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For a linear smoother, that is $\hat{y} = Sy$

$\begin{matrix} & \swarrow & \\ & \text{N} \times \text{N symmetric matrix} & \\ & \searrow & \\ \uparrow & & \nwarrow \\ \text{N} \times 1 & & \text{N} \times 1 \end{matrix}$

show that $\sum_{i=1}^N \text{Cov}(y_i, \hat{y}_i) = \text{trace}(S) \cdot \sigma_e^2$

Where did σ_e^2 come from? We assume an additive error model

$$y_i = f(x_i) + \epsilon_i \quad \text{where } E(\epsilon_i) = 0, \text{Var}(\epsilon_i) = \sigma_e^2$$

and ϵ_i and ϵ_j independent

(Have we seen this before? Yes, in MLR the Hat matrix is the linear smoother: $\hat{\beta} = (X^T X)^{-1} X^T y$ and $\hat{y} = X \hat{\beta} = \underbrace{X(X^T X)^{-1} X^T}_{H} y$
 (one possible version of S)

Start with:

$$\begin{matrix} \text{N} \times \text{N matrix} \\ \uparrow \quad \uparrow \\ \text{N} \times 1 \quad \text{N} \times 1 \end{matrix} \text{Cov}(y, \hat{y}) = \begin{bmatrix} \text{Cov}(y_1, \hat{y}_1) & \text{Cov}(y_1, \hat{y}_2) & \dots & \text{Cov}(y_1, \hat{y}_N) \\ \text{Cov}(y_2, \hat{y}_1) & & & \\ \vdots & & & \\ \text{Cov}(y_N, \hat{y}_1) & & & \text{Cov}(y_N, \hat{y}_N) \end{bmatrix}$$

Then to get

$$\begin{aligned} \sum_{i=1}^N \text{Cov}(y_i, \hat{y}_i) &= \underbrace{\text{sum of diagonal elements of } \text{Cov}(y, \hat{y})}_{\text{def. of the trace}} \\ &= \text{trace}(\text{Cov}(y, \hat{y})) \end{aligned}$$

$\begin{matrix} \uparrow & \uparrow \\ \text{N} \times 1 & \text{N} \times 1 \end{matrix}$

$$= \text{trace} \left(\underset{N \times 1}{\text{Cov} \left(\underset{N \times N}{y}, \underset{N \times 1}{S \hat{y}} \right)} \right) = \text{trace} \left(S \cdot \underbrace{\text{Cov}(y, y)}_{\substack{\text{Var}(y) \\ = \\ \sigma_e^2 \cdot I}} \right) = \text{trace}(S) \cdot \sigma_e^2$$

[formulas from TMA4267]

since y_i, y_j independent

trace $\begin{pmatrix} s_{11} & \dots & s_{1N} \\ \vdots & \ddots & \vdots \\ s_{N1} & \dots & s_{NN} \end{pmatrix} \cdot \begin{bmatrix} \sigma_e^2 & 0 & \dots & 0 \\ 0 & \ddots & & \\ \vdots & & \ddots & \\ 0 & \dots & 0 & \sigma_e^2 \end{bmatrix}$

$$\underset{a \times b}{\text{Cov} \left(\underset{a \times 1}{V}, \underset{b \times 1}{W} \right)} = E \left((V - \mu_V) (W - \mu_W)^T \right) \quad [\text{here } a=b]$$

and if $W = SV$:

$$\begin{aligned} \text{Cov}(V, SV) &= E \left((V - \mu_V) (SV - \mu_W)^T \right) = E \left((V - \mu_V) (SV - \hat{\mu}_V)^T \right) \\ &= E \left((V - \mu_V) (V - \mu_V)^T S^T \right) = \text{Cov}(V) \cdot S^T \end{aligned}$$

can be derived from Cov of linear combination

$$\text{trace}(AB^T)$$

$$= \text{trace}(A^T B)$$

see for example

Härdle, Smeier (2015) Applied Multivariate

Statistical Analysis, Springer

(ebook at NTNU)

chapter 4.2 (link in L1.html)
to book & notes

or classnotes from TMA4267

$$\begin{aligned} &\text{trace}(\text{Cov}(V) \cdot S^T) \\ &= \text{trace} \left(S \underset{\substack{\uparrow \\ \text{symmetric}}}{\text{Cov}(V)} \right) \end{aligned}$$