MA50259: Statistical Design of Investigations

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Lecture 4

Generalised Inverse and its properties

Let A be a square matix. A generalised inverse (G-inv) of A is any matrix A^{-} such that

$$AA^{-}A = A$$

Properties: Let $G = A^- = (X^T X)^-$

- \bullet G is also squared
- G^T is also a G-inverse of $A = X^T X$
- Let $H_X = XGX^T$ (Projection matrix or Hat matrix): Projection onto the column space of X i.e.

$$H_X X = X$$

- $H_X = XGX^T$ is invariant to the choice of G
- $H_X = XGX^T$ is symmetric
- $\operatorname{rank}(H_X) = \operatorname{rank}(XGX^T) = \operatorname{rank}(X)$

Properties of the Projection Matrix H_X

- $(I H_X)$ is invariant to the choice of G
- Both H_X and $(I H_X)$ are idempotent
- SSE = $(Y X\tilde{\beta})(Y X\tilde{\beta})^T = Y^T(I H_X)Y$ and is invariant to the choice of G
- If A is idempotent, then rank(A) = trace(A)
- $H_X \mathbf{1} = \mathbf{1}$, since $\mathbf{1}$ is a column of X

ANOVA for CRD

• Treatment effects model using matrix notation:

$$y = X\beta + \epsilon$$

where $\epsilon \sim MVN(\mathbf{0}, \sigma^2 \mathbf{I})$ and

$$\beta = \begin{pmatrix} \mu \\ \tau_1 \\ \tau_2 \\ \vdots \\ \tau_t \end{pmatrix}$$

• We test $H_0: \tau_1 = \tau_2 = \ldots = \tau_t$

Reduced model

$$y = X_0 \beta + \epsilon$$

where $\beta = \mu$ and

$$X_0 = \begin{pmatrix} 1 \\ 1 \\ \vdots \\ 1 \end{pmatrix}$$

Sum of Squares of Errors

 $\bullet\,$ SSE under the general model :

$$SSE = (\mathbf{y} - \mathbf{X}\tilde{\boldsymbol{\beta}})^T (\mathbf{y} - \mathbf{X}\tilde{\boldsymbol{\beta}})$$

where $\tilde{\beta}$ are the solutions to normal equations.

• SSE under the null model:

$$\mathrm{SSE}_0 = \left(\mathbf{y} - \mathbf{X_0} \hat{\boldsymbol{\mu}}\right)^T \left(\mathbf{y} - \mathbf{X_0} \hat{\boldsymbol{\mu}}\right) = \mathbf{y^T} \mathbf{A} \mathbf{y}$$

where $A = I_n - \frac{1}{n} \mathbf{1} \mathbf{1}^T$

Comparing SSE Under Null and Alternative

- A is idempotent and rank(A) = trace(A) = n 1
- $SSE_0 > SSE$ (Fit is always worse if less parameters) and

$$SSE_0 - SSE = \mathbf{y^TBy}$$

where $\mathbf{B} = \mathbf{X}\mathbf{G}\mathbf{X}^{\mathbf{T}} - \frac{1}{n}\mathbf{1}\mathbf{1}^{T}$

• B is also idempotent and rank(B) = trace(B) = t - 1

Chi-square Distributions in the ANOVA Table

• If $\boldsymbol{y} \sim \text{MVN}(\boldsymbol{\mu}, \sigma^2 \boldsymbol{I})$, Then

$$\frac{\mathbf{y^TAy}}{\sigma^2} \sim \chi^2_{(\mathrm{rank}(A),\lambda)}$$

where A is an idempotent matrix and λ is the non-centrality parameter

- under H_0 , $\lambda = 0$
- In general for a balanced design with r replicates, $\lambda = \frac{r}{\sigma^2} \sum_{i=1}^t (\mu_i \bar{\mu})^2$

ANOVA Table

Sources	df	Sum of Squares	Mean Squares	F
Treatment	t-1	SST	$MST = \frac{SST}{t-1}$	
Error	n- t	SSE	$MSE = \frac{\mathring{S}S\mathring{E}}{n-t}$	$F = \frac{MST}{MSE}$
Total	n-1	$SS_{ m Total}$		

- Under H_0 , $F \sim F_{t-1,n-t}$ (Central F-distribution)
- In general, $F \sim F_{t-1,n-t,\lambda}$ (Non Central F-distribution)