# STSCI 5080 Probability Models and Inference

Lecture 5: Continuous Random Variables

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## Binomial coefficients

For a positive integer n and k = 0, 1, ..., n,

$$\binom{n}{k} = \text{number of } k\text{-element subsets of } \{1, \dots, n\}$$
$$= \frac{n!}{(n-k)!k!},$$

where

$$n! = n(n-1)\cdots 1$$
 and  $0! = 1$ .

For example,  $3! = 3 \cdot 2 \cdot 1 = 6, 4! = 4 \cdot 3 \cdot 2 \cdot 1 = 24$ , and

$$\binom{4}{2} = \frac{4!}{2!2!} = \frac{24}{2 \cdot 2} = 6.$$

# PMF of Bin(n, p)

#### **Theorem**

The pmf of  $Y \sim Bin(n, p)$  is

$$p(k) = P(Y = k) = \binom{n}{k} p^k (1 - p)^{n-k}, \ k = 0, 1, \dots, n.$$

Proof?

## Poisson random variable

#### **Definition**

Let  $\lambda>0$ . X is a Possion random variable with parameter  $\lambda$  if its takes values in  $\{0,1,2,\dots\}$  and its pmf is

$$p(k) = \frac{\lambda^k}{k!} e^{-\lambda}, \ k = 0, 1, 2, \dots$$

"X follows the Poisson distribution with parameter  $\lambda$ "

$$X \sim Po(\lambda)$$
.

## Continuous random variable

## Definition (probability density function (pdf))

A function f on  $\mathbb R$  is a probability density function (pdf) if  $f(x) \geq 0$  for any real x and

$$\int_{-\infty}^{\infty} f(x)dx = 1.$$

## Definition (Continuous random variable)

A random variable X is continuous if there exists a pdf f such that

$$P(X \in B) = \int_{B} f(x)dx$$

for any  $B \subset \mathbb{R}$ .

# Some properties of continuous random variables

If X has pdf f, then for any fixed real x,

$$P(X = x) = \int_{x}^{x} f(y)dy = 0.$$

In addition, for any a < b,

$$P(a < X < b) = P(a \le X < b)$$
  
=  $P(a < X \le b) = P(a \le X \le b) = \int_{a}^{b} f(x)dx$ .

## Cumulative distribution function

#### **Definition**

Let X be a continuous random variable with pdf f. Then the cumulative distribution function (cdf) F(x) of X is defined by

$$F(x) = P(X \le x) = \int_{-\infty}^{x} f(y)dy$$

for any real x.

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• For any a < b,

$$P(a < X \le b) = \int_a^b f(x)dx = \int_{-\infty}^b f(x)dx - \int_{-\infty}^a f(x)dx$$
$$= F(b) - F(a).$$

• For  $h \neq 0$ ,

$$\frac{F(x+h) - F(x)}{h} = \frac{1}{h} \int_{x}^{x+h} f(y) dy.$$

So, as long as f is continuous at x, taking  $h \to 0$ , we have

$$F'(x) = f(x).$$

# Example 5.1

#### Example (Uniform distribution)

Let a < b. A function defined by

$$f(x) = \begin{cases} \frac{1}{b-a} & \text{if } a \le x \le b\\ 0 & \text{otherwise} \end{cases}$$

is a pdf. A random variable X with this pdf is called a uniform random variable on [a,b]. The variable X concentrates on [a,b], i.e.,  $P(X \in [a,b]) = 1$ .

"X follows the uniform distribution on [a, b]"

$$X \sim U[a,b]$$
.

What is the cdf of X?

# Example 5.2

## Example

Let f be a function defined by

$$f(x) = \begin{cases} cx^2 & \text{if } 0 \le x \le 1\\ 0 & \text{otherwise} \end{cases},$$

where c > 0 is a constant. If f is a pdf, find the value of c, and compute the corresponding cdf.

# Exponential random variable

#### **Definition**

Let  $\lambda > 0$ . A random variable *X* with pdf

$$f(x) = \begin{cases} \lambda e^{-\lambda x} & \text{if } x \ge 0\\ 0 & \text{otherwise} \end{cases}$$

is called an exponential random variable with parameter  $\lambda$ . The variable X concentrates on  $[0, \infty)$ , i.e.,  $P(X \in [0, \infty)) = 1$ .

"X follows the exponential distribution with parameter  $\lambda$ "

$$X \sim Exp(\lambda)$$
.

What is the cdf of an exponential random variable?

## Standard normal random variable

#### **Definition**

A random variable X with pdf

$$f(x) = \frac{1}{\sqrt{2\pi}}e^{-x^2/2}, -\infty < x < \infty,$$

is called a standard normal random variable.

"X follows the standard normal distribution"

$$X \sim N(0, 1)$$
.

# Normal random variable with mean $\mu$ and variance $\sigma^2$

#### **Definition**

Let  $-\infty < \mu < \infty, \sigma > 0$ , and let  $X \sim N(0, 1)$ . The random variable

$$Y = \mu + \sigma X$$

is called a normal random variable with mean  $\mu$  and variance  $\sigma^2$ .

"Y follows the normal distribution with mean  $\mu$  and variance  $\sigma^2$ "

$$Y \sim N(\mu, \sigma^2)$$
.

# PDF of $N(\mu, \sigma^2)$

#### **Theorem**

Let  $Y \sim N(\mu, \sigma^2)$ . Then Y has pdf

$$f_Y(y) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-(y-\mu)^2/(2\sigma^2)}, -\infty < y < \infty.$$

Proof?