

1. Title of the first chapter

Definition 1 (Simple random walk in \mathbb{Z})

Let $\{X_n\}_{n=0}^{+\infty}$ be a sequence of independent and identically distributed random variables with values in $\{-1, +1\}$. That $\forall n \in \mathbb{N}$ satisfy $P(X_n = 1) = p \in (0, 1)$ and $P(X_n = -1) = 1 - p = q$. Let $S_0 = 0$ and $S_n = \sum_{i=1}^n X_i$. We call the pair $(\{S_n\}_{n=0}^{+\infty}, p)$ Simple random walk in \mathbb{Z} . In case that $p = q = \frac{1}{2}$ we call the pair $(\{S_n\}_{n=0}^{+\infty}, p)$ Symmetric simple random walk in \mathbb{Z} .

Remark. Very often we refer to n as time, X_i as i -th step and S_n as position in time n . In simple random walk in \mathbb{Z} we refer to $X_i = +1$ as i -th step was rightwards. If not stated otherwise, we assume that $S_0 = 0$.

Definition 2 (Set of possible positions)

Let $(\{S_n\}_{n=0}^{+\infty}, p)$ be a random walk. We call the set $A_n = \{z \in \mathbb{Z}; |z| \leq n, \frac{z+n}{2} \in \mathbb{Z}\}$ set of all possible positions of random walk $(\{S_n\}_{n=0}^{+\infty}, p)$ in time n .

Theorem 1 (Probability of position x in time n)

Let $(\{S_n\}_{n=0}^{+\infty}, p)$ be a random walk and A_n its set of possible positions.

$$P(S_n = x) = \begin{cases} \binom{n}{\frac{x+n}{2}} p^{\frac{n+x}{2}} q^{\frac{n-x}{2}} & \text{for } x \in A_n, \\ 0, & \text{for } x \notin A_n. \end{cases}$$

Proof. Let us define new variables $r_i = \mathbf{1}_{[X_i=1]}, l_i = \mathbf{1}_{[X_i=-1]}, R_n = \sum_{i=1}^n r_i, L_n = \sum_{i=1}^n l_i$. r_i can be interpreted as indicator whether i -th step was rightwards. Then R_n is number of rightwards steps and L_n is number of leftwards steps. We can easily see that $R_n + L_n = n$ and $R_n - L_n = S_n$. Therefore we get by adding these two equations $R_n = \frac{S_n + n}{2}$.

r_i has alternative distribution with parameter p $Alt(p)$. Therefore R_n as a sum of independent and identically distributed random variables with $Alt(p)$ has binomial distribution with parameters n and p ($Bi(n, p)$). Therefore we get $P(R_n = x) = \binom{n}{\frac{x+n}{2}} p^{\frac{n+x}{2}} q^{\frac{n-x}{2}}$. Where we define $\binom{a}{x} := 0$ for $a \in \mathbb{N}, x \in \mathbb{R} \setminus \mathbb{N}, x < 0, x > n$. Therefore we get $P(S_n = x) = P(R_n = \frac{x+n}{2}) = \binom{n}{\frac{x+n}{2}} p^{\frac{n+x}{2}} q^{\frac{n-x}{2}}$. \square

Lemma 2 (Spatial homogeneity)

$$P(S_n = j \mid S_0 = a) = P(S_n = j + b \mid S_0 = a + b) \forall b \in \mathbb{Z}$$

Proof. $P(S_n = j \mid S_0 = a) = P\left(\sum_{i=1}^n X_i = j - a\right)$
 $= P\left(\sum_{i=1}^n X_i = (j + b) - (a + b)\right) = P(S_n = j + b \mid S_0 = a + b)$. \square

Lemma 3 (Temporal homogeneity)

$$P(S_n = j \mid S_0 = a) = P(S_{n+m} = j \mid S_m = a) \forall m \in \mathbb{N}$$

Proof. $P(S_n = j \mid S_0 = a) = P\left(\sum_{i=1}^n X_i = j - a\right)$
 $= P\left(\sum_{i=m+1}^{m+n} X_i = j - a\right) = P(S_{n+m} = j \mid S_m = a)$. \square

Lemma 4 (Markov property)

Let $n \geq m$ and $a_i \in \mathbb{Z}, i \in \mathbb{N}$. Then $P(S_n = j \mid S_0 = a_0, S_1 = a_1, \dots, S_m = a_m) = P(S_n = j \mid S_m = a_m)$

Proof. Once S_m is known, then distribution of S_n depends only on steps $X_{m+1}, X_{m+2}, \dots, X_n$ and therefore cannot be dependent on any information concerning values $X_1, X_2, \dots, X_m - 1$ and accordingly $S_1, S_2, \dots, S_m - 1$. \square

Remark. In symmetric random walk, everything can be counted by number of possible paths from point to point.

Definition 3 (Number of possible paths)

Let $N_n(a, b)$ be number of possible paths of random walk $(\{S_n\}_{n=0}^{+\infty}, p)$ from point $(0, a)$ to point (n, b) and $N_n^x(a, b)$ be number of possible paths from point $(0, a)$ to point (n, b) that visit point (z, x) for some $z \in \{0, \dots, n\}$.

Theorem 5

Let $a, b \in \mathbb{Z}, n \in \mathbb{N}$ then $N_n(a, b) = \binom{n}{\frac{1}{2}(n+b-a)}$.

Proof. Let us choose a path from $(0, a)$ to (n, b) and let α be number of rightwards steps and β be number of leftwards steps. Then $\alpha + \beta = n$ and $\alpha - \beta = b - a$. By adding these two equations we get that $\alpha = \frac{1}{2}(n + b - a)$. The number of possible paths is the number of ways of picking α rightwards steps from n steps. Therefore we get $N_n(a, b) = \binom{n}{\alpha} = \binom{n}{\frac{1}{2}(n+b-a)}$. \square

Theorem 6 (Reflection principle)

Let $a, b > 0$, then $N_n^0(a, b) = N_n(-a, b)$.

Proof. Each path from $(0, -a)$ to (n, b) has to intersect x -axis at least once at some point. Let k be the time of earliest intersection with x -axis. By reflexing the segment from $(0, -a)$ to $(k, 0)$ in the x -axis, we get a path from point $(0, a)$ to (n, b) which visits 0 at point k . Because reflection is bijective operation on sets of paths, we get the correspondence between the collections of such paths. \square

Definition 4 (Return to origin)

Let $(\{S_n\}_{n=0}^{+\infty}, p)$ be a random walk. Then if $\exists k \in \mathbb{N}$ such that $S_k = 0$ then we say that in k -th step occurred return to origin. Let $(\{S_n\}_{n=0}^{+\infty}, p)$ be a random walk. Then if $S_1, S_2, \dots, S_{2n-1} \neq 0$ and $S_{2n} = 0$

Theorem 7 (Ballot theorem)

Let $n, b \in \mathbb{N}$ Number of paths from point $(0, 0)$ to point (n, b) which do not return to origin is equal to $\frac{b}{n} N_n(0, b)$

Proof. Let us call N the number of paths we are referring to. Because the path ends at point (n, b) , the first step has to be rightwards. Therefore we now have $N = N_{n-1}(1, b) - N_{n-1}^0(1, b) = N_{n-1}(1, b) - N_{n-1}(-1, b)$. The last equation was acquired using Reflection principle (6). We now have: $N_{n-1}(1, b) - N_{n-1}(-1, b) =$

$$\begin{aligned} & \binom{n-1}{\frac{n+b}{2}-1} - \binom{n-1}{\frac{n-b}{2}} = \frac{(n-1)!}{(\frac{n+b}{2}-1)!(\frac{n-b}{2})!} - \frac{(n-1)!}{(\frac{n+b}{2})!(\frac{n-b}{2}-1)!} = \frac{(n-1)!}{(\frac{n+b}{2}-1)!(\frac{n-b}{2})(\frac{n-b}{2}-1)!} - \\ & \frac{(n-1)!}{(\frac{n+b}{2})(\frac{n+b}{2}-1)!(\frac{n-b}{2}-1)!} = \frac{(n-1)!}{(\frac{n+b}{2}-1)!(\frac{n-b}{2}-1)!} \left(\frac{1}{\frac{n-b}{2}} - \frac{1}{\frac{n+b}{2}} \right) = \frac{1}{n} \frac{n!}{(\frac{n+b}{2}-1)!(\frac{n-b}{2}-1)!} \left(\frac{(\frac{n+b}{2}-\frac{n-b}{2})}{(\frac{n-b}{2})(\frac{n+b}{2})} \right) \\ & \frac{b}{n} \frac{n!}{(\frac{n+b}{2})!(\frac{n-b}{2})!} = \frac{b}{n} \binom{n}{\frac{n+b}{2}} = \frac{b}{n} N_n(0, b) \end{aligned} \quad \square$$

Definition 5

$$M_n = \max\{S_i, i \in \{0, 1, \dots, n\}\}$$

Theorem 8 (Probability of maximum up to time n)

Let $(\{S_n\}_{n=0}^{+\infty}, p)$ be a random walk .

$$P(M_n \geq r, S_n = b) = \begin{cases} P(S_n = b) & \text{for } b \geq r, \\ P(S_n = 2r - b) \left(\frac{q}{p}\right)^{r-b}, & \text{for otherwise.} \end{cases}$$

Proof. Let us firstly consider the easier case in which $b \geq r$. Because we defined M_n as $\max\{S_i, i \in \{0, 1, \dots, n\}\}$ we get that $M_n \geq b \geq r$ therefore $[M_n \geq r] \subset [S_n = b]$ therefore we get $P(M_n \geq r, S_n = b) = P(S_n = b)$. Let $r \geq 1, b < r$. $N_n^r(0, b)$ stands for number of paths from point $(0, 0)$ to point (n, b) which reach up to r . Let $k \in \{0, 1, \dots, n\}$ denote the first time we reach r . By reflection principle (6), we can reflex the segment from (k, r) to (n, b) in the axis: $y = r$. Therefore we now have path from $(0, 0)$ to $(n, 2r - b)$ and we get that $N_n^r(0, b) = N_n(0, 2r - b)$. $P(S_n = b, M_n \geq r) = N_n^r(0, b) p^{\frac{n+b}{2}} q^{\frac{n-b}{2}} = N_n(0, 2r - b) p^{\frac{n+(2r-b)}{2}} q^{\frac{n-(2r-b)}{2}} p^{\frac{n+b}{2}} q^{\frac{n-b}{2}} p^{b-r} q^{r-b} = \left(\frac{q}{p}\right)^{r-b} P(S_n = 2r - b)$. \square

Definition 6 (Walk reaching new maximum at particular time)

Let $b > 0$. $f_b(n)$ denotes the probability that we reach new maximum b in time n . $f_b(n) = P(M_{n-1} = S_{n-1} = b - 1, S_n = b)$

Theorem 9 (Probability of reaching new maximum b in time n)

Let $b > 0$ then $f_b(n) = \frac{b}{n} P(S_n = b)$.

Proof. $f_b = P(M_{n-1} = S_{n-1} = b - 1, S_n = b) = p P(M_{n-1} = S_{n-1} = b - 1) = p (P(M_{n-1} \geq b - 1, S_{n-1} = b - 1) - P(M_{n-1} \geq b, S_{n-1} = b - 1)) = p (P(S_{n-1} = b - 1) - \frac{q}{p} P(S_{n-1} = b + 1)) = p P(S_{n-1} = b - 1) - q P(S_{n-1} = b + 1) = \binom{n-1}{\frac{n}{2} + \frac{b}{2} - 1} p^{\frac{n}{2} + \frac{b}{2}} q^{\frac{n}{2} - \frac{b}{2}} - \binom{n-1}{\frac{n}{2} + \frac{b}{2}} p^{\frac{n}{2} + \frac{b}{2}} q^{\frac{n}{2} - \frac{b}{2}} = p^{\frac{n}{2} + \frac{b}{2}} q^{\frac{n}{2} - \frac{b}{2}} \left(\frac{(n-1)!}{(\frac{n}{2} + \frac{b}{2} - 1)! (\frac{n}{2} - \frac{b}{2})!} - \frac{(n-1)!}{(\frac{n}{2} + \frac{b}{2})! (\frac{n}{2} - \frac{b}{2} - 1)!} \right) = p^{\frac{n}{2} + \frac{b}{2}} q^{\frac{n}{2} - \frac{b}{2}} \left(\frac{(n-1)!}{(\frac{n}{2} + \frac{b}{2})! (\frac{n}{2} - \frac{b}{2})!} \right) \left(\frac{1}{\frac{n}{2} - \frac{b}{2}} - \frac{1}{\frac{n}{2} + \frac{b}{2}} \right) = p^{\frac{n}{2} + \frac{b}{2}} q^{\frac{n}{2} - \frac{b}{2}} \frac{b}{n} \left(\frac{n!}{(\frac{n}{2} + \frac{b}{2})! (\frac{n}{2} - \frac{b}{2})!} \right) = \frac{b}{n} p^{\frac{n}{2} + \frac{b}{2}} q^{\frac{n}{2} - \frac{b}{2}} \binom{n}{\frac{n}{2} + \frac{b}{2}} = \frac{b}{n} P(S_n = b)$. \square

Theorem 10 (XXXMean number of visits to b before returning to origin in symmetric random walk)

Let $(\{S_n\}_{n=0}^{+\infty}, p)$ be a symmetric random walk . Mean number μ_b of visits of the walk to point b before returning to origin is equal to 1.

Proof. aa \square

Definition 7 (Return to origin)

Let $(\{S_n\}_{n=0}^{+\infty}, p)$ be a random walk . Let $k \in \mathbb{N}$. We say a return to origin occurred in time $2k$ if $S_{2k} = 0$. The probability that in time $2k$ occurred a return to origin shall be denoted by u_{2k} . We say that in time $2k$ occurred first return to origin if $S_1, S_2, \dots, S_{2k-1} \neq 0$ and $S_{2k} = 0$. The probability that in time $2k$ occurred first return to origin shall be denoted by f_{2k} . By definition $f_0 = 0$. Let $\alpha_{2n}(2k)$ denote $u_{2k} u_{2(n-k)}$

Lemma 11 (Binomial identity)

Let $n, k \in \mathbb{N}, n > k : \binom{n-1}{k} - \binom{n-1}{k-1} = \frac{n-2k}{n} \binom{n}{k}$

$$\begin{aligned} \text{Proof. } \binom{n-1}{k} - \binom{n-1}{k-1} &= \frac{(n-1)!}{k!(n-k-1)!} - \frac{(n-1)!}{(k-1)!(n-k)!} = \frac{(n-1)!}{(k-1)!(n-k-1)!} \left(\frac{1}{k} - \frac{1}{n-k} \right) \\ &= \frac{1}{n} \frac{n!}{(k-1)!(n-k-1)!} \frac{n-2k}{k(n-k)} = \frac{n-2k}{n} \frac{n!}{k!(n-k)!} = \frac{n-2k}{n} \binom{n}{k} \quad \square \end{aligned}$$

Lemma 12 (Main lemma)

Let $(\{S_n\}_{n=0}^{+\infty}, p)$ be a symmetrical random walk. Then $P(S_1 \cdot S_2, \cdot, \dots, S_{2n} \neq 0) = P(S_{2n} = 0)$.

$$\begin{aligned} \text{Proof. } P(S_1 \cdot S_2, \cdot, \dots, S_{2n} \neq 0) &= \sum_{i=-\infty}^{+\infty} P(S_1 \cdot S_2, \cdot, \dots, S_{2n-1} \neq 0, S_{2n} = 2i) = \\ &= \sum_{i=-n}^n P(S_1 \cdot S_2, \cdot, \dots, S_{2n-1} \neq 0, S_{2n} = 2i) = 2 \cdot \sum_{i=1}^n P(S_1 \cdot S_2, \cdot, \dots, S_{2n-1} \neq 0, S_{2n} = 2i) = \\ &= 2 \sum_{i=1}^n \frac{2i}{2n} P(S_{2n} = 2k) = 2 \sum_{i=1}^n \frac{2i}{2n} \binom{2n}{n+i} 2^{-2n} = 2 \cdot 2^{-2n} \sum_{i=1}^n \left(\binom{2n-1}{m+k-1} - \binom{2n-1}{m+k} \right) = 2 \cdot \\ &= 2^{-2n} \binom{2n-1}{n} = 2^{-2n} \frac{2n}{n} \binom{2n-1}{n} = 2^{-2n} \frac{2n(2n-1)!}{m(m-1)!m!} = 2^{-2n} \frac{(2n)!}{m!m!} = 2^{-2n} \binom{2n}{n} = P(S_{2n} = 0) \quad \square \end{aligned}$$

Theorem 13

Let $(\{S_n\}_{n=0}^{+\infty}, p)$ be a symmetric random walk. The probability that the last return to origin up to time $2n$ occurred in time $2k$ is $P(S_{2k} = 0) P(S_{2(n-k)} = 0)$.

$$\begin{aligned} \text{Proof. } \alpha_{2n}(2k) &= u_{2k} u_{2(n-k)} = P(S_{2k} = 0) P(S_{2k+1} \cdot S_{2k+2}, \dots, S_{2n} \neq 0 \mid S_{2k} = 0) = \\ &= P(S_{2k} = 0) P(S_1 \cdot S_2, \dots, S_{2(n-k)} \neq 0) = P(S_{2k} = 0) P(S_{2(n-k)} = 0) \quad \square \end{aligned}$$

Theorem 14

Let $b \in \mathbb{Z}$. $P(S_1 \cdot S_2, \dots, S_n \neq 0, S_n = b) = \frac{|b|}{n} P(S_n = b)$.

Proof. Let us Without loss of generality assume that $b > 0$. In that case, first step has to be rightwards ($X_1 = +1$). Now we have path from point $(1, 1)$ to point (n, b) that does not return to origin. By Ballot theorem 7 there are $\frac{b}{n} N_n(0, b)$ such paths. Each path consists of $\frac{n+b}{2}$ rightwards steps and $\frac{n-b}{2}$ leftwards steps. Therefore $P(S_1 \cdot S_2, \dots, S_n \neq 0, S_n = b) = \frac{b}{n} N_n(0, b) p^{\frac{n+b}{2}} q^{\frac{n-b}{2}} = \frac{b}{n} P(S_n = b)$. Case $b < 0$ is identical. \square

Lemma 15

Let $(\{S_n\}_{n=0}^{+\infty}, p)$ be a symmetric random walk. $P(S_1, S_2, \dots, S_{2n} > 0) = \frac{1}{2} P(S_{2n} = 0) = \frac{1}{2} u_{2n}$.

Proof. Because $S_i > 0 \forall i \in \mathbb{N}$ the first step has to be rightwards ($X_1 = S_1 = 1$). Therefore we get $P(S_1, S_2, \dots, S_{2n} > 0) = \sum_{r=1}^{n-1} P(S_1, S_2, \dots, S_{2n} > 0, S_{2n} = 2r)$.

The r -th term follows equation: $P(S_1, S_2, \dots, S_{2n} > 0, S_{2n} = 2r)$

$$\begin{aligned} &= P(X_1 = 1, S_2, \dots, S_{2n} > 0, S_{2n} = 2r) \\ &= \frac{1}{2} P(S_2, S_3, \dots, S_{2n} > 0, S_{2n} = 2r) \\ &= \frac{1}{2} (P(S_{2n} = 2r) - P(S_2 \cdot S_3 \cdot \dots \cdot S_{2n} = 0, S_{2n} = 2r)) \\ &= \frac{1}{2} \left(\frac{1}{2} {}^{2n-1}N_{2n-1}(1, 2r) - \frac{1}{2} {}^{2n-1}N_{2n-1}^0(1, 2r) \right) \\ &= \frac{1}{2} \frac{1}{2} {}^{2n-1}N_{2n-1}(1, 2r) - \frac{1}{2} {}^{2n-1}N_{2n-1}^0(1, 2r) \\ &= \frac{1}{2} \frac{1}{2} {}^{2n-1}N_{2n-1}(1, 2r) - N_{2n-1}(-1, 2r) \end{aligned}$$

$= \frac{1}{2} \frac{1}{2}^{2n-1} \left(\binom{2n-1}{n+r-1} - \binom{2n-1}{n+r} \right)$. Because of the fact that the negative parts of r -th terms cancel against the positive parts of $(r+1)$ -st terms and the sum reduces to just $\frac{1}{2} \frac{1}{2}^{2n-1} \binom{2n-1}{n} = \frac{1}{2} \cdot 2 \cdot \frac{1}{2}^{2n} \binom{2n-1}{n} = \frac{1}{2} \frac{1}{2}^{2n} \frac{2(2n-1)!}{n!(n-1)!} = \frac{1}{2} \frac{1}{2}^{2n} \frac{(2n)!}{n!n!} = \frac{1}{2} \frac{1}{2}^{2n} \binom{2n}{n} = \frac{1}{2} \mathbb{P}(S_{2n} = 0) = \frac{1}{2} u_{2n}$. \square

Theorem 16 (No return=return)

$$\mathbb{P}(S_1, S_2, \dots, S_{2n} \neq 0) = \mathbb{P}(S_{2n} = 0) = u_{2n}$$

Proof. The event $[S_1, S_2, \dots, S_{2n} \neq 0]$ can be split into two disjoint events: $= [S_1, S_2, \dots, S_{2n} < 0] \cup [S_1, S_2, \dots, S_{2n} > 0]$. By previous theorem (15) we get that probability of both of them is $\frac{1}{2} u_{2n}$. Because the events are disjoint we can sum their probabilities and we get the result. \square

Corollary. $\mathbb{P}(S_1, S_2, \dots, S_{2n} \geq 0) = \mathbb{P}(S_{2n} = 0) = u_{2n}$

Proof. $\frac{1}{2} u_{2n} = \mathbb{P}(S_1, S_2, \dots, S_{2n} > 0) = \mathbb{P}(X_1 = 1, S_2, S_3, \dots, S_{2n} \geq 1)$
 $= \frac{1}{2} \mathbb{P}(S_2, S_3, \dots, S_{2n} \geq 1 \mid S_1 = 1)$
 $= \frac{1}{2} \mathbb{P}(S_1, S_2, \dots, S_{2n-1} \geq 1 \mid S_0 = 1)$
 $= \frac{1}{2} \mathbb{P}(S_1, S_2, \dots, S_{2n-1} \geq 0)$
 $= \frac{1}{2} \mathbb{P}(S_1, S_2, \dots, S_{2n} \geq 0)$. Therefore $\mathbb{P}(S_1, S_2, \dots, S_{2n} \geq 0) = u_{2n}$. \square

Theorem 17

$$f_{2n} = u_{2n-2} - u_{2n}$$

Proof. The event $[S_1, S_2, \dots, S_{2n-1} \neq 0]$ can be split into two disjoint events: $[S_1, S_2, \dots, S_{2n-1} \neq 0, S_{2n} = 0]$ and $[S_1, S_2, \dots, S_{2n-1} \neq 0, S_{2n} \neq 0]$. Therefore $\mathbb{P}(S_1, S_2, \dots, S_{2n-1} \neq 0) = \mathbb{P}(S_1, S_2, \dots, S_{2n-1} \neq 0, S_{2n} = 0) + \mathbb{P}(S_1, S_2, \dots, S_{2n-1} \neq 0, S_{2n} \neq 0)$. Therefore we get $f_{2n} = \mathbb{P}(S_1, S_2, \dots, S_{2n-1} \neq 0, S_{2n} = 0) = \mathbb{P}(S_1, S_2, \dots, S_{2n-1} \neq 0) - \mathbb{P}(S_1, S_2, \dots, S_{2n} \neq 0)$. Because $2n-1$ is odd. $\mathbb{P}(S_{2n-1} = 0) = 0$. Therefore the first term is equal to $\mathbb{P}(S_1, S_2, \dots, S_{2n-2} \neq 0)$ which is by 16 equal to u_{2n-2} . Second term is by 16 equal to u_{2n} . Therefore we get the result. \square

Corollary. $f_{2n} = \frac{1}{2n-1} u_{2n}$

Proof. $u_{2n-2} = \frac{1}{2}^{2n-2} \binom{2n-2}{n-1} = 4 \cdot \frac{1}{2}^{2n} \frac{(2n-2)!}{(n-1)!(n-1)!} = \frac{4n^2}{(2n)(2n-1)} \frac{1}{2}^{2n} \binom{2n}{n} = \frac{2n}{2n-1} u_{2n}$. Therefore $u_{2n-2} - u_{2n} = u_{2n} \left(\frac{2n}{2n-1} - 1 \right) = u_{2n} \frac{1}{2n-1}$. \square

Theorem 18 (Arcsine law for last visits)

Let $k, n \in \mathbb{N}, k \leq n$. The probability that up to time $2n$ a return to origin occurred in time $2k$ is given by $\alpha_{2n}(2k) = u_{2n} u_{2(n-k)}$.

Proof. The probability involved can be rewritten as:

$$\begin{aligned} & \mathbb{P}(S_{2k+1}, S_{2k+2}, \dots, S_{2n} \neq 0, S_{2k} = 0) \\ &= * * \mathbb{P}(S_{2k+1}, S_{2k+2}, \dots, S_{2n} \neq 0 \mid S_{2k} = 0) \mathbb{P}(S_{2k} = 0) \\ &= * \mathbb{P}(S_1, S_2, \dots, S_{2(n-k)} \neq 0) \mathbb{P}(S_{2k} = 0) \\ &= * * u_{2(n-k)} u_{2k} \end{aligned}$$

\square

Definition 8 (Time spend on the positive and negative sides)

Let $(\{S_n\}_{n=0}^{+\infty}, p)$ be a random walk. We say that the walk spent τ time units of n on the positive side if $\sum_{i=1}^n \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = \tau$. Let $\beta_n(\tau)$ denote the probability of such an event. We say that the walk spent ζ time units of n on the negative side if $\sum_{i=1}^n \mathbf{1}_{[S_i < 0 \vee S_{i-1} < 0]} = \zeta$.

Theorem 19 (Arcsine law for sojourn times-OWN PROOF)

Let $(\{S_n\}_{n=0}^{+\infty}, p)$ be a symmetric random walk . Then $\beta_{2n}(2k) = \alpha_{2n}(2k)$.

Proof. Firstly let us start with degenerate cases. $*f\beta_{2n}(2n)$

$$= P\left(\sum_{i=1}^{2n} \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2n\right) = P(S_1, S_2, \dots, S_{2n} \geq 0) = *u_{2n}.$$

By symmetry $\beta_{2n}(0) = \beta_{2n}(2n) = u_{2n}$. Let $1 \leq k \leq v-1$, where $0 \leq v \leq n$.

For such k stands equation: $\beta_{2n}(2k) = P\left(\sum_{i=1}^n \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k\right)$

$$= *a \sum_{r=1}^n P\left(\sum_{i=1}^n \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k, S_1, S_2, \dots, S_{2r-1} \neq 0, S_{2r} = 0\right)$$

$$= *b \sum_{r=1}^n P\left(\sum_{i=1}^n \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k, S_1, S_2, \dots, S_{2r-1} < 0, S_{2r} = 0\right)$$

$$+ \sum_{r=1}^n P\left(\sum_{i=1}^n \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k, S_1, S_2, \dots, S_{2r-1} > 0, S_{2r} = 0\right)$$

$$= *c \sum_{r=1}^n P\left(\sum_{i=1}^n \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k \mid S_1, S_2, \dots, S_{2r-1} < 0, S_{2r} = 0\right)$$

$$P(S_1, S_2, \dots, S_{2r-1} < 0, S_{2r} = 0)$$

$$+ \sum_{r=1}^n P\left(\sum_{i=1}^n \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k \mid S_1, S_2, \dots, S_{2r-1} > 0, S_{2r} = 0\right)$$

$$P(S_1, S_2, \dots, S_{2r-1} > 0, S_{2r} = 0)$$

$$= *d \sum_{r=1}^n \frac{1}{2} f_{2r} P\left(\sum_{i=2r+1}^{2n} \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k \mid S_{2r} = 0\right)$$

$$+ \sum_{r=1}^n \frac{1}{2} f_{2r} P\left(\sum_{i=2r+1}^{2n} \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k - 2r \mid S_{2r} = 0\right)$$

$$= *e \sum_{r=1}^n \frac{1}{2} f_{2r} P\left(\sum_{i=1}^{2n-2r} \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k\right) + \sum_{r=1}^n \frac{1}{2} f_{2r} P\left(\sum_{i=1}^{2n-2r} \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k - 2r\right) =$$

$$\sum_{r=1}^n \frac{1}{2} f_{2r} \beta_{2n-2r}(2k) + \sum_{r=1}^n \frac{1}{2} f_{2r} \beta_{2n-2r}(2k - 2r) \text{ Now let us proceed by induction.}$$

Case for $v = 1$ is trivial because it implies degenerous case from $*f$. Let the

statment be true for $v \leq n-1$, then $\sum_{r=1}^n \frac{1}{2} f_{2r} \beta_{2n-2r}(2k) + \sum_{r=1}^n \frac{1}{2} f_{2r} \beta_{2n-2r}(2k - 2r)$

$$= *g \sum_{r=1}^n \frac{1}{2} f_{2r} \alpha_{2n-2r}(2k) + \sum_{r=1}^n \frac{1}{2} f_{2r} \alpha_{2n-2r}(2k - 2r)$$

$$= *h \sum_{r=1}^n \frac{1}{2} f_{2r} u_{2k} u_{2n-2r-2k} + \sum_{r=1}^n \frac{1}{2} f_{2r} u_{2k-2r} u_{2n-2k}$$

$$= *i \frac{1}{2} u_{2k} \sum_{r=1}^n f_{2r} u_{2n-2r-2k} + \frac{1}{2} u_{2n-2k} \sum_{r=1}^n f_{2r} u_{2k-2r} u_{2n-2k}$$

$$= *j \frac{1}{2} u_{2n-2k} u_{2k} + \frac{1}{2} u_{2n-2k} u_{2k} = u_{2n-2k} u_{2k}$$

$$= *h \alpha_{2n}(2k)$$

□

Definition 9 (Change of a sign)

Let $(\{S_n\}_{n=0}^{+\infty}, p)$ be a random walk . We say that in time n occurred a change of sign if $S_{n-1} \cdot S_{n+1} = -1$ in other words if $(S_{n-1} = +1 \wedge S_{n+1} = -1) \vee (S_{n-1} = -1 \wedge S_{n+1} = +1)$. We shall denote the probability that up to time n occurred r changes of sign by $\xi_{r,n}$.

Theorem 20 (Change of a sign)

Let $(\{S_n\}_{n=0}^{+\infty}, p)$ be a symmetric random walk . The probability $\xi_{r,2n+1} = 2 P(S_{2n+1} = 2r + 1)$

Proof. Feller

□

1.1 Problem chapter 9 Feller

Definition 10 (δ, ε)

Let $(\{S_n\}_{n=0}^{+\infty}, p)$ be a symmetric random walk. $\delta_n(k)$ shall denote $P\left(\sum_{i=1}^n \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = k, S_n = 0\right)$,
 $\varepsilon_n^r(k)$ shall denote $P\left(\sum_{i=1}^n \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = k, S_1, S_2, \dots, S_{r-1} \neq 0, S_r = 0, S_n = 0\right)$,
 $\varepsilon_n^{r,+}(k)$ shall denote $P\left(\sum_{i=1}^n \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = k, S_1, S_2, \dots, S_{r-1} > 0, S_r = 0, S_n = 0\right)$,
 $\varepsilon_n^{r,-}(k)$ shall denote $P\left(\sum_{i=1}^n \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = k, S_1, S_2, \dots, S_{r-1} < 0, S_r = 0, S_n = 0\right)$.

Lemma 21 (Factorization of $\delta_{2n}(2k)$)

$$\delta_{2n}(2k) = \frac{1}{2} \sum_{r=1}^n (f_{2r} \delta_{2n-2r}(2k-2r) + f_{2r} \delta_{2n-2r}(2r)).$$

Proof. Because $S_{2n} = 0$ a return to origin must have happened. Let $2r$ the time of first return to origin, where $r \in \{1, 2, \dots, n\}$. By the law of total probability:

$$\begin{aligned} \delta_{2n}(2k) &= P\left(\sum_{i=1}^{2n} \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = k, S_{2n} = 0\right) \\ &= \sum_{r=1}^n P\left(\sum_{i=1}^{2n} \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k, S_1, S_2, \dots, S_{2r-1} \neq 0, S_{2r} = 0, S_{2n} = 0\right) \text{ which can} \\ &\text{be again by the law of total probability factorized as:} \\ &= \sum_{r=1}^n P\left(\sum_{i=1}^{2n} \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k, S_1, S_2, \dots, S_{2r-1} \neq 0, S_{2r} = 0, S_{2n} = 0\right) \\ &= \sum_{r=1}^n P\left(\sum_{i=1}^{2n} \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k, S_1, S_2, \dots, S_{2r-1} > 0, S_{2r} = 0, S_{2n} = 0\right) \\ &+ \sum_{r=1}^n P\left(\sum_{i=1}^{2n} \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k, S_1, S_2, \dots, S_{2r-1} < 0, S_{2r} = 0, S_{2n} = 0\right) \\ &= \sum_{r=1}^n \varepsilon_{2n}^{2r,+}(2k) + \sum_{r=1}^n \varepsilon_{2n}^{2r,-}(2k). \end{aligned}$$

Now let us calculate $\varepsilon_{2n}^{2r,+}(2k)$

$$\begin{aligned} &= P\left(\sum_{i=1}^{2n} \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k, S_1, S_2, \dots, S_{2r-1} > 0, S_{2r} = 0, S_{2n} = 0\right) \\ &= *a P\left(\sum_{i=1}^{2n} \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k, S_{2n} = 0 \mid S_1, S_2, \dots, S_{2r-1} > 0, S_{2r} = 0\right) \\ &P(S_1, S_2, \dots, S_{2r-1} > 0, S_{2r} = 0) \\ &= P\left(\sum_{i=2r+1}^{2n} \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k-2r, S_{2n} = 0 \mid S_{2r} = 0\right) \frac{1}{2} f_{2r} \\ &= *b P\left(\sum_{i=1}^{2n-2r} \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k-2r, S_{2n-2r} = 0\right) \frac{1}{2} f_{2r} \\ &= *c \delta_{2n-2r}(2k-2r) \frac{1}{2} f_{2r}. \text{ Similarly } \varepsilon_{2n}^{2r,-}(2k) \\ &= P\left(\sum_{i=1}^{2n} \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k, S_1, S_2, \dots, S_{2r-1} < 0, S_{2r} = 0, S_{2n} = 0\right) \\ &= *a P\left(\sum_{i=1}^{2n} \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k, S_{2n} = 0 \mid S_1, S_2, \dots, S_{2r-1} < 0, S_{2r} = 0\right) \\ &P(S_1, S_2, \dots, S_{2r-1} < 0, S_{2r} = 0) \\ &= P\left(\sum_{i=2r+1}^{2n} \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k, S_{2n} = 0 \mid S_{2r} = 0\right) \frac{1}{2} f_{2r} \\ &= *b P\left(\sum_{i=1}^{2n-2r} \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k, S_{2n-2r} = 0\right) \frac{1}{2} f_{2r} = *c \delta_{2n-2r}(2k) \frac{1}{2} f_{2r}. \end{aligned}$$

$$\begin{aligned} \text{Therefore } \delta_{2n}(2k) &= \frac{1}{2} \sum_{r=1}^n f_{2r} \delta_{2n-2r}(2k-2r) + \frac{1}{2} \sum_{r=1}^n f_{2r} \delta_{2n-2r}(2k) \\ &= \frac{1}{2} \sum_{r=1}^n (f_{2r} \delta_{2n-2r}(2k-2r) + f_{2r} \delta_{2n-2r}(2r)) \end{aligned} \quad \square$$

Theorem 22 (Equidistributional theorem-ALMOST COMPLETE OWN PROOF) ||
Let $(\{S_n\}_{n=0}^{+\infty}, p)$ be a symmetric random walk and $n \in \mathbb{N}$, then $\forall k, l \in \{0, 1, \dots, n\}$:
 $\delta_{2n}(2k) = \delta_{2n}(2l) = \frac{u_{2n}}{n+1}$.

Proof. Let us prove this statement by induction in n . In case that $n = 1$ we have two options for k . Either $k = 0$ or $k = 1$. $\delta_2(0) = \mathbf{P}(S_1 < 0, S_2 = 0) = \frac{1}{2}f_2 = *a\frac{1}{2}u_{2\frac{1}{2-1}} = \frac{u_2}{2}\delta_2(2) = \mathbf{P}(S_1 > 0, S_2 = 0) = \frac{1}{2}f_2 = \frac{u_2}{2}$.

Let the statment be true for all $l \leq n-1$. In that case $\delta_{2(n-l)}(2k) = \frac{u_{2(n-l)}}{n-l+1}$. We want to show that $\delta_{2n} = \frac{u_{2n}}{n+1}$.

$$\begin{aligned} \text{Let us calculate } \delta_{2n}. \quad \delta_{2n} &= *b\frac{1}{2} \sum_{r=1}^n (f_{2r} \delta_{2n-2r}(2k-2r) + f_{2r} \delta_{2n-2r}(2r)) = \\ \frac{1}{2} \sum_{r=1}^n \left(f_{2r} u_{2n-2r} \frac{1}{n-r+1} + f_{2r} u_{2n-2r} \frac{1}{n-r+1} \right) &= \sum_{r=1}^n \left(\frac{f_{2r} u_{2n-2r}}{n-r+1} \right) = ? \frac{u_{2n}}{n+1} \text{ Because???} \end{aligned} \quad \square$$

Lemma 23

$$\sum_{r=1}^n \left(\frac{1}{n-r+1} \left(2^{-(2r-2)} \binom{2r-2}{r-1} - 2^{-2r} \binom{2r}{r} \right) 2^{-(2n-2r)} \binom{2n-2r}{n-r} \right) = \frac{1}{n+1} 2^{-2n} \binom{2n}{n}$$

Proof. □