Ch. 7 – Confidence Intervals

Standard Normal Probability

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Solving for *c* gives

$$c = -\Phi^{-1}(.025) = 1.96$$

Suppose a machine drills holes whose diameters are normally distributed with mean $\mu=3.0$ mm and standard deviation $\sigma=0.4$ mm. If a random sample of 16 holes are measured, find an interval [a,b] centered on 3.0 such that the sample mean \overline{X} diameter will be in the interval [a,b] with probability .95.

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Rearranging,

$$.95 = P(2.804 \le \overline{X} \le 3.196)$$

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So we may take $[A, B] = [\overline{X} - .196, \overline{X} + .196]$. For example, if $\overline{X} = 3.05$, then [A, B] = [2.854, 3.246].

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- In general, a $100(1-\alpha)\%$ confidence interval for θ is an interval [A,B] such that $P(A \le \theta \le B) = 1-\alpha$.
- We are usually interested in constructing **equal-tailed** confidence intervals, where $P(\theta < A) = P(\theta > B) = \alpha/2$.

Confidence Interval for Mean of Normal Distribution

If a random sample of size n is taken from a normal distribution with unknown mean μ and known standard deviation σ , then an equal-tailed $100(1-\alpha)\%$ confidence interval for μ is given by

$$\overline{X} \pm \frac{z_{\alpha/2} \cdot \sigma}{\sqrt{n}}$$

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Proof: The sample mean \overline{X} is normal with mean μ and standard deviation σ/\sqrt{n} , so $\frac{\overline{X}-\mu}{\sigma/\sqrt{n}}$ is standard normal, hence

$$\begin{split} 1 - \alpha &= P\left(-z_{\alpha/2} \leq \frac{\overline{X} - \mu}{\sigma/\sqrt{n}} \leq z_{\alpha/2}\right) \\ &= P\left(\overline{X} - \frac{z_{\alpha/2} \cdot \sigma}{\sqrt{n}} \leq \mu \leq \overline{X} + \frac{z_{\alpha/2} \cdot \sigma}{\sqrt{n}}\right) \end{split}$$

Example

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Here $\alpha = 1 - .99 = .01$, so the relevant critical value is

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The sample mean is $\overline{X} = 5.64$, so the confidence interval is given by

$$\overline{X} \pm \frac{z_{\alpha/2} \cdot \sigma}{\sqrt{n}} = 5.64 \pm \frac{2.57 \cdot 0.7}{\sqrt{9}} = 5.64 \pm 0.60$$

In the previous example, a random sample of 9 beads was used to estimate the mean diameter μ , with a margin of error of 0.60, at a 99% confidence level. What sample size would be required for a margin of error of no more than 0.10, at the 99% confidence level?

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Rounding up to an integer, a sample size of at least n = 324 would be required.

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- It does not mean that 95% of beads will have diameter between 4.9 and 5.1.

Large-sample Confidence Interval for Mean

If X_1,\ldots,X_n are a random sample from a distribution with unknown mean μ and *unknown* variance σ^2 , and if n is sufficiently large, then an approximate $100(1-\alpha)\%$ confidence interval for μ is given by

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• Rule of thumb: The large-sample confidence interval for μ may be used if n > 40.

$$\overline{X} \pm \frac{z_{\alpha/2} \cdot S}{\sqrt{n}} = 842 \pm \frac{1.96 \cdot 799}{\sqrt{600}}$$

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$$= [778.1, 905.9]$$

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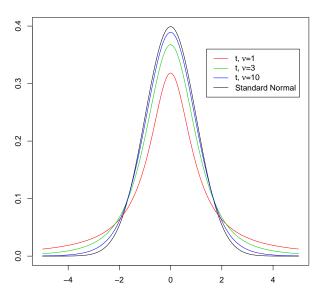
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Given a random sample X_1, \ldots, X_n from a normal distribution with unknown mean and variance, A $100(1-\alpha)\%$ confidence interval for the mean μ is

$$\overline{X} \pm \frac{t_{\alpha/2,\nu} \cdot S}{\sqrt{n}}$$

where $t_{\alpha/2,\nu}$ is a critical value from a **t distribution** with $\nu=n-1$ degrees of freedom.

PDF of t Distribution vs. Standard Normal



As $\nu \to \infty$, the t distribution approaches a standard normal.

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$$t_{\alpha/2,\nu} = t_{.005,8} = 3.355$$

The sample mean is $\overline{X} = 5.64$ and the sample standard deviation is S = .809, so the confidence interval is given by

$$\overline{X} \pm \frac{t_{\alpha/2} \cdot S}{\sqrt{n}} = 5.64 \pm \frac{3.355 \cdot 0.809}{\sqrt{9}} = 5.64 \pm 0.90$$

Prediction Interval for a Normal Population

Given a random sample X_1, \ldots, X_n from a normal distribution, suppose we want to construct an interval [A, B] which we can be 95% confident will contain a future observation X_{n+1} . Such an interval is called a **prediction interval**.

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The statistic $T=rac{\overline{X}-X_{n+1}}{S\sqrt{1+rac{1}{n}}}$ has a t distribution with $\nu=n-1$ degrees of freedom.

Given a random sample X_1, \ldots, X_n from a normal distribution, a $100(1-\alpha)\%$ prediction interval for an independent observation X_{n+1} is

$$\overline{X} \pm t_{\alpha/2,n-1} \cdot S\sqrt{1+rac{1}{n}}$$

An article reports the following data on the breakdown voltage of electrically stressed circuits, assumed to be normally distributed:

$$1470, 1510, 1690, 1740, 1900, 2000, 2030, 2100, 2200, \\ 2290, 2380, 2390, 2480, 2500, 2580, 2190, 2700$$

Find a 95% confidence interval for the mean μ . Then find a 95% prediction interval for a future observation.

We found $\overline{X}=2126.5$ and $S^2=137324.3$. The critical value is $t_{\alpha/2,\nu}=t_{.025,16}=2.120$, giving the confidence interval for μ :

$$\overline{X} \pm \frac{t_{\alpha/2,\nu} \cdot S}{\sqrt{n}} = 2126.5 \pm 190.5$$

Likewise we get the prediction interval:

$$\overline{X} \pm t_{\alpha/2,\nu} \cdot S\sqrt{1 + \frac{1}{n}} = 2126.5 \pm 808.4$$

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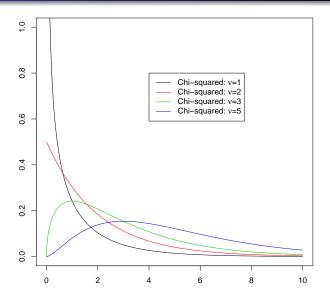
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Given a random sample X_1, \ldots, X_n from a normal distribution with unknown mean μ and variance σ^2 , A $100(1-\alpha)\%$ confidence interval for σ^2 is

$$\left[\frac{(n-1)S^2}{\chi^2_{\alpha/2,n-1}}, \frac{(n-1)S^2}{\chi^2_{1-\alpha/2,n-1}}\right]$$

where $\chi^2_{\alpha/2,n-1}$ and $\chi^2_{1-\alpha/2,n-1}$ are critical values from a χ^2 distribution with $\nu=n-1$ degrees of freedom.

Chi-squared distribution



Recall the breakdown voltage data, assumed to be normal:

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$$\left[\frac{(n-1)S^2}{\chi^2_{\alpha/2,n-1}}, \frac{(n-1)S^2}{\chi^2_{1-\alpha/2,n-1}}\right] = [76172.3, 318064.4]$$

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for σ^2 . The corresponding confidence interval for σ is

$$[\sqrt{76172.3}, \sqrt{318064.4}] = [276.0, 564.0]$$

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How do we construct a confidence interval for p based on \hat{p} ?

Suppose X is a binomial random variable counting the number of successes in n trials where each trial has probability p of success. If n is sufficiently large an approximate $100(1-\alpha)\%$ confidence interval is given by

$$\hat{p}\pm z_{lpha/2}\sqrt{rac{\hat{p}(1-\hat{p})}{n}}$$

where $\hat{p} = X/n$ is the sample proportion, and $z_{\alpha/2}$ is a critical value from the standard normal distribution.

Suppose we have a sequence of n Bernoulli trials, where the probability p of success is unknown. If we observe X successes, we know that the maximum likelihood estimator of p is the sample proportion $\hat{p} = X/n$.

How do we construct a confidence interval for p based on \hat{p} ?

Suppose X is a binomial random variable counting the number of successes in n trials where each trial has probability p of success. If n is sufficiently large an approximate $100(1-\alpha)\%$ confidence interval is given by

$$\hat{p} \pm z_{\alpha/2} \sqrt{rac{\hat{p}(1-\hat{p})}{n}}$$

where $\hat{p} = X/n$ is the sample proportion, and $z_{\alpha/2}$ is a critical value from the standard normal distribution.

Rule of thumb: This may be used if the number of successes X and the number of failures n-X are both at least 10.

A quality control team for a manufacturer tests 200 randomly selected devices, out of which 15 are defective. Assume that defective devices occur independently of one another. Find an approximate 95% confidence interval for the proportion defective.

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Here the sample proportion is $\hat{p}=15/200=.075$, and the critical value is $z_{.025}=1.96$, so the approximate 95% confidence interval is

$$\hat{p} \pm z_{\alpha/2} \sqrt{\frac{\hat{p}(1-\hat{p})}{n}} = .075 \pm 1.96 \sqrt{\frac{.075(1-.075)}{200}}$$
$$= .075 \pm .037$$

In the previous example, a 95% confidence interval for the proportion was .075 \pm .037. Estimate the required sample size to achieve a margin of error of .01.

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$$n \approx \frac{1.96^2(.075)(1 - .075)}{.01^2} = 2665.11 \approx 2666$$

Summary

| Confidence interval for mean μ of normal, σ known | $\overline{X} \pm \frac{z_{\alpha/2} \cdot \sigma}{\sqrt{n}}$ |
|--|---|
| Large-sample approximate confidence interval for mean μ | $\overline{X} \pm \frac{z_{\alpha/2} \cdot S}{\sqrt{n}}$ |
| Confidence interval for mean μ of normal, σ unknown | $\overline{X} \pm \frac{t_{\alpha/2,n-1} \cdot S}{\sqrt{n}}$ |
| | |
| Prediction interval for normal observation | $\overline{X} \pm t_{lpha/2,n-1} \cdot S\sqrt{1+rac{1}{n}}$ |
| | $ \overline{X} \pm t_{\alpha/2, n-1} \cdot S \sqrt{1 + \frac{1}{n}} $ $ \left[\frac{(n-1)S^2}{\chi^2_{\alpha/2, n-1}}, \frac{(n-1)S^2}{\chi^2_{1-\alpha/2, n-1}} \right] $ |