Advanced Econometrics 4EK608

Vysoká škola ekonomická v Praze

#### Outline

- Stochastic processes
- 2 Unit roots
- Cointegration
- Error correction models

# Czech terminology

Testy jednotkových kořenů, silně závislé časové řady, zdánlivá regrese, kointegrace, model korekce chyby, MA proces, AR proces, náhodná procházka s driftem, integrovaná časová řada, řád integrace, korelogram, ADF test, stochatický trend, deterministický trend, trendově stacionární časová řada,

#### Weakly dependent time series

- Moving average process of order one MA(1)  $x_t = e_t + \alpha_1 e_{t-1}$ , where  $e_t$  is *i.i.d.* time series. Observations with higher time distance than 1 are uncorrelated. This process is stationary.
- Autoregressive process of order 1: AR(1)  $y_t = \rho_1 y_{t-1} + e_t \Rightarrow Corr(y_t, y_{t+h}) = \rho_1^h$

If stability condition  $|\rho| < 1$  holds, the process is weakly dependent because correlation converges to zero with growing h. Also, this process is stationary for  $y_0 = 0$ .

Random walk:

$$y_{t} = y_{t-1} + e_{t}$$

$$y_{t} = y_{t-2} + e_{t-1} + e_{t}$$

$$y_{t} = y_{t-3} + e_{t-2} + e_{t-1} + e_{t}$$

$$\dots$$

$$y_{t} = y_{0} + e_{1} + \dots + e_{t-1} + e_{t}$$

Shocks have permanent effects, the series is not covariance stationary and is strongly dependent.

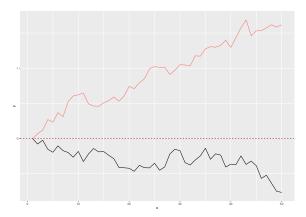
$$E(y_t) = E(y_0)$$

$$Var(y_t) = \sigma_e^2 t$$

$$Corr(y_t, y_{t+h}) = \sqrt{t/(t+h)}$$

Correlation decreases very slowly and speed depends on t.

• Two realizations of a random walk



• Random walk with a drift

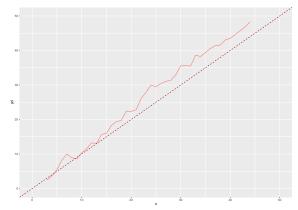
$$y_t = \alpha_0 + y_{t-1} + e_t \Rightarrow y_t = \alpha_0 t + e_t + e_{t-1} + \dots + e_1 + y_0$$

A linear trend with random walk around the trend. It is neither covariance stationary nor weakly dependent.

$$E(y_t) = \alpha_0 t + E(y_0)$$
$$Var(y_t) = \sigma_e^2 t$$
$$Corr(y_t, y_{t+h}) = \sqrt{t/(t+h)}$$

Correlation decreases very slowly and decline speed depends on t.

• Realization of random walk with a drift



• Different realizations of trending TS may produce similar time series.

$$y_t = |1 \cdot y_{t-1}| + |u_t| = y_{t-1} + u_t$$

- Unit root process:  $y_t = y_{t-1} + u_t$ ;  $u_t$  is a weakly dependent series.
- Random walk is a special case of the unit root process where:  $u_t \sim Distr(0, \sigma_u^2)$ , iid

We need to distinguish strongly and weakly dependent TS:

- Economic reasons:
   In strongly dependent series, shocks or policy changes have long or permanent effects; in weakly dependent series, their effect is only temporary.
- Statistical reasons:
   Analysis with strongly dependent series must be handled in specific ways.

### Integrated series

#### Terminology - Order of integration

- Weakly dependent TS are integrated of order zero: I(0).
- If we have to difference a TS once to get a weakly dependent TS, then it is integrated of order 1: I(1).
- Example of a I(1) process:

$$y_t = y_{t-1} + e_t \quad \Rightarrow \Delta y_t = y_t - y_{t-1} = e_t$$
$$\log y_t = \log y_{t-1} + e_t \Rightarrow \Delta \log y_t = e_t$$

• A time series is integrated of order d: I(d), if it becomes a weakly dependent TS after being differenced d times.

Unit root tests help to decide if a time series is I(0) or not

- Use either some informal procedure or a unit root test
- Informal procedures
  - Analyze autocorrelation of the first order

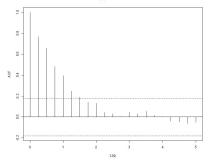
$$\hat{\rho}_1 = \hat{Corr}(y_t, y_{t-1})$$

- If  $\hat{\rho}_1$  approaches 1, it indicates that the series can have unit root. Alternatively, it could have a deterministic trend.
- We can analyze sample autocorrelations using a correlogram

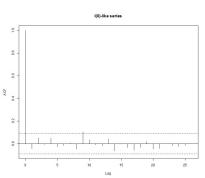
Correlogram:

$$\rho_h = \frac{cov(y_t, y_{t-h})}{\sigma_{y_t} \cdot \sigma_{y_{t-h}}}$$

I(1)-like series



I(1)-like series



I(0)-like series

#### Dickey-Fuller (DF) test – motivation

Unit root test in an ar(1) process:

$$y_t = \alpha + \rho y_{t-1} + e_t$$

$$H_0: \rho = 1, \ H_1: \rho < 1$$

- Under  $H_0$ ,  $y_t$  has a unit root.
  - For  $\rho = 1 \land \alpha = 0 \rightarrow y_t$  is a random walk.
  - For  $\rho = 1 \land \alpha \neq 0 \rightarrow y_t$  is a randomw walk with a drift and  $E(y_t)$  is a linear function of t.
- Under  $H_1$ ,  $y_t$  is a weakly dependent ar(1) process.

#### Dickey-Fuller (DF) test – motivation

Unit root test in an ar(1) process:

$$y_t = \alpha + \rho y_{t-1} + e_t$$

$$H_0: \rho = 1, \ H_1: \rho < 1$$

For DF tests,  $H_1: \rho < 1$  is a common simplification to the full space of alternatives to  $H_0: \rho = 1$ .

- For  $|\rho| < 1$ ,  $y_t$  is weakly dependent (as  $plim \, \rho^h = 0$ ) However, if unit root is likely to be present, the probability of  $\rho < 0$  is negligible.
- We usually ignore the possibility of  $\rho > 1$ , as it would lead to explosive behavior in  $y_t$ .
  - $\ldots |\rho| > 1$  would allow for explosive oscillations in  $y_t$ .

### Dickey Fuller (DF) test

- Basic equation for unit root test in an ar(1) process:  $y_t = \alpha + \rho y_{t-1} + e_t$
- For DF test, we apply a suitable transformation to  $y_t$ : we subtract  $y_{t-1}$  from both sides of the equation:

$$\Delta y_t = \alpha + (\rho - 1)y_{t-1} + e_t; \text{ apply substitution: } \theta = (\rho - 1)$$
 i.e. 
$$H_0 : \rho = 1 \Leftrightarrow H_0 : \theta = 0$$
$$\Delta y_t = \alpha + \theta y_{t-1} + e_t; \text{ now: } H_1 : \rho < 1 \Leftrightarrow H_1 : \theta < 0$$

- We use a t-ratio for testing  $H_0: \theta = 0$ . However: Under  $H_0$ , t-ratios don't have a t-distribution, but follow a DF-distribution. (-negative- critical values of the DF distribution are much farther from zero)
- Critical values for the DF distribution are available from statistical tables and implemented in most relevant SW packages.

#### DF test & ADF test

Unit root time series can manifest various levels of complexity. Hence, DF test is usually performed using the following three specifications:

$$\Delta y_t = \theta y_{t-1} + e_t$$
 random walk 
$$\Delta y_t = \alpha + \theta y_{t-1} + e_t$$
 random walk with a drift 
$$\Delta y_t = \alpha + \theta y_{t-1} + \delta t + e_t$$
 random walk with a drift and trend

DF test is the same  $(H_0: \theta = 0)$  for all specifications /critical values difffer/

Augmented Dickey-Fuller (ADF) test is a common generalization of DF test (example: Augmentation of the DF test for the  $2^{nd}$  specification)

$$\Delta y_t = \alpha + \left[\theta y_{t-1}\right] + \gamma_1 \Delta y_{t-1} + \dots + \gamma_p \Delta y_{t-p} + e_t$$

- When estimating  $\theta$ , we control for possible ar(p) behavior in  $\Delta y_t$ .
- ADF test has the same null hypothesis as a DF test  $\rightarrow H_0: \theta = 0$ .

# Unit root tests in R: package {urca}

Description of the options for the ur.df() function:

- type "none"  $\Delta y_t = \theta y_{t-1} + e_t$  tau1: we test for  $H_0: \theta = 0$  (unit root)
- ② type "drift"  $\Delta y_t = \alpha + \theta y_{t-1} + e_t$  tau2:  $H_0: \theta = 0$  (unit root) phi1:  $H_0: \theta = \alpha = 0$  (unit root and no drift)
- type "trend"  $\Delta y_t = \alpha + \theta y_{t-1} + \delta t + e_t$ tau3:  $H_0: \theta = 0$  (unit root) phi2:  $H_0: \theta = \alpha = \delta = 0$  (unit root, no drift, no trend) phi3:  $H_0: \theta = \delta = 0$  (unit root and no trend)

• ADF test for TS with trend

$$\Delta y_t = \alpha + \theta y_{t-1} + \delta t + \gamma_1 \Delta y_{t-1} + \dots + \gamma_p \Delta y_{t-p} + e_t$$

Under the alternative hypothesis of no unit root, the process is trend-stationary.

- The critical values in the ADF distribution with time trend are even more negative as compared to random walk and random walk with a drift.
- When using DF/ADF specification 1 or 2 (R-W, R-W with drift) to test for unit root in a clearly trending TS, the test would not have sufficient power (we would not reject  $H_0$  for trending weakly dependent TS).

### Unit roots and trend-stationary series

- $\Delta y_t = \alpha + \theta y_{t-1} + \delta t + \gamma_1 \Delta y_{t-1} + \cdots + \gamma_n \Delta y_{t-n} + e_t$
- Terminology:
  - Stochastic trend:  $\theta = 0$ Also called **difference-stationary process**:  $y_t$  can be turned into I(0) series by differencing. Terminology emphasizes stationarity after differencing  $y_t$  instead of weak dependence in differenced TS.
  - Deterministic trend:  $\delta \neq 0$ ,  $\theta < 0$ Also called **trend-stationary process**: has a linear trend, not a unit root.  $y_t$  is weakly dependent - I(0) - around its trend. We can use such series in LRMs, if trend is also used as regressor.
- DF/ADF tests are not precise tools. Distinguishing between stochastic and deterministic trend is not easy (sample size!).

# Handling trend-stationary time series

• Trend-stationary TS fulfill TS.1' assumption (look at Week1 presentation).

We can use them in regressions if we have time trend among regressors.

### Handling strongly dependent time series

- Strongly dependent time series do not fulfill TS.1' assumption (look at Week1 presentation). We cannot use them in regressions directly.
- Sometimes, taking logarithms helps.
- Sometimes, we can transform such series into weakly dependent time series.
- Differencing is popular, but it has drawbacks.

### Handling strongly dependent time series

#### Example

$$y_{t} = \beta_{0} + \beta_{1}x_{t} + \varepsilon_{t}$$
 
$$y_{t}, x_{t} \sim I(1)$$
 
$$y_{t-1} = \beta_{0} + \beta_{1}x_{t-1} + \varepsilon_{t-1}$$
 
$$\varepsilon_{t} \sim i.i.d.$$
 
$$(2)$$
 
$$\Delta y_{t} = \beta_{1}\Delta x_{t} + v_{t}$$
 
$$v_{t} = \varepsilon_{t} - \varepsilon_{t-1}$$
 
$$(3)$$

- ① If we work with logarithms, it has an additional advantage:  $\log y_t \log y_{t-1} = \log \left(\frac{y_t}{y_{t-1}}\right) \doteq \frac{y_t y_{t-1}}{y_{t-1}}$  i.e.: nice interpretation as rate of growth of  $y_t$
- 2 Three problems
  - $\mathbf{0}$   $v_t$  is no more i.i.d.
  - We loose information linked with the levels of variables, short term relation are stressed
  - **3** Estimates often generate bad long-term predictions:  $\Delta \hat{y}_t = \hat{\beta}_1 \Delta x_t$ ; ... what if  $\beta_0 \neq 0$ ?

### Handling strongly dependent time series

#### Some properties of integrated processes

- The sum of stationary and non-stationary series must be non-stationary.
- 2 Consider a process  $y_t = \alpha + \beta x_t$ :
  - · If  $x_t$  is stationary then  $y_t$  will be stationary.
  - If  $x_t$  is non-stationary then  $y_t$  will be non-stationary.
- 3 If two time series are integrated of different orders, then any linear combination of the series will be integrated at the higher of the two orders of integration.
- Sometimes it turns out a linear combination of two I(d)series is integrated of order less then d.

# Spurious regression or cointegration

- Spurious regression Regressing one I(1)-series on another I(1)-series may lead to extremely high t-statistics even if the series are completely independent. Similarly, the  $R^2$  of such regressions tend to be very high. Regression analysis involving time series that have a unit root may generate completely misleading inferences.
- Cointegration Fortunately, regressions with I(1)-variables are not always spurious: If there is a stable relationship between time series that, individually, display unit root behavior, these time series are called "cointegrated".

# Spurious regression or cointegration

#### General definition of cointegration

Two I(1)-time series  $y_t$ ,  $x_t$  are said to be cointegrated if there exists a stable relationship between them, where:

$$y_t = \alpha + \beta x_t + e_t, \quad e_t \sim I(0)$$

#### Cointegration (CI) test if CI parameters are known

For residuals of the known CI relationship:

$$e_t := y_t - \alpha - \beta x_t,$$

test whether the residuals have a unit root (DF/ADF and other unit root tests may be applied "directly").

If the unit root  $H_0$  is rejected,  $y_t$ ,  $x_t$  are cointegrated.

# Spurious regression or cointegration

- Testing for CI if the parameters are unknown If the potential relationship is unknown, it can be estimated by OLS. After that, we test whether the regression residuals have a unit root. If the unit root is rejected, this means that  $y_t$ ,  $x_t$  are cointegrated. Due to the pre-estimation of parameters, critical values are different than in the case of known parameters. (Software handles this automatically.)
- The CI relationship may include a time trend If the two series have differential time trends (drifts in this case), the deviation between them may still be I(0) but with a linear time trend. In this case one should include a time trend in the CI-regression. Also, we have to use different critical values when testing residuals. (Software handles this automatically.)

# Cointegration tests based on regression residuals

**Engle-Granger test** estimates a *p*-lag ADF equation:

$$\Delta \hat{u}_t = \theta \, \hat{u}_{t-1} + \sum_{j=1}^p \Delta \hat{u}_{t-j} + e_t$$

- Esentially, this is an ADF test on  $\hat{u}_t$  [ $\theta = (\rho 1)$ ]
- Specific critical values apply (farther from 0 than t or DF).

Phillips-Ouliaris test estimates a DF equation:

$$\Delta \hat{u}_t = \theta \, \hat{u}_{t-1} + e_t$$

• The t-ratio is based on robust standard errors, different estimator exist for the robust standard errors.

In both cases (EG and PO),  $H_0$  of unit root in  $\hat{u}$  i.e. "no-cointegration" is tested.

# Error correction model (ECM)

• It can be shown that when variables are cointegrated, i.e. when there exists a long-term relationship among them, their short-term dynamics are related as in a so-called error correction model (ECM).

#### Error Correction Models - motivation

#### Autoregressive distributed lag models

Autoregressive distributed lag model with one regressor

ADL
$$(p,q)$$
:  $y_t = \beta_0 + \sum_{i=1}^p \beta_i y_{t-i} + \sum_{j=0}^q \gamma_j x_{t-j} + u_t$ ,  $u_t \sim iid(0, \sigma^2)$ 

• There are many useful modifications/simplifications to the ADL(p,q) process. For example:

ADL(1,1): 
$$y_t = \beta_0 + \beta_1 y_{t-1} + \gamma_0 x_t + \gamma_1 x_{t-1} + u_t$$
. (4)

Additional ADL(1,1) restriction:  $\beta_1 = 1$  and  $\gamma_1 = -\gamma_0$ 

gives a model in 1<sup>st</sup> diffs.:  $\Delta y_t = \beta_0 + \gamma_0 \Delta x_t + u_t$ .

For ADL(1,1) model (4), suppose there is an equilibrium value  $x^{\circ}$  and in the absence of shocks,  $x_t \to x^{\circ}$  as  $t \to \infty$ . Then, assuming absence of  $u_t$  errors,  $y_t$  converges to steady state:  $y^{\circ}$ .

Hence, the ADL(1,1) model (4) can be re-written as:

$$y^{\circ} = \beta_0 + \beta_1 y^{\circ} + (\gamma_0 + \gamma_1) x^{\circ}$$

Solving this for  $y^{\circ}$  as a function of  $x^{\circ}$ , we get

$$y^{\circ} = \frac{\beta_0}{1 - \beta_1} + \frac{\gamma_0 + \gamma_1}{1 - \beta_1} x^{\circ} = \frac{\beta_0}{1 - \beta_1} + \lambda x^{\circ}$$

where  $\lambda \equiv \frac{\gamma_0 + \gamma_1}{1 - \beta_1}$  and  $|\beta_1| < 1$  is assumed.

#### Error Correction Models - motivation

$$y^{\circ} = \frac{\beta_0}{1 - \beta_1} + \lambda x^{\circ}$$
$$\lambda \equiv \frac{\gamma_0 + \gamma_1}{1 - \beta_1}$$

- $\lambda$  is the long-run derivative of  $y^{\circ}$  with respect to  $x^{\circ}$ .
- $\lambda$  is an elasticity if both  $y^{\circ}$  and  $x^{\circ}$  are in logs.
- $\hat{\lambda}$  can be computed directly from the estimated parameters of the ADL(1,1) model (4).

The ADL(1,1) equation (4) - repeated here for convenience:

$$y_t = \beta_0 + \beta_1 y_{t-1} + \gamma_0 x_t + \gamma_1 x_{t-1} + u_t,$$

can be equivalently rewritten as follows:

$$\Delta y_t = \beta_0 + (\beta_1 - 1)(y_{t-1} - \lambda x_{t-1}) + \gamma_0 \Delta x_t + u_t.$$
 (5)

Again,  $\lambda \equiv \frac{\gamma_0 + \gamma_1}{1 - \beta_1}$  and  $|\beta_1| < 1$  is assumed.

Equation (5) is an error-correction model (ECM).

### Error Correction Models

ECM: 
$$\Delta y_t = \beta_0 + (\beta_1 - 1)(y_{t-1} - \lambda x_{t-1}) + \gamma_0 \Delta x_t + u_t$$
.

- $(y_{t-1} \lambda x_{t-1})$  measures the extent to which the long run equilibrium between  $y_t$  and  $x_t$  is not satisfied (at t-1).
- Consequently,  $(\beta_1 1)$  can be interpreted as the proportion of the disequilibrium  $(y_{t-1} - \lambda x_{t-1})$  that is reflected in the movement of  $y_t$ , i.e. in  $\Delta y_t$ .
- $(\beta_1 1)(y_{t-1} \lambda x_{t-1})$  is the error-correction term.
- Many ADL(p,q) specifications can be re-written as ECMs.
- ECMs can be used with non stationary TS (Week 3).
- ECMs  $(\beta_1 1)$  is essentially the same as  $\theta$  from Partial adjustment model (see Week 4).

# Box 1: Partial Adjustment Models (Will be put into context in the 4th week)

$$Y_t^* = \alpha + \beta X_t + u_t$$
  $Y_t^*$ : optimal value or target value (6) or long-run equilibrium value

$$Y_t - Y_{t-1} = \theta(Y_t^* - Y_{t-1}) \quad 0 < \theta < 1$$

$$\theta \sim \text{coefficient of adjustment}$$

$$Y_t = \theta Y_t^* + (1 - \theta) Y_{t-1}$$

$$(7)$$

As  $\theta \to 1$ , the speed of contemporaneous adjustment of  $Y_t$  towards  $Y_t^*$  grows.

Substituting  $Y_t^*$  from (6) to (7) yields

$$Y_t = \alpha \theta + \beta \theta X_t + (1 - \theta) Y_{t-1} + \theta u_t \tag{8}$$

We estimate (8) and then calculate parameters in (6) and (7).

#### Error Correction Models

#### Some more complicated ECMs:

1) We can use higher order lags, e.g. ADL(2,2):

$$y_t = \beta_0 + \beta_1 y_{t-1} + \beta_2 y_{t-2} + \gamma_0 x_t + \gamma_1 x_{t-1} + \gamma_2 x_{t-2} + u_t,$$

to establish ECMs. It is again possible to rearrange and re-parametrize  $\mathrm{ADL}(2,2)$  to get an ECM. More than one re-parameterization is possible.

2) More than two variables can enter into an equilibrium relationship.