1. Title of the first chapter

Definition 1 (Simple random walk in \mathbb{Z}). Let $\{X_n\}_{n=0}^{+\infty}$ be a sequence of independent and identically distributed random variables with values in $\{-1, +1\}$. That $\forall n \in \mathbb{N}$ satisfy $P(X_n = 1) = p \in (0, 1)$ and $P(X_n = -1) = 1 - p = q$. Let $S_0 = 0$ and $S_n = \sum_{i=1}^n X_i$. We call the pair $(\{S_n\}_{n=0}^{+\infty}, p)$ Simple random walk in \mathbb{Z} . In case that $p = q = \frac{1}{2}$ we call the pair $(\{S_n\}_{n=0}^{+\infty}, p)$ Symmetric simple random walk in \mathbb{Z} .

Remark. Very often we refer to n as time, X_i as i-th step and S_n as position in time n. In simple random walk in \mathbb{Z} we refer to $X_i = +1$ as i-th step was rightwards. If not stated otherwise, we

Definition 2 (Set of possible positions). Let $(\{S_n\}_{n=0}^{+\infty}, p)$ be a random walk. We call the set $A_n = \{z \in \mathbb{Z}; |z| \le n, \frac{z+n}{2} \in \mathbb{Z}\}$ set of all possible positions of random walk $(\{S_n\}_{n=0}^{+\infty}, p)$ in time n.

Theorem 1 (Probability of position x in time n). Let $(\{S_n\}_{n=0}^{+\infty}, p)$ be a random walk and A_n its set of possible positions.

$$P(S_n = x) = \begin{cases} \binom{n}{\frac{x+n}{2}} p^{\frac{n+x}{2}} q^{\frac{n-x}{2}} & \text{for } x \in A_n, \\ 0, & \text{for } x \notin A_n. \end{cases}$$

Proof. Let us define new variables $r_i = \mathbf{1}_{[X_i=1]}, l_i = \mathbf{1}_{[X_i=-1]}, R_n = \sum_{i=1}^n r_i, L_n = \sum_{i=1}^n l_i$. r_i can be interpreted as indicator wether i-th step was rightwards. Then R_n is number of rightwards steps and L_n is number of leftwards steps. We can easily see that $R_n + L_n = n$ and $R_n - L_n = S_n$. Therefore we get by adding these two equations $R_n = \frac{S_n + n}{2}$.

 r_i has alternative distribution with parameter p (Alt(p)). Therefore R_n as a sum of independent and identically distributed and wariables with Alt(p) has binomial distribution with parameters n and p (Bi(n,p)). Therefore we get $P(R_n = x) = \binom{n}{\frac{x+n}{2}} p^{\frac{n+x}{2}} q^{\frac{n-x}{2}}$. Where we define $\binom{a}{x} := 0$ for $a \in \mathbb{N}, x \in \mathbb{R} \setminus \mathbb{N}, x < 0, x > n$. Therefore we get $P(S_n = x) = P\left(R_n = \frac{x+n}{2}\right) = \binom{n}{\frac{x+n}{2}} p^{\frac{n+x}{2}} q^{\frac{n-x}{2}}$. \square

Lemma 2 (Spatial homogeneity). $P(S_n = j | S_0 = a) = P(S_n = j + b | S_0 = a + b) \forall b \in \mathbb{Z}$

Proof.
$$P(S_n = j | S_0 = a) = P\left(\sum_{i=1}^n X_i = j - a\right) = P\left(\sum_{i=1}^n X_i = (j+b) - (a+b)\right) = P(S_n = j + b | S_0 = a + b).$$

Lemma 3 (Temporal homogeneity). $P(S_n = j | S_0 = a) = P(S_{n+m} = j | S_m = a) \forall m \in \mathbb{N}$

Proof.
$$P(S_n = j | S_0 = a) = P\left(\sum_{i=1}^n X_i = j - a\right) = P\left(\sum_{i=m+1}^{m+n} X_i = j - a\right) = P(S_{n+m} = j | S_m = a).$$

Lemma 4 (Markov property). Let $n \ge m$ and $a_i \in \mathbb{Z}$. Then $P(S_n = j | S_0 = a_0, S_1 = a_1, \dots, S_m = a_m) = P(S_n = j | S_m = a_m)$

Proof. Once S_m is known, then distribution of S_n depends only on steps $X_{m+1}, X_{m+2}, \ldots X_n$ and therefore cannot be dependent on any information concerning values X_1, X_2, \ldots, X_m-1 and accordingly S_1, S_2, \ldots, S_m-1 .

Remark. In symmetric random walk, everything can be counted by number of possible paths from point to point.

Definition 3 (Number of possible paths). Let $N_n(a,b)$ be number of possible paths of random walk $(\{S_n\}_{n=0}^{+\infty}, p)$ from point (0,a) to point (n,b) and $N_n^x(a,b)$ be number of possible paths from point (0,a) to point (n,b) that visit point (z,x) for some $z \in \{0,\ldots n\}$.

Theorem 5. Let $a, b \in \mathbb{Z}, n \in \mathbb{N}$ then $N_n(a, b) = \binom{n}{\frac{1}{2}(n+b-a)}$.

Proof. Let us choose a path from (0, a) to (n, b) and let α be number of rightwards steps and β be number of leftwards steps. Then $\alpha + \beta = n$ and $\alpha - \beta = b - a$. By adding these two equations we get that $\alpha = \frac{1}{2}(n+b-a)$. The number of possible paths is the number of ways of picking α rightwards steps from n steps. Therefore we get $N_n(a,b) = \binom{n}{\alpha} = \binom{n}{\frac{1}{2}(n+b-a)}$.

Theorem 6 (Reflection principle). Let a, b > 0, then $N_n^0(a, b) = N_n(-a, b)$.

Proof. Each path from (0, -a) to (n, b) has to intersect x-axis at least once at some point. Let k be the time of earliest intersection with x-axis. By reflexing the segment from (0, -a) to (k, 0) in the x-axis, we get a path from point (0, a) to (n, b) which visits 0 at point k. Because reflection is bijective operation on sets of paths, we get the correspondence between the collections of such paths. \square

Definition 4 (Return to origin). Let $(\{S_n\}_{n=0}^{+\infty}, p)$ be a random walk. Then if $\exists k \in \mathbb{N}$ such that $S_k = 0$ then we say that in k-th step occurred return to origin. Let $(\{S_n\}_{n=0}^{+\infty}, p)$ be a random walk. Then if $S_1, S_2, \ldots, S_{2n-1} \neq 0$ and $S_{2n} = 0$

Theorem 7 (Ballot theorem). Let $n, b \in N$ Number of paths from point (0,0) to point (n,b) which do not return to origin is equal to $\frac{b}{n}N_n(0,b)$

Proof. Let us call N the number of paths we are referring to. Because the path ends at point (n,b), the first step has to be rightwards. Therefore we now have $N = N_{n-1}(1,b) - N_{n-1}^0(1,b) = N_{n-1}(1,b) - N_{n-1}(-1,b)$. The last equation was aquired using Ballot theorem (6). We now have:

$$N_{n-1}(1,b) - N_{n-1}(-1,b) = \binom{n-1}{\frac{n}{2} + \frac{b}{2} - 1} - \binom{n-1}{\frac{n}{2} + \frac{b}{2}} = \frac{(n-1)!}{(\frac{n}{2} + \frac{b}{2} - 1)!(\frac{n}{2} - \frac{b}{2})!} - \frac{(n-1)!}{(\frac{n}{2} + \frac{b}{2})!(\frac{n}{2} - \frac{b}{2} - 1)!}$$

Definition 5. $M_n = \max\{S_i, i \in \{0, 1, ..., n\}\}$

Theorem 8 (Probability of maximum up to time n). Let $(\{S_n\}_{n=0}^{+\infty}, p)$ be a random walk.

$$P(M_n \ge r, S_n = b) = \begin{cases} P(S_n = b) & \text{for } b \ge r, \\ P(S_n = 2r - b) \left(\frac{q}{p}\right)^{r-b}, & \text{for otherwise.} \end{cases}$$

Proof. Let us firstly consider the easier case in which $b \geq r$. Because we defined M_n as $\max\{S_i, i \in \{0, 1, \dots, n\}\}$ we get that $M_n \geq b \geq r$ therefore $[M_n \geq r] \subset [S_n = b]$ therefore we get $\mathsf{P}(M_n \geq r, S_n = b) = \mathsf{P}(S_n = b)$. Let $r \geq 1, b < r$. $N_n^r(0,b)$ stands for number of paths from point (0,0) to point (n,b) which reach up to r. Let $k \in \{0,1,\ldots,n\}$ denote the first time we reach r. By reflection principle (6), we can reflex the segment from (k,r) to (n,b) in the axis:y = r. Therefore we now have path from (0,0) to (n,2r-b) and we get that $N_n^r(0,b) = N_n(0,2r-b)$. $\mathsf{P}(S_n = b, M_n \geq r) = N_n^r(0,b)p^{\frac{n+b}{2}}q^{\frac{n-b}{2}} = N_n(0,2r-b)p^{\frac{n+(2r-b)}{2}}q^{\frac{n-(2r-b)}{2}}p^{\frac{n+b}{2}}q^{\frac{n-b}{2}}p^{b-r}q^{r-b} = \left(\frac{q}{p}\right)^{r-b}\mathsf{P}(S_n = 2r-b)$. \square

Definition 6 (Walk reaching new maximum at particular time). Let b > 0. $f_b(n)$ denotes the probability that we reach new maximum b in time n. $f_b(n) = P(M_{n-1} = S_{n-1} = b - 1, S_n = b)$

Theorem 9 (Probability of reaching new maximum b in time n). Let b > 0 then $f_b(n) = \frac{b}{n} P(S_n = b)$.

$$\begin{array}{l} \textit{Proof.} \ \ f_b = \mathsf{P}(M_{n-1} = S_{n-1} = b-1, S_n = b) = p \, \mathsf{P}(M_{n-1} = S_{n-1} = b-1) = \\ p \, (\mathsf{P}(M_{n-1} \geq b-1, S_{n-1} = b-1) - \mathsf{P}(M_{n-1} \geq b, S_{n-1} = b-1)) = p \, \left(\mathsf{P}(S_{n-1} = b-1) - \frac{q}{p} \, \mathsf{P}(S_{n-1} = b-1) - \frac{q}{p} \,$$

Theorem 10 (XXXMean number of visits to b before returning to origin in symmetric random walk). Let $(\{S_n\}_{n=0}^{+\infty}, p)$ be a symmetric random walk. Mean number μ_b of visits of the walk to point b before returning to origin is equal to 1.

Proof. aa
$$\Box$$

Definition 7 (Return to origin). Let $(\{S_n\}_{n=0}^{+\infty}, p)$ be a random walk. Let $k \in \mathbb{N}$. We say a return to origin occurred in time 2k if $S_{2k} = 0$. The probability that in time 2k occurred a return to origin shall be denoted by u_{2k} . We say that in time 2k occurred first return to origin if $S_1, S_2, \ldots S_{2k-1} \neq 0$ and $S_{2k} = 0$. The probability that in time 2k occurred first return to origin shall be denoted by f_{2k} . By definition $f_0 = 0$.

Lemma 11 (Binomial identity). Let
$$n, k \in \mathbb{N}, n > k : \binom{n-1}{k} - \binom{n-1}{k-1} = \frac{n-2k}{n} \binom{n}{k}$$

Proof. $\binom{n-1}{k} - \binom{n-1}{k-1} = \frac{(n-1)!}{k!(n-k-1)!} - \frac{(n-1)!}{(k-1)!(n-k)!} = \frac{(n-1)!}{(k-1)!(n-k-1)!} \left(\frac{1}{k} - \frac{1}{n-k}\right) = \frac{1}{n} \frac{n!}{(k-1)!(n-k-1)!} \frac{n-2k}{k(n-k)} = \frac{n-2k}{n} \binom{n}{k}$

 $\frac{n-2k}{n} \frac{n!}{k!(n-k)!} = \frac{n-2k}{n} \binom{n}{k}$

Lemma 12 (Main lemma). Let $(\{S_n\}_{n=0}^{+\infty}, p)$ be a symmetrical random walk. Then $P(S_1 \cdot S_2, \cdot, \dots, S_{2n} \neq 0) = P(S_{2n} = 0)$

Proof.
$$P(S_1 \cdot S_2, \cdot, \dots, S_{2n} \neq 0) = \sum_{i=-\infty}^{+\infty} P(S_1 \cdot S_2, \cdot, \dots, S_{2n-1} \neq 0, S_{2n} = 2i) = \sum_{i=-n}^{n} P(S_1 \cdot S_2, \cdot, \dots, S_{2n-1} \neq 0, S_{2n} = 2i) = 2 \cdot \sum_{i=1}^{n} P(S_1 \cdot S_2, \cdot, \dots, S_{2n-1} \neq 0, S_{2n} = 2i) = 2 \cdot \sum_{i=1}^{n} P(S_1 \cdot S_2, \cdot, \dots, S_{2n-1} \neq 0, S_{2n} = 2i) = 2 \cdot \sum_{i=1}^{n} \frac{2i}{2n} P(S_{2n} = 2k) = 2 \cdot \sum_{i=1}^{n} \frac{2i}{2n} \binom{2n}{n+i} 2^{-2n} = 2 \cdot 2^{-2n} \sum_{i=1}^{n} \binom{2n-1}{m+k-1} - \binom{2n-1}{m+k} = 2 \cdot 2^{-2n} \binom{2n-1}{n} = 2 \cdot 2^{-2n} \binom{2n}{n} = 2$$

- 1.2 Title of the second subchapter of the first chapter

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List of Figures

List of Tables

List of Abbreviations

A. Attachments

A.1 First Attachment