

Draft: Quantile estimation with using auxiliary information

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1 Motivation

Suppose we have a population, we can always observe the information X . But we can only observe the sample of Y from the population. Our target is to estimate the population quantile of Y .

2 Method

2.1 Direct Method

A very naive method is to use the sample of Y . Define

$$F_n(y) = \frac{\sum_{i \in A} d_i I(y_i \leq y)}{\sum_{i \in A} d_i}$$

. Then we can get the estimated quantile by

$$\hat{\theta}_d = \inf \{y : F_n(y) \geq \tau\}$$

2.2 Regression estimator and difference estimate under pseudo empirical log-likelihood function

We put these two estimators together, because they are asymptotically same. The theorem 1 in the paper shows that.

$$\hat{\theta}_{df} = \hat{\theta}_d + N^{-1} \left\{ \sum_{i=1}^N q(\mathbf{x}_i; \hat{\beta}_{\tau_0}) - \sum_{i \in A} d_i q(\mathbf{x}_i; \hat{\beta}_{\tau_0}) \right\}$$

For this estimator, we can extend to multi-calibrations. We can use $\tau_0, \tau_1, \dots, \tau_m$ to do the calibration together.

2.3 Using empirical distribution to do calibration

Define

$$F_N(x) = \frac{\sum_i^N I(x_i \leq x)}{N}$$

$$F_w(x) = \sum_{i \in A} w_i I(x_i \leq x)$$

We want to minimize

$$\arg \min_w Var(\hat{\theta}_w) \cong \left(\frac{1}{f(\hat{\theta}_\tau)} \right)^2 Var(\hat{F}_w(\theta)) \quad (1)$$

Given constraints:

$$F_N(x) = F_w(x)$$

$$\sum_{i \in A} w_i = 1$$

But unfortunately, this constraint is not always holds for any $x \in \mathbb{R}$. So I just use the moments to do the calibration. Because the theorem below:

Theorem 1 (*Frechet-Shohat Therom*) Suppose $X_n, n \geq 0$ are random variables. If $\lim EX_n^r = \beta_r$ for all r and if all β_r are the moments of a unique random variable X_0 , then $X_n \rightarrow X_0$ in distribution.

So we can use the first four moments to do the calibration. But I don't get the minimized weights. I first to do the simulation to see if this works. Then we can minimize the w to see we can get the improvement.

2.4 Ratio estimator

For any quantile τ , we can get the true population quantile for X . So the ratio estimator is possible. Suppose the $\theta_{x,\tau}$ is the quantile of τ for population X .

The ratio estimator is

$$\hat{\theta}_{ratio} = \theta_{x,\tau} \frac{\hat{\theta}_{w,y}}{\hat{\theta}_{w,x}}$$

But in some cases if the $\hat{\theta}_{w,x}$ is 0 or closed to 0. That may cause problems. We may use the bias-corrected ratio estimator.

$$\hat{\theta}_{ratio} = \theta_{x,\tau} \frac{\hat{\theta}_{w,y} \hat{\theta}_{w,x} + \hat{C}(\hat{\theta}_{w,y}, \hat{\theta}_{w,x})}{\hat{\theta}_{w,x}^2 + \hat{V}(\hat{\theta}_{w,x})}$$

The problem is that can we use

$$\hat{C}(\hat{\theta}_{w,y}, \hat{\theta}_{w,x}) = ??$$

and

$$\hat{V}(\hat{\theta}_{w,x}) \cong \left(\frac{1}{f(\hat{\theta}_\tau)} \right)^2 \hat{Var}(\hat{F}_w(\theta))$$

3 A small simulation study

The simulation studies are conducted to compare the performance of different estimators. Two finite populations of size $N = 1000$ were generated from bivariate normal distribution respectively. The correlation between two variables in the first population is 0.9, then 0.6. For each simulation run, a simple random sample of size $n = 100$ was taken.

To compare with each other, we set the direct estimator as the base and define relative efficiency:

$$RE_* = \frac{MSE_*}{MSE_{Direct}}$$

Note: We use the same $\tau_0 = 0.5$ to get the difference estimator.

	$\tau = 0.1$	$\tau = 0.3$	$\tau = 0.5$	$\tau = 0.7$	$\tau = 0.9$
Difference estimator RE	0.8380	0.7203	0.7767	0.8469	0.8599
ratio estimator RE	1.2040	2.2434	7.7272	2.1759	1.9842
Moment estimator RE	0.8392	0.7665	0.8253	0.8852	0.9038

Table 1: RE for different estimators for $\rho = 0.6$

	$\tau = 0.1$	$\tau = 0.3$	$\tau = 0.5$	$\tau = 0.7$	$\tau = 0.9$
Difference estimator RE	0.7419	0.5571	0.4283	0.5278	0.7796
ratio estimator RE	0.5734	0.5104	15.6849	0.8469	0.8919
Moment estimator RE	0.7390	0.5520	0.4329	0.5392	0.7891

Table 2: RE for different estimators for $\rho = 0.9$

- Both difference estimator and moment estimator do better than direct estimator.
- Ratio estimator is worse than estimator than the direct estimator. Especially at $\tau = 0.5$. That is because the quantile at $\tau = 0.5$ is very closed to 0. So bias-corrected ratio estimator may have a better estimation.
- Difference estimator is best in these estimators.

4 Variance Estimation for difference estimator

In the paper, we have two method to get the variance estimator for quantile estimator.

For method one: I correct two mistakes in the paper. I get the

	source	$\tau = 0.1$	$\tau = 0.3$	$\tau = 0.5$	$\tau = 0.7$	$\tau = 0.9$
$\frac{E(\hat{V})}{V} - 1$	$\rho = 0.6$	5.84%	21.51%	34.79824%	11.16%	10.86%
$\frac{E(\hat{V})}{V} - 1$	$\rho = 0.9$	-11.51%	-8.18%	24.20%	23.66%	0.2718%

Table 3: Relative bias for variance estimations

For method two:

	source	$\tau = 0.1$	$\tau = 0.3$	$\tau = 0.5$	$\tau = 0.7$	$\tau = 0.9$
$\frac{E(\hat{V})}{V} - 1$	$\rho = 0.6$	344.69%	320.10%	322.11%	303.28%	495.60%
$\frac{E(\hat{V})}{V} - 1$	$\rho = 0.9$	355.76%	283.12%	315.44%	329.97%	388.80%

Table 4: Relative bias for variance estimations from Woodruff method