

4.1

P57 18.  $x \in [0, a]$ .  $F(x) = P(X \leq x) = kx$

$$F(a) = ka = 1. \quad k = \frac{1}{a}$$

$$\text{故 } F(x) = \begin{cases} 0, & x < 0 \\ \frac{1}{a}x, & 0 \leq x < a \\ 1, & x \geq a \end{cases}$$

19. (1)  $P(X \leq 3) = F(3) = 1 - e^{-1.2}$

(2)  $P(X \geq 4) = 1 - F(4-0) = 1 - (1 - e^{-1.6}) = e^{-1.6}$

(3)  $P(3 \leq X \leq 4) = F(4) - F(3-0) = 1 - e^{-1.6} - 1 + e^{-1.2} = e^{-1.2} - e^{-1.6}$

(4)  $P(X \leq 3 \cup X \geq 4) = P(X \leq 3) + P(X \geq 4) = 1 - e^{-1.2} + e^{-1.6}$

(5)  $P(X = 2.5) = 0$

20. (1)  $P(X < 2) = F(2-0) = \ln 2$

$$P(0 < X \leq 3) = F(3) - F(0) = 1 - 0 = 1$$

$$P(2 < X < \frac{5}{2}) = F(\frac{5}{2}-0) - F(2) = \ln \frac{5}{2} - \ln 2 = \ln \frac{5}{4} = \ln 5 - 2 \ln 2$$

(2)  $1 < x < e$  时,  $F_x(x)$  连续 又因  $F_x(x) = \int_{-\infty}^x f_x(t) dt$   
 所以  $f_x(x) = F'_x(x) = \frac{1}{x}$

$$\text{故 } f_x(x) = \begin{cases} \frac{1}{x}, & 1 < x < e \\ 0, & \text{其他} \end{cases}$$

21. (1).  $x < 1$ ,  $F(x) = 0$

$$1 \leq x \leq 2, F(x) = \int_1^x 2(1 - \frac{1}{t^2}) dt = 2(x + \frac{1}{x}) \Big|_1^x = 2x + \frac{2}{x} - 4$$

$$x > 2, F(x) = \int_1^2 2(1 - \frac{1}{t^2}) dt = 1$$

$$F(x) = \begin{cases} 0, & x < 1 \\ 2x + \frac{2}{x} - 4, & 1 \leq x \leq 2 \\ 1, & x > 2 \end{cases}$$



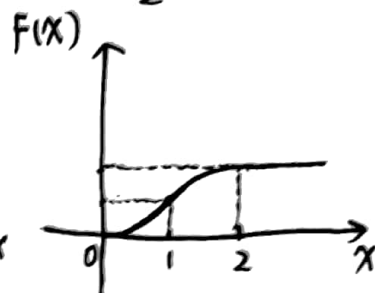
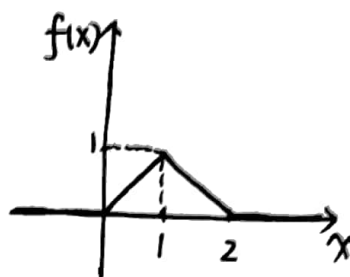
(2).  $x < 0, F(x) = 0$

$0 \leq x < 1, F(x) = \int_0^x x dx = \frac{1}{2} x^2 \Big|_0^x = \frac{1}{2} x^2$

$1 \leq x < 2, F(x) = \int_0^1 x dx + \int_1^x (2-x) dx = \frac{1}{2} + (2x - \frac{1}{2} x^2) \Big|_1^x = \frac{1}{2} + 2x - \frac{1}{2} x^2 - \frac{3}{2}$

$x \geq 2, F(x) = 1$

故  $F(x) = \begin{cases} 0 & x < 0 \\ \frac{1}{2} x^2 & 0 \leq x < 1 \\ -\frac{1}{2} x^2 + 2x - 1 & 1 \leq x < 2 \\ 1 & x \geq 2 \end{cases}$



P58 23.  $F(x) = \int_{-\infty}^x f(t) dt \quad (x > 1000)$

$= \int_{1000}^x \frac{1000}{t^2} dt = 1000 \left( -\frac{1}{t} \right) \Big|_{1000}^x = 1000 \left( -\frac{1}{x} + \frac{1}{1000} \right) = 1 - \frac{1000}{x}$

$P(X > 1500) = 1 - P(X \leq 1500) = 1 - F(1500) = 1 - \left( 1 - \frac{1000}{1500} \right) = \frac{2}{3}$

令  $Y$  表示 5 只中青命大于 1500 小时的只数, 则  $Y \sim B(5, \frac{2}{3})$

$P(X \geq 2) = 1 - P(X=0) - P(X=1) = 1 - \binom{5}{0} \left(\frac{2}{3}\right)^0 \left(\frac{1}{3}\right)^5 - \binom{5}{1} \left(\frac{2}{3}\right)^1 \left(\frac{1}{3}\right)^4$   
 $= 1 - \frac{232}{243}$

24.  $P(X > 10) = 1 - F(10) = 1 - \int_0^{10} \frac{1}{5} e^{-\frac{t}{5}} dt$

$= e^{-2}$

易知  $Y \sim B(5, e^{-2})$



Y	0	1	2	3	4	5
P	$(1-e^{-2})^5$	$5e^{-2}(1-e^{-2})^4$	$10e^{-4}(1-e^{-2})^3$	$10e^{-6}(1-e^{-2})^2$	$5e^{-8}(1-e^{-2})$	$e^{-10}$

$P(Y \geq 1) = 1 - P(Y=0) = 1 - (1-e^{-2})^5$

25.  $\Delta = 16k^2 - 16(k+2) = 16(k^2 - k - 2) = 16(k-2)(k+1) \geq 0 \quad k \leq -1 \text{ 或 } k \geq 2$

有实根概率  $P = \frac{3}{5}$

P15 18.  $E(X) = \int_{-\infty}^{+\infty} t f(t) dt = \int_0^{+\infty} \frac{t^2}{6^2} e^{-\frac{t^2}{36}} dt = \int_0^{+\infty} t d(e^{-\frac{t^2}{36}})$

$= - \int_0^{+\infty} e^{-\frac{t^2}{36}} dt - t e^{-\frac{t^2}{36}} \Big|_0^{+\infty} = \sqrt{2\pi} \cdot 6 \int_0^{+\infty} \frac{1}{\sqrt{2\pi} \cdot 6} e^{-\frac{t^2}{36}} dt = \frac{\sqrt{2\pi} \cdot 6}{2}$

同理  $E(X^2) = \int_{-\infty}^{+\infty} \frac{t^3}{6^2} e^{-\frac{t^2}{36}} dt = - \int_0^{+\infty} t^2 d(e^{-\frac{t^2}{36}}) = \int_0^{+\infty} e^{-\frac{t^2}{36}} 2t dt - t^2 e^{-\frac{t^2}{36}} \Big|_0^{+\infty}$

$= -26^2 e^{-\frac{t^2}{36}} \Big|_0^{+\infty} = 26^2$

$D(X) = E(X^2) - (E(X))^2 = \left(2 - \frac{\sqrt{2\pi}}{2}\right) 6^2$



4.2. 设长方形的宽为  $X$  米  $X \sim U(0,2)$

$$f(x) = \begin{cases} \frac{1}{2}, & x \in [0,2] \\ 0, & \text{其它} \end{cases}$$

$x \in [0,2]$  时,  $F(x) = \int_{-\infty}^x f(t) dt = \int_0^x \frac{1}{2} dt = \frac{1}{2}x$ .

长方形周长为  $2(X + \frac{10}{X})$  米

$$E(2X + \frac{20}{X}) = 2E(X) + 20E(\frac{1}{X})$$

$$= 2 \cdot \frac{0+2}{2} + 20 \cdot \int_{-\infty}^{+\infty} \frac{1}{t} \cdot \frac{1}{2} dt = 2 + 20 \cdot \frac{1}{2} \cdot \ln t \Big|_0^{+\infty}$$

~~$E(2X + \frac{20}{X})$~~  故期望不存在, 所以方差也不存在

~~$$\text{Var}(X) = E \text{Var}(2X + \frac{20}{X}) = E(2X + \frac{20}{X})^2 - (E(2X + \frac{20}{X}))^2 =$$~~

4.3.  $E(Y) = E(e^{-2X}) = \int_{-\infty}^{+\infty} e^{-2t} \cdot A e^{-t} dt$

由概率密度的规范性可得  $\int_{-\infty}^{+\infty} A e^{-t} dt = -A e^{-t} \Big|_0^{+\infty} = A = 1$ .

$$E(Y) = \int_0^{+\infty} e^{-2t} e^{-t} dt = \int_0^{+\infty} e^{-3t} dt = -\frac{1}{3} e^{-3t} \Big|_0^{+\infty} = \frac{1}{3}$$

4.4. 证:  $\left( \int_{-\infty}^{+\infty} e^{-\frac{x^2}{2}} dx \right)^2 = \int_{-\infty}^{+\infty} e^{-\frac{x^2}{2}} dx \int_{-\infty}^{+\infty} e^{-\frac{y^2}{2}} dy$

$$= \int_{-\infty}^{+\infty} e^{-\frac{x^2+y^2}{2}} dx dy$$

$$\begin{matrix} x=r\cos\theta \\ y=r\sin\theta \end{matrix} \quad = \int_0^{2\pi} d\theta \int_0^{+\infty} e^{-\frac{r^2}{2}} r dr$$

$$= \int_0^{2\pi} d\theta \int_0^{+\infty} e^{-\frac{r^2}{2}} d\frac{r^2}{2}$$

$$= 2\pi$$

所以  $\int_{-\infty}^{+\infty} e^{-\frac{x^2}{2}} dx = \sqrt{2\pi}$

令  $y = X = \frac{t-\mu}{\sigma}$   $t \in (-\infty, +\infty)$

则有  $\int_{-\infty}^{+\infty} e^{-\frac{(t-\mu)^2}{2\sigma^2}} d\left(\frac{t-\mu}{\sigma}\right) = \sqrt{2\pi}$

$$\int_{-\infty}^{+\infty} e^{-\frac{(t-\mu)^2}{2\sigma^2}} dt = \sqrt{2\pi} \sigma.$$

