

# Math 3070, Applied Statistics

## Section 1

October 4, 2019

## Section 5.1

- Joint Probability Density Functions
- Marginal PDFs
- Independence
- Conditional PDFs

# Joint Probability Density

Random variables  $X$  and  $Y$  are said to have a **joint probability density** (joint pdf)  $f(x, y)$  if

$$\begin{aligned} P(a \leq X \leq b, c \leq Y \leq d) &= P(a \leq X \leq b \cap c \leq Y \leq d) \\ &= \int_a^b \int_c^d f(x, y) \, dy \, dx \end{aligned}$$

for all real constants  $a \leq b$  and  $c \leq d$ .

Need  $f(x, y) \geq 0$  and  $\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y) \, dy \, dx = 1$

## Example, Check PDF

A bank operates a drive-up facility and a walk-up window. On a random day, let  $X$  be the proportion of time the drive-up facility is in use, and let  $Y$  be the proportion of time the walk-up window is in use. Check that a valid joint pdf for  $X$  and  $Y$  is given by

$$f(x, y) = \begin{cases} \frac{6}{5}(x + y^2), & 0 \leq x \leq 1, 0 \leq y \leq 1 \\ 0, & \text{otherwise} \end{cases}$$

$$f(x, y) \geq 0$$

$$\begin{aligned} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y) \, dy \, dx &= \int_0^1 \int_0^1 \frac{6}{5}(x + y^2) \, dy \, dx \\ &= \frac{6}{5} \int_0^1 \left( xy + \frac{y^3}{3} \right) \Big|_{y=0}^1 \, dx \\ &= \frac{6}{5} \int_0^1 \left( x + \frac{1}{3} \right) \, dx = \frac{6}{5} \left( \frac{1}{2} + \frac{1}{3} \right) = 1 \end{aligned}$$

# Marginal Probability Density Function

Let  $X$  and  $Y$  be random variables with joint pdf  $f(x, y)$ . The **marginal probability density function** (marginal pdf) of  $X = x$  is defined as

$$f_X(x) = \int_{-\infty}^{\infty} f(x, y) \, dy$$

Similarly, the marginal pdf of  $Y$  is

$$f_Y(y) = \int_{-\infty}^{\infty} f(x, y) \, dx$$

Note:  $f_X$  is simply the pdf of  $X$  considered as a random variable on its own. Same as  $f_X(x)$  for PDFs.

## Example, Marginal PDF

In the bank example, the joint pdf of two random variables  $X$  and  $Y$  was given by

$$f(x, y) = \begin{cases} \frac{6}{5}(x + y^2), & 0 \leq x \leq 1, 0 \leq y \leq 1 \\ 0, & \text{otherwise} \end{cases}$$

Find the marginal pdf of  $X$ .

$$\begin{aligned} f_X(x) &= \int_{-\infty}^{\infty} f(x, y) \, dy \\ &= \int_0^1 \frac{6}{5}(x + y^2) \, dy \\ &= \frac{6}{5} \left( xy + \frac{y^3}{3} \right) \Big|_{y=0}^1 = \frac{6}{5} \left( x + \frac{1}{3} \right) \end{aligned}$$

# Independence

$X$  and  $Y$  are **independent** continuous random variables if their joint pdf is the product of the marginal pdfs:

$$f(x, y) = f_X(x)f_Y(y)$$

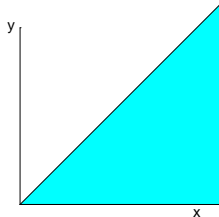
Their events are also independent,

$$P(a < X < b \cap c < Y < d) = P(a < X < b)P(c < Y < d).$$

## Example, Independence

A system has two components; assuming their lifetimes  $X$  and  $Y$  are independent exponential random variables with  $E(X) = 10$  and  $E(Y) = 20$ , what is the probability that the first component outlasts the second?

$$\begin{aligned}P(X \geq Y) &= \int_0^{\infty} \int_y^{\infty} f(x, y) \, dx \, dy \\&= \int_0^{\infty} \int_y^{\infty} \frac{1}{200} e^{-\frac{x}{10} - \frac{y}{20}} \, dx \, dy \\&= \frac{1}{200} \int_0^{\infty} e^{-y/20} \int_y^{\infty} e^{-x/10} \, dx \, dy \\&= \frac{1}{200} \int_0^{\infty} e^{-y/20} \cdot 10e^{-y/10} \, dy \\&= \frac{1}{20} \int_0^{\infty} e^{-\frac{3}{20}y} \, dy = \frac{1}{20} \cdot \frac{20}{3} = \frac{1}{3}\end{aligned}$$





# Conditional Probability Density Function

Let  $X$  and  $Y$  be random variables with joint pdf  $f(x, y)$ . The **conditional probability density function** of  $Y$  given  $X$  is

$$f_{Y|X}(y | x) = \frac{f(x, y)}{f_X(x)}$$

If  $X$  and  $Y$  are independent, then the conditional pdf of  $Y$  given  $X$  is simply the marginal pdf of  $Y$ :

$$f_{Y|X}(y | x) = \frac{f(x, y)}{f_X(x)} = \frac{f_X(x)f_Y(y)}{f_X(x)} = f_Y(y)$$

$$P(a < Y < b | X = x) = \int_a^b f_{Y|X}(y|x) dy$$

Note: The event  $X = x$  has zero probability. Typically, this is used when one value is already measured. Note, this would be a different calculation for events involving two random variables.

## Example, Conditional Probability Density Function

Recall the bank example.  $X$  is the proportion of time the drive-up facility is in use and  $Y$  is the proportion of time the walk-up window is in use. They have the following joint PDF,

$$f(x, y) = \begin{cases} \frac{6}{5}(x + y^2), & 0 \leq x \leq 1, 0 \leq y \leq 1 \\ 0, & \text{otherwise} \end{cases}$$

Given that drive-up facility is always in use, what is the probability that the walk-up facility is in use less than half the time?

We are conditioning on the event that  $X = 1$ .

We can use the marginal from the previous example.

$$f_X(x) = \frac{6}{5} \left( x + \frac{1}{3} \right)$$

$$f_{Y|X}(Y|1) = \frac{f(1, y)}{f_X(1)} = \frac{\frac{6}{5}(1 + y^2)}{\frac{6}{5}(1 + \frac{1}{3})} = \frac{1 + y^2}{1 + \frac{1}{3}} = \frac{3}{4}(1 + y^2)$$

## Example, Conditional Probability Density Function

$$f_{Y|X}(y|1) = \begin{cases} \frac{3}{4}(1 + y^2), & 0 \leq y \leq 1 \\ 0, & \text{otherwise} \end{cases}$$

$$\begin{aligned} P\left(Y < \frac{1}{2} \middle| X = 1\right) &= \int_{-\infty}^{\frac{1}{2}} f_{Y|X}(y|1) dy \\ &= \int_0^{\frac{1}{2}} \frac{3}{4}(1 + y^2) dy \\ &= \frac{3}{4} \left[ y + \frac{y^3}{3} \right]_{y=0}^{1/2} \\ &= \frac{3}{4} \left[ \frac{1}{2} + \frac{1}{24} \right] \approx 0.40625 \end{aligned}$$

# Non-Example, Conditional Probability Density Function

Recall the bank example.  $X$  is the proportion of time the drive-up facility is in use and  $Y$  is the proportion of time the walk-up window is in use. They have the following joint PDF,

$$f(x, y) = \begin{cases} \frac{6}{5}(x + y^2), & 0 \leq x \leq 1, 0 \leq y \leq 1 \\ 0, & \text{otherwise} \end{cases}$$

Given that drive-up facility is in use less than half the time, what is the probability that the walk-up facility is in use less than half the time?

# Non-Example, Conditional Probability Density Function

$$P(Y < 1/2 | X < 1/2) = \frac{P(Y < 1/2 \cap X < 1/2)}{P(X < 1/2)}$$

$$= \frac{\int_{-\infty}^{1/2} \int_{-\infty}^{1/2} f(x, y) dx dy}{\int_{-\infty}^{1/2} f_X(x) dx}$$

$$\int_{-\infty}^{1/2} \int_{-\infty}^{1/2} f(x, y) dx dy = \int_0^{1/2} \int_0^{1/2} \frac{6}{5}(x + y^2) dx dy = \frac{1}{5}$$

$$\int_{-\infty}^{1/2} f_X(x) dx = \int_0^{1/2} \frac{6}{5}(x + 1/3) dx = \frac{7}{20}$$

$$P(Y < 1/2 | X < 1/2) = \frac{1/5}{7/20} = \frac{4}{7}$$

# Joint PDF, Summary

- Joint PDF:  $f(x, y)$  such that
$$P(a \leq X \leq b, c \leq Y \leq d) = \int_a^b \int_c^d f(x, y) dy dx$$
- Marginal PDF:  $f_X(x) = \int_{-\infty}^{\infty} f(x, y) dy$ . Is a univariate PDF.
- If  $X$  and  $Y$  are independent,  $f(x, y) = f_X(x)f_Y(y)$ .  
Their events are also independent
$$P(a < X < b \cap c < Y < d) = P(a < X < b)P(c < Y < d)$$
- Conditional PDF:

$$f_{Y|X}(y | x) = \frac{f(x, y)}{f_X(x)}$$

Note: the conditioned variable  $X$  is fixed,  $X = x$ . Think of this as a function of  $y$ . Different from conditioning on events.

- Can define these with more variables than two.

## Example, Sum of Two Random Variables

Suppose that  $X$  and  $Y$  are independent continuous random variables. Determine the distribution of  $Z = X + Y$ .  
Determine the CDF.

$$P(Z < z) = P(X + Y < z) = \int_{-\infty}^{\infty} \int_{-\infty}^{z-y} f(x, y) dx dy$$

on the inside integral, use  $x = v - y \rightarrow dx = dv$

$$= \int_{-\infty}^{\infty} \int_{-\infty}^z f(v - y, y) dv dy = \int_{-\infty}^z \int_{-\infty}^{\infty} f(v - y, y) dy dv$$

differentiate to determine  $f_Z(z)$

$$\frac{d}{dz} P(Z < z) = \int_{-\infty}^{\infty} f(z - y, y) dy$$

$$f_Z(z) = \int_{-\infty}^{\infty} f(z - y, y) dy$$

With independence,  $f_Z(z) = \int_{-\infty}^{\infty} f_X(z - y) f_Y(y) dy$ .

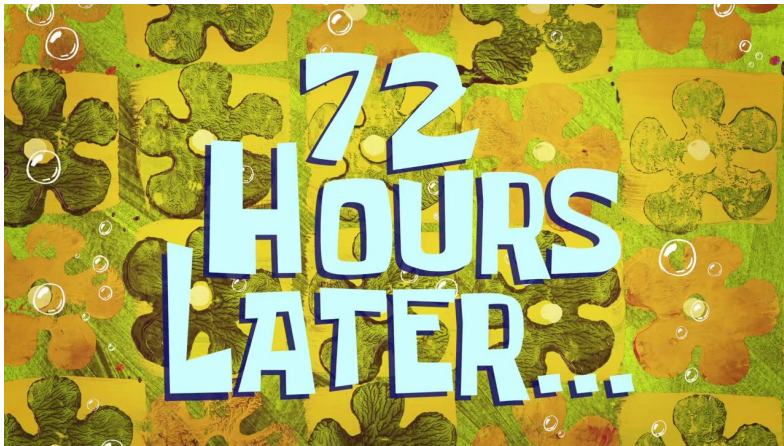
## Example, Sum of Two Independent Normals

Suppose that  $X \sim N(\mu_1, \sigma_1)$  and  $Y \sim N(\mu_2, \sigma_2)$  are independent. Determine the distribution of  $Z = X + Y$ .

$$\begin{aligned} f_Z(z) &= \int_{-\infty}^{\infty} f_X(z-y)f_Y(y)dy \\ &= \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}\sigma_1} \exp\left(\frac{-(z-y-\mu_1)^2}{2\sigma_1^2}\right) \frac{1}{\sqrt{2\pi}\sigma_2} \exp\left(\frac{-(y-\mu_2)^2}{2\sigma_2^2}\right) dy \end{aligned}$$



# Example, Sum of Two Independent Normals



## Example, Sum of Two Independent Normals

$$f(z) = \frac{1}{\sqrt{2\pi}\sqrt{\sigma_1^2 + \sigma_2^2}} \exp\left(\frac{-(z - (\mu_1 + \mu_2))}{2(\sigma_1^2 + \sigma_2^2)}\right)$$

$$Z \sim N\left(\mu_1 + \mu_2, \sqrt{\sigma_1^2 + \sigma_2^2}\right)$$

[https://en.wikipedia.org/wiki/Sum\\_of\\_normally\\_distributed\\_random\\_variables#Proof\\_using\\_convolutions](https://en.wikipedia.org/wiki/Sum_of_normally_distributed_random_variables#Proof_using_convolutions)