

Ch. 8 – Hypothesis Testing

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- Given a coin, one hypothesis is that each toss has probability $p = .5$ of coming up heads. Another hypothesis would be that $p \neq .5$.
- Given a certain type of candy bar, labeled as having a mass of 60 grams, one hypothesis is that the mean mass is as labeled, $\mu = 60$. Another hypothesis would be that the mean mass is smaller than labeled: $\mu < 60$.

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- H_0 , the **null hypothesis**, is the hypothesis which we initially presume to be true.
- The other hypothesis H_a is called the **alternative hypothesis**.
- The hypothesis H_0 is rejected only if the sample evidence strongly contradicts it. Otherwise we continue to believe that H_0 is plausible.
- The two possible outcomes of the analysis are that we **reject** the null hypothesis H_0 , or we **do not reject** the null hypothesis.

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For example, one procedure would be to toss the coin 10 times and reject the null hypothesis if we obtain 8 heads or more.

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For the example of testing for an unfair coin, the test statistic was the number of heads X out of 10 tosses, and the rejection region was $\{8, 9, 10\}$.

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We cannot eliminate the possibility of these errors. However, we can quantify their probability of occurring:

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Typically, the choice of rejection region involves a tradeoff between the two types of errors. But by using larger samples, both error probabilities may be reduced.

Example

Suppose we test a coin by tossing it 10 times and rejecting it if we get 8 or more heads. If in reality the coin is fair, what is the probability α of a Type I error? If the coin is unfair with probability $p = .75$ of being heads, what is the probability β of a Type II error?

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To find the Type I error, we assume $p = .5$ and calculate the probability that the number of heads X is at least 8:

$$\alpha = P(X \geq 8) = \sum_{x=8}^{10} \binom{10}{x} (.5)^x (1 - .5)^{10-x} = .055$$

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To find the Type II error in the case $p = .75$, we calculate the probability that the number of heads X is less than 8:

$$\beta = P(X < 8) = \sum_{x=0}^7 \binom{10}{x} (.75)^x (.25)^{10-x} = .474$$

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To find the Type II error in the case $p = .75$, we calculate the probability that the number of heads X is less than 7:

$$\beta = P(X < 7) = \sum_{x=0}^6 \binom{10}{x} (.75)^x (.25)^{10-x} = .224$$

By enlarging the rejection region, we increased α but decreased β .

Problem

A type of candy bar is labeled 60 grams. We decide to test the label's accuracy using a random sample X_1, \dots, X_5 by rejecting the null hypothesis $H_0 : \mu = 60$ if $\bar{X} < 59$. If the masses are normally distributed with $\sigma = 0.8$, what is the Type I error probability α ?

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If H_0 is true, then $Z = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} = \frac{\bar{X} - 60}{0.8/\sqrt{5}}$ is a standard normal random variable,

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If H_0 is true, then $Z = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} = \frac{\bar{X} - 60}{0.8/\sqrt{5}}$ is a standard normal random variable, so

$$\begin{aligned}\alpha &= P(\bar{X} < 59) \\ &= P\left(\frac{\bar{X} - 60}{0.8/\sqrt{5}} < \frac{59 - 60}{0.8/\sqrt{5}}\right) \\ &= P(Z < -2.80) \\ &= .0026\end{aligned}$$

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In this case, $Z = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} = \frac{\bar{X} - 58.5}{0.8/\sqrt{5}}$ is a standard normal random variable, so

$$\begin{aligned}\beta &= P(\bar{X} \geq 59) \\ &= P\left(\frac{\bar{X} - 58.5}{0.8/\sqrt{5}} \geq \frac{59 - 58.5}{0.8/\sqrt{5}}\right) \\ &= P(Z \geq 1.40) \\ &= .0808\end{aligned}$$

Problem: Two-tailed Test

A machine is specified to drill holes with diameter 4 mm. We test the hypothesis $H_0 : \mu = 4$, using a random sample X_1, \dots, X_{30} . We reject H_0 if $\bar{X} > 4.1$ or $\bar{X} < 3.9$. If the diameters are normally distributed with $\sigma = .2$, find the Type I error probability α .

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If H_0 is true, then $Z = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} = \frac{\bar{X} - 4}{0.2/\sqrt{30}}$ is a standard normal random variable, so

$$\begin{aligned}\alpha &= P(\bar{X} < 3.9) + P(\bar{X} > 4.1) \\ &= 2P(\bar{X} < 3.9) \\ &= 2P\left(\frac{\bar{X} - 4}{0.2/\sqrt{30}} < \frac{3.9 - 4}{0.2/\sqrt{30}}\right) \\ &= 2P(Z < -2.74) \\ &= .0062\end{aligned}$$

As we have seen, for a fixed sample size, selecting a rejection region for a test involves a tradeoff between the Type I error probabilities α and the Type II error probability β .

Significance Level

As we have seen, for a fixed sample size, selecting a rejection region for a test involves a tradeoff between the Type I error probabilities α and the Type II error probability β .

A common practice is to design the test to achieve a specified small value of α , such as $\alpha = .1, .05$, or $.01$. The choice for α is called the **significance level**.

z Test for Mean of Normal Distribution with Known σ

Given a random sample X_1, \dots, X_n from a normal distribution with known standard deviation σ , the z **test** for the null hypothesis $H_0 : \mu = \mu_0$, based on the test statistic $Z = \frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}}$, is given by the following rejection region, depending on whether a one-tailed or two-tailed test is desired:

Alternative hypothesis		Rejection region
(Upper-tailed test)	$H_a : \mu > \mu_0$	$Z \geq z_\alpha$
(Lower-tailed test)	$H_a : \mu < \mu_0$	$Z \leq -z_\alpha$
(Two-tailed test)	$H_a : \mu \neq \mu_0$	$ Z \geq z_{\alpha/2}$

Here α is the significance level (Type I error probability), and z_α is a critical value from the standard normal distribution.

Example

A machine is specified to drill holes with diameter 4 mm. We wish to test the null hypothesis $H_0 : \mu = 4$ against the alternative $H_a : \mu \neq 4$. If the diameters are normally distributed with $\sigma = .2$, and we observe $\bar{X} = 3.87$ in a sample of size 10, do we reject the null hypothesis at the significance level $\alpha = .05$?

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The test statistic is

$$Z = \frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}} = \frac{3.87 - 4}{.2/\sqrt{10}} = -2.06$$

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The rejection region is $\{|Z| > z_{\alpha/2}\}$ where $z_{\alpha/2} = z_{.025} = 1.96$.

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The rejection region is $\{|Z| > z_{\alpha/2}\}$ where $z_{\alpha/2} = z_{.025} = 1.96$.

Since $|Z| = 2.06 > 1.96$, the test statistic is in the rejection region, so we reject the null hypothesis.

Large-Sample z Test for Mean with Unknown σ

Given a random sample X_1, \dots, X_n from a distribution with unknown standard deviation, the z **test** for the null hypothesis $H_0 : \mu = \mu_0$, based on the test statistic $Z = \frac{\bar{X} - \mu_0}{S/\sqrt{n}}$, is given by the following rejection region:

Alternative hypothesis		Rejection region
(Upper-tailed test)	$H_a : \mu > \mu_0$	$Z \geq z_\alpha$
(Lower-tailed test)	$H_a : \mu < \mu_0$	$Z \leq -z_\alpha$
(Two-tailed test)	$H_a : \mu \neq \mu_0$	$ Z \geq z_{\alpha/2}$

Here α is the nominal significance level. If n is large, then under H_0 , Z is approximately standard normal by the Central Limit Theorem, so the true significance level is approximately α .

Note: Here we don't need to assume that the distribution of X_1, \dots, X_n is normal.

Example

A machine is specified to drill holes with diameter 4 mm. We wish to test the null hypothesis $H_0 : \mu = 4$ against the alternative $H_a : \mu \neq 4$. If we observe $\bar{X} = 3.97$ and $S = .21$ in a sample of size 100, do we reject the null hypothesis at the significance level $\alpha = .05$?

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The rejection region is $\{|Z| > z_{\alpha/2}\}$ where $z_{\alpha/2} = z_{.025} = 1.96$.

Since $|Z| = 1.43 < 1.96$, the test statistic is not in the rejection region, so we do not reject the null hypothesis. In other words, based on the data, it is plausible that the mean is $\mu = 4$ as specified.

t Test for Mean of Normal Distribution with Unknown σ

Given a random sample X_1, \dots, X_n from a normal distribution with unknown standard deviation, the t **test** for the null hypothesis $H_0 : \mu = \mu_0$, based on the test statistic $T = \frac{\bar{X} - \mu_0}{S/\sqrt{n}}$, is given by the following rejection region:

Alternative hypothesis		Rejection region
(Upper-tailed test)	$H_a : \mu > \mu_0$	$T \geq t_{\alpha, n-1}$
(Lower-tailed test)	$H_a : \mu < \mu_0$	$T \leq -t_{\alpha, n-1}$
(Two-tailed test)	$H_a : \mu \neq \mu_0$	$ T \geq t_{\alpha/2, n-1}$

Here α is the significance level (Type I error probability), and $t_{\alpha, n-1}$ is a critical value from the t distribution with $n - 1$ degrees of freedom.

Example

A type of candy bar is labeled 60 grams. Someone suggests that the candy bars weigh less than specified. To test this, we gather a random sample of size 5 and observe $\bar{X} = 58.8$ and $S = 0.9$. Do we reject the null hypothesis $H_0 : \mu = 60$ at the significance level $\alpha = .01$?

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We use the t test with the one-tailed alternative $H_a : \mu < 60$. The test statistic is

$$T = \frac{\bar{X} - \mu_0}{S/\sqrt{n}} = \frac{58.8 - 60}{0.9/\sqrt{5}} = -2.98$$

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$$T = \frac{\bar{X} - \mu_0}{S/\sqrt{n}} = \frac{58.8 - 60}{0.9/\sqrt{5}} = -2.98$$

The critical value is $t_{\alpha, n-1} = t_{.01, 4} = 3.747$. The rejection region is $\{T < -3.747\}$, whereas in our sample $T = -2.98 > -3.747$, so we do not reject the null hypothesis.

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The critical value is $t_{\alpha, n-1} = t_{.01, 4} = 3.747$. The rejection region is $\{T < -3.747\}$, whereas in our sample $T = -2.98 > -3.747$, so we do not reject the null hypothesis.

In other words, the data does *not* allow us to conclude that the average weight of the candy bars is less than specified.

Large-Sample z Test for Proportion

Given a random sample from a Bernoulli distribution with unknown parameter p , the z **test** for the null hypothesis $H_0 : p = p_0$ based on the test statistic $Z = \frac{\hat{p} - p_0}{\sqrt{p_0(1-p_0)/n}}$ is given by the following rejection region:

Alternative hypothesis		Rejection region
(Upper-tailed test)	$H_a : p > p_0$	$Z \geq z_\alpha$
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(Two-tailed test)	$H_a : p \neq p_0$	$ Z \geq z_{\alpha/2}$

Here α is the nominal significance level, and z_α is a critical value from the standard normal distribution.

Example

We are given a coin which someone suggests may give outcomes with unequal proportions when we spin it on a table. We test this by spinning the coin 80 times. If we observe 54 heads, do we reject the null hypothesis at the $\alpha = .01$ significance level?

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The rejection region is $\{|Z| > z_{\alpha/2}\}$ where $z_{\alpha/2} = z_{.005} = 2.58$. Since $|Z| = 3.13 > 2.58$, we reject the null hypothesis.

In other words, the test provides strong evidence that the coin indeed gives heads more often than tails when spun.

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For example, in the candy bar example, at the $\alpha = .01$ we failed to reject the null hypothesis; however, if we had used a less strict significance level, $\alpha = .05$, then the test would have rejected.

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For example, in the candy bar example, at the $\alpha = .01$ we failed to reject the null hypothesis; however, if we had used a less strict significance level, $\alpha = .05$, then the test would have rejected. The P-value would provide a more nuanced summary of the test result by indicating precisely at what significance level the test changes from rejecting to failing to reject.

P-Values

Loosely speaking, another way to describe the P-value is as follows:

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- This probability is calculated assuming that the null hypothesis is true.
- Beware: The P-value is not the probability that H_0 is true.
- The interpretation of “as extreme as” depends on the alternative hypothesis:
 - For an upper-tailed alternative, it means “as large as”.
 - For a lower-tailed alternative, it means “as small as”.
 - For a two-tailed alternative, it means “as large in absolute value as”.

P-Values for z tests and t test

With all of the z test procedures, the P-value may be calculated as follows, where z is the observed value of the test statistic Z (assumed to have a standard normal distribution):

Alternative hypothesis		P-value
(Upper-tailed test)	$H_a : \mu > \mu_0$	$P(Z \geq z)$
(Lower-tailed test)	$H_a : \mu < \mu_0$	$P(Z \leq z)$
(Two-tailed test)	$H_a : \mu \neq \mu_0$	$P(Z \geq z)$

Similarly, for the t test, the P-value may be calculated as follows, where t is the observed value of the test statistic T (assumed to have a t distribution with $n - 1$ degrees of freedom):

Alternative hypothesis		P-value
(Upper-tailed test)	$H_a : \mu > \mu_0$	$P(T \geq t)$
(Lower-tailed test)	$H_a : \mu < \mu_0$	$P(T \leq t)$
(Two-tailed test)	$H_a : \mu \neq \mu_0$	$P(T \geq t)$

Example

A type of candy bar is labeled 60 grams. Someone suggests that the candy bars weigh less than specified. To test this, we gather a random sample of size 5 and observe $\bar{X} = 58.8$ and $S = 0.9$. What is the P-value for the test?

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We use the t test with the one-tailed alternative $H_a : \mu < 60$. As before, the test statistic is

$$T = \frac{\bar{X} - \mu_0}{S/\sqrt{n}} = \frac{58.8 - 60}{0.9/\sqrt{5}} = -2.98$$

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So $P = .020$ is the P-value for the test.

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The P-value is the probability that we would observe a value of Z this extreme (i.e., a value of Z with $|Z| \geq 3.13$):

$$P = P(|Z| \geq 3.13) = 2\Phi(-3.13) \approx 2(.0009) = .0018$$

Summary

α = probability of Type I error, that H_0 is true but is rejected

β = probability of a Type II error, that H_0 is false but is not rejected

P = P-value = smallest α for which the test would reject H_0

Test	Null Hypothesis	Test Statistic
z test	$H_0 : \mu = \mu_0$	$Z = \frac{\bar{X} - \mu_0}{\sigma / \sqrt{n}}$
t test	$H_0 : \mu = \mu_0$	$T = \frac{\bar{X} - \mu_0}{S / \sqrt{n}}$
z test for a proportion	$H_0 : p = p_0$	$Z = \frac{\hat{p} - p_0}{\sqrt{p_0(1-p_0)/n}}$

Alternative hypothesis		P-value for z test
(Upper-tailed test)	$H_a : \mu > \mu_0$	$P(Z \geq z)$
(Lower-tailed test)	$H_a : \mu < \mu_0$	$P(Z \leq z)$
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Given a fixed significance level α , we reject H_0 if and only if $P \leq \alpha$.