

# Probability and Statistics

## Solutions to Tutorial 5

Q1: (a) Marginal distribution of  $X$ :

$$p_X(x) = \sum_{y=1}^{\infty} \frac{1}{2^{x+y}} = \frac{1}{2^x} \sum_{y=1}^{\infty} \frac{1}{2^y} = \frac{1}{2^x} \cdot \frac{1/2}{1-1/2} = \frac{1}{2^x}$$

Similarly,  $p_Y(y) = \frac{1}{2^y}$  for  $y = 1, 2, 3, \dots$

Since  $p_{XY}(x, y) = \frac{1}{2^x} \cdot \frac{1}{2^y} = p_X(x) \cdot p_Y(y)$ ,  $X$  and  $Y$  are independent.

$$(b) \mathbb{P}(X^2 + Y \leq 10) = \sum_{\substack{x, y \geq 1 \\ x^2 + y \leq 10}} \frac{1}{2^{x+y}}$$

For  $x = 1$ :  $y \leq 9$ , so  $\sum_{y=1}^9 \frac{1}{2^{1+y}} = \frac{1}{2} \sum_{y=1}^9 \frac{1}{2^y} = \frac{1}{2} \cdot \frac{1/2 - (1/2)^{10}}{1 - 1/2} = 1 - \frac{1}{2^9}$

For  $x = 2$ :  $y \leq 6$ , so  $\sum_{y=1}^6 \frac{1}{2^{2+y}} = \frac{1}{4} \cdot \frac{1/2 - (1/2)^7}{1/2} = \frac{1}{4} - \frac{1}{2^8}$

For  $x = 3$ :  $y = 1$ , so  $\frac{1}{2^4} = \frac{1}{16}$

Therefore:  $\mathbb{P}(X^2 + Y \leq 10) = 1 - \frac{1}{2^9} + \frac{1}{4} - \frac{1}{2^8} + \frac{1}{16} = \frac{3}{4} + \frac{1}{16} - \frac{3}{2^9} = \frac{3}{4} + \frac{1}{16} - \frac{3}{512} = 0.813$

Q2: We are asked to compute the covariance  $\text{Cov}(Z, W)$ , where:

$$Z = 1 + X + XY^2, \quad W = 1 + X$$

Using the properties of covariance:

$$\begin{aligned} \text{Cov}(Z, W) &= \text{Cov}(1 + X + XY^2, 1 + X) \\ &= \text{Cov}(X + XY^2, X) \\ &= \text{Cov}(X, X) + \text{Cov}(XY^2, X) \\ &= \text{Var}(X) + \mathbb{E}[X^2Y^2] - \mathbb{E}[XY^2]\mathbb{E}[X] \end{aligned}$$

Since  $X$  and  $Y$  are independent, we have:

$$\text{Var}(X) = 1, \quad \mathbb{E}[X^2] = 1, \quad \mathbb{E}[Y^2] = 1, \quad \mathbb{E}[X] = 0$$

Thus:

$$\text{Cov}(Z, W) = 1 + \mathbb{E}[X^2]\mathbb{E}[Y^2] - 0 = 1 + 1 - 0 = 2$$

Therefore, the covariance is:

$$\text{Cov}(Z, W) = 2$$

Q3: (1) Compute the marginals:

$$f_X(x) = \int_0^{\infty} 6e^{-(2x+3y)} dy = 6e^{-2x} \int_0^{\infty} e^{-3y} dy = 6e^{-2x} \cdot \frac{1}{3} = 2e^{-2x}, \quad x \geq 0,$$

$$f_Y(y) = \int_0^{\infty} 6e^{-(2x+3y)} dx = 6e^{-3y} \int_0^{\infty} e^{-2x} dx = 6e^{-3y} \cdot \frac{1}{2} = 3e^{-3y}, \quad y \geq 0.$$

Since

$$f_X(x)f_Y(y) = (2e^{-2x})(3e^{-3y}) = 6e^{-(2x+3y)} = f_{X,Y}(x, y) \quad \text{for } x, y \geq 0,$$

we conclude that  $X$  and  $Y$  are **independent**.

(2) Because of independence,

$$\mathbb{E}[Y \mid X > 2] = \mathbb{E}[Y].$$

Noting that  $f_Y(y) = 3e^{-3y}$ , i.e.  $Y \sim \text{Exp}(3)$ , we have

$$\mathbb{E}[Y] = \int_0^\infty y \cdot 3e^{-3y} dy.$$

Using the formula  $\int uv dx = u \int v dx - \int (u' \int v dx) dx$ , let  $u = y$  and  $v = 3e^{-3y}$ . Then

$$\int_0^\infty y \cdot 3e^{-3y} dy = y \int_0^\infty 3e^{-3y} dy - \int_0^\infty (1 \cdot \int 3e^{-3y} dy) dy.$$

Now,  $\int 3e^{-3y} dy = -e^{-3y}$ . So,

$$\int_0^\infty y \cdot 3e^{-3y} dy = \left[ -ye^{-3y} \right]_0^\infty + \int_0^\infty e^{-3y} dy = 0 + \frac{1}{3}.$$

Therefore,

$$\boxed{\mathbb{E}[Y \mid X > 2] = \frac{1}{3}.$$

Q4: Let  $A$  denote the event “producer produces  $b$  items”, so  $\mathbb{P}(A) = p$  and  $\mathbb{P}(A^c) = 1 - p$ . When  $A^c$  occurs the consumer gets 0. When  $A$  occurs the consumed amount is  $C = \min(T, b)$ , where  $T \sim \text{Exp}(\lambda)$  and  $T$  is independent of  $A$ .

### (1) Distribution of $C$ .

**Conditional distribution given  $A$ .** Given  $A$ ,  $C = \min(T, b)$  has a mixed distribution: for  $0 < c < b$  its density is

$$f_{C|A}(c) = f_T(c) = \lambda e^{-\lambda c}, \quad 0 < c < b,$$

and there is an atom at  $b$  with

$$p_{C|A}(b) = \bar{F}_C(b) = e^{-\lambda b}.$$

Given  $A^c$ ,  $C = 0$  with probability 1.

**Unconditional (marginal) distribution.** Using the law of total probability,

- Probability mass at 0:

$$\mathbb{P}(\{C = 0\}) = \mathbb{P}(A^c) = 1 - p.$$

- For  $0 < c < b$  the density of  $C$  is

$$f_C(c) = \mathbb{P}(A) f_{C|A}(c) = p \lambda e^{-\lambda c}, \quad 0 < c < b.$$

- Probability mass at  $b$ :

$$\mathbb{P}(\{C = b\}) = \mathbb{P}(A)p_{C|A}(b) = pe^{-\lambda b}.$$

Thus  $C$  is a mixed random variable with a mass  $1 - p$  at 0, a continuous density  $p\lambda e^{-\lambda c}$  on  $(0, b)$ , and a mass  $pe^{-\lambda b}$  at  $b$ .

**(2) Expected value  $\mathbb{E}[C]$ .** Since  $C = 0$  on  $A^c$  and  $C = \min(T, b)$  on  $A$ ,

$$\begin{aligned}\mathbb{E}[C] &= 0 \cdot \mathbb{P}(A^c) + \int_0^b t\lambda e^{-\lambda t} \cdot \mathbb{P}(A) dt + b \cdot \mathbb{P}(\{C = b\}) \\ \int_0^b t\lambda e^{-\lambda t} dt &= \left[-te^{-\lambda t}\right]_0^b + \int_0^b e^{-\lambda t} dt = -be^{-\lambda b} + \frac{1 - e^{-\lambda b}}{\lambda}.\end{aligned}$$

So the final expectation is

$$\boxed{\mathbb{E}[C] = p \frac{1 - e^{-\lambda b}}{\lambda}.$$

Q5: (a) The CDF of  $Z$  is defined as:

$$F_Z(z) = P(Z \leq z) = P\left(\frac{X}{Y} \leq z\right).$$

Since  $X$  and  $Y$  are independent and uniformly distributed over  $(0, 1)$ , we can write this as:

$$F_Z(z) = P(X \leq zY).$$

We now express the probability as an integral over the possible values of  $Y$ :

$$F_Z(z) = \int_0^1 P(X \leq zy | Y = y) f_Y(y) dy.$$

For  $X \sim \text{Uniform}(0, 1)$ , we have  $P(X \leq zy) = \min(1, zy)$ , so the CDF becomes:

$$F_Z(z) = \int_0^1 \min(1, zy) dy.$$

Now, let's evaluate the integral in two parts, based on the value of  $z$ .

- If  $z \leq 1$ , the integration of  $\min(1, zy)$  is over  $zy \leq 1$ , which simplifies to  $zy$  for  $y \in [0, 1]$ .
- If  $z > 1$ , the minimum value becomes 1 for  $y \in [0, 1]$ , so the integral is over the entire interval.

Therefore, for  $z \leq 1$ :

$$F_Z(z) = \int_0^1 zy dy = \frac{z}{2}.$$

For  $z > 1$ :

$$F_Z(z) = \int_0^{1/z} zy dy + \int_{1/z}^1 1 dy = \frac{1}{2z} + \left(1 - \frac{1}{z}\right).$$

Thus, the CDF of  $Z$  is:

$$F_Z(z) = \begin{cases} \frac{z}{2}, & \text{if } 0 \leq z \leq 1, \\ 1 - \frac{1}{2z}, & \text{if } z > 1. \end{cases}$$

(b) To find the PDF of  $Z$ , we differentiate the CDF:

$$f_Z(z) = \frac{d}{dz}F_Z(z).$$

For  $z \leq 1$ :

$$f_Z(z) = \frac{d}{dz} \left( \frac{z}{2} \right) = \frac{1}{2}.$$

For  $z > 1$ :

$$f_Z(z) = \frac{d}{dz} \left( 1 - \frac{1}{2z} \right) = \frac{1}{2z^2}.$$

Thus, the PDF of  $Z$  is:

$$f_Z(z) = \begin{cases} \frac{1}{2}, & \text{if } 0 \leq z \leq 1, \\ \frac{1}{2z^2}, & \text{if } z > 1. \end{cases}$$

Q6: (a) The event  $(X \leq 2, Y \leq 4)$  corresponds to the sum of the probabilities in the table where  $X$  is either 1 or 2, and  $Y$  is either 2 or 4.

$$\begin{aligned} \mathbb{P}(X \leq 2, Y \leq 4) &= \mathbb{P}(X = 1, Y = 2) + \mathbb{P}(X = 1, Y = 4) \\ &\quad + \mathbb{P}(X = 2, Y = 2) + \mathbb{P}(X = 2, Y = 4) \\ &= \frac{1}{12} + \frac{1}{24} + \frac{1}{6} + \frac{1}{12} \\ &= \frac{2}{24} + \frac{1}{24} + \frac{4}{24} + \frac{2}{24} = \frac{9}{24} = \frac{3}{8}. \end{aligned}$$

(b) To find the marginal PMFs, we sum the probabilities across the rows (for  $Y$ ) and down the columns (for  $X$ ).

**Marginal PMF of  $X$ :**

- $p_X(1) = \mathbb{P}(X = 1) = \frac{1}{12} + \frac{1}{24} + \frac{1}{24} = \frac{2+1+1}{24} = \frac{4}{24} = \frac{1}{6}.$
- $p_X(2) = \mathbb{P}(X = 2) = \frac{1}{6} + \frac{1}{12} + \frac{1}{8} = \frac{4+2+3}{24} = \frac{9}{24} = \frac{3}{8}.$
- $p_X(3) = \mathbb{P}(X = 3) = \frac{1}{4} + \frac{1}{8} + \frac{1}{12} = \frac{6+3+2}{24} = \frac{11}{24}.$
- $p_X(x) = 0$  otherwise

**Marginal PMF of  $Y$ :**

- $p_Y(2) = \mathbb{P}(Y = 2) = \frac{1}{12} + \frac{1}{6} + \frac{1}{4} = \frac{1+2+3}{12} = \frac{6}{12} = \frac{1}{2}.$
- $p_Y(4) = \mathbb{P}(Y = 4) = \frac{1}{24} + \frac{1}{12} + \frac{1}{8} = \frac{1+2+3}{24} = \frac{6}{24} = \frac{1}{4}.$
- $p_Y(5) = \mathbb{P}(Y = 5) = \frac{1}{24} + \frac{1}{8} + \frac{1}{12} = \frac{1+3+2}{24} = \frac{6}{24} = \frac{1}{4}.$
- $p_Y(y) = 0$  otherwise

(c) Using the conditional probability formula:  $\mathbb{P}(Y = y|X = x) = \frac{\mathbb{P}(X=x, Y=y)}{\mathbb{P}(X=x)}.$

$$\begin{aligned} \mathbb{P}(Y = 2|X = 1) &= \frac{\mathbb{P}(X = 1, Y = 2)}{\mathbb{P}(X = 1)} \\ &= \frac{1/12}{1/6} \\ &= \frac{1}{12} \cdot 6 = \frac{6}{12} = \frac{1}{2}. \end{aligned}$$

- (d) For two random variables to be independent, their joint PMF must be equal to the product of their marginal PMFs for all pairs of values  $(x, y)$ . That is,  $\mathbb{P}(X = x, Y = y) = \mathbb{P}(X = x) \cdot \mathbb{P}(Y = y)$  for all  $x$  and  $y$ .

Let's check the first entry:  $(X = 1, Y = 2)$ .

- Joint PMF:  $\mathbb{P}(X = 1, Y = 2) = \frac{1}{12}$ .
- Marginal PMFs:  $\mathbb{P}(X = 1) = \frac{1}{6}$  and  $\mathbb{P}(Y = 2) = \frac{1}{2}$ .
- Product of marginals:  $\mathbb{P}(X = 1) \cdot \mathbb{P}(Y = 2) = \frac{1}{6} \cdot \frac{1}{2} = \frac{1}{12}$ .

In this case,  $\mathbb{P}(X = 1, Y = 2) = \mathbb{P}(X = 1) \cdot \mathbb{P}(Y = 2)$ .

Let's check another entry, for example  $(X = 2, Y = 4)$ :

- Joint PMF:  $\mathbb{P}(X = 2, Y = 4) = \frac{1}{12}$ .
- Marginal PMFs:  $\mathbb{P}(X = 2) = \frac{3}{8}$  and  $\mathbb{P}(Y = 4) = \frac{1}{4}$ .
- Product of marginals:  $\mathbb{P}(X = 2) \cdot \mathbb{P}(Y = 4) = \frac{3}{8} \cdot \frac{1}{4} = \frac{3}{32}$ .

Here,  $\frac{1}{12} \neq \frac{3}{32}$ . Since the condition  $\mathbb{P}(X = x, Y = y) = \mathbb{P}(X = x) \cdot \mathbb{P}(Y = y)$  does not hold for all values, the random variables  $X$  and  $Y$  are **not independent**.

Q7: The joint PDF of  $X$  and  $Y$  is

$$f_{X,Y}(x, y) = \frac{1}{4}, \quad 0 < x < 2, 0 < y < 2.$$

We want to find

$$\mathbb{P}(\{XY < 1\}) = \iint_{xy < 1} f_{X,Y}(x, y) dx dy.$$

The region  $xy < 1$  in the  $xy$ -plane is bounded by the curves  $y = \frac{1}{x}$ ,  $x = 0$ ,  $y = 0$ , and the lines  $x = 2$  and  $y = 2$ . The area of integration is thus the region where  $0 < x < 2$  and  $0 < y < \frac{1}{x}$ . Thus,

$$\mathbb{P}(\{XY < 1\}) = \int_0^2 \int_0^{1/x} \frac{1}{4} dy dx = \int_0^2 \frac{1}{4x} dx = \frac{1}{4} \ln(2).$$

Q8: Conditional probability  $\mathbb{P}(X > Y | X < 1/2)$  is:

$$\mathbb{P}(X > Y | X < 1/2) = \frac{\mathbb{P}(X > Y, X < 1/2)}{\mathbb{P}(X < 1/2)}$$

We find the marginal probability density  $f_X(x)$  to calculate the denominator. Integrating the joint probability density function  $f_{X,Y}(x, y)$  with respect to  $y$ :

$$f_X(x) = \int_0^1 (x + y) dy = \left[ xy + \frac{y^2}{2} \right]_0^1 = x + \frac{1}{2} \quad \text{for } 0 \leq x \leq 1$$

Now we can calculate the probability  $\mathbb{P}(X < 1/2)$ :

$$\mathbb{P}(X < 1/2) = \int_0^{1/2} \left( x + \frac{1}{2} \right) dx = \left[ \frac{x^2}{2} + \frac{x}{2} \right]_0^{1/2} = \frac{(1/2)^2}{2} + \frac{1/2}{2} = \frac{1}{8} + \frac{1}{4} = \frac{3}{8}$$

We calculate the joint probability  $P(X > Y, X < 1/2)$  by integrating over the region where  $0 \leq y < x$  and  $0 \leq x < 1/2$ :

$$\mathbb{P}(X > Y, X < 1/2) = \int_0^{1/2} \int_0^x (x + y) dy dx$$

Inner integral:

$$\int_0^x (x + y) dy = \left[ xy + \frac{y^2}{2} \right]_0^x = x^2 + \frac{x^2}{2} = \frac{3x^2}{2}$$

Outer integral:

$$\int_0^{1/2} \frac{3x^2}{2} dx = \left[ \frac{3x^3}{6} \right]_0^{1/2} = \left[ \frac{x^3}{2} \right]_0^{1/2} = \frac{(1/2)^3}{2} = \frac{1/8}{2} = \frac{1}{16}$$

Finally, we can find the conditional probability:

$$\mathbb{P}(X > Y | X < 1/2) = \frac{\mathbb{P}(X > Y, X < 1/2)}{\mathbb{P}(X < 1/2)} = \frac{1/16}{3/8} = \frac{1}{16} \times \frac{8}{3} = \frac{1}{6}$$