#### MA50259: Statistical Design of Investigations

Dr. Sandipan Roy

Lecture 6: Blocking in Factorial Designs

#### Factorial Designs in Blocks

- ▶ Blocking is even more effective when combined with a factorial design in treatment factors
- ► The number of experimental units in each block must be equal to the product of levels of all the factors. This is called a randomized complete block factorial (RCBF)
- ► The model for analysis of a two-factor factorial in a randomized block design is given by

$$y_{ijk} = \mu + b_i + \alpha_j + \beta_k + (\alpha\beta)_{jk} + \epsilon_{ijk}$$

- Note there is an interaction between the two factorial factors  $(\alpha_j \text{ and } \beta_k)$
- ► There is no interactions between treatment factors and the block factors b<sub>i</sub> as there will be zero degrees of freedom for ssE

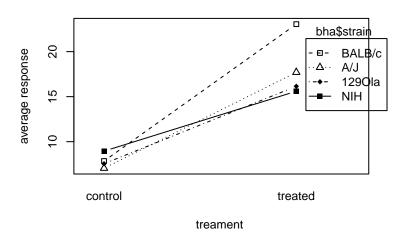
# Efficiency of factorial designs in Blocks

```
# data in RCBF data.frame
mod.blocks <- aov( response ~ block + factor1* factor2, data = RCBF)
summary(mod.blocks)
              Df Sum Sq Mean Sq F value
                                      Pr(>F)
                   47.6 47.6 18.372 0.00363 **
block
factor1
               3 33.0 11.0 4.240 0.05274 .
factor2
              1 422.3 422.3 162.961 4.19e-06 ***
factor1:factor2 3 40.3 13.4 5.189 0.03368 *
Residuals
            7 18.1 2.6
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
mod.no.blocks <- aov( response ~ factor1 * factor2, data = RCBF)
summary(mod.no.blocks)
              Df Sum Sq Mean Sq F value
                                      Pr(>F)
               3 33.0 11.0 1.337
                                       0.329
factor1
              1 422.3 422.3 51.383 9.54e-05 ***
factor2
factor1:factor2 3 40.3 13.4 1.636
                                         0.257
Residuals 8 65.8 8.2
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

- ▶ The estimated variance has been reduced from 8.2 to 2.6 so that 8.2/2.6 = 3.15
- ▶ It would take more than three times as many EUs to have the same power if the experiments were not blocked

# Factorial Designs in Blocks

```
interaction.plot(bha$treat, bha$strain, bha$y, type = "b", pch = c(18,24,22,15), leg.bty = "o",
main = "", xlab = "treament ",ylab = "average response")
```



- When experimental units (EUs) represent physical entities, smaller blocks of EUs usually result in greater homogeneity.
- ► The larger the block, the more likely it is to have many heterogeneous EUs. Then is not advised to have blocked designs with more the number of treatment levels or number of combination of levels of treatment factors
- However, sometimes the EUs represent trials rather than physical entities and the runs can be made quickly, then larger block sizes may not increase the variability of EUs within a block
- In that case, a design with replicates of each treatment level within a block, a generalized complete block design (GCB), can be used

Since there are replicate EUs for each treatment in each block, it is possible to fit the model with interactions between blocks and treatment factors. For an RCB we can use model:

$$y_{ijk} = \mu + b_i + \tau_j + (b\tau)_{ij} + \epsilon_{ijk}$$

- Interactions are meaninful if block factors are of interest!
- Problem: If block factors are not interest and the interaction between blocks and treatment factor is significant, its interpretation would imply that treatment effect is different for different blocks!
- ▶ In that case we need to use a **special test statistic** for the null hypothesis of no treatment effect  $H_0$ :  $\tau_i = 0$  for all i
- ▶ the test should use the block by treatment interaction mean square as the denominator of the F-test!

```
# data in GCB data.frame
inter<-interaction(GCB$block,GCB$treatment,sep=":")
X1<-model.matrix(~block-1.GCB):X2<-model.matrix(~treatment-1.GCB):X3<-model.matrix(~inter-1)
X<-cbind(1,X1,X2,X3); X0<-cbind(1,X1,X2); X00<-cbind(1,X1)
G < ginv(t(X)%*%X): GO < ginv(t(XO)%*%XO): GOO < ginv(t(XOO)%*%XOO)
v<-GCB$response; n<-length(v)</pre>
resid<-v-X%*%G%*%t(X)%*%v: ssE<-t(resid)%*%resid %>% as.vector:
ssE
[1] 7340.752
resid0<-y-X0%*%G0%*%t(X0)%*%y;ssE0<-t(resid0)%*%resid0 %>% as.vector;
ssE0-ssE
[1] 2356.1
resid00<-v-X00%*%G00%*%t(X00)%*%v: ssE00<-t(resid00)%*%resid00 %>% as.vector:
ssE00-ssE0
[1] 1723.932
Fval<-((ssE00-ssE0)/(2))/((ssE0-ssE)/(16)):Fval
[1] 5.85351
p_value<-pf(Fval,2,16,lower.tail = FALSE);p_value
[1] 0.01236634
# this analysis is incorrect
aov(response~block*treatment,GCB) %>% summary
                Df Sum Sq Mean Sq F value Pr(>F)
block
                 8 124741 15593 229,406 < 2e-16 ***
                2 1724
                              862 12.682 1.13e-05 ***
treatment
block:treatment 16 2356
                          147 2 166 0 0102 *
Residuals
          108 7341
                          68
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

#### Alternative codes

```
aov(response~ treatment + Error(block/treatment), data = GCB) %>% summary
Error: block
         Df Sum Sq Mean Sq F value Pr(>F)
Residuals 8 124741 15593
Error: block:treatment
         Df Sum Sq Mean Sq F value Pr(>F)
treatment 2 1724 862.0 5.854 0.0124 *
Residuals 16 2356 147.3
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Error: Within
          Df Sum Sq Mean Sq F value Pr(>F)
Residuals 108 7341 67.97
cellmeans <- tapply( GCB$response, list(GCB$block, GCB$treatment), mean)
# tidyverse version
summ GCB<-GCB%>%
group_by(block,treatment) %>% summarise(response = mean(response))
aov( response ~block+ treatment, summ_GCB) %>% summary
           Df Sum Sq Mean Sq F value Pr(>F)
           8 24948 3118 5 105 888 2 19e-12 ***
block
treatment 2 345 172.4 5.854 0.0124 *
Residuals 16 471 29.5
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```