

0.1 Solution:

(a)

$$n = (2 \times 3)^{15} = 470184984576$$

(b)

$$n = (2 \times 3)^{14} \times 15 \times 3 = 3526387384320$$

(c)

$$(2 \times 2)^{15} = 1073741824$$

0.2 Solution:

(a)

$$P(A^S \cap B^S) = 1 - 0.3 - 0.5 = 0.2$$

(b)

$$P(A \cap B^S) = P(A) = 0.3$$

(c)

$$P(A \cap B) = P(\emptyset) = 0$$

0.3 Solution:

$$\begin{aligned} P(\text{Accept}) &= P(\text{American Express}) + P(\text{Visa}) - P(\text{American Express}) \cap P(\text{Visa}) \\ &= 0.24 + 0.61 - 0.11 \\ &= 0.74 \end{aligned}$$

0.4 Solution:

$$P(\text{Blackjack}) = \frac{4 \times 16}{\binom{52}{2}} = \frac{64}{1326} = 0.048$$

0.5 Solution:

$$\begin{aligned} P(\text{No blackjack}) &= 1 - P(I \cup \text{Dealer}) \\ &= 1 - (P(I) + P(\text{Dealer}) - P(I) \cap P(\text{Dealer})) \\ &= 1 - (0.048 + 0.048 - \frac{2(4 \times 16)(3 \times 15)}{\binom{52}{4}}) \\ &= 1 - (0.096 - 5760/270725) = 1 - 0.075 = 0.925 \end{aligned}$$

0.6 Solution:

(a)

$$P(1) = 4/20 = 0.2$$

$$P(2) = 8/20 = 0.4$$

$$P(3) = 5/20 = 0.25$$

$$P(4) = 2/20 = 0.1$$

$$P(5) = 1/20 = 0.05$$

(b)

$$n = 4 \times 1 + 8 \times 2 + 5 \times 3 + 2 \times 4 + 5 = 48$$

$$P(1) = 4/48 = 1/12$$

$$P(2) = 16/48 = 1/3$$

$$P(3) = 15/48 = 5/16$$

$$P(4) = 8/48 = 1/6$$

$$P(5) = 5/48$$

0.7 Solution: In a single roll,

$$P(5) = 4/36 = 1/9$$

$$P(7) = 6/36 = 1/6$$

$$P(5 \cup 7) = 10/36 = 5/18$$

$$P(\text{not } 5 \text{ and not } 7) = 1 - 5/18 = 13/18$$

Thus,

$$P(E_n) = (13/18)^{n-1}/9$$

and

$$P(5 \text{ comes first}) = \sum_{n=1}^{\infty} P(E_n) = \lim_{n \rightarrow \infty} \frac{5}{6} [1 - (\frac{13}{18})^n] = \frac{5}{6}$$

0.8 Solution: In a single draw,

$$P(\text{red}) = 3/10$$

$$P(\text{black}) = 7/10$$

If denote the event that A will draw the first red ball in the nth cycle as E_n ,

$$\begin{aligned} P(E_n) &= \frac{3}{10 - 2(n-1)} \prod_{i=1}^{n-1} \left(\frac{9-2n}{12-2n} \frac{8-2n}{11-2n} \right) \\ &= \frac{3}{12-2n} \prod_{i=1}^{n-1} \left(\frac{9-2n}{12-2n} \frac{8-2n}{11-2n} \right) \end{aligned}$$

Since there are only 7 black balls, so A can only draw red ball in the first 4 cycles. Thus,

$$P(A) = \sum_{i=1}^4 P(E_n) = 3/10 + 7/40 + 1/12 + 1/40 = 7/12$$

0.9 Solution:

(a)

$$P(\text{same}) = P(\text{red}) + P(\text{green}) + P(\text{blue}) = (5/19)^3 + (6/19)^3 + (8/19)^3 = 0.124$$

(b)

$$P(\text{diff}) = (5/19)(6/19)(8/19) = 0.035$$

0.10 Solution: Having a girl on position i means to randomly choose a girl to interlope her into the rest $b+g-1$ people to divide any permutation of them into 2 part. Thus,

$$P(i) = \frac{(b+g-1)!g}{(b+g)!} = \frac{g}{b+g}$$

0.11 Proof: $EF^C = (E^C)^C F^C = (E^C \cup F)^C$, then $P(EF^C) = P((E^C \cup F)^C) = 1 - P(E^C \cup F) = 1 - (P(E^C) + P(F) - P(E^C F)) = 1 - ((1 - P(E)) + P(F) - P(E^C F)) = 1 - 1 + P(E) - P(F) + P(E^C F) = P(E) - (P(F) - P(E^C F)) = P(E) - P(EF)$ ■

0.12 Proof: $P(EF) = P((E^C)^C (F^C)^C) = P((E^C \cup F^C)^C) = 1 - P((E^C \cup F^C)) = 1 - (P(E^C) + P(F^C) - P(E^C F^C)) = 1 - ((1 - P(E)) + (1 - P(F) - P(E^C F^C))) = 1 - (2 - (P(E) + P(F) - P(E^C F^C))) = (P(E) + P(F)) - 1 + P(E^C F^C) \geq P(E) + P(F) - 1$ ■

Base case is proved above, we can suppose that when $n = k$

$$P\left(\bigcap_{i=1}^k E_i\right) \geq \sum_{i=1}^k P(E_i) - (k-1)$$

is true.

When $n = k + 1$,

$$\begin{aligned}
P\left(\bigcap_{i=1}^{k+1} E_i\right) &= P\left(\bigcap_{i=1}^k E_i \cap E_{k+1}\right) \\
&= P\left(\left(\left(\bigcap_{i=1}^k E_i\right)^C\right)^C \cap (E_{k+1}^C)^C\right) \\
&= P\left(\left(\left(\bigcap_{i=1}^k E_i\right)^C \cup E_{k+1}^C\right)^C\right) \\
&= 1 - P\left(\left(\bigcap_{i=1}^k E_i\right)^C \cup E_{k+1}^C\right) \\
&= 1 - \left(P\left(\left(\bigcap_{i=1}^k E_i\right)^C\right) + P(E_{k+1}^C) - P\left(\left(\bigcap_{i=1}^k E_i\right)^C E_{k+1}^C\right)\right) \\
&= 1 + P\left(\left(\bigcap_{i=1}^k E_i\right)^C E_{k+1}^C\right) - \left((1 - P\left(\bigcap_{i=1}^k E_i\right)) + (1 - P(E_{k+1}))\right) \\
&= P\left(\left(\bigcap_{i=1}^k E_i\right)^C E_{k+1}^C\right) + P\left(\bigcap_{i=1}^k E_i\right) + P(E_{k+1}) - 1
\end{aligned}$$

According to our hypothesis

$$P\left(\bigcap_{i=1}^k E_i\right) \geq \sum_{i=1}^k P(E_i) - (n - 1)$$

Then

$$\begin{aligned}
P\left(\bigcap_{i=1}^{k+1} E_i\right) &\geq \sum_{i=1}^k P(E_i) - (n - 1) + P(E_{k+1}) - 1 \\
&\geq \sum_{i=1}^k P(E_i) - (n - 1) + P(E_{k+1}) - 1 \\
&\geq \sum_{i=1}^{k+1} P(E_i) - n
\end{aligned}$$

We can conclude that $\forall n \in \mathbb{N}$

$$P\left(\bigcap_{i=1}^n E_i\right) \geq \sum_{i=1}^n P(E_i) - (n - 1)$$

■

0.13 Solution:

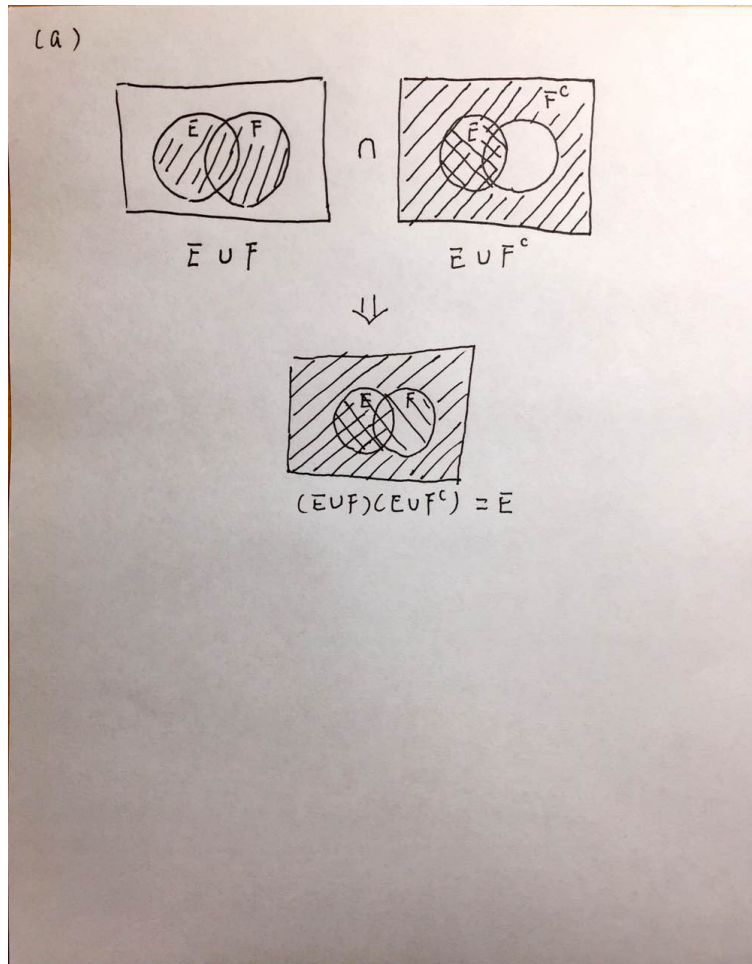


Figure 1: 0.13.a

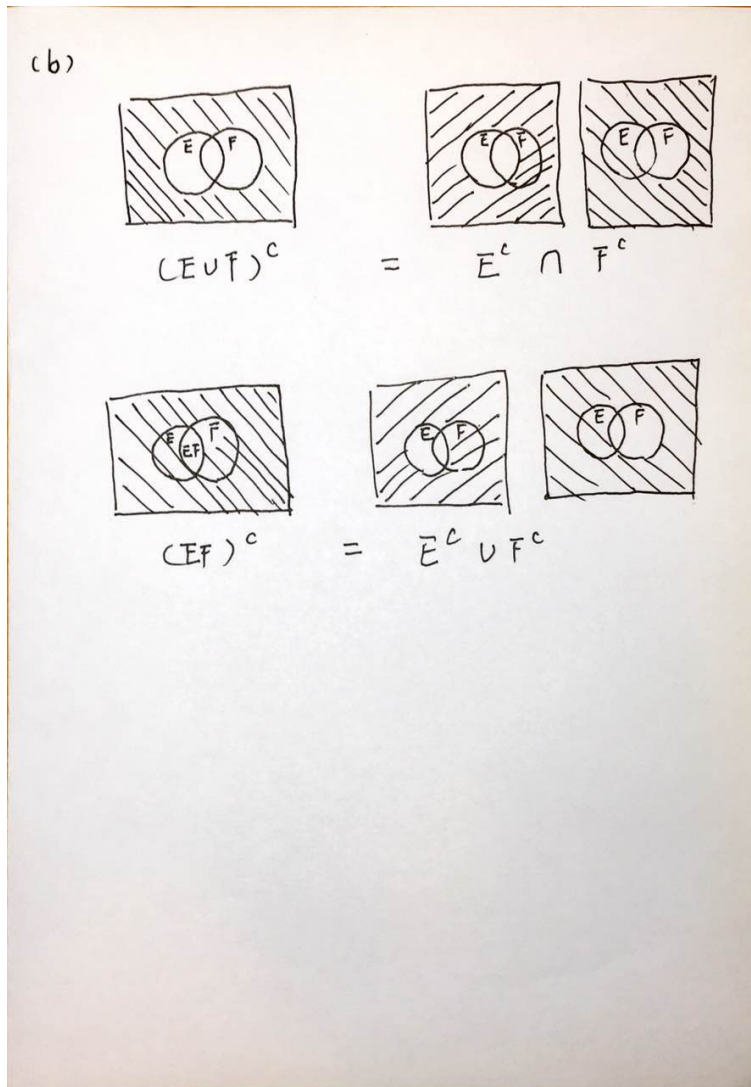


Figure 2: 0.13.b