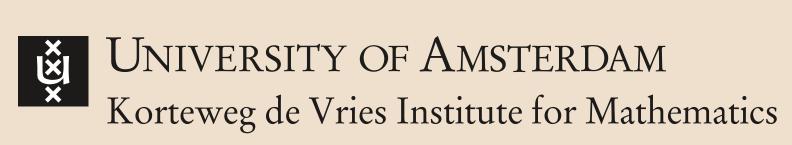
# Online Newton Method for Bandit Convex Optimisation

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#### Summary

We introduce a computationally efficient algorithm for zeroth-order bandit convex optimisation which achieves a regret of at most  $d^{3.5}\sqrt{n}$  polylog(n,d)in the adversarial setting and  $Md^2 \operatorname{polylog}(n,d)$ in the stochastic setting The parameter  $M \in$  $[d^{-1/2}, d^{-1/4}]$  depends on the geometry of the constraint set and the desired computational properties.

The previous best known rates were  $d^{10.5}\sqrt{n}$  for an efficient algorithm [1], and  $d^{2.5}\sqrt{n}$  for an inefficient algorithm using information-theoretic tools [2].

Our algorithm is built out of the following components:

- 1. The algorithm itself is adapted from [3] for the unconstrained case.
- 2. We introduce a bandit version of the extension proposed by [4].
- 3. To deal with the Adversarial case, negative bonuses, which can be seen as an increasing learning rate, are added.
- 4. A restart condition inspired by [1, 5] is added, which checks if the optimum is moving away and which guarantees negative regret on a restart.

## Bandit Convex Optimisation

In rounds  $t = 1, \ldots, n$ :

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- 1. Learner chooses action  $A_t \in K$
- 2. Learner suffers loss  $\ell_t(A_t)$
- 3. Learner observes  $Y_t = \ell_t(A_t) + \varepsilon_t$

Aim is to minimize the Regret with high probability:

$$\operatorname{Reg}_n = \sum_{t=1}^n \ell_t(A_t) - \min_{x \in K} \sum_{t=1}^n \ell_t(x).$$

#### Assumptions

Let

- $K \subset \mathbb{R}^d$ , a convex body with  $\mathbb{B}(1) \subseteq K \subseteq 2\mathbb{B}(d+1);$
- Convex  $\ell_1, ..., \ell_n : K \to [0, 1];$
- In the stochastic case:  $\ell = \ell_1 = \ldots = \ell_n$ ;
- The noise,  $\varepsilon_t$ , is conditionally sub-Gaussian and mean 0.

#### Theoretical Results

**Theorem 3.** There exists an algorithm such that with probability at least  $1 - \delta$ ,

$$\operatorname{Reg}_n \lesssim \begin{cases} d^{3.5}\sqrt{n}\operatorname{polylog}(n,d,1/\delta) & Adversarial; \\ d^{1.5}\sqrt{n}\operatorname{polylog}(n,d,1/\delta) & Stochastic, K \ in \ John's \ position \ or \ symmetric \ isotropic; \\ d^{1.75}\sqrt{n}\operatorname{polylog}(n,d,1/\delta) & Stochastic, K \ isotropic. \end{cases}$$

The parameters that achieve these regret results are

$$\gamma = \frac{1}{4dL}$$
  $\eta = \sqrt{\frac{d}{nL^3}}$   $\lambda = \frac{1}{d^3L^5}$   $\sigma^2 = \frac{1}{d^2}$   $\varepsilon = \frac{d^{3.5}L^{8.5}}{\sqrt{n}}$   $F_{\text{max}} = d^5L^8$ .

In the stochastic setting we choose

$$\gamma = 0$$
  $\eta = \frac{Md}{\sqrt{n}}$   $\lambda = \frac{5}{Md^{3/2}L^3}$   $\sigma^2 = \frac{1}{16M^2dL^3}$   $\varepsilon = \frac{Md^2L^5}{\sqrt{n}}$   $F_{\text{max}} = 25M^2d^3L^5$ .

#### Convex Extension, Bandit edition

Let  $\pi(x) = \inf\{t > 0 : x \in tK\}$ , the Minkowski functional

Shrink 
$$K \to K_{\varepsilon} = \{x \in \mathbb{R}^d : \pi_{\varepsilon}(x) = \frac{\pi(x)}{1 - \epsilon} \le 1\}$$
  
Define  $\pi_+(x) = \max(1, \pi_{\varepsilon}(x))$ , with  $\varepsilon = \Theta(1/\sqrt{n})$ 

**Definition 1 (Extension).** Let  $\ell$  be a convex function, define the extension  $f: \mathbb{R}^d \to \mathbb{R}$ :

$$f(x) = \pi_{+}(x)\ell\left(\frac{x}{\pi_{+}(x)}\right) + \frac{2(\pi_{+}(x) - 1)}{\varepsilon}$$
$$= \pi_{+}(x)\ell\left(\frac{x}{\pi_{+}(x)}\right) + 2v(x).$$

The output can be estimated with only one query X, set  $A = X/\pi_+(X)$ ,

$$Y = \pi_{+}(X)[\ell(A) + \varepsilon] + 2v(X).$$

Notice that f is  $\Theta(\sqrt{n})$ -Lipschitz, which renders existing analyses vacuous. However, in our case we can still prove that  $\sum_{t=1}^{n} Y_t^2 = \tilde{O}(n)$ 

#### Algorithm

1 input 
$$n$$
,  $\eta$ ,  $\lambda$ ,  $\gamma$ ,  $\sigma$  and  $K_0 = K_{\varepsilon}$ 

2 for 
$$t=1$$
 to  $n$ 

let 
$$\Phi_{t-1}(x) = \frac{1}{2} \|x\|^2 + \sum_{u=1}^{t-1} \flat_u(x) + \eta \sum_{u=1}^{t-1} \hat{q}_u(x)$$
  
compute  $\mu_t = \operatorname{arg\,min}_{x \in K_{t-1}} \Phi_{t-1}(x)$  and  $\Sigma_t^{-1} = \Phi_{t-1}''(\mu_t)$ 

sample 
$$X_t \sim \mathcal{N}(\mu_t, \Sigma_t)$$

play 
$$A_t = \frac{X_t}{\pi_+(X_t)}$$
 and observe  $Y_t = \pi_+(X_t)[\ell_t(A_t) + \varepsilon_t] + 2v(X_t)$ 

$$K_t = K_{t-1} \cap \{x : ||x - \mu_t||_t^2 \le F_{\text{max}}\}$$

compute 
$$z_t = \operatorname{arg\,min}_{z \in \mathbb{R}^d} \sum_{s=1}^{t-1} \mathbf{1}(\flat_s \neq \mathbf{0}) \|z - \mu_s\|_s^2$$

$$10 \qquad b_{t}(x) = \begin{cases} 0 & \text{if } \sum_{s=1}^{t-1} \mathbf{1}(b_{s} \neq \mathbf{0}) \|z_{t} - \mu_{s}\|_{s}^{2} \geq \frac{F_{\text{max}}}{16} \\ -\gamma \|x - \mu_{t}\|_{t}^{2} & \text{if } \|\cdot\|_{t}^{2} \not\leq \sum_{s=1}^{t-1} \mathbf{1}(b_{s} \neq \mathbf{0}) \|\cdot\|_{s}^{2} \\ -\gamma \|x - \mu_{t}\|_{t}^{2} & \text{if } \|\mu_{t} - z_{t}\|_{t}^{2} \geq \frac{F_{\text{max}}}{8} \end{cases}$$

11 **if** 
$$\max_{y \in K_t} \eta \sum_{u=1}^t (\hat{s}_u(\mu_u) - \hat{s}_u(y)) \le -\frac{\gamma F_{\text{max}}}{32}$$

then restart algorithm

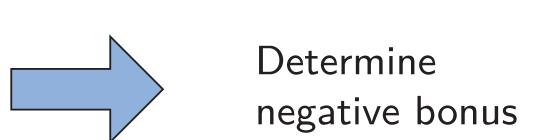
end if

14 end for

## FTRL on quadratic surrogate estimates

Sample action Observe extended loss

Update focus region





# Gaussian Smoothing

Definition 2 (Gaussian Smoothing). Let  $f: \mathbb{R}^d \to \mathbb{R}$  be a convex function, and  $X \sim \mathcal{N}(\mu, \Sigma)$ , given parameter  $\lambda \in (0, \frac{1}{1+d})$ , define

$$s(z) = \mathbb{E}\left[\left(1 - \frac{1}{\lambda}\right)f(X) + \frac{1}{\lambda}f((1 - \lambda)X + \lambda z)\right].$$

A quadratic approximation is defined as

$$q(z) = \langle s'(\mu), z - \mu \rangle + \frac{1}{4} \|z - \mu\|_{s''(\mu)}^{2}.$$

All necessary quantities can be estimated. Let  $r(X,z)=rac{p\left(\frac{X-\lambda z}{1-\lambda}\right)}{(1-\lambda)^d p(X)}$  with pthe density of the  $\mathcal{N}(\mu, \Sigma)$ -distribution. The relevant estimators are

• 
$$\hat{s}(z) = Y\left(1 + \frac{r(X,z) - 1}{\lambda}\right)$$

$$Yr(X,\mu) = 1 \left(X - \mu\right)$$
 $\{x \in A(x) \mid x \in A(x) = 1 \}$ 

Only concentrate well in the focus regions

• 
$$\hat{s}'(\mu) = \frac{Yr(X,\mu)}{1-\lambda} \Sigma^{-1} \left(\frac{X-\mu}{1-\lambda}\right) \qquad \{x \in K : \lambda \|z-\mu_t\|_{\Sigma_t^{-1}} \le \frac{1}{L}\}$$

• 
$$\hat{s}''(\mu) = \frac{\lambda Y r(X,\mu)}{(1-\lambda)} \left( \Sigma^{-1} \left[ \frac{X-\mu}{1-\lambda} \right] \left[ \frac{X-\mu}{1-\lambda} \right]^{\top} \Sigma^{-1} - \Sigma^{-1} \right)$$

• 
$$\hat{q}(z) = \langle \hat{s}'(\mu), z - \mu \rangle + \frac{1}{4} \|z - \mu\|_{\hat{s}''(\mu)}^2$$

# Geometry of the Constraint Set

The dimension dependence in the stochastic case of our algorithm depends on the mean width of the polar body  $K^{\circ}$  of the constraints set K:

$$M(K^{\circ}) = \int_{\mathbb{S}(1)} \pi(x) \mathrm{d}\rho(x).$$

The parameter of interest is  $M = \max (d^{-1/2}, M(K^{\circ}))$  and dimension dependence in the stochastic case is controlled by  $Md^2$ .

- (a) Without any assumption on K you can take  $M=d^{-1/2}$ , but the algorithm may be computationally inefficient.
- (b) Given access to sampling and membership oracles for K you can take
- $M=d^{-1/4}$  and the algorithm is efficient. (c) Given access to sampling and membership oracles for a symmetric K you can take  $M=d^{-1/2}$  and the algorithm is efficient.

#### References & Link

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