# Math 3070, Applied Statistics

Section 1

October 16, 2019

### Lecture Outline, 10/16

#### Section 5.3

- Random Samples and Statistics
- Sampling Distributions

### Random Samples

- $X_1, X_2, \dots, X_n$  are a called a **simple random sample** of size n if
  - 1 The  $X_i$ 's are independent.
  - 2 Every  $X_i$  has the same probability distribution.

Note: These two assumptions are frequently referred to as independently and identically distributed (i.i.d.).

For example, measurements of parts from an assembly line or outcomes of sufficently randomized polls could be considered simple random sample. Framework is extermely flexible, can describe most settings, and be used to compute statistical errors.

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Examples:

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Both  $\overline{X}$  and  $S^2$  functions of  $X_i$  so they are random. But, they estimate  $E[X_i] = \mu$  and  $Var(X_i) = \sigma^2$  which are numbers and not random.

Check the expected value of  $S^2$  assuming  $\mu = E[X_i]$ ,  $\sigma^2 = Var(X_i)$ , and the  $X_i$ 's are i.i.d.

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$$\sum_{i=1}^{n} \left( X_i - \frac{1}{n} \sum_{j=1}^{n} X_j \right)^2 = \sum_{i=1}^{n} \left( X_i^2 - \frac{2}{n} X_i \sum_{j=1}^{n} X_j + \left( \frac{1}{n} \sum_{k=1}^{n} X_k \right)^2 \right)$$

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$$\begin{split} &\sum_{i=1}^{n} \left( X_{i} - \frac{1}{n} \sum_{j=1}^{n} X_{j} \right)^{2} = \sum_{i=1}^{n} \left( X_{i}^{2} - \frac{2}{n} X_{i} \sum_{j=1}^{n} X_{j} + \left( \frac{1}{n} \sum_{k=1}^{n} X_{k} \right)^{2} \right) \\ &= \left( \sum_{i=1}^{n} X_{i}^{2} \right) - \frac{2}{n} \left( \sum_{i=1}^{n} \sum_{j=1}^{n} X_{i} X_{j} \right) + \frac{1}{n^{2}} \sum_{i=1}^{n} \left( \sum_{j=1}^{n} X_{j} \right)^{2} \end{split}$$

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As n becomes very large, dividing by n or n-1 is mostly irrelevant.

# Sampling Distributions

Sampling distributions are the PMF/PDF of a statistic.

Weights X of randomly selected seeds follow PMF f. Compute the sampling distribution of the sample mean of two seeds weights.

$$f(x) = \begin{cases} 0.25, & x = 1 \\ 0.50, & x = 2 \\ 0.25, & x = 3 \end{cases}$$

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 | 1 | 1.5 | 2 | 2.5 | 3   
  $f_X(x)$  | 0.0625 | 0.25 | 0.375 | 0.25 | 0.0625

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$$P(\overline{X} = 1) = P(X_1 = 1 \cap X_2 = 1)$$

$$= P(X_1 = 1)P(X_2 = 1) = 0.25^2 = 0.0625$$

$$P(\overline{X} = 1.5) = P(X_1 = 2 \cap X_2 = 1) + P(X_1 = 1 \cap X_2 = 2)$$

$$= 0.5(0.25) + 0.25(0.5) = 0.25$$

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Since they have the same distribution,  $X_i = Y_i + 1$  can be used.

$$\overline{X} = \frac{X_1 + X_2}{2} = \frac{Y_1 + 1 + Y_2 + 1}{2} = \frac{Y_1 + Y_2}{2} + 1$$

$$= \frac{U_{1,1} + U_{1,2} + U_{2,1} + U_{2,2}}{2} + 1 = \frac{Z}{2} + 1$$

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Since Z is the sum of four independent bernoulli random variables, Z is also a bionomial random variable with n=4 and p=0.5.

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$\overline{X} = t$	1	1.5	2	2.5	3
$f_X(x)$	0.0625	0.25	0.375	0.25	0.0625
z=2(t-1)	0	1	2	3	4
$f_Z(z)$	0.0625	0.25	0.375	0.25	0.0625

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$$= p + p + 1 = 2$$

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$$= E[X_{i}] = 2$$

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$$Var(\overline{X}) = Var\left(\frac{X_{1} + X_{2}}{2}\right) = \frac{Var(X_{1}) + Var(X_{2})}{4} = \frac{Var(X_{i})}{2} = 0.25$$

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Fun Fact:

$$Var(S^2) = \frac{\mu_4}{n} - \frac{\sigma^2(n-3)}{n(n-1)}$$

where

$$\mu_4 = E[(X_i - \mu)^4]$$

https://math.stackexchange.com/questions/2476527/variance-of-sample-variance

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$$\overline{X} = \frac{1}{n} \sum_{i=1}^{n} X_i \qquad \overline{X} \sim N\left(9, \frac{4}{\sqrt{n}}\right)$$

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$$= \int_{-\infty}^{\infty} \int_{-\infty}^{2t - y} f_{T_1}(x) f_{T_2}(y) dx dy$$

$$= \int_{-\infty}^{\infty} f_{T_2}(y) \int_{-\infty}^{2t - y} f_{T_1}(x) dx dy$$

Suppose that  $T_i$ , wait tines between redlights, follow an exponential distribution with mean 2. Compute the PDF of the sample average of two wait times.

$$\overline{T} = \frac{T_1 + T_2}{2}$$

$$P(\overline{T} < t) = P(T_1 + T_2 < 2t) = P(T_1 < 2t - T_2)$$

$$= \int_{-\infty}^{\infty} \int_{-\infty}^{2t - y} f_{T_1}(x) f_{T_2}(y) dx dy$$

$$= \int_{-\infty}^{\infty} f_{T_2}(y) \int_{-\infty}^{2t - y} f_{T_1}(x) dx dy$$

$$= \int_{-\infty}^{\infty} f_{T_2}(y) F_{T_1}(2t - y) dy$$

$$\frac{d}{dt}P(\overline{T} < t) = \frac{d}{dt} \int_{-\infty}^{\infty} f_{T_2}(y)F_{T_1}(2t - y)dy$$

$$\frac{d}{dt}P(\overline{T} < t) = \frac{d}{dt} \int_{-\infty}^{\infty} f_{T_2}(y)F_{T_1}(2t - y)dy$$
$$= \int_{-\infty}^{\infty} f_{T_2}(y)\frac{d}{dt}F_{T_1}(2t - y)dy$$

$$\frac{d}{dt}P(\overline{T} < t) = \frac{d}{dt} \int_{-\infty}^{\infty} f_{T_2}(y)F_{T_1}(2t - y)dy$$
$$= \int_{-\infty}^{\infty} f_{T_2}(y)\frac{d}{dt}F_{T_1}(2t - y)dy$$
$$= \int_{-\infty}^{\infty} 2f_{T_2}(y)f_{T_1}(2t - y)dy$$

$$\frac{d}{dt}P(\overline{T} < t) = \frac{d}{dt} \int_{-\infty}^{\infty} f_{T_2}(y)F_{T_1}(2t - y)dy 
= \int_{-\infty}^{\infty} f_{T_2}(y)\frac{d}{dt}F_{T_1}(2t - y)dy 
= \int_{-\infty}^{\infty} 2f_{T_2}(y)f_{T_1}(2t - y)dy 
\text{note: } y > 2t \to f_{T_1}(2t - y) = 0 
= \int_{0}^{2t} 2\frac{\exp(-y/2)}{2}\frac{\exp(-t + y/2)}{2}dy$$

$$\frac{d}{dt}P(\overline{T} < t) = \frac{d}{dt} \int_{-\infty}^{\infty} f_{T_2}(y)F_{T_1}(2t - y)dy 
= \int_{-\infty}^{\infty} f_{T_2}(y)\frac{d}{dt}F_{T_1}(2t - y)dy 
= \int_{-\infty}^{\infty} 2f_{T_2}(y)f_{T_1}(2t - y)dy 
\text{note: } y > 2t \to f_{T_1}(2t - y) = 0 
= \int_{0}^{2t} 2\frac{\exp(-y/2)}{2}\frac{\exp(-t + y/2)}{2}dy 
= \int_{0}^{2t} \frac{\exp(-t)}{2}dy = \frac{t}{2}\exp(-t)$$

$$\frac{d}{dt}P(\overline{T} < t) = \frac{d}{dt} \int_{-\infty}^{\infty} f_{T_2}(y)F_{T_1}(2t - y)dy 
= \int_{-\infty}^{\infty} f_{T_2}(y)\frac{d}{dt}F_{T_1}(2t - y)dy 
= \int_{-\infty}^{\infty} 2f_{T_2}(y)f_{T_1}(2t - y)dy 
\text{note: } y > 2t \to f_{T_1}(2t - y) = 0 
= \int_{0}^{2t} 2\frac{\exp(-y/2)}{2}\frac{\exp(-t + y/2)}{2}dy 
= \int_{0}^{2t} \frac{\exp(-t)}{2}dy = \frac{t}{2}\exp(-t) 
\overline{T} \sim Gamma(2, 1)$$

### Summary

- Simple random sample or independent identically distributed assumption applies in many cases, but not all.
- Sampling distribution can model the probablistic behavior of statistics, but are often hard to find.
- Sum of Normal Random Variables is a Normal Random Variable. It's mean and standard deviation must be identified.
- Assuming  $X_i$ 's are i.i.d.

$$E[\overline{X}] = E[X_i] = \mu$$
  $Var(\overline{X}) = \frac{Var(X_i)}{n} = \frac{\sigma}{n}$   
 $E[S^2] = Var(X_i) = \sigma^2$