

Computationally Efficient Optimization of Plackett-Luce Ranking Models for Relevance and Fairness

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Goal of this work:

- Optimize a Plackett-Luce (PL) model for relevance or fairness ranking metrics,
- with an unbiased method (no heuristic or bounding),
- in a computationally efficient way (avoid combinatorial problems).

Contribution: PL-Rank

- A novel sampling-based method for quickly estimating PL gradients.
- Derivation to prove the estimation is unbiased.

Background

Background: Problem Setting



For any ranking y, an arbitrary ranking metric uses the weights per rank θ_k , the relevance of the items $P(R=1\,|\,q,d)=\rho_d$, and the policy π with the probability of a ranking $\pi(y\,|\,q)$:

$$\mathcal{R}(q) = \sum_{y \in \pi} \pi(y \mid q) \sum_{k=1}^{K} \theta_k P(R = 1 \mid q, y_k) = \sum_{y \in \pi} \pi(y) \sum_{k=1}^{K} \theta_k \rho_{y_k} = \mathbb{E}_y \left[\sum_{k=1}^{K} \theta_k \rho_{y_k} \right].$$

This is taken in expectation over a query distribution:

$$\mathcal{R} = \mathbb{E}_q[\mathcal{R}(q)] = \sum_{q \in \mathcal{Q}} P(q)\mathcal{R}(q).$$

This description applies to well-known metrics: precision@k, recall@k, DCG, ARP.



A Plackett-Luce model applies a **SoftMax** to the scores of **unplaced items**:

$$\pi(d \mid y_{1:k}, D) = \underbrace{\frac{\mathbb{1}[d \not\in y_{1:k}]e^{m(d)}}{\mathbb{1}[d \not\in y_{1:k}]e^{m(d')}}}_{\text{sum of all unplaced item scores}}.$$

The probability of a ranking is the product over each item placement:

$$\pi(y) = \prod_{k=1}^{K} \pi(y_k \,|\, y_{1:k-1}, \mathcal{D}).$$

We can sample from a Plackett-Luce model by sampling Gumbell Noise: $\zeta_d \sim \textit{Gumbell}$, and sorting according to $m(d) + \zeta_d$.



The prevalent approach in existing work uses **policy-gradients** with the log-trick:

$$\frac{\delta}{\delta m}\pi(y) = \pi(y) \left[\frac{\delta}{\delta m} \log(\pi(y)) \right].$$

Given N samples from π : $y^{(i)} \sim \pi$, the gradient can be unbiasedly estimated:

$$\frac{\delta}{\delta m} \mathcal{R}(q) \approx \frac{1}{N} \sum_{i=1}^{N} \underbrace{\left[\frac{\delta}{\delta m} \log(\pi(y^{(i)})) \right]}_{\text{gradient w.r.t. log prob. of full ranking}} \underbrace{\left(\sum_{k=1}^{K} \theta_k \rho_{y_k^{(i)}} \right)}_{\text{observed reward}}.$$

Method: PL-Rank



A **reward before** rank *k* should **not influence** the probabilities of the ranking **after** *k*:

$$\mathcal{R}(q) = \sum_{y \in \pi} \pi(y) \sum_{k=1}^K \theta_k \rho_{y_k} = \sum_{k=1}^K \theta_k \sum_{y \in \pi} \pi(y) \rho_{y_k} = \sum_{k=1}^K \theta_k \sum_{y_{1:k} \in \pi} \pi(y_{1:k}) \rho_{y_k}.$$

Given N samples from π : $y^{(i)} \sim \pi$, the gradient can be unbiasedly estimated:

$$\frac{\delta}{\delta m} \mathcal{R}(q) \approx \frac{1}{N} \sum_{i=1}^{N} \sum_{k=1}^{K} \underbrace{\left[\frac{\delta}{\delta m} \log(\pi(y_k^{(i)}|y_{1:k-1}^{(i)})) \right]}_{\text{prob. of partial ranking up to } k} \underbrace{\sum_{x=k}^{K} \theta_x \rho_{y_x^{(i)}}}_{\text{reward received after } k} \; .$$



Using the fact that π is a Plackett-Luce model, we can estimate the gradient using:

$$\frac{\delta}{\delta m} \mathcal{R}(q) \approx \frac{1}{N} \sum_{d \in \mathcal{D}} \underbrace{\left[\frac{\delta}{\delta m} m(d)\right]}_{i=1} \underbrace{\sum_{i=1}^{N} \left(\underbrace{\left(\sum_{k=\mathrm{rank}(d,y^{(i)})}^{K} \theta_k \rho_{y_k^{(i)}}\right)}_{rank(d,y^{(i)})} - \underbrace{\sum_{k=1}^{K} \pi(d \mid y_{1:k-1}^{(i)}) \left(\sum_{x=k}^{K} \theta_x \rho_{y_x^{(i)}}\right)}_{risk \ imposed \ by \ placement} \right)}_{risk \ imposed \ by \ placement}$$

Given N samples, this can be computed in $\mathcal{O}(N \cdot K \cdot D)$.

Flaw: items that are **not** in the top-K of any of the N sampled rankings will always have a **negative** gradient.



We can avoid the flaw while maintaining the $\mathcal{O}(N \cdot K \cdot D)$ complexity:

$$\frac{\delta}{\delta m} \mathcal{R}(q) \approx \frac{1}{N} \sum_{d \in \mathcal{D}} \underbrace{\left[\frac{\delta}{\delta m} m(d)\right]}_{\text{i=1}} \sum_{i=1}^{N} \underbrace{\left(\sum_{k=\mathrm{rank}(d,y^{(i)})+1}^{K} \theta_k \rho_{y_k^{(i)}}\right)}_{\text{expected direct reward minus the risk of placement}}^{\text{future reward after placement}} + \underbrace{\sum_{k=1}^{K} \pi(d \mid y_{1:k-1}^{(i)}) \left(\theta_k \rho_d - \sum_{x=k}^{K} \theta_x \rho_{y_x^{(i)}}\right)}_{\text{expected direct reward minus the risk of placement}}^{\text{future reward after placement}}.$$



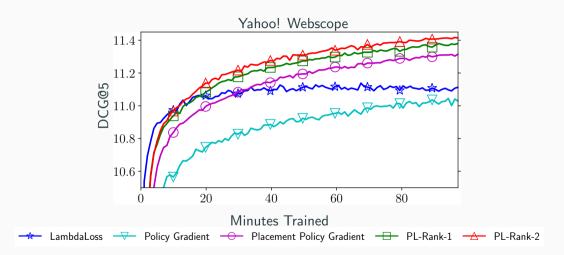
Fairness in exposure generally use rank-based exposure:

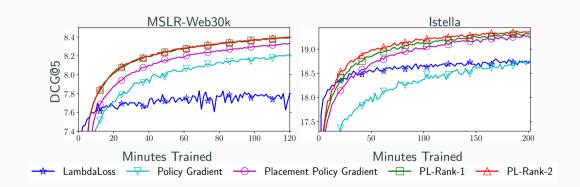
$$\mathcal{E}(q,d) = \mathbb{E}_y \left[\sum_{k=1}^K \theta_k \mathbb{1}[y_k = d] \right] = \sum_{y \in \pi} \pi(y) \sum_{k=1}^K \theta_k \mathbb{1}[y_k = d].$$

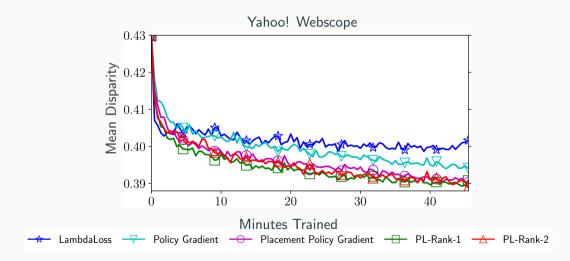
PL-Rank can be used to optimize a fairness based metric \mathcal{F} :

$$\begin{split} \frac{\delta}{\delta m} \mathcal{F}(q) &= \sum_{d \in \mathcal{D}} \bigg[\frac{\delta}{\delta m} m(d) \bigg] \mathbb{E}_y \Bigg[\left(\sum_{k = \mathsf{rank}(d, y) + 1}^K \theta_k \bigg[\frac{\delta \mathcal{F}(q)}{\delta \mathcal{E}_{y_k}} \bigg] \right) \\ &+ \sum_{k = 1}^{\mathsf{rank}(d, y)} \pi(d \, | \, y_{1:k-1}) \Bigg(\theta_k \bigg[\frac{\delta \mathcal{F}(q)}{\delta \mathcal{E}_d} \bigg] - \sum_{x = k}^K \theta_x \bigg[\frac{\delta \mathcal{F}(q)}{\delta \mathcal{E}_{y_x}} \bigg] \Bigg) \Bigg]. \end{split}$$

Experimental Results







Conclusion



PL-Rank: a novel LTR method for Plackett-Luce models:

- unbiased sample-based gradient estimation (no heuristic or bounding),
- computationally efficient (avoids combinatorial problems).
- applicable to relevance and fairness ranking metrics.

Continue our work: https://github.com/HarrieO/2021-SIGIR-plackett-luce

Future Work: StochasticRank



The **StochasticRank** algorithm (Ustimenko and Prokhorenkova, 2020) uses sampled noise to **stochastically smooth** a ranking function.

This algorithm has strong theoretical properties and could also be applied to Plackett-Luce models with comparable computational complexity.

Very promising direction for finding computationally efficient, effective and broadly applicable LTR.