Treatment Effect Analysis

Questions

Assumptions/Preconditions:

- Conditional Independence Assumptions (CIA)
 - $(BTUgas_i(0), BTUgas_i(1)) \perp Post_i | X_i$

1. Treatment Effect

- 1. Describe in words the potential outcomes
 - The energy unit BTU change between treatment $(Post_i)$ of year discrimination at a random assignment (\bot) in the independence of the other confounding variables X_i
- 2. Describe in words and mathematically the unit-level treatment Effect
 - Unit level treatment/ITE is defined by the difference in the treated and untreated (potential outcomes)
 - Mathematically: $ITE_i = Y_i(1) Y_i(0)$

2. CIA Assumptions

- 1. Explain why $E[BTUgas_i|Post_i = 1] E[BTUgas_i|Post_i = 0]$ is unlikely to identify ATE
 - During mean difference one could say omitted variable bias exists within the regression proven by:

$$\frac{Cov(BTUgas_i, Post_i)}{Var(Post_i)} = \frac{Cov(\alpha + \beta Post_i + \gamma X_i + \epsilon)}{Var(D_i)} = \beta + \frac{\gamma Cov(Post_i, X_i)}{Var(Post_i)}$$

- For true regression will always include X_i along with the D_i treatment
- 2. Describe in words the CIA Assumptions $(\perp Post_i|X_i)$
 - Only treatments are influenced through random assignment/sampling
 - No confounding variables should influence the outcome
- 3. Explain why the conditional independence assumption is more plausible than the unconditional independence assumption
 - Some observable factors are controlled to not influence the outcome in CIA
 - IA does not protect against confounding variable bias
 - One possible X_i could be: habitant eating habit (e.g., if eat late: cook more, use dishwasher more, use more lighting)

3. COC Assumptions

- 1. Express the COC assumption mathematically and in words
 - $0 < P(Post_i = 1 | X_i = x) < 1 \quad \forall x$
 - For all possible observations of x, there is covariation of treated and untreated, making comparison possible; no complete independence of treatment to control
- 2. Why COC may fail at high-dimensional X_i
 - COC may fail due to the treatment and control units may not bound to all the pretreatment variables
 - $Pr(D_i|X_i = x) = 1 \text{ or } 0$

```
df$Post_i <- ifelse(df$constYRS == 7, 1, 0)

# ATE Calculated by Difference in Means
treated_mean <- mean(df$BTUgas[df$Post_i == 1], na.rm = TRUE)
control_mean <- mean(df$BTUgas[df$Post_i == 0], na.rm = TRUE)
cat("Pre-code reform energy usage", control_mean, "\n")</pre>
```

Pre-code reform energy usage 49.80924

```
ATE_dim <- treated_mean - control_mean
cat('ATE (Difference in Means):', ATE_dim, '\n')</pre>
```

ATE (Difference in Means): -2.651527

ATE (Regression): -2.532691

```
for (k_index in seq_along(K_values)) {
    K <- K_values[k_index]</pre>
    Y match <- numeric(nrow(X std))</pre>
    for (i in 1:nrow(X_std)) {
        dist_i <- rowSums((X_std - X_std[i, ])^2)</pre>
        dist_i[D == D[i]] <- Inf</pre>
        neighbors <- order(dist i)[1:K]</pre>
        Y_match[i] <- mean(Y[neighbors])</pre>
    }
    ATE_results[k_index] <- mean(Y_match[D == 1]) - control_mean
print(data.frame(K = K_values, ATE = ATE_results))
##
      K
               ATE
## 1 1 -7.755631
## 2 5 -7.101051
## 3 10 -5.320479
# Reload to fix df being empty
df0 <- read_dta("table3-cols12.dta")</pre>
df <- df0[df0$constYRS==6| df0$constYRS ==7, ]</pre>
df <- df[!is.na(df$lnBTUgas), ]</pre>
df$BTUgas <- exp(df$lnBTUgas)</pre>
df$Post_i <- ifelse(df$constYRS == 7, 1, 0)</pre>
X <- c("cddzip30_00", "hddzip30_00", "lnsqftK", "numroom", "elecboth",
       "remodeled", "lnyrs_res", "lnrescnt", "lnhhinc", "nr0_5",
       "nr65_99", "college", "anydisabled", "hohblk1", "hohlat1",
       "own", "year09", "cecfast")
# Prepare Propensity Score
prop_model <- glm(as.formula(paste("Post_i", "~", paste(X, collapse = " + "))),</pre>
                  data = df, family = binomial)
df$prop_score <- plogis(fitted.values(prop_model))</pre>
print(range(df$prop_score))
## [1] 0.5741077 0.7023633
# Perform KNN with distance based on abs(p(x_i) - p(x_j))
K_{values} \leftarrow c(1, 5, 10)
ATE_results <- data.frame(K = K_values, ATE = NA)
ATE_results <- numeric(length(K_values))</pre>
for (k_index in seq_along(K_values)) {
    K <- K_values[k_index]</pre>
    Y_match <- numeric(nrow(X_std))</pre>
    for (i in 1:nrow(X std)) {
        dist_i <- abs(df$prop_score - df$prop_score[i])</pre>
```

```
dist_i[D == D[i]] <- Inf</pre>
        neighbors <- order(dist_i)[1:K]</pre>
        Y_match[i] <- mean(Y[neighbors])</pre>
    }
    ATE_results[k_index] <- mean(Y_match[D == 1]) - control_mean
}
print(data.frame(K = K_values, ATE = ATE_results))
##
      K
               ATE
## 1 1 -0.1695891
## 2 5 -0.9497518
## 3 10 -0.4213290
# ATE by Inverse Probability Weighting
df$weights_treated <- ifelse(df$Post_i == 1, 1 / df$prop_score, 0)</pre>
df$weights_control <- ifelse(df$Post_i == 0, 1 / (1 - df$prop_score), 0)</pre>
weighted_outcome_treated <- sum(df$Post_i * df$BTUgas * df$weights_treated) /</pre>
                           sum(df$weights treated)
weighted_outcome_control <- sum((1 - df$Post_i) * df$BTUgas * df$weights_control) /</pre>
                           sum(df$weights_control)
ATE_IPW <- weighted_outcome_treated - weighted_outcome_control
cat("The estimated ATE using IPW is:", ATE_IPW)
```

The estimated ATE using IPW is: -2.628553

Discussion

- 1. Outside of KNN and propensity scored KNN, the ATE estimates are fairly consistent at 5 percent in negative direction (e.g., -2.5 to -2.6 out of 49.08). We see a sharp decrease from KNN result at -7/49, with some decrease as we increased the K for the classification. Using the propensity score for the distance calculated KNN, we have a marginal difference at fraction of decreased units.
- 2. The KNN implementation is harder than expected. I believe the simplest may be Regression, IPW and the Mean Difference. Writing algorithms gets self-doubt. One might need to try different methods of ATE estimation to see their consistent outcome. I am not sure why the propensity-KNN is returning a marginal difference.