Statistical Methods in Image Processing EE-048954 Homework 3: Contrastive Divergence and Noise Contrastive Estimation Due Date: June 16, 2022 **Submission Guidelines** • Submission only in **pairs** on the course website (Moodle). Working environment: We encourage you to work in Jupyter Notebook online using Google Colab as it does not require any installation. You should submit two separated files: A .ipynb file, with the name: ee048954_hw3_id1_id2.ipynb which contains your code implementations. ■ A .pdf file, with the name: ee048954_hw3_id1_id2.pdf which is your report containing plots, answers, and discussions. ■ No handwritten submissions and no other file-types (.docx , .html , ...) will be accepted. Mounting your drive for saving/loading stuff #from google.colab import drive #drive.mount('/content/drive') Importing relevant libraries ## Standard libraries import os import math import time import numpy as np import random import copy ## Scipy optimization routines from scipy.optimize import minimize ## Progress bar import tqdm ## Imports for plotting import matplotlib.pyplot as plt import matplotlib.animation as animation %matplotlib inline import matplotlib matplotlib.rcParams['lines.linewidth'] = 2.0 plt.style.use('ggplot') Part I: Contrastive Divergence (50 points) Consider the following Gaussian Mixture Model (GMM) distribution $p(x; \{\mu_i\}) = \sum_{i=1}^{N} \frac{1}{N} \frac{1}{2\pi} \exp\left\{-\frac{1}{2}||x - \mu_i||^2\right\},$ where $x, \mu_i \in \mathbb{R}^2$. We will use N = 4, $\sigma = 1$, and $\{\mu_i\} = \{(0,0)^T, (0,3)^T, (3,0)^T, (3,3)^T\}$. Sampling from GMM **Task 1**. Direct sampling: Use your function from HW1 that accepts $\{\mu_i\}$, and returns a sample x from $p(x;\{\mu_i\})$. Draw J=1000samples $\{x\}$ from the distribution $p(x; \{\mu_i\})$ using this function. These will be our **real samples**. def plot scatter(samples, title, group=None): if group is not None: plt.scatter(samples[:, 0], samples[:, 1], c=group) plt.scatter(samples[:, 0], samples[:, 1]) plt.title(title) plt.show() def sub_plot_scatter(samples_list, titles_list, width, figsize=(15,5)): N = len(samples list)fig, axes = plt.subplots(1, 2, figsize=figsize) rows = int(N/width) for i in range(rows): for j in range(width): idx = rows * i + jif rows > 1: axes[i,j] = plt.scatter(samples_list[idx][:, 0], samples_list[idx][:, 1]) axes[i,j].set_title(titles_list[idx]) axes[idx].set title(titles list[idx]) axes[idx].scatter(samples_list[idx][:, 0], samples_list[idx][:, 1]) In [4]: def mix gauss draw(N=4, sigma=1, mu=[np.array([0,0]), np.array([0,3]), np.array([3,0]), $np.array([3,3])], J=1_000):$ Sigma = np.eye(len(mu[0])) * (sigma ** 2)X = np.zeros([J, len(mu[0])])G = np.zeros([J, 1])for i in range(J): m = np.random.choice(N, 1)Mu = mu[int(m)]x = np.random.multivariate normal(mean=Mu, cov=Sigma) X[i] = xG[i] = mreturn X, G real_samples, groups = mix_gauss_draw(sigma=1) plot scatter(samples=real samples, title='Real Samples Scatter', group=groups) Real Samples Scatter 6 2 **Task 2**. Sampling with MCMC: implement the MALA algorithm to draw samples from $p(x;\{\mu_i\})$. The function will get an initial guess $\{\hat{x}_i\}$ and will generate chains of length L. Use $\sqrt{2\varepsilon}=0.1$ and $N\sim\mathcal{N}(0,I)$. def log gradient_p(x, sigma_2=1, mu=[np.array([0,0]), np.array([0,3]), np.array([3,0]), np.array([3,3])]):N = len(mu)denominator = 0for i in range(N): denominator += np.exp((-1/(2*sigma_2)) * np.linalg.norm(x-mu[i], axis=1) ** 2) denominator = np.reshape(denominator, (denominator.shape[0],1)) nominator = 0for i in range(N): $s = (np.exp((-1/(2*sigma_2)) * np.linalg.norm(x-mu[i], axis=1) ** 2))$ d = (mu[i] - x).transpose()nominator += s * d return (1/sigma 2) * nominator.transpose() / denominator def p(x, sigma 2=1,mu=[np.array([0,0]), np.array([0,3]), np.array([3,0]), np.array([3,3])]): N = len(mu)sum exp = 0for n in range(N): log exp = -0.5 * (np.linalg.norm(x-mu[n], axis=1) ** 2)sum exp += np.exp(log exp) return (1/N) * (1 /(2*np.pi)) * sum_exp In [9]: **def** $q(x_{tag}, x, epsilon, mu=[np.array([0,0]), np.array([0,3]),$ np.array([3,0]), np.array([3,3])]): norm = np.linalg.norm(x_tag - x - epsilon * log_gradient_p(x, mu=mu), axis=1) ** 2 log q = -norm/(4*epsilon) $q = np.exp(log_q)$ return q def langevin mala dynamics (init guess, steps=5000, samples=1000, sigma 2=1, factored epsilon=0.1, mu=[np.array([0,0]), np.array([0,3]), np.array([3,0]), np.array([3,3])],MALA=False, verbos=False): epsilon = (factored epsilon ** 2) / 2 x = init guessmala reg update = 0 mala alpha update = 0 mala no update = 0 for s in tqdm.tqdm(range(steps)): noise = np.random.randn(x.shape[0], x.shape[1]) new x = x + epsilon * log gradient p(x, sigma 2=sigma 2, mu=mu) + factored epsilon * noiseif MALA: if p(new x, mu=mu).sum() > p(x, mu=mu).sum():x = new x.copy()mala reg update += 1 else: p x = p(x=x, mu=mu).sum() $p \times new = p(x=new \times, mu=mu).sum()$ $q \times x = q(x + tag=x, x=new x, epsilon=epsilon, mu=mu).sum()$ $q \times new \times = q(x \text{ tag=new } x, x=x, epsilon=epsilon, mu=mu).sum()$ alpha = (p x / p x new) * (q x x new / q x new x)r = np.random.uniform() if r < alpha:</pre> x = new x.copy()mala alpha update += 1 else: mala no update += 1 else: x = new x.copy()if MALA and verbos: print(mala reg update / steps) print(mala alpha update / steps) print(mala no update / steps) return x init_guess = np.random.uniform(low=0, high=3, size=(1_000, 2)) x = langevin_mala_dynamics(init_guess=init_guess, steps=10_000, samples=1 000, MALA=False, sigma 2=1) | 10000/10000 [00:05<00:00, 180 9.22it/splot scatter(samples=x, title='Langevin') Langevin init guess = np.random.uniform(low=0, high=3, size=(1 000, 2)) x = langevin mala dynamics(init guess=init guess, steps=10 000, samples=1 000, MALA=True, sigma 2=1) | 10000/10000 [00:16<00:00, 60 6.90it/s] In [14]: plot_scatter(samples=x, title='MALA') MALA **From now on**, we will refer to $\{\mu_i\}$ as **unknowns** and we will estimate them using different algorithms. Estimation of $\{\mu_i\}$ **Task 3**. Implement Maximum likelihood (ML) estimation of $\{\mu_i\}$ using direct sampling: • Step 1: Randomly initialize $\{\tilde{\mu}_i\}$ from $U([0,3]^2)$. Step 2: Use your function from Task 1 to draw 100 samples \tilde{x} from $p(x; \{\mu_i\})$ using $\{\tilde{\mu}_i\}$. Step 3: Update $\{\tilde{\mu}_i\}$ using the ML gradient descent step: $ilde{\mu}_i^{k+1} = ilde{\mu}_i^k + \eta \left(\langle
abla_{\mu_i} \log p(x; \{\mu_i\})
angle_x - \langle
abla_{\mu_i} \log p(x; \{\mu_i\})
angle_{ ilde{x}}
ight),$ where $\langle \cdot \rangle_x$ denotes averaging over the real samples from Task 1 and $\langle \cdot \rangle_{\tilde{x}}$ denotes averaging over the synthetically generated samples from Step 2. Use $\eta = 1$. Repeat Step 2 and Step 3 until convergence. def grad mu log p(x, mu): N = len(mu)denominator = 0for n in range(N): log exp = -0.5 * (np.linalg.norm(x-mu[n]) ** 2)denominator += np.exp(log exp) nominator = np.zeros([N,2])for n in range(N): log exp = -0.5 * (np.linalg.norm(x-mu[n]) ** 2)exp = np.exp(log exp)nominator[n] = exp * (x-mu[n])return (1/denominator) * nominator def avg_grad_mu_log_p(x, mu): grad = np.zeros(shape=mu.shape) for x_i in x: grad += grad_mu_log_p(x=x_i, mu=mu) return grad / x.shape[0] def ML DS(N, real samples, etta=1, draw mcmc=False, init task 1=False): mu hat = np.random.uniform(size=[N,2]) * 3 step idx = 0while True: if draw mcmc: = mix gauss draw(N=N, sigma=1, J=100) # using original mu's x hat = langevin mala dynamics(init guess=init guess, steps=10, samples=100, MALA=True, sigma 2 else: init guess = np.random.randn(100, 2) * np.sqrt(2) + 1.5 x hat = langevin mala dynamics(init guess=init guess, steps=1 000, samples=100, MALA=True, sign x_hat , _ = $mix_gauss_draw(N=N, sigma=1, mu=mu_hat$, J=100) avg mu grad estim = avg grad mu log p(x hat, mu hat) avg mu grad real = avg grad mu log p(real samples, mu hat) old mu = mu hat.copy() mu hat = mu hat + etta * (avg mu grad real - avg mu grad estim) step norm = np.linalg.norm(mu hat - old mu) if step norm < 1e-1:</pre> print(step norm) break **if** step idx % 10 == 0: print(step norm) step idx += 1 return mu hat In [18]: real_samples, groups = mix_gauss_draw(sigma=1) estim mu = ML DS(N=4, real samples=real samples, etta=1)print("Estimated MU:\n", estim_mu) 0.2558766679562691 0.21107447927951567 0.1263122438198493 0.08803442535979092 Estimated MU: [[3.01091859 0.03061885] $[-0.01691787 \quad 3.00878443]$ [-0.07696866 -0.04974094]] $^{f J}$ Task 4. Implement Maximum likelihood (ML) estimation of $\{\mu_i\}$ using MCMC: • Step 1: Randomly initialize $\{\tilde{\mu}_i\}$ from $U([0,3]^2)$. • Step 2: Use your function from Task 2 to draw 100 samples ilde x from $p(x;\{\mu_i\})$ using $\{ ilde\mu_i\}$. Initialize the chains with $\hat x_i\sim \mathcal{N}(1.5,2)$ and use chains length of L=1000. • Step 3: Update $\{\tilde{\mu}_i\}$ using the ML gradient descent step: $ilde{\mu}_i^{k+1} = ilde{\mu}_i^k + \eta \left(\langle
abla_{\mu_i} \log p(x; \{\mu_i\})
angle_x - \langle
abla_{\mu_i} \log p(x; \{\mu_i\})
angle_{ ilde{x}}
ight),$ where $\langle \cdot \rangle_x$ denotes averaging over the real samples from Task 1 and $\langle \cdot \rangle_{\tilde{x}}$ denotes averaging over the synthetically generated samples from Step 2. Use $\eta = 1$. Repeat Step 2 and Step 3 until convergence. estim mu mala = ML DS(N=4, real samples=real samples, etta=1, draw mcmc=True) print("Estimated MU:\n", estim mu mala) 1000/1000 [00:00<00:00, 115 3.55it/s]0.45575212224887335 100%| 1000/1000 [00:00<00:00, 119 4.50it/s]100%| 1000/1000 [00:00<00:00, 124 3.23it/s100%| 1000/1000 [00:00<00:00, 122 6.41it/s] 1000/1000 [00:00<00:00, 124 100%| 4.78it/s] 1000/1000 [00:00<00:00, 126 100%| 9.71it/s1000/1000 [00:00<00:00, 127 100%| 1.85it/s] 100%| 1000/1000 [00:00<00:00, 123 9.83it/s] 1000/1000 [00:00<00:00, 123 100%| 3.35it/s100%| 1000/1000 [00:00<00:00, 126 6.69it/s] 100%| 1000/1000 [00:00<00:00, 116 3.90it/s] 0.12004580897410269 100%| 1000/1000 [00:00<00:00, 109 7.99it/s] 0.09669773564672267 Estimated MU: [3.07690851 0.132218] $[-0.06140892 \quad 0.06579094]$ [-0.10931543 3.10781589]] **Task 5**. Implement Contrastive Divergence (CD) estimation of $\{\mu_i\}$ using MCMC sampling: Step 1: Randomly initialize $\{\tilde{\mu}_i\}$ from $U([0,3]^2)$. Step 2: Use your function from Task 2 to draw 100 samples \tilde{x} from $p(x; \{\mu_i\})$ using $\{\tilde{\mu}_i\}$. Initialize the chains with **100 samples** randomly chosen from the real set of examples from Task 1, and use only L=10 update steps. • Step 3: Update $\{\tilde{\mu}_i\}$ using the CD gradient descent step: $ilde{\mu}_i^{k+1} = ilde{\mu}_i^k + \eta \left(\langle
abla_{\mu_i} \log p(x; \{\mu_i\})
angle_x - \langle
abla_{\mu_i} \log p(x; \{\mu_i\})
angle_{ ilde{x}}
ight),$ where $\langle \cdot \rangle_x$ denotes averaging over the 100 real samples used for initialization of the chains in Step 2 and $\langle \cdot \rangle_{\tilde{x}}$ denotes averaging over the MCMC generated samples from Step 3. Use $\eta=1$. Repeat Step 2 and Step 3 until convergence. $\texttt{estim_mu_mala_init_task_1 = ML_DS (N=4, real_samples=real_samples, etta=1, draw_mcmc=True, init_task_1=True) }$ print("Estimated MU:\n", estim mu mala init task 1) 10/10 [00:00<00:00, 248 100%| 6.84it/s0.2707670236008407 10/10 [00:00<00:00, 248 5.66it/s100%| | 10/10 [00:00<00:00, 252 100%| | 10/10 [00:00<00:00, 248 4.92it/s] 100%| 10/10 [00:00<00:00, 200 5.02it/s] 100%| | 10/10 [00:00<00:00, 125 3.04it/s]100%| | 10/10 [00:00<00:00, 166 8.58it/s] 100%| | 10/10 [00:00<00:00, 245 8.70it/s] 100%| | 10/10 [00:00<00:00, 247 4.22it/s] 100%| | 10/10 [00:00<00:00, 250 6.61it/s] 100%| | 10/10 [00:00<00:00, 167 1.37it/s] 0.1205612773427741 100%| 10/10 [00:00<00:00, 143 2.33it/s] 100%| 10/10 [00:00<00:00, 200 9.83it/s] 0.09028178109519877 Estimated MU: [[3.29127853 0.60871475] [1.08688689 -0.55987471] [0.05760177 2.03691587]] **Task 6**. Present the estimated $\{\mu_i\}$ and the final random samples $\{\tilde{x}_i\}$ generated with each of the three algorithms in Tasks 3-5. Discuss the differences in convergence. import itertools **def** calc diff mu(mu hat, mu=[np.array([0,0]), np.array([0,3]), np.array([3,0]), np.array([3,3])]):mu = np.array(mu) permutations = np.array(list(itertools.permutations(mu hat))) min tot err = 1e6 for perm in permutations: tot_err = np.linalg.norm(mu-perm) ** 2 if tot err < min tot err:</pre> min tot err = tot err return min tot err In [24]: print("Task 3:\n", estim mu) print("Squared Error: ", calc diff mu(mu hat=estim mu)) print("\nTask 4:\n", estim mu mala) print("Squared Error: ", calc diff mu(mu hat=estim mu mala)) print("\nTask 5:\n", estim mu mala init task 1) print("Squared Error: ", calc diff mu(mu hat=estim mu mala init task 1)) Task 3: [[3.01091859 0.03061885] [-0.01691787 3.00878443] [-0.07696866 -0.04974094]] Squared Error: 0.010022005521795623 Task 4: [[2.73926735 3.1403476] [3.07690851 0.132218] [-0.06140892 0.06579094] [-0.10931543 3.10781589]] Squared Error: 0.1427491165494605 Task 5: [[3.29127853 0.60871475] [1.08688689 -0.55987471] [0.05760177 2.03691587]] Squared Error: 3.3977992119222185 gen mu 3, g 3=mix gauss draw(mu=estim mu) gen mu 4, g 4=mix gauss draw(mu=estim mu mala) gen mu 5, g 5=mix gauss draw(mu=estim mu mala init task 1) plot_scatter(samples=gen_mu_3, title='Generated samples MK_DS ', group=g_3) Generated samples MK_DS plot scatter(samples=gen mu 4, title='Generated samples MK_MCMC L=1000 ', group=g_4) Generated samples MK_MCMC L=1000 6 2 plot scatter(samples=gen mu 5, title='Generated samples MK MCMC L=10 ', group=g 5) Generated samples MK_MCMC L=10 **Discussion:** From the plots it's very hard to see any pattern even on the real data. So, to distinguish between the generated samples is even harder. In order to estimate the diffrences between the $\hat{\mu}$ we calculated the squared error between each one and the real μ . Results shows as excpected that the ML estimation based on the real results are the best and the generated samples based on L=10 are worse that generated MCMC samples based on L=1000 as excpected since the L=10 will not converge as good to the real data distribution as L=1000. Part II: Noise Contrastive Estimation (50 points) Consider the distribution $p_m(x;\{\mu_i\}) = rac{1}{Z}\sum_{i=1}^N \expiggl\{-rac{1}{2\sigma^2}||x-\mu_i||^2iggr\},$ where $Z \in \mathbb{R}$ is a normalization constant, and $x, \mu_i \in \mathbb{R}^2$. Sampling from GMM Task 7. What is the value of Z? Easy to see that: $\Sigma = I\sigma^2$ $\Rightarrow rac{1}{Z} = rac{1}{N} \cdot \det(2\pi\Sigma)^{-1/2} = rac{1}{2\pi N \sigma^2}$ $\Rightarrow Z = 2\pi N \sigma^2$ Task 8. Use N=4, $\sigma=1$, and $\{\mu_i\}=\{(0,0)^T,(0,3)^T,(3,0)^T,(3,3)^T\}$. Draw J=1000 samples $\{x_j\}$ from the distribution $p_m(x; \{\mu_i\})$ using the function from Task 1. real_samples, groups = mix_gauss_draw(N=4, sigma=1, mu=[np.array([0,0]), np.array([0,3]), np.array([3,0]),
np.array([3,3])], J=1_000) **From now on**, we will refer to $\{\mu_i\}$ as **unknowns** and we will estimate them using the Noise Contrastive Estimation method. Estimation of $\{\mu_i\}$ **Task 9**. Implement Noise Contrastive Estimation of $\{\mu_i\}$: • Step 1: Generating the artificial data-set of noise: Draw J=1000 samples $\{y_i\}$ from $p_n(y;\mu_n) = rac{1}{2\pi\sigma_\pi^2} \mathrm{exp}iggl\{ -rac{1}{2\sigma_\pi^2} ||y-\mu_n||^2 iggr\}$ using $\mu_n=(1,1)^T$ and $\sigma_n=2$. • Step 2: Randomly select an initial guess for the model means $\{\tilde{\mu}_i\}$ from $U([0,3]^2)$. • Step 3: Update $\{\tilde{\mu}_i\}$ by **maximizing**: $\{ ilde{\mu}_i\} = rg \max_{\{\mu_i\}} \ \sum_{i=1}^J \left[\ln(h(x_j; \{\mu_i\})) + \ln(1 - h(y_j; \{\mu_i\}))
ight],$ where $h(u;\{\mu_i\}) = rac{p_m(u;\{\mu_i\})}{p_m(u;\{\mu_i\}) + p_m(u;\mu_m)}.$ Implementation Tip: This step can be executed using the function scipy.optimize.minimize which finds the minimum of an (unconstrained) optimization problem (e.g. using the 'BFGS' method), given a function that calculates the objective and an initial guess (see scipy documentation for more details). In our case, for maximization, implement a function that calculates the minus of the objective above. def p(x, sigma squared, means vec, Z=None): N = len(means vec) if Z is None: Z = N*2*np.pi*sigma squaredsum exp = 0for n in range(N): log exp = -(0.5/sigma squared) * (np.linalg.norm(x-means vec[n], axis=1) ** 2)sum exp += np.exp(log exp)return 1/Z * sum exp **def** h(x, mu, Z = None): pm = p(x=x, sigma_squared=1, means vec=mu, Z=Z) pn = p(x=x, sigma_squared=4, means_vec=np.array([1,1])) return pm/(pm+pn) def objective(params, *args): sum = 0real_samples, noise_samples = args[0], args[1] mu = params.reshape(-1, 2)sum = np.sum(np.log(h(x=real samples, mu=mu)) + np.log(1-h(x=noise samples, mu=mu))) return -sum J = 1000sigma n=2 mu n=[np.array([1,1])]noise_samples_9 = np.random.randn(J, 2) * np.sqrt(sigma_n) + 1 $mu_hat_9 = np.random.uniform(size=[4,2]) * 3$ results = minimize(method='BFGS', fun=objective, x0=mu_hat_9, args=(real_samples, noise_samples_9)) #, options= mu task 9 = results.x.reshape(-1,2)print(mu_task_9) [[-0.37680239 -0.2867889] [-0.36224257 3.48330967] [3.22317244 3.04671973] [3.26963204 -0.16830731]] We will now regard both $\{\mu_i\}$ and the normalization constant Z as unknowns, and will estimate them using Noise Contrastive Estimation. Task 10. Implement Noise Contrastive Estimation with an un-normalized probability model: • Step 1: Generating the artificial data-set of noise: Draw J=1000 samples $\{y_i\}$ from $p_n(y;\mu_n) = rac{1}{2\pi\sigma_\pi^2} \mathrm{exp}iggl\{ -rac{1}{2\sigma_\pi^2} ||y-\mu_n||^2 iggr\}$ using $\mu_n=(1,1)^T$ and $\sigma_n=2$. • Step 2: Randomly select an initial guess for the model means $\{\tilde{\mu}_i\}$ from $U([0,3]^2)$, and for the normalization constant Z from • Step 3: Update $\{\tilde{\mu}_i\}$ and Z by **maximizing**: $\{ ilde{\mu}_i\}, Z = rgmax_{\{\mu_i\}, Z} \ \sum_{i=1}^J \left[\ln(h(x_j; \{\mu_i\}, Z)) + \ln(1 - h(y_j; \{\mu_i\}, Z))
ight],$ where $h(u;\{\mu_i\},Z) = rac{p_m(u;\{\mu_i\},Z)}{p_m(u;\{\mu_i\},Z) + p_n(u;\mu_n)}.$ def Objective Z(params, *args): obj sum = 0real samples, noise samples = args[0], args[1] $mu_hat = params[:-1].reshape(-1,2)$ Z = params[-1] $obj_sum = np.sum (np.log (h (x=real_samples, mu=mu_hat, Z=Z)) + np.log (1-h (x=noise_samples, mu=mu_hat, Z=Z)))$ return -obj sum In [46]: J = 1000sigma=2 mu=[np.array([1,1])] $noise_samples_10 = np.random.randn(J, 2) * np.sqrt(2) + 1$ mu hat 10 = np.random.uniform(size=[4,2]) * 3 $Z_10 = np.random.uniform() * 0.9 + 0.1$ $mu_hat_Z = np.append(mu hat 10, Z 10)$ In [47]: results Z = minimize (method = 'BFGS', fun = Objective Z, x0 = mu hat Z, args = (real samples, noise samples 10))mu task $10 = results_{Z.x[:8].reshape(-1,2)}$ $Z \text{ task } 10 = \text{results}_Z.x[8]$ print(Z_task_10) print(mu_task_10) 19.623383237463372 [[-0.19051958 - 0.55742932][-0.38375688 3.44490873] [3.50002984 -0.0390691]] **Evaluating the Results Task 11**. Visually: plot the estimates of $\{\mu_i\}$ of Tasks 9 and 10 (two separate figures). Include the model samples, the noise samples, the initial guess for the model means, and the final estimates of $\{\tilde{\mu}_i\}$. In [48]: fig, axes = plt.subplots(1,2, figsize=(15,5)) axes[0].scatter(noise_samples_9[:,0], noise_samples_9[:,1], label='Noise') axes[0].scatter(real_samples[0:500][:,0], real_samples[0:500][:,1], c='green', label='Real Samples') axes[0].scatter(mu_task_9[:,0], mu_task_9[:,1], marker="*", s=200, c='black', label='Final Mu') axes[0].scatter(mu_hat_9[:,0], mu_hat_9[:,1], marker="*", s=200, c='blue', label='Init Mu') axes[0].set_title('Task 9 ') axes[0].legend() axes[1].scatter(noise_samples_10[:,0], noise_samples_10[:,1], label='Noise') axes[1].scatter(real_samples[0:500][:,0], real_samples[0:500][:,1], c='green', label='Real Samples') axes[1].scatter(mu_task_10[:,0], mu_task_10[:,1], marker="*", s=200, c='black', label='Final Mu') axes[1].scatter(mu_hat_10[:,0], mu_hat_10[:,1], marker="*", s=200, c='blue', label='Init Mu') axes[1].set_title('Task 10 ') axes[1].legend() Out[48]: <matplotlib.legend.Legend at 0x24898400898> Task 10 6 2 Real Samples Real Samples Final Mu Final Mu Init Mu Init Mu -2 **Task 12**. Quantitatively: repeat Tasks 9 and 10, this time with J=100 imes[1,5,10,20,30,50]. For each value of J repeat the estimation process for 50 times, each time with different realizations for $\{x_j\}$ and $\{y_j\}$ and initial guesses for the estimands ($\{\mu_i\}$ in Task 9 and $\{\mu_i\}, Z$ in Task 10. For each value of J, calculate the MSE between the true parameter values and their estimates (the mean will be taken over the different realizations). Note that for the model means, the MSE should be calculated to the closest true μ_i for each estimation. If at the same run two estimated μ_i s pick the same true μ_i , then this run should be declared as a failure and should be disregarded. Report the number of failure def calc dist(mu hat, mu=np.array([[0,0],[0,3],[3,0],[3,3]])): map order = {} err = 0for idx, m in enumerate(mu hat): err += np.min(np.linalg.norm(mu-m, axis=1) ** 2) map order[idx] = np.argmin(np.linalg.norm(mu-m, axis=1) ** 2) if not len(set(map order.values())) == len(map order): return False, err return True, err dic res = {} for J in [100, 150, 1_000, 2_000, 3_000, 5_000]: num fail 9 = 0 $num_fail_10 = 0$ tot err 9 = 0tot_err_10 = 0 for i in tqdm.tqdm(range(50)): real_samples, groups = mix_gauss_draw(N=4, sigma=1, mu=[np.array([0,0]), np.array([0,3]), np.array([3,0]), np.array([3,3])], J=J) $noise_samples_9 = np.random.randn(J, 2) * np.sqrt(2) + 1$ mu hat 9 = np.random.uniform(size=[4,2]) * 3results = minimize(method='BFGS', fun=objective, x0=mu_hat_9, args=(real_samples, noise_samples_9)) mu_task_9 = results.x.reshape(-1,2) suc, err = calc_dist(mu_task_9) num fail 9 += int(not(suc)) tot_err_9 += err noise_samples_10 = np.random.randn(J, 2) * np.sqrt(2) + 1 $mu_hat_10 = np.random.uniform(size=[4,2]) * 3$ $Z_10 = np.random.uniform() * 0.9 + 0.1$ mu_hat_Z = np.append(mu_hat_10, Z_10) results_Z = minimize(method='BFGS', fun=Objective_Z, x0=mu_hat_Z, args=(real_samples, noise_samples_10) $mu_task_10 = results_z.x[:8].reshape(-1,2)$ $Z_{\text{task}} = results_{Z.x}[8]$ suc, err = calc_dist(mu_task_10) num_fail_10 += int(not(suc)) tot_err_10 += err $dic_{res}[(9, J)] = (num_{fail_9}, tot_{err_9})$ dic_res[(10, J)] = (num_fail_10, tot_err_10)

