

# 1. Title of the first chapter

**Definition 1** (Simple random walk in  $\mathbb{Z}$ )

Let  $\{X_n\}_{n=0}^{+\infty}$  be a sequence of independent and identically distributed random variables with values in  $\{-1, +1\}$ . That  $\forall n \in \mathbb{N}$  satisfy the conditions  $P(X_n = 1) = p \in (0, 1)$  and  $P(X_n = -1) = 1 - p = q$ . Let  $S_0 = 0$  and  $S_n = \sum_{i=1}^n X_i$ . We call the pair  $(\{S_n\}_{n=0}^{+\infty}, p)$  Simple random walk in  $\mathbb{Z}$ . In case that  $p = q = \frac{1}{2}$  we call the pair  $(\{S_n\}_{n=0}^{+\infty}, p)$  Symmetric simple random walk in  $\mathbb{Z}$ .

*Remark.* Very often we refer to  $n$  as time,  $X_i$  as  $i$ -th step and  $S_n$  as position in time  $n$ . In simple random walk in  $\mathbb{Z}$  we refer to  $X_i = +1$  as  $i$ -th step was rightwards. If not stated otherwise, we assume that  $S_0 = 0$ .

**Definition 2** (Set of possible positions)

Let  $(\{S_n\}_{n=0}^{+\infty}, p)$  be a random walk. We call the set  $A_n = \{z \in \mathbb{Z}; |z| \leq n, \frac{z+n}{2} \in \mathbb{Z}\}$  set of all possible positions of random walk  $(\{S_n\}_{n=0}^{+\infty}, p)$  in time  $n$ .

**Theorem 1** (Probability of position  $x$  in time  $n$ )

Let  $(\{S_n\}_{n=0}^{+\infty}, p)$  be a random walk and  $A_n$  its set of possible positions.

$$P(S_n = x) = \begin{cases} \binom{n}{\frac{x+n}{2}} p^{\frac{n+x}{2}} q^{\frac{n-x}{2}} & \text{for } x \in A_n, \\ 0, & \text{for } x \notin A_n. \end{cases}$$

*Proof.* Let us define new variables  $r_i = \mathbf{1}_{[X_i=1]}, l_i = \mathbf{1}_{[X_i=-1]}, R_n = \sum_{i=1}^n r_i, L_n = \sum_{i=1}^n l_i$ .  $r_i$  can be interpreted as indicator whether  $i$ -th step was rightwards. Then  $R_n$  is number of rightwards steps and  $L_n$  is number of leftwards steps. We can easily see that  $R_n + L_n = n$  and  $R_n - L_n = S_n$ . Therefore we get by adding these two equations  $R_n = \frac{S_n + n}{2}$ .

$r_i$  has alternative distribution with parameter  $p$   $Alt(p)$ . Therefore  $R_n$  as a sum of independent and identically distributed random variables with  $Alt(p)$  has binomial distribution with parameters  $n$  and  $p$  ( $Bi(n, p)$ ). Therefore we get  $P(R_n = x) = \binom{n}{\frac{x+n}{2}} p^{\frac{n+x}{2}} q^{\frac{n-x}{2}}$ . Where we define  $\binom{a}{x} := 0$  for  $a \in \mathbb{N}, x \in \mathbb{R} \setminus \mathbb{N}, x < 0, x > n$ . Therefore we get  $P(S_n = x) = P(R_n = \frac{x+n}{2}) = \binom{n}{\frac{x+n}{2}} p^{\frac{n+x}{2}} q^{\frac{n-x}{2}}$ .  $\square$

**Lemma 2** (Spatial homogeneity)

$$P(S_n = j \mid S_0 = a) = P(S_n = j + b \mid S_0 = a + b) \forall b \in \mathbb{Z}$$

*Proof.*  $P(S_n = j \mid S_0 = a) = P\left(\sum_{i=1}^n X_i = j - a\right)$   
 $= P\left(\sum_{i=1}^n X_i = (j + b) - (a + b)\right) = P(S_n = j + b \mid S_0 = a + b)$ .  $\square$

**Lemma 3** (Temporal homogeneity)

$$P(S_n = j \mid S_0 = a) = P(S_{n+m} = j \mid S_m = a) \forall m \in \mathbb{N}$$

*Proof.*  $P(S_n = j \mid S_0 = a) = P\left(\sum_{i=1}^n X_i = j - a\right)$   
 $= P\left(\sum_{i=m+1}^{m+n} X_i = j - a\right) = P(S_{n+m} = j \mid S_m = a)$ .  $\square$

**Lemma 4** (Markov property)

Let  $n \geq m$  and  $a_i \in \mathbb{Z}, i \in \mathbb{N}$ . Then  $P(S_n = j \mid S_0 = a_0, S_1 = a_1, \dots, S_m = a_m) = P(S_n = j \mid S_m = a_m)$

*Proof.* Once  $S_m$  is known, then distribution of  $S_n$  depends only on steps  $X_{m+1}, X_{m+2}, \dots, X_n$  and therefore cannot be dependent on any information concerning values  $X_1, X_2, \dots, X_m - 1$  and accordingly  $S_1, S_2, \dots, S_m - 1$ .  $\square$

*Remark.* In symmetric random walk, everything can be counted by number of possible paths from point to point.

**Definition 3** (Number of possible paths)

Let  $N_n(a, b)$  be number of possible paths of random walk  $(\{S_n\}_{n=0}^{+\infty}, p)$  from point  $(0, a)$  to point  $(n, b)$  and  $N_n^x(a, b)$  be number of possible paths from point  $(0, a)$  to point  $(n, b)$  that visit point  $(z, x)$  for some  $z \in \{0, \dots, n\}$ .

**Theorem 5**

Let  $a, b \in \mathbb{Z}, n \in \mathbb{N}$  then  $N_n(a, b) = \binom{n}{\frac{1}{2}(n+b-a)}$ .

*Proof.* Let us choose a path from  $(0, a)$  to  $(n, b)$  and let  $\alpha$  be number of rightwards steps and  $\beta$  be number of leftwards steps. Then  $\alpha + \beta = n$  and  $\alpha - \beta = b - a$ . By adding these two equations we get that  $\alpha = \frac{1}{2}(n + b - a)$ . The number of possible paths is the number of ways of picking  $\alpha$  rightwards steps from  $n$  steps. Therefore we get  $N_n(a, b) = \binom{n}{\alpha} = \binom{n}{\frac{1}{2}(n+b-a)}$ .  $\square$

**Theorem 6** (Reflection principle)

Let  $a, b > 0$ , then  $N_n^0(a, b) = N_n(-a, b)$ .

*Proof.* Each path from  $(0, -a)$  to  $(n, b)$  has to intersect  $x$ -axis at least once at some point. Let  $k$  be the time of earliest intersection with  $x$ -axis. By reflexing the segment from  $(0, -a)$  to  $(k, 0)$  in the  $x$ -axis, we get a path from point  $(0, a)$  to  $(n, b)$  which visits 0 at point  $k$ . Because reflection is bijective operation on sets of paths, we get the correspondence between the collections of such paths.  $\square$

**Definition 4** (Return to origin)

Let  $(\{S_n\}_{n=0}^{+\infty}, p)$  be a random walk. Then if  $\exists k \in \mathbb{N}$  such that  $S_k = 0$  then we say that in  $k$ -th step occurred return to origin. Let  $(\{S_n\}_{n=0}^{+\infty}, p)$  be a random walk. Then if  $S_1, S_2, \dots, S_{2n-1} \neq 0$  and  $S_{2n} = 0$

**Theorem 7** (Ballot theorem)

Let  $n, b \in \mathbb{N}$  Number of paths from point  $(0, 0)$  to point  $(n, b)$  which do not return to origin is equal to  $\frac{b}{n} N_n(0, b)$

*Proof.* Let us call  $N$  the number of paths we are referring to. Because the path ends at point  $(n, b)$ , the first step has to be rightwards. Therefore we now have  $N = N_{n-1}(1, b) - N_{n-1}^0(1, b) = N_{n-1}(1, b) - N_{n-1}(-1, b)$ . The last equation was acquired using Reflection principle (6). We now have:  $N_{n-1}(1, b) - N_{n-1}(-1, b) =$

$$\begin{aligned} & \binom{n-1}{\frac{n+b}{2}-1} - \binom{n-1}{\frac{n-b}{2}} = \frac{(n-1)!}{(\frac{n+b}{2}-1)!(\frac{n-b}{2})!} - \frac{(n-1)!}{(\frac{n+b}{2})!(\frac{n-b}{2}-1)!} = \frac{(n-1)!}{(\frac{n+b}{2}-1)!(\frac{n-b}{2})(\frac{n-b}{2}-1)!} - \\ & \frac{(n-1)!}{(\frac{n+b}{2})(\frac{n+b}{2}-1)!(\frac{n-b}{2}-1)!} = \frac{(n-1)!}{(\frac{n+b}{2}-1)!(\frac{n-b}{2}-1)!} \left( \frac{1}{\frac{n-b}{2}} - \frac{1}{\frac{n+b}{2}} \right) = \frac{1}{n} \frac{n!}{(\frac{n+b}{2}-1)!(\frac{n-b}{2}-1)!} \left( \frac{(\frac{n+b}{2}-\frac{n-b}{2})}{(\frac{n-b}{2})(\frac{n+b}{2})} \right) \\ & \frac{b}{n} \frac{n!}{(\frac{n+b}{2})!(\frac{n-b}{2})!} = \frac{b}{n} \binom{n}{\frac{n+b}{2}} = \frac{b}{n} N_n(0, b) \end{aligned} \quad \square$$

**Definition 5**

$$M_n = \max\{S_i, i \in \{0, 1, \dots, n\}\}$$

**Theorem 8** (Probability of maximum up to time  $n$ )

Let  $(\{S_n\}_{n=0}^{+\infty}, p)$  be a random walk .

$$P(M_n \geq r, S_n = b) = \begin{cases} P(S_n = b) & \text{for } b \geq r, \\ P(S_n = 2r - b) \left(\frac{q}{p}\right)^{r-b}, & \text{for otherwise.} \end{cases}$$

*Proof.* Let us firstly consider the easier case in which  $b \geq r$ . Because we defined  $M_n$  as  $\max\{S_i, i \in \{0, 1, \dots, n\}\}$  we get that  $M_n \geq b \geq r$  therefore  $[M_n \geq r] \subset [S_n = b]$  therefore we get  $P(M_n \geq r, S_n = b) = P(S_n = b)$ . Let  $r \geq 1, b < r$ .  $N_n^r(0, b)$  stands for number of paths from point  $(0, 0)$  to point  $(n, b)$  which reach up to  $r$ . Let  $k \in \{0, 1, \dots, n\}$  denote the first time we reach  $r$ . By reflection principle (6), we can reflex the segment from  $(k, r)$  to  $(n, b)$  in the axis:  $y = r$ . Therefore we now have path from  $(0, 0)$  to  $(n, 2r - b)$  and we get that  $N_n^r(0, b) = N_n(0, 2r - b)$ .  $P(S_n = b, M_n \geq r) = N_n^r(0, b) p^{\frac{n+b}{2}} q^{\frac{n-b}{2}} = N_n(0, 2r - b) p^{\frac{n+(2r-b)}{2}} q^{\frac{n-(2r-b)}{2}} p^{\frac{n+b}{2}} q^{\frac{n-b}{2}} p^{b-r} q^{r-b} = \left(\frac{q}{p}\right)^{r-b} P(S_n = 2r - b)$ .  $\square$

**Definition 6** (Walk reaching new maximum at particular time)

Let  $b > 0$ .  $f_b(n)$  denotes the probability that we reach new maximum  $b$  in time  $n$ .  $f_b(n) = P(M_{n-1} = S_{n-1} = b - 1, S_n = b)$

**Theorem 9** (Probability of reaching new maximum  $b$  in time  $n$ )

Let  $b > 0$  then  $f_b(n) = \frac{b}{n} P(S_n = b)$ .

*Proof.*  $f_b = P(M_{n-1} = S_{n-1} = b - 1, S_n = b) = p P(M_{n-1} = S_{n-1} = b - 1) = p (P(M_{n-1} \geq b - 1, S_{n-1} = b - 1) - P(M_{n-1} \geq b, S_{n-1} = b - 1)) = p (P(S_{n-1} = b - 1) - \frac{q}{p} P(S_{n-1} = b + 1)) = p P(S_{n-1} = b - 1) - q P(S_{n-1} = b + 1) = \binom{n-1}{\frac{n}{2} + \frac{b}{2} - 1} p^{\frac{n}{2} + \frac{b}{2}} q^{\frac{n}{2} - \frac{b}{2}} - \binom{n-1}{\frac{n}{2} + \frac{b}{2}} p^{\frac{n}{2} + \frac{b}{2}} q^{\frac{n}{2} - \frac{b}{2}} = p^{\frac{n}{2} + \frac{b}{2}} q^{\frac{n}{2} - \frac{b}{2}} \left( \frac{(n-1)!}{(\frac{n}{2} + \frac{b}{2} - 1)! (\frac{n}{2} - \frac{b}{2})!} - \frac{(n-1)!}{(\frac{n}{2} + \frac{b}{2})! (\frac{n}{2} - \frac{b}{2} - 1)!} \right) = p^{\frac{n}{2} + \frac{b}{2}} q^{\frac{n}{2} - \frac{b}{2}} \left( \frac{(n-1)!}{(\frac{n}{2} + \frac{b}{2})! (\frac{n}{2} - \frac{b}{2})!} \right) \left( \frac{1}{\frac{n}{2} - \frac{b}{2}} - \frac{1}{\frac{n}{2} + \frac{b}{2}} \right) = p^{\frac{n}{2} + \frac{b}{2}} q^{\frac{n}{2} - \frac{b}{2}} \frac{b}{n} \left( \frac{n!}{(\frac{n}{2} + \frac{b}{2})! (\frac{n}{2} - \frac{b}{2})!} \right) = \frac{b}{n} p^{\frac{n}{2} + \frac{b}{2}} q^{\frac{n}{2} - \frac{b}{2}} \binom{n}{\frac{n}{2} + \frac{b}{2}} = \frac{b}{n} P(S_n = b)$ .  $\square$

**Theorem 10** (XXXMean number of visits to  $b$  before returning to origin in symmetric random walk)

Let  $(\{S_n\}_{n=0}^{+\infty}, p)$  be a symmetric random walk . Mean number  $\mu_b$  of visits of the walk to point  $b$  before returning to origin is equal to 1.

*Proof.* aa  $\square$

**Definition 7** (Return to origin)

Let  $(\{S_n\}_{n=0}^{+\infty}, p)$  be a random walk . Let  $k \in \mathbb{N}$ . We say a return to origin occurred in time  $2k$  if  $S_{2k} = 0$ . The probability that in time  $2k$  occurred a return to origin shall be denoted by  $u_{2k}$ . We say that in time  $2k$  occurred first return to origin if  $S_1, S_2, \dots, S_{2k-1} \neq 0$  and  $S_{2k} = 0$ . The probability that in time  $2k$  occurred first return to origin shall be denoted by  $f_{2k}$ . By definition  $f_0 = 0$ . Let  $\alpha_{2n}(2k)$  denote  $u_{2k} u_{2(n-k)}$

**Lemma 11** (Binomial identity)

Let  $n, k \in \mathbb{N}, n > k : \binom{n-1}{k} - \binom{n-1}{k-1} = \frac{n-2k}{n} \binom{n}{k}$

$$\begin{aligned} \text{Proof. } \binom{n-1}{k} - \binom{n-1}{k-1} &= \frac{(n-1)!}{k!(n-k-1)!} - \frac{(n-1)!}{(k-1)!(n-k)!} = \frac{(n-1)!}{(k-1)!(n-k-1)!} \left( \frac{1}{k} - \frac{1}{n-k} \right) \\ &= \frac{1}{n} \frac{n!}{(k-1)!(n-k-1)!} \frac{n-2k}{k(n-k)} = \frac{n-2k}{n} \frac{n!}{k!(n-k)!} = \frac{n-2k}{n} \binom{n}{k} \quad \square \end{aligned}$$

**Lemma 12** (Main lemma)

Let  $(\{S_n\}_{n=0}^{+\infty}, p)$  be a symmetrical random walk. Then  $P(S_1 \cdot S_2, \cdot, \dots, S_{2n} \neq 0) = P(S_{2n} = 0)$ .

$$\begin{aligned} \text{Proof. } P(S_1 \cdot S_2, \cdot, \dots, S_{2n} \neq 0) &= \sum_{i=-\infty}^{+\infty} P(S_1 \cdot S_2, \cdot, \dots, S_{2n-1} \neq 0, S_{2n} = 2i) = \\ &= \sum_{i=-n}^n P(S_1 \cdot S_2, \cdot, \dots, S_{2n-1} \neq 0, S_{2n} = 2i) = 2 \cdot \sum_{i=1}^n P(S_1 \cdot S_2, \cdot, \dots, S_{2n-1} \neq 0, S_{2n} = 2i) = \\ &= 2 \sum_{i=1}^n \frac{2i}{2n} P(S_{2n} = 2k) = 2 \sum_{i=1}^n \frac{2i}{2n} \binom{2n}{n+i} 2^{-2n} = 2 \cdot 2^{-2n} \sum_{i=1}^n \left( \binom{2n-1}{m+k-1} - \binom{2n-1}{m+k} \right) = 2 \cdot \\ &= 2^{-2n} \binom{2n-1}{n} = 2^{-2n} \frac{2n}{n} \binom{2n-1}{n} = 2^{-2n} \frac{2n(2n-1)!}{m(m-1)!m!} = 2^{-2n} \frac{(2n)!}{m!m!} = 2^{-2n} \binom{2n}{n} = P(S_{2n} = 0) \quad \square \end{aligned}$$

**Theorem 13**

Let  $(\{S_n\}_{n=0}^{+\infty}, p)$  be a symmetric random walk. The probability that the last return to origin up to time  $2n$  occurred in time  $2k$  is  $P(S_{2k} = 0) P(S_{2(n-k)} = 0)$ .

$$\begin{aligned} \text{Proof. } \alpha_{2n}(2k) &= u_{2k} u_{2(n-k)} = P(S_{2k} = 0) P(S_{2k+1} \cdot S_{2k+2}, \dots, S_{2n} \neq 0 \mid S_{2k} = 0) = \\ &= P(S_{2k} = 0) P(S_1 \cdot S_2, \dots, S_{2(n-k)} \neq 0) = P(S_{2k} = 0) P(S_{2(n-k)} = 0) \quad \square \end{aligned}$$

**Theorem 14**

Let  $b \in \mathbb{Z}$ .  $P(S_1 \cdot S_2, \dots, S_n \neq 0, S_n = b) = \frac{|b|}{n} P(S_n = b)$ .

*Proof.* Let us Without loss of generality assume that  $b > 0$ . In that case, first step has to be rightwards ( $X_1 = +1$ ). Now we have path from point  $(1, 1)$  to point  $(n, b)$  that does not return to origin. By Ballot theorem 7 there are  $\frac{b}{n} N_n(0, b)$  such paths. Each path consists of  $\frac{n+b}{2}$  rightwards steps and  $\frac{n-b}{2}$  leftwards steps. Therefore  $P(S_1 \cdot S_2, \dots, S_n \neq 0, S_n = b) = \frac{b}{n} N_n(0, b) p^{\frac{n+b}{2}} q^{\frac{n-b}{2}} = \frac{b}{n} P(S_n = b)$ . Case  $b < 0$  is identical.  $\square$

**Lemma 15**

Let  $(\{S_n\}_{n=0}^{+\infty}, p)$  be a symmetric random walk.  $P(S_1, S_2, \dots, S_{2n} > 0) = \frac{1}{2} P(S_{2n} = 0) = \frac{1}{2} u_{2n}$ .

*Proof.* Because  $S_i > 0 \forall i \in \mathbb{N}$  the first step has to be rightwards ( $X_1 = S_1 = 1$ ). Therefore we get  $P(S_1, S_2, \dots, S_{2n} > 0) = \sum_{r=1}^{n-1} P(S_1, S_2, \dots, S_{2n} > 0, S_{2n} = 2r)$ .

The  $r$ -th term follows equation:  $P(S_1, S_2, \dots, S_{2n} > 0, S_{2n} = 2r)$

$$\begin{aligned} &= P(X_1 = 1, S_2, \dots, S_{2n} > 0, S_{2n} = 2r) \\ &= \frac{1}{2} P(S_2, S_3, \dots, S_{2n} > 0, S_{2n} = 2r) \\ &= \frac{1}{2} (P(S_{2n} = 2r) - P(S_2 \cdot S_3 \cdot \dots \cdot S_{2n} = 0, S_{2n} = 2r)) \\ &= \frac{1}{2} \left( \frac{1}{2} {}^{2n-1}N_{2n-1}(1, 2r) - \frac{1}{2} {}^{2n-1}N_{2n-1}^0(1, 2r) \right) \\ &= \frac{1}{2} \frac{1}{2} {}^{2n-1}N_{2n-1}(1, 2r) - \frac{1}{2} {}^{2n-1}N_{2n-1}^0(1, 2r) \\ &= \frac{1}{2} \frac{1}{2} {}^{2n-1}N_{2n-1}(1, 2r) - N_{2n-1}(-1, 2r) \end{aligned}$$

$= \frac{1}{2} \frac{1}{2}^{2n-1} \left( \binom{2n-1}{n+r-1} - \binom{2n-1}{n+r} \right)$ . Because of the fact that the negative parts of  $r$ -th terms cancel against the positive parts of  $(r+1)$ -st terms and the sum reduces to just  $\frac{1}{2} \frac{1}{2}^{2n-1} \binom{2n-1}{n} = \frac{1}{2} \cdot 2 \cdot \frac{1}{2}^{2n} \binom{2n-1}{n} = \frac{1}{2} \frac{1}{2}^{2n} \frac{2(2n-1)!}{n!(n-1)!} = \frac{1}{2} \frac{1}{2}^{2n} \frac{(2n)!}{n!n!} = \frac{1}{2} \frac{1}{2}^{2n} \binom{2n}{n} = \frac{1}{2} \mathbb{P}(S_{2n} = 0) = \frac{1}{2} u_{2n}$ .  $\square$

**Theorem 16** (No return=return)

$$\mathbb{P}(S_1, S_2, \dots, S_{2n} \neq 0) = \mathbb{P}(S_{2n} = 0) = u_{2n}$$

*Proof.* The event  $[S_1, S_2, \dots, S_{2n} \neq 0]$  can be split into two disjoint events:  $= [S_1, S_2, \dots, S_{2n} < 0] \cup [S_1, S_2, \dots, S_{2n} > 0]$ . By previous theorem (15) we get that probability of both of them is  $\frac{1}{2} u_{2n}$ . Because the events are disjoint we can sum their probabilities and we get the result.  $\square$

*Corollary.*  $\mathbb{P}(S_1, S_2, \dots, S_{2n} \geq 0) = \mathbb{P}(S_{2n} = 0) = u_{2n}$

*Proof.*  $\frac{1}{2} u_{2n} = \mathbb{P}(S_1, S_2, \dots, S_{2n} > 0) = \mathbb{P}(X_1 = 1, S_2, S_3, \dots, S_{2n} \geq 1)$   
 $= \frac{1}{2} \mathbb{P}(S_2, S_3, \dots, S_{2n} \geq 1 \mid S_1 = 1)$   
 $= \frac{1}{2} \mathbb{P}(S_1, S_2, \dots, S_{2n-1} \geq 1 \mid S_0 = 1)$   
 $= \frac{1}{2} \mathbb{P}(S_1, S_2, \dots, S_{2n-1} \geq 0)$   
 $= \frac{1}{2} \mathbb{P}(S_1, S_2, \dots, S_{2n} \geq 0)$ . Therefore  $\mathbb{P}(S_1, S_2, \dots, S_{2n} \geq 0) = u_{2n}$ .  $\square$

**Theorem 17**

$$f_{2n} = u_{2n-2} - u_{2n}$$

*Proof.* The event  $[S_1, S_2, \dots, S_{2n-1} \neq 0]$  can be split into two disjoint events:  $[S_1, S_2, \dots, S_{2n-1} \neq 0, S_{2n} = 0]$  and  $[S_1, S_2, \dots, S_{2n-1} \neq 0, S_{2n} \neq 0]$ . Therefore  $\mathbb{P}(S_1, S_2, \dots, S_{2n-1} \neq 0) = \mathbb{P}(S_1, S_2, \dots, S_{2n-1} \neq 0, S_{2n} = 0) + \mathbb{P}(S_1, S_2, \dots, S_{2n-1} \neq 0, S_{2n} \neq 0)$ . Therefore we get  $f_{2n} = \mathbb{P}(S_1, S_2, \dots, S_{2n-1} \neq 0, S_{2n} = 0) = \mathbb{P}(S_1, S_2, \dots, S_{2n-1} \neq 0) - \mathbb{P}(S_1, S_2, \dots, S_{2n} \neq 0)$ . Because  $2n-1$  is odd.  $\mathbb{P}(S_{2n-1} = 0) = 0$ . Therefore the first term is equal to  $\mathbb{P}(S_1, S_2, \dots, S_{2n-2} \neq 0)$  which is by 16 equal to  $u_{2n-2}$ . Second term is by 16 equal to  $u_{2n}$ . Therefore we get the result.  $\square$

*Corollary.*  $f_{2n} = \frac{1}{2n-1} u_{2n}$

*Proof.*  $u_{2n-2} = \frac{1}{2}^{2n-2} \binom{2n-2}{n-1} = 4 \cdot \frac{1}{2}^{2n} \frac{(2n-2)!}{(n-1)!(n-1)!} = \frac{4n^2}{(2n)(2n-1)} \frac{1}{2}^{2n} \binom{2n}{n} = \frac{2n}{2n-1} u_{2n}$ . Therefore  $u_{2n-2} - u_{2n} = u_{2n} \left( \frac{2n}{2n-1} - 1 \right) = u_{2n} \frac{1}{2n-1}$ .  $\square$

**Theorem 18** (Arcsine law for last visits)

Let  $k, n \in \mathbb{N}, k \leq n$ . The probability that up to time  $2n$  a return to origin occurred in time  $2k$  is given by  $\alpha_{2n}(2k) = u_{2n} u_{2(n-k)}$ .

*Proof.* The probability involved can be rewritten as:

$$\begin{aligned} & \mathbb{P}(S_{2k+1}, S_{2k+2}, \dots, S_{2n} \neq 0, S_{2k} = 0) \\ &= * * * \mathbb{P}(S_{2k+1}, S_{2k+2}, \dots, S_{2n} \neq 0 \mid S_{2k} = 0) \mathbb{P}(S_{2k} = 0) \\ &= * \mathbb{P}(S_1, S_2, \dots, S_{2(n-k)} \neq 0) \mathbb{P}(S_{2k} = 0) \\ &= * * u_{2(n-k)} u_{2k} \end{aligned} \quad \square$$

**Definition 8** (Time spend on the positive and negative sides)

Let  $(\{S_n\}_{n=0}^{+\infty}, p)$  be a random walk. We say that the walk spent  $\tau$  time units of  $n$  on the positive side if  $\sum_{i=1}^n \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = \tau$ . Let  $\beta_n(\tau)$  denote the probability of such an event. We say that the walk spent  $\zeta$  time units of  $n$  on the negative side if  $\sum_{i=1}^n \mathbf{1}_{[S_i < 0 \vee S_{i-1} < 0]} = \zeta$ .

**Theorem 19** (Arcsine law for sojourn times-OWN PROOF)

Let  $(\{S_n\}_{n=0}^{+\infty}, p)$  be a symmetric random walk . Then  $\beta_{2n}(2k) = \alpha_{2n}(2k)$ .

*Proof.* Firstly let us start with degenerate cases.  $*f\beta_{2n}(2n)$

$$= P\left(\sum_{i=1}^{2n} \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2n\right) = P(S_1, S_2, \dots, S_{2n} \geq 0) = *u_{2n}.$$

By symmetry  $\beta_{2n}(0) = \beta_{2n}(2n) = u_{2n}$ . Let  $1 \leq k \leq v-1$ , where  $0 \leq v \leq n$ .

For such  $k$  stands equation:  $\beta_{2n}(2k) = P\left(\sum_{i=1}^n \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k\right)$

$$= *a \sum_{r=1}^n P\left(\sum_{i=1}^n \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k, S_1, S_2, \dots, S_{2r-1} \neq 0, S_{2r} = 0\right)$$

$$= *b \sum_{r=1}^n P\left(\sum_{i=1}^n \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k, S_1, S_2, \dots, S_{2r-1} < 0, S_{2r} = 0\right)$$

$$+ \sum_{r=1}^n P\left(\sum_{i=1}^n \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k, S_1, S_2, \dots, S_{2r-1} > 0, S_{2r} = 0\right)$$

$$= *c \sum_{r=1}^n P\left(\sum_{i=1}^n \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k \mid S_1, S_2, \dots, S_{2r-1} < 0, S_{2r} = 0\right)$$

$$P(S_1, S_2, \dots, S_{2r-1} < 0, S_{2r} = 0)$$

$$+ \sum_{r=1}^n P\left(\sum_{i=1}^n \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k \mid S_1, S_2, \dots, S_{2r-1} > 0, S_{2r} = 0\right)$$

$$P(S_1, S_2, \dots, S_{2r-1} > 0, S_{2r} = 0)$$

$$= *d \sum_{r=1}^n \frac{1}{2} f_{2r} P\left(\sum_{i=2r+1}^{2n} \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k \mid S_{2r} = 0\right)$$

$$+ \sum_{r=1}^n \frac{1}{2} f_{2r} P\left(\sum_{i=2r+1}^{2n} \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k - 2r \mid S_{2r} = 0\right)$$

$$= *e \sum_{r=1}^n \frac{1}{2} f_{2r} P\left(\sum_{i=1}^{2n-2r} \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k\right) + \sum_{r=1}^n \frac{1}{2} f_{2r} P\left(\sum_{i=1}^{2n-2r} \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k - 2r\right) =$$

$$\sum_{r=1}^n \frac{1}{2} f_{2r} \beta_{2n-2r}(2k) + \sum_{r=1}^n \frac{1}{2} f_{2r} \beta_{2n-2r}(2k - 2r) \text{ Now let us proceed by induction.}$$

Case for  $v = 1$  is trivial because it implies degenerous case from  $*f$ . Let the

statment be true for  $v \leq n-1$ , then  $\sum_{r=1}^n \frac{1}{2} f_{2r} \beta_{2n-2r}(2k) + \sum_{r=1}^n \frac{1}{2} f_{2r} \beta_{2n-2r}(2k - 2r)$

$$= *g \sum_{r=1}^n \frac{1}{2} f_{2r} \alpha_{2n-2r}(2k) + \sum_{r=1}^n \frac{1}{2} f_{2r} \alpha_{2n-2r}(2k - 2r)$$

$$= *h \sum_{r=1}^n \frac{1}{2} f_{2r} u_{2k} u_{2n-2r-2k} + \sum_{r=1}^n \frac{1}{2} f_{2r} u_{2k-2r} u_{2n-2k}$$

$$= *i \frac{1}{2} u_{2k} \sum_{r=1}^n f_{2r} u_{2n-2r-2k} + \frac{1}{2} u_{2n-2k} \sum_{r=1}^n f_{2r} u_{2k-2r} u_{2n-2k}$$

$$= *j \frac{1}{2} u_{2n-2k} u_{2k} + \frac{1}{2} u_{2n-2k} u_{2k} = u_{2n-2k} u_{2k}$$

$$= *h \alpha_{2n}(2k)$$

□

**Definition 9** (Change of a sign)

Let  $(\{S_n\}_{n=0}^{+\infty}, p)$  be a random walk . We say that in time  $n$  occurred a change of sign if  $S_{n-1} \cdot S_{n+1} = -1$  in other words if  $(S_{n-1} = +1 \wedge S_{n+1} = -1) \vee (S_{n-1} = -1 \wedge S_{n+1} = +1)$ . We shall denote the probability that up to time  $n$  occurred  $r$  changes of sign by  $\xi_{r,n}$ .

**Theorem 20** (Change of a sign)

Let  $(\{S_n\}_{n=0}^{+\infty}, p)$  be a symmetric random walk . The probability  $\xi_{r,2n+1} = 2 P(S_{2n+1} = 2r + 1)$

*Proof.* Feller

□

## 1.1 Problem chapter 9 Feller

**Definition 10** ( $\delta, \varepsilon$ )

Let  $(\{S_n\}_{n=0}^{+\infty}, p)$  be a symmetric random walk.  $\delta_n(k)$  shall denote  $P\left(\sum_{i=1}^n \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = k, S_n = 0\right)$ ,  $\varepsilon_n^r(k)$  shall denote  $P\left(\sum_{i=1}^n \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = k, S_1, S_2, \dots, S_{r-1} \neq 0, S_r = 0, S_n = 0\right)$ ,  $\varepsilon_n^{r,+}(k)$  shall denote  $P\left(\sum_{i=1}^n \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = k, S_1, S_2, \dots, S_{r-1} > 0, S_r = 0, S_n = 0\right)$ ,  $\varepsilon_n^{r,-}(k)$  shall denote  $P\left(\sum_{i=1}^n \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = k, S_1, S_2, \dots, S_{r-1} < 0, S_r = 0, S_n = 0\right)$ .

**Lemma 21** (Factorization of  $\delta_{2n}(2k)$ )

$$\delta_{2n}(2k) = \frac{1}{2} \sum_{r=1}^n (f_{2r} \delta_{2n-2r}(2k-2r) + f_{2r} \delta_{2n-2r}(2r)).$$

*Proof.* Because  $S_{2n} = 0$  a return to origin must have happened. Let  $2r$  the time of first return to origin, where  $r \in \{1, 2, \dots, n\}$ . By the law of total probability:

$$\begin{aligned} \delta_{2n}(2k) &= P\left(\sum_{i=1}^{2n} \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = k, S_{2n} = 0\right) \\ &= \sum_{r=1}^n P\left(\sum_{i=1}^{2n} \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k, S_1, S_2, \dots, S_{2r-1} \neq 0, S_{2r} = 0, S_{2n} = 0\right) \text{ which can} \\ &\text{be again by the law of total probability factorized as:} \\ &= \sum_{r=1}^n P\left(\sum_{i=1}^{2n} \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k, S_1, S_2, \dots, S_{2r-1} \neq 0, S_{2r} = 0, S_{2n} = 0\right) \\ &= \sum_{r=1}^n P\left(\sum_{i=1}^{2n} \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k, S_1, S_2, \dots, S_{2r-1} > 0, S_{2r} = 0, S_{2n} = 0\right) \\ &+ \sum_{r=1}^n P\left(\sum_{i=1}^{2n} \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k, S_1, S_2, \dots, S_{2r-1} < 0, S_{2r} = 0, S_{2n} = 0\right) \\ &= \sum_{r=1}^n \varepsilon_{2n}^{2r,+}(2k) + \sum_{r=1}^n \varepsilon_{2n}^{2r,-}(2k). \end{aligned}$$

Now let us calculate  $\varepsilon_{2n}^{2r,+}(2k)$

$$\begin{aligned} &= P\left(\sum_{i=1}^{2n} \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k, S_1, S_2, \dots, S_{2r-1} > 0, S_{2r} = 0, S_{2n} = 0\right) \\ &= *a P\left(\sum_{i=1}^{2n} \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k, S_{2n} = 0 \mid S_1, S_2, \dots, S_{2r-1} > 0, S_{2r} = 0\right) \\ &P(S_1, S_2, \dots, S_{2r-1} > 0, S_{2r} = 0) \\ &= P\left(\sum_{i=2r+1}^{2n} \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k-2r, S_{2n} = 0 \mid S_{2r} = 0\right) \frac{1}{2} f_{2r} \\ &= *b P\left(\sum_{i=1}^{2n-2r} \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k-2r, S_{2n-2r} = 0\right) \frac{1}{2} f_{2r} \\ &= *c \delta_{2n-2r}(2k-2r) \frac{1}{2} f_{2r}. \text{ Similarly } \varepsilon_{2n}^{2r,-}(2k) \\ &= P\left(\sum_{i=1}^{2n} \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k, S_1, S_2, \dots, S_{2r-1} < 0, S_{2r} = 0, S_{2n} = 0\right) \\ &= *a P\left(\sum_{i=1}^{2n} \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k, S_{2n} = 0 \mid S_1, S_2, \dots, S_{2r-1} < 0, S_{2r} = 0\right) \\ &P(S_1, S_2, \dots, S_{2r-1} < 0, S_{2r} = 0) \\ &= P\left(\sum_{i=2r+1}^{2n} \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k, S_{2n} = 0 \mid S_{2r} = 0\right) \frac{1}{2} f_{2r} \\ &= *b P\left(\sum_{i=1}^{2n-2r} \mathbf{1}_{[S_i > 0 \vee S_{i-1} > 0]} = 2k, S_{2n-2r} = 0\right) \frac{1}{2} f_{2r} = *c \delta_{2n-2r}(2k) \frac{1}{2} f_{2r}. \end{aligned}$$

$$\begin{aligned} \text{Therefore } \delta_{2n}(2k) &= \frac{1}{2} \sum_{r=1}^n f_{2r} \delta_{2n-2r}(2k-2r) + \frac{1}{2} \sum_{r=1}^n f_{2r} \delta_{2n-2r}(2k) \\ &= \frac{1}{2} \sum_{r=1}^n (f_{2r} \delta_{2n-2r}(2k-2r) + f_{2r} \delta_{2n-2r}(2r)) \end{aligned} \quad \square$$

**Theorem 22** (Equidistributional theorem-ALMOST COMPLETE OWN PROOF) ||  
*Let  $(\{S_n\}_{n=0}^{+\infty}, p)$  be a symmetric random walk and  $n \in \mathbb{N}$ , then  $\forall k, l \in \{0, 1, \dots, n\}$  :*  
 $\delta_{2n}(2k) = \delta_{2n}(2l) = \frac{u_{2n}}{n+1}$ .

*Proof.* Let us prove this statement by induction in  $n$ . In case that  $n = 1$  we have two options for  $k$ . Either  $k = 0$  or  $k = 1$ .  $\delta_2(0) = \mathbf{P}(S_1 < 0, S_2 = 0) = \frac{1}{2}f_2 = *a\frac{1}{2}u_{2\frac{1}{2-1}} = \frac{u_2}{2}\delta_2(2) = \mathbf{P}(S_1 > 0, S_2 = 0) = \frac{1}{2}f_2 = \frac{u_2}{2}$ .

Let the statment be true for all  $l \leq n-1$ . In that case  $\delta_{2(n-l)}(2k) = \frac{u_{2(n-l)}}{n-l+1}$ . We want to show that  $\delta_{2n} = \frac{u_{2n}}{n+1}$ .

$$\begin{aligned} \text{Let us calculate } \delta_{2n}. \quad \delta_{2n} &= *b\frac{1}{2} \sum_{r=1}^n (f_{2r} \delta_{2n-2r}(2k-2r) + f_{2r} \delta_{2n-2r}(2r)) = \\ \frac{1}{2} \sum_{r=1}^n \left( f_{2r} u_{2n-2r} \frac{1}{n-r+1} + f_{2r} u_{2n-2r} \frac{1}{n-r+1} \right) &= \sum_{r=1}^n \left( \frac{f_{2r} u_{2n-2r}}{n-r+1} \right) = ? \frac{u_{2n}}{n+1} \text{ Because???} \end{aligned} \quad \square$$

**Lemma 23**

$$\sum_{r=1}^n \left( \frac{1}{n-r+1} \left( 2^{-(2r-2)} \binom{2r-2}{r-1} - 2^{-2r} \binom{2r}{r} \right) 2^{-(2n-2r)} \binom{2n-2r}{n-r} \right) = \frac{1}{n+1} 2^{-2n} \binom{2n}{n}$$

*Proof.* □



## 2. Multi dimensional random walk

**Definition 11** (Type I random walk in  $\mathbb{Z}^m$ )

Let  $m \in \mathbb{N}$ . Let  $e_i$  denote

$$\left( \underbrace{0 \ 0 \ \dots \ 0}_{i-1} \ 1 \ \underbrace{0 \ \dots \ 0 \ 0}_{m-i} \right)^T \forall i \in \{1, 2, \dots, m\}.$$

Let  $\{X_n\}_{n=0}^{+\infty}$  be a sequence of independent and identically distributed random variables with values in  $\{-e_1, +e_1, -e_2, +e_2, \dots, -e_m, +e_m\}$ . That  $\forall n \in \mathbb{N}$  satisfy the conditions  $P(X_n = e_i) = p_i \in (0, 1)$ ,  $P(X_n = -e_i) = q_i \in (0, 1)$  and  $\sum_{i=1}^m p_i + \sum_{i=1}^m q_i = 1$ . Let  $\mathbf{p} = (p_1, p_2, \dots, p_m)^T$ ,  $\mathbf{q} = (q_1, q_2, \dots, q_m)^T$  and  $\mathbf{0} =$

$(0, 0, \dots, 0)^T$ . Let  $S_0 = \mathbf{0}$  and  $S_n = \sum_{i=1}^n X_i$ . Then the **check english** triplet

$(\{S_n\}_{n=0}^{+\infty}, \mathbf{p}, \mathbf{q})$  is called Type I random walk in  $\mathbb{Z}^m$ .

If  $\forall i \in \{1, 2, \dots, m\} : p_i = q_i = \frac{1}{2m}$  we call the **check english** element  $\{S_n\}_{n=0}^{+\infty}$  Symmetric type I random walk  $\mathbb{Z}^m$ .

**Definition 12** (Type II random walk in  $\mathbb{Z}^m$ )

Let  $m \in \mathbb{N}, \forall n \in \mathbb{N} X_n = (x_{1,n} \ x_{2,n} \ \dots, x_{m,n})^T$ , where  $\forall i \in \{1, 2, \dots, m\} x_{i,n}$  has values in  $\{-1, +1\}$  with probabilities  $P(x_{i,n} = +1) = p_i \in (0, 1)$  and  $P(x_{i,n} = -1) = q_i = 1 - p_i \in (0, 1)$ . Let  $\{X_n\}_{n=0}^{+\infty}$  be a sequence of independent and identically distributed random variables. Let  $S_0 = \mathbf{0}$  and  $\forall n \in \mathbb{N} : S_n = \sum_{i=1}^n X_i$ . Then the

pair  $(\{S_n\}_{n=0}^{+\infty}, \mathbf{p})$  is called Type II random walk in  $\mathbb{Z}^m$ .

If  $\forall i \in \{1, 2, \dots, m\} : p_i = q_i = \frac{1}{2}$  we call the **check english** element  $\{S_n\}_{n=0}^{+\infty}$  Symmetric type II random walk  $\mathbb{Z}^m$ .