6.34

Consider a random point (X,Y) uniformly distributed over the quadrilateral region D with vertices at (0,0), (2,0), (1,1), and (0,1).

(a) Given that the area of D equals $\frac{3}{2}$, the joint probability density function is:

$$f_{X,Y}(x,y) = \begin{cases} \frac{2}{3} & \text{for } (x,y) \in D, \\ 0 & \text{otherwise.} \end{cases}$$

Note that the boundary of D includes a line segment from (1,1) to (2,0), described by y=2-x. We can derive the marginal density functions as follows: For the marginal density of X:

$$f_X(x) = \begin{cases} 0, & x \le 0 \text{ or } x \ge 2, \\ \int_0^1 \frac{2}{3} \, dy = \frac{2}{3}, & 0 < x \le 1, \\ \int_0^{2-x} \frac{2}{3} \, dy = \frac{4}{3} - \frac{2}{3}x, & 1 < x < 2. \end{cases} \tag{2}$$

To verify that this is a valid density function:

$$\int_{-\infty}^{\infty} f_X(x) dx = \int_0^1 \frac{2}{3} dx + \int_1^2 \left(\frac{4}{3} - \frac{2}{3}x\right) dx$$
$$= 1,$$

confirming that $f_X(x)$ is indeed a proper density function. Similarly, for the marginal density of Y:

$$f_Y(y) = \begin{cases} 0, & y \le 0 \text{ or } y \ge 1, \\ \int_0^{2-y} \frac{2}{3} dx = \frac{4}{3} - \frac{2}{3}y, & 0 < y < 1. \end{cases}$$

(b) To find $\mathbb{E}[X]$ and $\mathbb{E}[Y]$:

$$\mathbb{E}[X] = \int_{-\infty}^{\infty} x f_X(x) \, dx$$

$$= \int_0^1 \frac{2}{3} x \, dx + \int_1^2 \left(\frac{4}{3} x - \frac{2}{3} x^2 \right) \, dx$$

$$= \frac{7}{9},$$

and

$$\mathbb{E}[Y] = \int_{-\infty}^{\infty} y f_Y(y) \, dy$$
$$= \int_0^1 \left(\frac{4}{3}y - \frac{2}{3}y^2\right) \, dy$$
$$= \frac{4}{9}.$$

6.36

Suppose that the random variables X and Y have the joint probability density

$$f(x,y) = ce^{-\frac{x^2}{2} - \frac{(x-y)^2}{2}}, \quad x, y \in (-\infty, \infty),$$

where c is a constant.

(a) To determine the value of c, we require that

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y) \, dx dy = 1.$$

Evaluating this double integral:

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} c e^{-\frac{x^2}{2} - \frac{(x-y)^2}{2}} dx dy = c \int_{-\infty}^{\infty} e^{-\frac{x^2}{2}} \int_{-\infty}^{\infty} e^{-\frac{(x-y)^2}{2}} dy dx$$
$$= \sqrt{2\pi} c \int_{-\infty}^{\infty} e^{-\frac{x^2}{2}} dx = 2\pi c.$$

Setting this equal to 1 yields $c = \frac{1}{2\pi}$. (b) The marginal density function of X is given by

$$f_X(x) = \int_{-\infty}^{\infty} \frac{1}{2\pi} e^{-\frac{x^2}{2} - \frac{(x-y)^2}{2}} dy$$

$$= \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{(x-y)^2}{2}} dy$$

$$= \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}},$$

which is the density of a standard normal random variable. Similarly, for the marginal density of Y:

$$f_Y(y) = \int_{-\infty}^{\infty} \frac{1}{2\pi} e^{-\frac{x^2}{2} - \frac{(x-y)^2}{2}} dx$$
$$= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2} - \frac{(x-y)^2}{2}} dx.$$

Completing the square in the exponent and simplifying:

$$f_Y(y) = \frac{1}{\sqrt{4\pi}} e^{-y^2/4} \int_{-\infty}^{\infty} \frac{1}{\sqrt{\pi}} e^{-(x-y/2)^2} dx$$
$$= \frac{1}{\sqrt{4\pi}} e^{-y^2/4}.$$

The last step uses the fact that $\frac{1}{\sqrt{\pi}}e^{-(x-y/2)^2}$ is the pdf of an N(y/2,1) distributed random variable. Thus, $Y \sim N(0,2)$.

In summary, $X \sim N(0,1)$ and $Y \sim N(0,2)$.