

COMP2043.GRP INTERIM GROUP REPORT

SEQUENTIAL MONTE CARLO

TEAM 6

2017.12.7

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1 Introduction

1.1 Background of the Project

Sequential Monte Carlo (SMC, also called particle filter), is a widely-used statistical sampling method, which is used to sequentially sample from a sequence of target densities. It deals with the problem of recursively estimating the probability density function $p(x_t|Y_s)$.

Before the short journey of Introduction to Particle filter, a brief Background of particle filter will be explained in advanced. In our present world, the ability of solve the problem of estimating various quantities in nonlinear dynamic systems is of paramount importance in many practical applications. In order to understand how a system, for instance, a aircraft, a car, or a camera performs, we need to have access to certain important quantities related with the system. However, typically we have to estimate them based on various noisy measurements available from the system due to the fact that we do not have direct access to these factors. According to Thomas B.Schon, the state estimation problem is addressed mainly within a probabilistic framework. More specifically, the approach is heavily affected by the Bayesian view of estimation, which implies that the complete solution to the estimation problem is provided by the probability density function $p(x_t|Y_s)$. This density function contains all available information about the state variable.

1.2 Motivation & Wide implementation of state estimation

The automotive industry's focus is recently shifting from mechanics to electronics and software. This opens up factors a great deal of interesting applications and research opportunities within the field of estimation theory.

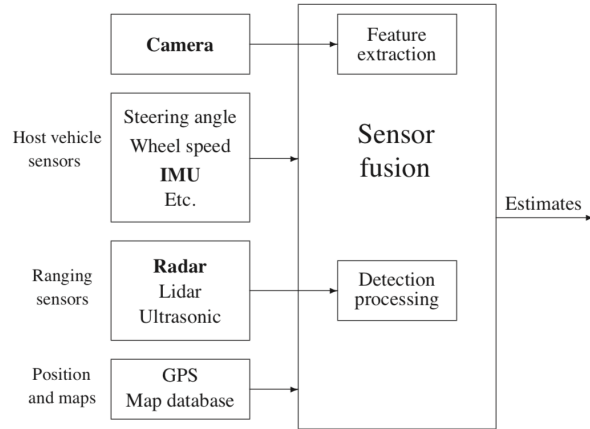


Figure 1: The most important factors enabling future automotive safety systems is the availability of accurate information about the state. The process of obtaining this information is to a large extent dependent on a unified treatment of the sensor information, as illustrated in this figure.

1.2.1 Automotive Navigation-Example

The objective of this study is to calculate estimates of the road geometry, which are important in several advanced control systems such as lane guidance and collision avoidance. The idea exemplified here follows from the general framework introduced in Figure 1.



Figure 2: when entering a curve, all vehicles start moving in the lateral direction. This information can be used to support the road geometry estimate.

Here information from several sensors is used to obtain better performance, than separate use of the sensors would allow for. The main assumption is that the leading vehicles will keep following their lane, and their lateral movement can thus be used to support the otherwise difficult process of road geometry estimation. For instance, when entering a curve as in Figure 2 the vehicles ahead will start moving to the right and thus there is a high probability that the road is turning to the right. Assuming that the surrounding vehicles will keep following the same lane, is in discrete-time expressed as $y_t + 1^i = y_t^i + w_t$, $w_t \sim \mathcal{N}(0, Q_{lat})$. Here, y^i denotes the lateral position of vehicle i and w_t denotes Gaussian white noise which is used to account for model uncertainties. The estimate of the road curvature during an exit phase of a curve is illustrated in Figure 3. The true reference signal was generated using the method proposed by Eidehall and Gustafsson (2006).

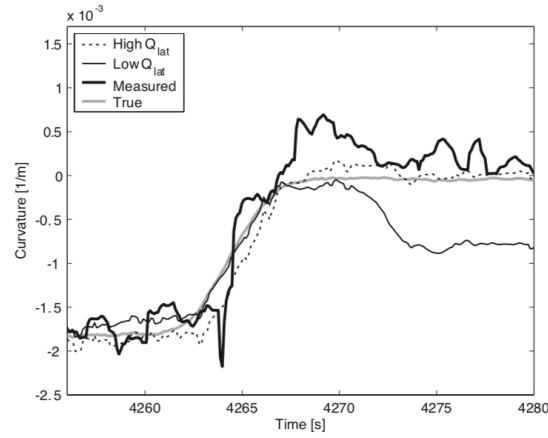


Figure 3: Comparison of estimation performance from two filters, one with a large Q_{lat} and one with a small Q_{lat} . The raw measurement signal from the image processing unit is also included. Comparing this raw vision measurement to the result from the filters clearly illustrates the power of a model based sensor fusion approach

1.2.2 Navigation for Augmented Reality



Figure 4: Some examples illustrating the concept of augmented reality.

1.3 Scope of the project

This subsection describes the scenario where the product will be used for.

1.4 Introducing two kinds of auxiliary particle filters(APF)

1.4.1 Sequential Monte Carlo Methods

As we have mentioned above that Sequential Monte Carlo methods, or particle methods, provide the solution for state estimation. Then, what does it constitute? It is constituted by a combination of sequential importance sampling and resampling. The key idea underlying the particle methods is to represent the probability density function by a set of samples(also referred to as particles, hence the name particle methods) and its associated weights. The density function $p(x_t|Y_s)$ is approximated with an empirical density function,

$$p(x_t|Y_s) \approx \sum_{i=1}^M \tilde{q}_t^{(i)} \delta(x_t - x_{t|s}^{(i)}), \quad \sum_{i=1}^M \tilde{q}_t^{(i)} = 1, \quad \tilde{q}_t^{(i)} \geq 0, \forall i \quad (1)$$

where $\delta(\cdot)$ represents the Dirac delta function and $\tilde{q}_t^{(i)}$ denotes the weight associated with particle $x_{t|s}^{(i)}$. For intuition we can think of each particle x_t^i as a possible system state and the corresponding weight $\tilde{q}_t^{(i)}$ contains information about how probable that particular state is.

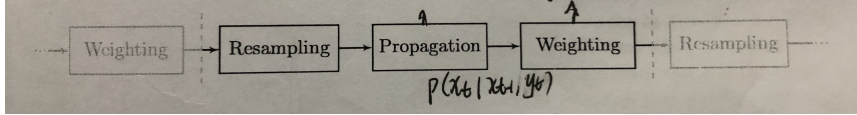


Figure 5: Illustrating the three parts making up sequential Monte Carlo

1.4.2 Auxiliary particle filter(APF)

As a principled solution to compute the filtering PDFs $(x_t|y_{1:t})$ sequentially in time is provided by the following two recursively equations:

$$p(x_t|y_{1:t}) = p(x_t|y_t, y_{1:t-1}) = \frac{p(y_t|x_t)p(x_t|y_{1:t-1})}{p(y_t|y_{1:t-1})} \quad (2)$$

$$p(x_t|y_{1:t-1}) = \int p(x_t|x_{t-1})p(x_{t-1}|y_{1:t-1})d_{x_{t-1}} \quad (3)$$

These two equations are coming from Forward filtering. The explanation will be included in the Appendixes in detail. Further, the main barrier here is the integral in equation (3), which is in general not analytically tractable.

Nevertheless, the integral can be approximated using an importance sampler targeting the filtering distribution at time $t - 1$. This give us the impetus to proceed in an inductive fashion. Thus, assuming we have an empirical approximation of the filtering distribution at time $t-1$, constituted by N weighted samples, $\{x_{t-1}^i, w_{t-1}^i\}_{i=1}^N$, i.e.

$$\hat{p}^N(x_{t-1}|y_{1:t-1}) = \sum_{i=1}^N w_{t-1}^i \delta_{x_{t-1}^i}(x_{t-1}) \quad (4)$$

At time $t = 1$, it is able to obtain a point-mass approximation according to equation (4), by targeting $p(x_1|y_1) \propto p(y_1|x_1)\mu(x_1)$ with an importance sampler. Inserting the approximation $\hat{p}^N(x_{t-1}|y_{1:t-1})$ into equation (3), result in

$$\hat{p}^N(x_t|y_{1:t-1}) = \int p(x_t|x_{t-1}) \sum_{i=1}^N w_{t-1}^i \delta_{x_{t-1}^i}(x_{t-1}) dx_{t-1} = \sum_{i=1}^N w_{t-1}^i p(x_t|x_{t-1}^i) \quad (5)$$

Using equation (5) and equation (2), we can evaluate an approximation of the filtering PDF $p(x_t|y_{1:t})$ up to proportionality

$$p(x_t|y_{1:t}) \approx \frac{1}{p(y_t|y_{1:t-1})} \sum_{i=1}^N w_{t-1}^i p(y_t|x_t) p(x_t|x_{t-1}^i) \quad (6)$$

As for this opens up for targeting $p(x_t|y_{1:t})$ with an importance sampler. Then, choosing a similar type of mixture as proposal density, namely

$$q(x_t|y_{1:t}) = \sum_{i=1}^N w_{t-1}^i q(x_t|x_{t-1}^i, y_t) \quad (7)$$

There are many different options when it comes to choosing the component $q(x_t|x_{t-1}^i, y_t)$ in this mixture. Note that, in general, the proposal density at time t is allowed to depend on the current observation y_t , as indicated by the notation used in equation (7). Therefore, Using auxiliary variable in the form of a discrete random variable a_t which takes values on the set of integers $\{1, \dots, N\}$ in the $q(x_t|y_{1:t})$ can take y_t into account, which make use of this information already when simulating the particles $\{x_t^i\}_{i=1}^N$, to increase the probability of producing samples in the most relevant parts of the state space. This is why we indicate a possible dependence on y_t in the mixture components of the proposal density in (7). The key in this development is to target the joint distribution of (x_t, a_t) with an importance sampler, instead of directly targeting the marginal distribution of x_t

Analogously to above, the mixture proposal (7) can be interpreted as a joint proposal distribution for the pair (x_t, a_t) , given by

$$q(x_t, a_t|y_{1:t}) = w_{t-1}^{a_t} q(x_t|x_{t-1}^{a_t}, y_t) \quad (8)$$

Here, a_t should be thought of as an index selecting one of the components in the sum (7). Generating, independently, N realizations from this joint proposal distribution can be done as follows.

1. Sample the ancestor indices $\{a_t^i\}_{i=1}^N$ according to

$$\mathbb{P}(A_t = i | \{x_{t-1}^j, w_{t-1}^j\}_{j=1}^N) = w_{t-1}^i \quad i = 1, \dots, N \quad (9)$$

This corresponds to the resampling step of the algorithm. The resampled particles are given as $\bar{x}_{t-1}^i = x_{t-1}^{a_t^i}$ for $i = 1, \dots, N$.

2. Propagate the particles to time t by simulating $x_t^i \sim q(x_t | \bar{x}_{t-1}^i, y_t)$ for $i = 1, \dots, N$

The advantage of explicitly introducing and implementing auxiliary variable lies in the computation of the importance weight. From (6), we have that the joint target distribution for (x_t, a_t) is proportional to what called unnormalized joint target density:

$$w_{t-1}^{a_t} p(y_t | x_t) p(x_t | x_{t-1}^{a_t}) \quad (10)$$

In fact, it is possible to use the information available in the current observation y_t , not only when proposing the new state x_t , but also when proposing its ancestor index a_t . The principle is that we can thereby increase the probability of resampling particles at time $t-1$ that are in agreement with the observation y_t .

It is free to use any proposal distribution that we find appropriate to obtain the pair (x_t, a_t) (Thomas B. Schon, Fredrik Lindsten, 2017). For example, let V be the function that will be used to adapt the proposal distribution for the ancestor indices. For each particle, i.e. for $i = 1, \dots, N$, we then compute the quantities

$$v_{t-1}^i := v(x_{t-1}^i, y_t), \quad (11)$$

referred to as adjustment multipliers. The adjustment multipliers are used to construct a proposal distribution for the ancestor index variable a_t

$$\mathbb{P}(A_t = i | \{x_{t-1}^j, w_{t-1}^j\}_{j=1}^N, y_t) = \frac{w_{t-1}^i V_{t-1}^i}{\sum_{l=1}^N w_{t-1}^l v_{t-1}^l} \quad i = 1, \dots, N \quad (12)$$

From the above expression, it is clear that v acts as adjusting the original weights w via multiplication. The underlying ideas of implementation of adjustment multipliers is that by carefully choosing the function v , we can adapt the sampling of the ancestor indices in (16) to make use of the information available in the current measurement y_t

Then, we get the following expression for the i importance weight as the use of adjustment multipliers:

$$\bar{w}_t^i = w(x_{t-1}, x_t, y_t) = \frac{w_{t-1}^{a_t^i} p(y_t | x_t^i) p(x_t^i | x_{t-1}^{a_t^i})}{w_{t-1}^{a_t^i} v(\bar{x}_{t-1}, y_t) q(x_t^i | x_{t-1}^{a_t^i}, y_t)} \quad (13)$$

This completes the algorithm, since these weighted particles in turn can be used to approximate the filtering PDF at time $t+1$, then at time $t+2$ and so on.

1.4.3 Bootstrap particle filter

After we have known the Auxiliary particle filter, a pragmatic solution for bootstrap particle filter is to choose the proposal density according to

$$w_{t-1}^{a_t^i} v_{t-1}^{a_t^i} = w_{t-1}^{a_t^i}, \quad (14)$$

$$q(x_t | x_{t-1}^{a_t^i}, y_t) = p(x_t | x_{t-1}^{a_t^i}), \quad (15)$$

$$q(x_t, a_t | y_{1:t}) = \sum_{j=1}^N w_{t-1}^j p(x_t | x_{t-1}^j) \quad (16)$$

Therefore, From (6) and (7) it follows that the weight function is given by

$$\bar{w}_t^i = \frac{p(x_t, a_t | y_{1:t})}{q(x_t, a_t | y_{1:t})} = p(y_t | x_t) \quad (17)$$

This completes the algorithm.

Bootstrap particle filter

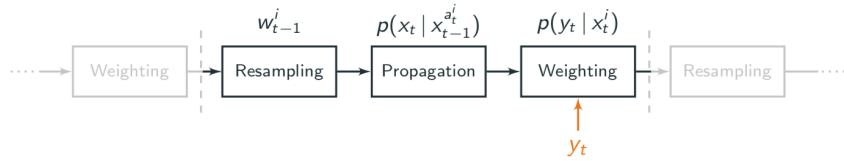


Figure 6:

1.4.4 Fully adapted particle filter

In order to adapt the proposals to the information that is available in the current observation y_t . A nature choice as for the proposal for the state x_t is to use

$$q(x_t | x_{t-1}, y_t) = p(x_t | x_{t-1}, y_t), \quad (18)$$

From Bayes'rule, we can write

$$p(x_t|x_{t-1}, y_t) = \frac{p(y_t|x_t)p(x_t|x_{t-1})}{p(y_t|x_{t-1})}, \quad (19)$$

Plugging this expression into the denominator of equation (13) we obtain

$$\bar{w}_t^i = w(x_{t-1}, x_t, y_t) = \frac{p(y_t|x_{t-1})}{v(x_{t-1}, y_t)}, \quad (20)$$

When it confront the choice to decide the adjustment multipliers. In particular, we see the choice $v(x_{t-1}, y_t) = p(y_t|x_{t-1})$ will lead to a weight function that is identically equal to 1. In other word, we obtain particles that are generated in such a way that they are equally informative about the target distribution.

$$\bar{w}_t^i = 1. \quad (21)$$

This completes the algorithm.

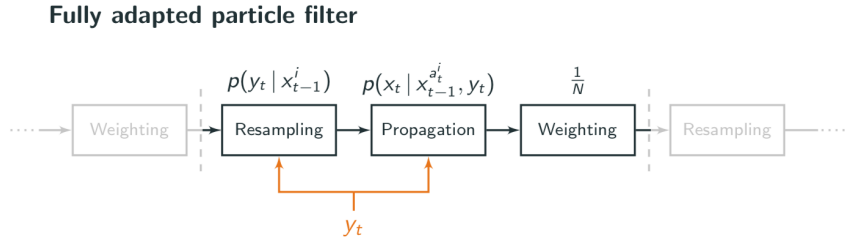


Figure 7: