

My Solutions for Exercises of Deep Learning Fundamentals by Bishop

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May 15, 2024

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1 The Deep Learning Revolution

1.1 No Exercises

2 Probabilities

Exercise 2.1

Bayes rule

$$P[C = 1|T = 1] = \frac{P[T = 1|C = 1] * P[C = 1]}{P[T = 1|C = 1] * P[C = 1] + P[T = 1|C = 0] * P[C = 0]} \quad (1)$$

$$P[C = 1|T = 1] = \frac{0.90 * 0.001}{0.90 * 0.001 + 0.03 * 0.999} = 0.0292 \quad (2)$$

Given that the test result was positive, there is a 2.92% chance that you have cancer.

2.1 Exercise 2.2

Not attempted

Exercise 2.3

$$p(\mathbf{y}) = \int p_{\mathbf{u}, \mathbf{v}}(\mathbf{u}, \mathbf{y} - \mathbf{u}) d\mathbf{u} \quad (3)$$

$$= \int p_{\mathbf{u}}(\mathbf{u}) p_{\mathbf{v}}(\mathbf{y} - \mathbf{u}) d\mathbf{u} \quad (4)$$

2.2 Exercise 2.4

Not attempted

Exercise 2.5

Exponential:

$$p(x|\lambda) = \lambda e^{-\lambda x} \quad (5)$$

Laplace:

$$p(x|\mu, \gamma) = \frac{1}{2\gamma} e^{-\frac{|x-\mu|}{\gamma}} \quad (6)$$

Verifying that the exponential distribution is normalized:

$$p(x|\lambda) = \lambda e^{-\lambda x} \quad (7)$$

$$\int_0^{\infty} \lambda e^{-\lambda x} = -e^{-\lambda x} \Big|_0^{\infty} = \frac{1}{e^{\infty}} + \frac{1}{e^0} \quad (8)$$

$$= 1 \quad (9)$$

Verifying the laplace distribution:

$$p(x|\mu, \gamma) = \frac{1}{2\gamma} e^{-\frac{|x-\mu|}{\gamma}} \quad (10)$$

$$\begin{cases} \frac{1}{2\gamma} e^{-\frac{x-\mu}{\gamma}} & \text{if } x \geq \mu \\ \frac{1}{2\gamma} e^{-\frac{-x+\mu}{\gamma}} & \text{if } x < \mu \end{cases} \quad (11)$$

$$\int_{\mu}^{\infty} \frac{1}{2\gamma} e^{-\frac{x-\mu}{\gamma}} = \frac{1}{2} e^{-\frac{x-\mu}{\gamma}} \Big|_{\mu}^{\infty} = -\frac{1}{2} (e^{-\infty} - e^0) \quad (12)$$

$$= \frac{1}{2} \quad (13)$$

$$\int_{-\infty}^{\mu} \frac{1}{2\gamma} e^{-\frac{-x+\mu}{\gamma}} = \frac{1}{2} e^{-\frac{-x+\mu}{\gamma}} \Big|_{-\infty}^{\mu} = \frac{1}{2} (e^0 - e^{-\infty}) \quad (14)$$

$$= \frac{1}{2} \quad (15)$$

$$\frac{1}{2} + \frac{1}{2} = 1 \quad (16)$$

2.3 Exercise 2.6

Not attempted

2.4 Exercise 2.7

$$P(x|D) = \frac{1}{N} \sum_{n=1}^N \delta(x - x_n) \quad (17)$$

$$E[f] = \int p(x) f(x) dx \quad (18)$$

$$\text{Substituting:} \quad (19)$$

$$E[f] = \int \frac{1}{N} \sum_{n=1}^N \delta(x - x_n) f(x) dx \quad (20)$$

$$E[f] = \frac{1}{N} \sum_{n=1}^N \int_{x_n-\varepsilon}^{x_n+\varepsilon} \delta(x - x_n) f(x) dx \quad (21)$$

$$E[f] = \frac{1}{N} \sum_{n=1}^N f(x_n) \int_{x_n-\varepsilon}^{x_n+\varepsilon} \delta(x - x_n) dx \quad (22)$$

$$E[f] = \frac{1}{N} \sum_{n=1}^N f(x_n) \quad (23)$$

2.5 Exercise 2.8

Not attempted

2.6 Exercise 2.9

$$\text{cov}[X, y] = E_{x,y}[xy] - E[x]E[y] \quad (24)$$

If x and y are independent, the joint distribution is equal to the product of the marginals. $p(x, y) = p(x)p(y)$. If $E_{x,y}[xy] = E[x]E[y]$, then the covariance will be zero.

2.7 Exercise 2.10

Not attempted

2.8 Exercise 2.11

Proving $E[x] = E_y[E_x[x|y]]$:

$$E_x[x|y] = \int p(x|y)xdx \quad (25)$$

$$\text{Substituting:} \quad (26)$$

$$E[x] = E_y[\int p(x|y)xdx] \quad (27)$$

$$E[x] = \int E_y[p(x|y)]xdx \quad (28)$$

$$E[x] = \int \int p(x|y)xp(y)dxdy \quad (29)$$

$$E[x] = \int \int \frac{p(x,y)}{p(y)}xp(y)dxdy \quad (30)$$

$$E[x] = \int \int p(x,y)dxdy \quad (31)$$

$$E[x] = E[x] \quad (32)$$

2.9 Exercise 2.12

Not attempted

2.10 Exercise 2.13

$$E[x] = \int_{-\infty}^{\infty} x \frac{1}{\sqrt{2\pi}\sigma^2} e^{\frac{1}{2\sigma^2}(x-\mu)^2} dx \quad (33)$$

$$\text{Change of variables } z = \frac{x-\mu}{\sigma}, \sigma dz = dx \quad (34)$$

$$E[x] = \int_{-\infty}^{\infty} \sigma \frac{\sigma z + \mu}{\sqrt{2\pi}\sigma^2} e^{\frac{1}{2}z^2} dz \quad (35)$$

$$E[x] = \frac{\sigma}{\sqrt{2\pi}} \int_{-\infty}^{\infty} ze^{\frac{1}{2}z^2} + \frac{\mu}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{\frac{1}{2}z^2} \quad (36)$$

$$E[x] = \frac{\sigma}{\sqrt{2\pi}} * 0 + \frac{\mu}{\sqrt{2\pi}} * \sqrt{2\pi} \quad (37)$$

$$E[x] = \mu \quad (38)$$

2.11 Exercise 2.14

Not attempted

2.12 Exercise 2.15

Solving for μ_{ml} :

$$\log p(x|\mu, \sigma^2) = \frac{-1}{2\sigma^2} \sum_{n=1}^N (x_n - \mu)^2 - \frac{N}{2} \log \sigma^2 - \frac{N}{2} \log 2\pi \quad (39)$$

$$\frac{d}{d\mu} \log p(x|\mu, \sigma^2) = \frac{1}{2\sigma^2} \sum_{n=1}^N 2(x_n - \mu) \quad (40)$$

$$0 = \frac{1}{\sigma^2} \sum_{n=1}^N (x_n - \mu) \quad (41)$$

$$0 = \sum_{n=1}^N x_n - \sum_{n=1}^N \mu \quad (42)$$

$$N\mu = \sum_{n=1}^N x_n \quad (43)$$

$$\mu_{ml} = \frac{1}{N} \sum_{n=1}^N x_n \quad (44)$$

Solving for σ_{ml} :

$$\log p(x|\mu, \sigma^2) = \frac{-1}{2\sigma^2} \sum_{n=1}^N (x_n - \mu)^2 - \frac{N}{2} \log \sigma^2 - \frac{N}{2} \log 2\pi \quad (45)$$

$$\frac{d}{d\sigma^2} \log p(x|\mu, \sigma^2) = \frac{1}{2\sigma^4} \sum_{n=1}^N (x_n - \mu)^2 - \frac{N}{2\sigma^2} \quad (46)$$

$$\frac{N}{2\sigma^2} = \frac{1}{2\sigma^4} \sum_{n=1}^N (x_n - \mu)^2 \quad (47)$$

$$\sigma_{ml}^2 = \frac{1}{N} \sum_{n=1}^N (x_n - \mu_{ml})^2 \quad (48)$$

2.13 Exercise 2.16

not attempted

2.14 Exercise 2.17

Finding expectation of $\hat{\sigma}^2$

$$E[\hat{\sigma}^2] = E\left[\frac{1}{N} \sum_{n=1}^N (x_n - \mu)^2\right] \quad (49)$$

$$= \frac{1}{N} \sum_{n=1}^N E[x_n^2 - 2x_n\mu + \mu^2] \quad (50)$$

$$= \frac{1}{N} \sum_{n=1}^N E[x_n^2] - E[2x_n\mu] + E[\mu^2] \quad (51)$$

$$= \frac{1}{N} \sum_{n=1}^N E[x_n^2] - 2E[x_n]E[\mu] + E[\mu^2] \quad (52)$$

$$= \frac{1}{N} \sum_{n=1}^N E[x_n^2] - E[x_n]^2 \quad (53)$$

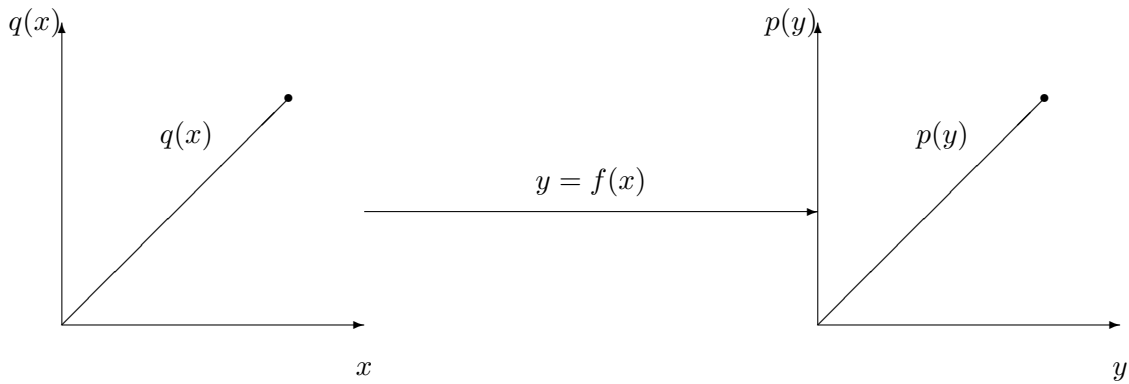
$$= \frac{1}{N} \sum_{n=1}^N \mu^2 + \sigma^2 - \mu^2 \quad (54)$$

$$= \sigma^2 \quad (55)$$

2.15 Exercise 2.18

Not attempted

2.16 Exercise 2.19



2.17 Exercise 2.20

Not attempted

2.18 Exercise 2.21

Showing $h(p^2) = 2h(p)$:

$$h(p) = h(p(x_1)) + h(p(x_2)) + \cdots + h(p(x_n)) \quad (56)$$

$$h(p^2) = h(x_1^2) + h(x_2^2) + \cdots + h(x_n^2) \quad (57)$$

$$\because h(x) = -\log_2 p(x), \quad (58)$$

$$h(p^2) = 2h(x_1) + 2h(x_2) + \cdots + 2h(x_n) \quad (59)$$

$$h(p^2) = 2h(p) \quad (60)$$

This can be applied to any exponent which includes any choice of n integer or $\frac{n}{m}$ positive rational number.

$$h(p^x) = xh(p) \quad \forall Q^+ \quad (61)$$

$$\therefore \quad (62)$$

$$h(p) \propto \ln p \quad (63)$$

2.19 Exercise 2.22

Not attempted

2.20 Exercise 2.23

Not attempted

2.21 Exercise 2.24

Not attempted

2.22 Exercise 2.25

Not attempted

2.23 Exercise 2.26

Not attempted

2.24 Exercise 2.27

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2.25 Exercise 2.28

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2.26 Exercise 2.29

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2.32 Exercise 2.35

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2.33 Exercise 2.36

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2.34 Exercise 2.37

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2.35 Exercise 2.38

Not attempted

2.36 Exercise 2.39

Not attempted

2.37 Exercise 2.40

Not attempted

2.38 Exercise 2.41

Not attempted

3 Standard Distributions

3.1 Exercise 3.1

Proving $\sum_{x=0}^1 p(x|\mu) = 1$:

$$\text{Bern}(x|\mu) = \mu^x(1-\mu)^{1-x} \quad (64)$$

$$\sum_{x=0}^1 p(x|\mu) = \mu^0(1-\mu)^1 + \mu^1(1-\mu)^0 \quad (65)$$

$$= 1 - \mu + \mu \quad (66)$$

$$= 1 \quad (67)$$

Proving $E[x] = \mu$:

$$E[x] = \sum_{x=0}^1 p(x|\mu) * x \quad (68)$$

$$E[x] = 0 * \mu^0(1-\mu)^1 + 1 * \mu^1(1-\mu)^0 \quad (69)$$

$$= 0 * (1-\mu) + 1 * \mu \quad (70)$$

$$= \mu \quad (71)$$

Proving $\text{var}[x] = \mu(1-\mu)$:

$$\text{Var}[x] = \sum_i (x_i - \mu)^2 p(x_i) \quad (72)$$

$$= (0 - \mu)^2(1-\mu) + (1 - \mu)^2\mu \quad (73)$$

$$= \mu^2(1-\mu) + (1-\mu)^2\mu \quad (74)$$

$$= \mu(1-\mu)(\mu + (1-\mu)) \quad (75)$$

$$= \mu(1-\mu)(1) \quad (76)$$

$$= \mu(1-\mu) \quad (77)$$

Solving for entropy of the bernoulli distribution proves simple as there are two probabilities, $p_1 = \mu, p_2 = 1 - \mu$. Therefore the entropy must be:

$$H[x] = -\mu \ln \mu - (1-\mu) \ln (1-\mu) \quad (78)$$

3.2 Exercise 3.2

Not attempted

3.3 Exercise 3.3

Normalizing the binomial distribution $\binom{N}{m}\mu^m(1-\mu)^{N-m}$:

$$\text{First step is proving } \binom{N}{m} + \binom{N}{m-1} = \binom{N+1}{m} : \quad (79)$$

$$\binom{N}{m} + \binom{N}{m-1} = \frac{N!}{(N-m)!m!} + \frac{N!}{(N-m+1)!(m-1)!} \quad (80)$$

$$= \frac{N!}{(N-m)!m!} * \frac{N-m+1}{N-m+1} + \frac{N!}{(N-m+1)!(m-1)!} * \frac{m}{m} \quad (81)$$

$$= \frac{N!(N+1-m)}{(N+1-m)!m!} + \frac{N!m}{(N-m+1)!m!} \quad (82)$$

$$= \frac{(N+1)! - N!m + N!m}{(N+1-m)!m!} \quad (83)$$

$$= \frac{(N+1)!}{(N+1-m)!m!} \quad (84)$$

$$= \binom{N+1}{m} \quad (85)$$

With this information we would like to prove that $(1+x)^N = \sum_{m=0}^N \binom{N}{m}x^m$
N=2:

$$\binom{2}{0}x^0 + \binom{2}{1}x^1 + \binom{2}{2}x^2 = 1 + 2x + x^2 = (1+x)^2 \quad (86)$$

N=3:

$$\binom{3}{0}x^0 + \binom{3}{1}x^1 + \binom{3}{2}x^2 + \binom{3}{3}x^3 = 1 + 3x + 3x^2 + x^3 = (1+x)^3 \quad (87)$$

N=4

$$\binom{4}{0}x^0 + \binom{4}{1}x^1 + \binom{4}{2}x^2 + \binom{4}{3}x^3 + \binom{4}{4}x^4 = 1 + 4x + 6x^2 + 4x^3 + x^4 = (1+x)^4 \quad (88)$$

By induction: $\forall x \in \mathbb{R}$ and $N \in \mathbb{N}$

$$(1+x)^N = \sum_{m=0}^N \binom{N}{m}x^m \quad (89)$$

Normalizing the binomial distribution:

$$\sum_{m=0}^N \binom{N}{m} \mu^m (1-\mu)^{N-m} = 1 \quad (90)$$

$$= (1-\mu)^N \sum_{m=0}^N \binom{N}{m} \mu^m (1-\mu)^{-m} \quad (91)$$

$$= (1-\mu)^N \sum_{m=0}^N \binom{N}{m} \left(\frac{\mu}{1-\mu}\right)^m \quad (92)$$

$$= (1-\mu)^N \left(1 + \frac{\mu}{1-\mu}\right)^N \quad (93)$$

$$= ((1-\mu)(1 + \frac{\mu}{1-\mu}))^N \quad (94)$$

$$= (1 + \frac{\mu}{1-\mu} - \mu - \frac{\mu^2}{1-\mu}) \quad (95)$$

$$= ((1-\mu) + \frac{\mu}{1-\mu}(1-\mu))^N \quad (96)$$

$$= (1-\mu + \mu)^N \quad (97)$$

$$= 1 \quad (98)$$

3.4 Exercise 3.4

Not attempted

3.5 Exercise 3.5

Finding the mode of the multivariate gaussian. The mode of a distribution is the same as the maximum of the probability density function. For the multivariate gaussian, this is:

$$\mathcal{N}(\mathbf{x}|\mu, \Sigma) = \frac{1}{(2\pi)^{D/2} |\Sigma|^{\frac{1}{2}}} e^{-\frac{1}{2}(\mathbf{x}-\mu)^T \Sigma^{-1}(\mathbf{x}-\mu)} \quad (99)$$

$$\frac{d}{d\mathbf{x}} \mathcal{N}(\mathbf{x}|\mu, \Sigma) \text{ ignoring normalization constant for brevity} \quad (100)$$

$$0 = -\frac{1}{2} e^{-\frac{1}{2}(\mathbf{x}-\mu)^T \Sigma^{-1}(\mathbf{x}-\mu)} * \frac{d}{d\mathbf{x}} (\mathbf{x}-\mu)^T \Sigma^{-1}(\mathbf{x}-\mu) \quad (101)$$

$$0 = -\frac{1}{2} e^{-\frac{1}{2}(\mathbf{x}-\mu)^T \Sigma^{-1}(\mathbf{x}-\mu)} * 2\Sigma^{-1}(\mathbf{x}-\mu) \quad (102)$$

$$\mathbf{x} = \mu \quad (103)$$

3.6 Exercise 3.6

Not attempted

3.7 Exercise 3.7

Finding Kullback-Leibler divergence between $q(\mathbf{x}) = \mathcal{N}(\mathbf{x}|\mu_{\mathbf{q}}, \Sigma_{\mathbf{q}})$ and $p(\mathbf{x}) = \mathcal{N}(\mathbf{x}|\mu_{\mathbf{p}}, \Sigma_{\mathbf{p}})$

$$KL(p||q) = - \int p(x) \ln \frac{q(x)}{p(x)} dx \quad (104)$$

$$KL(q(x)||p(x)) = - \int \mathcal{N}(\mathbf{x}|\mu_{\mathbf{q}}, \Sigma_{\mathbf{q}}) \ln \frac{\mathcal{N}(\mathbf{x}|\mu_{\mathbf{p}}, \Sigma_{\mathbf{p}})}{\mathcal{N}(\mathbf{x}|\mu_{\mathbf{q}}, \Sigma_{\mathbf{q}})} \quad (105)$$

$$= - \int \mathcal{N}(\mathbf{x}|\mu_{\mathbf{q}}, \Sigma_{\mathbf{q}}) (\ln \mathcal{N}(\mathbf{x}|\mu_{\mathbf{p}}, \Sigma_{\mathbf{p}}) - \ln \mathcal{N}(\mathbf{x}|\mu_{\mathbf{q}}, \Sigma_{\mathbf{q}})) \quad (106)$$

$$\begin{aligned} \text{definition: } \ln \mathcal{N}(\mathbf{x}|\mu_{\mathbf{p}}, \Sigma_{\mathbf{p}}) &= -\frac{D}{2} \ln(2\pi) - \frac{1}{2} \ln|\Sigma_{\mathbf{p}}| - \frac{1}{2} (\mathbf{x} - \mu_{\mathbf{p}})^T \Sigma_{\mathbf{p}}^{-1} (\mathbf{x} - \mu_{\mathbf{p}}) \\ &= -\frac{1}{2} \int \mathcal{N}(\mathbf{x}|\mu_{\mathbf{q}}, \Sigma_{\mathbf{q}}) (-\ln|\Sigma_{\mathbf{p}}| + \ln|\Sigma_{\mathbf{q}}| - (\mathbf{x} - \mu_{\mathbf{p}})^T \Sigma_{\mathbf{p}}^{-1} (\mathbf{x} - \mu_{\mathbf{p}}) + (\mathbf{x} - \mu_{\mathbf{q}})^T \Sigma_{\mathbf{q}}^{-1} (\mathbf{x} - \mu_{\mathbf{q}})) \end{aligned} \quad (107)$$

$$\text{Splitting integral into three} \quad (109)$$

$$\text{First: } -\frac{1}{2} \ln \frac{|\Sigma_{\mathbf{q}}|}{|\Sigma_{\mathbf{p}}|} \int \frac{1}{(2\pi)^{D/2} |\Sigma_{\mathbf{q}}|^{\frac{1}{2}}} e^{-\frac{1}{2} (\mathbf{x} - \mu_{\mathbf{q}})^T \Sigma_{\mathbf{q}}^{-1} (\mathbf{x} - \mu_{\mathbf{q}})} \quad (110)$$

$$\text{Second: } \frac{1}{2} \int \frac{1}{(2\pi)^{D/2} |\Sigma_{\mathbf{q}}|^{\frac{1}{2}}} e^{-\frac{1}{2} (\mathbf{x} - \mu_{\mathbf{q}})^T \Sigma_{\mathbf{q}}^{-1} (\mathbf{x} - \mu_{\mathbf{q}})} (\mathbf{x} - \mu_{\mathbf{p}})^T \Sigma_{\mathbf{p}}^{-1} (\mathbf{x} - \mu_{\mathbf{p}}) \quad (111)$$

$$\text{Third: } -\frac{1}{2} \int \frac{1}{(2\pi)^{D/2} |\Sigma_{\mathbf{q}}|^{\frac{1}{2}}} e^{-\frac{1}{2} (\mathbf{x} - \mu_{\mathbf{q}})^T \Sigma_{\mathbf{q}}^{-1} (\mathbf{x} - \mu_{\mathbf{q}})} (\mathbf{x} - \mu_{\mathbf{p}})^T \Sigma_{\mathbf{p}}^{-1} (\mathbf{x} - \mu_{\mathbf{p}}) \quad (112)$$

$$\text{Evaluating first integral: } -\frac{1}{2} \ln \frac{|\Sigma_{\mathbf{q}}|}{|\Sigma_{\mathbf{p}}|} * 1 \quad (113)$$

$$\text{Evaluating second integral: } \frac{1}{2} (\mu_{\mathbf{p}} - \mu_{\mathbf{q}})^T \Sigma_{\mathbf{p}}^{-1} (\mu_{\mathbf{p}} - \mu_{\mathbf{q}}) \quad (114)$$

$$\text{Evaluating third integral: } -\frac{1}{2} (\mu_{\mathbf{p}} - \mu_{\mathbf{q}})^T \Sigma_{\mathbf{p}}^{-1} (\mu_{\mathbf{p}} - \mu_{\mathbf{q}}) + \frac{1}{2} \text{Tr}(\Sigma_{\mathbf{p}}^{-1} \Sigma_{\mathbf{q}}) \quad (115)$$

I am not sure about this one it is quite hard

3.8 Exercise 3.8

Not attempted

3.9 Exercise 3.9

Not attempted

3.10 Exercise 3.10

Not attempted

3.11 Exercise 3.11

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3.12 Exercise 3.12

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3.34 Exercise 3.34

Not attempted

3.35 Exercise 3.35

Not attempted

3.36 Exercise 3.36

Not attempted

3.37 Exercise 3.37

Deriving a maximum likelihood estimator for a histogram-like density model. Space \mathbf{x} is divided into fixed regions with density $p(\mathbf{x}) = h_i \forall i$ regions. Volume i is denoted Δ_i . N total observations

such that n_i observations are in region i .

$$h_i = \frac{n_i}{N\Delta_i} \quad (116)$$

$$\sum_i h_i \Delta_i = 1 \quad (117)$$

$$L = \prod_{i=0}^K h_i^{n_i} \quad (118)$$

$$\log L = \prod_{i=0}^K n_i \log h_i \quad (119)$$

$$\max \sum_{i=0}^K n_i \log h_i - \lambda \left(\sum_i h_i \Delta_i - 1 \right) \quad (120)$$

$$0 = \frac{n_i}{h_i} - \lambda \Delta_i \quad (121)$$

$$\frac{n_i}{h_i} = \lambda \Delta_i \quad (122)$$

$$h_i = \frac{n_i}{\lambda \Delta_i} \quad (123)$$

$$\text{Plugging into normalization:} \quad (124)$$

$$\sum_i \frac{n_i}{\lambda} = 1 \quad (125)$$

$$\lambda = N \quad (126)$$

$$h_i = \frac{n_i}{N\Delta_i} \quad (127)$$

3.38 Exercise 3.38

Not attempted

4 Single-layer Networks: Regression

4.1 Exercise 4.1

Showing the coefficients $\mathbf{w} = \{w_i\}$ that minimize the error function $E[\mathbf{w}] = \frac{1}{2} \sum_{n=1}^N (y(x_n, \mathbf{w}) - t_n)^2$ are given by the solution to:

$$\sum_{j=0}^M A_{ij} w_j = T_i \quad (128)$$

where $A_{ij} = \sum_{n=1}^N x_n^{i+j}$, $T_i = \sum_{n=1}^N x_n^i t_n$ and $y(x, \mathbf{w}) = \sum_{j=0}^M w_j x^j$

$$\frac{1}{2} \sum_{n=1}^N (y(x_n, \mathbf{w}) - t_n)^2 = \frac{1}{2} \sum_{n=1}^N \left(\sum_{j=0}^M w_j x_n^j - t_n \right)^2 \quad (129)$$

$$\frac{d}{dw_i} = \sum_{n=1}^N \sum_{j=0}^M (w_j x_n^j - t_n) x_n^i \quad (130)$$

$$0 = \sum_{j=0}^M w_j \sum_{n=1}^N x_n^{i+j} - x_n^i t_n \quad (131)$$

$$\sum_{n=1}^N x_n^i t_n = \sum_{j=0}^M w_j \sum_{n=1}^N x_n^{i+j} \quad (132)$$

$$T_i = \sum_{j=0}^M A_{ij} w_j \quad (133)$$

4.2 Exercise 4.2

Not Attempted

4.3 Exercise 4.3

Showing that the \tanh function is related to the logistic sigmoid function by $\tanh(a) = 2\sigma(2a) - 1$

$$\tanh(a) = \frac{e^a - e^{-a}}{e^a + e^{-a}} \quad (134)$$

$$\sigma(a) = \frac{1}{1 + e^{-a}} \quad (135)$$

$$\frac{e^a - e^{-a}}{e^a + e^{-a}} + 1 = 2\sigma(2a) \quad (136)$$

$$\frac{\frac{e^{2a}-1}{e^a}}{\frac{e^{2a}+1}{e^a}} + 1 = 2\sigma(2a) \quad (137)$$

$$\frac{e^{2a}-1}{e^{2a}+1} + 1 = 2\sigma(2a) \quad (138)$$

$$\frac{2e^{2a}}{e^{2a}+1} = 2\sigma(2a) \quad (139)$$

$$\sigma(2a) = \frac{e^{2a}}{1+e^{2a}} \quad (140)$$

Showing that $w_0 + \sum_{j=1}^M w_j \sigma(\frac{x-\mu_j}{s})$ is a linear combination of $u_0 + \sum_{j=1}^M u_j \tanh(\frac{x-\mu_j}{2s})$

Defining $z_j = \frac{x-\mu_j}{s}$

$$u_0 + \sum_{j=1}^M u_j \tanh(\frac{z_j}{2}) \quad (141)$$

$$= u_0 + -1 + 2 \sum_{j=1}^M u_j \sigma(z_j) \quad (142)$$

$$y(x, \mathbf{u}) = u_0 - 1 - 2w_0 + 2y(x, \mathbf{w}) \quad (143)$$

4.4 Exercise 4.4

Not attempted

4.5 Exercise 4.5

The weighted sum of squares error function could attempt to weigh some data points more than others. This could tie to data-dependent noise variance by weighing the points that are outliers less than other points. The weighted part could also be a way to lower the weight of or completely negate the effect of duplicate data points. A statement could be set up where if a data point was duplicated, the weight will be set to 0.

Finding the solution \mathbf{w}^* that minimizes

$$E_D(\mathbf{w}) = \frac{1}{2} \sum_{n=1}^N r_n (t_n - \mathbf{w}^T \phi(\mathbf{x}_n))^2 \quad (144)$$

$$\frac{d}{d\mathbf{w}} = - \sum_{n=1}^N r_n \phi_n (t_n - \mathbf{w}^T \phi_n) \quad (145)$$

$$0 = \sum_{n=1}^N r_n \phi_n t_n - r_n \phi_n \mathbf{w}^T \phi_n \quad (146)$$

$$\sum_{n=1}^N \mathbf{w}^T \phi_n = \sum_{n=1}^N t_n \quad (147)$$

$$\mathbf{w} = (\Phi^T \Phi)^{-1} \Phi^T \mathbf{t} \quad (148)$$

4.6 Exercise 4.6

Not attempted

4.7 Exercise 4.7

Finding Σ_{ml} :

$$p(t|W, \Sigma) = \mathcal{N}(t|y(x, W), \Sigma) \quad (149)$$

$$y(x, W) = W^T \phi(x) \quad (150)$$

likelihood function of multivariate target variable t is:

$$\ln p(t|y(x, W), \Sigma) = -\frac{ND}{2} \ln(2\pi) - \frac{N}{2} \ln|\Sigma| - \frac{1}{2} \sum_{n=1}^N (t_n - y(x, W))^T \Sigma^{-1} (t_n - y(x, W)) \quad (151)$$

$$= -\frac{ND}{2} \ln(2\pi) - \frac{N}{2} \ln|\Sigma| - \frac{1}{2} \sum_{n=1}^N (t_n - W^T \phi(x))^T \Sigma^{-1} (t_n - W^T \phi(x)) \quad (152)$$

$$\frac{d}{d\Sigma} = -\frac{N}{2|\Sigma|} \frac{d}{d|\Sigma|} + \frac{1}{2} \sum_{n=1}^N (t_n - W^T \phi(x))^T \Sigma^{-1} I \Sigma^{-1} (t_n - W^T \phi(x)) \quad (153)$$

$$\frac{N}{2} \Sigma^{-1} = \Sigma^{-1} \frac{1}{2} \sum_{n=1}^N (t_n - W^T \phi(x)) (t_n - W^T \phi(x))^T \Sigma^{-1} \quad (154)$$

$$\Sigma = \frac{1}{N} \sum_{n=1}^N (t_n - W_{ML}^T \phi(x)) (t_n - W_{ML}^T \phi(x))^T \quad (155)$$

Important notes learned from this exercise: tr scalar = scalar and $X^T y X$ is a scalar for any dimension compatible matrices. Also, $\text{tr}(ABC) = \text{tr}(BCA) = \text{tr}(CAB)$

4.8 Exercise 4.8

Not attempted

4.9 Exercise 4.9

Finding $E[L(t, f(x))]$ for multiple target variables where:

$$E[L(t, f(x))] = \int \int ||f(x) - t||^2 p(x, t) dx dt \quad (156)$$

And showing that the function $f(x)$ that minimizes E is given by $E_t[t|x]$

$$E[L(t, f(x))] = \int \int ||f(x) - t||^2 p(x, t) dx dt \quad (157)$$

$$= \int \int ||f(x) - E[t|x] + E[t|x] - t||^2 p(x, t) dx dt \quad (158)$$

$$= \int \int ||f(x) - E[t|x]||^2 + 2(f(x) - E[t|x])(E[t|x] - t) + ||E[t|x] - t||^2 p(x, t) dx dt \quad (159)$$

$$\text{Integrating over } t \quad (160)$$

$$\int ||f(x) - E[t|x]||^2 p(x) dx + \int \text{var}(t|x) p(x) dx \quad (161)$$

$$\text{Minimizing this gives} \quad (162)$$

$$f(x) = E[t|x] \quad (163)$$

4.10 Exercise 4.10

Not attempted

4.11 Exercise 4.11

Normalizing

$$p(x|\sigma^2, q) = \frac{q}{2(2\sigma^2)^{\frac{1}{q}} \Gamma(\frac{1}{q})} e^{-\frac{|x|^q}{2\sigma^2}} \quad (164)$$

where

$$\Gamma(x) = \int_{-\infty}^{\infty} u^{x-1} e^{-u} du \quad (165)$$

$$\Gamma(\frac{1}{q}) = \int_{-\infty}^{\infty} u^{\frac{1}{q}-1} e^{-u} du \quad (166)$$

$$\Gamma(\frac{1}{q}) = \int_{-\infty}^{\infty} u^{\frac{1}{q}-1} e^{-u} du \quad (167)$$

4.12 Exercise 4.12

Not attempted