



University of L'Aquila

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Masters Degree in Applied Data Science

TIME SERIES WITH APPLICATIONS ON BIG DATA

Windspeed Forecast in Delhi

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Introduction

In this project, the daily climate time series data for Delhi is analysed using different **ARIMA** models in **GretL**. The best model is then identified and used to create a forecast.

The variables of interest are dates and wind speed (measured in km/h). The data was collected from the 1st of January to the 24th of April in 2017. Although this is daily data, it is assumed that the season is constant since it is for four months.

Firstly, the wind speed is plotted against time to identify the data trend. Time is the independent variable while windspeed is the dependent variable. To make the windspeed series stationary, the data is transformed by taking the first difference. Next, possible orders of p and q are selected using a correlogram and the *armax* function. This makes it easy to create the models.

Using only observations from the Train set, each of the models is fitted. The models are then tested with the Test data and used to produce a forecast for wind speed. The forecasted value is compared with the actual windspeed value to determine each model's RMSE (Root Mean Squared Error) and choose the best model.

Finally, the best model is used to create a prediction for the next 20 days.

The data used can be downloaded from Kaggle.

Preliminary Analysis

Here the general trend of the data is shown. It is seen that the windspeed has a general upward trend between January and April however there are periods of low wind speed particularly in February. It may also be seen that the highest windspeed was in the middle of April.

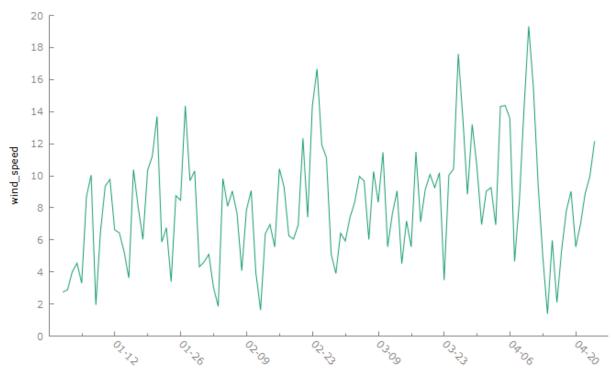


Figure 2.1: Windspeed against Time

Due to this systematic increase over time, the first difference is taken so that the data can achieve stationarity i.e. fluctuate around a constant mean. This makes it easier to model and forecast.

Data Transformation

In the previous chapter, a trend was recognised in the data. A key assumption for time series models is that they must be stationary hence the first difference of the data has to be calculated.

Given a time series x_t , and the lag operator $\Delta = 1 - L$

The first difference is defined as

$$\Delta x_t = x_t - x_{t-1}$$

where

 Δx_t is the differenced series

 x_t is the value of the windspeed at time t

 x_{t-1} is windspeed at t-1

Below is the graph of windspeed against time after the first difference is taken.

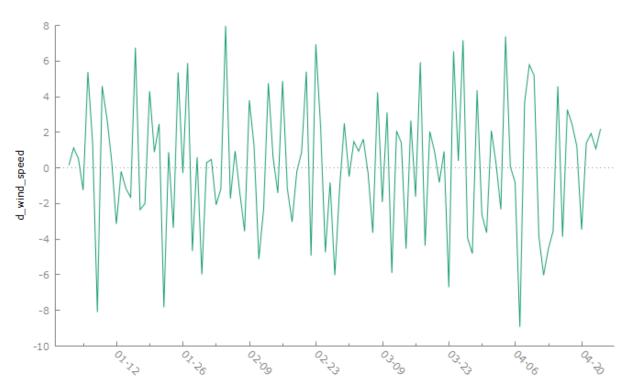


Figure 3.1: first difference of windspeed against time

It is noticed that the windspeed appears stationary around 0. The summary statistics for the first difference series are given below and it can be noticed that the mean of the data is different from 0.

Table 3.1: Summary statistics for the first difference of windspeed

Summary statistics, using the observations 2017-01-01 - 2017-04-24 for the variable 'd_wind_speed' (113 valid observations)

Mean	0.083306
Median	0.39170
Minimum	-8.9278
Maximum	7.9655
Standard deviation	3.7396

In addition, the Augmented Dicker Fuller (ADF) test was applied to the first differenced series. This is a test used to determine whether a time series data is stationary. The ADF test has a null hypothesis H0 which states that the time series has a unit root i.e non-stationary. The Python function *adfuller* from the *statsmodel* package was used to carry out this test and the result below was obtained.

Table 3.2: ADF test results

	adf_result	values
0	Test Statistic	-7.3679
1	p-value	9.1335e-11
2	Number of Lags Used	5
3	Number of Observations Used	107
4	Critical Values	{'1%': -3.493, '5%': -2.889, '10%': -2.5814}
5	IC Best	533.1258

The test statistic is less than the critical value at 1%, 5% and 10% significance values indicating that the null hypothesis can be rejected. Furthermore, the p-value is less than 0.05 significance level. Therefore, the first differenced series is stationary, and the data requires no further transformation. The number of lags used, 5, was chosen to minimise the Information Criteria (IC) value of 533 which in this case represents the Akaike IC (AIC) value.

Model Identification

Firstly, the correlogram with a maximum lag length of 10 is plotted for the first differenced data. This shows the graphs for the Autocorrelation Function (ACF) and Partial Autocorrelation Function (PACF). In the graph below, there is a spike at lag length 1 then it sharply diminishes at 2 for both ACF and PACF suggesting that q=1 and p=1.

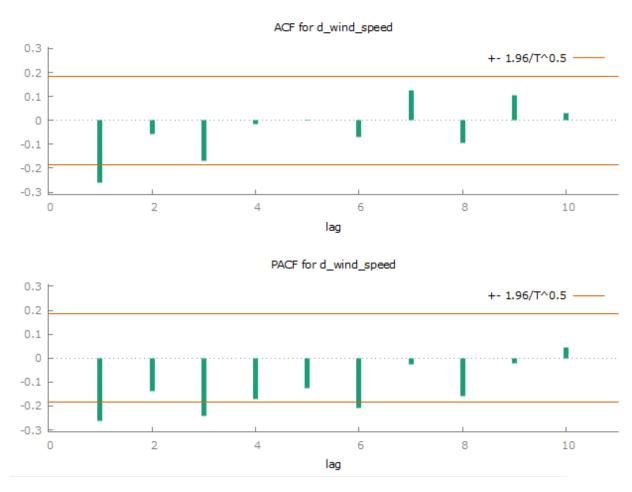


Figure 4.1: Correlogram for the first difference of windspeed

4.1 Information Criteria

Sometimes, visualising the correlogram is not enough to determine the ARIMA models to build. The GretL software has the armax function which can be used to check the Information Criteria (IC). This uses three different estimators to achieve a balance between model fit and complexity. The estimators are **AIC**, Bayesian IC (**BIC**) and Hannan-Quinn Criterion (**HQC**). The minimum values of **AIC**, **BIC** and **HQC** are usually chosen as the best model.

Based on the correlogram, the parameters for the *armax* function was set as: $(0 \le p \le 3, 0 \le q \le 3)$ and the following table was produced

? armax(3, 3, d_wind_speed, null, 1, 1, 0, 1, 0)

Information Criteria of ARMAX(p,q) for d wind speed

p,	q	AIC	BIC	HQC
ο,	0	621.7660	627.2208	623.9795
Ο,	1	609.8551	618.0373	613.1754
0,	2	594.6586	605.5682	599.0856
Ο,	3	591.3316	604.9686	596.8654
1,	0	615.7644	623.9466	619.0846
1,	1	591.1318*	602.0413*	595.5588*
1,	2	593.0435	606.6804	598.5772
1,	3	593.1997	609.5641	599.8402
2,	0	615.6510	626.5606	620.0780
2,	1	592.9886	606.6256	598.5224
2,	2	595.0509	611.4152	601.6914
2,	3	595.1964	614.2881	602.9436
3,	0	610.8831	624.5200	616.4168
3,	1	593.0676	609.4320	599.7081
3,	2	595.0554	614.1471	602.8026
3,	3	597.0478	618.8669	605.9018

^{*} indicates best models.

It can be observed that the *armax* function indicates p=1 and q=1 as the best model with minimum values of **AIC**, **BIC** and **HQC**. It was earlier established that d=1 since the first difference was applied once to make the series stationary so we have the model as

$$\Delta x_t \sim ARMA(1,1) \Rightarrow x_t \sim ARIMA(1,1,1)$$

Hence the model has the following equation:

$$(1 - \phi_1 L)(1 - L)x_t = c + (1 + \theta_1 L)u_t \tag{4.1}$$

Furthermore, two other models are considered; **AR**(1) and **MA**(1) to understand how the AutoRegressive and Moving Average parts work together to explain a given time series.

ARIMA(0,1,1)

$$(1 - L)x_t = c + (1 + \theta_1 L)u_t \tag{4.2}$$

ARIMA(1,1,0)

$$(1 - \phi_1 L)(1 - L)x_t = c + u_t \tag{4.3}$$

^{&#}x27;9999.9999' suggests failures to estimate the models.

Model Estimation

For this process, the data is divided into two parts:

Train set: 80% of total data (2017-01-01 to 2017-03-01)

Test set: remaining 20% (2017-04-01 to 2017-04-24)

5.1. First model: **ARIMA**(1,1,1)

Below is the result of applying the model to the data. The coefficients θ_1 and φ_1 and the constant are all significantly different from 0 with p-values < 0.05

Table 5.1.1: Model 1 ARIMA(1,1,1) result

Model 5: ARIMA, using observations 2017-01-02:2017-03-31 (T = 89) Estimated using AS 197 (exact ML)
Dependent variable: (1-L) wind_speed
Standard errors based on Hessian

co	efficient	std. error	z z	p-value
const 0	0.0423140	0.0175430	2.412	0.0159 **
phi_1 0	.323923	0.102340	3.165	0.0015 ***
theta_1 -1	1.00000	0.0313546	-31.89	3.31e-223 ***
Mean dependent Mean of innovat R-squared Log-likelihood Schwarz criteri	0.1615 0.1876 -225.08	578 S.D. 668 Adjus 835 Akail	dependent var of innovation sted R-squared se criterion an-Quinn	ns 2.969894

So we can substitute their values into equation 4.1

$$(1 - 0.324L)(1 - L)x_t = 0.042 + (1 - L)u_t$$

5.1.1 Model Checking

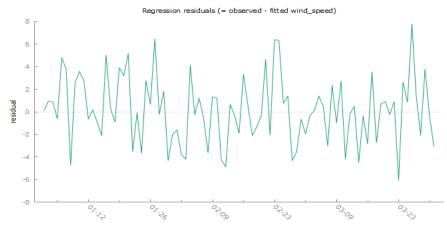


Figure 5.1.1: Residuals plotted against time

It is necessary to check the behaviour of the residuals to understand how good the model is. The residuals are analyzed to determine if it is a Gaussian White Noise (WN). In *fig 5.1.1* above, the residuals are plotted against time. It can be observed that the residuals are stationary with mean around 0.

Secondly, the correlogram below is plotted and it shows no obvious autocorrelation. All the values are close to 0 at all lags and lie within the 95% confidence interval. Hence, it can be concluded that the residuals behave like WN

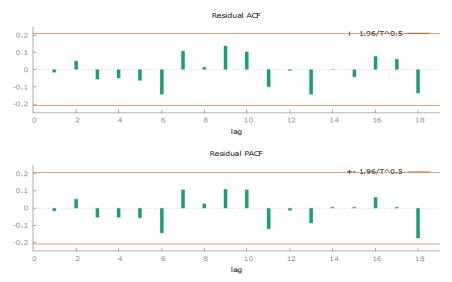


Figure 2.1.2: Residual Correlogram

Thirdly, the normality of the residuals are checked using a histogram and Q-Q plot.

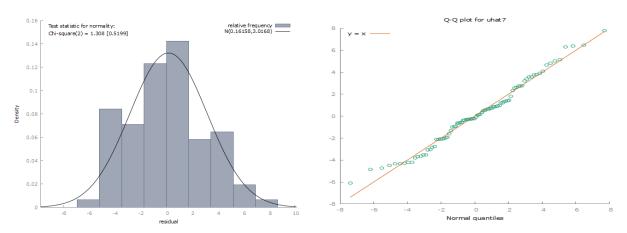


Figure 5.1.3: Normality tests for the residuals

The histogram in *fig* 5.1.3 shows a p-value of 0.5199 which is greater than 0.05 indicating that the null hypothesis (no significant autocorrelation) is not rejected. In the q-q plot, most values fall on the red line. This means that the residuals follow a Gaussian distribution.

However, creating a plot of the residuals versus the actual series below shows that there seems to be a correlation. But since the residuals passes the other tests, this can be considered a good model.

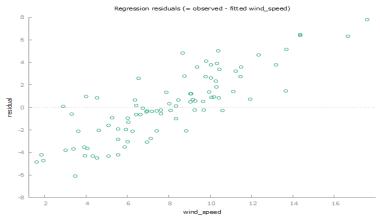


Figure 5.1.4: Residuals plotted against windspeed

5.2. Second model: ARIMA(0,1,1)

Applying an MA(1) model to the data yields the following result

Table 5.2.1: Model 2- ARIMA(0,1,1)

Model 3: ARIMA, using observations 2017-01-02:2017-03-31 (T = 89) Estimated using AS 197 (exact ML) Dependent variable: (1-L) wind_speed Standard errors based on Hessian

coeffic	cient	std.	erro	c z	p-value	
const 0.042	 1542	0.012	26700	3.327	0.0009	***
theta_1 -1.0000	00	0.037	78550	-26.42	8.84e-154	***
Mean dependent var	0.0472	264	S.D.	dependent va	r 3.6718	58
Mean of innovations	0.2079	986	S.D.	of innovatio	ns 3.1226	42
R-squared	0.1090	50	Adjus	sted R-square	d 0.1090	50
Log-likelihood	-229.87	779	Akail	ke criterion	465.75	58
Schwarz criterion	473.22	217	Hanna	n-Ouinn	468.76	51

Substitute the coefficient values into (4.2):

$$(1-L)x_t = 0.042 + (1-L)u_t$$

5.2.1. Model Checking



Figure 5.2.1: Residuals plotted against time

Taking a closer look at the plot of the residuals above, it appears to have a constant mean and variance.

However, at lag 1 in the correlogram, there is a value outside the confidence interval indicating that not all the residuals behave like white noise.

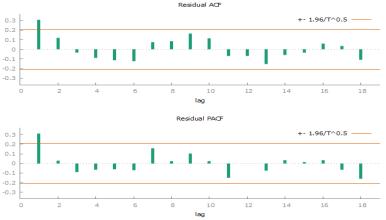


Figure 5.2.2: Correlogram of residuals

Observing the normality of the residuals below shows that they follow a Gaussian distribution and the p-value is greater than 0.05.

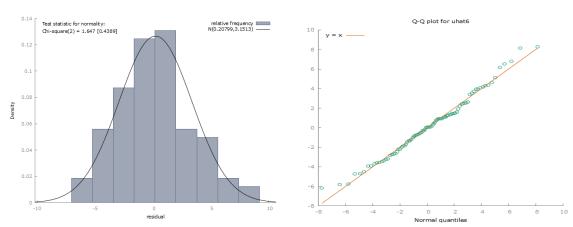


Figure 5.2.3: Normality tests for the residuals

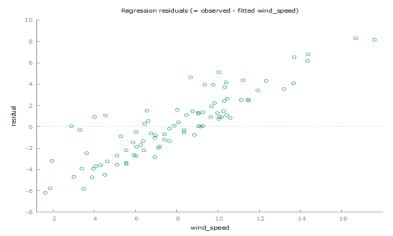


Figure 5.2.4: Residuals plotted against windspeed

However in the residuals versus the windspeed plotted above, there appears to be a correlation. The model does not pass the check for Residual correlogram and correlation with the actual series but it performs well for the other tests so it is a fairly good prediction model.

5.3. Third model: **ARIMA**(1,1,0)

Applying the AR(1) model to the data yields the result below

Table 5.3.1: Model 3 - ARIMA(1,1,0)

```
Model 9: ARIMA, using observations 2017-01-02:2017-03-31 (T = 89)
Estimated using AS 197 (exact ML)
Dependent variable: (1-L) wind speed
Standard errors based on Hessian
            coefficient
                                               p-value
                         std. error
                                        Z
  const
            0.0578589
                       0.267667
                                     0.2162
                                               0.8289
 phi_l
            -0.354815
                        0.0988847
                                      -3.588
                                               0.0003
Mean dependent var 0.047264 S.D. dependent var
Mean of innovations 0.000303 S.D. of innovations 3.410940
                   0.145096 Adjusted R-squared 0.145096
R-squared
                  -235.5547
Log-likelihood
                             Akaike criterion
                                                  477.1095
Schwarz criterion
                   484.5754
                              Hannan-Quinn
                                                  480.1187
```

Substitute the coefficient values into (4.3):

$$(1 + 0.355L)(1 - L)x_t = 0.058 + u_t$$

5.3.1. Model Checking

Once again, the characteristics of the residuals are checked to determine if it is a Gaussian WN. Firstly, the residuals are plotted against time.

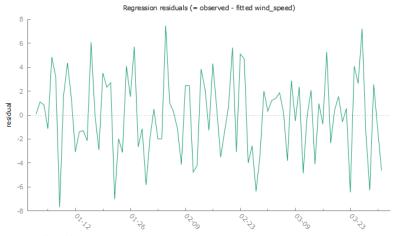


Figure 5.3.1: Residuals plotted against time

Next, the correlogram is observed below. In the Residual ACF, all values lie within the confidence interval but this is not the case in the PACF.

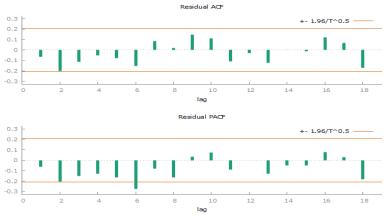


Figure 5.3.2: Correlogram of residuals

Looking at the results of the normality tests indicates that the residual is a Gaussain WN. The p-value is greater than 0.05 and majority of the residuals lie on the red line in the q-q plot.

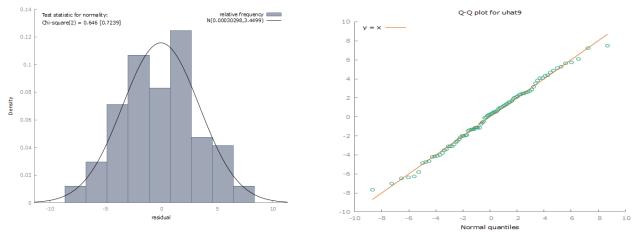


Figure 5.3.3: Normality tests for the residuals

Lastly, the residuals are plotted against the windspeed and unlike the previous models, there are no obvious correlations. Hence it can be concluded that this is a good prediction model.

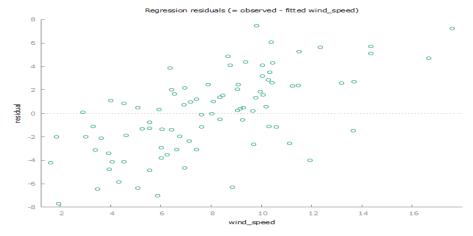


Figure 5.3.4: Residuals plotted against windspeed

Model Selection

In the previous chapter, three models were defined and fitted to the data. In this chapter, the best model is selected. The prediction of April's values is obtained by applying the models to the test set and the model with the lowest RMSE value is the best model.

Below are the graphs that compare the predicted values and the actual values and a table showing the forecast statistics for each model.

6.1 **ARIMA**(1,1,1) Forecast

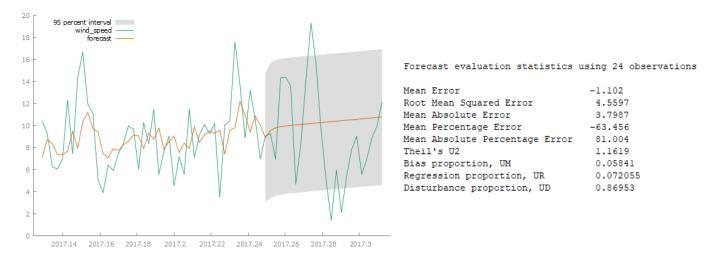


Figure 3.1: ARIMA(1,1,1) forecast result

The above graph shows that only a few points lie outside the confidence interval. The negative mean error indicates that a lower windspeed value was mostly forecasted.

6.2. **ARIMA**(0,1,1) Forecast

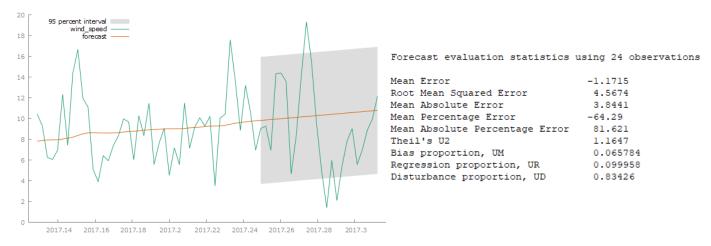


Figure 6.2: ARIMA(0,1,1) forecast result

Fig 6.2 is very similar to Fig 6.1 however the latter is better because it has a lower RMSE of 4.5597. **ARIMA**(0,1,1) had an RMSE of 4.5674.

6.3. **ARIMA**(1,1,0) Forecast

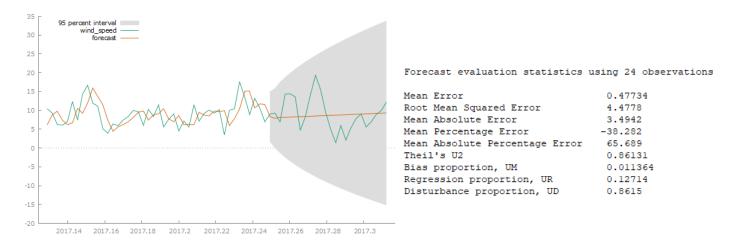


Figure 6.3: ARIMA(1,1,0) forecast result

The third model produces a better result than the other two models. Its RMSE value of **4.478** is the lowest of all models. This suggests that the **MA** part of the model can affect prediction performance. Hence, **ARIMA**(1,1,0) is identified as the best model and will be used for forecasting.

Forecast and Conclusion

The data is restored to the full range (2017-01-01 to 2017-04-24) and the model **ARIMA**(1,1,0) is applied.

Table 7.1: Best model ARIMA(1,1,0)

Model 10: ARIMA, using observations 2017-01-02:2017-04-24 (T = 113) Estimated using AS 197 (exact ML)
Dependent variable: (1-L) wind_speed
Standard errors based on Hessian

coe	fficient st	d. error	z p-	value
				7678 0040 ***
Mean dependent vo				3.739620 3.592408
R-squared	0.189298	Adjusted R	-squared	0.189298
Log-likelihood	-304.8822	Akaike cri	terion	615.7644
Schwarz criterio	n 623.9466	Hannan-Qui	nn	619.0846

Then the windspeed for the next 20 days is predicted.

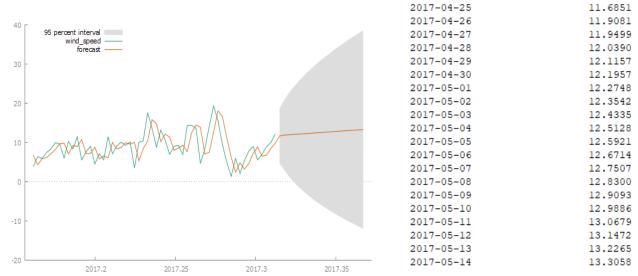


Figure 7.1: Graph of actual vs predicted windspeed

Table 7.2: Windspeed forecast

It can be observed that the windspeed values are expected to increase gradually over the next 20 days. However, this will likely be affected by other climatic factors such as rainfall and humidity which the forecast did not consider so the increase will not be a linear one and there will be periods of windspeed decrease.