

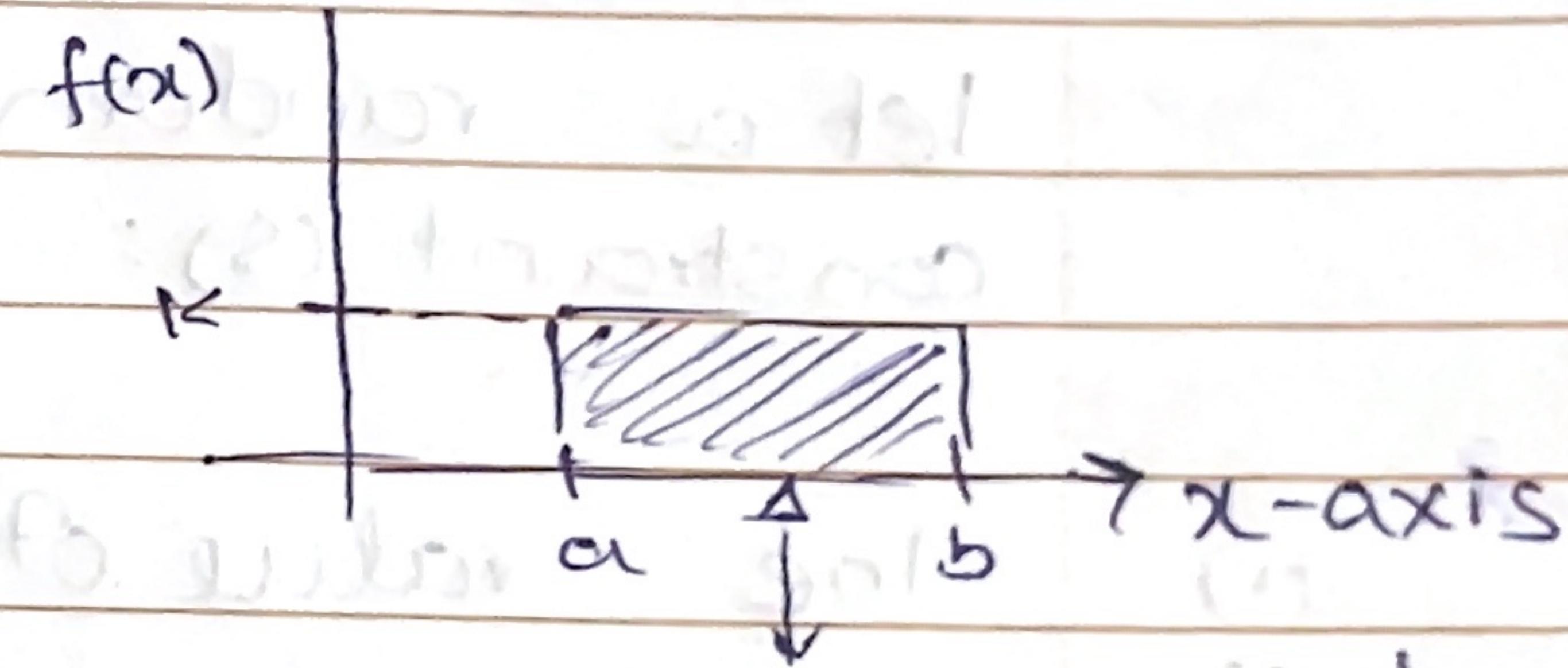
Post Mid-term!

26-03-24.

Continuous (\rightarrow uniform case) ^{DIST.}.

The type of distribution in which the probability of all the outcomes of an event are equally likely.

As by the definition
of probability



$\int_a^b k dx = 1$ mean: a point where k is a const. where $a \& b$ are balanced.

$$k \int_a^b dx = 1$$

$$k = \frac{1}{(b-a)} \quad \left\} \text{uniform's case!} \right.$$

Day / Date

(\rightarrow probability density funct)

Now writing PDF of uniform distribution (continuous).

$$f(x) = \begin{cases} \frac{1}{b-a} & a \leq x \leq b \\ 0 & \text{otherwise} \end{cases}$$

* Mean :- $\mu = \frac{a+b}{2}$

* Variance :- $\sigma^2 = \frac{(b-a)^2}{12}$

IDENTIFICATION :- $X \sim \text{uniform}$ (in Q statement).

$$X \sim U(a, b) \quad ("")$$

\rightarrow EXPONENTIAL:-

DST. CASE

let x random variable X that has(a)

constraint (s):

- (i) The value of X is never negative.
- (ii) Larger values of X have relative probability that drops exponentially.



Day / Date

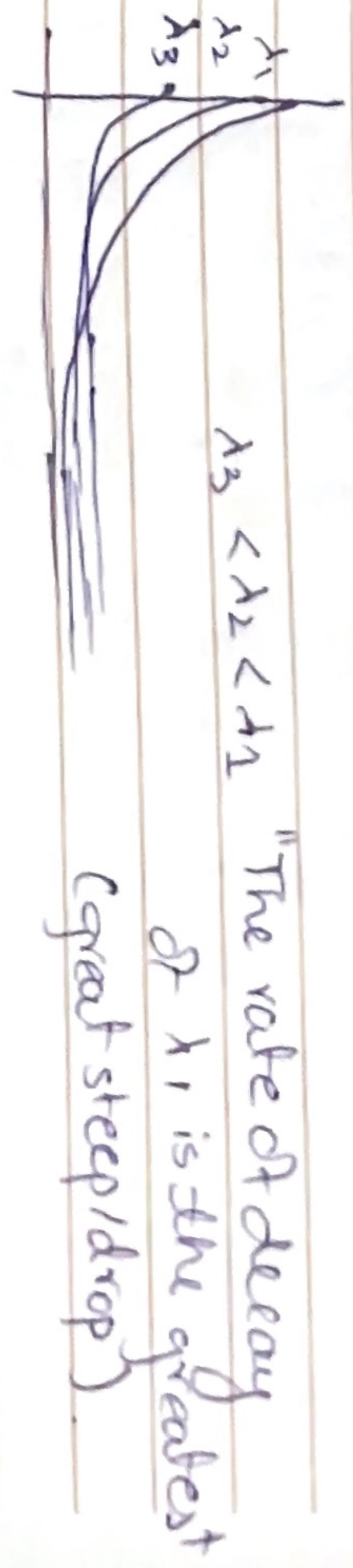
Designing probability Mass function.
density

$$f(x) = \begin{cases} \frac{1}{\beta} e^{-\frac{x}{\beta}} & x \geq 0 \\ 0 & x < 0 \end{cases}$$

also note $\lambda \rightarrow$ rate of decay.

$$\lambda = \frac{1}{\beta}$$

$$f(x) = \begin{cases} \lambda e^{-\lambda x} & x \geq 0 \\ 0 & x < 0 \end{cases}$$



* expected value / Mean :- $\mu / E(x) = \frac{1}{\lambda} \Rightarrow \beta$

$$\text{variance} := \sigma^2 / V[x] = \frac{1}{\lambda^2}$$

Simple linear Regression and Correlation.

$$y = mx + c .$$

$$y = \underset{\text{slope}}{b_1}x + b_0 . \quad *$$

$$b_1 = \frac{n \sum x_i y_i - (\sum x_i)(\sum y_i)}{n(\sum x_i^2) - (\sum x_i)^2}$$

$$\sum_{i=1}^{33} x_i = 1104 , \quad \sum_{i=1}^{33} y_i = 1124$$

$$\sum_{i=1}^{33} x_i y_i = 41355$$

$$\sum_{i=1}^{33} x_i^2 = 41086$$

$$B_1 = \frac{n \sum x_i y_i - (\sum x_i)(\sum y_i)}{n(\sum x_i^2) - (\sum x_i)^2}$$

$$b_1 = \frac{(33)(41355) - (1104)(1124)}{33(41086) - (1104)^2} \Rightarrow 0.903643 .$$

$$b_0 = \frac{\sum y_i - b_1 \sum x_i}{n} \Rightarrow \frac{1124 - (0.903643)(1104)}{33} .$$

$$\Rightarrow 3.829633$$

$$\therefore y = b_1 x + b_0 .$$

$$y = 0.903643x + 3.829633 \text{ Ans!}$$

Chapter no.2

بِسْمِ اللّٰهِ الرَّحْمٰنِ الرَّحِيْمِ

Rule no 2.1.

If an operation is performed in n_1 ways and for each of these ways another operation can be performed in n_2 ways. The total no. of ways an operation can be performed is:

$$n_1 \times n_2. \text{ (and so on).}$$

Permutations:-

Arrangement of all or some parts of a set of objects.

→ In general n distinct objects can be arranged as;

$$n(n-1)(n-2) \dots \text{ways. or } n!$$

→ No. of permutations of n distinct objects taken r at a time is given by.

$${}^n P_r = \frac{n!}{(n-r)!}$$

→ no. of permutation of n things where n_1 are of one kind, n_2 are of the second kind ... n_k of the k^{th} kind.

$$\frac{n!}{n_1! n_2! \dots n_k!}$$

→ the no. of ways of partitioning a set of n objects into r cells with n_1 elements in one cell, n_2 in the 2nd and so on

$$\frac{n!}{n_1! n_2! \dots n_r!} \Rightarrow \text{example 2.23}$$

combination.

The number of selection from n objects of r taken at a time, without the restriction of order and selection without replacement.

$${}^n C_r = \frac{n!}{r!(n-r)!}$$

example 2.22

Probability of an event.

(i) $0 \leq P(A) \leq 1$ probability of an event lies b/w 0 and 1.

(ii) $P(S) = 1$ probability of a sample space = 1.

Union of two events:-

overlapping sets.
 $A \cap B$

3 events' union.
 $P(A \cup B \cup C) = P(A) + P(B) + P(C) - P(A \cap B) - P(A \cap C) - P(B \cap C) + P(A \cap B \cap C)$.
mutually exclusive or disjoint sets.
 $A \cap B$.

$$P(A \cup B) = P(A) + P(B) - P(A \cap B) \quad P(A \cup B) = P(A) + P(B).$$

Theorem: ↑ Probability of an event and its complement is always equal 1.

$$P(A) + P'(A) = 1.$$

conditional Probability *

$$P(B|A) = \frac{P(B \cap A)}{P(A)} \Rightarrow \frac{P(A \cap B)}{P(A)}$$

always check for first event probability
for the second event if using combination.

When one event is independent of the other.

$$P(B|A) = P(B) \rightarrow \text{independent of } P(A)$$

$$P(A|B) = P(A) \rightarrow \text{independent of } P(B)$$

Independence criteria of two events!

* Two events are said to be independent

if $P(A \cap B) = P(A) \cdot P(B)$.

$$\left[P(A \cap B) = P(A|B) \times P(B) \right]$$

then shows dependence of
both events on one
another.

Total probability.

A probability of an event predicted through
probabilities of the outcomes of an event being
all the possible outcomes summed.

Chapter no.9. of P.

(i) Confidence interval of known $\mu \& \sigma^2$

If \bar{x} is the mean of a random sample of size n from a population of known σ^2 , $100(1-\alpha)\%$. interval

of μ is given by.

$$\frac{\bar{x} - z\alpha_{1/2} \frac{\sigma}{\sqrt{n}}}{\text{not lower bound}} < \mu < \frac{\bar{x} + z\alpha_{1/2} \frac{\sigma}{\sqrt{n}}}{\text{not upper bound}}$$

$$n = \left(\frac{z\alpha_{1/2} \sigma}{e}\right)^2 \rightarrow \begin{array}{l} \text{predicting sample size} \\ \text{for } \mu \text{ estimate.} \end{array}$$

$$\text{upper bound} = \bar{x} + z\alpha \frac{\sigma}{\sqrt{n}} \quad \left. \begin{array}{l} \text{take value of } z\alpha \\ \text{positive} \end{array} \right\}$$

$$\text{lower bound} = \bar{x} - z\alpha \frac{\sigma}{\sqrt{n}}.$$

(ii) Confidence interval when μ, σ^2 are unknown

If \bar{x} & s are the mean and variance of random sample of size n then the confidence interval $(1-\alpha)100\%$. for μ is given as.

$$\bar{x} - t\alpha_{1/2} \frac{s}{\sqrt{n}} < \mu < \bar{x} + t\alpha_{1/2} \frac{s}{\sqrt{n}}.$$

(iii) Confidence interval of σ^2

$$\frac{(n-1)s^2}{\chi^2_{1-\alpha/2}} < \sigma^2 < \frac{(n-1)s^2}{\chi^2_{\alpha/2}}$$

sample mean

Chapter 8. true mean = μ

$$\bar{X} \Rightarrow \frac{1}{n} \sum_{i=1}^n X_i$$

Sample variance.

true variance = σ^2 .

$$S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2 \Rightarrow \sqrt{\frac{1}{n-1} \sum_{i=1}^n [X_i^2 - (\sum X_i)^2]} / n(n-1)$$

95% confidence.

$$P(Z < -)$$

means to find area under the curve.

$$\chi^2_{0.975} - \chi^2_{0.025}$$

$$0.484 - 1.413$$

χ^2 distribution.

$$\chi^2 = \frac{(n-1)S^2}{\sigma^2}$$

$$P(\chi^2 = \chi^2_{\alpha}) = 0.99 \quad \nu = 4$$

$$\alpha = P$$

$\chi^2_{0.99}$ with $\nu=4$ is used!

t -distribution.

$$T = \frac{X - \mu}{S/\sqrt{n}}$$

confidence interval at 95%.

$$-t_{0.025} \leftrightarrow t_{0.025}$$

By def $P(T > t_{\alpha}) = \alpha$.

Chap 6:-

Dist. Uniform density funct.

Density funct of a continuous random variable on the interval $[A, B]$ is given as.

$$f(x; A, B) = \begin{cases} \frac{1}{B-A} & A \leq x \leq B \\ 0 & \text{elsewhere.} \end{cases}$$

$$\mu = \frac{A+B}{2} \quad \sigma^2 = \frac{(B-A)^2}{12}$$

Dist Exponential function.

The density function of a continuous random variable X with parameter β

$$f(x; \beta) = \begin{cases} \frac{1}{\beta} e^{-x/\beta} & x > 0 \\ 0 & \text{elsewhere} \end{cases}$$

$$\mu = \beta \quad \sigma^2 = \beta^2$$

↓ can be transformed
into decay
By $\lambda = \frac{1}{\beta}$.

Normal Distribution.

The density of a normally distributed random variable is;

$$n(x; \mu, \sigma^2) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2\sigma^2}(x-\mu)^2} \quad -\infty < x < \infty$$

it is σ^2 of normal.

$$Z = \frac{X-\mu}{\sigma}$$

Chapter no. 10.

	NULL Hypothesis True	NULL Hypothesis False
Reject H_0	False positive	True positive
Fail to Reject H_0	True negative	False negative

Type I error:- false positive

Null hypothesis is true but we rejected it.

Type II error : False negative

Null hypothesis is false but we reject it.

Z-test:- (two tail test).

► Check if the necessary data is complete.

i) Set $H_0 \& H_1$

ii) Set confidence interval from $Z_{\alpha/2}^{\checkmark} - Z_{\alpha/2}^{\checkmark}$

iii) find Z-value from given data

$$\text{using } Z = \frac{X - \mu}{\sigma}$$

iv) evaluate 'Z' in the confidence interval.

→ if it lies ⁱⁿ the critical region H_0 accept

→ otherwise H_0 reject.

* same for t-test if 'S' is given & $n < 30$.

Z/t-test (one tail):-

for $n < 30$
use t-
ib
ue

- (i) State H_0 & H_1 .
- (ii) Find the rejection & critical region but remember to use α now (either Z_{α} or as a one tail test. T_{α})
- (iii) Find the Z or T value by.
- (iv) Compare with the drawn regions.

$$Z = \frac{\bar{X} - \mu}{\sigma / \sqrt{n}} \text{ or } T = \frac{\bar{X} - \mu}{S / \sqrt{n}}$$

Calculating probability of type I error / type II error.

Problem 8.

$$H_0: \bar{X} = 68 \Rightarrow \mu = 68$$

$$H_1: \bar{X} \neq 68 \Rightarrow \mu \neq 68$$

$$\sigma = 3.6$$

$$n = 36$$

$\alpha \Rightarrow$ probability of the type I error.

not accepting the true hypothesis.

$$\alpha = P(\bar{X} < 67 \text{ when } \mu = 68) + P(\bar{X} > 69 \text{ when } \mu = 68)$$

$$P(Z < -1.67) + P(Z > 1.67)$$

$$2P(Z < -1.67) \Rightarrow 0.095$$

$$H_0 \geq 70$$

$$67 < \bar{x} < 69$$

$$H_1 \neq 70$$

β = Probability of accepting a false hypothesis.

$$P(\bar{x} < 69 \text{ when } \mu = 70)$$

$$z_1 = \frac{67 - 70}{0.45} = -6.67 \quad z_2 = \frac{69 - 70}{0.45} = -2.22$$

$$P(-6.67 < z < -2.22)$$

$$P(z < -2.22) - P(z < -6.67) \Rightarrow 0.0132$$

Change is usually in a alternate hypothesis.
Always identify what we are checking for in the question.

$$\begin{array}{ll} P(z > -a) & P(z < -a) \\ P(z < a) & \text{or} & P(z > a). \end{array}$$

$$\text{also } P(a < z < b)$$

$$\boxed{P(z < b) - P(z < a)}$$