice, it's unrealistic to expect a perfect separation of data, especially when the data is noisy or not linearly separable. To address this, we can **allow for some prediction errors while still striving to find an optimal decision boundary**.

One approach is to minimize the norm of the weight vector, while penalizing the number of errors N_e . The optimization problem can be formulated as follows:

$$\begin{aligned} & \min \frac{1}{2}\|\mathbf{w}\|^2 + C \cdot N_e, \text{ subject to} \\ & y_i(\mathbf{w} \cdot \mathbf{x}_i + b) \geq 1 \quad \forall i. \end{aligned}$$

Here, C is a regularization parameter that controls the trade-off between minimizing the weight vector and the number of errors. The penalty approach described here is known as *0-1 loss*, where all errors are treated equally. However, this approach is not commonly used. Instead, a more practical approach introduces a *slack variable* with *hinge loss*. The slack variable (ξ_j) measures the degree of misclassification or how much a point is violating the margin. This leads to the following problem:

$$\begin{aligned} & \min \frac{1}{2}\|\mathbf{w}\|^2 + C \sum_j \xi_j, \quad \text{subject to} \\ & \begin{cases} y_i(\mathbf{w} \cdot \mathbf{x}_i + b) \geq 1 - \xi_j \quad \forall i, \\ \xi_j \geq 0, \quad \forall j. \end{cases} \end{aligned}$$

Note that $\xi_j > 1$, when SVMs make prediction errors:

$$\xi_j = (1 - (\mathbf{w}\mathbf{x}_j + b) \cdot y_j)$$

Let's look at the new constraint. If some data points are misclassified, then $\xi_j > 1$ and $y_i(\mathbf{w} \cdot \mathbf{x}_i + b) \leq 0$. This approach is called **soft-margin SVM**. Lastly, how do we set C ? A large value of C places a higher penalty on errors, leading to a narrower margin but fewer misclassifications (*i.e.*, the SVM will try to classify all data points correctly), whereas a smaller value of C allows for a wider margin but potentially more misclassifications. The optimal value of C is typically chosen through cross-validation.

15.3 SVM Optimization: Lagrange Multipliers

15.3.1 Lagrange Multipliers

Consider the optimization problem:

$$\begin{aligned} & \min_{\mathbf{x}} f(\mathbf{x}) \\ & \text{subject to } g(\mathbf{x}) = 0. \end{aligned}$$

To find the minimum of f under the constraint $g(\mathbf{x})$, we use the method of *Lagrange multipliers*. The key idea is that at the optimal point, the gradient of $f(\mathbf{x})$ must be parallel to the gradient of $g(\mathbf{x})$. Mathematically, this condition is expressed as:

$$\nabla f(\mathbf{x}) = \lambda \nabla g(\mathbf{x}).$$

Example: Consider a simple 2D example where you want to minimize the function $f(x, y) = x^2 + y^2$, which represents a circle centered at the origin. This function increases as you move away from the origin, so the minimum is at the origin.

Now, consider the constraint: $g(x, y) = x + y - 1 = 0$. This constraint is a line that runs through the xy -plane. Our goal is to find the point on this line that minimizes $f(x, y)$.

A Lagrange multiplier is like a balancing factor that adjusts the direction and magnitude of your search along the constraint. As you move along the constraint line $g(x, y)$, λ ensures that the solution also respects the shape of the function $f(x, y)$ that you are trying to minimize. To solve the constraint optimization problem, we define the Lagrangian function:

$$\mathcal{L}(\mathbf{x}, \lambda) = f(\mathbf{x}) - \lambda g(\mathbf{x}).$$

To find the minimum, we take the partial derivatives of $\mathcal{L}(\mathbf{x}, \lambda)$ with respect to both \mathbf{x} and λ , and set them equal to zero.

15.3.2 SVM Optimization

Recall that we want to solve the following convex quadratic optimization problem:

$$\begin{aligned} & \min \frac{1}{2} \|\mathbf{w}\|^2, \quad \text{subject to} \\ & y_i(\mathbf{w} \cdot \mathbf{x}_i + b) \geq 1 \quad \forall i. \end{aligned}$$

The objective is to find the optimal hyperplane that maximizes the margin between two classes of data points.

We can reformulate this optimization problem using the method of Lagrange multipliers, which introduces a set of multipliers α_i (c.f., one for each constraint). The Lagrangian function for this problem is given by:

$$\mathcal{L}(\mathbf{w}, b, \alpha) = \frac{1}{2} \|\mathbf{w}\|^2 - \sum_{i=1}^N \alpha_i [y_i(\mathbf{w} \cdot \mathbf{x}_i + b) - 1]$$

15.3.3 Duality and the Lagrangian Problem

While we could attempt to solve the primal optimization problem directly, it is often more practical, especially for large datasets, to reformulate the problem using the duality principle. The dual form is advantageous because it depends only on the inner products of the data points, which allows the use of *kernel* methods for non-linear classification.

To find the solution to the primal problem, we solve the following problem:

$$\begin{aligned} & \max_{\mathbf{w}, b} \min_{\alpha} \mathcal{L}(\mathbf{w}, b, \alpha) \\ & \text{subject to } \alpha_i \geq 0, \forall i. \end{aligned}$$

Here, we maximize the Lagrangian with respect to the multipliers α_i , while minimizing with respect to the primal variables \mathbf{w} and b .

15.3.4 Handling Inequality Constraints with KKT Conditions

You may observe that the method of Lagrange multipliers is used for equality constraints. However, it can be extended to handle inequality constraints through the use of additional conditions known as the **Karush-Kuhn-Tucker (KKT) conditions**. These conditions ensure that the solution satisfies the necessary optimality criteria for problems with inequality constraints. For more details on the KKT conditions.

15.4 The Wolfe Dual Problem

The Lagrangian problem for SVM optimization involves N inequality constraints, where N is the number of training examples. This problem is typically tackled by using its *dual form*. The duality principle provides a powerful framework, stating that **an optimization problem can be approached from two perspectives**:

1. The *primal problem*, which in our context is a minimization problem.
2. The *dual problem*, which is a maximization problem.

An important aspect of duality is that **the maximum value of the dual problem is always less than or equal to the minimum value of the primal problem**. This relationship means that the dual problem **provides a lower bound to the solution of the primal problem**.

In the context of SVM optimization, we are dealing with a convex optimization problem. According to *Slater's condition*, which applies to problems with affine constraints, strong duality holds. Strong duality implies that the optimal values of the primal and dual problems are equal, meaning the maximum value of the dual problem equals the minimum value of the primal problem. This equality allows us to solve the dual problem instead of the primal problem, often leading to computational advantages.

Recall that we aim to solve the following optimization problem:

$$\mathcal{L}(\mathbf{w}, b, \alpha) = \frac{1}{2} \|\mathbf{w}\|^2 - \sum_{i=1}^N \alpha_i [y_i (\mathbf{w} \cdot \mathbf{x}_i + b) - 1]$$

The minimization problem involves solving the partial derivatives of \mathcal{L} with respect to \mathbf{w} and b and set them equal to zero:

$$\begin{aligned}\nabla_{\mathbf{w}} \mathcal{L}(\mathbf{w}, b, \alpha) &= \mathbf{w} - \sum_i \alpha_i y_i \mathbf{x}_i \\ \nabla_b \mathcal{L}(\mathbf{w}, b, \alpha) &= - \sum_i \alpha_i y_i\end{aligned}$$

Form the first equation, we obtain:

$$\mathbf{w} = \sum_{i=1}^m \alpha_i y_i \mathbf{x}_i$$

Next, we substitute the objective function with \mathbf{w} :

$$\begin{aligned}\mathbf{W}(\alpha, b) &= \frac{1}{2} \left(\sum_i \alpha_i y_i \mathbf{x}_i \right) \cdot \left(\sum_j \alpha_j y_j \mathbf{x}_j \right) - \sum_i \alpha_i \left[y_i \left(\left(\sum_j \alpha_j y_j \mathbf{x}_j \right) \cdot \mathbf{x}_i + b \right) - 1 \right] \\ &= \frac{1}{2} \left(\sum_i \sum_j \alpha_i \alpha_j y_i y_j \mathbf{x}_i \cdot \mathbf{x}_j \right) - \sum_i \alpha_i \left[y_i \left(\left(\sum_j \alpha_j y_j \mathbf{x}_j \right) \cdot \mathbf{x}_i + b \right) \right] + \sum_i \alpha_i \\ &= \frac{1}{2} \sum_i \sum_j \alpha_i \alpha_j y_i y_j \mathbf{x}_i \cdot \mathbf{x}_j - \sum_i \sum_j \alpha_i \alpha_j y_i y_j \mathbf{x}_i \cdot \mathbf{x}_j - \sum_i \alpha_i y_i b + \sum_i \alpha_i \\ &= \sum_i \alpha_i - \frac{1}{2} \sum_i \sum_j \alpha_i \alpha_j y_i y_j \mathbf{x}_i \cdot \mathbf{x}_j - \sum_i \alpha_i y_i b\end{aligned}$$

Note that we use two indices, i and j when substituting \mathbf{W} . This is clear when we consider a simple example. Let's say you have two data points:

$$\begin{aligned}\mathbf{x}_1, y_1 &= (1, 2), 1 \\ \mathbf{x}_2, y_2 &= (2, 1), 1\end{aligned}$$

Then,

$$\|\mathbf{w}\|^2 = \mathbf{w} \cdot \mathbf{w} = \underbrace{(\alpha_1 y_1 \mathbf{x}_1 + \alpha_2 y_2 \mathbf{x}_2)}_{\sum_i} \cdot \underbrace{(\alpha_1 y_1 \mathbf{x}_1 + \alpha_2 y_2 \mathbf{x}_2)}_{\sum_j}.$$

This simplification shows that the optimization problem can be reformulated purely in terms of the Lagrange multipliers α_i . Note that the term involving b can be removed by setting $b = 0$, simplifying our equation further:

$$\mathbf{W}(\alpha, b) = \sum_i \alpha_i - \frac{1}{2} \sum_i \sum_j \alpha_i \alpha_j y_i y_j (\mathbf{x}_i \cdot \mathbf{x}_j) \quad (15.1)$$

This expression is known as the **Wolfe dual Lagrangian function**. We have transformed the problem into one involving only the multipliers α_i , resulting in a quadratic programming problem, commonly referred to as the **Wolfe dual problem**:

$$\max_{\alpha} \mathbf{W}(\alpha, b) = \sum_i \alpha_i - \frac{1}{2} \sum_i \sum_j \alpha_i \alpha_j y_i y_j (\mathbf{x}_i \cdot \mathbf{x}_j)$$

subject to $\alpha_i \geq 0$ for any $i = 1, \dots, m$

$$\sum_{i=1}^m \alpha_i y_i = 0$$

Once we get the value of α , the optimal \mathbf{w} and b can be computed using

$$\alpha_i [y_i(\mathbf{w} \cdot \mathbf{x}_i + b) - 1] = 0.$$

One important aspect of the dual problem is that it only involves the dot product of the input vectors \mathbf{x} . This property allows us to use of the *kernel trick* to handle non-linearly separable data by mapping it to a higher-dimensional space.

15.5 Karush-Kuhn-Tucker conditions

When dealing with optimization problems that involve inequality constraints, such as those encountered in Support Vector Machines (SVMs), an additional requirement must be met: the solution must satisfy the **Karush-Kuhn-Tucker (KKT) conditions**.

The KKT conditions are a set of first-order necessary conditions that must be satisfied for a solution to be optimal. These conditions extend the method of Lagrange multipliers to handle inequality constraints and are particularly useful in non-linear programming. For the KKT conditions to apply, the problem must also satisfy certain regularity conditions. Fortunately, one of these regularity conditions is *Slater's condition*, which we have already established holds true for SVMs.

15.5.1 KKT Conditions and SVM Optimization

In the context of SVMs, the optimization problem is convex, meaning that the KKT conditions are not only necessary but also sufficient for optimality. This implies that if a solution satisfies the KKT conditions, it is guaranteed to be the optimal solution for both the primal and dual problems. Moreover, in this case, there is no duality gap, meaning the optimal values of the primal and dual problems are equal.

15.6 Prediction

Firstly, for a Linear SVM with no kernel, the primal weight vector is given by

$$\mathbf{w} = \sum_{i \in \mathcal{S}} \alpha_i y_i \mathbf{x}_i$$

Then, the decision function is

$$f(x) = \mathbf{w}^\top \mathbf{x} + b$$

Since the feature map $\phi(\cdot)$ may live in a huge or even infinite-dimensional space, we never form \mathbf{w} explicitly. Instead we keep the *kernel trick* inside the decision function:

$$f(x) = \sum_{i \in \mathcal{S}} \alpha_i y_i K(x_i, x) + b,$$

where $K(x_i, x) = \langle \phi(x_i), \phi(x) \rangle$. Typical kernels are the RBF $K(u, v) = \exp(-\frac{\|u-v\|^2}{2\sigma^2})$ or the polynomial $K(u, v) = (u^\top v + c)^p$.

Finally, we need to turn the decision value into a class label. For binary classification the prediction is simply the sign of the decision function:

$$\hat{y} = \text{sign}(f(x)) = \begin{cases} +1 & \text{if } f(x) > 0, \\ -1 & \text{if } f(x) \leq 0. \end{cases}$$

In sum,

$$\textbf{Predict}(x) : \quad f(x) = \sum_{i \in S} \alpha_i y_i K(x_i, x) + b, \quad \hat{y} = \text{sign}(f(x))$$

Chapter 16

Least Square SVM

16.1 Introduction

Least Squares Support Vector Machine (LS-SVM) is a modified version of the traditional Support Vector Machine (SVM) that simplifies the quadratic optimization problem by using a *least squares cost function*. LS-SVM transforms the quadratic programming problem in classical SVM into a set of linear equations, which are easier and faster to solve.

16.1.1 Optimization Problem (Primal Problem)

$$\begin{aligned} & \min_{w,b,e} \frac{1}{2} \|w\|^2 + \frac{\gamma}{2} \sum_{i=1}^N e_i^2, \\ & \text{subject to } y_i(w^T \phi(x_i) + b) = 1 - e_i, \quad \forall i \end{aligned}$$

where:

- w is the weight vector.
- b is the bias term.
- e_i are the error variables.

$$\begin{aligned} \sum_{i=1}^N e_i^2 &= \sum_{i=1}^N (1 - y_i(w^T \phi(x_i) + b))^2 \\ &= \sum_{i=1}^N (y_i^2 - y_i(w^T \phi(x_i) + b))^2 \quad , \text{ since } y_i = \pm 1, y_i^2 = 1 \\ &= \sum_{i=1}^N y_i^2 (y_i - (w^T \phi(x_i) + b))^2 \\ &= \sum_{i=1}^N (y_i - (w^T \phi(x_i) + b))^2 \end{aligned}$$

- γ is a regularization parameter.

- $\phi(x_i)$ is the feature mapping function.
- Note that $y_i^{-1} = y_i$, since $y_i = \pm 1$.

16.1.2 Lagrangian Function

To solve the constraint optimization problem, we define the Lagrangian function:

$$L(w, b, e, \alpha) = \min_{w, b, e} \frac{1}{2} \|w\|^2 + \frac{\gamma}{2} \sum_{i=1}^N e_i^2 - \sum_{i=1}^n \alpha_i [y_i(w^T \phi(x_i) + b) - 1 + e_i],$$

where α_i are Lagrange multipliers. Then, by setting the partial derivatives of the Lagrangian with respect to w , b , e , and α to zero, we get the KKT conditions.

- w :

$$\frac{\partial L}{\partial w} = w - \sum_{i=1}^n \alpha_i y_i \phi(x_i) = 0 \implies w = \sum_{i=1}^n \alpha_i y_i \phi(x_i)$$

- b :

$$\frac{\partial L}{\partial b} = - \sum_{i=1}^n \alpha_i y_i = 0$$

- e_i :

$$\frac{\partial L}{\partial e_i} = \gamma e_i - \alpha_i = 0 \implies \alpha_i = \gamma e_i$$

Thus, $e_i = \frac{\alpha_i}{\gamma}$

- α_i :

$$\frac{\partial L}{\partial \alpha_i} = - [y_i(w^T \phi(x_i) + b) - 1 + e_i] = 0 \implies y_i(w^T \phi(x_i) + b) = 1 - e_i, i = 1, \dots, N.$$

Let's substitute w and e :

- K : kernel matrix
- $\alpha = [\alpha_1, \alpha_2, \dots, \alpha_n]^T$
- $y = [y_1, y_2, \dots, y_n]^T$.
- $\Omega = YKY$, where $\Omega_{kl} = y_k y_l \phi(x_k)^T \phi(x_l)$
- $Y = \text{diag}(y)$.
- b : 1×1

Then, we can express it compactly

$$\begin{aligned} Y(KY^T\alpha + b\mathbf{1}) - \mathbf{1} + \frac{\alpha}{2\gamma} &= 0 \\ \mathbf{1}^T Y \alpha &= 0 \end{aligned}$$

By using the expression of α and b , we get

$$\begin{bmatrix} 0 & y^T \\ y & \Omega + \frac{1}{\gamma}I \end{bmatrix} \begin{bmatrix} b \\ \alpha \end{bmatrix} = \begin{bmatrix} 0 \\ \mathbf{1} \end{bmatrix}$$

Note that the dimension of the matrix on the left-hand side is $(N+1) \times (N+1)$. Once we have b and α by solving the linear system, the decision function for a new input x can be obtained by:

$$f(x) = \sum_{i=1}^n \alpha_i y_i K(x_i, x) + b.$$

Example: Suppose we have three training examples with feature vectors x_1, x_2 , and x_3 , and corresponding labels y_1, y_2 , and y_3 . The kernel matrix Ω is defined as:

$$\Omega_{ij} = y_i y_j K(x_i, x_j)$$

The dual form is:

$$\begin{bmatrix} 0 & y^T \\ y & \Omega + \frac{1}{\gamma}I \end{bmatrix} \begin{bmatrix} b \\ \alpha \end{bmatrix} = \begin{bmatrix} 0 \\ \mathbf{1} \end{bmatrix}$$

- $y = \begin{bmatrix} y_1 \\ y_2 \\ y_3 \end{bmatrix}$
- $\alpha = \begin{bmatrix} \alpha_1 \\ \alpha_2 \\ \alpha_3 \end{bmatrix}$
- $e = \begin{bmatrix} 1 \\ 1 \\ 1 \end{bmatrix}$
- I is a 3×3 identity matrix

Then, the Ω is given by

$$\Omega = \begin{bmatrix} y_1 y_1 K(x_1, x_1) & y_1 y_2 K(x_1, x_2) & y_1 y_3 K(x_1, x_3) \\ y_2 y_1 K(x_2, x_1) & y_2 y_2 K(x_2, x_2) & y_2 y_3 K(x_2, x_3) \\ y_3 y_1 K(x_3, x_1) & y_3 y_2 K(x_3, x_2) & y_3 y_3 K(x_3, x_3) \end{bmatrix}$$

Now, the complete matrix equation is:

$$\begin{bmatrix} 0 & y_1 & y_2 & y_3 \\ y_1 & \Omega_{11} + \frac{1}{\gamma} & \Omega_{12} & \Omega_{13} \\ y_2 & \Omega_{21} & \Omega_{22} + \frac{1}{\gamma} & \Omega_{23} \\ y_3 & \Omega_{31} & \Omega_{32} & \Omega_{33} + \frac{1}{\gamma} \end{bmatrix} \begin{bmatrix} b \\ \alpha_1 \\ \alpha_2 \\ \alpha_3 \end{bmatrix} = \begin{bmatrix} 0 \\ 1 \\ 1 \\ 1 \end{bmatrix}$$

This can be written explicitly as:

$$\begin{bmatrix} 0 & y_1 & y_2 & y_3 \\ y_1 & y_1^2 K(x_1, x_1) + \frac{1}{\gamma} & y_1 y_2 K(x_1, x_2) & y_1 y_3 K(x_1, x_3) \\ y_2 & y_2 y_1 K(x_2, x_1) & y_2^2 K(x_2, x_2) + \frac{1}{\gamma} & y_2 y_3 K(x_2, x_3) \\ y_3 & y_3 y_1 K(x_3, x_1) & y_3 y_2 K(x_3, x_2) & y_3^2 K(x_3, x_3) + \frac{1}{\gamma} \end{bmatrix} \begin{bmatrix} b \\ \alpha_1 \\ \alpha_2 \\ \alpha_3 \end{bmatrix} = \begin{bmatrix} 0 \\ 1 \\ 1 \\ 1 \end{bmatrix}$$

The solution to this matrix equation provides the values of b and $\alpha_1, \alpha_2, \alpha_3$, which are then used to construct the decision function:

$$f(x) = \sum_{i=1}^3 \alpha_i y_i K(x, x_i) + b$$

16.2 LS-SVM with Asymmetric Kernels

Recall that the dual form of LS-SVM is given by

$$\begin{bmatrix} 0 & y^T \\ y & \Omega + \frac{1}{\gamma} I \end{bmatrix} \begin{bmatrix} b \\ \alpha \end{bmatrix} = \begin{bmatrix} 0 \\ e \end{bmatrix}$$

An interesting point here is that using an asymmetric kernel in LS-SVM will not reduce to its symmetrization and asymmetric information can be learned. Then we can develop asymmetric kernels in the LS-SVM framework in a straightforward way.

Asymmetric kernels are particularly useful in capturing directional relationships in data that symmetric kernels cannot. For instance, in scenarios involving directed graphs or conditional probabilities, the relationship from x to y is inherently different from the relationship from y to x .

16.2.1 AsK-LS Primal Problem Formulation

We first define a generalized kernel trick for an inner product of two mappings ϕ_s and ϕ_t .

$$K(\mathbf{u}, \mathbf{v}) = \langle \phi_s(\mathbf{u}), \phi_t(\mathbf{v}) \rangle, \forall \mathbf{u} \in \mathbb{R}^{d_s}, \mathbf{v} \in \mathbb{R}^{d_t},$$

where $\phi_s : \mathbb{R}^{d_s} \rightarrow \mathbb{R}^p$, $\phi_t : \mathbb{R}^{d_t} \rightarrow \mathbb{R}^p$, and \mathbb{R}^p is a high-dimensional or even an infinite-dimensional space. Note that d_s and d_t can be different.

The primal problem for AsK-LS is formulated to handle these asymmetric relationships. The goal is to find the weight vectors ω and ν , and bias terms b_1 and b_2 , that minimize the following objective function:

$$\min_{\omega, \nu, b_1, b_2, e, h} \frac{1}{2} \omega^T \nu + \frac{\gamma}{2} \sum_{i=1}^m e_i^2 + \frac{\gamma}{2} \sum_{i=1}^m h_i^2,$$

subject to the constraints:

$$\begin{aligned} y_i(\omega^T \phi_s(x_i) + b_1) &= 1 - e_i \\ y_i(\nu^T \phi_t(x_i) + b_2) &= 1 - h_i \end{aligned}$$

Here:

- ω and ν are weight vectors for the source and target features.
- $\phi_s(x)$ and $\phi_t(x)$ are the source and target feature mappings.
- e_i and h_i are error terms for the source and target constraints.
- γ is a regularization parameter.

Note that this formulation is almost the same as the LS-SVM except that this considers both the source and target feature spaces simultaneously.

16.2.2 Dual Form

Let's transform it into a *dual* form. The dual problem involves solving a system of linear equations derived from the primal problem's Lagrangian. The Lagrangian function for the primal problem is:

$$\begin{aligned}\mathcal{L}(\omega, \nu, b_1, b_2, e, h, \alpha, \beta) = & \\ \frac{1}{2}\omega^T \nu + \frac{\gamma}{2} \sum_{i=1}^m e_i^2 + \frac{\gamma}{2} \sum_{i=1}^m h_i^2 + & \\ \sum_{i=1}^m \alpha_i(1 - e_i - y_i(\omega^T \phi_s(x_i) + b_1)) + \sum_{i=1}^m \beta_i(1 - h_i - y_i(\nu^T \phi_t(x_i) + b_2)) &\end{aligned}$$

The KKT conditions are derived by setting the partial derivatives of the Lagrangian with respect to ω, ν, b_1, b_2, e , and h to zero. The dual problem leads to the following linear system:

$$\begin{bmatrix} 0 & 0 & Y^T & 0 \\ 0 & 0 & 0 & Y^T \\ Y & 0 & \frac{I}{\gamma} & H \\ 0 & Y & H^T & \frac{I}{\gamma} \end{bmatrix} \begin{bmatrix} b_1 \\ b_2 \\ \alpha \\ \beta \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \\ 1 \\ 1 \end{bmatrix}$$

where:

- Y is a vector of class labels.
- H is the kernel matrix with elements $H_{ij} = y_i K(x_i, x_j) y_j$, where $K(x_i, x_j) = \langle \phi_s(x_i), \phi_t(x_j) \rangle$ is the asymmetric kernel function.
 - For an asymmetric kernel K , the kernel function $K(x_i, x_j) \neq K(x_j, x_i)$. This asymmetry is directly incorporated into the matrix H , where:

$$\begin{aligned}H_{ij} &= y_i K(x_i, x_j) y_j \\ H_{ji} &= y_j K(x_j, x_i) y_i\end{aligned}$$

AsK-LS uses two different feature mappings ϕ_s and ϕ_t for the source and target features. This approach allows capturing more information compared to symmetric kernels. The dual solution provides weight vectors ω and ν , which span the target and source feature spaces, respectively.

The decision functions for classification from the source and target perspectives are given by

$$\begin{aligned}f_t(x) &= \sum_{i=1}^m \alpha_i y_i K(x_i, x) + b_1 \\ f_s(x) &= \sum_{i=1}^m \beta_i y_i K(x, x_i) + b_2\end{aligned}$$

These functions utilize the learned asymmetric relationships in the data.

16.2.3 Asymmetric Kernels

We can consider asymmetric kernels like these:

$$K_{SNE}(\mathbf{x}, \mathbf{y}) = \frac{\exp(-\|\mathbf{x} - \mathbf{y}\|_2^2)/\sigma^2}{\sum_{\mathbf{z} \in \mathbf{X}_{train}} \exp(-\|\mathbf{x} - \mathbf{z}\|_2^2)/\sigma^2}$$
$$K_T(\mathbf{x}, \mathbf{y}) = \frac{(1 + \|\mathbf{x} - \mathbf{y}\|_2^2)^{-1}}{\sum_{\mathbf{z} \in \mathbf{X}_{train}} (1 + \|\mathbf{x} - \mathbf{z}\|_2^2)^{-1}}$$

Part IV

Generative Modeling

Chapter 17

Introduction

17.1 KL Divergence

The KL divergence can be defined as follows:

$$D_{KL}(P||Q) = \mathbb{E}_{x \sim P} \left[\log \frac{P(X)}{Q(X)} \right]$$

17.1.1 Properties

- Non symmetric
- $D_{KL} \in [0, \infty]$
- In order for the KL divergence to be finite, the support of P needs to be in the support of Q .

17.1.2 Rewriting the Objective

$$\begin{aligned} D_{KL}(P||Q) &= \mathbb{E}_{x \sim P} \left[\log \frac{P(X)}{Q(X)} \right] \\ &= \mathbb{E}_{x \sim P}[-\log Q(X)] - \mathcal{H}(P(X)) \end{aligned}$$

- $\mathbb{E}_{x \sim P}[-\log Q(X)]$: Cross entropy
- $\mathcal{H}(P(X))$: Entropy of P

17.1.3 Forward and Reverse KL

Let's say there is a true distribution $P(X)$ with two modes and our approximation $Q(X)$ has one mode. Then,

- Forward KL: $D_{KL}(P||Q)$
- Reverse KL: $D_{KL}(Q||P)$

Forward KL: Mean-Seeking Behavior

$$\begin{aligned}\arg \min_{\theta} D_{KL}(P||Q) &= \arg \min_{\theta} \mathbb{E}_{x \sim P}[-\log Q_{\theta}(X)] - \mathcal{H}(P(X)) \\ &= \arg \min_{\theta} \mathbb{E}_{x \sim P}[-\log Q_{\theta}(X)] \\ &= \arg \max_{\theta} \mathbb{E}_{x \sim P}[\log Q_{\theta}(X)]\end{aligned}$$

Intuition: x will be sampled from the distribution P , and its value will be estimated from Q . Thus, there will be higher chance that x will be sampled from a space with higher probability in P . Therefore, Q_{θ} has to consider all modes, which have high probabilities.

To use the forward KL, we have to have an access to the true model $P(X)$ for sampling.

Reverse KL: Mode-Seeking Behavior

$$\begin{aligned}\arg \min_{\theta} D_{KL}(Q||P) &= \arg \min_{\theta} \mathbb{E}_{x \sim Q_{\theta}}[-\log P(X)] - \mathcal{H}(Q_{\theta}(X)) \\ &= \arg \max_{\theta} \mathbb{E}_{x \sim Q_{\theta}}[\log P(X)] + \mathcal{H}(Q_{\theta}(X))\end{aligned}$$

Intuition: x will be sampled from the distribution Q , and its value will be estimated from P . Thus, there will be higher chance that x will be sampled from a space with higher probability in Q . Therefore, to maximize the value, we need to focus on a single mode.

To use the reverse KL, we have to be able to evaluate the true model $P(X)$.

Chapter 18

Sampling Based Inference

18.1 Basic Sampling Methods

18.1.1 Inverse Transform Sampling

Inverse transform sampling is a basic method for pseudo-random number sampling, *i.e.*, for generating sample numbers at random from any probability distribution given its cumulative distribution function (CDF).

Assume that we already have a uniformly distributed random number generator, *e.g.*, `np.random.randn()`

1. Generate a random number $u \sim \text{Unif}[0, 1]$
2. Find the inverse of the desired CDF, $F_X^{-1}(x)$.
3. Compute $X = F_X^{-1}(u)$. The computed random variable X has distribution $F_X(x)$

However, it is **hard to compute the inverse of CDF ($F_X(x)$)**

- $F_X(x) : \mathbb{R} \mapsto [0, 1]$ is any CDF.
- CDF is a non-negative and non-decreasing (monotone) function that is continuous.
- Our objective is to simulate a random variable X distributed as F ; that is, we want to simulate a X such that $P(X \leq x) = F(x)$.
- F is invertible since it is continuous and strictly increasing.

18.1.2 Ancestral Sampling

$$p(\mathbf{x}) = p(\mathbf{x}_1)p(\mathbf{x}_2|\mathbf{x}_1)p(\mathbf{x}_3|\mathbf{x}_2)\cdots$$

Sampling steps:

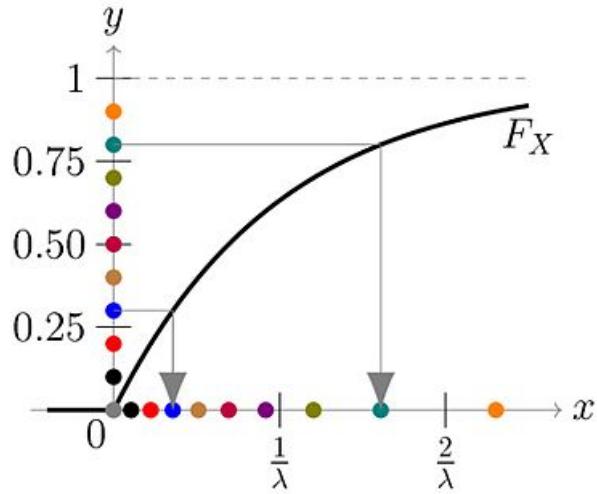
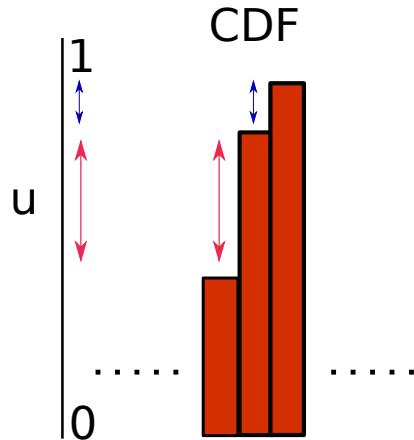
Figure 18.1: y -axis: Uniform distribution, x -axis: sample value

Figure 18.2: How can this sampling method recover the original distribution?

1. sample \mathbf{x}_1
2. sample \mathbf{x}_2 conditioned by \mathbf{x}_1
3. sample \mathbf{x}_3 conditioned by \mathbf{x}_2

18.1.3 Rejection Sampling

Rejection sampling is a simple method. It rejects samples violating a given condition (*e.g.*, conditions of conditional probability.). Let's see its theory.

Rejection sampling is a method for sampling from a distribution $p(x) = \frac{1}{Z}p'(x)$ that is difficult to sample directly, but its unnormalized pdf $p'(x)$ is easy to evaluate (Z is hard to compute). In rejection sampling, we need some simpler distribution $q(x)$, called a **proposal distribution**.

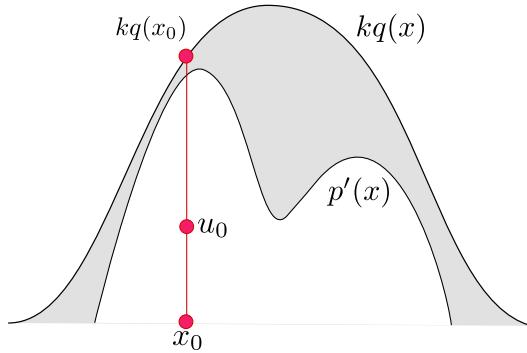
The intuition of rejection sampling is actually similar to Monte-Carlo estimation. By setting a large area (proposal distribution), we can sample points and take them that are inside the our

target distribution

To run the rejection sampling, introduce a constant k whose value is chosen such that $kq(x) \geq p'(x)$ for all values of x . The function $kq(x)$ is called a comparison function. Each step of the rejection sampler involves generating two random variables:

1. Sample $x_0 \sim q$
2. Sample $u_0 \sim U[0, kq(x_0)]$.

Finally, If $u_0 > p'(x_0)$, then the sample x_0 will be rejected, otherwise we add the sample x_0 to our set of samples $\{x^r\}$.



The original values of x are generated from the distribution $q(x)$ and these samples are then accepted with probability $p'(x)/kq(x)$ (see the figure above. The acceptance probability (*i.e.*, length) is the p' divided by kq). Then, the probability that a sample will be accepted is given by

$$\begin{aligned} p(\text{accept}) &= p\left(u \leq \frac{p'(x)}{kq(x)}\right) \\ &= \int p\left(u \leq \frac{p'(x)}{kq(x)} \mid x\right) q(x) dx \\ &= \int \frac{p'(x)}{kq(x)} q(x) dx \\ &= \frac{1}{k} \int p'(x) dx \end{aligned}$$

Thus, the sampling will be more efficient if we choose small k to increase the chance of acceptance.

18.1.4 Importance Sampling

We want to estimate an expectation of function $f(x)$, where $x \sim p(x)$, but it is hard to estimate the distribution $p(x)$. Again, the importance sampling is not a method for generating samples from $p(x)$. In this case, we can use a simple distribution $q(x)$ by

$$\begin{aligned}
\mathbb{E}_p[f(\mathbf{x})] &= \int p(\mathbf{x})f(\mathbf{x})d\mathbf{x} \\
&= \int p(\mathbf{x})f(\mathbf{x})\frac{q(\mathbf{x})}{q(\mathbf{x})}d\mathbf{x} \\
&= \int q(\mathbf{x})\left[f(\mathbf{x})\frac{p(\mathbf{x})}{q(\mathbf{x})}\right]d\mathbf{x} \\
&= \mathbb{E}_q\left[f(\mathbf{x})\frac{p(\mathbf{x})}{q(\mathbf{x})}\right]
\end{aligned}$$

- Assume that $p(\mathbf{x})$ is known and too complicated to be sampled directly.
- Samples are independently drawn from a **proposal density** $Q(\mathbf{x})$, which is designed to be close to the true density $p(\mathbf{x})$ and **simpler**
- Generate R samples from $Q(\mathbf{x})$

By applying the Monte-Carlo method, we can get

$$\mathbb{E}_q\left[f(\mathbf{x})\frac{p(\mathbf{x})}{q(\mathbf{x})}\right] \approx \frac{1}{N} \sum_{i=1}^N f(\mathbf{x}_i)\frac{p(\mathbf{x}_i)}{q(\mathbf{x}_i)}, \quad \mathbf{x}_i \sim p(\mathbf{x}) \quad (18.1)$$

- Unbiased estimation
- (Potentially) Smaller variance compared to the vanilla Monte-Carlo method above.

- $Var_q\left[f(\mathbf{x})\frac{p(\mathbf{x})}{q(\mathbf{x})}\right] < Var_p[f(\mathbf{x})]$
- When $q(\mathbf{x})$ is high where $|p(\mathbf{x})f(\mathbf{x})|$ is high.

Chapter 19

Markov Chain Monte Carlo

19.1 Gibbs Sampling

The phrase “Markov chain Monte Carlo” encompasses a broad array of techniques that have in common a few key ideas. The setup for all the techniques that we will discuss in this book is as follows:

1. We want to sample from a some complicated density or probability mass function π . Often, this density is the result of a Bayesian computation so it can be interpreted as a posterior density. The presumption here is that we can evaluate π but we cannot sample from it.
2. We know that certain stochastic processes called Markov chains will converge to a stationary distribution (if it exists and if specific conditions are satisfied). Simulating from such a Markov chain for a long enough time will eventually give us a sample from the chain’s stationary distribution.
3. Given the functional form of the density π , we want to construct a Markov chain that has π as its stationary distribution.
4. We want to sample values from the Markov chain such that the sequence of values $\{x_n\}$ generated by the chain converges in distribution to the density π .

In order for all these ideas to make sense, we need to first go through some background on Markov chains. The rest of this chapter will be spent defining all these terms, the conditions under which they make sense, and giving examples of how they can be implemented in practice.

19.2 Markov Chain

[Reference Link](#)

A Markov chain is a stochastic process that evolves over time by transitioning into different states. The sequence of states is denoted by the collection $\{X_i\}$ and the transition between states is random. Formally,

Definition 1 Let D be a finite set. A random process X_1, X_2, \dots with values in D is called a Markov chain if

$$P(X_t = x_{t+1} | X_t = x_t, \dots, X_0 = x_0) = P(X_{t+1} = x_{t+1} | X_t = x_t)$$

We can think of X_t as a random state at time t , and the Markovian assumption is that the probability of transitioning from x_t to x_{t+1} only depends on x_t . In other words, the future state depends only on the present. Let p_{ij} be the probability of transitioning from state i to state j . A Markov chain can be defined by a transition probability matrix:

Definition 2 The matrix $\mathbf{P} = (p_{ij})_{i,j} \in D$ is called the transition probability matrix.

Thus, P is a $D \times D$ matrix, where $|D|$ denotes the cardinality of D , and the cell value p_{ij} is the probability of transitioning from state i to state j , and the rows of P must sum to one. We will restrict ourselves to time *homogeneous Markov chains*:

Definition 3 A Markov chain is called time homogeneous if

$$\mathbb{P}\{X_{t+1} = j | X_n = i\} = p_{ij}, \forall n.$$

It state that *the transition probabilities are not changing as a function of time*. Finally, let's introduce some useful notation for the initial state of the Markov chain. Let

$$\mathbb{P}_{x_0}\{\cdot\} \triangleq \mathbb{P}\{\cdot | X_0 = x_0\}.$$

For example, I will write $\mathbb{P}_a\{X_1 = b\}$ rather than $\mathbb{P}\{X_1 = b | X_0 = a\}$ for simplicity.

Consider a simple Markov chain modeling the weather. The weather has two states: rainy and sunny. Thus, $D = \{r, s\}$ and X_n is the “weather on day n ”. The Markov chain model is as follows. If today is rainy (r), tomorrow it is sunny (s) with probability p . If today it is sunny, tomorrow it is rainy with probability q . Then, transition matrix is given by

$$\mathbf{P} = \begin{bmatrix} 1-p & p \\ q & 1-q \end{bmatrix}.$$

The state diagram of the chain can be represented as follows: As a consequence of the Markovian

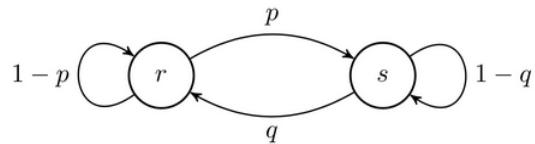


Figure 19.1: A sample Markov chain.

assumptions, the probability of any path on a Markov chain is just the multiplication of the numbers along the path. For example, for some Markov chain with states $D = \{a, b, c, d\}$, the probability of a particular path, say $a \rightarrow b \rightarrow b \rightarrow d \rightarrow c$ factorize as

$$p_{ab} \cdots p_{dc}$$

Let's try to formulate this by

$$\mathbb{P}_i\{X_n = j\}.$$

For instance, let's say $n = 2$. Then, we have

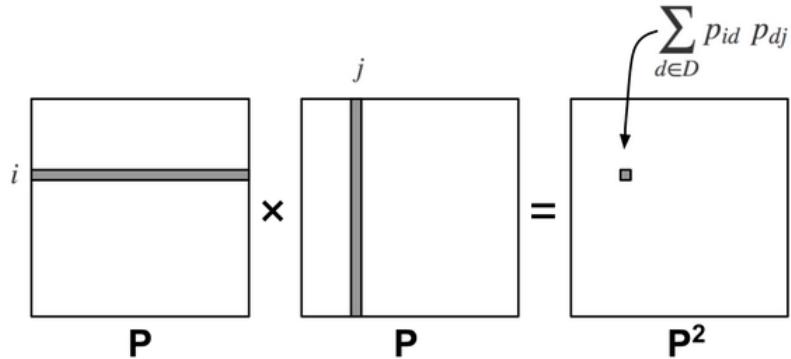
$$\begin{aligned}\mathbb{P}_i\{X_2 = j\} &= \sum_{d \in D} \mathbb{P}_i\{X_2 = j, X_1 = d\} \\ &= \sum_{d \in D} \mathbb{P}_i\{X_2 = j | X_1 = d\} \mathbb{P}_i\{X_1 = d\} \\ &= \sum_{d \in D} p_{id} p_{dj}\end{aligned}$$

This is equivalent to a dot product of the i -th row to j -th column of the transition matrix \mathbb{P} . This can be expressed as follows:

$$\mathbb{P}\{X_2 = j | X_0 = i\} = \sum_{d \in D} p_{id} p_{dj} = (\mathbf{P}_{ij}^2).$$

This can be generalized to

$$\mathbb{P}_i\{X_n = j\} = (\mathbf{P}_{ij}^n).$$



In sum, the dot product performs marginalization, and everything works out nicely thanks to the Markovian assumption. If the D -dimensional vector v represents a discrete distribution over initial states X_0 , then $\mathbf{v}^T \mathbf{P}$ is a D -dimensional vector representing the probability distribution over X_1 .

19.2.1 Ergodicity

Let's discuss about *ergodicity*, which is a property of a random process in which its time average is the same as its probability space average. It can be defined as follows:

Definition 4 A Markov chain $\{X_n\}$ is called ergodic if the limit

$$\pi(j) = \lim_{n \rightarrow \infty} \mathbb{P}_i\{X_n = j\}$$

exists for every state j and does not depend on the initial state i . The D -dimensional vector $\pi(j)$ is called the stationary probability.

In other words,

- The probability $\pi(j)$ of reaching at state j (i.e., $\mathbb{P}_i\{X_n = j\}$)
- After a long time (i.e., $\lim_{n \rightarrow \infty}$)
- Regardless of the initial state i (i.e., $\mathbb{P}_i\{X_n = j\}$).

Equivalently, it can be expressed as

$$\pi(j) = \lim_{n \rightarrow \infty} (\mathbf{P}^n)_{ij}.$$

The ergodicity gives the following property:

$$\begin{aligned}\pi(j) &= \lim_{n \rightarrow \infty} (\mathbf{P}^n)_{ij} \\ &\stackrel{*}{=} \lim_{n \rightarrow \infty} (\mathbf{P}^{n+1})_{ij} \\ &= \lim_{n \rightarrow \infty} (\mathbf{P}^n \mathbf{P})_{ij} \\ &= \lim_{n \rightarrow \infty} \sum_{d \in D} (\mathbf{P}^n)_{id} \mathbf{P}_{dj} \\ &= \sum_{d \in D} \pi(d) \mathbf{P}_{dj}\end{aligned}$$

- The step $*$ holds since the step n and $n + 1$ does not matter under the limit.

Finally, we can write this as

$$\pi^T = \pi^T \mathbf{P},$$

where π is a column vector. Hence, the name “stationary probability distribution” denotes that it is a distribution that does not change over time. Note that it can be also written as:

$$\mathbf{P}^T \pi = \pi.$$

Then, we can say π is an eigenvector of \mathbf{P}^T with eigenvalue of 1. Thus, we can obtain the stationary distribution π from an eigenvectors and eigenvalues of \mathbf{P}^T .

Since we are interested in ergodicity, let's now introduce some properties related to reachability and long-term behavior.

Definition 5 We say that there is a path from i to j ($i \rightsquigarrow j$) if there is a nonzero probability that starting at i , we can reach j at some point in the future.

Definition 6 A state i is called **transient** if there exists a state j such that $i \rightsquigarrow j$ but $j \not\rightsquigarrow i$.

Definition 7 A state i is called **recurrent** if for all states j there exist $i \rightsquigarrow j$ and $j \rightsquigarrow i$.

Definition 8 A Markov chain is **irreducible** if $i \leftrightarrow j, \forall i, j \in S$. Simply, if all states are able to visit other states, it is irreducible.

Definition 9 State i has a **period** d (i.e., periodically visit the state i) \leftrightarrow aperiodic.

Intuitively, a *transient* state is a state that cannot return to itself; while a *recurrent* state can return. In other words, these two conditions are opposite. A transient state is not recurrent, and a recurrent state is not transient. Note that if a state is recurrent, every reachable state is also recurrent. We think of a set of recurrent states as a "class" or a "recurrent class". We can notice

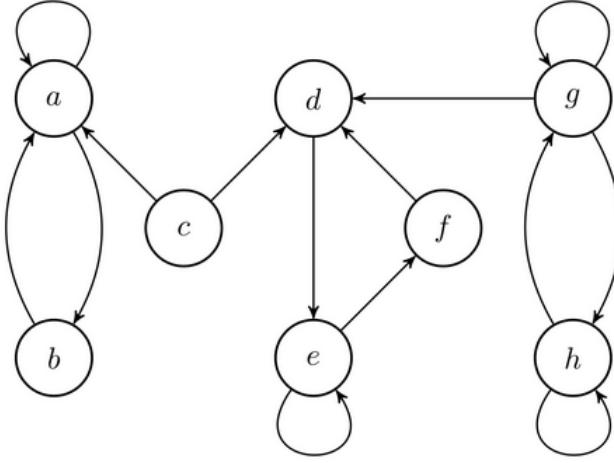


Figure 19.2: A Markov chain with $D = \{a, b, c, d, e, f, g, h\}$. The state c is a transient state. Note that g and h are also transient as there exist states that they cannot return. There are two recurrent classes: $\{a, b\}$ and $\{d, e, f\}$.

that Fig. 19.1 is ergodic but Fig. 19.2 is not since some states are not recurrent.

In sum a state is ergodic if the state is recurrent and aperiodic. Markov chain is ergodic if all states are ergodic.

19.2.2 Limit Theorem of Markov chain

Reference Link

For a Markov chain with a discrete state space and transition matrix P , let π_* be such that $\pi_*P = \pi_*$. Then π_* is a stationary distribution of the Markov chain and the chain is said to be stationary if it reaches this distribution.

The basic limit theorem for Markov chains says that, under a specific set of assumptions that we will detail below, we have

$$\|\pi_* - \pi_n\| \rightarrow 0$$

as $n \rightarrow \infty$, where $\|\cdot\|$ is the total variation distance between the two densities. Therefore, no matter where we start the Markov chain (π_0), π_n will eventually approach the stationary distribution. Another way to think of this is that

$$\lim_{n \rightarrow \infty} \pi_n(i) = \pi_*(i).$$

for all states i in the state space. Note that π_0 is the probability distribution of the Markov chain at time 0. Also, π_n denote the distribution of the chain at time n .

19.2.3 Time Reversibility

Consider a stationary ergodic Markov chain with transition probability $p(i, j)$ and stationary distribution $\pi(i)$, if we reverse the process, we will get a reversed Markov chain with transition probability $q(i, j)$:

$$\begin{aligned} q(j, i) &= P(X_m = i | X_{m+1} = j) \\ &= \frac{P(X_m = i, X_{m+1} = j)}{P(X_{m+1} = j)} \\ &= \frac{P(X_m = i | X_{m+1} = j)P(X_{m+1} = j)}{P(X_{m+1} = j)} \\ &= \frac{\pi(i)p(i, j)}{\pi(j)} \\ \pi(i)p(i, j) &= \pi(j)q(j, i) \end{aligned}$$

If $p(i, j) = q(j, i)$, it is called time-reversible Markov chain.

19.3 Markov Chain Monte Carlo

MCMC aims to generate samples from some complex probability distribution $p(x)$ that is difficult to directly sample from.

The basic sampling methods we have learnt so far do not leverage past information, which assumes all samples are independent. In Markov chain based sampling, we will treat random variables as a sequence of sampling process.

In Markov Chain Monte Carlo(MCMC), we assume that a stationary distribution is already known. We are more interested in estimating a transition rule that describing the stationary distribution.

Ground rules for MCMCs:

- MCMCs stochastically explore the parameter space in such a way that the histogram of their samples produces the target distribution.
- Markovian: Evolution of the chain (*i.e.*, collections of samples from one iteration to the other) only depends on the current position and some transition probability distribution (*i.e.*, how we move from one point in parameter space to another). This means that the chain has no memory and past samples cannot be used to determine new positions in parameter space.
- The chain will converge to the target distribution if the transition probability is:
 - Irreducible: From any point in parameter space, we must be able to reach any other point in the space in a finite number of steps.
 - Positive recurrent: For any point in parameter space, the expected number of steps for the chain to return to that point is finite. This means that the chain must be able to re-visit previously explored areas of parameter space.
 - Aperiodic: The number of steps to return to a given point must not be a multiple of some value k . This means that the chain cannot get caught in cycles.

19.3.1 Metropolis-Hasting Algorithm

Suppose we have a target posterior distribution $\pi(x)$, where x here can be any collection of parameters (not a single parameter). In order to move around this parameter space we must formulate some proposal distribution:

$$q(x_{i+1} | x_i),$$

that specifies the probability of moving to a point in parameter space, x_{i+1} , given that we are currently at x_i . The Metropolis Hastings algorithm accepts a “jump” to x_{i+1} with the following probability

$$\kappa(x_{i+1} | x_i) = \min \left(1, \frac{\pi(x_{i+1})q(x_i | x_{i+1})}{\pi(x_i)q(x_{i+1} | x_i)} \right) = \min(1, H),$$

where the fraction above is called the Hastings ratio, H . The above expression represents that the probability of transitioning from point x_{i+1} given the current position x_i is a function of the ratio of the value of the posterior at the new point to the old point (*i.e.*, $\pi(x_{i+1})/\pi(x_i)$) and the ratio of the transition probabilities at the new point to the old point (*i.e.*, $q(x_i | x_{i+1})/q(x_{i+1} | x_i)$). Firstly, it is clear that if the ratio is bigger than 1 then the jump will be accepted. Secondly, the ratio of the target posteriors ensures that the chain will gradually move to high probability regions. Lastly, the ratio of the transition probabilities ensures that the chain is not “favored” toward certain locations by the proposal distribution function. Note that many proposal distributions are symmetric (*i.e.*, $q(x_{i+1} | x_i) = q(x_i | x_{i+1})$).

The Metropolis-Hastings algorithm is then:

```

1 def mh_sampler(x0, lnprob_fn, prop_fn, prop_fn_kwargs={}, iterations=100000):
2     """Simple metropolis hastings sampler.
3
4     :param x0: Initial array of parameters.
5     :param lnprob_fn: Function to compute log-posterior.
6     :param prop_fn: Function to perform jumps.
7     :param prop_fn_kwargs: Keyword arguments for proposal function
8     :param iterations: Number of iterations to run sampler. Default=100000
9
10    :returns:
11        (chain, acceptance, lnprob) tuple of parameter chain , acceptance rate
12        and log-posterior chain.
13    """
14
15    # number of dimensions
16    ndim = len(x0)
17
18    # initialize chain, acceptance rate and lnprob
19    chain = np.zeros((iterations, ndim))
20    lnprob = np.zeros(iterations)
21    accept_rate = np.zeros(iterations)
22
23    # first samples
24    chain[0] = x0
25    lnprob0 = lnprob_fn(x0)
26    lnprob[0] = lnprob0
27
28    # start loop
29    naccept = 0
30    for ii in range(1, iterations):
31
32        # propose
33        x_star, factor = prop_fn(x0, **prop_fn_kwargs)
34
35        # draw random uniform number
36        u = np.random.uniform(0, 1)
37
38        # compute hastings ratio
39        lnprob_star = lnprob_fn(x_star)
40        H = np.exp(lnprob_star - lnprob0) * factor
41
42        # accept/reject step (update acceptance counter)
43        if u < H:
44            x0 = x_star
45            lnprob0 = lnprob_star
46            naccept += 1
47
48        # update chain
49        chain[ii] = x0
50        lnprob[ii] = lnprob0

```

```
51     accept_rate[ii] = naccept / ii  
52  
53     return chain, accept_rate, lnprob
```

Chapter 20

Topic Modeling

20.1 Latent Semantic Allocation

Topic model provides insights like

- Can analyze topics. *e.g.*,, a certain topic includes specific words more than other topics.
- A certain document's topic distribution (simplex)

pLSA: Latent Variable Model

$$P_{LSA}(w|d) = \sum_z P(w|z; \theta)P(z|d; \pi)$$

- w : word
- d : document
- z : latent variable

Equivalently,

$$P_{LSA}(d, w) = \sum_z P(w|z)P(d|z)p(z) = p(d) \sum_z P(w|z)P(z|d)$$

The probability of observing $n(w_i, d_j)$ occurrences of word w_i in document d_j is given by

$$p(w_i, d_j)^{n(w_i, d_j)}$$

The probability of observing the complete document collection is then given by the product of probabilities of observing every single word in every document with corresponding number of occurrences.

Then, the likelihood function becomes

$$L = \prod_{i=1}^m \prod_{j=1}^n p(w_i, d_j)^n (w_i, d_j)$$

The log-likelihood is then

$$\begin{aligned} \mathcal{L} &= \sum_{i=1}^m \sum_{j=1}^n n(w_i, d_j) \log(p(w_i, d_j)) \\ &= \sum_{l=1}^k p(w_i|z_l) p(d_j|z_l) p(z_l) \end{aligned}$$

Parameter Inference:

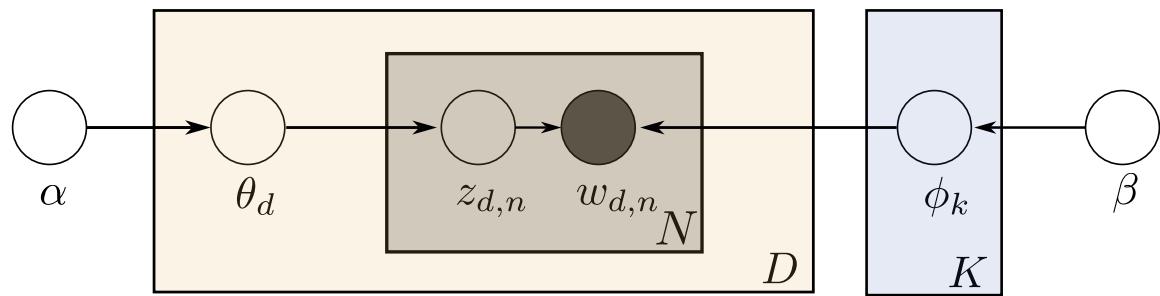
- We cannot maximize the likelihood analytically because of the logarithm of the sum
- A standard procedure is to use *EM*.

20.2 Latent Dirichlet Allocation

The assumptions of LDA:

- Each topic is a distribution over words.
- Each document is a mixture of corpus-wide topics.
- Each word is sampled from one of topics.

The LDA attempts to model the document generation process stochastically. However, we have to infer the latent structure (the distributions) of documents.



- $\theta_d \sim Dir(\alpha)$: For each document, draw topic distribution.
 - α : Dirichlet parameter
- $z_{d,n} \sim Mult(\theta_d)$: per-word topic assignment. The n -th word of document d is from which topic?
- $w_{d,n} \sim Mult(\phi_{z_{d,n}}, n)$: observed word. The n -th word in a document d is from a certain topic ($z_{d,n}$) distribution $\phi_{z_{d,n}}$.

- $\phi_k \sim Dir(\beta), i = \{1, \dots, K\}$: topics.
- β : topic hyperparameter (Dirichlet parameter).

We can immediately notice that the LDA's modeling approach is quite far from the way we write texts. However, LDA works quite well.

The document generation process can be modelled as follows:

$$p(\phi_{1:K}, \theta_{1:D}, z_{1:D}, w_{1:D}) = \prod_{i=1}^K p(\phi_i | \beta) \prod_{d=1}^D p(\theta_d | \alpha) \left(\prod_{n=1}^N p(w_{d,n} | \phi_{1:K}, z_{d,n}) p(z_{d,n} | \theta_d) \right).$$

20.2.1 LDA Inference

The posterior of the latent variables given the document is

$$p(\phi, \theta, \mathbf{z} | \mathbf{w}) = \frac{p(\phi, \theta, \mathbf{z}, \mathbf{w})}{\int_{\phi} \int_{\theta} \sum_{\mathbf{z}} p(\phi, \theta, \mathbf{z}, \mathbf{w})}$$

Computing the posterior is intractable:

- The denominator is intractable
- We cannot compute the denominator, the marginal $p(\mathbf{w})$.
- We are going to use collapsed Gibbs sampling.

We want to estimate the topic distribution \mathbf{z} .

20.2.2 Dirichlet Distribution

The Dirichlet Distribution can be considered as a extension of the *beta distribution*.

$$p(P = \{p_i\} | \alpha_i) = \frac{\Gamma(\sum_i \alpha_i)}{\prod_i \Gamma(\alpha_i)} \prod_i p_i^{\alpha_i - 1} \quad (20.1)$$

- $\sum_i p_i = 1$
- The posterior distribution of Dirichlet distribution is also Dirichlet distribution.

Chapter 21

Latent Variable Models

21.1 Motivation of Latent Variable Models

Let's say we want to classify some data. If we had access to a corresponding latent variable for each observation \mathbf{x}_i , modeling would be more straightforward. To illustrate this, consider the challenge of finding the latent variable (*i.e.*, the true class of \mathbf{x}) $z^* = \text{argmax}_z p(\mathbf{x}|z)$, as shown in Fig. 21.1(b).

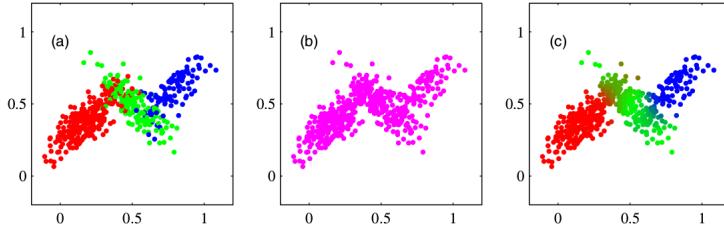


Figure 21.1: (a) Complete dataset $p(\mathbf{x}|z)$. (b) Incomplete dataset $p(\mathbf{x})$. (c) Inference result.

Consider modeling the complete data set $p(\mathbf{x}|z)$ under the assumption that the observations are independent and identically distributed (i.i.d.). Based on the Fig. 21.1(a), the joint distribution for a single observation $(\mathbf{x}_i, \mathbf{z}_i)$ given the model parameters $\boldsymbol{\theta}$ can be expressed:

$$p(\mathbf{x}_i, \mathbf{z}_i | \boldsymbol{\theta}) = \begin{cases} p(\mathcal{C}_1)p(\mathbf{x}_i | \mathcal{C}_1) & \text{if } z_i = 0 \\ p(\mathcal{C}_2)p(\mathbf{x}_i | \mathcal{C}_2) & \text{if } z_i = 1 \\ p(\mathcal{C}_3)p(\mathbf{x}_i | \mathcal{C}_3) & \text{if } z_i = 2 \end{cases}$$

Given N observations, the joint distribution for the entire dataset $\{\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_N\}$ along with their corresponding latent variables $\{\mathbf{z}_1, \mathbf{z}_2, \dots, \mathbf{z}_N\}$ is:

$$p(\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_N, \mathbf{z}_1, \mathbf{z}_2, \dots, \mathbf{z}_N | \boldsymbol{\theta}) = \prod_{n=1}^N \prod_{k=1}^K \pi_k^{z_{nk}} \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)^{z_{nk}}$$

Here, $\pi_k = p(\mathcal{C}_k)$ represents the prior probability of the k -th component, and $p(\mathbf{x}_n | \mathcal{C}_k) =$

$\mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$ denotes the Gaussian distribution associated with component \mathcal{C}_k . Also, $z_{nk} \in \{0, 1\}$ and $\sum_k z_{nk} = 1$.

However, in practice, the latent variables \mathbf{z}_k are often not directly observable, which complicates the modeling process.

In the following sections, we present various methods for identifying and handling these latent variables to improve the classification and modeling of data.

Chapter 22

Clustering

22.1 K-Means Clustering

Suppose that we have a data set $\mathbf{X} = \{\mathbf{x}_1, \dots, \mathbf{x}_n\}$ consisting of N observations of a random D -dimensional variable $\mathbf{x} \in \mathbb{R}^D$. Our goal is to partition the data into K of clusters. Intuitively, a cluster can be thought as **a group of data points whose inter-point distances are small compared with the distances to points outside of the cluster**.

This notion can be formalized by introducing a set of D -dimensional vectors $\boldsymbol{\mu}_k$, which represents the centers of the clusters. Our goal is to find an assignment of data points to clusters, as well as a set of vectors $\{\boldsymbol{\mu}_k\}$. Objective function of K -means clustering (*distortion measure*) can be defined as follows:

$$J = \sum_{n=1}^N \sum_{k=1}^K r_{nk} \|\mathbf{x}_n - \boldsymbol{\mu}_k\|^2$$

, where $r_{nk} \in \{0, 1\}$ is a binary indicator variable which represents the **membership of data \mathbf{x}_n** . It can be expressed as follows:

$$r_{nk} = \begin{cases} 1 & \text{if } k = \operatorname{argmin}_j \|\mathbf{x}_n - \boldsymbol{\mu}_j\|^2 \\ 0 & \text{otherwise} \end{cases}$$

Our goal is to find values for the $\boldsymbol{\mu}_k$ and the r_{nk} that minimize J .

We can minimize J through an iterative procedure in which each iteration involves two successive steps corresponding to successive optimizations with respect to the $\boldsymbol{\mu}_k$ and the r_{nk} . First we choose some initial values for the $\boldsymbol{\mu}_k$. Then in the first phase we minimize J with respect to the r_{nk} , keeping the $\boldsymbol{\mu}_k$ fixed. In the second phase we minimize J with respect to the $\boldsymbol{\mu}_k$, keeping r_{nk} fixed. This two-stage optimization is then repeated until convergence.

Now consider the optimization of the $\boldsymbol{\mu}_k$ with the r_{nk} held fixed. The objective function J is a quadratic function of $\boldsymbol{\mu}_k$, and it can be minimized by setting its derivative with respect to $\boldsymbol{\mu}_k$ to zero giving

$$2 \sum_{n=1}^N r_{nk} (\mathbf{x}_n - \boldsymbol{\mu}_k) = 0.$$

We can arrange as

$$\boldsymbol{\mu}_k = \frac{\sum_n r_{nk} \mathbf{x}_n}{\sum_n r_{nk}}.$$

The denominator of $\boldsymbol{\mu}_k$ is equal to the number of points assigned to cluster k . The mean of cluster k is essentially the same as the mean of data points \mathbf{x}_n assigned to cluster k . For this reason, the procedure is known as the *K-means clustering algorithm*.

The two phases of re-assigning data points to clusters and re-computing the cluster means are repeated in turn until there is no further change in the assignments. These two phases reduce the value of the objective function J , so the convergence of the algorithm is assured. However, it may converge to a local rather than global minimum of J .

We can also sequentially update the μ_k as follows:

$$\mu_{k+1} = \mu_k + \eta(\mathbf{x}_k - \mu_k)$$

There are some properties to note:

- It is a hard clustering algorithm (\leftrightarrow soft clustering)
- It is sensitive to the initialization of centroid.
- The number of clusters is uncertain.
- Sensitive to distance metrics (*e.g.*, Euclidean?)

22.2 Gaussian Mixture Models

K-means clustering is a form of hard clustering, where each data point is assigned to exactly one cluster. However, in some cases, soft clustering—where data points can belong to multiple clusters with varying degrees of membership—provides a better model in practice. A Gaussian Mixture Model (GMM) assumes a linear superposition of Gaussian components, offering a richer class of density models than a single Gaussian distribution.

In essence, rather than assuming that all data points are generated by a single Gaussian distribution, we assume that the data is generated by a mixture of K different Gaussian distributions, where each Gaussian represents a different component in the mixture.

For a single sample, the Gaussian Mixture Model can be expressed as a weighted sum of these individual Gaussian distributions:

$$\begin{aligned} p(\mathbf{x}) &= \sum_{\mathbf{z}} p(\mathbf{z})p(\mathbf{x}|\mathbf{z}) \\ &= \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \end{aligned}$$

Here, \mathbf{x} is a data point, π_k represents the mixing coefficients, $\mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$ is a Gaussian distribution with mean $\boldsymbol{\mu}_k$ and covariance $\boldsymbol{\Sigma}_k$, and K is the number of Gaussian components.

A key quantity in GMMs is the conditional probability of \mathbf{z} given \mathbf{x} , denoted as $p(z_k = 1|\mathbf{x})$ or $\gamma(z_k)$. This is also known as the responsibility or assignment probability, which represents the probability that a given data point \mathbf{x} belongs to component k of the mixture. Essentially, this can be thought of as the **classification result** for \mathbf{x} .

This responsibility is updated using Bayes' Theorem, and can be expressed as:

$$\begin{aligned} \gamma(z_k) &\equiv p(z_k = 1|\mathbf{x}) \equiv \frac{p(z_k = 1)p(\mathbf{x}|z_k = 1)}{\sum_{j=1}^K p(z_j = 1)p(\mathbf{x}|z_j = 1)} \\ &= \frac{\pi_k \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)} \end{aligned}$$

In this expression, π_k is the prior probability (or mixing coefficient) for component k , and $\mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$ is the likelihood of the data point \mathbf{x} under the Gaussian distribution corresponding to component k . The denominator is a normalization factor that ensures the responsibilities sum to 1 across all components for a given data point.

This framework allows for a soft classification of data points, where each point is associated with a probability of belonging to each cluster, rather than being strictly assigned to a single cluster as in K-means.

22.2.1 Maximum Likelihood

Suppose we have a data set of observations $\mathbf{X} = \{\mathbf{x}_1, \dots, \mathbf{x}_n\}^T \in \mathbb{R}^{N \times D}$ and we want to model the data distribution $p(\mathbf{X})$ using GMM. Assuming the data is independent and identically distributed

(i.i.d.), the likelihood of the entire dataset can be expressed as:

$$p(\mathbf{X}|\boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \prod_{n=1}^N \left(\sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \right).$$

To simplify the optimization process, we consider the log-likelihood function, which is given by:

$$\ln p(\mathbf{X}|\boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \sum_{n=1}^N \ln \left(\sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \right)$$

To solve the Maximum Likelihood Estimation (MLE) for Gaussian Mixture Models (GMMs), we typically consider the iterative *Expectation-Maximization* (EM) algorithm due to the non-convex nature of the problem. Before discussing how to maximize the likelihood, it is important to emphasize two significant issues that arise in GMMs: *singularities* and *identifiability*.

Singularity A major challenge in applying the maximum likelihood framework to Gaussian Mixture Models is the presence of singularities. This problem arises because the likelihood function can become unbounded under certain conditions, leading to an ill-posed optimization problem.

For simplicity, consider a Gaussian mixture model where each component has a covariance matrix of the form $\boldsymbol{\Sigma}_k = \sigma_k^2 I$, where I is the identity matrix. Suppose one of the mixture components, say the j -th component, has its mean $\boldsymbol{\mu}_j$ exactly equal to one of the data points \mathbf{x}_n , so that $\boldsymbol{\mu}_j = \mathbf{x}_n$ for some value of n . The contribution of this data point to the likelihood function would then be:

$$\mathcal{N}(\mathbf{x}_n | \mathbf{x}_n, \sigma_j^2 I) = \frac{1}{\sqrt{2\pi}\sigma_j} \cdot \exp^0$$

As σ_j approaches 0, this term goes to infinity, causing the log-likelihood function to also diverge to infinity. Therefore, maximizing the log-likelihood function becomes an ill-posed problem because such singularities can always be present. These singularities occur whenever one of the Gaussian components **collapses** onto a specific data point, leading to a covariance matrix with a determinant approaching zero. This issue did not arise with a single Gaussian distribution because the variance cannot be zero by definition.

Identifiability Another issue in finding MLE solutions for GMMs is related to identifiability. For any given maximum likelihood solution, a GMM with K components has a total of $K!$ equivalent solutions. This arises from the fact that the $K!$ different ways of permuting the K sets of parameters (means, covariances, and mixing coefficients) yield the same likelihood.

In other words, for any point in the parameter space that represents a maximum likelihood solution, there are $K! - 1$ additional points that produce exactly the same probability distribution. This lack of identifiability means that the solution is not unique, complicating both the interpretation of the model and the optimization process.

22.2.2 Expectation Maximization for GMM

The goal of the Expectation-Maximization (EM) algorithm is to find maximum likelihood solutions for models that involve latent variables.

- Suppose that directly optimizing the likelihood $p(\mathbf{X}|\boldsymbol{\theta})$ is difficult.
- However, it is easier to optimize the complete-data likelihood function $p(\mathbf{X}, \mathbf{Z}|\boldsymbol{\theta})$ as discussed in the previous sections.
- In such cases, we can use the **EM algorithm**. EM algorithm is a general technique for finding maximum likelihood solutions in latent variable models.

Let's begin by writing down the conditions that must be satisfied at a maximum of the likelihood function. By setting the derivatives of $\ln p(\mathbf{X}|\boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma})$ with respect to the means $\boldsymbol{\mu}_k$ of the Gaussian components to zero, we obtain

$$0 = - \sum_{n=1}^N \frac{\pi_k \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)} \boldsymbol{\Sigma}_k (\mathbf{x}_n - \boldsymbol{\mu}_k)$$

Multiplying by $\boldsymbol{\Sigma}_k^{-1}$ (which we assume to be non-singular) and rearranging we obtain

$$\boldsymbol{\mu}_k = \frac{1}{N_k} \sum_{n=1}^N \gamma(z_{nk}) \mathbf{x}_n,$$

where we have defined

$$N_k = \sum_{n=1}^N \gamma(z_{nk}).$$

We can interpret N_k as the effective number of points assigned to cluster k . We can obtain the MLE solutions for other variables similarly.

Algorithm 2: EM algorithm for GMM

Initialize the means $\boldsymbol{\mu}_k$, covariances $\boldsymbol{\Sigma}_k$ and mixing coefficients π_k and evaluate the initial value of the log likelihood.

for n **do**

E-step: evaluate the responsibilities of \mathbf{x}_n based on the current parameter values with the given parameters

$$\gamma(z_{nk}) = p(z_k = 1 | \mathbf{x}_n) = \frac{\pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)}$$

where z_{nk} denote the k -th component of \mathbf{z}_n

M-step: maximize expectation

• $\boldsymbol{\mu}_k^{\text{new}} = \frac{1}{N_k} \sum_{n=1}^N \gamma(z_{nk}) \mathbf{x}_n$

• $\boldsymbol{\Sigma}_k^{\text{new}} = \frac{1}{N_k} \sum_{n=1}^N \gamma(z_{nk}) (\mathbf{x}_n - \boldsymbol{\mu}_k^{\text{new}})(\mathbf{x}_n - \boldsymbol{\mu}_k^{\text{new}})^T$

• $\pi_k^{\text{new}} = p(z_k = 1) = \frac{N_k}{N}$

Evaluate the log likelihood to check for convergence of parameters

$$\ln p(\mathbf{X}|\boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \sum_{n=1}^N \ln \left(\sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \right)$$

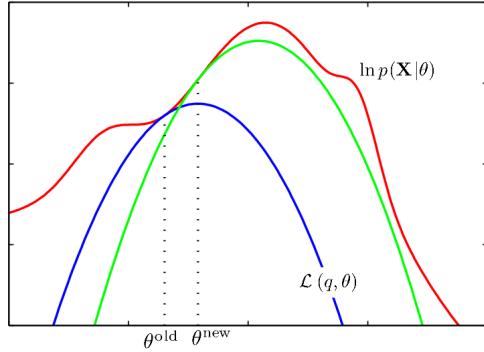


Figure 22.1: M-step of EM algorithm

22.3 Alternative View of EM

The goal of the EM algorithm is to find maximum likelihood (loglikelihood) solutions for models having latent variables.

$$\ln p(\mathbf{X}|\theta) = \ln \sum_Z p(\mathbf{X}, \mathbf{Z}|\theta).$$

We are not given the complete data set X, Z , but only the incomplete data X . Our state of knowledge of the values of the latent variables in Z is given only by the posterior distribution $p(Z|X, \theta)$. Because we cannot use the complete-data log likelihood function, we consider instead its expected value under the posterior distribution of the latent variable, which corresponds (as we shall see) to the E-step of the EM algorithm.

In the subsequent M-step, we maximize this expectation. If the current estimate for the parameters is denoted θ_{old} , then a pair of successive E- and M-steps gives rise to a revised estimate θ^{new} .

The algorithm is initialized by choosing some starting value for the parameters θ_0 . The use of the expectation may seem somewhat arbitrary.

In the E-step, we use the current parameter values θ^{old} to find the posterior distribution of the latent variables given by $p(Z|X, \theta^{\text{old}})$. We then use this posterior distribution to find the **expectation of the complete-data log likelihood function** evaluated for some general parameter value θ . In other words, this is a expectation over a some function. This expectation, denoted $Q(\theta, \theta^{\text{old}})$, is given by

$$Q(\theta, \theta^{\text{old}}) = \sum_Z p(Z|X, \theta^{\text{old}}) \ln p(\mathbf{X}, \mathbf{Z}|\theta).$$

In the M-step, we determine the revised parameter estimate θ^{new} by maximizing this function

$$\theta^{\text{new}} = \operatorname{argmax} Q(\theta, \theta^{\text{old}}).$$

Algorithm 3: General EM algorithm

The goal is to maximize the likelihood function $p(X|\theta)$ with respect to θ given a joint distribution $p(X, Z|\theta)$.

1. Init θ^{old}
2. E-Step: evaluate $p(Z|X, \theta^{old})$
3. M-Step: evaluate θ^{new} given by

$$\theta^{new} = \operatorname{argmax} Q(\theta, \theta^{old}),$$

where

$$Q(\theta, \theta^{old}) = \sum_Z p(Z|X, \theta^{old}) \ln p(X, Z|\theta).$$

4. Check for convergence of either the log likelihood or the parameter values. If the convergence criterion is not satisfied, then let

$$\theta^{old} \leftarrow \theta^{new}.$$

Return to the step 2.

22.4 Latent Variable Modeling

For each object x_i , we establish additional latent variable z_i which denotes the index of Gaussian from which i -th object was generated. Then our model is

$$p(X, Z|\theta) = \prod_{i=1}^n p(x_i, z_i|\theta) = \prod_{i=1}^n p(x_i|z_i, \theta)p(z_i|\theta) = \prod_{i=1}^n \mathcal{N}(x_i|\mu_{z_i}, \sigma_{z_i}^2)\pi_{z_i},$$

where $\pi_j = p(z_i = j)$ are prior probability of j -th gaussian and $\theta = \{\mu_j, \sigma_j, \pi_j\}_{j=1}^K$. If we know both X and Z then we can obtain explicit ML-solution:

$$\theta_{ML} = \underset{\theta}{\operatorname{argmax}} p(X, Z|\theta) = \underset{\theta}{\operatorname{argmax}} \log p(X, Z|\theta).$$

However, in practice, we don't know Z , but only know X . Thus, we need to maximize w.r.t. θ the log of incomplete likelihood

$$\log p(X|\theta) = \ln \int p(X, Z|\theta) dZ \tag{22.1}$$

$$= \ln \int q(Z|X) \frac{p(X, Z|\theta)}{q(Z|X)} dZ \tag{22.2}$$

$$\geq \underbrace{\int q(Z|X) \ln \frac{p(X, Z|\theta)}{q(Z|X)} dZ}_{\text{ELBO, } \mathcal{L}(q, \theta)} \quad \text{by Jensen's Inequality.} \tag{22.3}$$

$$= \int q(Z|X) \ln p(X, Z|\theta) - q(Z|X) \ln q(Z|X) dZ \tag{22.4}$$

$$= \int q(Z|X) [\ln p(X|Z, \theta) + \ln p(Z|\theta)] - q(Z|X) \ln q(Z|X) dZ \tag{22.5}$$

$$= \int q(Z|X) \ln p(X|Z, \theta) - q(Z|X) \ln \frac{q(Z|X)}{p(Z|\theta)} dZ \tag{22.6}$$

$$= \mathbb{E}_{q(Z|X)} \ln p(X|Z, \theta) - KL(q(Z|X) \| p(Z|\theta)) \tag{22.7}$$

To maximize the above equation, we need to minimize KL divergence.

To get more intuition about ELBO, we can express ELBO as follows:

$$\begin{aligned} \mathcal{L}(\phi, \theta) &= \mathbb{E}_{q_\phi(z|x)} \left[\log \frac{p_\theta(x, z)}{q_\phi(z|x)} \right] \\ &= \mathbb{E}_{q_\phi(z|x)} \left[\log p_\theta(x, z) - \log q_\phi(z|x) \right] \\ &= \mathbb{E}_{q_\phi(z|x)} \left[\log p_\theta(x) + \log p_\theta(z|x) - \log q_\phi(z|x) \right] \\ &= \log p_\theta(x) - D_{KL}(q_\phi(z|x) \| p_\theta(z|x)) \\ &\leq \log p_\theta(x) \end{aligned}$$

We can get a conclusion that maximizing ELBO is equivalent to minimizing the KL divergence through the above equation. Finally, the log-likelihood can be rewritten as follows:

$$\log p_\theta(x) = \mathcal{L}(\phi, \theta) + D_{KL}(q_\phi(z|x) \| p_\theta(z|x))$$

22.4.1 Expectation Maximization

We want to maximize ELBO, $\mathcal{L}(q, \theta)$ to minimize KL divergence between $q(Z)$ and $\log p(Z|X, \theta)$.

$$\max_{q, \theta} \mathcal{L}(q, \theta) = \max_{q, \theta} \int q(Z) \log \frac{p(X, Z|\theta)}{q(Z)} dZ.$$

We start from initial point θ_0 and iteratively repeat (i) E-step and (ii) M-step, iteratively:

- E-Step: θ_0 is fixed.

$$q(Z) = \operatorname{argmax}_q \mathcal{L}(q, \theta) = \operatorname{argmin}_q \text{KL}(q(Z)|p(Z|X, \theta)) = p(Z|X, \theta_0).$$

- This is because, maximizing ELBO is equal to minimizing KL divergence and the minimum q can be achieved when q is equal to $p(Z|X, \theta_0)$.
- Now, we just have to evaluate $p(Z|X, \theta_0)$.

- M-Step: q is fixed.

$$\theta_* = \operatorname{argmax}_{\theta} \mathcal{L}(q, \theta) = \operatorname{argmax}_{\theta} \mathbb{E}_{q(Z)} [\log p(X, Z|\theta)]$$

- Can be accomplished by taking derivatives
- Set $\theta_0 = \theta_*$ and go to the E-Step until convergence

22.4.2 Categorical Latent Variables

$$z_i \in \{1, \dots, K\}$$

$$p(x_i|\theta) = \sum_{k=1}^K p(x_i|k, \theta)p(z_i = k|\theta)$$

is simply a finite mixture of distributions.

E-Step:

$$q(z_i = k) = p(z_i = k|x_i, \theta) = \frac{p(x_i|z_i = k, \theta)p(z_i = k|\theta)}{\sum_{l=1}^K p(x_i|z_i = l, \theta)p(z_i = l|\theta)}$$

M-Step:

$$\operatorname{argmax}_{\theta} \mathbb{E}_{q(Z)} [\log p(X, Z|\theta)] = \sum_{i=1}^n \mathbb{E}_{q(z_i)} [\log p(x_i, z_i|\theta)] = \sum_{i=1}^n \sum_{k=1}^K q(z_i = k) \log p(x_i, k|\theta)$$

For GMM, we model $p(x|z)$ as Gaussian.

Chapter 23

Hidden Markov Models

23.1 Introduction

The HMM is based on the Markov chain assumption. A Markov chain is a model that tells us something about the probabilities of sequences of random variables, states, each of which can take on values from some set. These sets can be words, or tags, or symbols representing anything, like the weather.

There are two important assumptions:

- Markov assumption
- Output independence: $p(x_i|z_1, \dots, z_i, \dots, z_T, x_1, \dots, x_i, \dots, x_T) = p(x_i|z_i)$

23.1.1 Conditional Independence

If two events A and B are **conditionally independent** given an event C then,

- $P(A \cap B|C) = P(A|C)P(B|C).$
- $P(A|B, C) = P(A|C)$

23.1.2 Notation

- $X = (x_1, x_2, \dots, x_T)$
- Initial state probabilities: $p(z_1) \sim \text{Multinomial}(\pi_1, \dots, \pi_k)$, need to learn π
- Transition probability:

$$p(z_t|z_{t-1} = i) \sim \text{Multinomial}(a_{i,1}, \dots, a_{i,k})$$

, where $a_{i,j} = p(z_t = j|z_{t-1} = i)$ and i and j denote clusters or states, respectively.

- Emission probability:

$$p(x_t|z_t = i) \sim \text{Multinomial}(b_{i,1}, \dots, b_{i,m})$$

, where $b_{i,j} = p(x_t = j|z_t = i)$

23.2 Bayesian Network

23.2.1 Bayes Ball

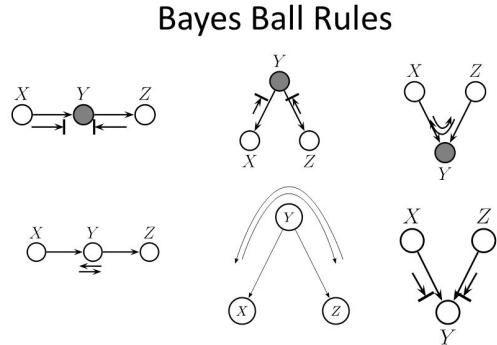


Figure 23.1: Bayes ball

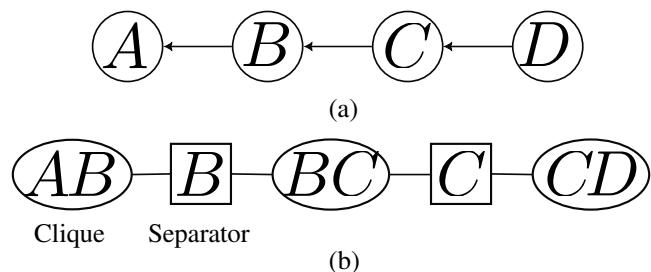
- Cascading: $P(Z|Y, X) = P(Z|Y)$. The information of Y decouples X and Z .
- Common parent: $P(X, Z|Y) = P(X|Y)P(Z|Y)$. The information of Y decouples X and Z .
- V-Structure (common child): Unlike the above two cases, the information of Y couples X and Z .

$$P(X, Y, Z) = P(X)P(Y)P(Y|X, Z).$$

23.2.2 Potential Function

Potential function is a function which is not a probability function, but it can become a probability function by normalizing it.

$$P(A, B, C, D) = P(A|B)P(B|C)P(C|D)P(D)$$



- Cliques: $\Psi(a, b), \Psi(b, c), \Psi(c, d)$
- Separators $\phi(b), \phi(c)$

Given a clique tree with cliques and separators, the joint probability distribution is defined as follows:

$$P(A, B, C, D) = P(U) = \frac{\prod_N \Psi(N)}{\prod_L \phi(L)} = \frac{\Psi(a, b)\Psi(b, c)\Psi(c, d)}{\phi(b)\phi(c)}$$

An effect of an observation propagates through the clique graph → **Belief propagation**. How to propagate the belief? **Absorption rule!**

Let's say we have some new observations about A , then it affects the clique $\Psi(a, b)$. The updated clique is now $\Psi^*(a, b)$. Similarly, $\phi^*(b) = \sum_A \Psi^*(a, b)$. Subsequently, $\Psi^*(b, c) = \Psi(b, c) \frac{\phi^*(b)}{\phi(b)}$.

23.3 Hidden Markov Models

Hidden Markov Models (HMMs) are a powerful statistical tool used for modeling sequences of observable events that are believed to be generated by underlying hidden states. The core idea behind HMMs is that the system being modeled can be represented as a **Markov process with unobservable (hidden) states z , where each state emits observable outputs x according to specific probability distributions**.

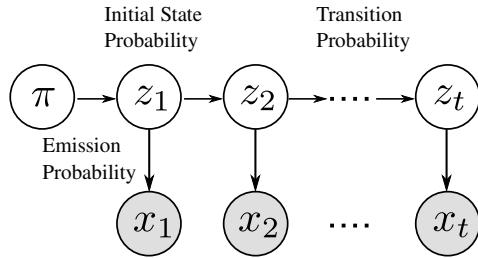


Figure 23.2: HMM Structure

In an HMM, the hidden states are connected by transition probabilities, which dictate the likelihood of moving from one state to another. Meanwhile, each state is associated with an emission probability, which defines the likelihood of observing a particular output given that state. The challenge and beauty of HMMs lie in their ability to infer the most probable sequence of hidden states that could have produced a given sequence of observations, even when the underlying states are not directly observable.

Through algorithms such as the Forward-Backward algorithm, the Viterbi algorithm, and the Baum-Welch algorithm, HMMs provide a framework for decoding sequences, estimating model parameters, and predicting future events based on past observations. In this post, we'll explore the fundamentals of Hidden Markov Models, their key components, and how they can be applied to solve real-world problems.

It's important to note that the observations in an HMM can be either discrete or continuous. When the latent factors are continuous, the model is often referred to as a *Kalman filter*.

23.3.1 Key Components of HMM

The key components of an HMM can be listed as follows:

- Initial state probability: $P(z_1) \sim \text{Mult}(\pi_1, \dots, \pi_k)$. In other words, z is belong to one of k classes.
- Transition probability: $P(z_t|z_{t-1}^i = 1) \sim \text{Mult}(a_{i,1}, \dots, a_{i,k})$,
where $P(z_t^j = 1|z_{t-1}^i = 1) = a_{i,j}$
- Emission probability: $P(x_t|z_t^i = 1) \sim \text{Mult}(b_{i,1}, \dots, b_{i,m}) \sim f(x_t|\theta_i)$,
where $P(x_t^j = 1|z_t^i = 1) = b_{i,j}$. The probability of observing x_j at the i -th cluster.

Note that i and j are indices of clusters (*e.g.*, classes).

There are three main problems in HMM:

1. Evaluation Questions (likelihood):

- Given $\pi, \mathbf{a}, \mathbf{b}, X$
- Find $p(X|M, \pi, \mathbf{a}, \mathbf{b})$
- How much are X likely to be observed by a model M ?

2. Decoding Questions:

- Given $\pi, \mathbf{a}, \mathbf{b}, X$
- Find $\text{argmax}_Z p(Z|X, M, \pi, \mathbf{a}, \mathbf{b})$
- What is the most probable sequence of Z (latent states)?

3. Learning Questions: Forward-Backward (Baum-Welch)

- Given X
- Find $\text{argmax}_{\pi, \mathbf{a}, \mathbf{b}} p(X|M, \pi, \mathbf{a}, \mathbf{b})$
- What would be the optimal model parameters?

23.4 Evaluation: Forward-Backward Probability

23.4.1 Joint Probability

We can factorize the joint distribution of HMM in Fig. 23.2 by using a Bayesian approach as follows::

$$p(X, Z) = p(x_1, \dots, x_t, z_1, \dots, z_t) \quad (23.1)$$

$$= p(z_1)p(x_1|z_1)p(z_2|z_1), \dots, p(x_t|z_t)p(z_t|z_{t-1}) \quad (23.2)$$

The key assumption involved in factorizing the Markov chain within a Hidden Markov Model (HMM) is *conditional independence* among certain components of the state variables. Here's a detailed breakdown of what this assumption means:

- Independence of State Components: The transition of each component z_t^k only depends on its corresponding previous component z_{t-1}^k and is independent of other components.

As the number of latent factor increases, it is getting harder to decode the latent factors.

23.4.2 Marginal Probability

We want to compute the likelihood of sequence X which is given by

$$p(X|\pi, \mathbf{a}, \mathbf{b}) = \sum_Z p(X, Z|\pi, \mathbf{a}, \mathbf{b})$$

The computation can be done as follows:

$$\begin{aligned} p(X) &= \sum_Z p(X, Z) \\ &= \sum_{z_1} \cdots \sum_{z_t} p(x_1, \dots, x_t, z_1, \dots, z_t) \\ &= \sum_{z_1} \cdots \sum_{z_t} \pi_{z_1} \prod_{t=2}^T a_{z_{t-1}, z_t} \prod_{t=1}^T b_{z_t, x_t} \end{aligned}$$

The last step is done by using the factorization above. The computation of this equation requires lots of computations, so we will change it into a **recursive form** by using the factorization rule

$$p(a, b, c) = p(a)p(b|a)p(c|a, b).$$

$$p(x_1, \dots, x_t, z_t^k = 1) = \sum_{z_{t-1}} p(x_1, \dots, x_{t-1}, x_t, z_{t-1}, z_t^k = 1) \quad (23.3)$$

$$= \sum_{z_{t-1}} p(\underbrace{x_1, \dots, x_{t-1}}_a, \underbrace{z_{t-1}}_c, \underbrace{x_t}_b, \underbrace{z_t^k = 1}_b) \quad (23.4)$$

$$= \sum_{z_{t-1}} p(x_1, \dots, x_{t-1}, z_{t-1}) p(z_t^k = 1 | x_1, \dots, x_{t-1}, z_{t-1}) p(x_t | z_t^k = 1, x_1, \dots, x_{t-1}, z_{t-1}) \quad (23.5)$$

$\because p(a, b, c) = p(a)p(b|a)p(c|a, b)$ or by the structure of HMM

$$= \sum_{z_{t-1}} p(x_1, \dots, x_{t-1}, z_{t-1}) p(z_t^k = 1 | z_{t-1}) p(x_t | z_t^k = 1) \quad (23.6)$$

$$= p(x_t | z_t^k = 1) \sum_{z_{t-1}} p(x_1, \dots, x_{t-1}, z_{t-1}) p(z_t^k = 1 | z_{t-1}) \quad (23.7)$$

$$= b_{z_t^k, x_t} \sum_{z_{t-1}} p(x_1, \dots, x_{t-1}, z_{t-1}) a_{z_{t-1}, z_t^k} \quad (23.8)$$

- In the second line, the x_{t-1} and z_{t-1} are grouped together.
- Then, we can find the HMM structure by factorizing the equation.
- In the fourth line, x terms are removed, since z_t only relies on z_{t-1} by the Markov assumption. Similarly, x_t only depends on z_t . We can interpret this by using Bayes ball too.

Now we can find a recursive structure of $p(x_1, \dots, x_t, z_t^k = 1)$ as follows:

$$\alpha_t^k = p(x_1, \dots, x_t, z_t^k = 1) = b_{k, x_t} \sum_i \alpha_{t-1}^i a_{i, k}$$

, where α_t^k is the probabilities of being in state k after observing the first t observations. Thus,

$$\begin{aligned} p(x_1, \dots, x_t) &= \sum_z p(x_1, \dots, x_t, z) \\ &= \sum_k \alpha_t^k \end{aligned}$$

Note that α_t^k is also called **Forward probability**.

23.4.3 Forward Algorithm

Forward probability solves the evaluation problem. Essentially, this is a dynamic programming, so it calculates required values in a bottom-up manner.

- Forward probability: α_t^k , Time \times States

Note again that

$$p(X) = p(x_1, \dots, x_T) = \sum_i \alpha_T^i = \sum_i p(x_1, \dots, x_T, z_T^i = 1)$$

Note also that the forward-algorithm returns $p(X)$ and forward probability is the probability of being in state k after observing the first t observations without Z .

Algorithm 4: Forward Algorithm

Create a probability matrix $forward[M, T] = \alpha_t^k$

Initialization:

for each state $k=1, \dots, M$ **do**

$$\quad \quad \quad \alpha_1^k \leftarrow \pi_k b_{k,x_1}$$

for time step $t=2, \dots, T$ **do**

for each step $k=1, \dots, M$ **do**

$$\quad \quad \quad \alpha_t^k = b_{k,x_t} \sum_i \alpha_{t-1}^i a_{i,k}$$

$$\text{Return } p(X) = \sum_i^M \alpha_T^i$$

23.4.4 Backward Probability

The forward probability only considers an observation at t . To determine the z_t , we need to leverage the future observations. **The backward probability β is the probability of seeing the observations from time $t + i$ to the end, given that we are in state k at time t .**

$$\beta_t^k = p(x_{t+1}, \dots, x_T | z_t^k = 1)$$

We want to compute $p(z_t^k = 1 | X)$ rather than $p(x_1, \dots, x_t, z_t^k = 1)$. In other words, we will leverage the whole observations X .

$$\begin{aligned} p(z_t^k = 1, X) &= p(x_1, \dots, x_t, z_t^k = 1, x_{t+1}, \dots, x_T) \\ &= p(x_1, \dots, x_t, z_t^k = 1)p(x_{t+1}, \dots, x_T | x_1, \dots, x_t, z_t^k = 1) \\ &= p(x_1, \dots, x_t, z_t^k = 1)p(x_{t+1}, \dots, x_T | z_t^k = 1) \\ &= \alpha_t^k \beta_t^k \end{aligned}$$

We already know that $p(x_1, \dots, x_t, z_t^k = 1) = \alpha_t^k$. We just need to compute backward probability as follows:

$$\begin{aligned} \beta_t^k &= p(x_{t+1}, \dots, x_T | z_t^k = 1) \\ &= \sum_{z_{t+1}} p(\underbrace{z_{t+1}}, \underbrace{x_{t+1}}, \underbrace{x_{t+2}, \dots, x_T}_{c} | z_t^k = 1) \\ &= \sum_i p(z_{t+1}^i = 1 | z_t^k = 1)p(x_{t+1} | z_{t+1}^i = 1, z_t^k = 1)p(x_{t+2}, \dots, x_T | x_{t+1}, z_{t+1}^i = 1, z_t^k = 1) \\ &\because p(a, b, c) = p(a)p(b|a)p(c|a, b) \\ &= \sum_i p(z_{t+1}^i = 1 | z_t^k = 1)p(x_{t+1} | z_{t+1}^i = 1)p(x_{t+2}, \dots, x_T | z_{t+1}^i = 1) \\ &= \sum_i a_{k,i} b_{i,x_{t+1}} \beta_{t+1}^i \end{aligned}$$

Another recursive structure:

$$\begin{aligned} p(z_t^k = 1, X) &= \alpha_t^k \beta_t^k \\ &= b_{k,x_t} \sum_i \alpha_{t-1}^i a_{i,k} \times \sum_i a_{k,i} b_{i,x_{t+1}} \beta_{t+1}^i \end{aligned}$$

This means at time t , the latent label is belong to some class k and this can be computed by using the forward probability and the backward probability. Now we can compute

$$p(z_t^k = 1 | X) = \frac{p(z_t^k = 1, X)}{p(X)} = \frac{\alpha_t^k \beta_t^k}{p(X)}$$

Then,

$$k_t = \operatorname{argmax}_k p(z_t^k = 1 | X)$$

Note that this is for a single latent variable at a single time step given the whole observation X , but we want to decode a sequence of latent variables. Thus, we need some decoding algorithm.

23.5 Decoding: Viterbi Algorithm

For any model, such as an HMM, that contains hidden variables, **the task of determining which sequence of variables is the underlying source of some sequence of observations is called the decoding task.**

We might propose to find the best sequence as follows:

1. For each possible hidden state sequence (HHH, HHC, HCH, etc.), we could run the forward algorithm and compute the likelihood of the observation sequence given that hidden state sequence.
2. Then, we could choose the hidden state sequence with the maximum observation likelihood.

However, this is not a feasible solution, because there are an exponentially large number of state sequences.

Instead, the most common decoding algorithms for HMMs is the **Viterbi algorithm**. Like the forward algorithm, **Viterbi** is a kind of **dynamic programming algorithm**.

Note that the Viterbi algorithm is identical to the forward algorithm except that it takes the **max** over the previous path probabilities whereas the forward algorithm takes the **sum**. This is because, we want to obtain **the most probable latent variable sequence**. Note also that the Viterbi algorithm has one component that the forward algorithm doesn't have: **backpointers**. The reason is that while the forward algorithm needs to produce an observation likelihood, the Viterbi algorithm must produce a probability and also the most likely state sequence. We compute this best state sequence by keeping track of the path of hidden states that led to each state and then at the end backtracing the best path to the beginning (the Viterbi backtrace).

We can leverage the forward-backward probabilities:

$$\bullet k^* = \operatorname{argmax}_k p(z_t^k = 1 | X) = \operatorname{argmax}_k p(z_t^k = 1, X) = \operatorname{argmax}_k \alpha_t^k \beta_t^k$$

We will use a forward approach:

$$V_t^k = \max_{z_1, \dots, z_{t-1}} p(x_1, \dots, x_{t-1}, z_1, \dots, z_{t-1}, x_t, z_t^k = 1) \quad (23.9)$$

$$= \max_{z_1, \dots, z_{t-1}} p(x_t, z_t^k = 1 | x_1, \dots, x_{t-1}, z_1, \dots, z_{t-1}) p(x_1, \dots, x_{t-1}, z_1, \dots, z_{t-1}) \quad (23.10)$$

$$= \max_{z_1, \dots, z_{t-1}} p(x_t, z_t^k = 1 | z_{t-1}) p(x_1, \dots, x_{t-2}, z_1, \dots, z_{t-2}, x_{t-1}, z_{t-1}) \quad (23.11)$$

$$= \max_{z_{t-1}} p(x_t, z_t^k = 1 | z_{t-1}) \max_{z_1, \dots, z_{t-2}} p(x_1, \dots, x_{t-2}, z_1, \dots, z_{t-2}, x_{t-1}, z_{t-1}) \quad (23.12)$$

$$= \max_{i \in z_{t-1}} p(x_t, z_t^k = 1 | z_{t-1}^i = 1) V_{t-1}^i \quad (23.13)$$

$$= \max_{i \in z_{t-1}} p(x_t | z_t^k = 1) p(z_t^k = 1 | z_{t-1}^i = 1) V_{t-1}^i \quad (23.14)$$

$$= p(x_t | z_t^k = 1) \max_{i \in z_{t-1}} p(z_t^k = 1 | z_{t-1}^i = 1) V_{t-1}^i \quad (23.15)$$

$$= b_{k, x_t} \max_{i \in z_{t-1}} a_{i, k} V_{t-1}^i \quad (23.16)$$

- V_t^k is Viterbi variable which denotes the probability that the HMM is in state k at t after observing the first t observations and $t - 1$ latent variables. In another words, this is the probability of most likely sequence of states ending at state $z_t = k$.
- The first line assumes that the observation at time t and the latent variable are fixed and also the fourth line has the recursive structure.
- The third step, only z_{t-1} can affect the z_t , so we can remove all other unnecessary variables.
- The step six can be derived by the HMM structure.
- $i \in z_{t-1}$ simply denotes the index of potential cluster at $t - 1$.
- We have already computed the backward and the forward probabilities. So we just need to apply the Viterbi algorithm.

Note that Also note that we present the most probable path by taking the maximum over all possible previous state sequences $\max_{z_1, \dots, z_{t-1}}$. Like other DP-algorithm, Viterbi fills each cell recursively.

Algorithm 5: Viterbi Algorithm

```

 $V_t^k = \text{viterbi}[M, T]$ , where  $M$  is the number states
for  $k=1, \dots, M$  do
   $V_1^k \leftarrow \pi_{z_k} b_{k,x_1}$ 
   $\text{backpointer}[k, 1] \leftarrow 0$ 
for  $t=2, \dots, T$  do
  for  $k=1, \dots, M$  do
     $V_t^k \leftarrow b_{k,x_t} \max_{k'} V_t^{k'} a_{k',k}$ , where  $k'$  is the previous state.
     $\text{backpointer}[k, t] \leftarrow b_{k,x_t} \arg\max_{k'} V_t^{k'} a_{k',k}$ 
bestpathprob  $\leftarrow \max_k V_T^k$  //termination step
bestpathpointer  $\leftarrow \arg\max_k V_T^k$  //termination step
bestpath  $\leftarrow$  the path starting at state bestpathpointer, that follows backpointer[] to
  states back in time
Return bestpathpointer, bestpathprob

```

Viterbi algorithm typically shows some technical issues:

- Underflow problems $\rightarrow \log V$.

23.6 Learning: Baum-Welch Algorithm

We have to learn HMM parameters with only X . Baum-Welch algorithm or Forward-Backward Algorithm is a standard training algorithm for HMM. The algorithm let us train both the transition and the emission probabilities of the HMM. If we do not have the information about Z , then we can assign the most probable Z given X .

- Given X , estimate parameters π, a, b .
- Then, find the most probable Z given the parameters.

We will use EM algorithm!

23.6.1 EM Algorithm

$$P(X|\theta) = \sum_Z P(X, Z|\theta) \rightarrow \ln P(X|\theta) = \ln \sum_Z P(X, Z|\theta).$$

We cannot directly estimate the log-likelihood function, so we will estimate the expectation of it.

$$\begin{aligned} Q(\theta, \theta^{old}) &= \mathbb{E}_Z \ln P(X, Z|\theta) \\ &= \sum_Z p(Z|X, \theta^{old}) \ln P(X, Z|\theta) \\ &= \sum_Z p(Z|X, \pi^t, a^t, b^t) \ln P(X, Z|\pi, a, b). \end{aligned}$$

Note that $p(X, Z) = \pi_{z_1} \prod_{t=2}^T a_{z_{t-1}, z_t} \prod_{t=2}^T b_{z_t, x_t}$. Thus, $\ln p(X, Z) = \ln \pi_{z_1} + \sum_{t=2}^T \ln a_{z_{t-1}, z_t} + \sum_{t=1}^T \ln b_{z_t, x_t}$. Therefore

$$Q(\theta, \theta^{old}) = \sum_Z p(Z|X, \theta^{old}) \left(\ln \pi_{z_1} + \sum_{t=2}^T \ln a_{z_{t-1}, z_t} + \sum_{t=1}^T \ln b_{z_t, x_t} \right).$$

To optimize the above function we will use the Lagrange method as follows:

$$\mathcal{L}(\pi, a, b) = Q(\theta, \theta^{old}) - \lambda_\pi \left(\sum_{i=1}^K \pi_i - 1 \right) - \sum_i \lambda_{a_i} \left(\sum_{j=1}^K a_{i,j} - 1 \right) - \sum_i \lambda_{b_i} \left(\sum_{j=1}^K b_{i,j} - 1 \right).$$

The constraints are for forcing the sum of each probability is equal to 1.

Now, take a partial derivative for each parameter. Let's take a derivative with regard to π_i first. Then,

$$\begin{aligned} \frac{\partial \mathcal{L}}{\partial \pi_i} &= \frac{\partial Q(\theta, \theta^{old})}{\partial \pi_i} - \lambda_\pi \\ &= \frac{\partial}{\partial \pi_i} \sum_Z p(Z|X, \theta^{old}) \ln \pi_{z_1} - \lambda_\pi \\ &= \frac{p(z_1^i = 1|X, \theta^{old})}{\pi_i} - \lambda_\pi \\ \frac{\partial \mathcal{L}}{\partial \lambda_{\pi_i}} &= \sum_{i=1}^K \pi_i - 1 = 0 \rightarrow \sum_{i=1}^K \pi_i = 1. \end{aligned}$$

By setting the derivative is equal to zero,

$$\pi_i = \frac{p(z_1^i = 1 | X, \theta^{old})}{\lambda_\pi}.$$

By using the constraint of π , the Lagrange multiplier λ_π must be a normalizer.

$$\pi_i = \frac{p(z_1^i = 1 | X, \theta^{old})}{\sum_{j=1}^K p(z_1^j = 1 | X, \theta^{old})}.$$

Similarly, we can compute other parameters too.

$$a_{i,j}^{t+1} = \frac{\sum_{t=2}^T p(z_{t-1}^i = 1, z_t^j = 1 | X, \theta^{old})}{\sum_{t=2}^T p(z_{t-1}^i = 1 | X, \theta^{old})}.$$

$$b_{i,j}^{t+1} = \frac{\sum_{t=1}^T p(z_{t1}^i = 1 | X, \theta^{old}) I(x_t = j)}{\sum_{t=1}^T p(z_t^i = 1 | X, \theta^{old})},$$

where $I(x)$ is an indicator function which returns 1 if x is true and 0, otherwise.

23.7 Python Implementation

23.7.1 Viterbi Algorithm

The Viterbi algorithm is a dynamic programming algorithm used to determine the most probable sequence of hidden states in a Hidden Markov Model (HMM) based on a sequence of observations.

The algorithm works by recursively computing the probability of the most likely sequence of hidden states that ends in each state for each observation.

At each time step, the algorithm computes the probability of being in each state and emits the current observation based on the probabilities of being in the previous states and making a transition to the current state.

Assuming we have an HMM with N hidden states and T observations, the Viterbi algorithm can be summarized as follows:

Initialization: At time $t=1$, we set the probability of the most likely path ending in state i for each state i to the product of the initial state probability π_i and the emission probability of the first observation given state i . This is denoted by: $\delta[1,i] = \pi_i * b[i,1]$.

Recursion: For each time step t from 2 to T , and for each state i , we compute the probability of the most likely path ending in state i at time t by considering all possible paths that could have led to state i . This probability is given by:

$$\delta[t, i] = \max_j (\delta[t - 1, j] * a[j, i] * b[i, t])$$

Here, $a[j,i]$ is the probability of transitioning from state j to state i , and $b[i,t]$ is the probability of observing the t -th observation given state I .

We also keep track of the most likely previous state that led to the current state i , which is given by:

$$\psi[t, i] = \operatorname{argmax}_j (\delta[t - 1, j] * a[j, i])$$

- **Termination:** The probability of the most likely path overall is given by the maximum of the probabilities of the most likely paths ending in each state at time T . That is, $P_* = \max_i (\delta[T, i])$.
- **Backtracking:** Starting from the state i_* that gave the maximum probability at time T , we recursively follow the ψ values back to time $t = 1$ to obtain the most likely path of hidden states.

The Viterbi algorithm is an efficient and powerful tool that can handle long sequences of observations using dynamic programming.

23.8 Summary

- Forward-probability: probability of being in state k after observing the first t observations.

$$\alpha_t^k = p(x_1, \dots, x_t, z_t^k = 1)$$

- Backward-probability: probability of observations from time $t + 1$ to the end, given that we are in state k

$$\beta_t^k = p(x_{t+1}, \dots, x_T | z_t^k = 1)$$

- These two sets of probability distributions can then be combined to obtain the distribution over states at any specific point in time given the entire observation sequence

$$\begin{aligned} p(z_t^k = 1, X) &= p(x_1, \dots, x_t, z_t^k = 1, x_{t+1}, \dots, x_T) \\ &= p(x_1, \dots, x_t, z_t^k = 1)p(x_{t+1}, \dots, x_T | x_1, \dots, x_t, z_t^k = 1) \\ &= p(x_1, \dots, x_t, z_t^k = 1)p(x_{t+1}, \dots, x_T | z_t^k = 1) \\ &= \alpha_t^k \beta_t^k \end{aligned}$$

In short, if we know the forward and backward probability, we could know the cluster of state at time t given our observations.

- Forward-algorithm: return a marginal likelihood of the observed sequence
- Forward-backward: predict a single hidden state
- Viterbi: predict an entire sequence of hidden states
- Baum-Welch: unsupervised training (EM)

There are two shortcomings of HMM:

- HMM models capture dependences between each state and only its corresponding observation: Most NLP cases, many tasks needs not only local but also global feature (sentence level).
- Mismatch between learning objective function and prediction objective function: HMM learns a joint distribution of states and observations $p(Y, X)$, but we are more interested in $p(Y|X)$

Chapter 24

Explicit Generative Models

24.1 Variational Autoencoder

Our goal is to find the data distribution $p(X)$. Fig. 24.1 represents a general structure of deep generative model. As you can see, we first sample $z \sim p(z)$ and feed it into a deep neural network $f(z)$ and output x .

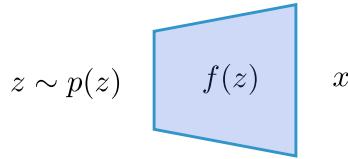


Figure 24.1: General structure of deep generative models. This model does not infer z from x .

VAE performs an inference by introducing a probabilistic encoder, called inference network. VAEs are generative model with a latent variable distributed according to some distribution $p(z_i)$. The observed variable is distributed according to a conditional distribution

$$p_\theta(x_i|z_i)$$

This conditioning means the latent variable values are the one most likely given the observations. We also create a distribution $q_\phi(z_i|x_i)$. We would like to be able to encode our data into the latent variable space. Let's model the distribution.

- $p_\theta(x_i|z_i) \sim \mathcal{N}(x_i|\mu(z_i), \sigma^2(z_i))$: A probabilistic decoder (or generative network, θ)
- $q_\phi(z_i|x_i)$: A probabilistic encoder (or inference network ϕ). We can choose a family of distributions for our conditional distribution q (e.g., standard Gaussian distribution).

$$q_\phi(z_i|x_i) = \mathcal{N}(z_i|\mu(x_i, W_1), \sigma^2(x_i, W_2)I),$$

where W_1 and W_2 are network weights and collectively denoted as ϕ . We create a neural network to model the distribution q from our data in a non-linear manner. The outputs of the network are μ and σ .

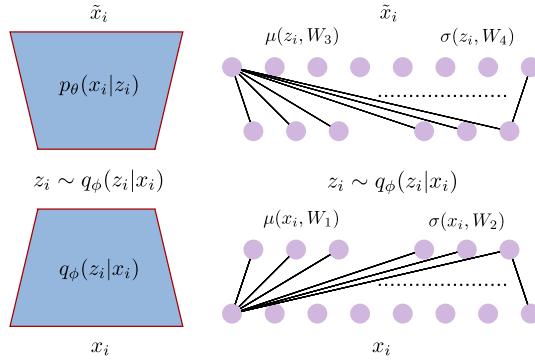


Figure 24.2: Overview of variational autoencoder.

$$\begin{aligned}
 p(X, Z|\theta) &= \prod_{i=1}^n \underbrace{p(x_i|z_i, \theta)}_{\text{Likelihood, Generator Prior on latent variable}} \underbrace{p(z_i|\theta)}_{\text{Non-linear}} \\
 &= \prod_{i=1}^n \mathcal{N}(x_i | \underbrace{\mu(z_i), \sigma^2(z_i)}_{\text{Non-linear}}) \mathcal{N}(z_i | 0, I)
 \end{aligned}$$

Subsequently, marginal distributions can be expressed as follows under i.i.d. assumption:

$$\begin{aligned}
 p(X|\theta) &= \prod_{i=1}^n p(x_i|\theta) \\
 &= \prod_{i=1}^n \int p(x_i, z_j|\theta) dz_j \\
 &= \prod_{i=1}^n \int p(x_i|z_i, \theta) p(z_i|\theta) dz_i \\
 &= \prod_{i=1}^n \int \mathcal{N}(x_i | \mu(z_i), \sigma^2(z_i)) \underbrace{\mathcal{N}(z_i | 0, I)}_{\text{Mixture weight}} dz_i
 \end{aligned}$$

- As you can see, the marginal distribution $p(X|\theta)$ becomes a mixture of Gaussian (infinite mixture of Gaussian).
- Even though $p(x|z)$ and $p(z)$ are normal, $p(x)$ is not normal, because it is a mixture distribution.
- The non-linearity of Gaussian parameters (modeled by a neural network), conjugacy between the prior and the likelihood does not hold anymore.
- Again, μ and σ is non-linear function of z modeled by some non-linear neural network. The neural network works as a powerful non-linear parameter approximator (based on universal approximation theorem).
- Simple prior is used. Let's consider the data x is an image of 100×100 pixels. Then the covariance matrix has to be 10000×10000 . Thus, it is common to set a simple prior such as the standard Gaussian (covariance matrix is diagonal matrix). However, even if we set a simple distribution, with the infinite mixture of Gaussian, we can model any distribution.

- VAE uses a global parametric model to predict the local variational parameters for each data point (**amortized inference**).
- It allows to convert complicated large-dimensional data distributions into simple lower-dimensional latent variable representations.

24.1.1 VAE Optimization

We can train VAE using variational inference with the following objective function, ELBO:

$$\mathcal{L}(\phi, \theta) = \mathbb{E}_{q_\phi(z|x)}[\log p_\theta(x|z)] - D_{\text{KL}}(q_\phi(z|x)||p_\theta(z))$$

Let's closely look at this objective function:

- In $q_\phi(z|x)$, x is a given data, so it is not stochastic. How to sample z ?
- q has to be deterministic and differentiable.

→ **Reparameterization trick!**

$$\tilde{z} \sim q_\phi(z|x) \rightarrow \tilde{z} \sim g_\phi(\epsilon, x)$$

, where $\epsilon \sim p(\epsilon)$.

- Estimated by using Monte-Carlo estimation

$$\mathbb{E}_{q_\phi(z|x)}[\log p_\theta(x|z)] \approx \frac{1}{N} \sum_j \log p_\theta(x_i|z_j).$$

24.1.2 Conditional VAE

If we have label information about data, then it would provide a better optimization of VAE model. Recall that the following objective function is the objective of the original VAE:

$$\mathcal{L}(\phi, \theta) = \mathbb{E}_{q_\phi(z|x)}[\log p_\theta(x|z)] - D_{\text{KL}}(q_\phi(z|x)||p_\theta(z))$$

In conditional VAE,

$$\mathcal{L}(\phi, \theta) = \mathbb{E}_{q_\phi(z|x,y)}[\log p_\theta(x|y,z)] - D_{\text{KL}}(q_\phi(z|x,y) || p_\theta(z|y))$$

$$\log p(X|Y) = \ln \int q(Z|X,Y) \frac{p(X,Z|Y)}{q(Z|X,Y)} dZ \quad (24.1)$$

$$\geq \underbrace{\int q(Z|X,Y) \ln \frac{p(X,Z|Y)}{q(Z|X,Y)} dZ}_{\text{ELBO, } \mathcal{L}(q,\theta)} \quad \text{by Jensen's Inequality.} \quad (24.2)$$

$$\dots \quad (24.3)$$

$$\dots \quad (24.4)$$

$$= \mathbb{E}_{q(Z|X,Y)}[\ln p(X|Z,Y)] - KL(q(Z|X,Y)||p(Z|Y)) \quad (24.5)$$

Note that not we have a prior $p_\theta(z|y)$. However, we have no idea about latent variable z , so we simply assume that we cannot impact the z by y . Thus, we typically set it as a standard normal distribution. Also, we can simply concatenate the input X with Y .

24.1.3 Variational Deep Embedding (VaDE)

The generative process of VADE $p(x, z, c) = p(x|z)p(z|c)p(c)$:

- Choose a cluster $c \sim Cat(\pi)$
- Choose a latent vector $z \sim \mathcal{N}(\mu_c, \sigma_c^2 I)$
- Choose a sample x :

$$x \sim \begin{cases} Ber(\mu_x) & \text{If } x \text{ is binary} \\ \mathcal{N}(\mu_x, \sigma_x^2 I) & \text{else} \end{cases}$$

ELBO of VaDE:

$$\log p(X) = \ln \int \sum_c p(X, Z, C) dz \quad (24.6)$$

$$\geq \underbrace{\int q(Z, C|X) \ln \frac{p(X, Z, C)}{q(Z, C|X)} dZ}_{\text{ELBO}} \quad (24.7)$$

The ELBO can be decomposed as follows:

$$\begin{aligned} \mathcal{L}_{ELBO} &= \mathbb{E}_q(z, c|x) \left[\ln \frac{p(x, z, c)}{q(z, c|x)} \right] \\ &= \mathbb{E}_q(z, c|x) [\ln p(x, z, c) - \ln q(z, c|x)] \\ &= \mathbb{E}_q(z, c|x) [\ln p(x|z) + \ln p(z|c) + \ln p(c) - \ln q(z|x) - \ln q(c|x)] \end{aligned}$$

By using two factorizations:

- $p(x, z, c) = p(x|z)p(z|c)p(c)$
- $q(z, c|x) \approx q(z|x)q(c|x)$ (Mean-field assumption)
 - $q(z|x) \sim \mathcal{N}$: encoder, estimate mean and variance.
 - $q(c|x)$: assignment probability of Gaussian mixture model

24.1.4 Importance Weighted VAE

Chapter 25

Implicit Generative Models

25.1 Generative Adversarial Networks

- Generator's distribution: p_g
- Prior on input noise: $p_z(z)$
- Mapping to data space: $z \rightarrow x$ through $G(z; \theta_g)$
a differentiable multilayer perceptron with parameter θ_g
- $D(x; \theta_d)$: a differentiable multilayer perceptron with parameter θ_d . It outputs a single scalar
- $D(x)$: probability that x (real) came from the data rather than p_g (fake)

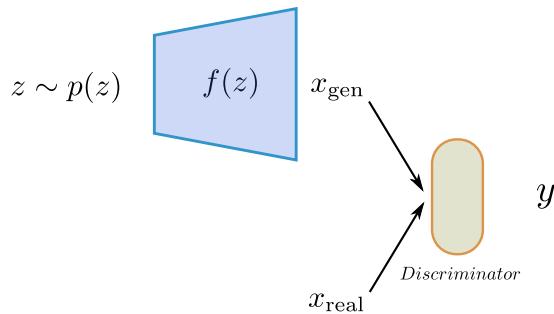


Figure 25.1: GAN structure

25.1.1 Discriminator

The discriminator's goal is to maximize the following equation given G

$$\mathbb{E}_{x \sim p_{\text{data}}(x)} \log(D(x)) + \mathbb{E}_{z \sim p_z(z)} \log(1 - D(G(z)))$$

The optimal discriminator given G can be denoted as D_G^* . To get the optimal discriminator, define a value function

$$V(G, D) := \mathbb{E}_{x \sim p_{\text{data}}(x)} \log(D(x)) + \mathbb{E}_{z \sim p_z(z)} \log(1 - D(G(z))).$$

Then, $D_G^* = \operatorname{argmax}_D V(G, D)$

However, the generator G wants to minimize the value function given $D = D_G^*$.

$$G^* = \operatorname{argmin}_G V(G, D_G^*).$$

$$\min_G \max_D V(D, G) = \mathbb{E}_{x \sim p_{\text{data}}(x)} [\log D(x)] + \mathbb{E}_{z \sim p_z(z)} [\log(1 - D(G(z)))]$$

- $\min_G \rightarrow$ try to generate fake data that is similar to real data
- $\max_D \rightarrow$ try to assign correct label ¹

At this point, we must show that this optimization problem has a unique solution G^* and that this solution satisfies $p_G = p_{\text{data}}$.

One big idea from the GAN paper–, which is different from other approaches is that G **need not be invertible**. Many pieces of notes online miss this fact when they try to replicate the proof and incorrectly use the change of variables formula from calculus (which would depend on G being invertible). Rather, the whole proof relies on this equality:

$$\mathbb{E}_{z \sim p_z(z)} \log(1 - D(G(z))) = \mathbb{E}_{x \sim p_G(x)} \log(1 - D(x)).$$

With the above equality,

$$\begin{aligned} & \mathbb{E}_{x \sim p_{\text{data}}(x)} \log(D(x)) + \mathbb{E}_{z \sim p_z(z)} \log(1 - D(G(z))) \\ &= \int_x p_{\text{data}}(x) \log D(x) dx + \int_z p(z) \log(1 - D(G(z))) dz \\ &= \int_x p_{\text{data}}(x) \log D(x) + p_G(x) \log(1 - D(x)) dx \end{aligned}$$

Additionally, we will use the following property:

$$f(y) = a \log y + b \log(1 - y).$$

To find a critical point,

$$f'(y) = 0 \Rightarrow \frac{a}{y} - \frac{b}{1-y} = 0 \Rightarrow y = \frac{a}{a+b}$$

If $a + b \neq 0$, do the second derivative test:

$$f''\left(\frac{a}{a+b}\right) = -\frac{a}{(\frac{a}{a+b})^2} - \frac{b}{(1-\frac{a}{a+b})^2} < 0$$

If $a, b \in (0, 1)$, $\frac{a}{a+b}$ is a maximum.

By rewriting the equation,

$$\begin{aligned} V(G, D) &= \int_x p_{\text{data}}(x) \log D(x) + p_G(x) \log(1 - D(x)) dx \\ &\leq \int_x \max_y p_{\text{data}}(x) \log y + p_G(x) \log(1 - y) dx \end{aligned}$$

Thus, if $D(x) = \frac{p_{\text{data}}}{p_{\text{data}} + p_G}$, then we can achieve the maximum $V(G, D)$.

¹The above equation is trained separately at the same time, don't get confused

25.1.2 Generator

If we achieve the optimal G (i.e., $p_G = p_{data}$), then D would be completely confused and $D_G^*(x) = \frac{p_{data}}{p_{data} + p_G} = \frac{1}{2}$ (it means that D cannot make a clear decision.).

The global minimum of the virtual training criterion $C(G) = \max_D V(G, D)$ is achieved if and only if $p_G = p_{data}$. Let's plug $D_G^*(x)$ into the criterion then,

$$C(G) = \int_x p_{data}(x) \log \left(\frac{p_{data}(x)}{p_G(x) + p_{data}(x)} \right) + p_G(x) \log \left(\frac{p_G(x)}{p_G(x) + p_{data}(x)} \right) dx.$$

To get the minimum $C(G)$, we can use the Jansen-Shannon divergence:

$$\begin{aligned} D_{JS}(p_{data} || p_G) &= \frac{1}{2} \left[D_{KL}\left(p_{data} \middle\| \frac{p_{data} + p_G}{2}\right) + D_{KL}\left(p_G \middle\| \frac{p_{data} + p_G}{2}\right) \right] \\ &= \frac{1}{2} \left[\left(\int_x p_{data}(x) \log \left(\frac{2p_{data}(x)}{p_{data}(x) + p_G(x)} \right) dx \right) + \left(\int_x p_G(x) \log \left(\frac{2p_G(x)}{p_{data}(x) + p_G(x)} \right) dx \right) \right] \\ &= \frac{1}{2} \left[\left(\int_x p_{data}(x) \log 2 + p_{data}(x) \log \left(\frac{p_{data}(x)}{p_{data}(x) + p_G(x)} \right) dx \right) + \right. \\ &\quad \left. \left(\int_x p_G(x) \log 2 + p_G(x) \log \left(\frac{p_G(x)}{p_{data}(x) + p_G(x)} \right) dx \right) \right] \\ &= \frac{1}{2} \left[\left(\log 2 + \int_x p_{data}(x) \log \left(\frac{2p_{data}(x)}{p_{data}(x) + p_G(x)} \right) dx \right) + \right. \\ &\quad \left. \left(\log 2 + \int_x p_G(x) \log \left(\frac{2p_G(x)}{p_{data}(x) + p_G(x)} \right) dx \right) \right] \\ &= \frac{1}{2} (\log 4 + C(G)) \end{aligned}$$

Thus,

$$C(G) = -\log 4 + 2D_{JS}(p_{data} || p_G)$$

Since the Jensen-Shannon divergence between two distributions is always non-negative and zero only when they are equal, we have shown that $C^* = -\log(4)$ is the global minimum of $C(G)$ and that the only solution is $p_G = p_{data}$, i.e., the generative model perfectly replicating the data generating process.

25.2 Some notes

What would be the optimal discriminator that separates the two different distributions $p(x)$ and $q(x)$? It turns out that it is

$$f(x) = \frac{q(x)}{p(x) + q(x)}$$

Actually, there are many choices for classifiers e.g., KL-divergence

Algorithm 1 Minibatch stochastic gradient descent training of generative adversarial nets. The number of steps to apply to the discriminator, k , is a hyperparameter. We used $k = 1$, the least expensive option, in our experiments.

```

for number of training iterations do
  for  $k$  steps do
    • Sample minibatch of  $m$  noise samples  $\{\mathbf{z}^{(1)}, \dots, \mathbf{z}^{(m)}\}$  from noise prior  $p_g(\mathbf{z})$ .
    • Sample minibatch of  $m$  examples  $\{\mathbf{x}^{(1)}, \dots, \mathbf{x}^{(m)}\}$  from data generating distribution  $p_{\text{data}}(\mathbf{x})$ .
    • Update the discriminator by ascending its stochastic gradient:
      
$$\nabla_{\theta_d} \frac{1}{m} \sum_{i=1}^m \left[ \log D(\mathbf{x}^{(i)}) + \log (1 - D(G(\mathbf{z}^{(i)}))) \right].$$

  end for
  • Sample minibatch of  $m$  noise samples  $\{\mathbf{z}^{(1)}, \dots, \mathbf{z}^{(m)}\}$  from noise prior  $p_g(\mathbf{z})$ .
  • Update the generator by descending its stochastic gradient:
    
$$\nabla_{\theta_g} \frac{1}{m} \sum_{i=1}^m \log (1 - D(G(\mathbf{z}^{(i)}))).$$

end for
The gradient-based updates can use any standard gradient-based learning rule. We used momentum in our experiments.

```

Figure 25.2: Training GAN

1. What do we need to learn a classifier?

- Only samples from $p(x)$ and $q(x)$

2. How do we parameterize $q(x)$?

- Parametric density function (Gaussian)
- Define implicitly (GANs approach): define mapping from one (noise) to another (data or image)

The original GAN does not learn the data distributions.

25.3 Wasserstein Generative Adversarial Networks

25.3.1 KL Divergence

Definition:

$$D_{\text{KL}}(q(x) \parallel p(x)) = \int q(x) \log \frac{q(x)}{p(x)} dx$$

- Forward KL:

- If $q(z) \rightarrow 0$, Forward KL $\rightarrow \infty$
- Zero avoiding for $q(z)$

- Reverse KL:

- If $p(z) \rightarrow 0$, Reverse KL $\rightarrow \infty$
- Zero forcing: $q(z) \rightarrow 0$

Typically, $p(x)$ and $q(x)$ are far apart at the initial state.

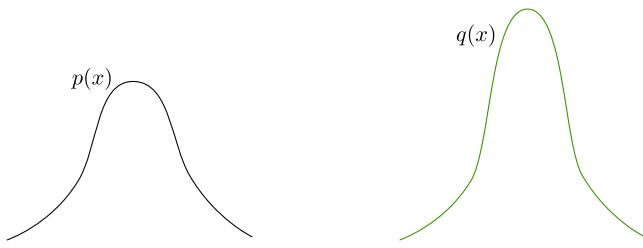


Figure 25.3: Two distributions: $p(x)$ and $q(x)$

Thus, both the forward KL and the reverse KL suffers an instability issue. Specifically, in each case, if the denominator goes to zero, then the divergence goes to infinity.

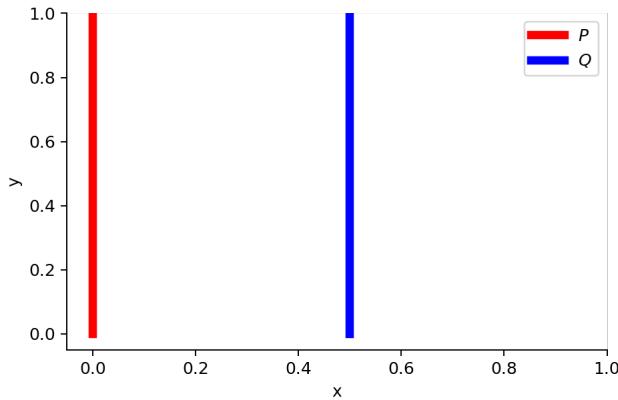
25.3.2 Jensen-Shannon Divergence

Definition:

$$D_{JS}(p_{\text{data}} \parallel p_G) = \frac{1}{2} \left[D_{\text{KL}}\left(p_{\text{data}} \parallel \frac{p_{\text{data}} + p_G}{2}\right) + D_{\text{KL}}\left(p_G \parallel \frac{p_{\text{data}} + p_G}{2}\right) \right]$$

The KL divergence's issue can be alleviated by JS-divergence. Consider a simple example in Fig. 25.46

$$\begin{aligned} & \forall (x, y) \in P, x = 0 \text{ and } y \sim U(0, 1) \\ & \forall (x, y) \in Q, x = \theta, 0 \leq \theta \leq 1 \text{ and } y \sim U(0, 1) \end{aligned}$$

Figure 25.4: Two distributions: $p(x)$ and $q(x)$

$$D_{\text{KL}}(q(x)||p(x)) = \infty$$

$$D_{\text{KL}}(p(x)||q(x)) = \infty$$

$$\begin{aligned} D_{JS}(p_{\text{data}}||p_G) &= \frac{1}{2} \left[D_{KL}\left(p_{\text{data}} \middle\| \frac{p_{\text{data}} + p_G}{2}\right) + D_{KL}\left(p_G \middle\| \frac{p_{\text{data}} + p_G}{2}\right) \right] \\ &= \frac{1}{2} \left[D_{KL}\left(p_{\text{data}} \middle\| \frac{p_{\text{data}}}{2}\right) + D_{KL}\left(p_G \middle\| \frac{p_G}{2}\right) \right] \\ &= \frac{1}{2} [\log 2 + \log 2] = \log 2 \\ W(p, q) &= |\theta| \end{aligned}$$

Therefore, Jensen-Shannon divergence is more stabler than KL divergence. This is one of the reasons why GAN, which uses JS divergence works better than VAE, which uses KL divergence.

However, JS divergence also has some problem. If the value is close to $\frac{1}{2} \log 2$, then the gradient will be very small or close to zero, because the divergence is close to constant. It means that a training speed is very slow. Thus, we need a better metric.

25.3.3 Wasserstein Distance

Wasserstein Distance is a measure of the distance between two probability distributions. It is also called Earth Mover's distance, short for EM distance, because informally it can be interpreted as the minimum energy cost of moving and transforming a pile of dirt in the shape of one probability distribution to the shape of the other distribution.

$$W(p_r, p_g) = \inf_{\gamma \sim \Pi(p_r, p_g)} \mathbb{E}_{(x,y) \sim \gamma} [\|x - y\|]$$

- Π : is the transportation plan and the set of all possible joint probability distributions between p_r and p_g . One joint distribution $\gamma \sim \Pi(p_r, p_g)$ describes one transport plan.
- $\mathbb{E}_{x,y \sim \gamma} \|x - y\| = \sum_{x,y} \gamma(x,y) \|x - y\|$

- Finally, we take the minimum one among the costs of all dirt moving solutions as the EM distance (by infimum).

25.4 WGAN

However, consider all possible joint distribution is intractable, so dual solution can be used.

$$W(p_r, p_g) = \frac{1}{K} \sup_{\|f\|_L \leq K} \mathbb{E}_{x \sim p_r}[f(x)] - \mathbb{E}_{x \sim p_g}[f(x)]$$

So to calculate the Wasserstein distance, we just need to find a 1-Lipschitz function. To enforce the constraint, WGAN applies a very simple clipping to restrict the maximum weight value in f , i.e. the weights of the discriminator

Suppose this function f comes from a family of K -Lipschitz continuous functions, $\{f_w\}_{w \in W}$, parameterized by w . In the modified Wasserstein-GAN, the “discriminator” model is used to learn w to find a good f_w and the loss function is configured as measuring the Wasserstein distance between p_r and p_g .

$$L(p_r, p_g) = W(p_r, p_g) = \max_{w \in W} \mathbb{E}_{x \sim p_r}[f_w(x)] - \mathbb{E}_{z \sim p_r(z)}[f_w(g_\theta(z))]$$

There are two ways to satisfy the Lipschitz continuity:

- Weight clipping
- Gradient Penalty

25.4.1 Lipschitz continuity

The function f in the new form of Wasserstein metric is demanded to satisfy $\|f\|_L \leq K$, meaning it should be K -Lipschitz continuous.

A real-valued function $f : \mathbb{R} \rightarrow \mathbb{R}$ is called K -Lipschitz continuous if there exists a real constant $K \geq 0$ such that, for all $x_1, x_2 \in \mathbb{R}$

$$|f(x_1) - f(x_2)| \leq K|x_1 - x_2|$$

Here K is known as a Lipschitz constant for function $f(\cdot)$. Functions that are everywhere continuously differentiable is Lipschitz continuous, because the derivative, estimated as $\frac{|f(x_1) - f(x_2)|}{|x_1 - x_2|}$, has bounds. However, a Lipschitz continuous function may not be everywhere differentiable, such as $f(x) = |x|$

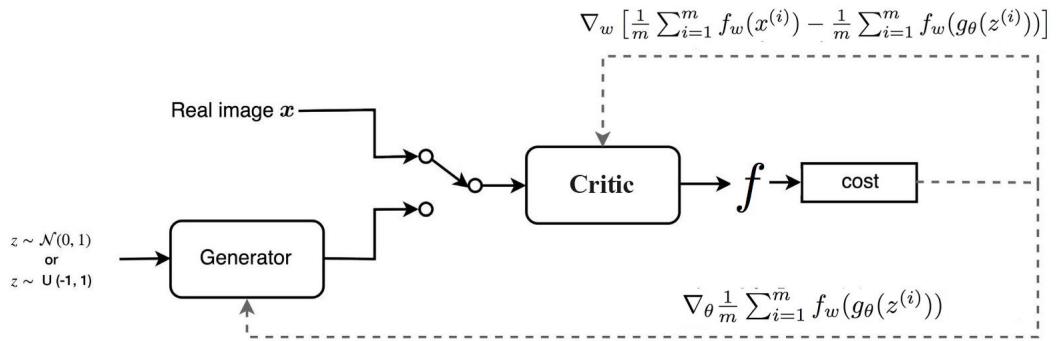


Figure 25.5: WGAN

	Discriminator/Critic	Generator
GAN	$\nabla_{\theta_d} \frac{1}{m} \sum_{i=1}^m [\log D(G(z^{(i)})) + \log (1 - D(x^{(i)}))]$	$\nabla_{\theta_g} \frac{1}{m} \sum_{i=1}^m \log (D(G(z^{(i)})))$
WGAN	$\nabla_w \frac{1}{m} \sum_{i=1}^m [f(x^{(i)}) - f(G(z^{(i)}))]$	$\nabla_\theta \frac{1}{m} \sum_{i=1}^m f(G(z^{(i)}))$

Figure 25.6: WGAN

25.5 InfoGAN: Interpretable Representation Learning by Information Maximizing Generative Adversarial Nets

25.5.1 Joint Entropy

$$H(X, Y) = \mathbb{E}_{X,Y}[-\log p(x, y)] = - \sum_{x,y} p(x, y) \log p(x, y)$$

25.5.2 Conditional Entropy

$$\begin{aligned} H(X|Y) &= \mathbb{E}_Y[H(X, Y)] \\ &= - \sum_{y \sim p_Y(y)} p(y) \sum_{x \sim p_X(x)} p(x|y) \log p(x|y) \\ &= - \sum_{y \sim p_Y(y)} \sum_{x \sim p_X(x)} p(y)p(x|y) \log p(x|y) \\ &= - \sum_{y \sim p_Y(y)} \sum_{x \sim p_X(x)} p(x, y) \log p(x|y) = -\mathbb{E}_{x,y}[\log p(x|y)] \\ &= - \sum_{y \sim p_Y(y)} \sum_{x \sim p_X(x)} p(x, y) \log \frac{p(x, y)}{p(y)} \\ &= - \sum_{y \sim p_Y(y)} \sum_{x \sim p_X(x)} p(x, y) \log p(x, y) + \sum_{y \sim p_Y(y)} \sum_{x \sim p_X(x)} p(x, y) \log p(y) \\ &= H(X, Y) - H(Y) \end{aligned}$$

25.5.3 Variational Mutual Information Maximization

$$\begin{aligned} I(c; G(z, c)) &= H(c) - H(c|G(z, c)) \\ &= H(c) + \int \int p(c = c', x = G(z, c)) \log p(c = c'|x = G(z, c)) dc' dz \\ &= H(c) + \mathbb{E}_{x \sim G(z, c), c' \sim p(c|x)} [\log p(c'|x)] \\ &= H(c) + \mathbb{E}_{x \sim G(z, c)} \mathbb{E}_{c' \sim p(c|x)} [\log p(c'|x)] \\ &= H(c) + \mathbb{E}_{x \sim G(z, c)} \mathbb{E}_{c' \sim p(c|x)} \left[\log \frac{p(c'|x)Q(c'|x)}{Q(c'|x)} \right] \\ &= H(c) + \mathbb{E}_{x \sim G(z, c)} \mathbb{E}_{c' \sim p(c|x)} \left[\log \frac{p(c'|x)}{Q(c'|x)} \right] + \mathbb{E}_{x \sim G(z, c)} \mathbb{E}_{c' \sim p(c|x)} \left[\log Q(c'|x) \right] \\ &= H(c) + \mathbb{E}_{x \sim G(z, c)} \left[D_{KL}(p(c'|x)||Q(c'|x)) \right] + \mathbb{E}_{x \sim G(z, c)} \mathbb{E}_{c' \sim p(c|x)} \left[\log Q(c'|x) \right] \\ &\geq H(c) + \mathbb{E}_{x \sim G(z, c)} \mathbb{E}_{c' \sim p(c|x)} \left[\log Q(c'|x) \right]^2 \end{aligned}$$

Thus we get a lower bound for the mutual information as follows:

$$I(c; G(z, c)) \geq H(c) + \mathbb{E}_{x \sim G(z, c)} \mathbb{E}_{c' \sim p(c|x)} [\log Q(c'|x)]$$

However, we still have a problem. We need to sample c from $p(c|x)$. Thus, we need to replace it with a known distribution. Firstly, with the reasoning that $x \sim G(z, c)$ means sample c from $p(c)$ then sample x from $G(z, c)$. So we can express $\mathbb{E}_{x \sim G(z, c)}$ with $\mathbb{E}_{c \sim p(c)} \mathbb{E}_{x \sim G(z, c)}$. and by the Lemma 1,

$$\begin{aligned} I(c; G(z, c)) &\geq H(c) + \mathbb{E}_{x \sim G(z, c)} \mathbb{E}_{c' \sim p(c|x)} [\log Q(c'|x)] \\ &= H(c) + \mathbb{E}_{c \sim p(c)} \mathbb{E}_{x \sim G(z, c)} \mathbb{E}_{c' \sim p(c|x)} [\log Q(c'|x)] \\ &= H(c) + \mathbb{E}_{c \sim p(c)} \mathbb{E}_{x \sim G(z, c)} [\log Q(c'|x)]^3 \end{aligned}$$

Thus, we can directly sample c from the known distribution instead of $p(c|x)$.

lemma 1 *For random variables X, Y and function $f(x, y)$ under suitable regularity conditions:*

$$\mathbb{E}_{x \sim X, y \sim Y|x} [f(x, y)] = \mathbb{E}_{x \sim X, y \sim Y|x, x' \sim X|y} [f(x', y)]$$

proof 1

$$\begin{aligned} \mathbb{E}_{x \sim X, y \sim Y|x} [f(x, y)] &= \int_x P(x) \int_y P(y|x) f(x, y) dy dx \\ &= \int_x \int_y P(x, y) f(x, y) dy dx \\ &= \int_{x'} \int_y P(x', y) f(x', y) dy dx' \\ &= \int_{x'} \int_y P(y) P(x'|y) f(x', y) dy dx' \\ &= \int_{x'} \int_y \int_x P(x, y) P(x'|y) f(x', y) dx dy dx' \\ &= \int_x P(x) \int_y P(x|y) \int_{x'} P(x'|y) f(x', y) dx dy dx' \\ &= \mathbb{E}_{x \sim X, y \sim Y|x, x' \sim X|y} [f(x', y)] \end{aligned}$$

Chapter 26

Diffusion Model

26.1 Introduction

(Denoising) Diffusion models have emerged as the new SOTA family of deep generative models.

- First proposed in 2015.
- Outperform GANs on image synthesis (DDPM) \sim 2020.
- Stable training dynamics.
- Image synthesis, super resolution, text-to-image, and so on.

The overall idea is to construct a Markov chain of progressively less noisy samples. Each transition denoises a noisy sample. Diffusion models consist of two Markov chains:

1. Forward: A Markov chain of diffusion steps to **slowly add random noise** to data.

$$\underbrace{\mathbf{x}_0}_{\text{data}} \rightarrow \mathbf{x}_1 \cdots \rightarrow \underbrace{\mathbf{x}_T}_{\text{noise}}$$

- $\mathcal{D}_{data} \rightarrow \mathcal{N}$.

2. Backward (Reverse): Learn to **reverse the diffusion process** to construct desired data samples from the noise.

$$\underbrace{\mathbf{x}_T}_{\text{noise}} \rightarrow \mathbf{x}_{T-1} \cdots \rightarrow \underbrace{\mathbf{x}_0}_{\text{data}}$$

- $\mathcal{N} \rightarrow \mathcal{D}_{data}$.

Breaking the overall process into smaller steps allows us to use simple distributions at each step. As will be discussed in the following subsections, we can use Gaussian distributions for the transitions. Thanks to the properties of a Gaussian, the posterior will remain a Gaussian if the likelihood and the prior are both Gaussians. Therefore, if each transitional distribution above is a Gaussian, the joint distribution is also a Gaussian. Since a Gaussian is fully characterized by the first two moments (mean and variance), the computation is highly tractable.

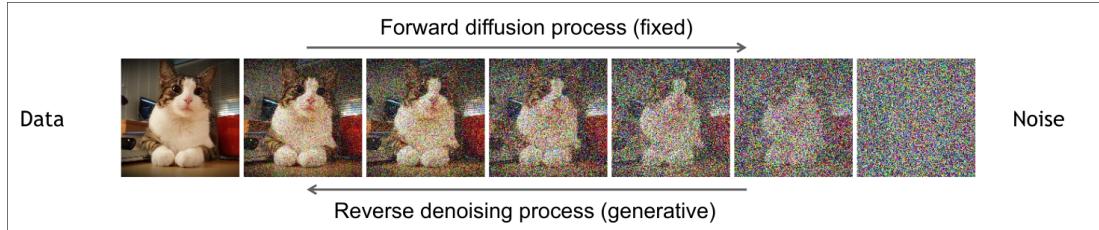


Figure 26.1: [CVRP 2022 Tutorial: Denoising Diffusion-based Generative Modeling: Foundations and Applications](#)

This particular structure is called the *variational diffusion model*.

Some properties of diffusion models:

1. Diffusion model has a pre-defined sampling equation.
 - The equation relies on a random noise.
 - Noise is all we need → Predict noise at a time step t .
2. Fit a model via forward and backward processes.
3. **Iterative transform** of one distribution into another via **Makov Chain**.
 - Diffusion model≈Generative Markov Chain.
4. We want to learn a transition model:

$$p_{\theta}(\mathbf{x}_{t-1}|\mathbf{x}_t) = \mathcal{N}(\mathbf{x}_{t-1}|\boldsymbol{\mu}_{\theta}(\mathbf{x}_t, t), \boldsymbol{\Sigma}_{\theta}(\mathbf{x}_t, t)).$$

5. Base case: $p(\mathbf{x}_T) = \mathcal{N}(0, I)$
6. Marginal distribution over \mathbf{x}_0 can be expressed as follows:

$$p_{\theta}(\mathbf{x}_0) = \int p_{\theta}(\mathbf{x}_0, \dots, \mathbf{x}_T) d\mathbf{x}_1, \dots, \mathbf{x}_T$$

- Similar to the VAE, the $\mathbf{x}_1, \dots, \mathbf{x}_T$ can be considered as latent variables, since we don't know.
- 7. The goal is to learn the parameters so that

$$p(\mathbf{x}_0) \approx p_{\theta}(\mathbf{x}_0)$$

26.2 Forward Diffusion

Let's first model a forward trajectory by using the *Markov property*:

$$q(\mathbf{x}_{0:T}) = q(\mathbf{x}_0) \prod_{t=1}^T \underbrace{q(\mathbf{x}_t | \mathbf{x}_{t-1})}_{\text{Transition kernel}}$$

- Slow transform with a large T : $\mathbf{x}_0 \rightarrow \mathbf{x}_1 \cdots \rightarrow \mathbf{x}_T$
 - To track the transitions of data, we will try to model as many steps as possible.
- The question is how to model $q(\mathbf{x}_t | \mathbf{x}_{t-1})$?

In a continuous case (*e.g.*, image), forward diffusion $q(\mathbf{x}_t | \mathbf{x}_{t-1})$ can be parameterized as follows:

$$q(\mathbf{x}_t | \mathbf{x}_{t-1}) = \mathcal{N}(\mathbf{x}_t; \sqrt{1 - \beta_t} \mathbf{x}_{t-1}, \beta_t \mathbf{I}) \quad (26.1)$$

- $\beta_t \in (0, 1)$ is a variance at time t .
- $\sqrt{1 - \beta_t}$ downscals \mathbf{x}_{t-1} to be 0, $\beta_1 < \cdots < \beta_t$. Thus, \mathbf{x}_t will become more noisier than \mathbf{x}_{t-1} . Then, \mathbf{x}_t can be sampled as:

$$\mathbf{x}_t = \sqrt{1 - \beta_t} \mathbf{x}_{t-1} + \sqrt{\beta_t} \odot \epsilon$$

1. Sample $\mathbf{x}_t \sim q(\mathbf{x}_t)$ and scale it by $\sqrt{1 - \beta_t}$
2. Adds noise $\epsilon \sim \mathcal{N}(0, I)$ with variance β_t .

- The above process is autoregressive (*i.e.*, ancestral sampling), but we can sample \mathbf{x}_t directly from $q(\mathbf{x}_t | \mathbf{x}_0)$ in an analytic form:

$$\mathbf{x}_t = \sqrt{1 - \beta_t} \mathbf{x}_{t-1} + \sqrt{\beta_t} \odot \epsilon_{t-1} \quad (26.2)$$

$$= \sqrt{\alpha_t} \mathbf{x}_{t-1} + \sqrt{1 - \alpha_t} \odot \epsilon_{t-1} \quad (26.3)$$

$$= \sqrt{\alpha_t} \left(\sqrt{\alpha_{t-1}} \mathbf{x}_{t-2} + \sqrt{1 - \alpha_{t-1}} \odot \epsilon_{t-2} \right) + \sqrt{1 - \alpha_t} \odot \epsilon_{t-1} \quad (26.4)$$

$$= \sqrt{\alpha_t \alpha_{t-1}} \mathbf{x}_{t-2} + \sqrt{\alpha_t - \alpha_t \alpha_{t-1}} \odot \epsilon_{t-2} + \sqrt{1 - \alpha_t} \odot \epsilon_{t-1} \quad (26.5)$$

$$= \sqrt{\alpha_t \alpha_{t-1}} \mathbf{x}_{t-2} + \sqrt{\alpha_t - \alpha_t \alpha_{t-1} + 1 - \alpha_t} \odot \epsilon_{t-2} \quad (26.6)$$

$$= \sqrt{\alpha_t \alpha_{t-1}} \mathbf{x}_{t-2} + \sqrt{1 - \alpha_t \alpha_{t-1}} \odot \epsilon_{t-2} \quad (26.7)$$

$$= \dots \quad (26.8)$$

$$= \sqrt{\prod_t \alpha_t} \mathbf{x}_0 + \sqrt{1 - \prod_t \alpha_t} \odot \epsilon_{t_0} \quad (26.9)$$

$$= \sqrt{\bar{\alpha}_t} \mathbf{x}_0 + \sqrt{1 - \bar{\alpha}_t} \odot \epsilon \quad (26.10)$$

$$\sim \mathcal{N}(\mathbf{x}_t; \sqrt{\bar{\alpha}_t} \mathbf{x}_0, (1 - \bar{\alpha}_t) \mathbf{I}), \quad (26.11)$$

where $\alpha_t = 1 - \beta_t$ and $\bar{\alpha}_t = \prod_{s=1}^t \alpha_s$. Thus, $\mathbf{x}_t = \sqrt{\bar{\alpha}_t} \mathbf{x}_0 + \sqrt{1 - \bar{\alpha}_t} \odot \epsilon$. Note that the fifth step is done by using the property of sum of two Gaussian distributions (*e.g.*, $\mathcal{N}(0, \sigma_1^2 I) + \mathcal{N}(0, \sigma_2^2 I) = \mathcal{N}(0, (\sigma_1^2 + \sigma_2^2) I)$).

We can get some intuitions:

- The original input \mathbf{x}_0 **gradually loses all info** during the forward diffusion process.
- This Markov chain has a **stationary distribution** *i.e.*, as $t \rightarrow \infty$, $q(\mathbf{x}_t) \approx \mathcal{N}(0, I)$.
 - In practice, T is set to be a very high number *e.g.*, 1,000.
 - The high T minimizes the information loss at each step, which allows a smooth training.
- The forward diffusion process does not require any training.

26.3 Backward Process

In the forward process, we model the diffusion process, which turns the input data into the noisy data. However, a noisy data is not what we want to obtain. What we want is to generate a new data with high quality.

To generate data, we will reverse the forward process. Then, we can reconstruct the true data. We call this process *Backward process!* The backward process works as follows:

- Start from the noise distribution $q(\mathbf{x}_T) \approx \mathcal{N}(0, I)$.
- Then, sample a noise $\mathbf{x}_T \sim \mathcal{N}(\mathbf{x}_T | 0, I)$
- Iteratively sample $\mathbf{x}_{t-1} \sim q(\mathbf{x}_{t-1} | \mathbf{x}_t)$.
 - Note that $q(\mathbf{x}_{t-1} | \mathbf{x}_t)$ is the **true denoising distribution** that we have no access.
 - In general, $q(\mathbf{x}_{t-1} | \mathbf{x}_t)$ is intractable.
- We can **approximate** $q(\mathbf{x}_{t-1} | \mathbf{x}_t)$ as **Normal distribution** if β_t is small enough in each forward diffusion step.
- We will approximate $q(\mathbf{x}_{t-1} | \mathbf{x}_t)$ using a neural network, $p_\theta(\mathbf{x}_{t-1} | \mathbf{x}_t)$.
- $p_\theta(\mathbf{x}_{0:T}) = p(\mathbf{x}_T) \prod_{t=1}^T p_\theta(\mathbf{x}_{t-1} | \mathbf{x}_t)$.
 - $p(\mathbf{x}_T) = \mathcal{N}(\mathbf{x}_T; 0, I)$.
 - $p_\theta(\mathbf{x}_{t-1} | \mathbf{x}_t) = \mathcal{N}(\mathbf{x}_{t-1}; \boldsymbol{\mu}_\theta(\mathbf{x}_t, t), \boldsymbol{\Sigma}_\theta(\mathbf{x}_t, t))$.
 - We can model the denoising distribution as Normal distribution like above (*c.f.*, [Eq. \(26.35\)](#)).
 - Note that the reverse conditional probability $q(\mathbf{x}_{t-1} | \mathbf{x}_t)$ is tractable when it is conditioned on \mathbf{x}_0 as shown in [Eq. \(26.35\)](#). This allows us to train a neural network to model this denoising distribution.
- Key to the success of this sampling process is training the reverse Markov chain to match the actual time reversal of the forward Markov chain.
- After optimizing the backward process, the sampling procedure is that just sample Gaussian noise from $p(\mathbf{x}_T)$ and then iteratively running the denoising transitions (backward process) for T steps to generate a novel \mathbf{x}_0 .

26.4 Distribution Modeling

What we want to learn (or model) is $p_\theta(\mathbf{x}_0) \approx p(\mathbf{x}_0)$ (approximate data distribution).

- $p_\theta(\mathbf{x}_0) = \int p_\theta(\mathbf{x}_{0:T}) d\mathbf{x}_{1:T}$
- It is intractable to compute all trajectories (*i.e.*, integral over all $\mathbf{x}_{1:T}$).
- Thus, we will deviate the issue by leveraging a variational lower bound with KL-Divergence as follows:

The log-likelihood of p_θ is given by

$$\log p_\theta(\mathbf{x}_0) = \log \int p(\mathbf{x}_{0:T}) d\mathbf{x}_{1:T} \quad (26.12)$$

$$= \log \int p(\mathbf{x}_{0:T}) \frac{q(\mathbf{x}_{1:T}|\mathbf{x}_0)}{q(\mathbf{x}_{1:T}|\mathbf{x}_0)} d\mathbf{x}_{1:T} \quad (26.13)$$

$$= \log \int q(\mathbf{x}_{1:T}|\mathbf{x}_0) \frac{p(\mathbf{x}_{0:T})}{q(\mathbf{x}_{1:T}|\mathbf{x}_0)} d\mathbf{x}_{1:T} \quad (26.14)$$

$$\geq \int q(\mathbf{x}_{1:T}|\mathbf{x}_0) \log \frac{p(\mathbf{x}_{0:T})}{q(\mathbf{x}_{1:T}|\mathbf{x}_0)} d\mathbf{x}_{1:T} \quad \text{by Jensen's Inequality} \quad (26.15)$$

$$= \mathbb{E}_{q(\mathbf{x}_{1:T}|\mathbf{x}_0)} \left[\log \frac{p(\mathbf{x}_{0:T})}{q(\mathbf{x}_{1:T}|\mathbf{x}_0)} \right] \quad (\text{ELBO}) \quad (26.16)$$

$$= \mathbb{E}_{q(\mathbf{x}_{1:T}|\mathbf{x}_0)} \left[\log p(\mathbf{x}_T) \prod_{t=1}^T \frac{p(\mathbf{x}_{t-1}|\mathbf{x}_t)}{q(\mathbf{x}_t|\mathbf{x}_{t-1})} \right] \quad \text{by Markov Property} \quad (26.17)$$

$$= \mathbb{E}_{q(\mathbf{x}_{1:T}|\mathbf{x}_0)} \left[\log \frac{p(\mathbf{x}_T)p(\mathbf{x}_0|\mathbf{x}_1)\prod_{t=2}^T p(\mathbf{x}_{t-1}|\mathbf{x}_t)}{q(\mathbf{x}_T|\mathbf{x}_{T-1})\prod_{t=1}^{T-1} q(\mathbf{x}_t|\mathbf{x}_{t-1})} \right] \quad (26.18)$$

$$= \mathbb{E}_{q(\mathbf{x}_{1:T}|\mathbf{x}_0)} \left[\log \frac{p(\mathbf{x}_T)p(\mathbf{x}_0|\mathbf{x}_1)}{q(\mathbf{x}_T|\mathbf{x}_{T-1})} \right] + \underbrace{\mathbb{E}_{q(\mathbf{x}_{1:T}|\mathbf{x}_0)} \left[\log \prod_{t=1}^{T-1} \frac{p(\mathbf{x}_t|\mathbf{x}_{t+1})}{q(\mathbf{x}_t|\mathbf{x}_{t-1})} \right]}_{\text{Consistency term}} \quad (26.19)$$

The first term of the last line can be further decomposed:

$$\underbrace{\mathbb{E}_{q(\mathbf{x}_{1:T}|\mathbf{x}_0)} [\log p(\mathbf{x}_0|\mathbf{x}_1)]}_{\text{Reconstruction}} + \underbrace{\mathbb{E}_{q(\mathbf{x}_{1:T}|\mathbf{x}_0)} \left[\log \frac{p(\mathbf{x}_T)}{q(\mathbf{x}_T|\mathbf{x}_{T-1})} \right]}_{\text{Prior matching}} + \mathbb{E}_{q(\mathbf{x}_{1:T}|\mathbf{x}_0)} \left[\sum_{t=1}^{T-1} \log \frac{p(\mathbf{x}_t|\mathbf{x}_{t+1})}{q(\mathbf{x}_t|\mathbf{x}_{t-1})} \right]$$

The *reconstruction term* can be simplified as follows:

$$\mathbb{E}_{q(\mathbf{x}_{1:T}|\mathbf{x}_0)} [\log p(\mathbf{x}_0|\mathbf{x}_1)] = \mathbb{E}_{q(\mathbf{x}_1|\mathbf{x}_0)} [\log p(\mathbf{x}_0|\mathbf{x}_1)].$$

This is because we can ignore the steps $[2, \dots, T]$, which are unrelated to the calculation of the expectation.

The *prior matching term* ensures that the latent distribution at the end of the diffusion step is similar to a prior distribution, which is a Gaussian distribution. Note that the *prior matching term* requires no optimization, as it has no trainable parameters. Recall that we have assumed a large enough T such that the final distribution is Gaussian, this term effectively becomes zero.

Now, let's simplify the prior matching term. First, we will only consider the $\mathbf{x}_T, \mathbf{x}_{T-1}$ since the conditional probability depends only on them. Then, the expectation term $q(\mathbf{x}_{1:T}|\mathbf{x}_0)$ becomes:

$$\begin{aligned} q(\mathbf{x}_{1:T}|\mathbf{x}_0) &= q(\mathbf{x}_T, \mathbf{x}_{T-1}|\mathbf{x}_0) \\ &= q(\mathbf{x}_T, \mathbf{x}_{T-1}, \mathbf{x}_0)/q(\mathbf{x}_0) \\ &= q(\mathbf{x}_{T-1}|\mathbf{x}_0)q(\mathbf{x}_T|\mathbf{x}_{T-1}, \mathbf{x}_0) \quad \text{by Chain Rule} \\ &= q(\mathbf{x}_{T-1}|\mathbf{x}_0)q(\mathbf{x}_T|\mathbf{x}_{T-1}) \quad \text{by Markov Property} \end{aligned}$$

By using the above simplification,

$$\begin{aligned} \mathbb{E}_{q(\mathbf{x}_{1:T}|\mathbf{x}_0)} \left[\log \frac{p(\mathbf{x}_T)}{q(\mathbf{x}_T|\mathbf{x}_{T-1})} \right] &= \mathbb{E}_{q(\mathbf{x}_{T-1}|\mathbf{x}_0)} \left[\mathbb{E}_{q(\mathbf{x}_T|\mathbf{x}_{T-1})} \log \frac{p(\mathbf{x}_T)}{q(\mathbf{x}_T|\mathbf{x}_{T-1})} \right] \\ &= -\mathbb{E}_{q(\mathbf{x}_{T-1}|\mathbf{x}_0)} [D_{KL}[q(\mathbf{x}_T|\mathbf{x}_{T-1})||p(\mathbf{x}_T)]] . \end{aligned}$$

Finally, the *consistency term* can be expressed as follows:

$$\begin{aligned} \mathbb{E}_{q(\mathbf{x}_{1:T}|\mathbf{x}_0)} \left[\sum_{t=1}^{T-1} \log \frac{p(\mathbf{x}_t|\mathbf{x}_{t+1})}{q(\mathbf{x}_t|\mathbf{x}_{t-1})} \right] &= \sum_{t=1}^{T-1} \mathbb{E}_{q(\mathbf{x}_{t-1}, \mathbf{x}_t, \mathbf{x}_{t+1}|\mathbf{x}_0)} \left[\log \frac{p(\mathbf{x}_t|\mathbf{x}_{t+1})}{q(\mathbf{x}_t|\mathbf{x}_{t-1})} \right] \\ &= -\sum_{t=1}^{T-1} \mathbb{E}_{q(\mathbf{x}_{t-1}, \mathbf{x}_{t+1}|\mathbf{x}_0)} [D_{KL}[q(\mathbf{x}_t|\mathbf{x}_{t-1})||p(\mathbf{x}_t|\mathbf{x}_{t+1})]] \end{aligned}$$

Note that the last step is done by the chain rule:

$$q(\mathbf{x}_t, \mathbf{x}_{t-1}|\mathbf{x}_0) = \frac{q(\mathbf{x}_t, \mathbf{x}_{t-1}, \mathbf{x}_0)}{q(\mathbf{x}_0)} = \frac{q(\mathbf{x}_t|\mathbf{x}_{t-1}, \mathbf{x}_0)q(\mathbf{x}_{t-1}, \mathbf{x}_0)}{q(\mathbf{x}_0)} = q(\mathbf{x}_t|\mathbf{x}_{t-1})q(\mathbf{x}_{t-1}|\mathbf{x}_0)$$

In sum, we derive the ELBO and its factorized form is given by

$$\begin{aligned} \log p_\theta(\mathbf{x}_0) &\geq \mathbb{E}_{q(\mathbf{x}_{1:T}|\mathbf{x}_0)} \left[\log \frac{p(\mathbf{x}_{0:T})}{q(\mathbf{x}_{1:T}|\mathbf{x}_0)} \right] \\ &= \mathbb{E}_{q(\mathbf{x}_1|\mathbf{x}_0)} [\log p(\mathbf{x}_0|\mathbf{x}_1)] - \mathbb{E}_{q(\mathbf{x}_{T-1}|\mathbf{x}_0)} [D_{KL}[q(\mathbf{x}_T|\mathbf{x}_{T-1})||p(\mathbf{x}_T)]] \\ &\quad - \sum_{t=1}^{T-1} \mathbb{E}_{q(\mathbf{x}_{t-1}, \mathbf{x}_{t+1}|\mathbf{x}_0)} [D_{KL}[q(\mathbf{x}_t|\mathbf{x}_{t-1})||p(\mathbf{x}_t|\mathbf{x}_{t+1})]] . \end{aligned}$$

Let's closely look at the consistency term.

- It endeavors to make the distribution at x_t consistent, from both forward and backward processes. That is, a denoising step from a noisier image should match the corresponding noising step from a cleaner image, for every intermediate timestep. This term is minimized when we train $p_\theta(\mathbf{x}_t|\mathbf{x}_{t+1})$ to match the Gaussian distribution $q(\mathbf{x}_t|\mathbf{x}_{t-1})$, which is defined in Eq. (26.1).
- Under this derivation, all terms of the ELBO are computed as expectations, and can therefore be approximated using Monte Carlo estimates.
- The challenge of the above variational diffusion model is that we need to draw samples (x_{t-1}, x_{t+1}) from a joint distribution $q(x_{t-1}, x_{t+1}|x_0)$ for every time step t . Thus, the variance of its Monte Carlo estimate could potentially be higher than a term that is estimated using only one random variable per time step. As it is computed by summing up $T - 1$ consistency terms, the final estimated value of the ELBO may have high variance for large T values.

Let us instead try to derive a new ELBO, where each term is computed as an expectation over one random variable at a time. The key insight is that we can rewrite encoder transitions as $q(x_t|x_{t-1}) = q(x_t|x_{t-1}, x_0)$, where the extra conditioning term is superfluous due to the Markov property. Then, according to Bayes rule, we can rewrite each transition as:

$$q(\mathbf{x}_t|\mathbf{x}_{t-1}, \mathbf{x}_0) = \frac{q(\mathbf{x}_{t-1}|\mathbf{x}_t, \mathbf{x}_0)q(\mathbf{x}_t|\mathbf{x}_0)}{q(\mathbf{x}_{t-1}|\mathbf{x}_0)}.$$

Armed with this equation, we can factorize the ELBO again as follows:

$$L = \mathbb{E}_{q(\mathbf{x}_{1:T}|\mathbf{x}_0)} \left[\log \frac{p(\mathbf{x}_{0:T})}{q(\mathbf{x}_{1:T}|\mathbf{x}_0)} \right] \rightarrow \text{ELBO} \quad (26.20)$$

$$= \mathbb{E}_{q(\mathbf{x}_{1:T}|\mathbf{x}_0)} \left[\log p(\mathbf{x}_T) \prod_{t=1}^T \frac{p(\mathbf{x}_{t-1}|\mathbf{x}_t)}{q(\mathbf{x}_t|\mathbf{x}_{t-1})} \right] \quad (26.21)$$

$$= \mathbb{E}_{q(\mathbf{x}_{1:T}|\mathbf{x}_0)} \left[\log \frac{p(\mathbf{x}_T)p(\mathbf{x}_0|\mathbf{x}_1)\prod_{t=2}^T p(\mathbf{x}_{t-1}|\mathbf{x}_t)}{q(\mathbf{x}_1|\mathbf{x}_0)\prod_{t=2}^T q(\mathbf{x}_t|\mathbf{x}_{t-1})} \right] \quad (26.22)$$

$$= \mathbb{E}_{q(\mathbf{x}_{1:T}|\mathbf{x}_0)} \left[\log \frac{p(\mathbf{x}_T)p(\mathbf{x}_0|\mathbf{x}_1)\prod_{t=2}^T p(\mathbf{x}_{t-1}|\mathbf{x}_t)}{q(\mathbf{x}_1|\mathbf{x}_0)\prod_{t=2}^T q(\mathbf{x}_t|\mathbf{x}_{t-1}, \mathbf{x}_0)} \right] \quad \text{by Markov Property} \quad (26.23)$$

$$= \mathbb{E}_{q(\mathbf{x}_{1:T}|\mathbf{x}_0)} \left[\log \frac{p(\mathbf{x}_T)p(\mathbf{x}_0|\mathbf{x}_1)}{q(\mathbf{x}_1|\mathbf{x}_0)} + \log \prod_{t=2}^T \frac{p(\mathbf{x}_{t-1}|\mathbf{x}_t)}{q(\mathbf{x}_t|\mathbf{x}_{t-1}, \mathbf{x}_0)} \right] \quad (26.24)$$

$$= \mathbb{E}_{q(\mathbf{x}_{1:T}|\mathbf{x}_0)} \left[\log \frac{p(\mathbf{x}_T)p(\mathbf{x}_0|\mathbf{x}_1)}{q(\mathbf{x}_1|\mathbf{x}_0)} + \log \prod_{t=2}^T \frac{p(\mathbf{x}_{t-1}|\mathbf{x}_t)}{\frac{q(\mathbf{x}_{t-1}|\mathbf{x}_t, \mathbf{x}_0)q(\mathbf{x}_t|\mathbf{x}_0)}{q(\mathbf{x}_{t-1}|\mathbf{x}_0)}} \right] \quad (26.25)$$

$$= \mathbb{E}_{q(\mathbf{x}_{1:T}|\mathbf{x}_0)} \left[\log \frac{p(\mathbf{x}_T)p(\mathbf{x}_0|\mathbf{x}_1)}{q(\mathbf{x}_1|\mathbf{x}_0)} + \log \frac{q(\mathbf{x}_1|\mathbf{x}_0)}{q(\mathbf{x}_T|\mathbf{x}_0)} + \log \prod_{t=2}^T \frac{p(\mathbf{x}_{t-1}|\mathbf{x}_t)}{q(\mathbf{x}_{t-1}|\mathbf{x}_t, \mathbf{x}_0)} \right] \quad (26.26)$$

$$= \mathbb{E}_{q(\mathbf{x}_{1:T}|\mathbf{x}_0)} \left[\log \frac{p(\mathbf{x}_T)p(\mathbf{x}_0|\mathbf{x}_1)}{q(\mathbf{x}_T|\mathbf{x}_0)} + \log \prod_{t=2}^T \frac{p(\mathbf{x}_{t-1}|\mathbf{x}_t)}{q(\mathbf{x}_{t-1}|\mathbf{x}_t, \mathbf{x}_0)} \right] \quad (26.27)$$

$$= \mathbb{E}_{q(\mathbf{x}_1|\mathbf{x}_0)} [\log p(\mathbf{x}_0|\mathbf{x}_1)] + \mathbb{E}_{q(\mathbf{x}_T|\mathbf{x}_0)} \left[\log \frac{p(\mathbf{x}_T)}{q(\mathbf{x}_T|\mathbf{x}_0)} \right] \quad (26.28)$$

$$+ \sum_{t=2}^T \mathbb{E}_{q(\mathbf{x}_t, \mathbf{x}_{t-1}|\mathbf{x}_0)} \left[\log \frac{p(\mathbf{x}_{t-1}|\mathbf{x}_t)}{q(\mathbf{x}_{t-1}|\mathbf{x}_t, \mathbf{x}_0)} \right] \quad (26.29)$$

$$= \mathbb{E}_{q(\mathbf{x}_1|\mathbf{x}_0)} [\log p(\mathbf{x}_0|\mathbf{x}_1)] - D_{KL}(q(\mathbf{x}_T|\mathbf{x}_0)||p(\mathbf{x}_T)) \quad (26.30)$$

$$- \sum_{t=2}^T \mathbb{E}_{q(\mathbf{x}_t|\mathbf{x}_0)} [D_{KL}(q(\mathbf{x}_{t-1}|\mathbf{x}_t, \mathbf{x}_0)||p(\mathbf{x}_{t-1}|\mathbf{x}_t))] \quad (26.31)$$

Let's delve into the last three terms:

- $\mathbb{E}_{q(\mathbf{x}_1|\mathbf{x}_0)} [\log p(\mathbf{x}_0|\mathbf{x}_1)]$: Reconstruction term.
- $D_{KL}[q(\mathbf{x}_T|\mathbf{x}_0)||p(\mathbf{x}_T)]$: Prior matching term.
 - No trainable parameters, since it can be fully characterized by \mathbf{x}_t and \mathbf{x}_0 , so we don't need a neural network to estimate it.
- $\sum_{t=2}^T \mathbb{E}_{q(\mathbf{x}_t|\mathbf{x}_0)} [D_{KL}(q(\mathbf{x}_{t-1}|\mathbf{x}_t, \mathbf{x}_0)||p(\mathbf{x}_{t-1}|\mathbf{x}_t))]$: Consistency term (or denoising matching term).

We can notice that the consistency term requires to estimate three different probabilities:

1. $q(\mathbf{x}_t | \mathbf{x}_0)$: we can sample \mathbf{x}_t by only using \mathbf{x}_0 and ϵ as we derived before:

$$\mathbf{x}_t = \sqrt{\bar{\alpha}_t} \mathbf{x}_0 + \sqrt{1 - \bar{\alpha}_t} \odot \epsilon$$

2. $q(\mathbf{x}_{t-1} | \mathbf{x}_t, \mathbf{x}_0)$: This distribution says we want to sample a less noisy sample \mathbf{x}_{t-1} using the clean sample \mathbf{x}_0 and its noisy version \mathbf{x}_t . In other words, we want to denoise the sample by using the two samples.

3. $p(\mathbf{x}_{t-1} | \mathbf{x}_t)$: We are going to use a distribution parameterized by θ , which is a neural network. We can rewrite this as follows:

$$p_\theta(\mathbf{x}_{t-1} | \mathbf{x}_t)$$

In sum, the consistency term or denoising matching term is the term for training a neural network to make it learn how to denoise samples. Let's compute the KL divergence:

By using the chain rule, $q(\mathbf{x}_{t-1} | \mathbf{x}_t, \mathbf{x}_0)$ can be factorized as follows:

$$q(\mathbf{x}_{t-1} | \mathbf{x}_t, \mathbf{x}_0) = \frac{q(\mathbf{x}_t | \mathbf{x}_{t-1}, \mathbf{x}_0) q(\mathbf{x}_{t-1} | \mathbf{x}_0)}{q(\mathbf{x}_t | \mathbf{x}_0)} = \frac{q(\mathbf{x}_t | \mathbf{x}_{t-1}) q(\mathbf{x}_{t-1} | \mathbf{x}_0)}{q(\mathbf{x}_t | \mathbf{x}_0)}.$$

Then, $q(\mathbf{x}_t | \mathbf{x}_{t-1}, \mathbf{x}_0) = q(\mathbf{x}_t | \mathbf{x}_{t-1})$ by Markov property. Then we get

$$q(\mathbf{x}_{t-1} | \mathbf{x}_t, \mathbf{x}_0) = \frac{q(\mathbf{x}_t | \mathbf{x}_{t-1}) q(\mathbf{x}_{t-1} | \mathbf{x}_0)}{q(\mathbf{x}_t | \mathbf{x}_0)} \quad (26.32)$$

$$= \frac{\mathcal{N}(\mathbf{x}_t; \sqrt{\bar{\alpha}_t} \mathbf{x}_{t-1}, (1 - \bar{\alpha}_t) \mathbf{I}) \mathcal{N}(\mathbf{x}_{t-1}; \sqrt{\bar{\alpha}_{t-1}} \mathbf{x}_0, (1 - \bar{\alpha}_{t-1}) \mathbf{I})}{\mathcal{N}(\mathbf{x}_t; \sqrt{\bar{\alpha}_t} \mathbf{x}_0, (1 - \bar{\alpha}_t) \mathbf{I})} \quad (26.33)$$

$$\vdots \quad (26.34)$$

$$\propto \mathcal{N}\left(\mathbf{x}_{t-1}; \underbrace{\frac{\sqrt{\bar{\alpha}_t}(1 - \bar{\alpha}_{t-1})\mathbf{x}_t + \sqrt{\bar{\alpha}_{t-1}}(1 - \bar{\alpha}_t)\mathbf{x}_0}{1 - \bar{\alpha}_t}}_{\mu_q(\mathbf{x}_t, \mathbf{x}_0)}, \underbrace{\frac{(1 - \bar{\alpha}_t)(1 - \bar{\alpha}_{t-1})}{1 - \bar{\alpha}_t} \mathbf{I}}_{\Sigma_q(t)}\right), \quad (26.35)$$

where $\alpha_t = 1 - \beta_t$ and $\bar{\alpha}_t = \prod_{s=1}^t \alpha_s$. We have shown that at each step, $\mathbf{x}_{t-1} \sim q(\mathbf{x}_{t-1} | \mathbf{x}_t, \mathbf{x}_0)$ is still normally distributed, with mean $\mu_q(\mathbf{x}_t, \mathbf{x}_0)$ that is a function of \mathbf{x}_t and \mathbf{x}_0 , and variance $\Sigma_q(t)$ as a function of α coefficients. We can rewrite the variance as $\Sigma_q(t) = \sigma_q^2(t) \mathbf{I}$.

We can also notice that the mean $\mu_q(\mathbf{x}_t, \mathbf{x}_0)$ is a linear combination of \mathbf{x}_t and \mathbf{x}_0 . Geometrically, it lives on the straight line connecting \mathbf{x}_t and \mathbf{x}_0 . In other words, the denoised sample \mathbf{x}_{t-1} stays somewhere between the \mathbf{x}_t and \mathbf{x}_0 .

Another observation is that the coefficient of \mathbf{x}_t increases and that of \mathbf{x}_0 decreases. Conversely, \mathbf{x}_t decreases and that of \mathbf{x}_0 increases as t goes to 0 from T .

In order to match approximate denoising transition step $p_\theta(\mathbf{x}_{t-1} | \mathbf{x}_t)$ to the denoising transition step $q(\mathbf{x}_{t-1} | \mathbf{x}_t, \mathbf{x}_0)$ as closely as possible, we can also model it as a Gaussian. Let's simplify the KL divergence:

$$\begin{aligned} D_{KL}(q(\mathbf{x}_{t-1} | \mathbf{x}_t, \mathbf{x}_0) \| p_\theta(\mathbf{x}_{t-1} | \mathbf{x}_t)) \\ = D_{KL}(\mathcal{N}(\mathbf{x}_{t-1}; \mu_q, \Sigma_q(t)) \| \mathcal{N}(\mathbf{x}_{t-1}; \mu_\theta, \Sigma_q(t))) \\ \vdots \\ = \frac{1}{2\sigma_q^2(t)} \|\mu_\theta(\mathbf{x}_t) - \mu_q(\mathbf{x}_t, \mathbf{x}_0)\|_2^2. \end{aligned}$$

By using this, we can rewrite the ELBO that we want to maximize:

$$\text{ELBO}_\theta(\mathbf{x}) = \mathbb{E}_{q(\mathbf{x}_1|\mathbf{x}_0)}[\log p(\mathbf{x}_0|\mathbf{x}_1)] - \sum_{t=2}^T \mathbb{E}_{q(\mathbf{x}_t|\mathbf{x}_0)} \left[\frac{1}{2\sigma_q^2(t)} \|\boldsymbol{\mu}_\theta(\mathbf{x}_t) - \boldsymbol{\mu}_q(\mathbf{x}_t, \mathbf{x}_0)\|_2^2 \right]$$

Note that the reconstruction term can be analytically computed, so we can make simplify further.

26.5 Training and Inference

Recall that the ELBO we derived suggests that we need to minimize the following equation:

$$\frac{1}{2\sigma_q^2(t)} \|\boldsymbol{\mu}_\theta(\mathbf{x}_t) - \boldsymbol{\mu}_q(\mathbf{x}_t, \mathbf{x}_0)\|_2^2$$

We already know that $\boldsymbol{\mu}_q(\mathbf{x}_t, \mathbf{x}_0)$ is given by

$$\boldsymbol{\mu}_q(\mathbf{x}_t, \mathbf{x}_0) = \frac{\sqrt{\alpha_t}(1 - \bar{\alpha}_{t-1})\mathbf{x}_t + \sqrt{\bar{\alpha}_{t-1}}(1 - \alpha_t)\mathbf{x}_0}{1 - \bar{\alpha}_t}.$$

Therefore, $\boldsymbol{\mu}_q(\mathbf{x}_t, \mathbf{x}_0)$ can be fully characterized by \mathbf{x}_t and \mathbf{x}_0 .

Next, we need to design $\boldsymbol{\mu}_\theta$ and we can set it to follow a simple form, which is similar to $\boldsymbol{\mu}_q$. Then, we get

$$\boldsymbol{\mu}_\theta(\mathbf{x}_t) = \frac{\sqrt{\alpha_t}(1 - \bar{\alpha}_{t-1})\mathbf{x}_t + \sqrt{\bar{\alpha}_{t-1}}(1 - \alpha_t)\hat{\mathbf{x}}_\theta(\mathbf{x}_t)}{1 - \bar{\alpha}_t},$$

where $\hat{\mathbf{x}}_\theta$ is another network that seeks to predict \mathbf{x}_0 from noisy image \mathbf{x}_t at t -th step. Then, the KL divergence becomes

$$\frac{1}{2\sigma_q^2(t)} \|\boldsymbol{\mu}_\theta(\mathbf{x}_t) - \boldsymbol{\mu}_q(\mathbf{x}_t, \mathbf{x}_0)\|_2^2 = \frac{1}{2\sigma_q^2(t)} \frac{\bar{\alpha}_{t-1}(1 - \alpha_t)^2}{(1 - \bar{\alpha}_t)^2} \|\hat{\mathbf{x}}_\theta(\mathbf{x}_t) - \mathbf{x}_0\|_2^2$$

This indicates that optimizing a VDM boils down to learning a neural network to predict the original ground truth sample from an arbitrarily noisy version of it.

Finally, we get

$$\text{ELBO}_\theta(\mathbf{x}) = \mathbb{E}_{q(\mathbf{x}_1|\mathbf{x}_0)}[\log p(\mathbf{x}_0|\mathbf{x}_1)] - \sum_{t=2}^T \mathbb{E}_{q(\mathbf{x}_t|\mathbf{x}_0)} \left[\frac{1}{2\sigma_q^2(t)} \frac{\bar{\alpha}_{t-1}(1 - \alpha_t)^2}{(1 - \bar{\alpha}_t)^2} \|\hat{\mathbf{x}}_\theta(\mathbf{x}_t) - \mathbf{x}_0\|_2^2 \right].$$

This implies that optimizing VDM is equivalent to a denoising problem as we need to find a network $\hat{\mathbf{x}}_\theta(\mathbf{x}_t)$ such that \mathbf{x}_t will be close to the ground truth \mathbf{x}_0 .

- We are not going to predict arbitrary noisy samples.
- The noisy sample \mathbf{x}_t is carefully crafted by the forward sampling equation.

26.6 Noise Prediction

$$\mathbf{x}_0 = \frac{\mathbf{x}_t - \sqrt{1 - \bar{\alpha}_t} \boldsymbol{\epsilon}_0}{\sqrt{\bar{\alpha}_t}}$$

By plugging the above equation, we get:

$$\boldsymbol{\mu}_{\theta}(\mathbf{x}_t) = \frac{1}{\sqrt{\alpha_t}} \mathbf{x}_t - \frac{1 - \alpha_t}{\sqrt{1 - \bar{\alpha}_t} \sqrt{\alpha_t}} \hat{\boldsymbol{\epsilon}}_{\theta}(\mathbf{x}_t)$$

$$\begin{aligned} D_{KL}(q(\mathbf{x}_{t-1}|\mathbf{x}_t, \mathbf{x}_0) \| p(\mathbf{x}_{t-1}|\mathbf{x}_t)) \\ \vdots \\ = \frac{1}{2\sigma_q^2(t)} \frac{(1 - \alpha_t)^2}{(1 - \bar{\alpha}_t)\alpha_t} [\|\hat{\boldsymbol{\epsilon}}_{\theta}(\mathbf{x}_t) - \boldsymbol{\epsilon}_0\|_2^2]. \end{aligned}$$

The $\hat{\boldsymbol{\epsilon}}_{\theta}(\mathbf{x}_t)$ is a neural network that learns to predict the source noise $\boldsymbol{\epsilon}_0 \sim \mathcal{N}(\boldsymbol{\epsilon}; \mathbf{0}, \mathbf{I})$. This indicates that learning a VDM by predicting the original image \mathbf{x}_0 is equivalent to learning to predict the noise; empirically, however, some works have found that predicting the noise resulted in better performance.

26.7 Summary

- The loss function can be decomposed.

$$L_{\text{VLB}} = L_T + L_{T-1} + \dots + L_0.$$

- $L_T = D_{\text{KL}}(q(\mathbf{x}_T | \mathbf{x}_0) \| p_{\theta}(\mathbf{x}_T))$
 - Constant ≈ 0 since x_T is a Gaussian noise.
- $L_t = D_{\text{KL}}(q(\mathbf{x}_{t-1} | \mathbf{x}_t, \mathbf{x}_0) \| p_{\theta}(\mathbf{x}_{t-1} | \mathbf{x}_t))$ for $t > 1$
 - This is the main part.
- $L_0 = -\log p_{\theta}(\mathbf{x}_0 | \mathbf{x}_1)$
 - Can be modeled by a separate decoder.

- $q(\mathbf{x}_t | \mathbf{x}_0) = \mathcal{N}(\sqrt{\bar{\alpha}_t} \mathbf{x}_0, (1 - \bar{\alpha}_t)I)$
- $q(\mathbf{x}_t | \mathbf{x}_{t-1}) = \mathcal{N}(\mathbf{x}_t; \sqrt{1 - \beta_t} \mathbf{x}_{t-1}, \beta_t \mathbf{I}),$

We can sample by $\mathbf{x}_t = \sqrt{1 - \beta_t} \mathbf{x}_{t-1} + \sqrt{\beta_t} \epsilon$

- $p_{\theta}(\mathbf{x}_{t-1} | \mathbf{x}_t) = \mathcal{N}(\mathbf{x}_{t-1}; \mu_{\theta}(\mathbf{x}_t, t), \Sigma_{\theta}(\mathbf{x}_t, t)).$
 - We need to learn mean and variance.
 - DDPM kept the variance fixed and let the neural network only learn the mean μ_{θ} .
 - $\Sigma_{\theta}(\mathbf{x}_t, t) = \sigma_t^2 \mathbf{I}$ and set $\sigma_t^2 = \beta_t$.
 - Improved DDPM model trains σ also.
- One can reparameterize the mean to make the neural network learn the added noise via a network ϵ_{θ} .
$$\mu_{\theta}(\mathbf{x}_t, t) = \frac{1}{\sqrt{\bar{\alpha}_t}} \left(\mathbf{x}_t - \frac{\beta_t}{\sqrt{1 - \bar{\alpha}_t}} \underbrace{\epsilon_{\theta}(\mathbf{x}_t, t)}_{\text{Network}} \right)$$
- Final objective function L_t is

$$\|\epsilon - \epsilon_{\theta}(\mathbf{x}_t, t)\|^2 = \|\epsilon - \epsilon_{\theta}(\sqrt{\bar{\alpha}_t} \mathbf{x}_0 + \sqrt{(1 - \bar{\alpha}_t)} \epsilon, t)\|^2$$

- $t \sim \text{Unif}[\{1, \dots, T\}]$
- $\mathbf{x}_t = \sqrt{\bar{\alpha}_t} \mathbf{x}_0 + \sqrt{(1 - \bar{\alpha}_t)} \epsilon \sim q(\mathbf{x}_t | \mathbf{x}_0)$
- $\epsilon \sim \mathcal{N}(0, I)$
- \mathbf{x}_t is perturbed by ϵ and the noise prediction network ϵ_{θ} predicts ϵ .

Algorithm 6: Training

```

repeat
   $\mathbf{x}_0 \sim q(\mathbf{x}_0)$ 
   $t \sim \text{Unif}\{1, \dots, T\}$ 
   $\epsilon \sim \mathcal{N}(0, I)$ 
  Take gradient descent step on  $\nabla_{\theta} ||\epsilon - \epsilon_{\theta}(\sqrt{\bar{\alpha}_t} \mathbf{x}_0 + \sqrt{(1 - \bar{\alpha}_t)} \epsilon, t)||^2$ 
until converged;

```

The training process is given by

1. $\mathbf{x}_0 \sim q(\mathbf{x}_0)$
2. Sample a noise level t between 1 and T (*i.e.*, random time step).
3. Sample a noise from a Gaussian distribution and perturb the input by the sampling equation.
4. NN is trained to predict this noise ϵ used for generating \mathbf{x}_t .
5. β is often scheduled linearly.
6. Σ is set equal to β .

The sampling process is given by

Algorithm 7: Sampling

```

 $\mathbf{x}_T \sim \mathcal{N}(0, I)$ 
for  $t = T, \dots, 1$  do
   $\mathbf{z} \sim \mathcal{N}(0, I)$ 
   $\mathbf{x}_{t-1} = \frac{1}{\sqrt{\bar{\alpha}_t}} \left( \mathbf{x}_t - \frac{1 - \bar{\alpha}_t}{\sqrt{1 - \bar{\alpha}_t}} \epsilon_{\theta}(\mathbf{x}_t, t) \right) + \Sigma_t \mathbf{z}$ 
return  $\mathbf{x}_0$ 

```

- Ancestral sampling.
- T is typically around 1,000

26.8 Score Matching

- Suppose $\{\mathbf{x}_0, \dots, \mathbf{x}_N\}$, where each data point (*e.g.*, image, video, or text)) is sampled independently from a data distribution $p(\mathbf{x})$.
- Given the dataset, the goal of generative modeling is to fit a model to the data distribution such that we can synthesize new data points at will by sampling from the model.
- One way is to directly model the distribution function as in likelihood-based models. Let $f_\theta(\mathbf{x}) \in \mathbb{R}^d$, then we can define a density function:

$$p_\theta(\mathbf{x}) = \frac{\exp^{-f_\theta(\mathbf{x})}}{Z_\theta}$$

- $f_\theta(\mathbf{x}) \in \mathbb{R}^d$ is often called unnormalized probabilistic model or energy-based model.
- Energy-based model originates from the Gibbs distribution in statistical physics.
- $p_\theta(\mathbf{x})$ can be trained by maximizing the log-likelihood of the data.

$$\max_{\theta} \sum_i^N \log p_\theta(\mathbf{x}_i).$$

- The gradient of the loglikelihood is given by

$$\begin{aligned} \nabla_{\theta} \log p_{\theta}(\mathbf{x}) &= \nabla_{\theta} f_{\theta}(\mathbf{x}) - \nabla_{\theta} Z_{\theta} \\ \nabla_{\theta} Z_{\theta} &= \frac{\nabla_{\theta} Z_{\theta}}{Z_{\theta}} \\ &= \frac{1}{Z_{\theta}} \nabla_{\theta} \int \exp(f_{\theta}(\mathbf{x})) d\mathbf{x} \\ &= \frac{1}{Z_{\theta}} \int \exp(f_{\theta}(\mathbf{x})) \nabla_{\theta} f_{\theta}(\mathbf{x}) d\mathbf{x} \\ &= \int \frac{1}{Z_{\theta}} \exp(f_{\theta}(\mathbf{x})) \nabla_{\theta} f_{\theta}(\mathbf{x}) d\mathbf{x} \\ &= \int p_{\theta}(\mathbf{x}) \nabla_{\theta} f_{\theta}(\mathbf{x}) d\mathbf{x} \\ &= \mathbb{E}_{p_{\theta}(\mathbf{x})} [\nabla_{\theta} f_{\theta}(\mathbf{x})] \\ \nabla_{\theta} \log p_{\theta}(\mathbf{x}) &= \nabla_{\theta} f_{\theta}(\mathbf{x}) - \mathbb{E}_{p_{\theta}(\mathbf{x})} [\nabla_{\theta} f_{\theta}(\mathbf{x})] \end{aligned}$$

- However, it is undesirable, since Z_θ is intractable.
 - For instance, a gray scale image of 100×100 has $256^{10,000}$ space.
- Thus, we have to sidestep the issue by using some solutions, for instance:
 - Approximate by using VAE or MCMC

Instead, we can leverage ***Stein Score***:

- By modeling a score function, instead of the density function, we can sidestep the difficulty of computing the intractable normalizing constants.

- *Stein Score* function: $\nabla_{\mathbf{x}} \log p(\mathbf{x})$.
 - Not a gradient w.r.t. model parameters.
 - Gradient of the log probability density function.
 - Not same as the score in stat.
- It is a direction that maximizes a log data density.
- A model for approximating the score function is called a *score-based model* $s_{\theta}(\mathbf{a})$.
- Score-based models does not have to compute the intractable normalizing constant, Z_{θ} .

$$s_{\theta}(\mathbf{x}) = \nabla_{\mathbf{x}} \log p_{\theta}(\mathbf{x}) = -\nabla_{\mathbf{x}} f_{\theta}(\mathbf{x}) - \underbrace{\nabla_{\mathbf{x}} \log Z_{\theta}}_{\text{Constant}}.$$

26.8.1 Fisher Divergence

We need to know about *Fisher Divergence*:

- Given *i.i.d.* samples $\{\mathbf{x}_1, \dots, \mathbf{x}_N\} \sim p_{\text{data}}(\mathbf{x}) = p(\mathbf{x})$.
- Estimating the score function $\nabla_{\mathbf{x}} \log p(\mathbf{x})$.
- Score model $s_{\theta}(\mathbf{x}) : \mathbb{R}^D \rightarrow \mathbb{R}^D$.
- Use score estimator $s_{\theta}(x)$:

$$\mathcal{L}_{\theta} = \frac{1}{2} \mathbb{E}_{p(\mathbf{x})} [||\nabla_{\mathbf{x}} \log p(\mathbf{x}) - s_{\theta}(\mathbf{x})||_2^2].$$

- It is called *Fisher divergence*.
- Intuitively, the Fisher divergence compares the squared distance between the ground-truth data score and the score-based model.
 - It changes the problem into a *regression* problem.
- Direct computation of the divergence is **infeasible** due to the unknown data score $\nabla_{\mathbf{x}} \log p(\mathbf{x})$.
 - Since we have no access to the true data distribution $p(\mathbf{x})$.

Fortunately, there exists a family of methods called **score matching** that minimize the Fisher divergence without knowledge of the ground-truth data score.

- Score matching objectives can directly be estimated on a dataset and optimized with stochastic gradient descent, analogous to the log-likelihood objective for training likelihood-based models (with known normalizing constants).
- We can train the score-based model by minimizing a score matching objective, without requiring adversarial optimization.

$$\begin{aligned}\mathcal{L}_\theta &= \mathbb{E}_{p(\mathbf{x})} \left[\frac{1}{2} \|s_\theta(x)\|_2^2 + \text{tr}(\nabla_{\mathbf{x}} s_\theta(x)) \right] \\ &\approx \frac{1}{N} \sum_{i=1}^N \left[\frac{1}{2} \|s_\theta(x)\|_2^2 + \text{tr}(\nabla_{\mathbf{x}} s_\theta(x)) \right]\end{aligned}$$

- $\{\mathbf{x}_1, \dots, \mathbf{x}_N\} \sim p(\mathbf{x})$
- $\nabla_{\mathbf{x}} s_\theta(x)$: Jacobian
- Remove the dependency of $p(\mathbf{x})$

$$\mathcal{L}_\theta = \frac{1}{2} \mathbb{E}_{p(x)} [\| \nabla_x \log p(x) - s_\theta(x) \|_2^2] \quad (26.36)$$

$$= \frac{1}{2} \mathbb{E}_{p(x)} [(\nabla_x \log p(x) - s_\theta(x))^2] \quad (26.37)$$

$$= \frac{1}{2} \int p(x) (\nabla_x \log p(x) - s_\theta(x))^2 dx \quad (26.38)$$

$$= \underbrace{\frac{1}{2} \int p(x) (\nabla_x \log p(x))^2 dx}_{\text{independent from theta}} + \frac{1}{2} \int p(x) s_\theta(x)^2 dx - \int p(x) s_\theta(x) \nabla_x \log p(x) dx \quad (26.39)$$

$$= \dots - \int p(x) s_\theta(x) \nabla_x \log p(x) dx \quad (26.40)$$

$$= \dots - \int p(x) s_\theta(x) \frac{\nabla_x p(\mathbf{x})}{p(x)} dx \quad (26.41)$$

$$= \dots - \int \nabla_{\mathbf{x}} p(x) s_\theta(x) dx \quad (26.42)$$

$$= \dots - \left[p(x) s_\theta(x) \right]_{x=-\infty}^{\infty} + \int p(x) \nabla_x s_\theta(x) dx \quad (26.43)$$

$$= \frac{1}{2} \int p(x) s_\theta(x)^2 dx + \int p(x) \nabla_x s_\theta(x) dx + \text{const} \quad (26.44)$$

$$= \frac{1}{2} \mathbb{E}_{p(x)} [s_\theta(x)^2] + \mathbb{E}_{p(x)} [\nabla_x s_\theta(x)] + \text{const}. \quad (26.45)$$

- The second last term used the integration by parts.
- The last step is done by a boundary condition assumption which makes score function to be zero (*c.f.*, Sliced score matching paper).
 - $p_{\text{data}}(x) \rightarrow 0$ as $|x| \rightarrow \infty$.
 - In other words, gradient vanishes on the boundary.

26.8.2 Langevin Dynamics

- Once we have trained a score-based model $s_\theta(\mathbf{x}) \approx \nabla_{\mathbf{x}} \log p(\mathbf{x})$, we can use an iterative procedure called *Langevin Dynamics* (LD) [?] to draw samples from it.
- LD provides an MCMC procedure to sample from a distribution $p(\mathbf{x})$ using only its score function.

$$\mathbf{x}_t \leftarrow \mathbf{x}_{t-1} + \frac{\epsilon}{2} \nabla_{\mathbf{x}} \log p(\mathbf{x}_{t-1}) + \sqrt{\epsilon} \mathbf{z}_t$$

- Specifically, it initializes the chain from an arbitrary prior distribution $\mathbf{x}_0 \sim \pi(\mathbf{x})$, and then iterates the following
 - Sample from $p(x)$ using only the score $\nabla_x \log p(x)$.
 - $\mathbf{z}_t \sim \mathcal{N}(\mathbf{0}, \mathbf{I})$.
 - ϵ is the step size.
- As $T \rightarrow \infty$ and $\epsilon \rightarrow 0$, \mathbf{x}_T will become the true probability density $p(\mathbf{x})$.

Chapter 27

Flow Models

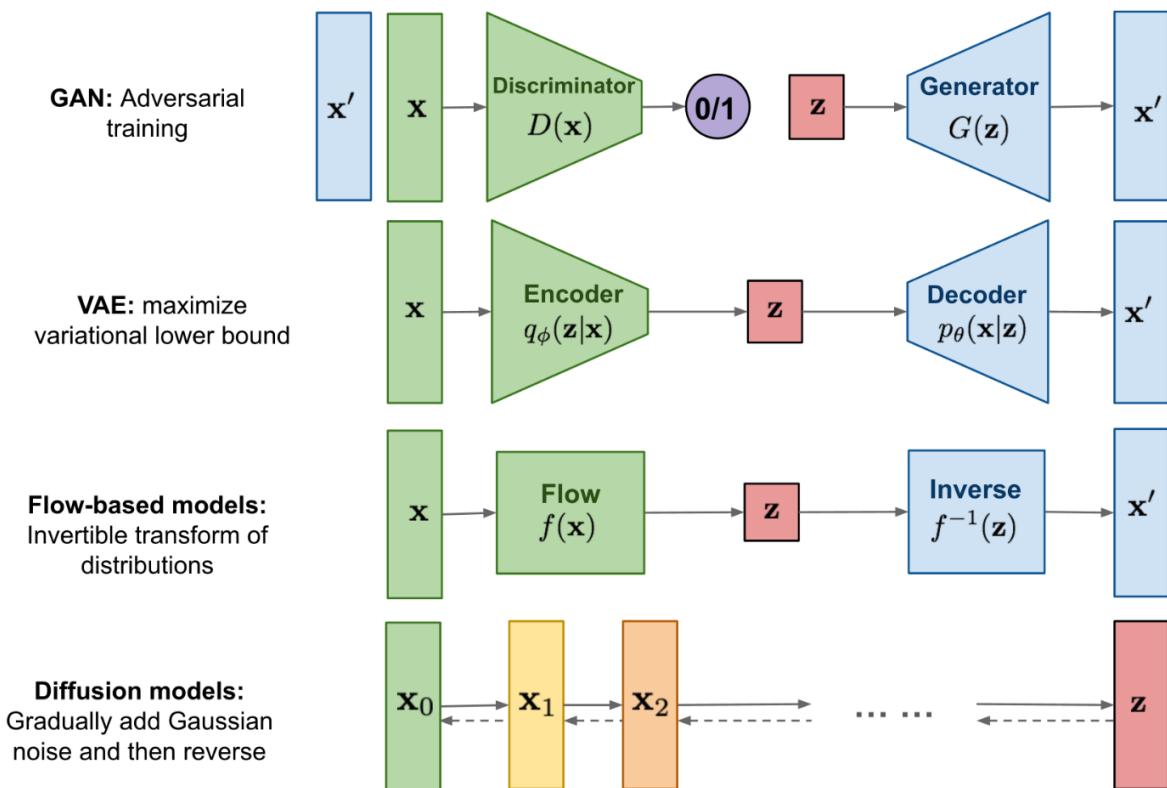
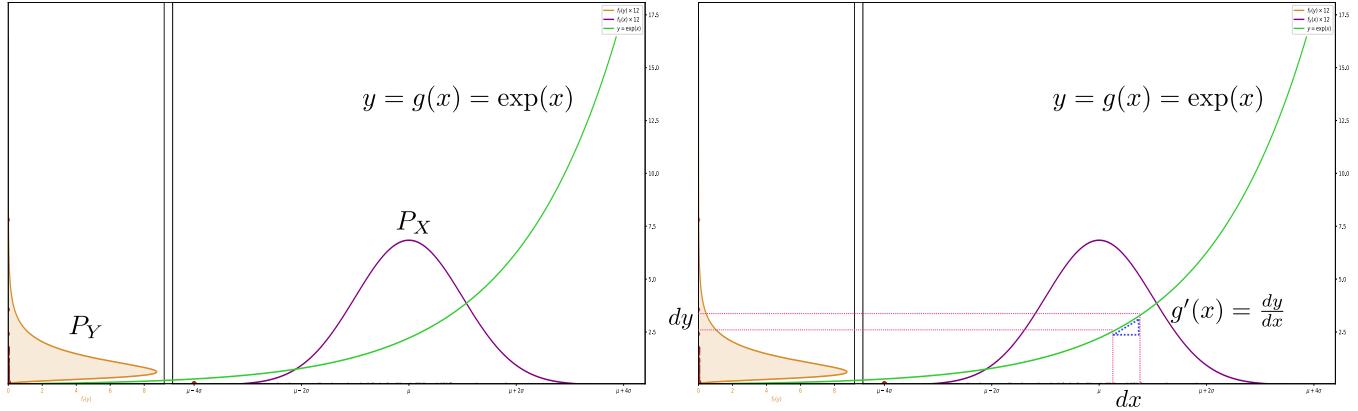


Figure 27.1: An overview of generative models.

- Flow-models get an exact estimate of the likelihood of your sample, as well as in the reverse direction.
- VAEs optimize a lower bound on the (log) likelihood
- GANs minimize a discrepancy between your input and transformed noise distributions.



27.1 The Method of Transformations of Random Variables

If we are interested in finding the PDF of $Y = g(X)$, where $g(\cdot)$ is some deterministic transformation of X , and the function g satisfies following properties, we can utilize a method called the method of transformations.

- $g(x)$ is differentiable;
- $g(x)$ is a strictly (or monotonically) increasing function, that is, if $x_1 < x_2$, then $g(x_1) < g(x_2)$.

How to derive the PDF of the random variable $Y = g(X)$ when one knows the PDF of the random variable X ? If X is discrete, we can derive the pmf for Y by simply summing up the probability mass for all the x 's such that $f(x) = y$. For a general function g , there is no direct formula to get the PDF of the random variable $Y = g(X)$ knowing $p(X)$. There is a formula in case when h is a differentiable one-to-one mapping from the range (*i.e.*, the support) of X to the range of Y .

Take for example a random variable $X \sim \mathcal{N}(\mu, \sigma)$ and set $Y = \exp(X)$. The figure below shows some simulations of X and the corresponding values of Y . The density of X is shown in blue and the one of Y is shown in orange in the vertical direction.

Now the question is: knowing the density of X , what is the density of Y ? Taking a point y in the range of Y , the PDF f_Y provides the probability of Y , belong to a small area dy around y by the formula below

$$P(Y \in dy) \approx f_Y(y)|dy|,$$

where $P(Y \in dy)$ is the area below the curve. Similarly, we can define

$$P(X \in dx) \approx f_X(x)|dx|$$

The above two areas are approximately the same in case of very small region. Note that if dy and dx are very small, we can approximate the derivative of $g'(x) = \frac{|dy|}{|dx|}$. Compactly, this can be expressed as follows:

$$P(X \in dx) = f_X(x) \frac{|dy|}{g'(x)}$$

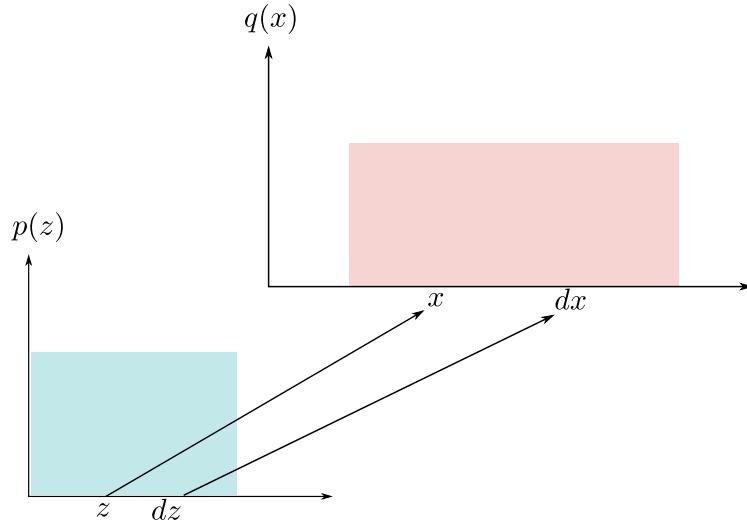


Figure 27.2: Area would be approximately $p(z)dz = q(z)dx$. Thus, $q(x) = p(z)\left|\frac{dz}{dx}\right|$

With $y = g(x)$ we can get

$$\begin{aligned} P(X \in dx) &\approx P(Y \in dy) = f_X(x) \frac{|dy|}{g'(x)} \\ &= f_X(g^{-1}(y)) \frac{|dy|}{g'(g^{-1}(y))} \\ &= f_X(g^{-1}(y)) |dy| (g^{-1})'(y) \end{aligned}$$

The last line is by the derivative of inverse function which is

$$\frac{d}{dx}f^{-1}(x) = \frac{1}{f'(f^{-1}(x))}$$

Then,

$$P(Y \in dy) \approx f_Y(y) |dy| = f_X(g^{-1}(y)) |(g^{-1})'(y)| \cdot |dy|$$

Finally, we can get

$$f_Y(y) = f_X(g^{-1}(y)) |(g^{-1})'(y)|$$

Note that the absolute value is determined by the function h . It is a scaling factor tells us how a tiny volume element (like a little square or cube) is stretched or compressed under the transformation.

27.1.1 Vector to Vector

Z and X be random variables which are related by a mapping $f : \mathbb{R}^n \rightarrow \mathbb{R}^n$ such that $X = f(Z)$ and $Z = f^{-1}(X)$. Then

$$p_X(\mathbf{x}) = p_Z(f^{-1}(\mathbf{x})) \left| \det \left(\frac{\partial f^{-1}(\mathbf{x})}{\partial \mathbf{x}} \right) \right|$$

Note that for any invertible square matrix A over a field (*e.g.*, real or complex numbers),

$$\det(A^{-1}) = \frac{1}{\det(A)}.$$

Let

$$A = \begin{bmatrix} 2 & 1 \\ 3 & 4 \end{bmatrix}.$$

Then, determinant of A is given by

$$\det(A) = 2 \cdot 4 - 3 \cdot 1 = 8 - 3 = 5.$$

The inverse of A is

$$A^{-1} = \frac{1}{5} \begin{bmatrix} 4 & -1 \\ -3 & 2 \end{bmatrix}.$$

Then, the determinant of A^{-1} is

$$\det(A^{-1}) = \frac{1}{5^2} (4 \cdot 2 - (-3)(-1)) = \frac{1}{25} (8 - 3) = \frac{5}{25} = \frac{1}{5}.$$

You can confirm the property:

$$\det(A^{-1}) = \frac{1}{5} = \frac{1}{\det(A)}.$$

For example, we can transform (x_1, x_2) to (r, θ) via $x_1 = r \cos \theta$ and $x_2 = r \sin \theta$.

Then

$$J_{y \rightarrow x} = \begin{pmatrix} \frac{\partial x_1}{\partial r} & \frac{\partial x_1}{\partial \theta} \\ \frac{\partial x_2}{\partial r} & \frac{\partial x_2}{\partial \theta} \end{pmatrix} = \begin{pmatrix} \cos \theta & -r \sin \theta \\ \sin \theta & r \cos \theta \end{pmatrix}, \quad (3)$$

so

$$|\det J_{y \rightarrow x}| = |r \cos^2 \theta + r \sin^2 \theta| = |r|. \quad (4)$$

Hence

$$p_{\mathbf{y}}(\mathbf{y}) = p_{\mathbf{x}}(\mathbf{x}) |\det J_{y \rightarrow x}| \implies p_{R,\Theta}(r, \theta) = p_{X_1, X_2}(x_1, x_2) r. \quad (5-6)$$

For a two dimensional random vector (X, Y) with density $p_{X,Y}$,

$$Pr((X, Y) \in A) = \int \int_A p_{X,Y}(x, y) dx dy.$$

For a infinitesimally small region, we can approximate as follows:

$$\int_x^{x+dx} p_X(x) dx \approx p_X(x) dx.$$

Similarly, we can get the probability as follows:

$$P(r \leq R \leq r + dr, \theta \leq \Theta \leq \theta + d\theta) = p_{R,\Theta}(r, \theta) dr d\theta. \quad (7)$$

Note that the length of arc is $r \times d\theta$. Thus, the area $r dr d\theta$ or probability is given by

$$P(r \leq R \leq r + dr, \theta \leq \Theta \leq \theta + d\theta) = p_{X,Y}(r \cos \theta, r \sin \theta) r dr d\theta, \quad (8-9)$$

so that finally

$$p_{R,\Theta}(r, \theta) = p_{X,Y}(r \cos \theta, r \sin \theta) r. \quad (10)$$

27.2 Normalizing Flows

For a **univariate** case,

- Z : source random variable and density $p(z)$
- X : target random variable and density $q(x)$

$$T : Z \rightarrow X$$

•

$$q(x) = p(z) \left| \frac{\partial T(z)}{\partial z} \right|^{-1}$$

For a **multivariate** case

$$\mathbf{T} : \mathbb{R}^d \rightarrow \mathbb{R}^d$$

So,

$$q(\mathbf{x}) = p(\mathbf{z}) |\det \nabla_{\mathbf{z}} \mathbf{T}(\mathbf{z})|^{-1}$$

For example,

- $\mathbf{X} \in \mathbb{R}^d$ and $\mathbf{Z} \in \mathbb{R}^d$
- $\mathbf{X} = (\mathbf{x}_1, \dots, \mathbf{x}_d)$ and $\mathbf{Z} = (\mathbf{z}_1, \dots, \mathbf{z}_d)$
- $\mathbf{T} = (T_1, T_2, \dots, T_d)$
- $\mathbf{x}_1 = T_1(\mathbf{z}), \mathbf{x}_2 = T_2(\mathbf{z}), \mathbf{x}_d = T_d(\mathbf{z})$

Jacobian matrix of \mathbf{T} is given by

$$\nabla_{\mathbf{z}} \mathbf{T}(\mathbf{z}) = \begin{bmatrix} \frac{\partial T_1}{\partial z_1} & \frac{\partial T_1}{\partial z_2} & \dots & \frac{\partial T_1}{\partial z_d} \\ \frac{\partial T_2}{\partial z_1} & \frac{\partial T_2}{\partial z_2} & \dots & \frac{\partial T_2}{\partial z_d} \\ \vdots & \vdots & \ddots & \vdots \\ \frac{\partial T_d}{\partial z_1} & \frac{\partial T_d}{\partial z_2} & \dots & \frac{\partial T_d}{\partial z_d} \end{bmatrix}$$

Our goal is to estimate the dataset $\mathcal{D} = \{\mathbf{x}_1, \dots, \mathbf{x}_n\} \sim q(\mathbf{x})$

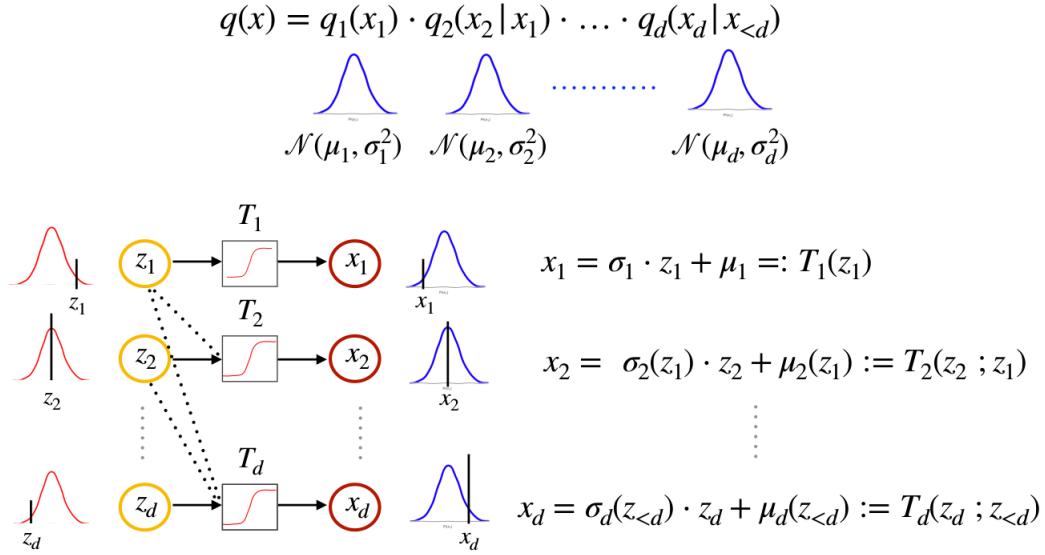
- Choose a simple density $p(z)$
- We want to estimate $q(\mathbf{x})$ by using MLE:

$$\prod_{i=1}^n q(\mathbf{x}_i) = \prod_{i=1}^n p(\mathbf{z}_i) |\det \nabla_{\mathbf{z}} \mathbf{T}(\mathbf{z})|^{-1} \quad \text{by change of variable}$$

$$\hat{\mathbf{T}} := \underset{\mathbf{T}}{\operatorname{argmax}} \prod_{i=1}^n p(\mathbf{z}_i) |\det \nabla_{\mathbf{z}} \mathbf{T}(\mathbf{z})|^{-1}$$

$$\hat{\mathbf{T}} := \underset{\mathbf{T}}{\operatorname{argmax}} \sum_{i=1}^n \log p(\mathbf{z}_i) - \log |\det \nabla_{\mathbf{z}} \mathbf{T}(\mathbf{z})|$$

- $\mathbf{z}_i = \mathbf{T}^{-1}(\mathbf{x}_i)$
- We need to compute the inverse \mathbf{T}^{-1} , which is computationally expensive
- We want to avoid the issue by using Jacobian trainagular matrix



27.2.1 Triangular Maps

$$\mathbf{T} : \mathbb{R}^d \rightarrow \mathbb{R}^d$$

$$\nabla_{\mathbf{z}} \mathbf{T}(\mathbf{z}) = \begin{bmatrix} \frac{\partial T_1}{\partial z_1} & 0 & \cdots & 0 \\ \frac{\partial T_2}{\partial z_1} & \frac{\partial T_2}{\partial z_2} & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ \frac{\partial T_d}{\partial z_1} & \frac{\partial T_d}{\partial z_2} & \cdots & \frac{\partial T_d}{\partial z_d} \end{bmatrix}$$

Increasing: T_j is w.r.t z_j . Determinant is the product of diagonal elements.

- $x_1 = T_1(z_1)$
- $x_2 = T_2(z_1, z_2)$
- $x_3 = T_2(z_1, z_2, z_3)$
- $x_d = T_d(z_1, z_2, \dots, z_d)$

$$\min_{\mathbf{T}} \sum_{i=1}^n \left[-\log p(\mathbf{T}^{-1}(\mathbf{x}_i)) + \sum_j \log \frac{\partial T_j}{\partial z_j} \right]$$

Part V

Causal Inference

Chapter 28

Causal Inference

In this chapter, we move from *correlation-based* learning to genuinely *causal* reasoning. We will:

- Distinguish correlation from causation and relate this to standard machine learning,
- Discuss the difference between observing the world and intervening in it,
- Use the Berkeley admissions case to illustrate the limits of purely observational analysis,
- Introduce causal models as a bridge between data and assumptions.

28.1 Correlation vs Causation

Before we talk about observations and actions, it is useful to clarify a basic but crucial distinction: *correlation* versus *causation*.

28.1.1 Correlation: how things move together

Two variables are said to be *correlated* if they tend to move together in the data:

- When one variable is high, the other also tends to be high, or
- When one is high, the other tends to be low.

Formally, we can measure this with quantities such as *covariance* or *correlation coefficients*. Intuitively, correlation captures how much two variables *co-vary* in the data we observe.

A key point is that correlation is purely about **patterns in the data**, not about mechanisms. If ice-cream sales and drowning accidents go up together in summer, they are correlated; but it would be absurd to say that buying ice cream *causes* drowning. Both are driven by a third factor: the weather.

28.1.2 Causation: how one variable affects another

By contrast, *causation* is about what happens to one variable if we *change* another variable:

If we force X to change, how does Y respond (on average)?

Here we are interested in the **exact impact** of one variable on another under *interventions*, not just how they happened to move together in the past. Causal questions are of the form:

- What is the effect of raising the legal driving age on traffic accidents?
- What is the effect of a new drug on blood pressure?
- What is the effect of a new recommendation algorithm on user engagement?

These questions all ask about the consequences of a hypothetical action, not just about existing correlations.

28.1.3 What most machine learning actually learns

Most standard machine learning algorithms—linear regression, logistic regression, random forests, neural networks, transformers, and so on—are trained to learn patterns like

$$P(Y | X) \quad \text{or} \quad f(x) \approx \mathbb{E}[Y | X = x],$$

that is, they learn to predict Y from X based on *observed co-variation* in the data. In other words:

- They learn **correlations** (or more generally, statistical dependencies) between inputs and outputs.
- They do not, by default, distinguish whether those dependencies are *causal* or simply due to confounding factors.

We can think of standard ML as **correlation-based learning**:

- It focuses on learning *how variables change together* in the data,
- not on what would happen if we actively intervened and changed one variable while holding everything else fixed.

This distinction is crucial:

- A model that is excellent at exploiting correlations may be very good at prediction in environments similar to its training data,
- but it can fail badly when we change the environment or policy, because it does not know the *underlying causal structure*.

Causal inference, by contrast, aims precisely at learning (or using assumptions about) the *causal impact* of one variable on another.

In short, correlation-based learning is about capturing *how variables change together* in the observed world, while causal inference aims to characterize the *direct impact* of changing one variable on another under a specified intervention.

28.1.4 From correlation to intervention

In the rest of this note, we will move from correlation-based questions to truly causal ones. To do this, we will introduce the distinction between:

- **Observations:** passively watching the world as it is, and
- **Actions (interventions):** asking what would happen if we changed something in the world.

This will lead naturally to the language of *causal reasoning*, and to examples—such as driving-age policies and the Berkeley admissions case—that illustrate why correlation alone is not enough.

28.2 Observations vs Actions

Let's start with the difference between just watching the world and actually changing something in it.

- When we passively observe, we simply watch how people normally behave, following their habits, routines, and preferences.
- The data we collect in this way are like a snapshot of the world as it currently is, summarized in whatever features we decided to record (age, income, accident history, etc.).

This kind of data is called *observational data*. So what we can answer with observational data? Some questions fit perfectly into this observational world. For example:

Example

Do 16-year-old drivers have more traffic accidents than 18-year-old drivers?

This is a question about how often something happens in the world as it is. Mathematically, we can:

- Compute the probability of an accident given that the driver is 16.
- Compute the probability of an accident given that the driver is 18.
- Then compare those two numbers (*e.g.*, subtract them).

If we have a large enough sample with both 16- and 18-year-old drivers, we can estimate these probabilities directly from the data using standard observational statistics.

Now consider a slightly different question:

What would happen to traffic fatalities if we raised the legal driving age by two years?

This sounds similar, but it's actually a different type of question.

- The previous question was about **how often something happens in the current world**.
- This new question is about **what would happen if we changed the rules of the world**.

That is, it's asking about the effect of a *hypothetical action (an intervention)*: "Raise the driving age → what changes?"

You can't answer that reliably just by looking at current accident rates by age, because:

- Maybe older drivers crash less simply because they have more experience, not because they're older.
- An 18-year-old with only 2 months of driving experience might be no safer than a 16-year-old with 2 months of experience.

We could try to control for experience, for example, we can compare accident rates for people with the same number of months of driving, but different ages.

But even then, we hit complications:

- Maybe 18-year-olds who only just started driving are unusually cautious people, who drive less and more carefully.
- Maybe they tend to live in areas with different traffic conditions than places where people start driving younger.

So even when you try to adjust for obvious factors like months of experience, other hidden differences can still mess up your conclusions.

We might try another strategy: Compare countries with different legal driving ages, like the US and Germany. These countries differ in many ways besides driving age:

- Road design, public transport, culture, enforcement,
- and even other laws, like the legal drinking age.

So differences in accident rates could be caused by any of those factors, not just the driving age.

28.3 What causal reasoning is about

This is where causal reasoning comes in. Causal reasoning is a framework—both conceptual and technical—for answering questions like:

- What is the effect of doing X? (e.g., raising the driving age)
- What action caused Y? (e.g., what policy likely reduced accidents?)

It focuses on *interventions*:

- Not just what is the world like?
- But what would the world look like if we changed something?

Once we understand how to define and estimate *the effect of an action*, we can turn questions around and ask:

Given that we observed this outcome, what actions or causes are likely responsible?

To do this systematically, we use *causal models*, which:

- Help design better studies (which variables to measure, control, or ignore),
- And give us a formal language to connect assumptions (how we think the world works) to conclusions (what we think would happen under different actions).

28.4 The limitations of observation

In 1973, researchers looked at graduate school admissions at the University of California, Berkeley. They had data on 12,763 applicants across 101 departments and programs.

- About 4,321 of these applicants were women, and roughly 35% of them were admitted.
- About 8,442 were men, and around 44% of them were admitted.

Just looking at these totals, it seems like **men were more likely to be admitted than women** (44% vs 35%). Standard statistical tests say this difference is too big to be explained by random chance alone, so it looks like a real gap.

The same pattern appeared when the researchers focused on the six largest departments:

- Across these six, men again had an overall acceptance rate of about 44%,
- while women had an overall rate of about 30%.

Again, this suggests that men are doing better than women when you look at the data in aggregate (all combined).

However, each department decides who to admit by itself, and departments can be very different — different fields, different standards, different competitiveness. So the researchers drilled down and looked at the acceptance rates within each of those six big departments.

What they found was surprising:

- In four of the six departments, women actually had a higher acceptance rate than men.
- In the other two departments, men had the higher acceptance rate.

But those two departments weren't big enough or different enough to explain the large overall gap in the combined data.

We can find a *reversal*:

- Overall across departments: men seem to be favored.
- Inside most departments: women do as well or better than men.

This is an example of what's often called *Simpson's paradox*: a situation where a pattern that appears in overall (aggregate) data reverses when you break the data into subgroups.

In this case:

- Event (Y): the applicant is accepted.
- Event (A): the applicant is female (gender treated as a binary variable).
- Variable (Z): which department the applicant applied to.

Simpson's paradox means it can happen that:

- For each department (Z), women might do as well as or better than men:

$$P(\text{accepted} \mid \text{female}, Z) \geq P(\text{accepted} \mid \text{male}, Z), \text{ but}$$

- Overall, women still have a lower acceptance rate than men:

$$P(\text{accepted} \mid \text{female}) < P(\text{accepted} \mid \text{male}).$$

This happens because men and women apply to different departments in different proportions, and those departments have different levels of competitiveness.

From the data, one thing is very clear: **Gender affects which departments people apply to**. Men and women have **different patterns of department choice**.

We also know that **departments differ in how hard it is to get in**. Some have low acceptance rates (very competitive), others have higher acceptance rates.

So one plausible explanation is:

- Women tended to apply more to highly competitive departments, while men applied more to less competitive ones.
- As a result, women were rejected more often overall, even though departments themselves may have treated individual male and female applicants fairly.

This was essentially the conclusion of the original study. They argued:

- The bias seen in the combined statistics does not come from admissions committees systematically discriminating against women.
- Instead, it comes from earlier stages in the pipeline: the way society and the education system have steered women toward certain fields.

They suggested that: Women were **shunted** by their upbringing and education into fields that:

- are more crowded,
- have fewer resources and funding,
- have lower completion rates,
- and often lead to poorer job prospects.

In other words, they said the gender bias was mainly a *pipeline problem*: by the time women reached graduate applications, they were already concentrated in less favorable, more competitive fields, through no fault of the departments themselves. However, it's hard to fully defend or criticize that conclusion using this data alone, because key information is missing.

For example, we don't know, why women chose those more competitive departments:

- Maybe some less competitive departments (like certain engineering programs) were unwelcoming to women.
- Maybe some departments had a bad reputation for how they treated women, so women avoided them.
- Maybe the way departments advertised or described themselves discouraged women from applying.

We also don't know anything about the qualifications of applicants:

- It could be that, because of social barriers, women who applied to engineering in 1973 were on average more qualified than the men who applied.
- In that case, if the acceptance rate for men and women is equal, it might actually mean women are being held to a higher bar, which would be discrimination.

So, the observed acceptance rates alone cannot tell us whether there was discrimination or not. They leave us with many plausible stories, and we can't distinguish between them without more information.

Given this uncertainty, the author says there are two main options:

1. Design a new study and collect better data.

- Measure more variables (like applicant qualifications, department culture, prior experiences, etc.).
 - This might allow a more conclusive answer about discrimination.
2. Stay with the current data and argue using assumptions and background knowledge.
- Use what we know about the social context of the 1970s, academic culture, and gender norms.
 - Then argue about which explanation is more likely:
 - Is it mostly a neutral pipeline effect?
 - Or a mix of pipeline factors and discrimination at various stages?

28.5 Causal models as a bridge between data and assumptions

Up to this point, we have seen two recurring themes:

- Observational data alone can be ambiguous: the same pattern can be explained by multiple causal stories.
- Policy questions (“what if we changed X ?”) are fundamentally about interventions, not just about correlations.

Causal models provide a way to organize both of these issues. They give us:

- A way to distinguish between *seeing* and *doing*,
- A structured way to encode assumptions about how variables influence one another,
- A toolkit to derive the effect of hypothetical actions from observational data when possible.

28.5.1 Observation vs intervention in notation

In ordinary statistics, we are used to working with conditional probabilities of the form

$$P(Y | X = x),$$

which answer questions of the form:

Among the individuals for whom we see $X = x$, how often do we see Y ?

For example, $P(\text{accident} | \text{age} = 16)$ is the accident rate among all drivers who happen to be 16 in our data.

Causal questions, by contrast, refer to the effect of an *action*. For this we introduce the notation

$$P(Y | \text{do}(X = x)),$$

which answers:

If we were to force X to take the value x (by intervention), how often would we see Y ?

In the driving-age example, $P(\text{accident} \mid \text{do}(\text{min. age} = 18))$ describes the accident rate we would expect under a policy that sets the legal driving age to 18, taking into account all downstream changes (who drives, when they start, how much experience they accumulate, etc.).

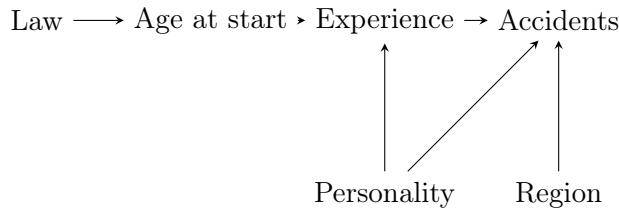
In general,

$$P(Y \mid X = x) \neq P(Y \mid \text{do}(X = x)),$$

because individuals who *happen* to have $X = x$ in the observational world may differ in systematic ways from those who would end up with $X = x$ under an intervention. This discrepancy is precisely what we saw in the Berkeley example and the driving-age discussion: simple conditioning on X mixes together many other influences.

28.5.2 Causal diagrams (informally)

One convenient way to encode assumptions about how variables influence each other is via *causal diagrams*, also called directed acyclic graphs (DAGs). For the driving-age example, a highly simplified diagram might look like:



Read informally, this says:

- The **Law** constrains the age at which people can start driving.
- **Age at start** influences how much **Experience** they can accumulate.
- **Experience** influences the risk of **Accidents**.
- **Personality** (cautious vs reckless) affects both how much **Experience** people seek and their risk of **Accidents**.
- **Region** (urban vs rural, road conditions, etc.) also affects **Accidents**.

This kind of diagram helps us see why simply comparing $P(\text{Accidents} \mid \text{Age})$ can be misleading:

- Age is entangled with Experience, Personality, and Region.
- Changing the Law changes Age, which in turn changes Experience and who drives at all.

Causal calculus (Pearl's do-calculus) tells us when and how we can recover $P(Y \mid \text{do}(X = x))$ from purely observational data given such a diagram, and when we cannot.

28.5.3 Designing better studies

Causal models are not only a tool for interpreting existing data; they also guide the *design* of new studies. Given a causal diagram, we can ask:

- Which variables must we measure in order to block spurious associations (confounding paths)?
- Which variables should we *not* condition on, because they may introduce new biases (colliders)?
- Under what conditions can we estimate causal effects from purely observational data?

In the Berkeley example, a causal model could tell us that we are missing key variables, such as:

- Applicant qualifications (grades, test scores, research experience),
- Department climate or reputation for treating women,
- Broader social forces shaping field choice.

The model would then make explicit that, without additional measurements or assumptions, certain causal questions (e.g., “Was there discrimination at the department level?”) cannot be answered definitively from the existing data.

28.5.4 Reasoning with incomplete data

Even when we cannot collect more data, causal models still have value. They allow us to:

- Enumerate different plausible causal stories consistent with the observed data,
- Explore the implications of each story (“If the world worked this way, what would the effect of changing X be?”),
- Identify which assumptions are doing the real work in our conclusions.

In other words, a causal model acts as a *logic engine* for cause-and-effect reasoning: given explicit assumptions, it yields explicit conclusions, and it clarifies which parts of our reasoning depend on empirical evidence and which parts depend on judgment and prior knowledge.

28.6 Summary

We can now summarize the main points:

- Observational data show us how the world *is*, not automatically how it would be if we *changed* something.

- Questions about policies or interventions (raising the driving age, changing admissions rules, etc.) are inherently *causal* and require reasoning about $\text{do}(\cdot)$, not just conditional probabilities.
- The Berkeley admissions example illustrates how aggregated patterns can reverse when we condition on a relevant variable (department), a phenomenon known as Simpson's paradox.
- Observational patterns alone can be compatible with multiple causal explanations (pipeline effects, discrimination at various stages, or both); data without a causal model cannot fully resolve such ambiguities.
- Causal models provide a structured way to:
 - Distinguish observation from intervention,
 - Design better studies and decide which variables to measure or control,
 - Connect domain knowledge and assumptions to concrete, testable implications.

In this sense, causal inference is not just an add-on to statistics but a complementary framework that lets us talk rigorously about actions, policies, and the mechanisms that generate the data we see.

Part VI

Advanced Topics

Chapter 29

Neural Ordinary Differential Equations

29.1 Preliminary

29.1.1 Euler Method

The Euler method is a simple numerical technique used to solve ordinary differential equations (ODEs) of the form $\frac{dy}{dt} = f(t, y)$. It is an initial value problem where we seek to find the function $y(t)$ given an initial condition $y(t_0) = y_0$. Here's a step-by-step explanation of the Euler method:

Problem Setup: Given,

- A differential equation $\frac{dy}{dt} = f(t, y)$
- An initial condition $y(t_0) = y_0$

Discretization: The idea is to approximate the solution at discrete points. Let's denote:

- t_n as the n -th time step
- y_n as the approximation of $y(t_n)$

We define a step size h such that $t_{n+1} = t_n + h$.

Euler's Approximation: Using the first-order Taylor series expansion, we can approximate $y(t)$ at t_{n+1} as:

$$y_{n+1} \approx y_n + h \cdot f(t_n, y_n)$$

Iterative Process: Starting from the initial condition (t_0, y_0) :

1. Calculate the next value using the formula:

$$y_{n+1} = y_n + h \cdot f(t_n, y_n)$$

2. Repeat the process for $n = 0, 1, 2, \dots$ until the desired value of t is reached.

Example: Let's solve the differential equation $\frac{dy}{dt} = y$ with the initial condition $y(0) = 1$ using the Euler method.

- Set the step size h (e.g., $h = 0.1$).
- Start with $t_0 = 0$ and $y_0 = 1$.

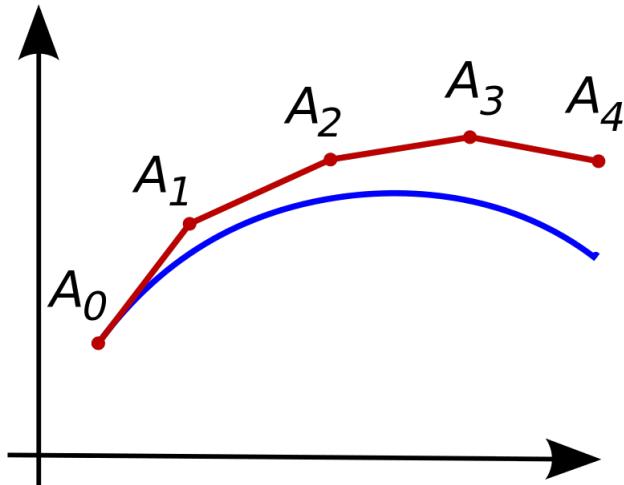
Using the Euler formula:

$$y_1 = y_0 + h \cdot f(t_0, y_0) = 1 + 0.1 \cdot 1 = 1.1$$

$$y_2 = y_1 + h \cdot f(t_1, y_1) = 1.1 + 0.1 \cdot 1.1 = 1.21$$

$$y_3 = y_2 + h \cdot f(t_2, y_2) = 1.21 + 0.1 \cdot 1.21 = 1.331$$

Euler's method can be visualized as taking small steps along the curve defined by the differential equation, using the slope at the current point to determine the direction of the next step.



Advantages:

- Simple to understand and implement.
- Requires only basic arithmetic operations.

Disadvantages:

- Low accuracy for large step sizes.
- Can become unstable if the step size is not chosen appropriately.
- Errors accumulate over time, leading to less accurate solutions.

Euler's method is often used as a basic introduction to numerical methods for solving ODEs, and more sophisticated methods like the *Runge-Kutta* methods are used for more accurate solutions.

29.2 Neural ODE

Models such as residual networks, recurrent neural network decoders, and normalizing flows build complicated transformations by composing a sequence of transformations to a hidden state:

$$\mathbf{h}_{t+1} = \mathbf{h}_t + f(\mathbf{h}_t, \theta_t).$$

The \mathbf{h} is iteratively updated as follows:

$$\begin{aligned}\mathbf{h}_2 &= \mathbf{h}_1 + f(\mathbf{h}_1, \theta) \\ \mathbf{h}_3 &= \mathbf{h}_2 + f(\mathbf{h}_2, \theta) = \mathbf{h}_1 + f(\mathbf{h}_1, \theta) + f(\mathbf{h}_2, \theta) \\ &\vdots\end{aligned}$$

These iterative update can be seen as *Euler discretization* of the *continuous* transformation. Think of traditional neural networks as a sequence of steps. You give it some input, it goes through several steps (layers), and you get an output. Instead of thinking in steps, Neural ODEs think in **continuous change over time**. Note that this is the key contribution of this approach.

Euler method can be expressed as follows:

$$y_n = y_{n-1} + h \frac{\partial y_{n-1}}{\partial x_{n-1}},$$

where h is the step size. In NODE, they view the f as an ordinary differential equation, which depends on the state at time t and parameter θ . The following equation is the shape of Euler method:

$$y_n = y_1 + h \frac{\partial y_1}{\partial x_1} + h \frac{\partial y_2}{\partial x_2} + \cdots + h \frac{\partial y_{n-1}}{\partial x_{n-1}}.$$

In NODE,

Chapter 30

State Estimations

30.1 Introduction to State-Space Model

Reference: [State-Space Models](#).

A state-space model is a mathematical framework used to describe a system by a set of input, output, and state variables related by first-order differential (or difference) equations. It's widely used in control theory, signal processing, and time series analysis.

For a continuous-time system, the state-space model is typically expressed as follows:

$$\begin{aligned}\dot{x}(t) &= Ax(t) + Bu(t), \\ y(t) &= Cx(t) + Du(t).\end{aligned}$$

The first and the second equations are known as *state equation* and *output equation*, respectively. The state equation tells us that how the state vector changes with the state vector and the (external) input. The state space model is **linear** and **time invariant**, since the equations are linear and the parameter matrices do not change over time. Such systems are referred as a LTI system.

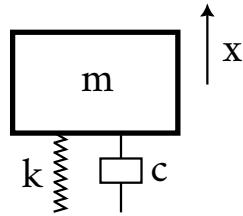
- $x(t) \in \mathbb{R}^n$: A state vector. This variable describes the state of the system at any given time.
- $\dot{x} \in \mathbb{R}^n$: state derivative *i.e.*, $(\frac{dX}{dt})$ represents the changes of the state vector. You can notice that this is a linear combination of state vector and the input vector.
- $u \in \mathbb{R}^m$: An external input affecting the system.
- $y \in \mathbb{R}^p$: An (Observed) Output.
- $A \in \mathbb{R}^{n \times n}$: A state matrix. This matrix describes how the state vector influences the changes of the state vector.
- $B \in \mathbb{R}^{n \times m}$: An input matrix. This matrix describes how the (external) input vector influences the changes of the state vector.
- $C \in \mathbb{R}^{p \times n}$: An output matrix. Typically, an identity matrix (I).
- $D \in \mathbb{R}^{p \times m}$: A feedthrough (or direct transmission) matrix. Typically, zeros

30.1.1 Example: Mass-Spring-Damper System

Let's consider a simple example of a mass-spring-damper system, which can be described by the second-order differential equation.

The mass-spring-damper system is a common mechanical system that consists of three main components:

- Mass (m): A mass that can move along a straight line.
- Spring (k): A spring that exerts a force proportional to its displacement from its equilibrium position (Hooke's Law).
- Damper (c): A damping element (like a shock absorber) that exerts a force proportional to the velocity of the mass (damping force).



System Dynamics The dynamics of this system can be described by Newton's second law of motion, which states that the sum of forces acting on the mass is equal to the mass times its acceleration ($F = ma$).

Differential Equation For a mass-spring-damper system, the forces are:

- Spring Force: $F_{\text{spring}} = -kx(t)$, where $x(t)$ is the displacement of the mass from its equilibrium position. The proportional constant k is called the spring constant. It is a measure of the spring's stiffness.
- Damping Force: Damping forces are a special type of force that are used to slow down or stop a motion. $F_{\text{damper}} = -c\dot{x}(t)$, where $\dot{x}(t)$ is the velocity of the mass (or object). Note that $\dot{x}(t) = dx(t)/dt$.
- External Force: $F(t)$, an external force applied to the mass.

By summing these forces, we get:

$$m\ddot{x}(t) = -kx(t) - c\dot{x}(t) + F(t).$$

To explain, the external force is stretching the spring, and the damper and the spring force are pulling the mass. Rearranging this, we get the second-order differential equation:

$$m\ddot{x}(t) + c\dot{x}(t) + kx(t) = F(t)$$

State-Space Representation To convert this second-order differential equation into a state-space representation, we need to express it as a system of first-order differential equations.

Defining State Variables We introduce two state variables:

- $x_1(t) = x(t)$: the position of the mass.
- $x_2(t) = \dot{x}(t)$: the velocity of the mass.

Now, we can write the original second-order equation as two first-order equations:

$$\begin{aligned}\dot{x}_1(t) &= x_2(t) \\ \dot{x}_2(t) &= \frac{1}{m}F(t) - \frac{k}{m}x_1(t) - \frac{c}{m}x_2(t)\end{aligned}$$

Matrix Form We can express these equations in matrix form:

$$\begin{bmatrix} \dot{x}_1(t) \\ \dot{x}_2(t) \end{bmatrix} = \begin{bmatrix} 0 & 1 \\ -\frac{k}{m} & -\frac{c}{m} \end{bmatrix} \begin{bmatrix} x_1(t) \\ x_2(t) \end{bmatrix} + \begin{bmatrix} 0 \\ \frac{1}{m} \end{bmatrix} F(t)$$

This is the state-space form:

$$\dot{x}(t) = Ax(t) + Bu(t)$$

Where:

- $x(t) = \begin{bmatrix} x_1(t) \\ x_2(t) \end{bmatrix}$ is the state vector.
- $u(t) = F(t)$ is the input (external force).
- $A = \begin{bmatrix} 0 & 1 \\ -\frac{k}{m} & -\frac{c}{m} \end{bmatrix}$ is the state matrix.
- $B = \begin{bmatrix} 0 \\ \frac{1}{m} \end{bmatrix}$ is the input matrix.

Output Equation If we consider the output to be the position of the mass ($x_1(t)$), the output equation is:

$$y(t) = [1 \ 0] \begin{bmatrix} x_1(t) \\ x_2(t) \end{bmatrix}$$

So, the output equation is:

$$y(t) = Cx(t),$$

Where $C = [1 \ 0]$.

Full State-Space Model Combining the state and output equations, we get the full state-space representation:

$$\begin{aligned}\dot{x}(t) &= \begin{bmatrix} 0 & 1 \\ -\frac{k}{m} & -\frac{c}{m} \end{bmatrix} x(t) + \begin{bmatrix} 0 \\ \frac{1}{m} \end{bmatrix} u(t) \\ y(t) &= [1 \ 0] x(t)\end{aligned}$$

In sum,

- State variables: $x_1(t) = x(t)$ (position), $x_2(t) = \dot{x}(t)$ (velocity)
- Input: $u(t) = F(t)$ (external force)
- Output: $y(t) = x(t)$ (position)
- State equation: $\dot{x}(t) = Ax(t) + Bu(t)$
- Output equation: $y(t) = Cx(t)$

This state-space model describes how the position and velocity of the mass change over time in response to an external force.

30.1.2 Stability

The linear state space model is stable if all eigenvalues of A are negative real numbers or have negative real parts to complex number eigenvalues. If all real parts of the eigenvalues are negative then the system is stable, meaning that any initial condition converges exponentially to a stable attracting point. If any real parts are zero then the system will not converge to a point and if the eigenvalues are positive the system is unstable and will exponentially diverge.

Chapter 31

Kalman Filter

31.1 Propagation of States and Covariances

Suppose we have a linear discrete-time system:

$$\boldsymbol{\theta}_k = F_{k-1}\boldsymbol{\theta}_{k-1} + G_{k-1}\mathbf{u}_{k-1} + \mathbf{w}_{k-1},$$

where \mathbf{w}_k is a Gaussian zero-mean white noise with covariance Q_k and \mathbf{u}_k is a known input. The F and G are state and input matrices, respectively. How does the mean of the state $\boldsymbol{\theta}_k$ changes over time? The expectation is given by

$$\mathbb{E}[\boldsymbol{\theta}_k] = \mathbb{E}[F_{k-1}\boldsymbol{\theta}_{k-1}] + \mathbb{E}[G_{k-1}\mathbf{u}_{k-1}] + \mathbb{E}[\mathbf{w}_{k-1}]$$

For simplicity, we can write it as

$$\bar{\boldsymbol{\theta}}_k = F_{k-1}\bar{\boldsymbol{\theta}}_{k-1} + G_{k-1}\mathbf{u}_{k-1}.$$

How does the state covariance of $\boldsymbol{\theta}_k$ change with time? The state covariance matrix propagation is given by

$$P_k = \mathbb{E}[(\boldsymbol{\theta}_k - \bar{\boldsymbol{\theta}}_k)(\boldsymbol{\theta}_k - \bar{\boldsymbol{\theta}}_k)^T]$$

and then compute the expectation of every term in that expression.

$$\begin{aligned} (\boldsymbol{\theta}_k - \bar{\boldsymbol{\theta}}_k)(\boldsymbol{\theta}_k - \bar{\boldsymbol{\theta}}_k)^T &= (F_{k-1}(\boldsymbol{\theta}_{k-1} - \bar{\boldsymbol{\theta}}_{k-1}) + \mathbf{w}_{k-1})(F_{k-1}(\boldsymbol{\theta}_{k-1} - \bar{\boldsymbol{\theta}}_{k-1}) + \mathbf{w}_{k-1})^T \\ &= F_{k-1}(\boldsymbol{\theta}_{k-1} - \bar{\boldsymbol{\theta}}_{k-1})(\boldsymbol{\theta}_{k-1} - \bar{\boldsymbol{\theta}}_{k-1})^T F_{k-1}^T + F_{k-1}(\boldsymbol{\theta}_{k-1} - \bar{\boldsymbol{\theta}}_{k-1})\mathbf{w}_{k-1}^T \\ &\quad + \mathbf{w}_{k-1}(\boldsymbol{\theta}_{k-1} - \bar{\boldsymbol{\theta}}_{k-1})^T F_{k-1}^T + \mathbf{w}_{k-1}\mathbf{w}_{k-1}^T \end{aligned}$$

Since $(\boldsymbol{\theta}_k - \bar{\boldsymbol{\theta}}_k)$ is uncorrelated to \mathbf{w}_{k-1} , we have

$$\begin{aligned} \mathbb{E}[F_{k-1}(\boldsymbol{\theta}_{k-1} - \bar{\boldsymbol{\theta}}_{k-1})\mathbf{w}_{k-1}^T] &= 0 \\ \mathbb{E}[\mathbf{w}_{k-1}(\boldsymbol{\theta}_{k-1} - \bar{\boldsymbol{\theta}}_{k-1})^T F_{k-1}^T] &= 0 \end{aligned}$$

Also, we have

$$\mathbb{E}[\mathbf{w}_{k-1}\mathbf{w}_{k-1}^T] = Q_{k-1}$$

$$P_{k-1} = \mathbb{E}[(\boldsymbol{\theta}_{k-1} - \bar{\boldsymbol{\theta}}_{k-1})(\boldsymbol{\theta}_{k-1} - \bar{\boldsymbol{\theta}}_{k-1})^T]$$

By using these expressions, we obtain the final equation for the propagation of the state covariance matrix

$$P_k = F_{k-1} P_{k-1} F_{k-1}^T + Q_{k-1}$$

This is called a discrete time *Lyapunov equation*, or a *Stein equation*. It is interesting to consider the conditions under which the discrete time Lyapunov equation has a *steady-state solution*.

31.2 Kalman Filtering

We are considering the following state-space model of a dynamical system:

$$\begin{aligned}\boldsymbol{\theta}_k &= F_{k-1} \boldsymbol{\theta}_{k-1} + G_{k-1} \mathbf{u}_{k-1} + \mathbf{w}_{k-1} \\ \mathbf{y}_k &= \mathbf{X}_k \boldsymbol{\theta}_k + \boldsymbol{\eta}_k,\end{aligned}$$

where

- $\mathbf{w}_k \sim (0, Q_k)$: process noise vector
- $E[\mathbf{w}_k \mathbf{w}_j^T] = Q_k \delta_{k-j}$
- $\boldsymbol{\eta}_k \sim (0, R_k)$: measurement noise vector
- $E[\boldsymbol{\eta}_k \boldsymbol{\eta}_j^T] = R_k \delta_{k-j}$
- $E[\boldsymbol{\eta}_k \mathbf{w}_j^T] = 0$
- \mathbf{u}_{k-1} : control unit vector
- F and G are state and input matrices.

Our primary goal is to estimate the state $\boldsymbol{\theta}_k$ based on our knowledge of the system dynamics and the availability of the noisy measurements \mathbf{y}_k .

If we have all of the measurements up to and including time k available for use in our estimate of $\boldsymbol{\theta}_k$, then we can form an a *posteriori estimate*, which we denote as $\hat{\boldsymbol{\theta}}_k^+$. The + superscript denotes that the estimate is a posteriori. One way to form the a posteriori state estimate is to compute the expected value of $\boldsymbol{\theta}_k$ conditioned on all of the measurements up to and including time k :

$$\hat{\boldsymbol{\theta}}_k^+ = E[\boldsymbol{\theta}_k | \mathbf{y}_1, \dots, \mathbf{y}_{k-1}, \mathbf{y}_k].$$

If we have all of the measurements before (but not including) time k available for use in our estimate of $\boldsymbol{\theta}_k$, then we can form an a priori estimate, which we denote

$$\hat{\boldsymbol{\theta}}_k^-.$$

One way to form the a *priori state estimate* is to compute the expected value of $\boldsymbol{\theta}_k$, conditioned on all of the measurements before (but not including) time k :

$$\hat{\boldsymbol{\theta}}_k^- = E[\boldsymbol{\theta}_k | \mathbf{y}_1, \dots, \mathbf{y}_{k-1}].$$

By their nature, it is natural to expect that $\hat{\boldsymbol{\theta}}_k^+$ to be a better estimate than $\hat{\boldsymbol{\theta}}_k^-$, since we can leverage more information to compute it.

If we have measurements after time k available for use in our estimate of $\boldsymbol{\theta}_k$, then we can form a *smoothed estimate*. One way to form the smoothed state estimate is to compute the expected value of $\boldsymbol{\theta}_k$ conditioned on all of the measurements that are available:

$$\hat{\boldsymbol{\theta}}_{k|k+N} = E[\boldsymbol{\theta}_k | \mathbf{y}_1, \dots, \mathbf{y}_k, \dots, \mathbf{y}_{k+N}],$$

where N is some positive integer. If we want to find the best prediction of $\boldsymbol{\theta}_k$ more than one time step ahead of the available measurements, then we can form a predicted estimate. One way to form the *predicted state estimate* is to compute the expected value of $\boldsymbol{\theta}_k$ conditioned on all of the measurements that are available:

$$\hat{\boldsymbol{\theta}}_{k|k-M} = E[\boldsymbol{\theta}_k | \mathbf{y}_1, \dots, \mathbf{y}_{k-M}].$$

The relationship between the a posteriori, a priori, smoothed, and predicted state estimates is represented in the following figure:

As done in RLS (*c.f.*, §10.1), we should define covariance matrices of the a priori and the a posteriori estimation errors. The covariance matrices are defined as follows:

$$\begin{aligned} P_k^- &= \mathbb{E}[(\boldsymbol{\theta}_k - \boldsymbol{\theta}_k^-)(\boldsymbol{\theta}_k - \boldsymbol{\theta}_k^-)^T] \\ P_k^+ &= \mathbb{E}[(\boldsymbol{\theta}_k - \boldsymbol{\theta}_k^+)(\boldsymbol{\theta}_k - \boldsymbol{\theta}_k^+)^T], \end{aligned}$$

where P_k^- is the covariance matrix of the a priori estimation error and P_k^+ is the covariance matrix of the a posteriori estimation error.



After the discrete-time instant $k-1$, we compute the a posteriori state estimate $\boldsymbol{\theta}_{k-1}^+$ and the covariance matrix of the estimation error P_{k-1}^+ . Then, we propagate these quantities through the equations describing the system dynamics and covariance matrix propagation (that is also derived on the basis of the system dynamics). That is, we propagate the a posteriori state and covariance through our model. By propagating these quantities, we obtain the a priori state estimate $\boldsymbol{\theta}_k^-$ and the covariance matrix of the estimation error P_k^- for the time step k . Then, after the measurement vector \mathbf{y}_k is observed at the discrete-time step k , we use this measurement and the recursive-least squares method to compute the a posteriori estimate $\boldsymbol{\theta}_k^+$ and the covariance matrix of the estimation error P_k^+ for the time instant k .

This is how Kalman filter works. Now let's derive the equations of Kalman filter. At the initial time instant $k = 0$, we need to set an initial guess of the state estimate, $\hat{\boldsymbol{\theta}}_0$. This will be our initial a posteriori state estimate, that is

$$\hat{\boldsymbol{\theta}}_0^+ = \hat{\boldsymbol{\theta}}_0$$

Then, the question is how to compute the a priori estimate $\hat{\boldsymbol{\theta}}_1^-$ at $k = 1$. The natural answer is that the system states, as well as the estimates, need to satisfy the system dynamics (1), and consequently,

$$\hat{\boldsymbol{\theta}}_1^- = F_0 \hat{\boldsymbol{\theta}}_0^+ + F_0 \mathbf{u}_0$$

where we excluded the disturbance part since the disturbance vector is not known. Besides the initial guess of the estimate, we also need to select an initial guess of the covariance matrix of the estimation error. That is, we need to select P_0^+ . If we have perfect knowledge about the initial state, then we select P_0^+ as a zero matrix, that is $P_0^+ = 0 \cdot I$, where I is an identity matrix. This is because the covariance matrix of the estimation error is the measure of uncertainty, and if we have perfect knowledge, then the measure of uncertainty should be zero. On the other hand, if we do not have any a priori knowledge of the initial state, then $P_0^+ = c \cdot I$, where c is a large number. Next, we need to compute the covariance matrix of the a priori state estimation error, that is, we need to compute P_1^- . In our previous post, which can be found here, we derived the expression for the time propagation of the covariance matrix of the state estimation error:

$$P_k = F_{k-1} P_{k-1} F_{k-1}^T + Q_{k-1}$$

By using this expression, we obtain the following equation for P_1^- .

31.2.1 Python Implementation

```

1  class KalmanFilter(object):
2
3      # x0 - initial guess of the state vector
4      # P0 - initial guess of the covariance matrix of the state estimation error
5      # A,B,C - system matrices describing the system model
6      # Q - covariance matrix of the process noise
7      # R - covariance matrix of the measurement noise
8
9      def __init__(self,x0,P0,A,B,C,Q,R):
10
11         # initialize vectors and matrices
12         self.x0=x0
13         self.P0=P0
14         self.A=A
15         self.B=B
16         self.C=C
17         self.Q=Q
18         self.R=R
19
20         # this variable is used to track the current time step k of the
21         # estimator
22         # after every measurement arrives, this variables is incremented for +1
23         self.currentTimeStep=0
24
25         # this list is used to store the a posteriori estimates xk^{+} starting
26         # from the initial estimate
27         # note: list starts from x0^{+}=x0 - where x0 is an initial guess of
28         # the estimate
29         self.estimates_aposteriori=[]
30         self.estimates_aposteriori.append(x0)
31
32         # this list is used to store the a apriori estimates xk^{-} starting
33         # from x1^{-}
34         # note: x0^{-} does not exist, that is, the list starts from x1^{-}
35         self.estimates_apriori=[]
36
37         # this list is used to store the a posteriori estimation error
38         covariance matrices Pk^{+}
39         # note: list starts from P0^{+}=P0, where P0 is the initial guess of
40         # the covariance
41         self.estimationErrorCovarianceMatricesApriori=[]

```

```

36         self.estimationErrorCovarianceMatricesApriori.append(P0)
37
38     # this list is used to store the a priori estimation error covariance
39     # matrices Pk^[-]
40     # note: list starts from P1^[-], that is, P0^[-] does not exist
41     self.estimationErrorCovarianceMatricesApriori=[]
42
43     # this list is used to store the gain matrices Kk
44     self.gainMatrices=[]
45
46     # this list is used to store prediction errors error_k=y_k-C*xk^[-]
47     self.errors=[]
48
49     # this function propagates x_{k-1}^[+] through the model to compute x_{k
50     #}^[-]
51     # this function also propagates P_{k-1}^[+] through the covariance model to
52     # compute P_{k}^[-]
53     # at the end this function increments the time index currentTimestep for +1
54     def propagateDynamics(self,inputValue):
55
56         xk_minus=self.A*self.estimates_aposteriori[self.currentTimestep]+self.B
57         *inputValue
58         Pk_minus=self.A*self.estimationErrorCovarianceMatricesApriori[self.
59         currentTimestep]*(self.A.T)+self.Q
60
61         self.estimates_apriori.append(xk_minus)
62         self.estimationErrorCovarianceMatricesApriori.append(Pk_minus)
63
64         self.currentTimestep=self.currentTimestep+1
65
66     # this function should be called after propagateDynamics() because the time
67     # step should be increased and states and covariances should be propagated
68     def computeAprioriEstimate(self,currentMeasurement):
69         import numpy as np
70         # gain matrix
71         Kk=self.estimationErrorCovarianceMatricesApriori[self.currentTimestep
72         -1]*(self.C.T)*np.linalg.inv(self.R+self.C*self.
73         estimationErrorCovarianceMatricesApriori[self.currentTimestep-1]*(self.C.T))
74
75         # prediction error
76         error_k=currentMeasurement-self.C*self.estimates_apriori[self.
77         currentTimestep-1]
78         # a posteriori estimate
79         xk_plus=self.estimates_apriori[self.currentTimestep-1]+Kk*error_k
80
81         # a posteriori matrix update
82         IminusKkC=np.matrix(np.eye(self.x0.shape[0]))-Kk*self.C
83         Pk_plus=IminusKkC*self.estimationErrorCovarianceMatricesApriori[self.
84         currentTimestep-1]*(IminusKkC.T)+Kk*(self.R)*(Kk.T)
85
86         # update the lists that store the vectors and matrices
87         self.gainMatrices.append(Kk)
88         self.errors.append(error_k)
89         self.estimates_aposteriori.append(xk_plus)
90         self.estimationErrorCovarianceMatricesApriori.append(Pk_plus)

```

Chapter 32

Deep State Space Models

32.1 Efficiently Modeling Long Sequences with Structured State-Spaces

The Linear State-Space Layer (LSSL) is a simple sequence model that maps a one-dimensional function or sequence $u(t) \rightarrow y(t)$ through an implicit state $x(t)$ by simulating a linear continuous-time state-space representation in discrete-time

$$\begin{aligned}x'(t) &= \mathbf{A}x(t) + \mathbf{B}u(t), \\y(t) &= \mathbf{C}x(t) + \mathbf{D}u(t).\end{aligned}$$

The first equation maps a single dimensional input signal $u(t)$ (or sequence) to an N -dim latent state (or hidden state) $x'(t)$ with the current state $x(t)$. The \mathbf{A} and \mathbf{B} can be considered as a non-linear mapping or transition matrices (*i.e.*, learnable parameters) to reflect the impact of the current state and the input, respectively. Finally, we project the input and the updated state to a one-dim output signal $y(t)$ (*i.e.*, sequence).

Our goal is to simply use the SSM as a black-box representation in a deep sequence model, where \mathbf{A} , \mathbf{B} , \mathbf{C} , and \mathbf{D} are parameters learned by gradient descent. We will omit the parameter \mathbf{D} for exposition (or equivalently, assume $\mathbf{D} = 0$, because the term $\mathbf{D}u$ can be viewed as a *skip connection* that doesn't depend on the hidden state (x)).

An SSM maps a input $u(t)$ to a state representation vector $x(t)$ and an output $y(t)$. For simplicity, we assume the input and output are one-dimensional, and the state representation is N -dimensional. The first equation defines the change in $x(t)$ over time.

Discretization We want to find a discrete-time state-space model. We can represent it by approximating a continuous model (*i.e.*, $h(t_k) \approx h(k\Delta)$) as follows:

$$\begin{aligned}x'(t) &= \overline{\mathbf{A}}x(t) + \overline{\mathbf{B}}u(t), \\y(t) &= \mathbf{C}x(t) + \underbrace{\mathbf{D}u(t)}_{=0},\end{aligned}$$

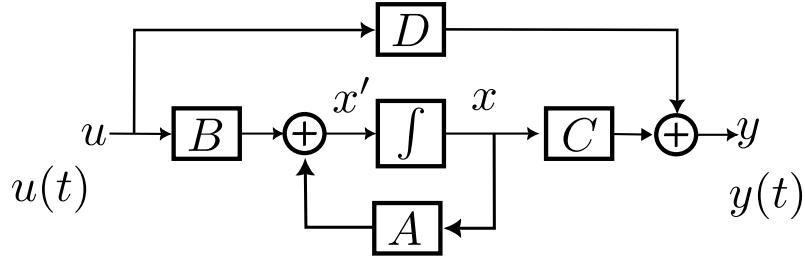


Figure 32.1: S4 Model

where $\bar{\mathbf{A}} = \mathbf{I} + \Delta \mathbf{A}$ and $\bar{\mathbf{B}} = \Delta \mathbf{B}$, respectively. They are derived by

$$\begin{aligned} x'(t) &= \bar{\mathbf{A}}x(t) + \bar{\mathbf{B}}u(t) \\ &= \lim_{\Delta \rightarrow 0} \frac{x(t + \Delta) - x(t)}{\Delta}, \end{aligned}$$

where Δ is a step size. Note that Δ is a learnable parameter to be determined during a training phase. Subsequently, we get

$$x(t + \Delta) - x(t) = \Delta x'(t).$$

Equivalently,

$$x(t + \Delta) = \Delta x'(t) + x(t).$$

Plugging $x'(t)$ into the above equation, we get:

$$\begin{aligned} x(t + \Delta) &= \Delta(\mathbf{A}x(t) + \mathbf{B}u(t)) + x(t) \\ &= \underbrace{(\mathbf{I} + \Delta \mathbf{A})}_{\bar{\mathbf{A}}} x(t) + \underbrace{\Delta \mathbf{B}}_{\bar{\mathbf{B}}} u(t). \end{aligned}$$

We can say $x(t + \Delta) = x(t + 1)$, since it is the next state we care. Thus,

$$x(t + 1) = (\mathbf{I} + \Delta \mathbf{A})x(t) + \Delta \mathbf{B}u(t).$$

Note that the original paper utilizes a special rule called *Zero-Order Hold* to approximate the $\bar{\mathbf{A}}$ and $\bar{\mathbf{B}}$.

Alternatively, we can use the trapezoidal method. The trapezoidal rule works by approximating the region under the graph of the function $f(x)$ as a trapezoid and calculating its area. It follows that

$$\int_a^b f(x) dx \approx (b - a) \cdot \frac{1}{2}(f(a) + f(b)).$$

Thus, the integral of $x'(t)$ from t_n to t_{n+1} can be approximated using the trapezoidal rule. The exact integral is:

$$\begin{aligned} x(t_{n+1}) - x(t_n) &= \int_{t_n}^{t_{n+1}} x'(t) dt \\ &\approx \frac{1}{2}\Delta(\mathbf{A}x(t_{n+1}) + \mathbf{B}u(t_{n+1}) + \mathbf{A}x(t_n) + \mathbf{B}u(t_n)), \end{aligned}$$

where $\Delta = t_{n+1} - t_n$. Then, we have

$$x(t_{n+1}) - \frac{\Delta}{2}\mathbf{A}x(t_{n+1}) = x(t_n) + \frac{\Delta}{2}\mathbf{A}x(t_n) + \frac{\Delta}{2}\mathbf{B}u(t_n) + \frac{\Delta}{2}\mathbf{B}u(t_{n+1})$$

$$\begin{aligned} \left(\mathbf{I} - \frac{\Delta}{2}\mathbf{A}\right)x(t_{n+1}) &= \left(\mathbf{I} - \frac{\Delta}{2}\mathbf{A}\right)x(t_n) + \Delta\mathbf{B} \frac{(u(t_n) + u(t_{n+1}))}{2} \\ x(t_{n+1}) &= \left(\mathbf{I} - \frac{\Delta}{2}\mathbf{A}\right)^{-1} \left(\mathbf{I} + \frac{\Delta}{2}\mathbf{A}\right)x(t_n) + \left(\mathbf{I} - \frac{\Delta}{2}\mathbf{A}\right)^{-1} \frac{\Delta}{2}\mathbf{B}(u(t_{n+1})) \end{aligned}$$

Finally, we get

$$\begin{aligned} \overline{\mathbf{A}} &= \left(\mathbf{I} - \frac{\Delta}{2}\mathbf{A}\right)^{-1} \left(\mathbf{I} + \frac{\Delta}{2}\mathbf{A}\right) \\ \overline{\mathbf{B}} &= \left(\mathbf{I} - \frac{\Delta}{2}\mathbf{A}\right)^{-1} \Delta\mathbf{B} \end{aligned}$$

Note that we assume that $u(t_{n+1}) \approx u(t_n)$. We can represent the update process as follows: At $t = 0$

$$\begin{aligned} x(0) &= \overline{\mathbf{B}}u(0), \\ y(0) &= \mathbf{C}x(0) \end{aligned}$$

At $t = 1$

$$\begin{aligned} x(1) &= \overline{\mathbf{A}}x(0) + \overline{\mathbf{B}}u(1), \\ y(1) &= \mathbf{C}x(1) \end{aligned}$$

At $t = 2$

$$\begin{aligned} x(2) &= \overline{\mathbf{A}}x(1) + \overline{\mathbf{B}}u(2), \\ y(2) &= \mathbf{C}x(2). \end{aligned}$$

Note that this update process is equivalent to the RNN's update process.

As a Convolution The above process can be viewed as an one-dimensional convolution.

$$\begin{aligned} x(0) &= \overline{\mathbf{B}}u(0), \\ y(0) &= \mathbf{C}x(0) = \mathbf{C}\overline{\mathbf{B}}u(0) \end{aligned}$$

$$\begin{aligned} x(1) &= \overline{\mathbf{A}}x(0) + \overline{\mathbf{B}}u(1) = \overline{\mathbf{A}}\overline{\mathbf{B}}u(0) + \overline{\mathbf{B}}u(1) \\ y(1) &= \mathbf{C}x(1) = \mathbf{C}(\overline{\mathbf{A}}\overline{\mathbf{B}}u(0) + \overline{\mathbf{B}}u(1)) = \mathbf{C}\overline{\mathbf{A}}\overline{\mathbf{B}}u(0) + \mathbf{C}\overline{\mathbf{B}}u(1) \end{aligned}$$

$$\begin{aligned} x(2) &= \overline{\mathbf{A}}x(1) + \overline{\mathbf{B}}u(2) = \overline{\mathbf{A}}(\overline{\mathbf{A}}\overline{\mathbf{B}}u(0) + \overline{\mathbf{B}}u(1)) + \overline{\mathbf{B}}u(2) = \overline{\mathbf{A}}^2\overline{\mathbf{B}}u(0) + \overline{\mathbf{A}}\overline{\mathbf{B}}u(1) + \overline{\mathbf{B}}u(2) \\ y(2) &= \mathbf{C}x(2) = \mathbf{C}(\overline{\mathbf{A}}^2\overline{\mathbf{B}}u(0) + \overline{\mathbf{A}}\overline{\mathbf{B}}u(1) + \overline{\mathbf{B}}u(2)) = \mathbf{C}\overline{\mathbf{A}}^2\overline{\mathbf{B}}u(0) + \mathbf{C}\overline{\mathbf{A}}\overline{\mathbf{B}}u(1) + \mathbf{C}\overline{\mathbf{B}}u(2) \end{aligned}$$

We get a general formula:

$$\begin{aligned} y(t) &= \mathbf{C}\overline{\mathbf{A}}^t\overline{\mathbf{B}}u(0) + \mathbf{C}\overline{\mathbf{A}}^{t-1}\overline{\mathbf{B}}u(1) + \cdots + \mathbf{C}\overline{\mathbf{B}}u(t) \\ &= \sum_{t=0}^T \mathbf{C}\overline{\mathbf{A}}^{T-t}\overline{\mathbf{B}}u(t) \end{aligned}$$

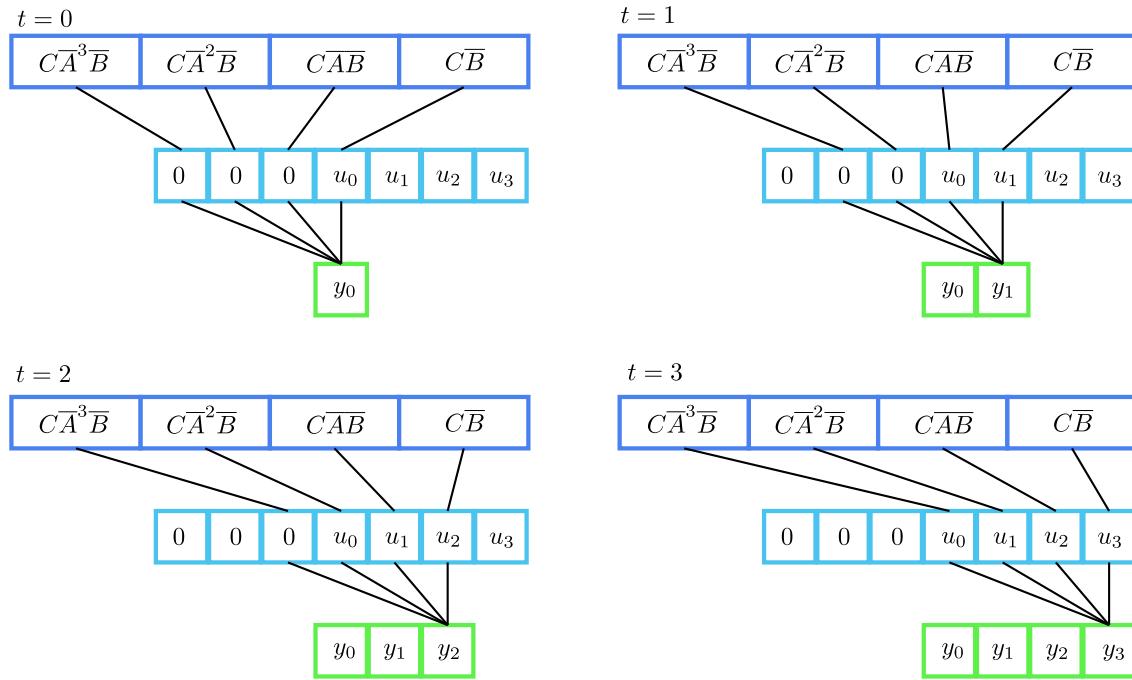


Figure 32.2:

It turns out that the above equation is a one-dimensional convolution by a kernel $\bar{\mathbf{K}}$:

$$y = x * \bar{\mathbf{K}}.$$

It is generally referred to as the SSM convolution kernel in the literature, and its size is equivalent to the entire input sequence. This convolution kernel is calculated by Fast Fourier Transform (FFT)

Let's say the kernel size is 4 with zero-padding, (See Fig. 32.2). During training, we can train the model as a convolutional neural network so that we can leverage the parallel training. During inference (*i.e.*, decoding stage), we can switch to the recurrent mode for near-constant time inference. Please note here, that if you look at the kernels you can see that they are fixed.

In the convolution kernel developed above, $\bar{\mathbf{C}}$ and $\bar{\mathbf{B}}$, are learnable scalars. Concerning $\bar{\mathbf{A}}$, we've seen that in our convolution kernel, it's expressed as a power of k at time k . This can be very time-consuming to calculate, so we're looking for a fixed $\bar{\mathbf{A}}$. For this, the best option is to have it diagonal:

Note that we can scale the single-dimensional SSM to multi-dimensional vector by putting a SSM for each dimension. For instance, there is going to be 256 SSMs if we want to learn a 256 embeddings.

32.2 Mamba: Linear-Time Sequence Modeling with Selective State Spaces

<https://blog.premai.io/s4-and-mamba/>

We can immediately notice the importance of the matrix \mathbf{A} . The Mamba leverages a special matrix called HiPPO. The HiPPO is a $N \times N$ matrix specifically designed to approximate all the

input signals by using Legendre polynomials (like Fourier series).

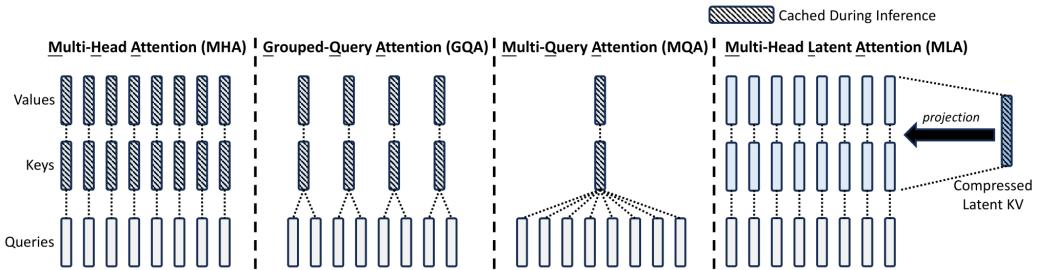
$$\mathbf{A}_{nk} = \begin{cases} (2n+1)^{1/2}(2k+1)^{1/2} & \text{if } n > k \\ n+1 & \text{if } n = k \\ 0 & \text{else} \end{cases}$$

This matrix helps to compress the history of the information by masking the next tokens, which is similar to the masked self-attention. Note that the matrix just needs to be computed once.

Chapter 33

DeepSeek

33.1 Multi-Head Latent Attention



The query, key, and value of the vanilla multi-head attention can be expressed as follows:

$$\begin{aligned}\mathbf{q}_t &= W^Q \mathbf{h}_t \\ \mathbf{k}_t &= W^K \mathbf{h}_t \\ \mathbf{v}_t &= W^V \mathbf{h}_t\end{aligned}$$

- $\mathbf{q}_t, \mathbf{k}_t, \mathbf{v}_t \in \mathbb{R}^{d_h n_h}$
- $\mathbf{h}_t \in \mathbb{R}^d$: Attention input of the t -th token at an layer.
- d_h : the attention head's dimension
- n_h : the number of attention heads

During inference, all keys and values need to be cached to accelerate inference, so MHA needs to cache $2n_h d_h l$ elements (*i.e.*, key, value for each layer and head) for each token. In model deployment, this heavy KV cache is a large bottleneck that limits the maximum batch size and sequence length.

In DeepSeek The key idea of Multi-Head Latent Attention (MLA) is the low-rank joint compression

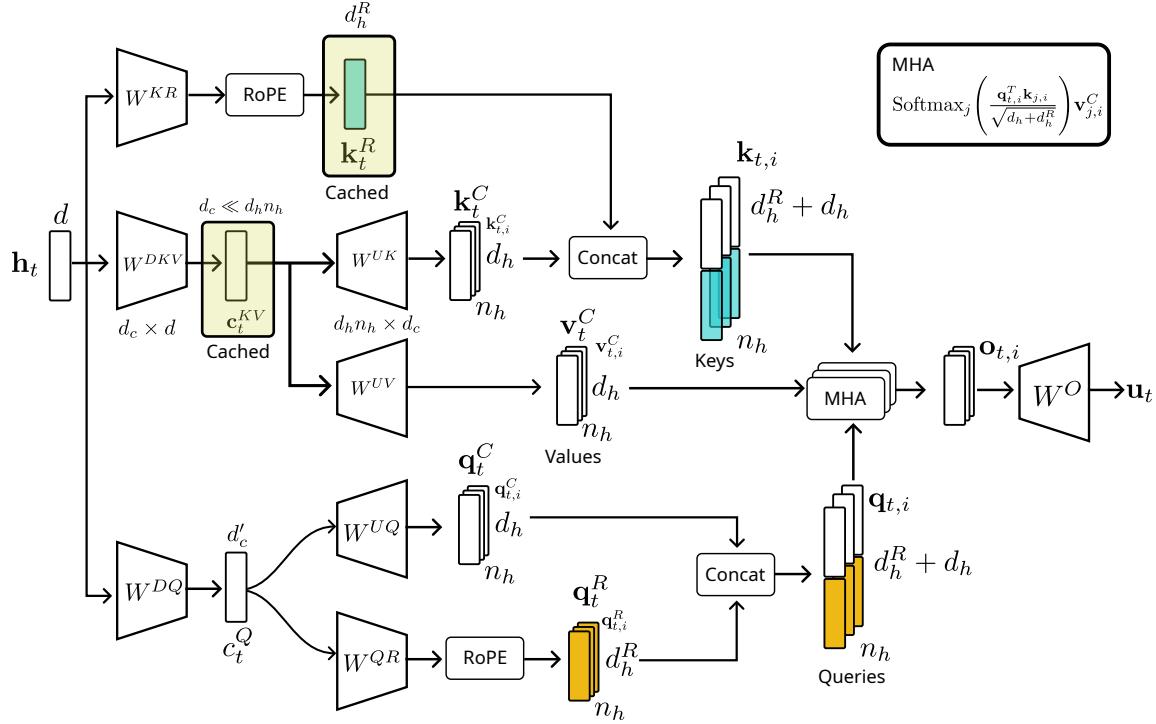


Figure 33.1: An overview of MLA.

for attention keys and values to reduce Key-Value (KV) cache during inference:

$$\begin{aligned}
 \mathbf{c}_t^{KV} &= W^{DKV} \mathbf{h}_t \\
 [\mathbf{k}_{t,1}^C; \mathbf{k}_{t,2}^C; \dots; \mathbf{k}_{t,n_h}^C] &= \mathbf{k}_t^C = W^{UK} \mathbf{c}_t^{KV} \\
 \mathbf{k}_t^R &= \text{RoPE}(W^{KR} \mathbf{h}_t) \\
 \mathbf{k}_{t,i} &= [\mathbf{k}_{t,i}^C; \mathbf{k}_t^R] \\
 [\mathbf{v}_{t,1}^C; \mathbf{v}_{t,2}^C; \dots; \mathbf{v}_{t,n_h}^C] &= \mathbf{v}_t^C = W^{UV} \mathbf{c}_t^{KV}
 \end{aligned}$$

- D and U superscripts denote the up and down projection
- $\mathbf{c}_t^{KV} \in \mathbb{R}^{d_c}$ is the compressed (*i.e.*, c) *latent vector* for keys and values, where $d_c \ll d_h n_h$. **Note that this is not a query vector.**
- $W^{DKV} \in \mathbb{R}^{d_c \times d}$ is the down-projection matrix
- $W^{UK}, W^{UV} \in \mathbb{R}^{d_h n_h \times d_c}$ are the up-projection matrices for keys and values, respectively.
- $W^{KR} \in \mathbb{R}^{d_h^R \times d}$ is the matrix used for generating the decoupled key of RoPE.
- During inference, **MLA only needs to cache \mathbf{c}_t^{KV} and \mathbf{k}_t^R , which is a compressed latent vector, rather than high-dimensional key and value vector.** Thus, it can still have the multiple heads. This alleviates the issue of GQA or MQA while keeping their benefits.
- Also, we do not have to explicitly compute the key and the values for attention.

$$\begin{aligned}
 q_t^T k_t &= (W^{UQ} c_t^Q)^T (W^{UK} c_t^{KV}) \\
 &= (c_t^Q)^T (W^{UQT} W^{UK}) c_t^{KV}
 \end{aligned}$$

- Here, $W^{UQT}W^{UK}$ is a combined matrix of the projection matrices.

Similarly, for values,

$$o_{t,i} = \text{AttnScore} \cdot v_t^C$$

The final output is

$$\begin{aligned} u_t &= W^O[o_{t,1}, \dots, o_{t,n_h}] \\ &= W^O[\text{AttnScore} \cdot (W^{UV}c_t^{KV})] \\ &= W^O W^{UV}[\text{AttnScore} \cdot (c_t^{KV})] \end{aligned}$$

- To leverage this advantage, we need to use the *decoupled RoPE*. Otherwise, the RoPE would be coupled with projection matrices.

$$\mathbf{q}_t^T \mathbf{k}_t = (\mathbf{c}_t^Q)^T ((W^{UQ})^T W^{UK}) \mathbf{c}_T^{KV}$$

Moreover, in order to reduce the activation memory during training, we also perform low-rank compression for the queries, even if it cannot reduce the KV cache:

$$\begin{aligned} \mathbf{c}_t^Q &= W^{DQ} \mathbf{h}_t \\ [\mathbf{q}_{t,1}^C; \mathbf{q}_{t,2}^C; \dots; \mathbf{q}_{t,n_h}^C] &= \mathbf{q}_t^C = W^{UQ} \mathbf{c}_t^Q \\ [\mathbf{q}_{t,1}^R; \mathbf{q}_{t,2}^R; \dots; \mathbf{q}_{t,n_h}^R] &= \mathbf{q}_t^R = \text{RoPE}(W^{QR} \mathbf{c}_t^Q) \\ \mathbf{q}_{t,i} &= [\mathbf{q}_{t,i}^C; \mathbf{q}_{t,i}^R] \end{aligned}$$

- $\mathbf{c}_t^Q \in \mathbb{R}^{d'_c}$ is the compressed latent vector for queries, where $d'_c \ll d_h n_h$
- $W^{DQ} \in \mathbb{R}^{d'_c \times d}$ and $W^{UQ} \in \mathbb{R}^{d_h n_h \times d'_c}$ are the down- and up- projection matrices for queries, respectively.
- $W^{QR} \in \mathbb{R}^{d_h n_h \times d'_c}$ is the matrix for decoupled queries of RoPE.

Finally, the attention queries ($\mathbf{q}_{t,i}$), keys ($\mathbf{k}_{j,i}$), and values ($\mathbf{v}_{j,i}^C$) are combined to yield the final attention output \mathbf{u}_t :

$$\begin{aligned} \mathbf{o}_{t,i} &= \sum_{j=1}^t \text{Softmax}_j \left(\frac{\mathbf{q}_{t,i}^T \mathbf{k}_{j,i}}{\sqrt{d_h + d_h^R}} \right) \mathbf{v}_{j,i}^C, \\ \mathbf{u}_t &= W^O[\mathbf{o}_{t,1}; \mathbf{o}_{t,2}; \dots; \mathbf{o}_{t,n_h}], \end{aligned}$$

where $W^O \in \mathbb{R}^{d \times d_h n_h}$ is the output projection matrix.

33.2 DeepSeek MoE

This paper mentioned that the conventional TopK MoE has Knowledge Hybridity and Knowledge Redundancy. Knowledge Hybridity: existing MoE practices often employ a limited number of experts (e.g., 8 or 16), and thus tokens assigned to a specific expert will be likely to cover diverse knowledge. Consequently, the designated expert will intend to assemble vastly different

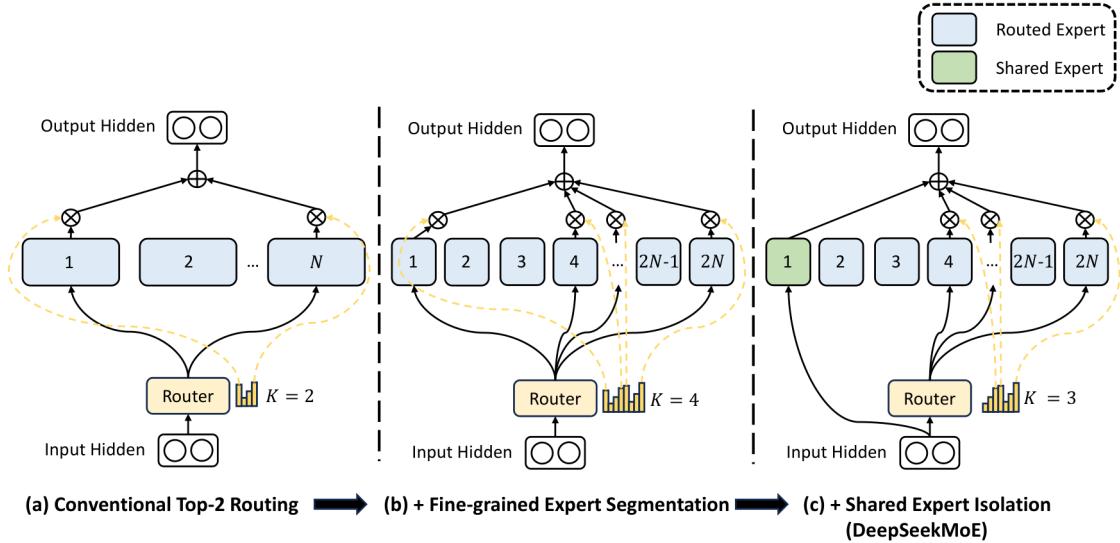


Figure 33.2: An overview of DeepSeek MoE.

types of knowledge in its parameters, which are hard to utilize simultaneously. (2) Knowledge Redundancy: tokens assigned to different experts may require common knowledge. As a result, multiple experts may converge in acquiring shared knowledge in their respective parameters, thereby leading to redundancy in expert parameters. These issues collectively hinder the expert specialization in existing MoE practices, preventing them from reaching the theoretical upper-bound performance of MoE models. By finely segmenting to more experts and introducing shared experts, DeepSeekMoE mitigated above two issues.

A typical practice to construct an MoE language model usually substitutes FFNs in a Transformer with MoE layers at specified intervals. An MoE layer is composed of multiple experts, where each expert is structurally identical to a standard FFN. Then, each token will be assigned to one or two experts. If the t -th FFN is substituted with an MoE layer, the computation for its output hidden state \mathbf{h}'_t is expressed as:

$$\mathbf{h}'_t = \mathbf{u}_t + \sum_{i=1}^{N_s} FFN_i^{(s)}(\mathbf{u}_t) + \sum_{i=1}^{N_r} g_{i,t} FFN_i^{(r)}(\mathbf{u}_t),$$

$$g_{i,t} = \frac{g'_{i,t}}{\sum_{j=1}^{N_r} g'_{j,t}},$$

where

$$g'_{i,t} = \begin{cases} s_{i,t} & s_{i,t} \in TopK(\{s_{j,t} | 1 \leq j \leq N_r\}, K_r) \\ 0 & \text{otherwise,} \end{cases}$$

$$s_{i,t} = \sigma(\mathbf{u}_t^T \mathbf{e}_i)$$

- \mathbf{u}_t : FFN input of the t -th token.
- \mathbf{h}'_t FFN output
- N_s and N_r are the number of *shared* experts and the *routed* experts, respectively.

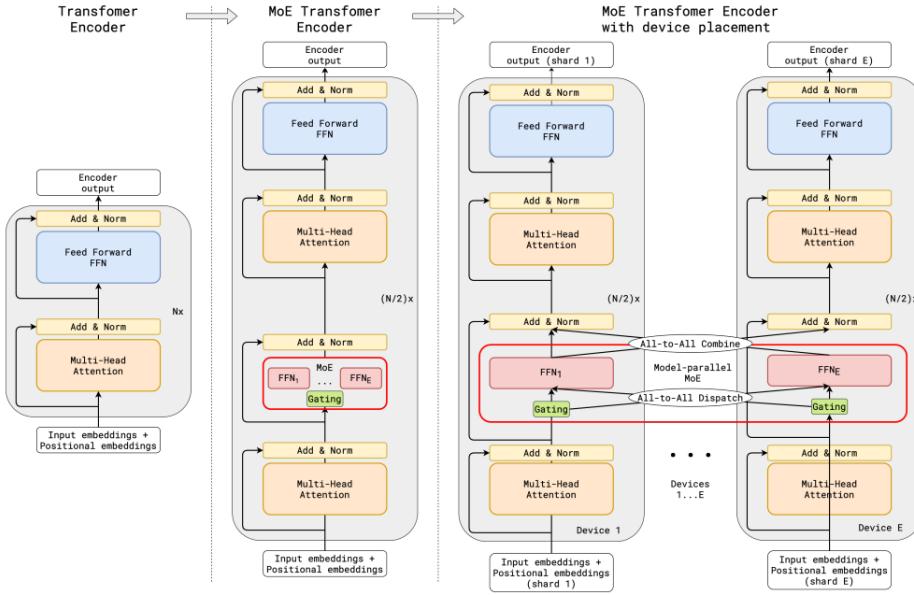


Figure 33.3: MoE Transformer Encoder from the GShard Paper

- $FFN_i^{(s)}(\cdot)$ and $FFN_i^{(r)}(\cdot)$ are the i -th *shared* expert and the i -th *routed* expert, respectively.
 - The shared experts are always activated, aiming at capturing and consolidating common knowledge across varying contexts.
 - Through compressing common knowledge into these shared experts, redundancy among other routed experts will be mitigated.
- K_r is the number of activated routed experts
- $g_{i,t}$: gating value for the i -th expert, which is sparse by its nature.
- $s_{i,t}$: a token assigned to an expert (*i.e.*, affinity score)
- \mathbf{e}_i : the centroid vector of the i -th (routed) expert
 - This is an (learnable) expert embedding representing assignment scores for each token claimed by each expert. The shared expert would get a vector like $\mathbf{1}$, which is a vector of ones.
- $TopK(\cdot, K)$: the set comprising K highest scores among the affinity scores calculated for the t -th token and all routed experts.
- σ : Sigmoid function.

Let's closely look at the equation

$$\mathbf{h}'_t = \mathbf{u}_t + \sum_{i=1}^{N_s} FFN_i^{(s)}(\mathbf{u}_t) + \sum_{i=1}^{N_r} g_{i,t} FFN_i^{(r)}(\mathbf{u}_t),$$

- $\mathbf{u}_t + \dots$: this is just a skip connection
- $\sum_{i=1}^{N_s} FFN_i^{(s)}(\mathbf{u}_t)$: Always activated

- $\sum_{i=1}^{N_r} g_{i,t} \text{FFN}_i^{(r)}(\mathbf{u}_t)$
 - Here, we can notice that $g_{i,t}$ is the weight for the corresponding (routed) experts

To balance the load of the experts, DeepSeek also introduces a bias term b_i for each expert and add it to the corresponding scores $s_{i,t}$ as follows:

$$g'_{i,t} = \begin{cases} s_{i,t} & s_{i,t} + b_i \in \text{TopK}(\{s_{j,t} + b_i | 1 \leq j \leq N_r\}, K_r) \\ 0 & \text{otherwise,} \end{cases}$$

During training, the bias term decreases by γ at the end of each step if its corresponding expert is overloaded.

33.3 Multi-Token Prediction

DeepSeek also adopts a multi-token prediction approach and they show that it can improve the output quality and generalization behaviors. One of potential reasons would be that MTP may mitigate the distributional discrepancy between training time teacher forcing and inference time autoregressive generation.

Unlike the conventional MTP approach, they keep the complete causal chain at each prediction depth by introducing MTP modules.

The input tokens $[t_1, t_2, t_3, t_4]$ go through the main model's transformer blocks and then go through the output head of main model to produce next predicted token t_5 . Meanwhile the representation of the input tokens $[t_1, t_2, t_3, t_4]$ (*i.e.*, output of main model's transformer blocks) will be passed to the MTP module and combine them with new input tokens' embedding $[t_2, t_3, t_4, t_5]$ to help produce additional token t_6 . In DeepSeek-V3, the model is designed to predict next 2 tokens.

$$\mathbf{h}'^k_i = M_k[\text{RMSNorm}(\mathbf{h}_i^{k-1}); \text{RMSNorm}(\text{Emb}(t_{i+k}))]$$

- \mathbf{h}_i^{k-1} : i -th token at the $(k-1)$ -th depth
- $(i+k)$ -th token embedding $\text{Emb}(t_{i+k})$

33.4 Reinforcement Learning

33.4.1 Backgrounds: Proximal Policy Optimization

The objective function of Proximal Policy Optimization (PPO) can be represented as follows:

$$\theta_{k+1} = \underset{\theta}{\operatorname{argmax}} \mathbb{E}_{s,a \sim \pi_{\theta_k}} [L(s, a, \theta_k, \theta)],$$

where θ_k is a parameter of a policy network at k -th step, θ is the current policy we want to update, and the A is the advantage (*i.e.*, reward). Finally, the L is given by

$$L(s, a, \theta_k, \theta) = \min \left(\frac{\pi_{\theta}(a|s)}{\pi_{\theta_k}(a|s)} A^{\pi_{\theta_k}}(s, a), \text{Clip} \left(\frac{\pi_{\theta}(a|s)}{\pi_{\theta_k}(a|s)}, 1 - \varepsilon, 1 + \varepsilon \right) A^{\pi_{\theta_k}}(s, a) \right).$$

Roughly, ε is a hyperparameter which says how far away the new policy is allowed to go from the old one. A simpler expression of the above expression is

$$L(s, a, \theta_k, \theta) = \min \left(\frac{\pi_\theta(a|s)}{\pi_{\theta_k}(a|s)} A^{\pi_{\theta_k}}(s, a), g(\varepsilon, A^{\pi_{\theta_k}}(s, a)) \right), \quad (33.1)$$

where

$$g(\varepsilon, A) = \begin{cases} (1 + \varepsilon)A & A \geq 0 \\ (1 - \varepsilon)A & A < 0. \end{cases} \quad (33.2)$$

1. Positive Advantage: Suppose the advantage for that state-action pair is positive, in which case its contribution to the objective reduces to

$$L(s, a, \theta_k, \theta) = \min \left(\frac{\pi_\theta(a|s)}{\pi_{\theta_k}(a|s)}, 1 + \varepsilon \right) A^{\pi_{\theta_k}}(s, a). \quad (33.3)$$

As the advantage is positive, the objective will increase if the action becomes more likely that is, if $\pi_\theta(a|s)$ increases. But the min in this term puts a limit to how much the objective can increase. Once $\pi_\theta(a|s) > (1 + \varepsilon)\pi_{\theta_k}(a|s)$, the min kicks in and this term hits a ceiling of $(1 + \varepsilon)A^{\pi_{\theta_k}}(s, a)$. Thus, the new policy does not benefit by going far away from the old policy.

2. Negative Advantage: Suppose the advantage for that state-action pair is negative, in which case its contribution to the objective reduces to

$$L(s, a, \theta_k, \theta) = \max \left(\frac{\pi_\theta(a|s)}{\pi_{\theta_k}(a|s)}, 1 - \varepsilon \right) A^{\pi_{\theta_k}}(s, a). \quad (33.4)$$

Since the advantage is negative, the objective will increase if the action becomes less likely. In other words, if $\pi_\theta(a|s)$ decreases. But the max in this term puts a limit to how much the objective can increase. Once $\pi_\theta(a|s) < (1 - \varepsilon)\pi_{\theta_k}(a|s)$, the max kicks in and this term hits a ceiling of $(1 - \varepsilon)A^{\pi_{\theta_k}}(s, a)$. Thus, again, the new policy does not benefit by going far away from the old policy.

In sum, **clipping serves as a regularizer** by restricting the rewards to the policy, which change it dramatically with the hyperparameter ε corresponds to how far away the new policy can go from the old while still profiting the objective.

33.4.2 Group Relative Policy Optimization

Similarly, in DeepSeek, they leveraged a RL objective called Group Relative Policy Optimization (GRPO), which is a variant of PPO.

$$\mathcal{J} = \frac{1}{G} \sum_{i=1}^G \min \left(\frac{\pi_\theta(o_i|q)}{\pi_{\theta_k}(o_i|q)} A_i, \text{Clip} \left(\frac{\pi_\theta(o_i|q)}{\pi_{\theta_k}(o_i|q)}, 1 - \varepsilon, 1 + \varepsilon \right) A_i \right) - \beta D_{KL}(\pi_\theta \| \pi_{\text{ref}}).$$

- GRPO samples a group of outputs $\{o_1, o_2, \dots, o_G\}$ from the old policy π_{θ_k} and then optimizes the policy model by maximizing the objective

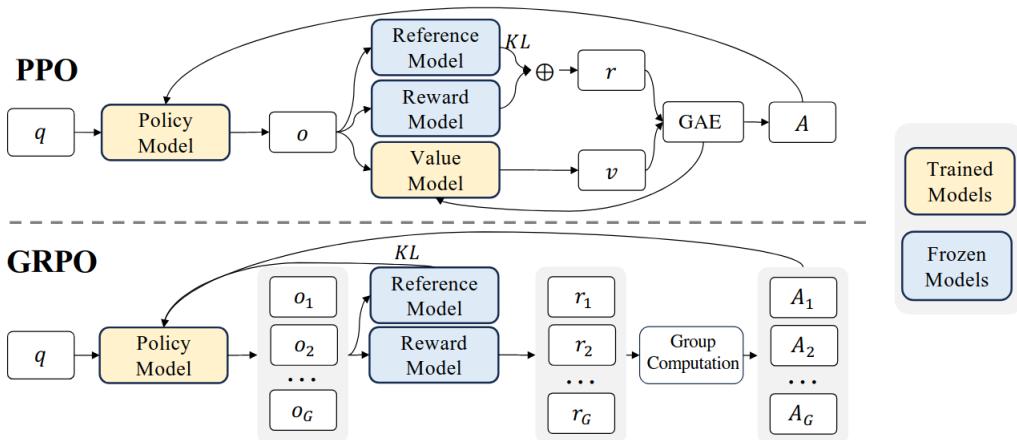


Figure 33.4: An illustration of GRPO. The policy model π takes a question (*i.e.*, observation, q) and generates answers (*i.e.*, actions, o_1) of the question.

- The KL-divergence is used to restrict the sudden change of policy.
- The advantage can be calculated by averaging and normalizing the rewards.
- Instead of using a value model explicitly, GRPO computes a value of the states (*i.e.*, o_i) by averaging them.
 - $A(s, a) = Q(s, a) - V(s)$
- π_{ref} is the reference model, which is an initial SFT model.

33.5 DeepSeek-R1

- DeepSeek-R1-Zero, which applies RL directly to the base model without any SFT data,
- DeepSeek-R1, which applies RL starting from a checkpoint fine-tuned with thousands of long Chain-of-Thought (CoT) examples.
- Distill the reasoning capability from DeepSeek-R1 to small dense models.

33.5.1 DeepSeek-R1-Zero

The reward is the source of the training signal, which decides the optimization direction of RL. To train DeepSeek-R1-Zero, we adopt a rule-based reward system that mainly consists of two types of rewards:

- Accuracy rewards: The accuracy reward model evaluates whether the response is correct. For example, in the case of math problems with deterministic results, the model is required to provide the final answer in a specified format (*e.g.*, within a box), enabling reliable rule-based verification of correctness. Similarly, for LeetCode problems, a compiler can be used to generate feedback based on predefined test cases.

Question: If $a > 1$, then the sum of the real solutions of $\sqrt{a - \sqrt{a+x}} = x$ is equal to

Response: <think>

To solve the equation $\sqrt{a - \sqrt{a+x}} = x$, let's start by squaring both ...

$$(\sqrt{a - \sqrt{a+x}})^2 = x^2 \implies a - \sqrt{a+x} = x^2.$$

Rearrange to isolate the inner square root term:

$$(a - x^2)^2 = a + x \implies a^2 - 2ax^2 + (x^2)^2 = a + x \implies x^4 - 2ax^2 - x + (a^2 - a) = 0$$

...

Wait, wait. Wait. That's an aha moment I can flag here.

Let's reevaluate this step-by-step to identify if the correct sum can be ...

We started with the equation:

$$\sqrt{a - \sqrt{a+x}} = x$$

First, let's square both sides:

$$a - \sqrt{a+x} = x^2 \implies \sqrt{a+x} = a - x^2$$

Next, I could square both sides again, treating the equation: ...

...

- Format rewards: In addition to the accuracy reward model, we employ a format reward model that enforces the model to put its thinking process between <think> and </think> tags.

Note that *R1 does not use a neural reward model*.

Aha Moment of DeepSeek-R1-Zero A particularly intriguing phenomenon observed during the training of DeepSeek-R1-Zero is the occurrence of an aha moment. This moment, as illustrated in Table 3, occurs in an intermediate version of the model. During this phase, **DeepSeek-R1-Zero learns to allocate more thinking time to a problem by reevaluating its initial approach.** This behavior is not only a testament to the model's growing reasoning abilities but also a captivating example of how reinforcement learning can lead to unexpected and sophisticated outcomes.

This moment is not only an aha moment for the model but also for the researchers observing its behavior. It underscores the power and beauty of reinforcement learning: rather than explicitly teaching the model on how to solve a problem, we simply provide it with the right incentives, and it autonomously develops advanced problem-solving strategies. The aha moment serves as a powerful reminder of the potential of RL to unlock new levels of intelligence in artificial systems, paving the way for more autonomous and adaptive models in the future.

33.5.2 DeepSeek-R1

We construct and collect a small amount of long CoT data (thousands) to fine-tune the model as the initial RL actor.

- A key limitation of DeepSeek-R1-Zero is that its content is often not suitable for reading
- In contrast, when creating cold-start data for DeepSeek-R1, we design a readable pattern that includes a summary at the end of each response and filters out responses that are not reader-friendly. Here, we define the output format as |special-token|<reasoning-process>|special-token|<summary>, where the reasoning process is the CoT for the query, and the summary is used to summarize the reasoning results.

When reasoning-oriented RL converges, we utilize the resulting checkpoint to collect SFT (Supervised Fine-Tuning) data for the subsequent round. Unlike the initial cold-start data, which

primarily focuses on reasoning, this stage incorporates data from other domains to enhance the model’s capabilities in writing, role-playing, and other general-purpose tasks. Specifically, we generate the data and fine-tune the model as described below.

We curate reasoning prompts and generate reasoning trajectories by performing rejection sampling from the checkpoint from the above RL training. In the previous stage, we only included data that could be evaluated using rule-based rewards. However, in this stage, we expand the dataset by incorporating additional data, some of which use a generative reward model by feeding the ground-truth and model predictions into DeepSeek-V3 for judgment. Additionally, because the model output is sometimes chaotic and difficult to read, we have filtered out chain-of-thought with mixed languages, long paragraphs, and code blocks. For each prompt, we sample multiple responses and retain only the correct ones. In total, we collect about 600k reasoning related training samples.

To equip more efficient smaller models with reasoning capabilities like DeepSeek-R1, we directly fine-tuned open-source models like Qwen (Qwen, 2024b) and Llama (AI@Meta, 2024) using the 800k samples curated with DeepSeek-R1.

Part VII

Appendix

Chapter 34

Vector Calculus

1.1 Differentiate

1.1.1 Differentiation Rules

- Product rule:

$$(f(x)g(x))' = f'(x)g(x) + f(x)g'(x)$$

- Quotient rule:

$$\left(\frac{f(x)}{g(x)}\right)' = \frac{f'(x)g(x) + f(x)g'(x)}{(g(x))^2}$$

- Sum rule:

$$(f(x) + g(x))' = f'(x) + g'(x)$$

- Chain rule:

$$(g(f(x)))' = g'(f(x))f'(x)$$

The generalization of the derivative to functions of several variables is the *gradient*. We find the gradient of the function f with respect to x by varying one variable at a time and keeping the others constant. **The gradient is then the collection of these partial derivatives.**

For example, partial derivatives using the chain rule of $f(x, y) = (x + 2y^3)^2$ is given by

$$\frac{\partial f(x, y)}{\partial x} = 2(x + 2y^3) \frac{\partial}{\partial x}(x + 2y^3) = 2(x + 2y^3)$$

Basic rules of Partial Differentiation:

- Product rule:

$$\frac{\partial}{\partial \mathbf{x}} [f(\mathbf{x})g(\mathbf{x})] = \frac{\partial f}{\partial \mathbf{x}}g(\mathbf{x}) + f(\mathbf{x})\frac{\partial g}{\partial \mathbf{x}}$$

- Sum rule:

$$\frac{\partial}{\partial \mathbf{x}} [f(\mathbf{x}) + g(\mathbf{x})] = \frac{\partial f}{\partial \mathbf{x}} + \frac{\partial g}{\partial \mathbf{x}}$$

- Chain rule:

$$\frac{\partial}{\partial \mathbf{x}} (g \circ f)(\mathbf{x}) = \frac{\partial}{\partial \mathbf{x}} [g(f(\mathbf{x}))] = \frac{\partial g}{\partial f} \frac{\partial f}{\partial \mathbf{x}}$$

2.2 Chain Rule

Consider a function $f : \mathbb{R}^2 \rightarrow \mathbb{R}$ of two variables x_1 and x_2 . They are functions of t , $x_1(t)$ and $x_2(t)$. To compute the gradient of f with respect to t , we need to apply the chain rule for multivariate functions as

$$\frac{\partial f}{\partial t} = \begin{bmatrix} \frac{\partial f}{\partial x_1} & \frac{\partial f}{\partial x_2} \end{bmatrix} \begin{bmatrix} \frac{\partial x_1(t)}{\partial t} \\ \frac{\partial x_2(t)}{\partial t} \end{bmatrix} = \frac{\partial f}{\partial x_1} \frac{\partial x_1(t)}{\partial t} + \frac{\partial f}{\partial x_2} \frac{\partial x_2(t)}{\partial t}$$

Given that $f(x_1, x_2)$ is a function of x_1 and x_2 , where $x_1 = x_1(s, t)$ and $x_2 = x_2(s, t)$ are themselves functions of two variables s and t , the chain rule can be used to find the partial derivatives of f with respect to s and t .

$$\frac{\partial f}{\partial s} = \frac{\partial f}{\partial x_1} \frac{\partial x_1}{\partial s} + \frac{\partial f}{\partial x_2} \frac{\partial x_2}{\partial s}$$

$$\frac{\partial f}{\partial t} = \frac{\partial f}{\partial x_1} \frac{\partial x_1}{\partial t} + \frac{\partial f}{\partial x_2} \frac{\partial x_2}{\partial t}$$

The gradient of f is obtained by the matrix multiplication as follows:

$$\frac{df}{d(s, t)} = \frac{\partial f}{\partial x} \frac{\partial x}{\partial (s, t)} = \begin{pmatrix} \frac{\partial f}{\partial x_1} & \frac{\partial f}{\partial x_2} \end{pmatrix} \begin{pmatrix} \frac{\partial x_1}{\partial s} & \frac{\partial x_1}{\partial t} \\ \frac{\partial x_2}{\partial s} & \frac{\partial x_2}{\partial t} \end{pmatrix}$$

3.3 Vector Notations

- $\mathbf{x} = (x_1, x_2, \dots, x_n)$ or

-

$$\mathbf{x} = \begin{bmatrix} x_1 \\ \vdots \\ x_n \end{bmatrix}$$

Differentiate a vector y by a scalar \mathbf{x} :

$$\frac{\partial \mathbf{y}}{\partial x} = \begin{bmatrix} \frac{\partial y_1}{\partial x} \\ \vdots \\ \frac{\partial y_n}{\partial x} \end{bmatrix}$$

Differentiate a scalar y by a vector \mathbf{x} :

$$\frac{\partial y}{\partial \mathbf{x}} = \left[\frac{\partial y}{\partial x_1}, \dots, \frac{\partial y}{\partial x_n} \right]$$

Differentiate a vector y by a vector \mathbf{x} :

$$\frac{\partial \mathbf{y}}{\partial \mathbf{x}} = \begin{bmatrix} \frac{\partial y_1}{\partial x_1} & \cdots & \frac{\partial y_1}{\partial x_n} \\ \vdots & \ddots & \vdots \\ \frac{\partial y_n}{\partial x_1} & \cdots & \frac{\partial y_n}{\partial x_n} \end{bmatrix}$$

$\mathbf{a}^T \mathbf{x}$ is a scalar value, so

$$\begin{aligned} \frac{\partial \mathbf{a}^T \mathbf{x}}{\partial \mathbf{x}} &= \left[\frac{\partial(\mathbf{a}^T \mathbf{x})}{\partial x_1} \cdots \frac{\partial(\mathbf{a}^T \mathbf{x})}{\partial x_n} \right] = \left[\frac{\partial(a_1 x_1 + \cdots + a_n x_n)}{\partial x_1}, \dots, \frac{\partial(a_1 x_1 + \cdots + a_n x_n)}{\partial x_n} \right] \\ &= [a_1, \dots, a_n] = \mathbf{a}^T \end{aligned}$$

$$A\mathbf{x} = \begin{bmatrix} a_{11} & \cdots & a_{1n} \\ \vdots & \ddots & \vdots \\ a_{m1} & \cdots & a_{mn} \end{bmatrix} \times \begin{bmatrix} x_1 \\ \vdots \\ x_n \end{bmatrix} = \begin{bmatrix} \sum_{i=1}^n a_{1i} x_i \\ \vdots \\ \sum_{i=1}^n a_{mi} x_i \end{bmatrix}$$

Thus,

$$\frac{\partial A\mathbf{x}}{\partial \mathbf{x}} = \begin{bmatrix} \frac{\partial \sum_{i=1}^n a_{1i} x_i}{\partial x_1} & \cdots & \frac{\partial \sum_{i=1}^n a_{1i} x_i}{\partial x_n} \\ \vdots & \ddots & \vdots \\ \frac{\partial \sum_{i=1}^n a_{mi} x_i}{\partial x_1} & \cdots & \frac{\partial \sum_{i=1}^n a_{mi} x_i}{\partial x_n} \end{bmatrix} = A$$

Chapter 35

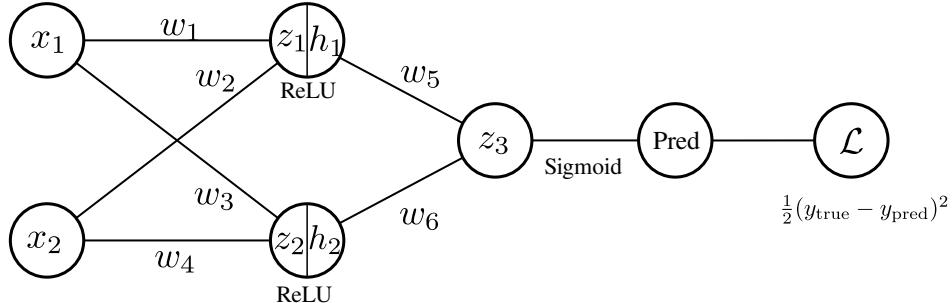
Backpropagation

1.1 Introduction

An area where the chain rule is used extensively is deep learning, where the function value y is computed as a many-level function composition:

$$y = (f_K \circ f_{K-1} \circ \cdots \circ f_1)(x) = f_K(f_{K-1}(\cdots(f_1(x))\cdots)),$$

where x are the inputs (e.g., images), y are the observations (e.g., class labels), and every function f_i , $i = 1, \dots, K$, possesses its own parameters.



We have:

- Inputs: x_1, x_2 .
- Hidden layer: 2 neurons, each with no bias (for simplicity).
- Weights $w_1, w_2, w_3, w_4, w_5, w_6$
- Output layer: 1 neuron, also with no bias.

Where:

- $z_1 = w_1 x_1 + w_2 x_2, \quad h_1 = \text{ReLU}(z_1)$
- $z_2 = w_3 x_1 + w_4 x_2, \quad h_2 = \text{ReLU}(z_2)$
- $z_3 = w_5 h_1 + w_6 h_2, \quad y_{\text{pred}} = \sigma(z_3)$

$$\sigma(z) = \frac{1}{1 + e^{-z}}.$$

We will train on one example $\{x = [x_1, x_2], y_{\text{true}}\}$.

Loss function: Mean Squared Error (MSE) for one sample:

$$L = \frac{1}{2} (y_{\text{pred}} - y_{\text{true}})^2.$$

We want to find the partial derivatives:

$$\frac{\partial L}{\partial w_1}, \frac{\partial L}{\partial w_2}, \frac{\partial L}{\partial w_3}, \frac{\partial L}{\partial w_4}, \frac{\partial L}{\partial w_5}, \frac{\partial L}{\partial w_6}.$$

Useful Derivatives:

- Derivative of the MSE loss (single sample):

$$\frac{\partial L}{\partial y_{\text{pred}}} = (y_{\text{pred}} - y_{\text{true}}).$$

- Sigmoid derivative:

$$\frac{d}{dz} \sigma(z) = \sigma(z) (1 - \sigma(z)).$$

So,

$$\frac{\partial y_{\text{pred}}}{\partial z_3} = y_{\text{pred}} (1 - y_{\text{pred}}).$$

- Chain rule:

$$\frac{\partial L}{\partial z_3} = \frac{\partial L}{\partial y_{\text{pred}}} \times \frac{\partial y_{\text{pred}}}{\partial z_3}.$$

The output node has:

$$z_3 = w_5 h_1 + w_6 h_2, \quad y_{\text{pred}} = \sigma(z_3).$$

First, find $\frac{\partial L}{\partial z_3}$:

$$\frac{\partial L}{\partial z_3} = (y_{\text{pred}} - y_{\text{true}}) \times [y_{\text{pred}} (1 - y_{\text{pred}})].$$

Then,

$$\frac{\partial L}{\partial w_5} = \frac{\partial L}{\partial z_3} \times \frac{\partial z_3}{\partial w_5} = \frac{\partial L}{\partial z_3} \times h_1.$$

Since $z_3 = w_5 h_1 + w_6 h_2$, so $\partial z_3 / \partial w_5 = h_1$.

Each hidden neuron influences the output via h_1 or h_2 . Let's do them one neuron at a time.

4.3.1. Hidden Neuron1 (w_1, w_2)

$$z_1 = w_1 x_1 + w_2 x_2, \quad h_1 = \sigma(z_1).$$

To get $\frac{\partial L}{\partial w_1}$, we chain through:

$$\frac{\partial L}{\partial w_1} = \frac{\partial L}{\partial z_3} \times \frac{\partial z_3}{\partial h_1} \times \frac{\partial h_1}{\partial z_1} \times \frac{\partial z_1}{\partial w_1}.$$

- We already have $\frac{\partial L}{\partial z_3} = -0.12058$. - $\frac{\partial z_3}{\partial h_1} = w_5$ (since $z_3 = w_5 h_1 + w_6 h_2$). - $\frac{\partial h_1}{\partial z_1} = \sigma(z_1)(1 - \sigma(z_1)) = h_1(1 - h_1)$. - $\frac{\partial z_1}{\partial w_1} = x_1$.

Numerically:

1. $\frac{\partial z_3}{\partial h_1} = w_5 = 0.2$.
2. $h_1(1 - h_1) = 0.62246 \times (1 - 0.62246) = 0.62246 \times 0.37754 \approx 0.2350$.
3. $x_1 = 1.0$.

So,

$$\frac{\partial L}{\partial w_1} = (-0.12058) \times (0.2) \times (0.2350) \times (1.0).$$

Compute step by step:

$$- (-0.12058) \times 0.2 = -0.024116. \quad - (-0.024116) \times 0.2350 \approx -0.005667.$$

Thus,

$$\frac{\partial L}{\partial w_1} \approx -0.005667.$$

For w_2 , the only difference is $\frac{\partial z_1}{\partial w_2} = x_2$. Since $x_2 = 0.0$ in our example, it follows immediately:

$$\frac{\partial L}{\partial w_2} = 0.$$

4.3.2. Hidden Neuron2 (w_3, w_4)

$$z_2 = w_3 x_1 + w_4 x_2, \quad h_2 = \sigma(z_2).$$

By the same chain rule:

$$\frac{\partial L}{\partial w_3} = \frac{\partial L}{\partial z_3} \times \frac{\partial z_3}{\partial h_2} \times \frac{\partial h_2}{\partial z_2} \times \frac{\partial z_2}{\partial w_3}.$$

- We have $\frac{\partial L}{\partial z_3} = -0.12058$. - $\frac{\partial z_3}{\partial h_2} = w_6 = -0.1$. - $\frac{\partial h_2}{\partial z_2} = h_2(1 - h_2) = 0.57444 \times 0.42556 \approx 0.24450$.
- $\frac{\partial z_2}{\partial w_3} = x_1 = 1.0$.

So,

$$\frac{\partial L}{\partial w_3} = (-0.12058) \times (-0.1) \times (0.24450) \times (1.0).$$

Compute step by step:

$$1. (-0.12058) \times (-0.1) = 0.012058. \quad 2. 0.012058 \times 0.24450 \approx 0.002950.$$

Hence,

$$\frac{\partial L}{\partial w_3} \approx 0.002950.$$

For w_4 , again $\partial z_2 / \partial w_4 = x_2 = 0$, so:

$$\frac{\partial L}{\partial w_4} = 0.$$

—

5. Weight Update

Let's do **one step** of Gradient Descent with learning rate $\eta = 0.1$. The update rule is:

$$w_i \leftarrow w_i - \eta \frac{\partial L}{\partial w_i}.$$

5.1. Final Gradients Recap

$$-\frac{\partial L}{\partial w_1} \approx -0.005667 - \frac{\partial L}{\partial w_2} = 0 - \frac{\partial L}{\partial w_3} \approx 0.002950 - \frac{\partial L}{\partial w_4} = 0 - \frac{\partial L}{\partial w_5} \approx -0.07509 - \frac{\partial L}{\partial w_6} \approx -0.06930$$

5.2. Updated Weights

1. w_1 :

$$w_1 \leftarrow 0.5 - 0.1 \times (-0.005667) = 0.5 + 0.0005667 \approx 0.50057.$$

2. w_2 :

$$w_2 \leftarrow -0.5 - 0.1 \times 0 = -0.5.$$

3. w_3 :

$$w_3 \leftarrow 0.3 - 0.1 \times (0.002950) = 0.3 - 0.000295 \approx 0.299705.$$

4. w_4 :

$$w_4 \leftarrow 0.1 - 0.1 \times 0 = 0.1.$$

5. w_5 :

$$w_5 \leftarrow 0.2 - 0.1 \times (-0.07509) = 0.2 + 0.007509 \approx 0.20751.$$

6. w_6 :

$$w_6 \leftarrow -0.1 - 0.1 \times (-0.06930) = -0.1 + 0.00693 \approx -0.09307.$$

Recap

1. **Forward pass**: - Calculated hidden neurons z_1, z_2 , then h_1, h_2 . - Calculated output z_3 , then y_{pred} . - Computed the MSE loss L .
2. **Backward pass**: - Used the chain rule to find $\frac{\partial L}{\partial w_i}$ for each weight. - Observed how zero input ($x_2 = 0$) caused some gradients to be zero for this training example.
3. **Gradient Descent update**: - Applied $w_i \leftarrow w_i - \eta (\partial L / \partial w_i)$ with $\eta = 0.1$.

While we only did one training example and one update step, the principle is the same for multiple samples (you would sum or average the gradients across the batch). Larger networks just repeat the same chain-rule logic on more layers and neurons.

That's it! This is the backpropagation process laid out in a small, fully traceable network with two hidden neurons. By doing the arithmetic yourself, you can see exactly how each weight nudges the output and how the gradients flow backward to update those weights.

Chapter 36

Divergence

1.1 KL Divergence between Two Normal Distribution

$$D_{\text{KL}}(P||Q) = \mathbb{E}_P \left[\log \frac{P}{Q} \right]$$

Consider two multivariate Gaussians in \mathbb{R}^n , P_1 and P_2

$$\begin{aligned} D_{\text{KL}}(P||Q) &= \int \left[\frac{1}{2} \log \frac{|\Sigma_2|}{|\Sigma_1|} - \frac{1}{2}(x - \mu_1)^T \Sigma_1^{-1} (x - \mu_1) + \frac{1}{2}(x - \mu_2)^T \Sigma_2^{-1} (x - \mu_2) \right] \times p(x) dx \\ &= \frac{1}{2} \log \frac{|\Sigma_2|}{|\Sigma_1|} - \frac{1}{2} \text{tr} \{ E[(x - \mu_1)(x - \mu_1)^T] \Sigma_1^{-1} \} + \frac{1}{2} E[(x - \mu_2)^T \Sigma_2^{-1} (x - \mu_2)] \\ &= \frac{1}{2} \log \frac{|\Sigma_2|}{|\Sigma_1|} - \frac{1}{2} \text{tr} \{ I_n \} + \frac{1}{2} (\mu_1 - \mu_2)^T \Sigma_2^{-1} (\mu_1 - \mu_2) + \frac{1}{2} \text{tr} \{ \Sigma_2^{-1} \Sigma_1 \} \\ &= \frac{1}{2} \left[\log \frac{|\Sigma_2|}{|\Sigma_1|} - n + \text{tr} \{ \Sigma_2^{-1} \Sigma_1 \} + (\mu_2 - \mu_1)^T \Sigma_2^{-1} (\mu_2 - \mu_1) \right] \end{aligned}$$

Trace tricks:

$$\begin{aligned} x^T A x &= \text{tr}[x^T A x] = \text{tr}[x x^T A] \\ \text{tr}[A + B] &= \text{tr}[A] + \text{tr}[B] \\ E[(x - \mu)^T \Sigma^{-1} (x - \mu)] &= \text{tr}(E[(x - \mu)(x - \mu)^T] \Sigma^{-1}) \\ \Sigma &= E[(X - \mu)(X - \mu)^T] = E[XX^T] - \mu\mu^T \\ E[XX^T] &= \Sigma + \mu\mu^T \end{aligned}$$

Note that the determinant of a diagonal matrix could be computed as product of its diagonal.

2.2 Various Tricks

2.2.1 Spectral Normalization

A persisting challenge in the training of GANs is the performance control of the discriminator. The derivative of discriminator could be unbounded and even incomputable, so they introduced a regularization on the derivative of discriminator called, Lipchitz continuity, which bound the gradient.

Neural network is actually a composite function. So if we make each function to satisfy the Lipchitz continuity, then we can make whole network satisfy it. Lipchitz continuity of a linear operator can be seen as

$$\begin{aligned} \|f(x_1) - f(x_2)\|_2 &\leq L\|x_1 - x_2\|_2 \\ \|Ax_1 - Ax_2\|_2 &\leq L\|x_1 - x_2\|_2 \\ \frac{\|Ax\|_2}{\|x\|_2} &\leq L \\ \sigma_{max} \underbrace{\sup_x \frac{\|Ax\|_2}{\|x\|_2}}_{SpectralNorm} &\leq L, \quad \text{Since inequality holds for all } x \end{aligned}$$

, where σ_{max} is the maximum singular value. Note that the spectral norm is from the linear algebra. We can make the matrix A Lipchitz continuous by

$$1 = \sup_x \underbrace{\frac{\|\frac{A}{\sigma_{max}}x\|_2}{\|x\|_2}}_{SpectralNorm} \leq L$$

2.2.2 Moving Averaging

2.2.3 Weight Averaging

2.2.4 Quality Measurements

3.3 f-Divergence

In probability theory, an f -divergence is a function $D_f(p||q)$ that measures the difference between two probability distributions p and q . It helps the intuition to think of the divergence as an average, weighted by the function f , of the odds ratio given by p and q .

For distributions p and q , f -divergence is defined as:

$$D_f(p||q) = \int_{\mathcal{X}} f\left(\frac{p(x)}{q(x)}\right) q(x) dx$$

- KL-divergence: $f(t) = t \log t$

- Reversed KL-divergence: $f(t) = -\log t$
- Total variation: $f(t) = \frac{1}{2}|t - 1|$

$$D_f(p||q) = \int_{\mathcal{X}} |p(x) - q(x)| dx$$

4.4 Lipchitz Continuous

The function f in the new form of Wasserstein metric is demanded to satisfy $\|f\|_L \leq K$, meaning it should be K -Lipschitz continuous.

A real-valued function $f : \mathbb{R} \rightarrow \mathbb{R}$ is called K -Lipschitz continuous if there exists a real constant $K \geq 0$ such that, for all $x_1, x_2 \in \mathbb{R}$

$$|f(x_1) - f(x_2)| \leq K|x_1 - x_2|$$

5.5 Singular Value

All singular values can be calculated via the singular value decomposition (SVD)

$$A = U\Sigma V^T$$

, where U is the left singular vectors and V is the right singular vectors. However, if we just want the maximum singular value then we just need to find corresponding vectors

$$\sigma = uAv^T$$

Actually, there is a simpler way to find the maximum singular value e.g., power iteration.

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