

The present paper is devoted to the study of the asymptotic behavior of the "tails" of the distribution of the random series

$$\psi = \eta_1 + \eta_2 \xi_1 + \eta_3 \xi_1 \xi_2 + \dots + \eta_n \xi_1 \dots \xi_{n-1} + \dots,$$

arising in a limit theorem for the product of independent random elements of the group of linear transformations of the line (see U. Grenander [1] and A. Grincevičius [6]), in schemes of random walks on the simplest geometric objects, i.e., straight lines (see V. Maksimov [5]), and also in the theory of semi-Markov processes of birth and death (see G. Lev [2, 3]).

In what follows we shall assume that the random matrices

$$\begin{pmatrix} \xi_j & \eta_j \\ 0 & 1 \end{pmatrix}, \quad j=1, 2, 3, \dots$$

are independent and identically distributed, and that $\xi_j > 0$ with probability 1. In this paper we shall prove

THEOREM 1. Let $M\xi_1^q < 1$, $M\xi_1^q < \infty$ and $P\{\eta_1 > x\} \sim x^{-\alpha} L(x)$ as $x \rightarrow \infty$ where $L(x)$ is a slowly varying function and $0 < \alpha < \beta$. Then

$$P\{\psi > x\} \sim x^{-\alpha} L(x) \sum_{n=0}^{\infty} (M\xi_1^q)^n.$$

THEOREM 2. Let $M\xi_1^q = 1$, $P\{|\eta_1| > x\} = o(x^{-\alpha})$ as $x \rightarrow \infty$, where $0 < \beta < \alpha$, $M\xi_1^q \ln \xi_1 < \infty$ and random variable ξ_1 is not degenerate. Then,

a) if random variable $\ln \xi_1$ has a nonarithmetic distribution, then there exists limit

$$\lim_{x \rightarrow \infty} \frac{P\{\psi > x\}}{x^{-\beta}} = C_1,$$

where constant C_1 is known beforehand to be positive if, for some real number c , $P\{\eta_1 \geq c(1 - \xi_1)\} = 1$ and $P\{\eta_1 > c(1 - \xi_1)\} > 0$;

b) if random variable $\ln \xi_1$ has an arithmetic distribution with step λ then, with the possible exception of a no more than countable number of real values of h , there exists a limit $\lim_{x \rightarrow \infty} \frac{P\{\psi > e^{h+\lambda n}\}}{e^{-(h+\lambda n)\beta}} = C_2(h)$, where $C_2(h)$ is a periodic function with period λ ; if, moreover, $P\{\eta_1 \geq 0\} = 1$ and $P\{\eta_1 > 0\} > 0$, then the limit exists for all real h , and function $C_2(h)$ is strictly positive.

In the proofs of Theorems 1 and 2 we shall use the ideas in the proofs of the analogous results of G. Lev [3] for the probability of degeneracy of semi-Markov processes of birth, but, in the paper of G. Lev, on the one hand, there is assumed, and essentially utilized, the independence of the sequences (in our notation) η_j , $j = 1, 2, 3, \dots$, and ξ_k , $k = 1, 2, 3$, while, on the other hand, in the paper of G. Lev [3, p. 785], there is an erroneous application of the renewal theorem to the entire line and, in connection with this,

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there is no separate consideration of the cases of arithmetic and nonarithmetic distributions of the random variable $\ln \xi_1$.

In the proof of Theorem 1 we shall need two lemmas.

LEMMA 1. Let $M \xi_1^\beta < \infty$, $P\{\eta_1 > x\} \sim x^{-\alpha} L(x)$, where $L(x)$ is a slowly varying function, as $x \rightarrow \infty$, $0 < \alpha < \beta$, and $\lim_{x \rightarrow \infty} \frac{P\{\psi_1 > x\}}{x^{-\alpha} L(x)} = C$, where random variable ψ_1 does not depend on random matrix $\begin{pmatrix} \xi_1 & \eta_1 \\ 0 & 1 \end{pmatrix}$. Then,

$$P\{\eta_1 + \xi_1 \psi_1 > x\} \sim (1 + C \cdot M \xi_1^\alpha) x^{-\alpha} L(x) \quad \text{as } x \rightarrow \infty.$$

Proof. We shall show that

$$\lim_{x \rightarrow \infty} \frac{P\{\xi_1 \psi_1 > x\}}{x^{-\alpha} L(x)} = C \cdot M \xi_1^\alpha. \quad (1)$$

Indeed, by virtue of the independence of random variables ξ_1 and ψ_1 we have

$$\lim_{x \rightarrow \infty} \frac{P\{\xi_1 \psi_1 > x\}}{x^{-\alpha} L(x)} = \lim_{x \rightarrow \infty} \int \frac{P\{\psi_1 y > x\}}{x^{-\alpha} L(x)} F_{\xi_1}\{dy\} = \lim_{x \rightarrow \infty} \int_0^{\frac{x}{N}} \frac{P\{\psi_1 y > x\}}{x^{-\alpha} L(x)} F_{\xi_1}\{dy\} \quad (2)$$

for any positive number N , since

$$\lim_{x \rightarrow \infty} \int_{\frac{x}{N}}^{\infty} \frac{P\{\psi_1 y > x\}}{x^{-\alpha} L(x)} F_{\xi_1}\{dy\} \leq \lim_{x \rightarrow \infty} \frac{P\left\{\xi_1 \geq \frac{x}{N}\right\}}{x^{-\alpha} L(x)} = 0$$

by virtue of the conditions of the lemma and the inequality

$$x^\beta P\{\xi_1 > x\} \leq M \xi_1^\beta < \infty.$$

The integrand on the right of Eq. (2) tends as $x \rightarrow \infty$, to Cy^α , so that it is only necessary to justify the transition to the limit under the integral sign.

We shall construct a majorant.

If $0 < y < 1$, then $\frac{P\{\psi_1 y > x\}}{x^{-\alpha} L(x)} \leq \frac{P\{\psi_1 > x\}}{x^{-\alpha} L(x)} < C_4$ for all x greater than some fixed value.

If, however, $1 \leq y < \frac{x}{N}$, we then use the representation of a slowly varying function (see W. Feller [4], Chap. 8, §9),

$$L(x) = a(x) \exp \left\{ \int_1^x \frac{\varepsilon(t)}{t} dt \right\},$$

where $a(x) \rightarrow C_3 > 0$, $\varepsilon(x) \rightarrow 0$ as $x \rightarrow \infty$. We can choose positive number N such that $x \geq N$, $\frac{C_3}{2} < a(x) < 2C_3$, while $|\varepsilon(x)| < \frac{\beta - \alpha}{2}$. Now we have

$$\begin{aligned} \frac{P\{\psi_1 y > x\}}{x^{-\alpha} L(x)} &= \frac{P\left\{\psi_1 > \frac{x}{y}\right\}}{\left(\frac{x}{y}\right)^{-\alpha} L\left(\frac{x}{y}\right)} \cdot y^\alpha \cdot \frac{L\left(\frac{x}{y}\right)}{L(x)} \leq C_4 y^\alpha \frac{L\left(\frac{x}{y}\right)}{L(x)} \leq \\ &\leq 4C_4 y^\alpha \exp \left\{ - \int_{\frac{x}{y}}^x \frac{\varepsilon(t)}{t} dt \right\} \leq 4C_4 y^\alpha \exp \left\{ \int_{\frac{x}{y}}^x \frac{dt}{t} \cdot \frac{\beta - \alpha}{2} \right\} = 4C_4 y^{\alpha - \frac{\beta - \alpha}{2}} \leq 4C_4 y^\beta. \end{aligned}$$

Thus,

$$\frac{P\{\psi_1 y > x\}}{x^{-\alpha} L(x)} < g(y),$$

where

$$g(y) = \begin{cases} C_4, & \text{if } 0 < y < 1, \\ 4C_4 y^0, & \text{if } y \geq 1 \end{cases}$$

for all x starting with some fixed one and $y < x/N$, so that, on the right of Eq. (2), we can go to the limit under the integral sign. Equation (1) is proven.

Furthermore, if $x > 1$, while $1 > \varepsilon > 0$, then, on one hand,

$$\begin{aligned} P\{\eta_1 + \xi_1 \psi_1 > x\} &\geq P\{\eta_1 > x(1+\varepsilon), |\xi_1 \psi_1| < x\varepsilon\} + \\ &+ P\{|\eta_1| < x\varepsilon, \xi_1 \psi_1 > x(1+\varepsilon)\} = P\{\eta_1 > x(1+\varepsilon) + \\ &+ P\{\xi_1 \psi_1 > x(1+\varepsilon)\} - P\{\eta_1 > x(1+\varepsilon), |\xi_1 \psi_1| \geq x\varepsilon\} - \\ &- P\{|\eta_1| > x\varepsilon, \xi_1 \psi_1 > x(1+\varepsilon)\}. \end{aligned}$$

Since

$$\frac{P\{\eta_1 > x(1+\varepsilon), |\xi_1 \psi_1| \geq x\varepsilon\}}{x^{-\alpha} L(x)} \leq \frac{P\{\xi_1 > x^{\frac{\alpha+\beta}{2\beta}}\}}{x^{-\alpha} L(x)} + \frac{P\{\eta_1 > x(1+\varepsilon)\} \cdot P\{|\psi_1| \geq x^{\frac{\beta-\alpha}{2\beta}} \varepsilon\}}{x^{-\alpha} L(x)} \rightarrow 0$$

as $x \rightarrow \infty$, and analogously

$$\frac{P\{|\eta_1| > x\varepsilon, \xi_1 \psi_1 > x(1+\varepsilon)\}}{x^{-\alpha} L(x)} \rightarrow 0,$$

then

$$\limsup_{x \rightarrow \infty} \frac{P\{\eta_1 + \xi_1 \psi_1 > x\}}{x^{-\alpha} L(x)} \geq \frac{1 + CM\xi_1^\alpha}{(1+\varepsilon)^\alpha}. \quad (3)$$

On the other hand,

$$P\{\eta_1 + \xi_1 \psi_1 > x\} \leq P\{\eta_1 > x(1-\varepsilon)\} + P\{\xi_1 \psi_1 > x(1-\varepsilon)\} + P\{\eta_1 > x\varepsilon, \xi_1 \psi_1 > x\varepsilon\},$$

whence, we obtain an inequality contrary to (3) (see the proof analogous to the one here: W. Feller [4, Chap. 8, §9]). Proof of Lemma 1 is completed.

LEMMA 2. Let $M\xi_1^\alpha < 1$, $\alpha > 0$. Then, $M(1 + \xi_1 + \dots + \xi_1 \dots \xi_n + \dots)^\alpha < \infty$.

Proof. If $\alpha \geq 1$, then, by virtue of the triangle inequality in space L_α ,

$$M(1 + \xi_1 + \dots + \xi_1 \dots \xi_n)^\alpha \leq \left[1 + (M\xi_1^\alpha)^{\frac{1}{\alpha}} + \dots + (M\xi_1^\alpha)^{\frac{n}{\alpha}} \right]^\alpha;$$

consequently the series $1 + \xi_1 + \dots + \xi_1 \dots \xi_n + \dots$ converges in metric L_α and

$$M(1 + \xi_1 + \dots + \xi_1 \dots \xi_n + \dots)^\alpha \leq \left[1 + (M\xi_1^\alpha)^{\frac{1}{\alpha}} + \dots + (M\xi_1^\alpha)^{\frac{n}{\alpha}} + \dots \right]^\alpha < \infty.$$

If $\alpha < 1$ then

$$M(1 + \xi_1 + \dots + \xi_1 \dots \xi_n + \dots)^\alpha \leq M(1 + \xi_1^\alpha + \dots + \xi_1^\alpha \dots \xi_n^\alpha + \dots) < \infty.$$

Proof of Theorem 1. By induction as $x \rightarrow \infty$

$$P\{\eta_1 + \eta_2 \xi_1 + \dots + \eta_n \xi_1 \dots \xi_{n-1} > x\} \sim x^{-\alpha} L(x) [1 + M\xi_1^\alpha + \dots + (M\xi_1^\alpha)^{n-1}]. \quad (4)$$

Since $M\xi_1^\alpha < 1$ and $M\xi_1^\beta < \infty$, we can find real numbers $\delta \in (0, \alpha)$ and $b > 1$ such that $M(b\xi_1)^{\alpha+\delta} < 1$.

Furthermore,

$$\begin{aligned} P\{\eta_1 + \eta_2 \xi_1 + \dots + \eta_n \xi_1 \dots \xi_{n-1} > x\} &\leq \\ &\leq \sum_{k=1}^{\infty} P\left\{\eta_k \xi_1 \dots \xi_{k-1} > \frac{b-1}{b^k} x, \xi_1 \dots \xi_{k-1} b^{k-1} \leq \frac{(b-1)x}{bN}\right\} + \\ &+ P\left\{\sup(1, b\xi_1, \dots, b^{k-1} \xi_1 \dots \xi_{k-1}, \dots) > \frac{x(b-1)}{Nb}\right\}. \end{aligned}$$

By Lemma 2,

$$\frac{P\left\{\sup(1, b\xi_1, \dots, b^k \xi_1 \dots \xi_k, \dots) > \frac{x(b-1)}{Nb}\right\}}{x^{-\alpha} L(x)} \leq \frac{P\left\{1 + b\xi_1 + \dots + b^k \xi_1 \dots \xi_k + \dots > \frac{x(b-1)}{Nb}\right\}}{x^{-\alpha} L(x)} = \frac{O(x^{-\alpha-\delta})}{x^{-\alpha} L(x)} \rightarrow 0$$

as $x \rightarrow \infty$.

Just the same as in the proof of Lemma 1, we can choose number $N > 0$ such that, when $x > N$ and $\frac{b-1}{b^k} \leq y \leq \frac{x(b-1)}{Nb^k}$, the following inequality is met

$$\frac{\mathbf{P}\left\{\eta_k y > \frac{b-1}{b^k} x\right\}}{x^{-\alpha} L(x)} \leq 4 C_4 \left(\frac{yb^k}{b-1}\right)^{\alpha+\delta},$$

while, if $0 < y < (b-1)/b^k$,

$$\frac{\mathbf{P}\left\{\eta_k y > \frac{b-1}{b^k} x\right\}}{x^{-\alpha} L(x)} \leq 4 C_4 \left(\frac{yb^k}{b-1}\right)^{\alpha-\delta}.$$

Therefore,

$$\begin{aligned} \limsup_{x \rightarrow \infty} \frac{\mathbf{P}\{\eta_1 + \eta_2 \xi_1 + \dots + \eta_n \xi_1 \dots \xi_{n-1} + \dots > x\}}{x^{-\alpha} L(x)} &\leq \\ &\leq \limsup_{x \rightarrow \infty} \sum_{k=1}^{\infty} \int_0^{\frac{x(b-1)}{Nb^k}} \frac{\mathbf{P}\left\{\eta_k y > \frac{b-1}{b^k} x\right\}}{x^{-\alpha} L(x)} F_{\xi_1, \dots, \xi_{k-1}}\{dy\} \leq \\ &\leq \sum_{k=1}^{\infty} 4 C_4 \left[\mathbf{M} \left(\frac{\xi_1 \dots \xi_{k-1} b^k}{b-1} \right)^{\alpha+\delta} + \mathbf{M} \left(\frac{\xi_1 \dots \xi_{k-1} b^k}{b-1} \right)^{\alpha-\delta} \right] = C_5 < \infty. \end{aligned} \quad (5)$$

Using (4) and (5) we can easily show (just as in the proof of Lemma 1) that, for any natural n

$$\begin{aligned} \limsup_{x \rightarrow \infty} \frac{\mathbf{P}\{\eta_1 + \eta_2 \xi_1 + \dots + \eta_n \xi_1 \dots \xi_{n-1} + \dots > x\}}{x^{-\alpha} L(x)} &\leq \\ &\leq [1 + \mathbf{M} \xi_1^\alpha + \dots + (\mathbf{M} \xi_1^\alpha)^{n-1} + C_5 (\mathbf{M} \xi_1^\alpha)^n], \end{aligned}$$

while

$$\liminf_{x \rightarrow \infty} \frac{\mathbf{P}\{\psi > x\}}{x^{-\alpha} L(x)} \geq [1 + \mathbf{M} \xi_1^\alpha + \dots + (\mathbf{M} \xi_1^\alpha)^{n-1}].$$

Whence also follows the assertion of Theorem 1.

Theorem 2 is of great interest since its conditions automatically hold if the bearer of random element $\begin{pmatrix} \xi_1 & \eta_1 \\ 0 & 1 \end{pmatrix}$ is lumped in some compact group if linear transformations of the line (in representation as the group of all real matrices of the form $\begin{pmatrix} X & Y \\ 0 & 1 \end{pmatrix}$ $X > 0$), $\mathbf{P}\{\xi_1 > 1\} > 0$ and (as is necessary for the convergence almost certainly of the series ψ) $\mathbf{M} \ln \xi_1 < 0$.

We mention also that when there hold, both the conditions of Theorem 1 and the conditions of Theorem 2, series ψ converges almost certainly (see A. Grincevicius [6]).

Proof of Theorem 2. We have $\psi = \eta_1 + \xi_1 \psi_1$, where random variable ψ_1 is distributed the same as ψ and does not depend on the random matrix $\begin{pmatrix} \xi_1 & \eta_1 \\ 0 & 1 \end{pmatrix}$. Therefore, for $x > 0$,

$$\mathbf{P}\{\psi > x\} = \mathbf{P}\{\xi_1 \psi_1 > x\} + \mathbf{P}\{x - \eta_1 < \xi_1 \psi_1 \leq x\} - \mathbf{P}\{x < \xi_1 \psi_1 \leq x - \eta_1\}. \quad (6)$$

From Eq. (6) by an obvious transformation we obtain that, when $x > 0$,

$$\begin{aligned} \frac{1}{x} \int_0^x t^\beta \mathbf{P}\{\psi > t\} dt &= \frac{1}{x} \int_0^x t^\beta \mathbf{P}\{\xi_1 \psi_1 > t\} dt + \\ &+ \frac{1}{x} \int_0^x t^\beta (\mathbf{P}\{t - \eta_1 < \xi_1 \psi_1 \leq t\} - \mathbf{P}\{t < \xi_1 \psi_1 \leq t - \eta_1\}) dt. \end{aligned} \quad (7)$$

We use the notation

$$Z(x) = \frac{1}{x} \int_0^x t^\beta \mathbf{P}\{\psi > t\} dt$$

and

$$z(x) = \frac{1}{x} \int_0^x t^\beta [\mathbf{P}\{t - \eta_1 < \xi_1 \psi_1 \leq x\} - \mathbf{P}\{t < \xi_1 \psi_1 \leq t - \eta_1\}] dt;$$

and then identity (7) gives

$$Z(x) = \int_0^\infty Z\left(\frac{x}{y}\right) y^\beta F_\xi\{dy\} + z(x). \quad (8)$$

It is easy to see that identity (8) is the renewal equation (see W. Feller [4]) for the multiplicative group of positive real numbers. By the change of variables $x = e^s$, we take this renewal equation to the usual form

$$Z(e^s) = \int_{-\infty}^\infty Z(e^{s-t}) F\{dt\} + z(e^s), \quad (9)$$

where the probability distribution function is

$$F(t) = \int_{-\infty}^t e^{t\beta} F_{\ln \xi_1}\{dt\}.$$

$$(\text{We note that } F(+\infty) = \int_{-\infty}^{+\infty} e^{t\beta} F_{\ln \xi_1}\{dt\} = \mathbf{M}\xi_1^\beta = 1.)$$

We now show that function $z(e^s)$ is directly Riemann-integrable (see W. Feller [4]). Let δ be a positive number less than 1 such that $\delta\alpha > \beta$; then

$$0 \leq \int_0^x t^\beta \mathbf{P}\{t - \eta_1 < \xi_1 \psi_1 \leq t\} dt \leq \int_0^x t^\beta \mathbf{P}\{\eta_1 > t^\delta\} dt + \int_0^x t^\beta \mathbf{P}\{t - t^\delta < \xi_1 \psi_1 \leq t\} dt. \quad (10)$$

We now estimate each of the integrals on the right of inequality (10):

$$\int_0^x t^\beta \mathbf{P}\{\eta_1 > t^\delta\} dt = O\left(\int_1^x t^{\beta-\delta\alpha} dt + 1\right) = o(x^{1-\varepsilon_1}) \quad \text{as } x \rightarrow \infty$$

for some positive number $\varepsilon_1 < 1$.

Since

$$\mathbf{M}|\eta_1|^\beta = \beta \int_0^\infty x^{\beta-1} \mathbf{P}\{|\eta_1| > x\} dx = O\left(1 + \int_1^\infty x^{\beta-1} x^{-\alpha} dx\right) = O(1)$$

and, for any positive number $\varepsilon_2 < \beta$, $\mathbf{M}\xi_1^{\beta-\varepsilon_2} < 1$, then, just as in Lemma 2, we readily show that $\mathbf{M}|\xi_1 \psi_1|^{\beta-\varepsilon_2} < \infty$, so that, consequently,

$$\mathbf{P}\{\xi_1 \psi_1 > x\} = o(x^{-\beta+\varepsilon_2}) \quad \text{as } x \rightarrow \infty.$$

Now, using the previous inequality, the inequality

$$\begin{aligned} t^\beta - (t - t^\delta)^\beta (1 - \delta t^{\delta-1}) &= \delta (t - t^\delta)^\beta t^{\delta-1} + t^\beta - (t - t^\delta)^\beta = \\ &= \delta (t - t^\delta)^\beta t^{\delta-1} + (y^\beta)'_{y=t-t^\delta} \cdot t^\delta = O(t^{\beta+\delta-1}), \end{aligned}$$

when $t \geq x_0$, for positive number x_0 such that, when $x \geq x_0$, $x^\delta \leq \frac{x}{2}$, and, making the change of variable $t = y - y^\delta$ in one of the following integrals, we obtain

$$\int_0^x t^\beta \mathbf{P}\{t - t^\delta < \xi_1 \psi_1 \leq t\} dt = \int_0^x t^\beta \mathbf{P}\{\xi_1 \psi_1 > t - t^\delta\} dt -$$

$$\begin{aligned}
& - \int_0^{x-x^\delta} t^\beta \mathbf{P}\{\xi_1 \psi_1 > t\} dt - \int_{x-x^\delta}^x t^\beta \mathbf{P}\{\xi_1 \psi_1 > t\} dt \leq \\
& \leq \int_0^x t^\beta \mathbf{P}\{\xi_1 \psi_1 > t-t^\delta\} dt - \int_1^x (y-y^\delta)^\beta \mathbf{P}\{\xi_1 \psi_1 > y-y^\delta\} (1-\delta y^{\delta-1}) dy = \\
& = O(1) + \int_1^x [t^\beta - (t-t^\delta)^\beta (1-\delta t^{\delta-1})] \mathbf{P}\{\xi_1 \psi_1 > t-t^\delta\} dt = \\
& = O\left(1 + \int_{x_0}^x t^{\beta+\delta-1} \mathbf{P}\{\xi_1 \psi_1 > t-t^\delta\} dt\right) = O\left(1 + \int_{x_0}^x t^{\beta+\delta-1} \cdot t^{-\beta+\varepsilon_1} dt\right) = \\
& = O(1+x^{\delta+\varepsilon_1}).
\end{aligned}$$

Thus, we have shown that, for some positive number,

$$\varepsilon_3 \int_0^x t^\beta \mathbf{P}\{t-\eta_1 < \xi_1 \psi_1 \leq t\} dt = o(x^{1-\varepsilon_3}) \text{ as } x \rightarrow \infty.$$

Similarly, we estimate the integral

$$\int_0^x t^\beta \mathbf{P}\{t < \xi_1 \psi_1 \leq t-\eta_1\} dt,$$

therefore, by virtue of the definition of function $z(x)$ we find that, as $x \rightarrow \infty$, for some positive number ε_4

$$z(x) = o(x^{-\varepsilon_4}). \quad (11)$$

On the other hand, as $x \rightarrow 0$,

$$|z(x)| \leq \frac{2 \int_0^x t^\beta dt}{x} = \frac{2x^\beta}{\beta+1}. \quad (12)$$

From inequalities (11) and (12) readily follows the direct Riemann-integrability of the function $z(e^s)$, so that, from the renewal theorem on the entire line it now follows (see W. Feller [4, Chap. 11, §9, Theorem 1], and the proof of Theorem 2 in Chap. 11, §1) that if the distribution of random variable $\ln \xi_1$ is nonarithmetic then, as $s \rightarrow \infty$,

$$Z(e^s) \rightarrow \frac{1}{\mu} \int_{-\infty}^{\infty} z(e^s) ds = C_1,$$

where

$$\mu = \int_{-\infty}^{\infty} t F\{dt\} = M \xi_1^\beta \ln \xi_1 < \infty.$$

Whence,

$$\lim_{x \rightarrow \infty} \frac{1}{x} \int_0^x t^\beta \mathbf{P}\{\psi > t\} dt = C_1$$

and, consequently,

$$\lim_{x \rightarrow \infty} \frac{\mathbf{P}\{\psi > x\}}{x^{-\beta}} = C_1$$

(see W. Feller [4, Chap. 13, §5, proof of lemma] or G. Lev [7, p. 51]).

We now assume that random variable η_1 is (almost certainly) nonnegative and, with positive probability, is greater than zero. In this case, Eq. (6) takes the form

$$\mathbf{P}\{\psi > x\} = \mathbf{P}\{\xi_1 \psi_1 > x\} + \mathbf{P}\{x - \eta_1 < \xi_1 \psi_1 \leq x\}. \quad (13)$$

It is not difficult to obtain

$$\begin{aligned} 2C_1 &= \int_{-\infty}^{\infty} \frac{1}{x^2} \left[\int_0^{e^s} t^\beta \mathbf{P} \{t - \tau_1 < \xi_1 \psi_1 \leq t\} dt \right] ds = \lim_{N \rightarrow \infty} \int_0^N \frac{1}{x^2} \left[\int_0^x t^\beta \mathbf{P} \{t - \tau_1 < \xi_1 \psi_1 \leq t\} dt \right] dx = \\ &= \lim_{N \rightarrow \infty} \int_0^N \int_0^N \frac{1}{x^2} dx t^\beta \mathbf{P} \{t - \tau_1 < \xi_1 \psi_1 \leq t\} dt = \lim_{N \rightarrow \infty} \int_0^N t^{\beta-1} \mathbf{P} \{t - \tau_1 < \xi_1 \psi_1 \leq t\} dt - \\ &- \lim_{N \rightarrow \infty} \frac{1}{N} \int_0^N t^\beta [\mathbf{P} \{\psi > t\} - \mathbf{P} \{\xi_1 \psi_1 > t\}] dt = \int_0^\infty t^{\beta-1} \mathbf{P} \{t - \tau_1 < \xi_1 \psi_1 \leq t\} dt, \end{aligned}$$

therefore equation $C_1 = 0$ means identically equal to zero for all $x > 0$

$$\mathbf{P} \{x - \tau_1 < \xi_1 \psi_1 \leq x\} \equiv 0$$

and, consequently, $\mathbf{P} \{\psi > x\} \equiv \mathbf{P} \{\xi_1 \psi_1 > x\}$ for all $x > 0$.

The last identity can hold only if $\mathbf{P} \{\psi > 0\} = 0$, or $\mathbf{P} \{\xi_1 = 1\} = 1$, and both these equations contradict the conditions of the theorem.

If $\mathbf{P} \{\tau_1 \geq c(1 - \xi_1)\} = 1$ and $\mathbf{P} \{\tau_1 > c(1 - \xi_1)\} > 0$, then, obviously, $\psi = c + [\tau_1 - c(1 - \xi_1)] + [\tau_2 - c(1 - \xi_2)]\xi_1 + [\tau_3 - c(1 - \xi_3)]\xi_1\xi_2 + \dots$ and, consequently,

$$\lim_{x \rightarrow \infty} \frac{\mathbf{P} \{\psi > x\}}{x^{-\beta}} = \lim_{x \rightarrow \infty} \frac{\mathbf{P} \{\psi - c > x\}}{x^{-\beta}} = C_1 > 0;$$

with this selection, the case of a nonarithmetic distribution of the random variable $\ln \xi_1$ is completely finished.

We now turn to the case of an arithmetic distribution of random variable $\ln \xi_1$ with step λ . The renewal equation of (9) gives, for any real number h ,

$$\lim_{n \rightarrow \infty} Z(e^{h+\lambda n}) = \frac{1}{\lambda} \sum_{k=-\infty}^{\infty} z(e^{h+\lambda k}) = C(h). \quad (14)$$

For two real numbers $h_1 \geq h_2$ we have

$$e^{h_1} Z(e^{h_1+\lambda n}) - e^{h_2} Z(e^{h_2+\lambda n}) \rightarrow e^{h_1} C(h_1) - e^{h_2} C(h_2)$$

as $n \rightarrow \infty$, or

$$\lim_{n \rightarrow \infty} e^{-h_2} \int_{e^{h_2+\lambda n}}^{e^{h_1+\lambda n}} t^\beta \mathbf{P} \{\psi > t\} dt = e^{h_1} C(h_1) - e^{h_2} C(h_2).$$

By the change of variable $t = e^{\lambda n \cdot s}$ we obtain

$$\lim_{n \rightarrow \infty} \int_{e^{h_2}}^{e^{h_1}} e^{\lambda n \beta} s^\beta \mathbf{P} \{\psi > e^{\lambda n \cdot s}\} ds = e^{h_1} C(h_1) - e^{h_2} C(h_2). \quad (15)$$

By virtue of the monotonicity of function $\mathbf{P} \{\psi > x\}$ it follows from this that, for fixed s , the integrand in (15) remains bounded. In particular, there exists a positive number C_3 such that

$$x^\beta \mathbf{P} \{\psi > x\} < C_3 \text{ for all } x > 0. \quad (16)$$

By the selection theorem (see W. Feller [4]), there exists a sequence $n_1, n_2, \dots \rightarrow \infty$ such that, when n runs through it, then $e^{\lambda n \beta} s^\beta \mathbf{P} \{\psi > e^{\lambda n \cdot s}\} \rightarrow q(s)$ at all points of continuity of the latter. It follows from (15) that

$$\int_{e^{h_2}}^{e^{h_1}} q(s) ds = e^{h_1} C(h_1) - e^{h_2} C(h_2). \quad (17)$$

As the limit of nonincreasing functions, function $q(s)/s^\beta$ is also nonincreasing, so that, consequently, function $q(s)$ has no more than a countable number of points of discontinuity. Equation (17) uniquely defines function $q(s)$ at all points of continuity, which entails its

independence on the selection of sequence n_1, n_2, \dots . From this we conclude that, with the possible exception of a countable number of values of h , there exists the limit

$$e^{(h+\lambda n)\beta} \mathbf{P}\{\psi > e^{h+\lambda n}\} \rightarrow q(e^h) = C_2(h) \text{ as } n \rightarrow \infty.$$

If, now, η_1 is a nonnegative random variable, then the strict positivity of function $C_2(h)$ is proven the same as in the case of the nonarithmetic distribution of random variable $\ln \xi_1$. We now show that the limit exists for all real h .

From identity (13) we readily obtain the renewal equation

$$Z_1(e^s) = \int_{-\infty}^{\infty} Z_1(e^{s-t}) F\{dt\} + z_1(e^s),$$

where

$$\begin{aligned} Z_1(x) &= x^\beta \mathbf{P}\{\psi > x\}, \text{ while } z_1(x) = x^\beta \mathbf{P}\{x - \eta_1 < \xi_1 \psi \leq x\} = \\ &= x^\beta [\mathbf{P}\{\psi > x\} - \mathbf{P}\{\xi_1 \psi > x\}]. \end{aligned}$$

We should emphasize the difference between the renewal equations of (18) and (9), amounting to the following, that function $z(e^s)$ of Eq. (9) is directly Riemann-integrable which has not been proven for function $z_1(e^s)$ of Eq. (18).

However, in the case of an arithmetic distribution of random variable $\ln \xi_1$, for justifying the convergence of $\lim_{n \rightarrow \infty} Z_1(e^{h+\lambda n}) = C_1(h)$ it suffices to prove only the (absolute) convergence of the series

$$\sum_{k=-\infty}^{\infty} z_1(e^{h+\lambda k}),$$

and, by virtue of the nonnegativity of the terms of the series, it suffices to show the boundedness (uniformly in N) of the partial sums

$$\sum_{k=-\infty}^N z_1(e^{h+\lambda k}), \quad N=1, 2, \dots$$

We now prove this. We have

$$\begin{aligned} \sum_{k=-\infty}^N z_1(e^{h+\lambda k}) &= \sum_{k=-\infty}^N e^{\beta(h+\lambda k)} [\mathbf{P}\{\psi > e^{h+\lambda k}\} - \mathbf{P}\{\psi \xi_1 > e^{h+\lambda k}\}] = \\ &= \sum_{k=-\infty}^N e^{\beta(h+\lambda k)} \mathbf{P}\{\psi > e^{h+\lambda k}\} - \sum_{k=-\infty}^N e^{\beta(h+\lambda k)} \mathbf{P}\left\{\psi_1 > e^{h+\lambda\left(k - \frac{\ln \xi_1}{\lambda}\right)}\right\}. \end{aligned}$$

Changing the index of summation in the second sum we obtain

$$\begin{aligned} \sum_{k=-\infty}^N z_1(e^{h+\lambda k}) &= \sum_{k=-\infty}^N e^{\beta(h+\lambda k)} \mathbf{P}\{\psi > e^{h+\lambda k}\} - \\ &- \mathbf{M} \sum_{k=-\infty}^{N - \frac{\ln \xi_1}{\lambda}} e^{\beta(h+\lambda k + \ln \xi_1)} \mathbf{P}\{\psi > e^{h+\lambda k}\} = \\ &= \sum_{k=-\infty}^N e^{\beta(h+\lambda k)} \mathbf{P}\{\psi > e^{h+\lambda k}\} - \mathbf{M} \xi_1^\beta \sum_{k=-\infty}^N e^{\beta(h+\lambda k)} \mathbf{P}\{\psi > e^{h+\lambda k}\} + \\ &+ \mathbf{M} \sum_{k=-\infty}^N e^{\beta(h+\lambda k + \ln \xi_1)} \mathbf{P}\{\psi > e^{h+\lambda k}\} - \end{aligned}$$

$$\begin{aligned}
& - M \sum_{k=-\infty}^{N - \frac{\ln \xi_1}{\lambda}} e^{\beta(h + \lambda k + \ln \xi_1)} P\{\psi > e^{h + \lambda k}\} = \\
& = M \left[\sum_{k=-\infty}^N e^{\beta(h + \lambda k + \ln \xi_1)} P\{\psi > e^{h + \lambda k}\} - \right. \\
& \quad \left. - \sum_{k=-\infty}^{N - \frac{\ln \xi_1}{\lambda}} e^{\beta(h + \lambda k + \ln \xi_1)} P\{\psi > e^{h + \lambda k}\} \right] = \\
& = M \sum_{N - \frac{\ln \xi_1}{\lambda} < k \leq N} e^{\beta(h + \lambda k + \ln \xi_1)} P\{\psi > e^{h + \lambda k}\} - \\
& - M \sum_{N < k \leq N - \frac{\ln \xi_1}{\lambda}} e^{\beta(h + \lambda k + \ln \xi_1)} P\{\psi > e^{h + \lambda k}\}.
\end{aligned}$$

Using inequality (16), we find that the right side of the last equation is no greater than

$$C_3 M \sum_{N - \frac{\ln \xi_1}{\lambda} < k \leq N} \xi_1^\beta = C_3 M \frac{\ln^+ \xi_1}{\lambda} \xi_1^\beta < \infty$$

by virtue of the conditions of the theorem.

The periodicity of function $C_2(h)$ is obvious. With this, Theorem 2 is completely proven.

With weaker conditions than in Theorem 2, we have the true

Assertion 1. Let $M \xi_1^\beta = 1$, $P\{\gamma_1 > x\} = O(x^{-\alpha})$ as $x \rightarrow \infty$, where $0 < \beta < \alpha$ and random variable ξ_1 is nondegenerate. Then, $P\{\psi > x\} = O(x^{-\beta})$ for all $x > 0$.

Assertion 1 is an immediate corollary to Theorem 2 and its proof.

We note that constant C_1 and function $C_2(h)$ in Theorem 2 can be identically equal to zero; more than that, the case is possible when $\psi < N$ almost certainly for some real number $N < \infty$.

Assertion 2. Let random series ψ converge almost everywhere. In order that $P\{\psi > N\} = 0$ for some real number $N < \infty$, it is necessary and sufficient that there exist real numbers $N_1 \leq N$ such that $P\{\gamma_1 + N_1 \xi_1 \leq N_1\} = 1$.

Proof. Let $P\{\psi > N\} = 0$. We assume that $P\{\gamma_1 + \xi_1 N_1 > N_1\} > 0$ where N_1 is an exact upper bound to the carrier of random variable ψ . Then, there exist positive numbers Q and ε such that $P\{\gamma_1 + \xi_1 N_1 > N_1 + \varepsilon, \xi_1 < Q\} > 0$, while, since random elements $\begin{pmatrix} \xi_1 & \gamma_1 \\ 0 & 1 \end{pmatrix}$ and ψ_1 are independent and

$$P\left\{\psi_1 > N_1 - \frac{\varepsilon}{2Q}\right\} = P\left\{\psi > N_1 - \frac{\varepsilon}{2Q}\right\} > 0,$$

then also

$$\begin{aligned}
P\left\{\psi > N_1 + \frac{\varepsilon}{2}\right\} &= P\left\{\gamma_1 + \xi_1 \psi_1 > N_1 + \frac{\varepsilon}{2}\right\} \geq \\
&\geq P\left\{\gamma_1 + \xi_1 \left(\psi_1 + \frac{\varepsilon}{2Q}\right) > N_1 + \varepsilon, \xi_1 < Q\right\} \geq \\
&\geq P\left\{\gamma_1 + \xi_1 N_1 > N_1 + \varepsilon, \xi_1 < Q\right\} \cdot P\left\{\psi_1 + \frac{\varepsilon}{2Q} > N_1\right\} > 0.
\end{aligned}$$

We have arrived at a contradiction, so that $P\{\gamma_1 + \xi_1 N_1 \leq N_1\} = 1$. The converse assertion is obvious. Assertion 2 has been proven.

We study analogously the asymptotic behavior of $P\{\psi < x\}$ as $x \rightarrow -\infty$.

LITERATURE CITED

1. U. Grenander, "Stochastic groups," Arkiv. Mat., 4, 189-207 (1961).
2. G. Lev, "Semi-Markov birth and death processes," Teoriya Veroyatn. i Ee Primen., 17, No. 1, 160-166 (1972).
3. G. Lev, "Asymptotic properties of probability of degeneration for semi-Markov birth processes," Teoriya Veroyatn. i Ee Primen., 18, No. 4, 778-789 (1973).
4. W. Feller, Introduction to Probability Theory and Its Applications, Vol. 2, Wiley (1968).
5. V. Maksimov, "Generalized Bernoulli scheme and its limit distribution," Teoriya Veroyatn. i Ee Primen., 18, No. 3, 547-556 (1973).
6. A. Grincevicius, "On continuity of distributions of one sum of dependent quantities, connected by independent walks on the line," Teoriya Veroyatn. i Ee Primen., 19, No. 1, 163-168 (1974).
7. G. Lev, Limit Theorems for Semi-Markov Birth Processes, Candidate's Dissertation, Barnaul (1973).