



Regressions and Gradient Descent

Machine Learning Decal

Hosted by Machine Learning at Berkeley

Agenda

Linear Regression

Optimization via Gradient Descent

5-minute break

Logistic Regression

Multinomial Regression

Questions

Extras

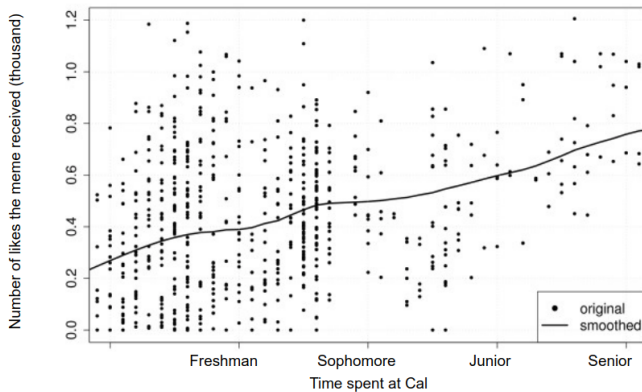
Linear Regression

Suppose you are applying for an internship at UCBMFET Corp and they gave you this technical challenge:

Suppose you are applying for an internship at UCBMFET Corp and they gave you this technical challenge:

- Investigate how the dankness of a meme is correlated with the time the meme creator has spent at UC Berkeley

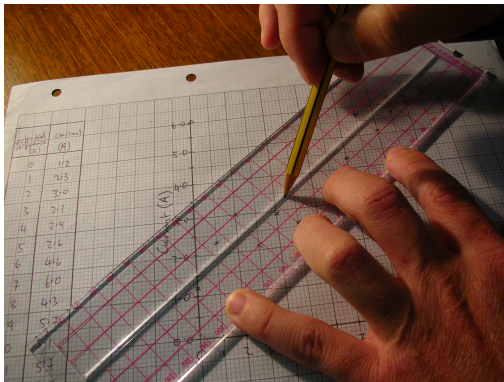
Dankness of meme vs time spent at Cal



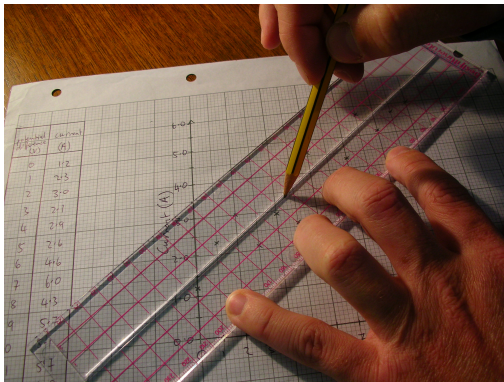
Question: Given any arbitrary time a student spent at Cal, can you tell me how dank the meme he/she creates will be?

Easy. You reach to your pocket and take out a

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and squint real hard to draw a best fit line.





You gotta step up your game with machine learning.

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- the y-intercept, b_1

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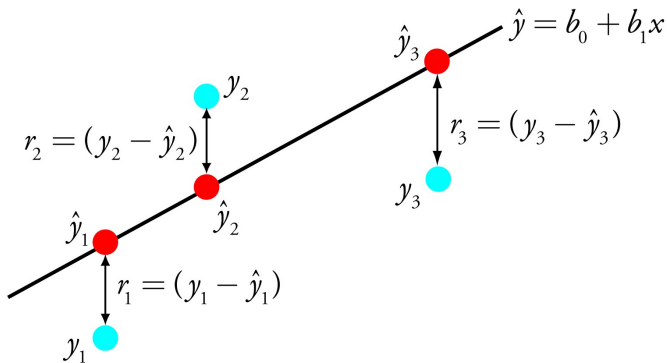
$$h(x) = b_0 + b_1x$$

Now the question becomes:

Given (x_i, y_i) pairs, how do you find b_0 and b_1 that give you the best-fit line?

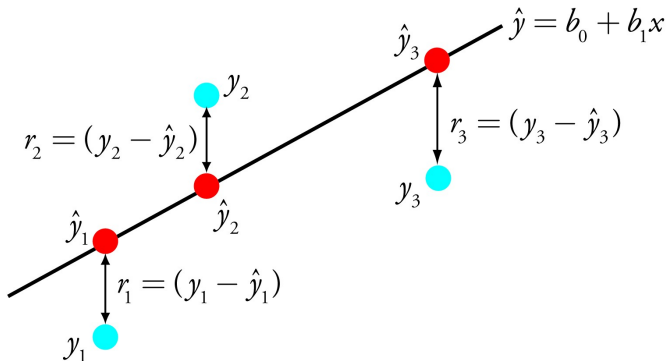
Let $\hat{y}_i = h(x) = b_0 + b_1x$

$$\min J(b_0, b_1)$$



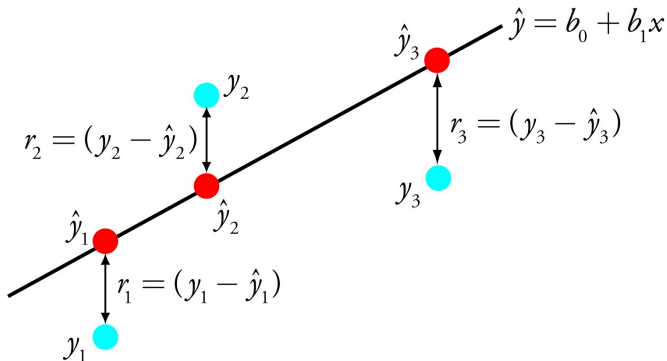
Let $\hat{y}_i = h(x) = b_0 + b_1x$

$$\min J(b_0, b_1) = \frac{1}{2m} \sum_{i=1}^m (y_i - \hat{y}_i)^2$$

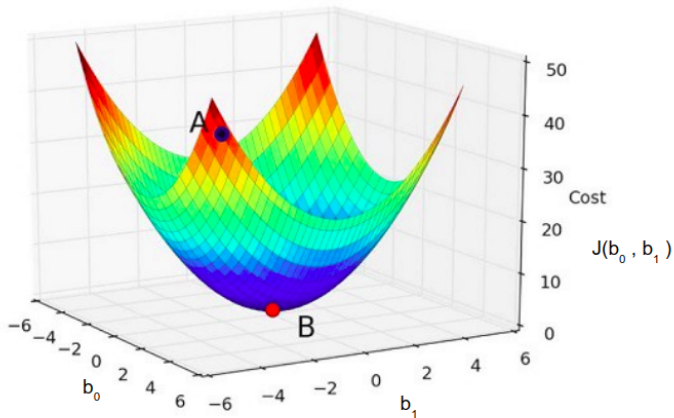


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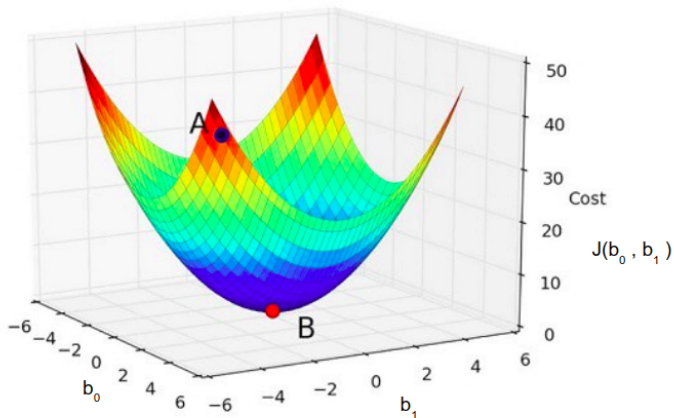
$$\min J(b_0, b_1) = \frac{1}{2m} \sum_{i=1}^m (y_i - \hat{y}_i)^2 = \frac{1}{m} \sum_{i=1}^m (y_i - b_0 - b_1 x_i)^2$$



Visualizing the Cost Function



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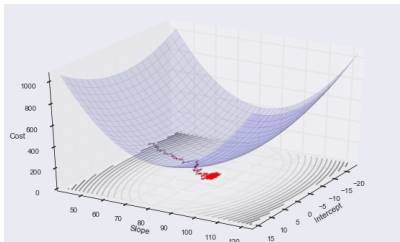


Get point $B = (b_0, b_1)$ such that $J(b_0, b_1)$ is the smallest.

Optimization via Gradient Descent

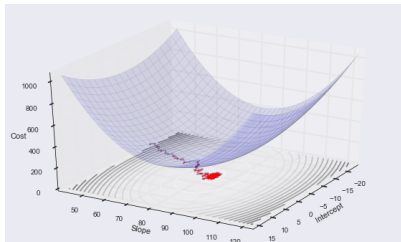
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- Choose an initial guess for b_0, b_1



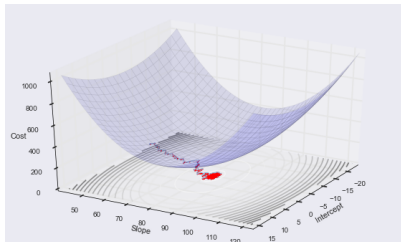
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- Choose an initial guess for b_0, b_1
- Repeatedly update b_0, b_1 to make $J(b_0, b_1)$ smaller
- Keep doing this until $J(b_0, b_1)$ reaches its minimum



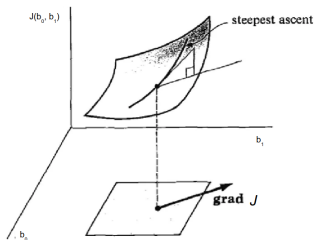
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grad J is the vector that points in the direction with the largest increase (steepest ascent)

Cost Function:

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Derivatives (to determine the direction of descent):

$$\frac{\partial}{\partial b_0} J(b_0, b_1) = \frac{1}{m} \sum_{i=1}^m (h(x^{(i)}) - y^{(i)})$$

$$\frac{\partial}{\partial b_1} J(b_0, b_1) = \frac{1}{m} \sum_{i=1}^m (h(x^{(i)}) - y^{(i)}) \cdot x^{(i)}$$

1. Initialize random b_0, b_1
2. Repeat until convergence {

$$b_0 := b_0 - \alpha \frac{\partial}{\partial b_0} J(b_0, b_1)$$

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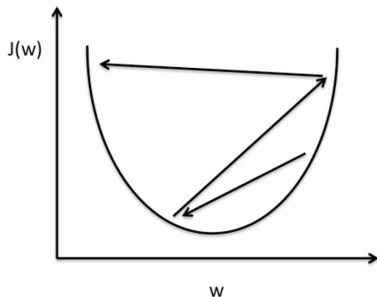
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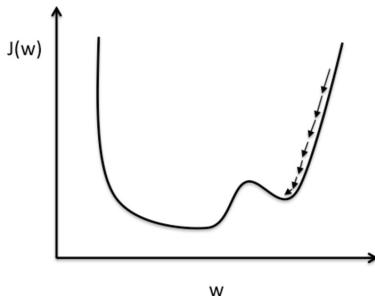
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Large learning rate: Overshooting.



Small learning rate: Many iterations until convergence and trapping in local minima.

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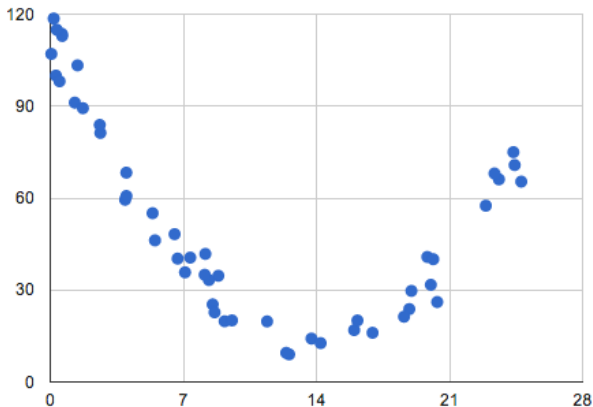
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So now given any arbitrary value x , we have a model $h(x)$ that can predict what the corresponding best prediction y will be.

Questions:

- But is a straight line always the line of best fit?
- Can the cost function $J(b)$ be smaller?

What if the interviewer decides to give you this dataset instead?



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Note: This method can be generalized to any polynomials!

$$h(x) = b_0 + b_1x_1 + b_2x_2^2 + \dots + b_nx_n^n$$

Now the interviewer gives you one more set of data (x_1, x_2, y)

x_1 = The time spent at Cal

x_2 = The time spent on Facebook

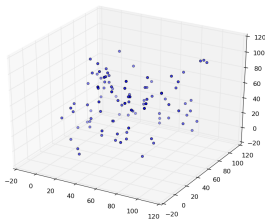
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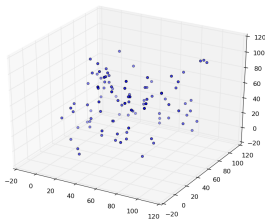
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Try visualizing the cost function, $J(b_0, b_1, b_2)$

Impossible! Beyond 3-dimensional, we need Linear Algebra

$x_j^{(i)}$: i^{th} sample, j^{th} feature

$$h(x_1^{(i)}, x_2^{(i)}) = b_0 + b_1 x_1^{(i)} + b_2 x_2^{(i)}$$

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$$\begin{pmatrix} \hat{y}^{(1)} \\ \hat{y}^{(2)} \\ \vdots \\ \hat{y}^{(m)} \end{pmatrix} = \begin{pmatrix} 1 & x_1^{(1)} & x_2^{(1)} \\ 1 & x_1^{(2)} & x_2^{(2)} \\ \vdots & \vdots & \vdots \\ 1 & x_1^{(m)} & x_2^{(m)} \end{pmatrix} \begin{pmatrix} b_0 \\ b_1 \\ b_2 \end{pmatrix}$$

$$\vec{\hat{y}} = \mathbf{X} \vec{b}$$

Let $\vec{e} = \vec{y} - \hat{\vec{y}}$

$$\vec{y} = \hat{\vec{y}} + \vec{e}$$

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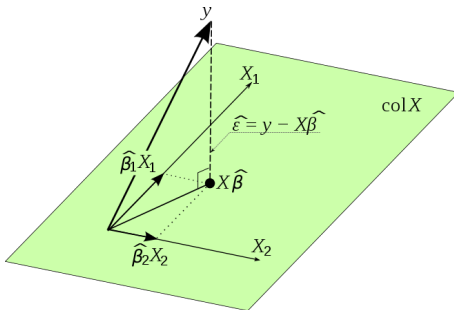


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5-minute break

Logistic Regression

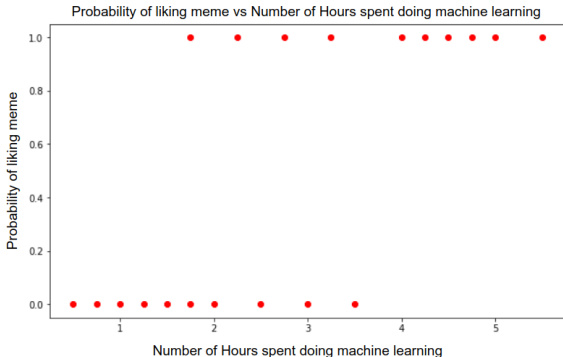
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Given this meme and the time an individual spent doing machine learning, predict whether he/she will like this meme

Why is linear regression a sub-optimal algorithm for this problem?



- Recall that linear regression is for **continuous** dependent variables, e.g. Dankness of memes vs. Time spent at Cal.

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- The current problem has **categorical** values for the dependent variable, i.e. Like/No like - only two **classes**
- We are regressing on (the likelihood of) **membership to a class**.

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- What is its domain and range?

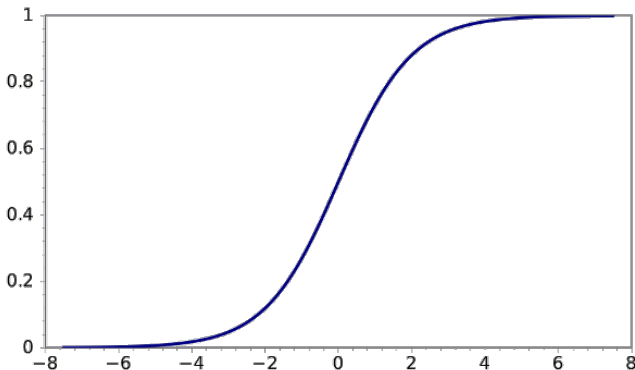
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- What is its domain and range?
- What is an interesting property regarding $s(x = 0)$?

Here's a picture. Notice that $s(x) - \frac{1}{2}$ intuitively appears to be an odd function. Hence, we have the following property for all x :

$$s(x) + s(-x) = 1$$



Recall from Linear Regression, that our hypothesis has the form:

$$h(x_1, x_2, \dots, x_n) = b_0 + b_1x_1 + \dots + b_nx_n = \vec{b}^T \vec{x}$$

And, in our new problems, our dependent variables y are in $\{0, 1\}$.

Recall from Linear Regression, that our hypothesis has the form:

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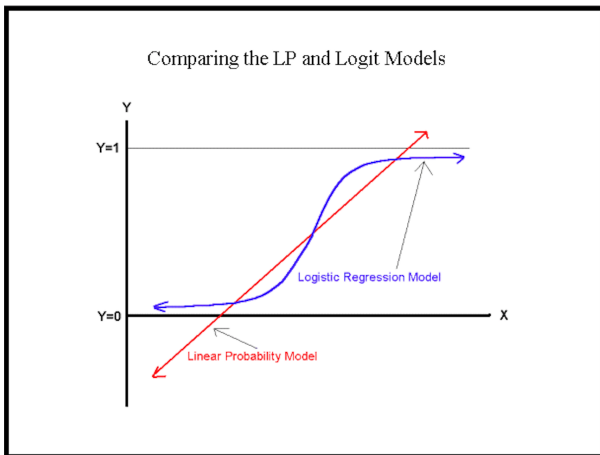
And, in our new problems, our dependent variables y are in $\{0, 1\}$.
So, we propose that the hypothesis for Logistic Regression be:

$$h(\vec{x}) = s(\vec{b}^T \vec{x}) = \frac{1}{1 + e^{-\vec{b}^T \vec{x}}}$$

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- Finally, we note that $h(x)$ is also known as an **expit**. It converts log-probability into a probability. The inverse function is known as a **logit**.



Our hypothesis is still parameterized by $b : [b_0 \dots b_n]$.

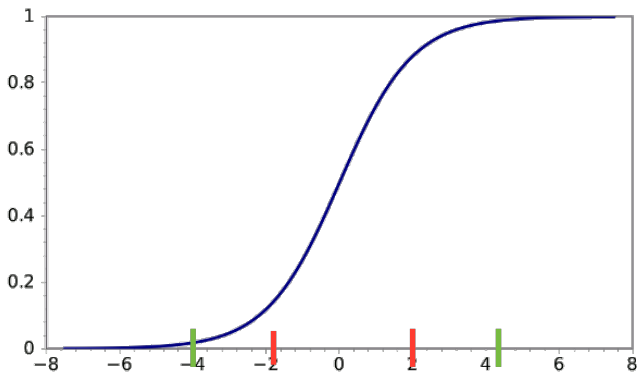
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The ideal b would be such that for all x such that $y = 1$, we have

$$h(x) \approx 1 \rightarrow \vec{b}^T \vec{x} \gg 0$$

A similar conclusion follows for $y = 0$.

Where would **ideal** values of $\vec{b}^T \vec{x}$ fall on the x-axis relative to the colored bands?



To improve our weights, we will seek to minimize the Logistic Loss Function, with $z = h(x) = \frac{1}{1 + e^{-\vec{b}^T \vec{x}}}$:

$$J(b) = - \sum_{i=1}^m \left(y^{(i)} \cdot \ln z^{(i)} + (1 - y^{(i)}) \cdot \ln (1 - z^{(i)}) \right)$$

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Because $\forall i, y_i \in \{0, 1\}$, only one of the addends is nonzero. That addend will be minimized when the expression involving z_i is closest to zero. (Remember $\ln(1) = 0$).

Again, we need to find the partial derivative with respect to b

$$\frac{\partial J(b)}{\partial b} = \sum_{i=1}^m -\frac{\partial z^{(i)}}{\partial b} \cdot \frac{\partial}{\partial z^{(i)}} J(b)$$

(see appendix for details)

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Expressed in Matrix form:

$$\frac{\partial J(\vec{b})}{\partial \vec{b}} = -X^T (\vec{y} - s(X\vec{b}))$$

The (mini-)batch gradient descent update rule then becomes:

$$b \leftarrow b + \alpha X^T (y^{(i)} - s(Xb))$$

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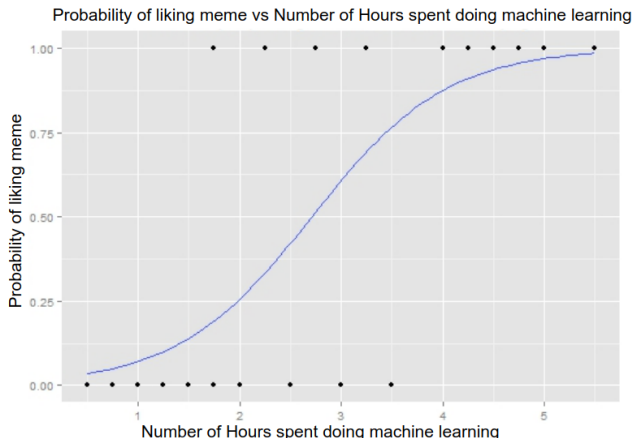
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For stochastic gradient descent, the update is:

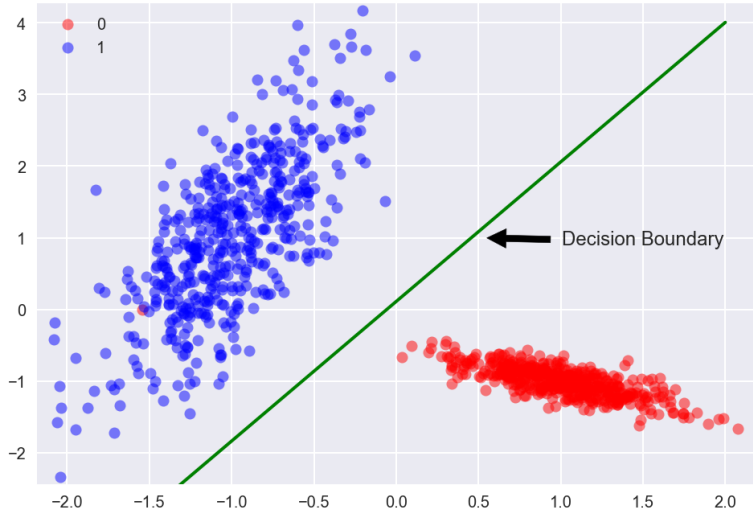
$$b \leftarrow b + \alpha (y^{(i)} - s(X_i b)) X_i$$

Take the time to understand how the summation is converted into a matrix calculation.

Just for reference, here is the logistic regression solution to the initial problem that was posed:



Reminder: Linear Decision Boundary



Here we have some of the key distinctions between the two kinds of regression:

| | Linear | Logistic |
|---------------------|-------------------|-------------------------|
| Label Type | Continuous | Categorical |
| Problem Type | Actual Regression | Actually Classification |
| Hypothesis | $\theta^T x$ | $s(\theta^T x)$ |
| Loss | Mean Squared | Logistic |
| Analytical Solution | Yes | No |

Multinomial Regression

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If logistic regression is answering True/False, then multinomial regression is answering a multiple choice question. The likelihood of any of the choices being correct forms a probability distribution.

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- Then we will need parameter vectors $\vec{b}_1, \dots, \vec{b}_k$.
- We can store these in parameter matrix: $\mathbf{B} \in \mathcal{R}^{k \times d}$, with each parameter vector being a row \vec{b}_i

We can give each class c_i a likelihood score:

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The product of the matrix **B** and vector x results in a vector z which needs to be **normalized**. We used sigmoid before.

$$\begin{bmatrix} y_1 \\ y_2 \\ y_3 \end{bmatrix} = \text{softmax} \left(\begin{bmatrix} W_{1,1} & W_{1,2} & W_{1,3} \\ W_{2,1} & W_{2,2} & W_{2,3} \\ W_{3,1} & W_{3,2} & W_{3,3} \end{bmatrix} \cdot \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix} + \begin{bmatrix} b_1 \\ b_2 \\ b_3 \end{bmatrix} \right)$$

Assuming that $h(x)$ still captures log-likelihood, we use the **softmax function** $S(z_i)$ to normalize multinomial regression scores:

Assuming that $h(x)$ still captures log-likelihood, we use the **softmax function** $S(z_i)$ to normalize multinomial regression scores:

$$P(y = c_i) = S(z_i) = \frac{e^{b_i x}}{\sum_{j=1}^k e^{b_j x}}$$

Note the denominator is the sum of all the exponentiated scores in the vector z .

The softmax loss function $J(b)$ is actually just a generalization of the logistic loss function:

$$J(b) = \sum_{i=1}^m \sum_{j=1}^k 1\{y^{(i)} = c_j\} \ln(z_j)$$

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The gradient and gradient update are as follows:

$$\nabla_{b_i} J(b) = x(z_i - y_i)$$

$$b_i := b_i - \alpha x(z_i - y_i)$$

In the first half of the lecture, we covered the normal equations solution to least squares regression. Now, we will do so for the regularized version:

$$\begin{aligned}\hat{b} &= \operatorname{argmin}_b ||y - Xb||_2^2 + \lambda ||b||_2^2 \\ &= \operatorname{argmin}_b y^T y - 2b^T X^T y + b^T X^T X b + \lambda b^T b\end{aligned}$$

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Taking the derivative and setting equal to zero is allowed since the function is convex in b

$$\begin{aligned}-2X^T y + 2X^T X \hat{b} + 2\lambda \hat{b} &= 0 \\ (X^T X + \lambda I_d) \hat{b} &= X^T y \\ \hat{b} &= (X^T X + \lambda I_d)^{-1} X^T y\end{aligned}$$

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Questions

Questions?

Extras

- Regression is a good summary of data, assuming the data has some key properties
- We need to know what those assumptions are, where they come from, and what to do when they fall apart

- Linearity
- Normality of errors

$$\epsilon_i \sim N(0, \sigma^2)$$

- Homoscedasticity (constant variance)

$$\text{Var}(\epsilon_i) = \text{Var}(\epsilon_j)$$

- Independence of errors

$$\epsilon_i \epsilon_j = 0 \quad \forall i \neq j$$

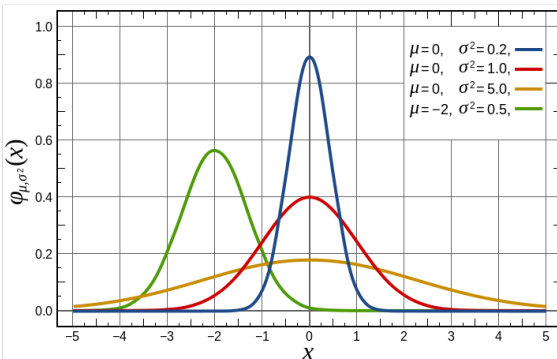
- The real world has natural processes. These processes can be collected/observed as data
- We try to build a model that mimics the real world model as close as possible
- Problem: there are some factors we can directly observe, and others that we can't
- Problem: we don't know what model the world uses
- In order to reason this process more rigorously, we can construct a "toy universe" to analyze our models

In our toy universe, data is generated through a linear model:

$$Y = X\theta$$

We can observe these Y values, but our observations have some noise in them. So our collected data looks like this:

$$Y = X\theta + Z \text{ where } Z \sim \mathcal{N}(0, \sigma^2)$$



So we have our noisy data:

$$Y = X\theta + Z$$

- We want to find the θ that can best replicate the data we've already observed
- Aka we want to **increase the likelihood** of the observed data being generated by the model

$$\hat{\theta} = \operatorname{argmax}_{\theta} P(y_1, y_2, \dots, y_n | \theta, x_1, x_2, \dots, x_n)$$

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$$\hat{\theta} = \operatorname{argmax}_{\theta} \prod_i P(y_i | \theta, x_i) = \operatorname{argmax}_{\theta} \sum_i \log P(y_i | \theta, x_i)$$

Remember that $y_i \sim \mathcal{N}(\theta^T x_i, \sigma^2)$:

$$\hat{\theta} = \operatorname{argmax}_{\theta} \sum_i \log \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{(y_i - \theta^T x_i)^2}{2\sigma^2}}$$

$$\hat{\theta} = \operatorname{argmax}_{\theta} \sum_i \log e^{-\frac{(y_i - \theta^T x_i)^2}{2\sigma^2}} = \operatorname{argmax}_{\theta} \sum_i -\frac{(y_i - \theta^T x_i)^2}{2\sigma^2}$$

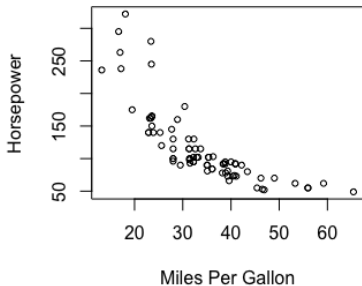
$$\hat{\theta} = \operatorname{argmin}_{\theta} \sum_i (y_i - \theta^T x_i)^2$$

- If the data is nonlinear...
 - Try performing a transformation on the independent or dependent variables such as squaring it, taking the log or square root, or ...
- If the errors are not normal...
 - Often, this isn't a big problem
 - Transformations help here too
 - Maybe subsets of the data are more normal than the overall set
 - Outliers and/or high leverage points may contribute to this issue

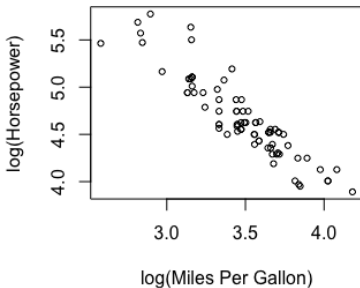
Example of the beauty of a log transform



MPG vs HP



Log(MPG) vs Log(HP)



First, we arrange all the data points into a **design matrix**, X , where point $x^{(i)}$ is the i^{th} row: X_i . We then find the appropriate gradient to solve $\min J(b)$.

$$\frac{\partial J(b)}{\partial b} = \sum_{i=1}^m -\frac{\partial z^{(i)}}{\partial b} \cdot \frac{\partial}{\partial z^{(i)}} J(b)$$

The first half of the chain rule:

$$\frac{\partial z^{(i)}}{\partial b} = \frac{\partial}{\partial b} s(b^T X_i) = s(b^T X_i)(1 - s(b^T X_i))X_i = z^{(i)}(1 - z^{(i)})X_i$$

You may want to check for yourself that $s'(x) = s(x)(1 - s(x))$

The other half of the chain rule:

$$\begin{aligned}\frac{\partial}{\partial z^{(i)}} J(b) &= \frac{\partial}{\partial z^{(i)}} (y^{(i)} \cdot \ln z^{(i)} + (1 - y^{(i)}) \cdot \ln (1 - z^{(i)})) \\ &= \frac{y^{(i)}}{z^{(i)}} - \frac{1 - y^{(i)}}{1 - z^{(i)}}\end{aligned}$$

Hence, we have:

$$\frac{\partial J(b)}{\partial b} = - \sum_{i=1}^n \left(\frac{y^{(i)}}{z^{(i)}} - \frac{1 - y^{(i)}}{1 - z^{(i)}} \right) z^{(i)} (1 - z^{(i)}) X_i = - \sum_{i=1}^n (y^{(i)} - z^{(i)}) X_i$$

Expressed in Matrix form:

$$\frac{\partial J(\theta)}{\partial \theta} = -X^T (y - s(X\theta))$$