

# Time Series Analysis

Homework of week 7

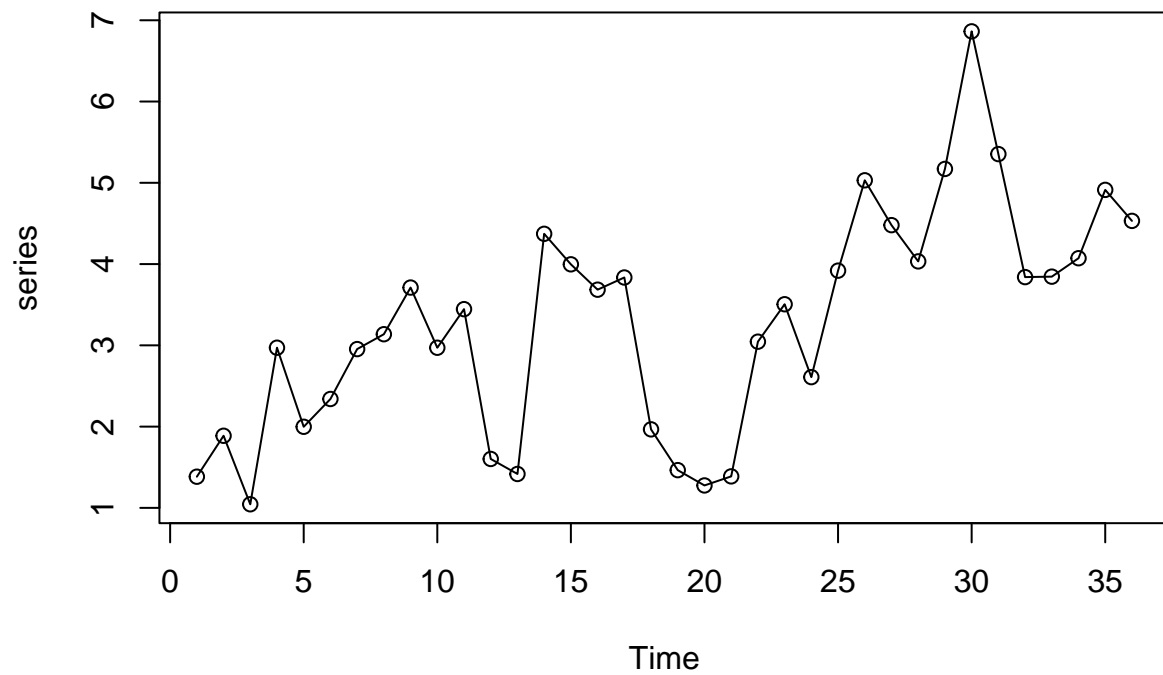
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6.32

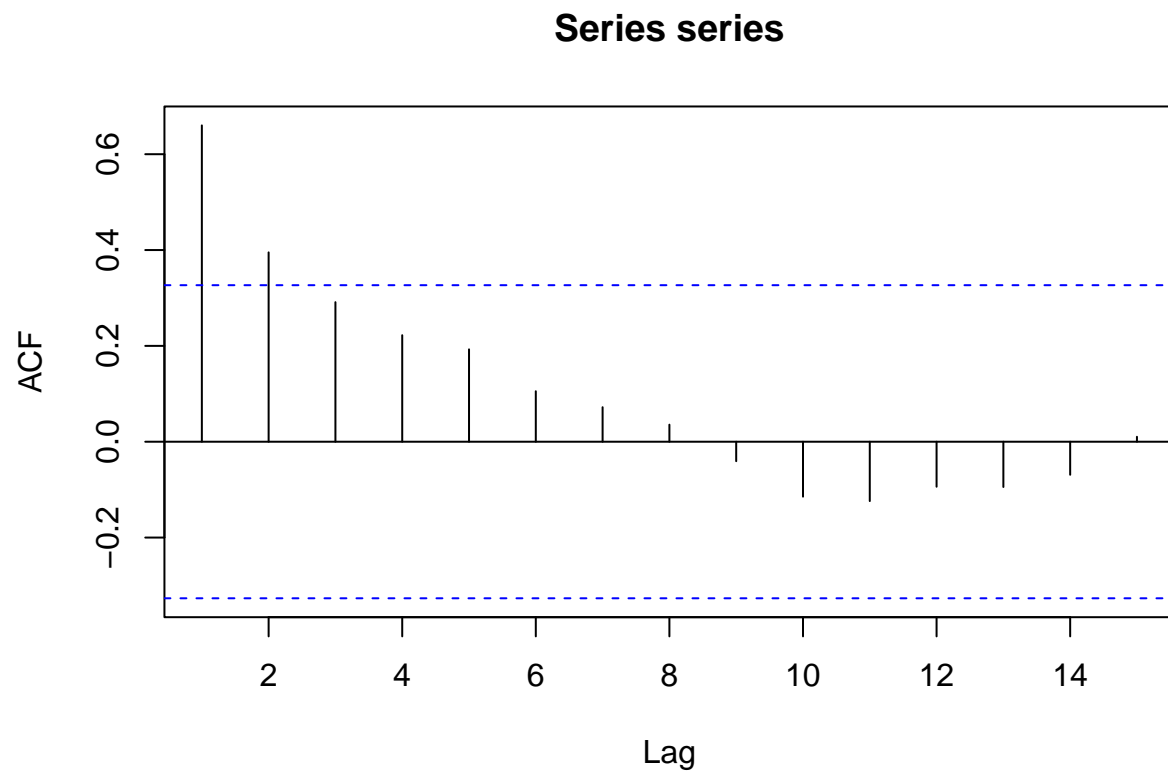
```
set.seed(274135)
series <- arima.sim(n = 36, list(ar = 0.95))
```

(a)

```
plot(series, type = 'o')
```

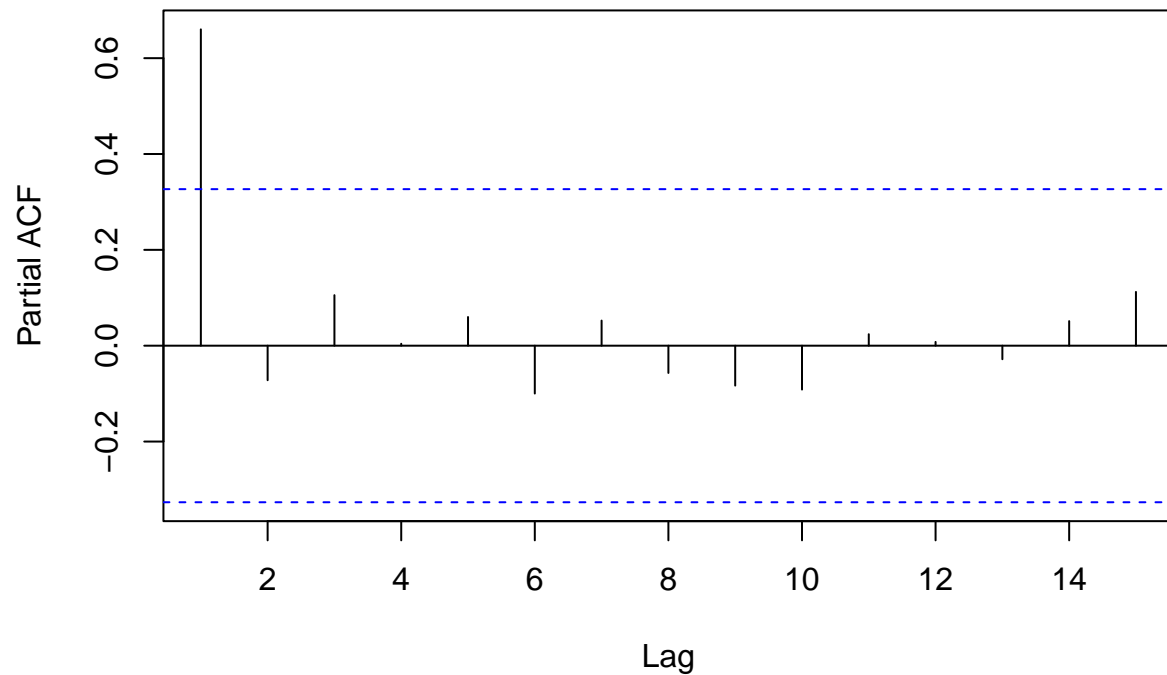


```
acf(series)
```



```
pacf(series)
```

## Series series



The ACF and PACF graphs would lead us to at least entertain an AR(1) model. However, the “upward trend” in the time series plot suggests nonstationarity of some kind

(b)

```
library(uroot)
ADF.test(series, selectlags=list(Pmax=0), itsd=c(1,0,0))
```

```
## Warning in interpval(code = code, stat = adfreg[, 3], N = N): p-value is
## greater than printed p-value
```

```
## -----
## Augmented Dickey & Fuller test
## -----
##
## Null hypothesis: Unit root.
## Alternative hypothesis: Stationarity.
##
## ----
## ADF statistic:
##
##      Estimate Std. Error t value Pr(>|t|)
## adf.reg    -0.323     0.124  -2.597    0.1
##
```

```
## Lag orders: 0
## Number of available observations: 35
```

The Dickey-Fuller test results suggest that we consider a nonstationary model for these data.

(c)

```
ar(diff(series))
```

```
##
## Call:
## ar(x = diff(series))
##
## Coefficients:
##      1      2
## -0.1573 -0.3090
##
## Order selected 2  sigma^2 estimated as 1.115
```

```
ADF.test(series, selectlags=list(mode=c(1,2)), itsd=c(1,0,0))
```

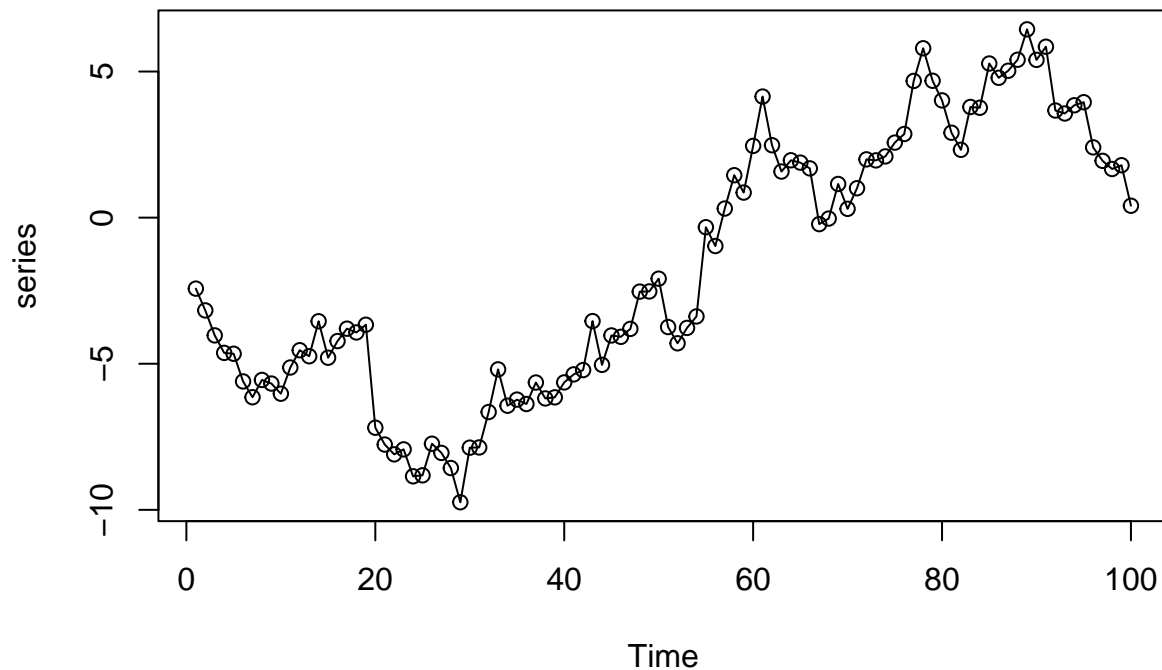
```
## Warning in interpval(code = code, stat = adfreg[, 3], N = N): p-value is
## greater than printed p-value
```

```
## ----- - -----
## Augmented Dickey & Fuller test
## ----- - -----
##
## Null hypothesis: Unit root.
## Alternative hypothesis: Stationarity.
##
## ----
## ADF statistic:
##
##      Estimate Std. Error t value Pr(>|t|)
## adf.reg    -0.335     0.155  -2.156    0.1
##
## Lag orders: 1 2
## Number of available observations: 33
```

The augmented Dickey-Fuller test, also suggests a unit root for this series.

(d)

```
set.seed(274153)
series=arima.sim(n=100,list(ar=0.95))
plot(series,type='o')
```



```
ADF.test(series, selectlags=list(Pmax=0), itsd=c(1,0,0))
```

```
## Warning in interpval(code = code, stat = adfreg[, 3], N = N): p-value is
## greater than printed p-value
```

```
## -----
## Augmented Dickey & Fuller test
## -----
##
## Null hypothesis: Unit root.
## Alternative hypothesis: Stationarity.
##
## ----
## ADF statistic:
##
##      Estimate Std. Error t value Pr(>|t|)
## adf.reg   -0.024     0.023  -1.073    0.1
##
## Lag orders: 0
## Number of available observations: 99
```

```
# The Dickey-Fuller test results suggest that we consider a nonstationary model for these data.
ar(diff(diff(series)))
```

```
##
```

```
## Call:
## ar(x = diff(diff(series)))
##
## Coefficients:
##      1      2      3      4
## -0.8531 -0.6263 -0.4173 -0.2979
##
## Order selected 4  sigma^2 estimated as  1.235
```

```
ADF.test(series, selectlags=list(mode=c(1,2,3,4)), itsd=c(1,0,0))
```

```
## Warning in interpval(code = code, stat = adfreg[, 3], N = N): p-value is
## greater than printed p-value
```

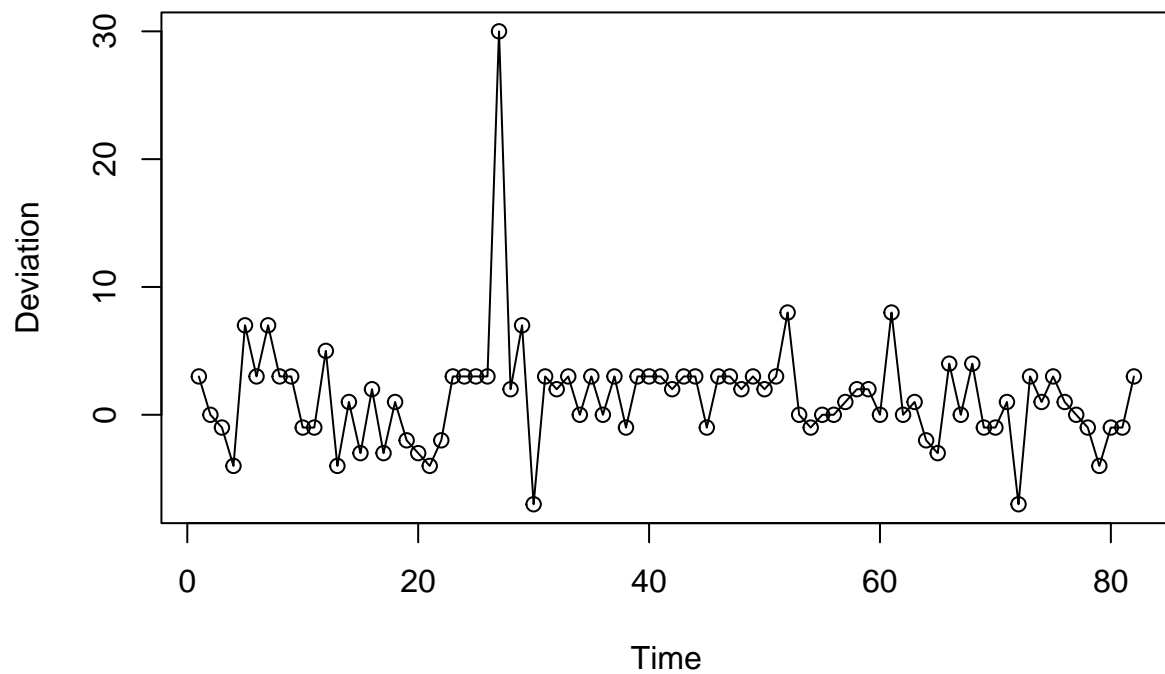
```
## ----- - - - - -
## Augmented Dickey & Fuller test
## ----- - - - - -
##
## Null hypothesis: Unit root.
## Alternative hypothesis: Stationarity.
##
## ----
## ADF statistic:
##
##      Estimate Std. Error t value Pr(>|t|)
## adf.reg    -0.025     0.024   -1.05    0.1
##
## Lag orders: 1 2 3 4
## Number of available observations: 95
```

```
# The augmented Dickey-Fuller test, also suggests a unit root for this series.
```

## 6.33

(a)

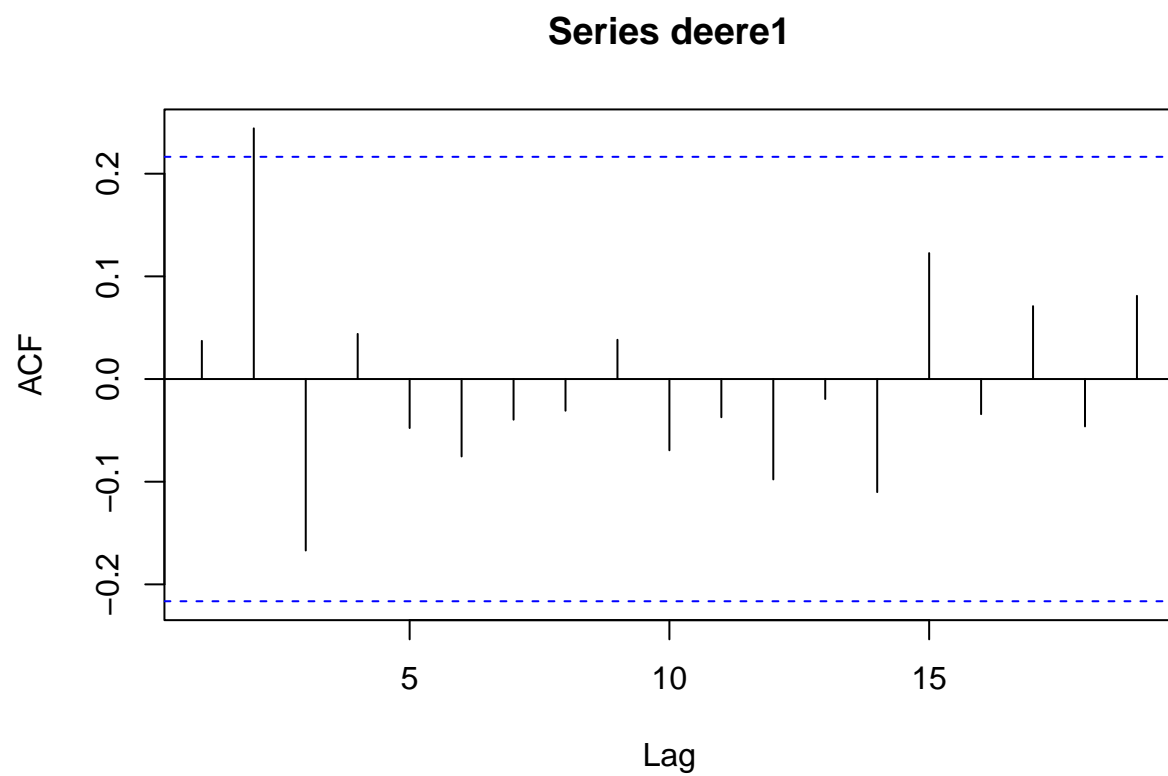
```
data(deere1)
plot(deere1, type = 'o', ylab = 'Deviation')
```



Except for one point of 30 at  $t = 27$  the process seems relatively stable and stationary.

(b)

```
acf(deere1)
```



The graph indicates a statistically significant autocorrelation at lag 2.

(c)

```
t <- which(deere1 == max(deere1)); t; deere1[t]
```

```
## [1] 27
```

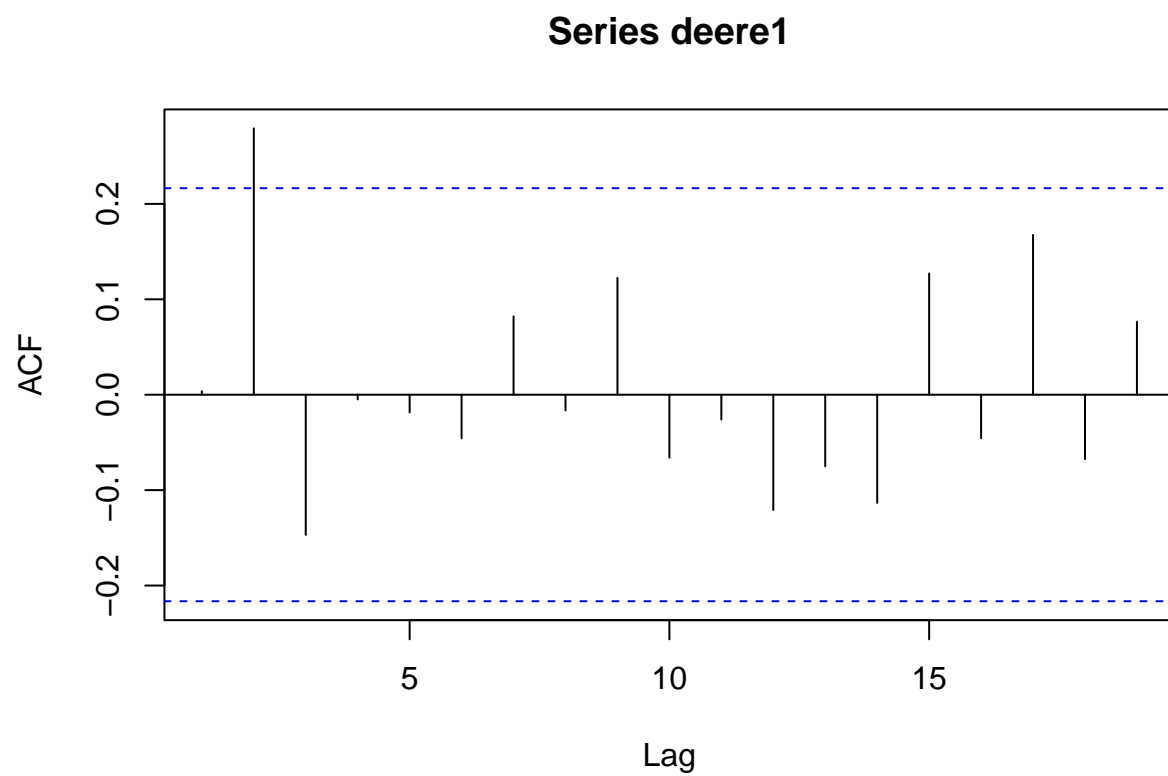
```
## [1] 30
```

```
deere1[t] <- max(deere1[-t]); deere1[t]
```

```
## [1] 8
```

```
acf(deere1)
```

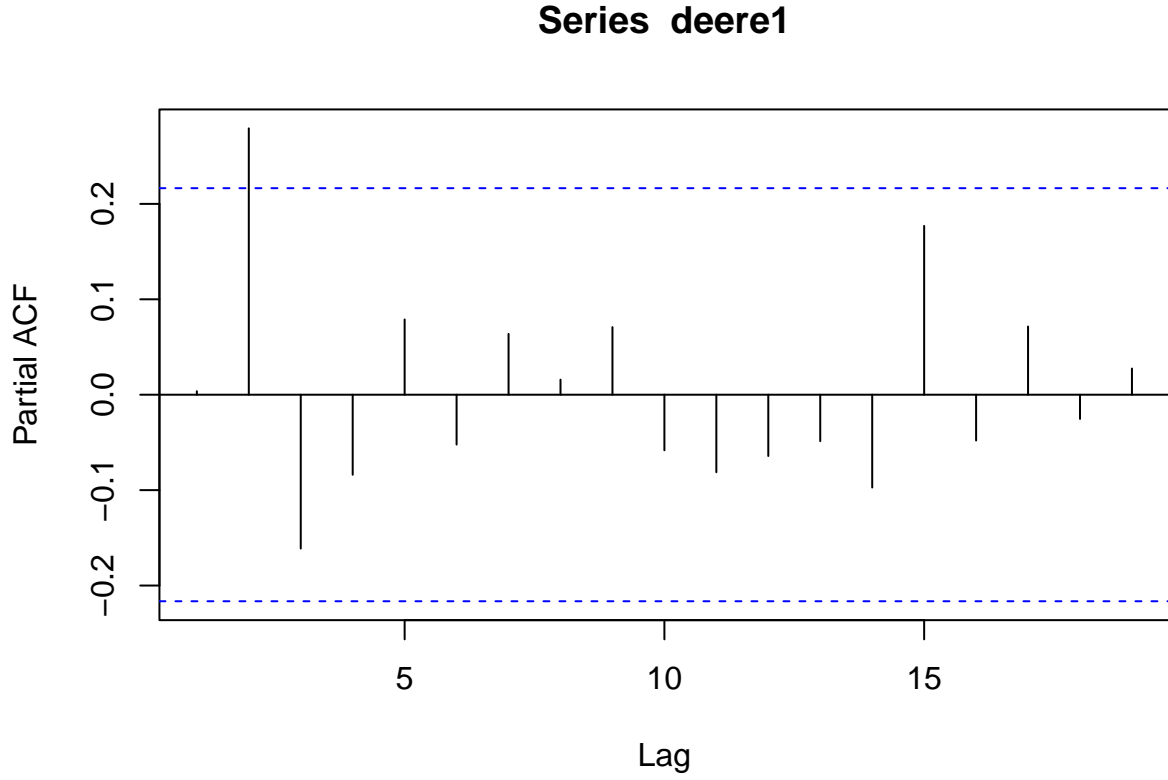




We replaced the unusual value of 30 at time 27 with the next largest value of 8. This had only a small effect on the sample autocorrelation function.

(d)

```
pacf(deere1)
```



This pacf suggests an AR(2) model for the series.

## 7.4

(a)

We have  $e_1 = Y_1 = 0, e_2 = Y_2 + \theta Y_1 = -1, e_3 = Y_3 + \theta Y_2 + \theta^2 Y_1 = \frac{1}{2} - \theta$ . Then

$$\sum_t e_t^2 = 1 + (\theta - \frac{1}{2})^2 = \theta^2 - \theta + \frac{5}{4},$$

which implies that  $\hat{\theta} = \frac{1}{2}$ .

(b)

The estimate of the noise variance is given by

$$\hat{\sigma}_e^2 = \frac{1}{n-1} \sum_t e_t^2 = \frac{1}{3-1} \cdot 1 = \frac{1}{2}.$$

## 7.7

MA(1)model:  $Y_t = e_t - \theta e_{t-1}$ . Then  $\rho_1 = \frac{-\theta}{1+\theta^2}$ . Replace  $\rho_1$  with  $r_1$ , we have

$$r_1 = \frac{-\theta}{1+\theta^2},$$

which has two roots  $\theta_1 = \frac{-1+\sqrt{1-4r_1^2}}{2r_1}, \theta_2 = \frac{-1-\sqrt{1-4r_1^2}}{2r_1}$ . Note that  $\theta_1\theta_2 = 1$ . If  $\theta_i > 0$ , then  $\theta_2 > 1 > \theta_1$ . If  $\theta_i < 0$ , then  $\theta_2 < -1 < \theta_1$ . Thus,  $\hat{\theta} = \theta_1 = \frac{-1+\sqrt{1-4r_1^2}}{2r_1}$  since  $|\theta_1| < 1$ .

## Additional problem

The conditional maximum likelihood function is given by

$$\begin{aligned} L(\phi_0, \phi_1, \phi_2, \sigma_e^2 | Y_1, Y_2) &= \prod_{t=3}^T f(Y_t | Y_{t-1}, \dots, Y_2, Y_1) \\ &= \prod_{t=3}^T (2\pi\sigma_e^2)^{-\frac{1}{2}} \exp \left\{ -\frac{1}{2\sigma_e^2} (Y_t - \phi_0 - \phi_1 Y_{t-1} - \phi_2 Y_{t-2})^2 \right\} \\ &= (2\pi\sigma_e^2)^{-\frac{T-2}{2}} \exp \left\{ -\frac{1}{2\sigma_e^2} \sum_{t=3}^T (Y_t - \phi_0 - \phi_1 Y_{t-1} - \phi_2 Y_{t-2})^2 \right\}. \end{aligned}$$

The log-likelihood is

$$l(\phi_0, \phi_1, \phi_2, \sigma_e^2 | Y_1, Y_2) = -\frac{T-2}{2} \log(2\pi\sigma_e^2) - \frac{1}{2\sigma_e^2} \sum_{t=3}^T (Y_t - \phi_0 - \phi_1 Y_{t-1} - \phi_2 Y_{t-2})^2.$$

We have

$$\begin{cases} 0 = \frac{\partial l}{\partial \sigma_e^2} = -\frac{T-2}{2} \cdot \frac{1}{\sigma_e^2} + \frac{1}{2\sigma_e^4} \sum_{t=3}^T (Y_t - \phi_0 - \phi_1 Y_{t-1} - \phi_2 Y_{t-2})^2 \\ 0 = \frac{\partial l}{\partial \phi_0} = -\frac{1}{\sigma_e^2} \sum_{t=3}^T (Y_t - \phi_0 - \phi_1 Y_{t-1} - \phi_2 Y_{t-2})(-1) \\ 0 = \frac{\partial l}{\partial \phi_1} = -\frac{1}{\sigma_e^2} \sum_{t=3}^T (Y_t - \phi_0 - \phi_1 Y_{t-1} - \phi_2 Y_{t-2})(-Y_{t-1}) \\ 0 = \frac{\partial l}{\partial \phi_2} = -\frac{1}{\sigma_e^2} \sum_{t=3}^T (Y_t - \phi_0 - \phi_1 Y_{t-1} - \phi_2 Y_{t-2})(-Y_{t-2}) \end{cases}$$

Thus, the equation of the MLE  $\phi_0, \phi_1, \phi_2, \sigma_e^2$  should satisfy is

$$\begin{cases} \sigma_e^2 = \frac{1}{T-2} \sum_{t=3}^T (Y_t - \phi_0 - \phi_1 Y_{t-1} - \phi_2 Y_{t-2})^2 \\ 0 = \sum_{t=3}^T (Y_t - \phi_0 - \phi_1 Y_{t-1} - \phi_2 Y_{t-2}) \\ 0 = \sum_{t=3}^T (Y_t - \phi_0 - \phi_1 Y_{t-1} - \phi_2 Y_{t-2}) Y_{t-1} \\ 0 = \sum_{t=3}^T (Y_t - \phi_0 - \phi_1 Y_{t-1} - \phi_2 Y_{t-2}) Y_{t-2} \end{cases}$$