## MTH-222, MTH-6031: Probability and Statistics

## Tutorial # 5 (Expectation)

- 1. Let X be a continuous random variable with pdf f. Suppose X is symmetric (i.e., X and -X have the same distribution) and  $X^2 \sim \exp(\lambda)$ . Find f.
- 2. Let X be a random variable such that  $P\{|X-1|=2\}=0$ . Is X a continuous random variable? Also write down  $P\{|X-1|\geq 2\}$  in terms of distribution function of X.
- 3. Let X be Poisson with parameter  $\lambda$ . Find mean of  $\frac{1}{1+X}$ .
- 4. Find the variance of Binomial(n, p), Poisson $(\lambda)$ , exponential $(\lambda)$  and  $N(\mu, \sigma^2)$  random variables.
- 5. Let X be exponentially distributed with parameter  $\lambda$ . Find the 4th moment.
- 6. Let X be standard normal random variable. Find E|X|.
- 7. Let X be a Binomial random variable with parameters  $n=4, p=\frac{1}{2}$ . Find  $E\left[\sin\left(\frac{\pi x}{2}\right)\right]$ .
- 8. Let X be a geometrically distributed random variable with parameter p and M be a positive integer. Find  $E[\min\{X,M\}]$ .
- 9. Let X be a discrete random variable with pmf

$$f_X(k) = P(X = k) = \frac{1}{9}$$
, for  $k = -4, -3, -2, \dots, 3, 4$ 

Find var(X).

10. Let

$$P_X(x) = \begin{cases} \frac{1}{2^x}, & x = 1, 2, 3, \dots, \\ 0, & otherwise. \end{cases}$$

be the pmf of the random variable X, then find all the possible values of t such that  $E[e^{tX}]$  exist.

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1. Let X be a continuous random variable with pdf f. Suppose X is symmetric (i.e., X and -X have the same distribution) and  $X^2 \sim \exp(\lambda)$ . Find f.

Solution ->

$$F_{x}(x) = P(x \le x)$$

$$= F_{-x}(x) = P(-x \le x) = P(xx - x)$$

$$= 1 - P(x < -x)$$

$$= 1 - F_{x}(-x)$$

$$= F_{x}(x) = F_{x}(x) =$$

$$F_{x}(x) = 1 - F_{x}(-x) + x \in \mathbb{R}$$

$$F_{x^{2}}(x) = P(x^{2} + x)$$

$$= \begin{cases} 0 & \text{if } x \neq 0 \\ P(-5x \neq x \neq 5x) & \text{if } x \neq 0 \end{cases}$$

$$F_{x^{2}}(x) = \begin{cases} 0 & \text{if } x \neq 0 \\ F_{x}(5x) - F_{x}(5x) & \text{if } x \neq 0 \end{cases}$$

$$CDF \text{ of } exp(\lambda) \text{ distribution is}$$

$$= \begin{cases} 0 & \text{if } x \neq 0 \\ 1 - e^{-\lambda x} & \text{if } x \neq 0 \end{cases}$$

$$So \text{ for } x \neq 0 \text{, we have}$$

$$F_{x}(5x) - F_{x}(-5x) = 1 - e^{-\lambda x}$$

$$F_{\chi}(\sqrt{2}) - (1 - F_{\chi}(\sqrt{2}))$$

$$2 F_{\chi}(\sqrt{2}) - 1 = 1 - e^{-\lambda n}$$

$$F_{x}(5x) = \frac{2 - e^{-\lambda x}}{2}, \quad n > 0$$

$$=\sum_{x} f_{x}(x) = \frac{2-e^{-\lambda x^{2}}}{2}, \quad x > 0$$

$$= \int_{X} (x) = + \lambda (2x) e^{-\lambda x^{2}}; \quad x \geqslant 0$$

$$\int_{\Lambda} (\chi) = \chi \chi e^{-\chi \chi^2}$$

For x 600

$$F_{\chi}(\chi) = 1 - F_{\chi}(-\chi)$$

$$= 1 - \left(\frac{2 - e^{-\lambda \chi^2}}{2}\right)$$

$$= e^{-\lambda \chi^2}$$

$$f_{\chi}(\chi) = -\lambda \chi e^{-\lambda \chi^2}$$
;  $\chi < 0$ 

$$f_{x}(\pi) = \lambda |\chi| e^{-\lambda \pi^{2}}$$
 $g_{x} \in \mathbb{R}$ 

2. Let X be a random variable such that  $P\{|X-1|=2\}=0$ . Is X a continuous random variable? Also write down  $P\{|X-1|\geq 2\}$  in terms of distribution function of X.

Soln ->
$$P(X \in \{3,-1\}) = 0$$

$$-\Omega = \{3,-3,1,-1\}$$

$$X(\omega) = \omega$$

$$P(1) = 1/2 \qquad P(-3) = 1/2$$

$$P(3) = 0 = P(-1)$$

$$L > Discrete s. v.$$

$$P(|X-1| > 2)$$
=  $P(|X-1| > 2) + P(-(|X-1| > 2))$ 
=  $P(|X| > 3) + P(|X| < -1)$ 
=  $|I-P(|X| < 3) + F_{x}(-1)$ 
=  $|I-F_{x}(|3-|) + F_{x}(-1)$ 

3. Let X be Poisson with parameter  $\lambda$ . Find mean of  $\frac{1}{1+X}$ .

$$\frac{1}{1+x} = \sum_{k=0}^{\infty} \frac{1}{1+k} P(x=k)$$

$$= \sum_{k=0}^{\infty} \frac{1}{1+k} e^{-\lambda} \frac{1}{k!}$$

$$= \frac{e^{-\lambda}}{\lambda} \sum_{k=0}^{\infty} \frac{\lambda^{k+1}}{(k+1)!}$$

$$= \frac{e^{-\lambda}}{\lambda} (e^{\lambda} - 1)$$

4. Find the variance of Binomial(n, p), Poisson $(\lambda)$ , exponential $(\lambda)$  and  $N(\mu, \sigma^2)$  random variables.

• Variance 
$$B(n, b)$$
  
 $f(x) = nb$ 

$$EX = np$$

$$E(\chi^2) = \sum_{k=0}^{\infty} k^2 P(\chi=k)$$

$$= \sum_{k=1}^{n} k^{2} \binom{n}{k} p^{k} (1-p)^{n-k}$$

$$= \frac{\frac{n}{k}}{\frac{k!}{k!}} \frac{k^2 n!}{k! (n-k)!} p^k (1-p)^{n-k}$$

$$= \frac{\sum_{k=1}^{n} \frac{k}{(k-1)!} \frac{h}{(n-k)!} p^{k} (1-p)^{n-k}}{(k-1)! (n-k)!}$$

$$= \sum_{k=1}^{n} (k-1+1) \frac{n!}{(k-1)!} (n-k)! p^{k} (1-k)^{n-k}$$

$$= \sum_{k=1}^{n} \frac{(k-1)}{(k-1)!} \frac{h!}{(n-k)!} \frac{h!}{(k-1)!} \frac{h!}{(n-k)!} + \sum_{k=1}^{n} \frac{n!}{(k-1)!} \frac{h!}{(n-k)!}$$

$$= \frac{\sum_{k=2}^{n} \frac{n!}{(k-2)!} \frac{p^{k} (1-p)^{n-k}}{(n-k)!} + n^{k}$$

$$= n(n-1)/p^2 \sum_{k=2}^{n} \frac{(n-2)!}{(k-2)! (n-k)!} p^{k-2} (1-p)^{n-k}$$

$$= \eta (n-1) p^{2} \left( \sum_{k=0}^{n-2} \frac{(n-2)!}{k! (n-k)!} \right)^{k} (1-p)^{n-k}$$

$$= \eta (n-1) p^{2} \left( \sum_{k=0}^{n-2} \frac{(n-k)!}{k! (n-k)!} \right)^{n-k}$$

$$E X^2 = n^2 p^2 - n p^2 + n p$$

$$\sqrt{ar}(X) = E(X^{2}) - (EX)^{2}$$

$$= n^{2}b^{2} - nb^{2} + nb - n^{2}b^{2}$$

$$= nb(1-b)$$

ii) Var Poisson (N)

$$E(x^2) = \sum_{k=0}^{\infty} k^2 e^{-\lambda} \lambda^k$$

$$= e^{-\lambda} \sum_{k=0}^{\infty} \frac{b(\lambda^{k})}{(k-1)!}$$

$$= e^{-\lambda} \left( \sum_{k=1}^{\infty} \frac{(k-1+1)\lambda^{k}}{(k-1)!} \right)$$

$$= e^{-\lambda} \left( \sum_{k=2}^{\infty} \frac{(k-1)\lambda^{k}}{(k-1)!} + \sum_{k=1}^{\infty} \frac{\lambda^{k}}{(k-1)!} \right)$$

$$= e^{-\lambda} \left( \sum_{k=2}^{\infty} \frac{\lambda^{k}}{(k-2)!} + \lambda \sum_{k=1}^{\infty} \frac{\lambda^{k-1}}{(k-1)!} \right)$$

$$= e^{-\lambda} \left( \lambda^{2} \sum_{k=2}^{\infty} \frac{\lambda^{k-2}}{(k-2)!} + \lambda e^{\lambda} \right)$$

$$= e^{-\lambda} \left( \lambda^{2} e^{\lambda} + \lambda e^{\lambda} \right)$$

$$= e^{-\lambda} \left( \lambda^{2} e^{\lambda} + \lambda e^{\lambda} \right)$$

$$= (\lambda^{2}) = \lambda^{2} + \lambda$$

$$= (\lambda^{2}) = (\lambda^{2}) - (\xi \lambda)^{2}$$

$$Var(X) = E(X^2) - (EX)^2$$
$$= \lambda^2 + \lambda - \lambda^2 = \lambda$$

5. Let X be exponentially distributed with parameter  $\lambda$ . Find the 4th moment.

$$X \sim exp(\lambda)$$

$$E(x^{4}) = \int_{0}^{\infty} x^{4} f_{x}(x) dx$$

$$= \int_{0}^{\infty} x^{4} \lambda e^{-\lambda x} dx$$

$$= \int_{0}^{\infty} x^{4} \lambda e^{-\lambda x} dx$$

$$= \frac{4!}{x^{4}}$$

6. Let X be standard normal random variable. Find E[X].

$$\sum_{n=1}^{\infty} |x| = \int_{-\infty}^{\alpha} |x| \frac{1}{\sqrt{2\pi}} e^{-n^{2}/2} dn$$

$$= \lim_{n \to \infty} \int_{-\alpha}^{\alpha} \frac{|x|}{\sqrt{2\pi}} e^{-x^{2}/2} dx$$

$$= 2 \int_{0}^{\infty} \frac{x e^{-x^{2}/2}}{\sqrt{2}\pi} dx$$

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$$= 2 \int_{0}^{\infty} \frac{x e^{-x^{2}/2}}{\sqrt{2}\pi} dx$$

$$= \sqrt{2} \int_{0}^{\infty} \frac{x e^{-x^{2}/2}}{\sqrt{2}\pi} dx$$

$$= \sqrt{2} \int_{0}^{\infty} \frac{x e^{-x^{2}/2}}{\sqrt{2}\pi} dx$$

8. Let X be a geometrically distributed random variable with parameter p and M be a positive integer. Find  $E[\min\{X, M\}]$ .

$$\begin{array}{ll}
X \sim \text{ geometric } (b) \\
E \left[ \text{ min } \left\{ X, m \right\} \right] \\
&= \sum_{k=1}^{\infty} \min \left\{ k, m \right\} P(X = k) \\
&= \sum_{k=1}^{m} \min \left\{ k, m \right\} p(1-p)^{k-1} + \sum_{k=m+1}^{\infty} \min \left\{ k, m \right\} p(1-p)^{k-1} \\
&= \sum_{k=1}^{m} k p(1-p)^{k-1} + m \sum_{k=m+1}^{\infty} p(1-p)^{k-1} \\
&= \sum_{k=1}^{m} k (1-p)^{k-1} + m \sum_{k=m+1}^{\infty} p(1-p)^{k-1} \\
&= \sum_{k=1}^{m} k (1-p)^{k-1} + m \sum_{k=m+1}^{\infty} p(1-p)^{k-1}
\end{array}$$

$$= \oint \sum_{k=1}^{\infty} \frac{d}{dp} \left( (1-p)^k \right) + mp$$

10. Let

$$P_X(x) = \begin{cases} \frac{1}{2^x}, & x = 1, 2, 3, \dots, \\ 0, & otherwise. \end{cases}$$

be the pmf of the random variable X, then find all the possible values of t such that  $E[e^{tX}]$  exist.

moment generating function af r. v X.

$$E\left[e^{tx}\right] = \sum_{k=1}^{\infty} e^{tk} \beta(x=k)$$

$$= \sum_{k=1}^{\infty} e^{tk} \frac{1}{2^k}$$

$$= \sum_{k=1}^{\infty} e^{tk} \frac{1}{2^k}$$

 $\frac{e^{t(k+1)}}{k\rightarrow\infty} \times \frac{2^{k}}{2^{tk}} = e^{t/2}$ 

The series converges if

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diverges if et/2 >1

et 22  $t \leq \ln 2 \Rightarrow \text{for these values}$ expectation of  $E[e^{tx}] \leq \infty$   $t \leq \ln 2 \quad E[e^{tx}] = \infty$