

第一次作业

1. 设 X_1, X_2, \dots, X_m 和 X'_1, X'_2, \dots, X'_m 分别为来自总体 $X \sim N(u, \sigma^2)$ 的两个样本
根据正态分布的性质有

$$\bar{X}_1 = \frac{1}{m} \sum_{i=1}^m X_i \sim N(u, \frac{\sigma^2}{m}), \quad \bar{X}_2 = \frac{1}{n} \sum_{i=1}^n X'_i \sim N(u, \frac{\sigma^2}{n})$$

进一步根据正态分布的性质有:

$$\bar{X}_1 - \bar{X}_2 \sim N(0, \frac{\sigma^2}{m} + \frac{\sigma^2}{n})$$

$$\text{于是可得 } \Pr(|\bar{X}_1 - \bar{X}_2| > \epsilon) = 2 - 2\Phi(\epsilon / \sqrt{\frac{\sigma^2}{m} + \frac{\sigma^2}{n}})$$

2. 设随机变量 $Z = X + Y$

$$\begin{aligned} \text{有 } f_Z(z) &= \int_{-\infty}^{+\infty} f_X(x) f_Y(z-x) dx \\ &= \int_0^z \frac{\lambda^{\alpha_1}}{\Gamma(\alpha_1)} x^{\alpha_1-1} e^{-\lambda x} \frac{\lambda^{\alpha_2}}{\Gamma(\alpha_2)} (z-x)^{\alpha_2-1} e^{-\lambda(z-x)} dx \\ &= \frac{\lambda^{\alpha_1+\alpha_2}}{\Gamma(\alpha_1)\Gamma(\alpha_2)} e^{-\lambda z} \int_0^z x^{\alpha_1-1} (z-x)^{\alpha_2-1} dx \end{aligned}$$

今变量替换 $x = zt$ 有

$$\int_0^z x^{\alpha_1-1} (z-x)^{\alpha_2-1} dx = z^{\alpha_1+\alpha_2-1} \int_0^1 t^{\alpha_1-1} (1-t)^{\alpha_2-1} dt = z^{\alpha_1+\alpha_2-1} B(\alpha_1, \alpha_2)$$

利用Beta函数的性质有:

$$B(\alpha_1, \alpha_2) = \int_0^1 t^{\alpha_1-1} (1-t)^{\alpha_2-1} dt = \frac{\Gamma(\alpha_1)\Gamma(\alpha_2)}{\Gamma(\alpha_1+\alpha_2)}$$

代入完成证明

3. 设随机变量 $Y = X^2$

当 $y \leq 0$ 时分布函数 $F_Y(y) = 0$

当 $y > 0$ 时

$$\begin{aligned} F_Y(y) &= P(X^2 \leq y) = P(-\sqrt{y} \leq X \leq \sqrt{y}) = \int_{-\sqrt{y}}^{\sqrt{y}} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx \\ &= \frac{1}{\sqrt{2\pi}} \int_{-\sqrt{y}}^{\sqrt{y}} e^{-\frac{x^2}{2}} dx \\ &= \frac{1}{\sqrt{2\pi}} \times 2 \int_0^{\sqrt{y}} e^{-\frac{x^2}{2}} dx \\ &= \frac{1}{\sqrt{2\pi}} \times \frac{1}{\sqrt{y}} e^{-\frac{y}{2}} \end{aligned}$$

$$\Rightarrow X^2 \sim \Gamma(\frac{1}{2}, \frac{1}{2})$$

当 k 为偶数时

$$\begin{aligned} E(X^k) &= \int_{-\infty}^{+\infty} \frac{x^k}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx = - \int_{-\infty}^{+\infty} \frac{x^{k-1}}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx = (k-1) \int_{-\infty}^{+\infty} \frac{x^{k-2}}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx \\ &= -(k-1) \int_{-\infty}^{+\infty} \frac{x^{k-3}}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx = (k-1)(k-3) \frac{x^{k-4}}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx \\ &= \dots = (k-1)(k-3)\dots 1 \int_{-\infty}^{+\infty} \frac{x^2}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx = (k-1)!! \end{aligned}$$

当 k 为奇数时

$$E(X^k) = \int_{-\infty}^{+\infty} \frac{x^k}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx = \dots = (k-1)!! \int_{-\infty}^{+\infty} \frac{x}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx$$

由 $(-\infty, +\infty)$ 关于 y 轴对称, 且 $\frac{x}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}$ 为奇函数
 $\Rightarrow E(X^k) = 0.$

$$\Rightarrow E(X^k) = \begin{cases} (k-1)!! & k \text{ 为偶数} \\ 0 & k \text{ 为奇数} \end{cases}$$

4. 根据正态分布的性质有:

$$X_1 + 2X_2 + \dots + nX_n \sim N(0, a^2 + 4a^2 + \dots + n^2 a^2)$$

$$\text{即 } X_1 + 2X_2 + \dots + nX_n \sim N(0, \frac{n(n+1)(2n+1)}{6} a^2) \quad \text{令 } 1 \sim n \text{ 的平方和为 } t$$

$$\text{即 } \frac{X_1 + 2X_2 + \dots + nX_n}{\sqrt{t} a} \sim N(0, 1).$$

$$\text{同理可得 } \frac{Y_1 + 2Y_2 + \dots + mY_m}{\sqrt{S} a} \sim N(0, 1), \quad 1 \sim m \text{ 的平方和为 } S.$$

$$\Rightarrow \text{当 } a = \frac{1}{n(n+1)(2n+1)a^2}, b = \frac{1}{m(m+1)(2m+1)a^2} \text{ 时服从 } \chi^2 \text{ 分布}$$

且自由度为 2, 即 $Y \sim \chi^2(2)$ 成立.

$\Rightarrow Y$ 的概率密度函数为

$$f_Y(y) = \begin{cases} \frac{1}{\Gamma(1)} y^0 e^{-\frac{y}{2}}, & y > 0 \\ 0, & y \leq 0. \end{cases} \quad \text{即 } f_Y(y) = \begin{cases} \frac{1}{2} e^{-\frac{y}{2}}, & y > 0 \\ 0, & y \leq 0. \end{cases}$$

5. Y_1, \dots, Y_n 是 $X \sim N(0, n)$ 的样本, 由正态分布的性质得

$$\Rightarrow \frac{Y_1}{\sqrt{n}}, \dots, \frac{Y_n}{\sqrt{n}} \sim N(0, 1).$$

$$\Rightarrow \sum_{i=1}^n \left(\frac{Y_i}{\sqrt{n}} \right)^2 = \frac{1}{n} \sum_{i=1}^n Y_i^2 \sim \chi^2(n)$$

X_1, \dots, X_n 是 $X \sim N(0, n)$ 的样本, 由正态分布的性质得

$$\Rightarrow \frac{X_1}{\sqrt{n}}, \dots, \frac{X_n}{\sqrt{n}} \text{ 是 } X' \sim N(0, 1) \text{ 的样本}$$

$$\Rightarrow \sum_{i=1}^n \frac{X_i}{\sqrt{n}} \sim N(0, 1). \text{ 即 } \frac{1}{\sqrt{n}} \sum_{i=1}^n X_i \sim N(0, 1).$$

$$\text{令随机变量 } T = \frac{\sum_{i=1}^n X_i}{\sqrt{\sum_{i=1}^n Y_i^2}} \text{ 也即 } T = \frac{\frac{1}{\sqrt{n}} \sum_{i=1}^n X_i}{\sqrt{\frac{1}{n} \sum_{i=1}^n Y_i^2}}$$

$$\Rightarrow T \sim t(n)$$

b. 由 $\frac{(n-1)S^2}{\sigma^2} \sim \chi^2(n-1)$

而 $\sigma^2 = \frac{1}{4}$

$\Rightarrow 4(n-1)S^2 \sim \chi^2(n-1)$

由 $S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2$

$\Rightarrow 4 \sum_{i=1}^n (X_i - \bar{X})^2$

$\Rightarrow \Pr\left[\sum_{i=1}^n (X_i - \bar{X})^2 \geq \varepsilon\right]$

$= \Pr\left[4 \sum_{i=1}^n (X_i - \bar{X})^2 \geq 4\varepsilon\right]$

$= 1 - \Pr\left[4 \sum_{i=1}^n (X_i - \bar{X})^2 \leq 4\varepsilon\right]$

令 $4 \sum_{i=1}^n (X_i - \bar{X})^2 = Y$, 分布函数为 $F(y) = \int_0^{+\infty} \frac{(\frac{1}{2})^{\frac{n}{2}}}{\Gamma(\frac{n}{2})} y^{\frac{n}{2}-1} e^{-\frac{y}{2}}$

$\Rightarrow \Pr\left[\sum_{i=1}^n (X_i - \bar{X})^2 \geq \varepsilon\right] = 1 - F(4\varepsilon)$

7. X_1, \dots, X_n 是 $X \sim N(12, \sigma^2)$ 的样本. 由正态分布的性质

$\Rightarrow \frac{\sum_{i=1}^n (X_i - 12)}{\sigma} \sim N(0, 1)$

而又有 $\frac{(n-1)S^2}{\sigma^2} \sim \chi^2(n-1)$

\Rightarrow 有随机变量 $T = \frac{\frac{\sum_{i=1}^n (X_i - 12)}{\sigma}}{\sqrt{\frac{(n-1)S^2}{(n-1)\sigma^2}}} = \frac{\frac{1}{n} \sum_{i=1}^n (X_i - 12)}{\frac{S}{\sigma}} = \frac{\frac{1}{n} \sum_{i=1}^n X_i - 12}{\frac{S}{\sigma}} \sim t(n)$

$\Rightarrow \Pr\left[\frac{1}{n} \sum_{i=1}^n X_i \geq \varepsilon\right] = \Pr\left[T \geq \frac{\varepsilon - 12}{\frac{S}{\sigma}}\right]$
 $= 1 - \Pr\left[T \leq \frac{\varepsilon - 12}{\frac{S}{\sigma}}\right]$

令 T 的分布函数为 $F(t)$

$1 - \Pr\left[T \leq \frac{\varepsilon - 12}{\frac{S}{\sigma}}\right] = 1 - F\left(\frac{\varepsilon - 12}{\frac{S}{\sigma}}\right)$

8. 书147-148页 2(1), 4(1), 7, 9

2(1).

$$\begin{aligned}
 & P\{\max(X_1, X_2, X_3, X_4, X_5) > 15\} \\
 &= 1 - P\{\max(X_1, X_2, X_3, X_4, X_5) \leq 15\} \\
 &= 1 - P(X_1 \leq 15, X_2 \leq 15, X_3 \leq 15, X_4 \leq 15, X_5 \leq 15) \\
 &= 1 - \prod_{i=1}^5 P(X_i \leq 15) \\
 &= 1 - \prod_{i=1}^5 P\left\{\frac{X_i - 10}{2} \geq \frac{10 - 15}{2}\right\} \\
 &= 1 - \prod_{i=1}^5 P\left\{\frac{X_i - 10}{2} \geq -1\right\} \\
 &= 1 - [1 - \Phi(-1)]^5 \\
 &= 1 - [\Phi(1)]^5 = 0.5785.
 \end{aligned}$$

4(1)

X_1, X_2, \dots, X_5 是总体 $X \sim N(0, 1)$ 的样本

$$\begin{aligned}
 & \Rightarrow X_1 + X_2 \sim N(0, 2), \quad X_3^2 + X_4^2 + X_5^2 \sim \chi^2(3) \\
 & \Rightarrow \frac{X_1 + X_2}{\sqrt{2}} \sim N(0, 1) \\
 & \text{又由 } X_1 + X_2, X_3^2 + X_4^2 + X_5^2 \text{ 相互独立} \\
 & \Rightarrow \frac{\frac{X_1 + X_2}{\sqrt{2}}}{\sqrt{\frac{X_3^2 + X_4^2 + X_5^2}{3}}} \sim t(3) \\
 & \Rightarrow \sqrt{\frac{3}{2}} \frac{X_1 + X_2}{\sqrt{X_3^2 + X_4^2 + X_5^2}} \sim t(3) \\
 & \Rightarrow C = \sqrt{\frac{3}{2}}
 \end{aligned}$$

7. $X \sim \chi^2(n)$

$$\Rightarrow E(X) = n, D(X) = 2n$$

$$\text{设 } \bar{X} = \frac{1}{10} \sum_{i=1}^{10} X_i$$

$$\Rightarrow E(\bar{X}) = E\left(\frac{1}{10} \sum_{i=1}^{10} X_i\right) = \frac{1}{10} \sum_{i=1}^{10} E(X_i) = E(X) = n$$

$$D(\bar{X}) = D\left(\frac{1}{10} \sum_{i=1}^{10} X_i\right) = \frac{1}{100} \sum_{i=1}^{10} D(X_i) = \frac{D(X)}{10} = \frac{n}{5}$$

$$E(S^2) = E\left[\frac{1}{9} \sum_{i=1}^{10} (X_i - \bar{X})^2\right] = D(X) = 2n$$

9(1)

由题意

$$\frac{(n-1)S^2}{\sigma^2} \sim \chi^2(n-1)$$

当 $n=16$ 时

$$\frac{15S^2}{\sigma^2} \sim \chi^2(15)$$

$$\begin{aligned} \Rightarrow P\left\{\frac{S^2}{\sigma^2} \leq 2.041\right\} &= P\left\{\frac{15S^2}{\sigma^2} \leq 2.041 \times 15\right\} \\ &= P\left\{\frac{15S^2}{\sigma^2} \leq 30.615\right\} \\ &= 1 - P\left\{\frac{15S^2}{\sigma^2} > 30.615\right\} \\ &= 0.99 \end{aligned}$$

12) 由卡方分布的性质

$$D\left[\frac{(n-1)S^2}{\sigma^2}\right] = \frac{(n-1)^2}{\sigma^4} D(S^2) = 2(n-1)$$

$$\begin{aligned} \Rightarrow D(S^2) &= 2(n-1) \times \frac{\sigma^4}{(n-1)^2} \\ &= \frac{2\sigma^4}{n-1} \end{aligned}$$

$$\text{当 } n=16 \text{ 时 } D(S^2) = \frac{2}{15} \sigma^4$$