

HW 9, Harry Luo

8.6 (d)

$$\begin{aligned} E[(X+Y)^2] &= E[X^2 + 2XY + Y^2] = E[X^2] + E[XY] + E[Y^2] \\ &= \frac{2-p}{p^2} + \frac{2nr}{p} + n(n-1)r^2 + nr \end{aligned} \quad (1)$$

8.12

(a) By definition, the MGF can be found by

$$\begin{aligned} M_Z(t) &= E(e^{tZ}) = \int_0^\infty e^{tz} \lambda^2 z e^{-\lambda z} dz \\ &= \lambda^2 \int_0^\infty z e^{(t-\lambda)z} dz \end{aligned} \quad (2)$$

• if $\lambda - t > 0$, noticing the following expectation of an exp r.v. with parameter $\lambda - t$:

$$\int_0^\infty z(\lambda - t) e^{-(\lambda-t)z} dz = \frac{1}{\lambda - t} \quad (3)$$

Equation 2 becomes:

$$M_Z(t) = \frac{\lambda^2}{(\lambda - t)^2} \quad (4)$$

• if $\lambda - t \leq 0$, the integral converges to infinity.

$$\Rightarrow M_Z = \begin{cases} \frac{\lambda}{\lambda-t} & t < \lambda \\ \infty & t \geq \lambda \end{cases} \quad (5)$$

(b) for X and Y, example 5.6 on textbook suggests that

$$M_X(t) = M_Y(t) = \begin{cases} \frac{\lambda}{\lambda-t} & t < \lambda \\ \infty & t \geq \lambda \end{cases} \quad (6)$$

Given that X and Y are independent,

$$M_{X+Y} = M_{X(t)} M_{Y(t)} = \begin{cases} \frac{\lambda^2}{(\lambda-t)^2} & t < \lambda \\ \infty & t \geq \lambda \end{cases} \quad (7)$$

It is obvious that Equation 7 is the same as Equation 5, so the MGF of X+Y is the same as the MGF of Z. Thus X+Y has the same distribution as Z.

8.22

Considering the indicator method: let $I_j = 1$ or $0, j \in [1, 89]$ be the indicator of “among the five numbers, both j and $j+1$ are chosen”.

Since every time j and $j+1$ are chosen, they are next to each other in the ordered sample, so $X = \sum_{j=1}^{89} I_j$

$$E[X] = E\left[\sum_{j=1}^{89} I_j\right] = \sum_{j=1}^{89} E[I_j] \quad (8)$$

recognizing that

$$E[I_j] = P(j \text{ and } j+1 \text{ are chosen}) = \frac{\binom{88}{3}}{\binom{90}{5}} = \frac{2}{801} \quad (9)$$

$$E[X] = \frac{89 * 2}{801} = \frac{20}{89} \quad (10)$$

8.23

(a) Denote the color of the i th pick as C_i , then C_1, C_2, \dots, C_{50} are exchangeable. Thus,

$$P(C_{28} \neq C_{29}) = P(C_{28} = R, C_{29} = G) = 2P(C_1 = R, C_2 = G) = 2 * \frac{20 * 30}{50 * 49} = \frac{24}{49} \quad (11)$$

(b) Let I_j be the indicator that $Y_j \neq Y_{j+1}, j = 1, 2, \dots, 49$

$X = I_1 + \dots + I_{49}$, thus,

$$E[X] = E[I_1] + \dots + E[I_{49}] = 49 * P(C_1 \neq C_2) = 49 * \frac{24}{49} = 24 \quad (12)$$

8.36

(a) let I_j be the indicator that the number j appears at least once in the four rolls. Then $X = I_1 + \dots + I_6$. By exchangeability,

$$E[X] = E[I_1] + \dots + E[I_6] = 6 * E[I_1] \quad (13)$$

Since, $E[I_1] = P(1 \text{ appears in roll}) = 1 - P(\text{no rolls with } 1) = 1 - \left(\frac{5}{6}\right)^4$, we have

$$E[X] = 6 * \left(1 - \left(\frac{5}{6}\right)^4\right) = \frac{671}{216} \quad (14)$$

(b) noticing that $I_j^2 = I_j$ since I_j is either 1 or zero, and using exchangeability,

$$\begin{aligned} E[X^2] &= E[(I_1 + \dots + I_6)^2] = E[I_1^2 + \dots + I_6^2 + 2I_1I_2 + \dots + 2I_5I_6] \\ &= \sum_{j=1}^6 E[I_j^2] + 2 \sum_{i < j \leq 6} E[I_i I_j] \\ &= 6E[I_1^2] + 30E[I_1 I_2] \end{aligned} \quad (15)$$

TO find $E[I_1 I_2]$, we use the inclusion-exclusion principle,

$$\begin{aligned} E[I_1 I_2] &= P(1 \text{ and } 2 \text{ both show up at least once}) \\ &= 1 - P(1 \text{ does not show up}) - P(2 \text{ does not show up}) + P(1 \text{ and } 2 \text{ both do not show up}) \\ &= 1 - \left(\left(\frac{5}{6}\right)^4 + \left(\frac{5}{6}\right)^4 - \left(\frac{4}{6}\right)^4 \right) = \frac{151}{648} \end{aligned} \quad (16)$$

Thus,

$$\begin{aligned} E[X^2] &= \frac{671}{216} + \frac{151}{648} = \frac{541}{162} \\ \text{var}[X] &= E[X^2] - E[X]^2 \approx 0.447 \end{aligned} \quad (17)$$

8.40 (a)

Using indicator method, let I_k is the indicator for the event that the number k is drawn at least once in the 4 weeks. Then $X = I_1 + \dots + I_{90}$. By exchangeability,

$$E[X] = E[I_1] + \dots + E[I_{90}] = 90E[I_1] \quad (18)$$

Considering that

$$E[I_1] = P(1 \text{ is drawn at least once in 4 weeks}) = 1 - P(1 \text{ is not drawn in 4 weeks}) = 1 - \left(\frac{85}{90}\right)^4 \quad (19)$$

$$\Rightarrow E[X] = 90 \left(1 - \left(\frac{85}{90} \right)^4 \right) \approx 18.39 \quad (20)$$

8.42

the fourth moment can be found as

$$E[\overline{X_n^4}] = E \left[\left(\frac{X_1 + X_2 + \dots + X_n}{n} \right)^4 \right] = \frac{1}{n^3} E[(X_1 + \dots + X_n)^4] \quad (21)$$

According to binomial theorem,

$$\begin{aligned} E[\overline{X_n^4}] &= \frac{1}{n^4} E \left[\sum_k = 1^n X_k^4 + 24 \sum_{i < j < k < l} X_i X_j X_k X_l + 12 \sum_{k < l, j \neq k, j \neq l} X_j^2 X_k^2 + 4 \sum_{j \neq k} X_j^3 X_k \right] \\ &= \frac{1}{n^3} E[X_1^4] + 24 \binom{n}{4} E[X_1 X_2 X_3 X_4] + 12 \binom{n}{3} E[X_1^2 X_2 X_3] + 4 \binom{n}{2} E[X_1^3 X_2] + 6 \binom{n}{2} E[X_1^2 X_2^2] \end{aligned} \quad (22)$$

by independence, we know

$$\begin{aligned} E[X_1 X_2 X_3 X_4] &= E[X_1] E[X_2] E[X_3] E[X_4] = (E[X_1])^4 = 0 \\ E[X_1^2 X_2 X_3] &= E[X_1^2] E[X_2] E[X_3] = 0 \\ E[X_1^3 X_2] &= E[X_1^3] E[X_2] = 0 \\ E[X_1^2 X_2^2] &= E[X_1^2] E[X_2^2] = E[X_1^2]^2 \end{aligned} \quad (23)$$

Therefore,

$$\begin{aligned} E[\overline{X_n^4}] &= \frac{1}{n^3} E[X_1^4] + \frac{3n(n-1)}{n^4} E[X_1^2]^2 \\ &= \frac{c}{n^3} + \frac{3(n-1)a^2}{n^3} \end{aligned} \quad (24)$$

8.48

We can graph the joint pmf as follows:

X \ Y	0	1	2
1	9/100	0	0
2	81/100	9/100	0
3	0	0	1/100

From which we can find the marginals:

$$\begin{aligned} p_X(1) &= \frac{9}{100}, p_X(2) = \frac{90}{100}, p_X(3) = \frac{1}{100} \\ p_Y(0) &= \frac{90}{100}, p_Y(1) = \frac{9}{100}, p_Y(2) = \frac{1}{100} \end{aligned} \quad (25)$$

The expectance can be naturally found as

$$\begin{aligned}
 E[X] &= \frac{48}{25}, E[Y] = \frac{11}{100}, E[XY] = \frac{6}{25} \\
 \text{Cov}(X, Y) &= E[XY] - E[X]E[Y] = \frac{18}{625}
 \end{aligned}
 \tag{26}$$

8.54

(a)

$$\begin{aligned}
 \text{var}(X) &= E[X^2] - E[X]^2 = 5 - 2^2 = 1 \\
 \text{var}(Y) &= E[Y^2] - E[Y]^2 = 10 - 1 = 9 \\
 \text{Cov}(X, Y) &= E[XY] - E[X]E[Y] = 1 - 2 = -1 \\
 \Rightarrow \text{Corr}(X, Y) &= \text{Cov} \frac{X, Y}{\sqrt{\text{var}(X) \text{var}(Y)}} = -\frac{1}{3}
 \end{aligned}
 \tag{27}$$

(b) Noticing

$$\text{Cov}(X, X + cY) = \text{Var}(X) + c \text{Cov}(X, Y) = 1 + c \tag{28}$$

so when $c = -1$, the covariance is 0, and the r.v. X and $X + cY$ are not correlated