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3. Probability

Stephen Covey, 7 Habits of Highly Effective People

Habit 2: Begin with the end in mind.

What lies behind us and what lies ahead of us are tiny compared with what lies within us.

Oliver Wendell Holmes

Overview

- We will not repeat the lecture material here, but move forward using that as a foundation.
- We will start with discrete random variables and their properties, and introduce some standard probability distributions.
- We will then consider continuous random variables and their properties. Some standard distributions will be introduced.

3.1 Discrete random variables

- A *random variable* (rv) is mapping from the sample space of a random experiment to the real numbers.
- A *discrete random variable* is one that takes *countably infinite* values.
- The distribution of a discrete rv X is described by a *probability mass function* (pmf), defined as

$$p_X(x) = P(X = x).$$

- A pdf satisfies

$$p_X(x) \geq 0 \quad \text{and} \quad \sum_{i=1}^{\infty} p_X(x) = 1.$$

3.2 Mean and variance

The *mean* of a random X , denoted $\mathbb{E}(X)$, is defined as

$$\mathbb{E}(X) = \sum_{i=-\infty}^{\infty} x_i p_X(x_i).$$

- ❶ $\mathbb{E}(X)$ is also called the *expectation* or the *mean* of X .
- ❷ We also use the symbol μ_X for $\mathbb{E}(X)$. (μ is the Greek letter mu.)
- ❸ In general, $\mathbb{E}(X^2) \neq [\mathbb{E}(X)]^2$.

Properties of Mean

E1 $\mathbb{E}(c) = c$ for any constant c .

E2 $\mathbb{E}(aX + b) = a \mathbb{E}(X) + b$ for any constants a and b .

E3 $\mathbb{E}(X + Y) = \mathbb{E}(X) + \mathbb{E}(Y)$. (The mean of a sum is the sum of the means.)

E4. $\mathbb{E}(X - \mu_X) = \mathbb{E}[X - \mathbb{E}(X)] = 0$. (Mean shifting.)

Proof

3.3 variance

For a random variable X , we define the *variance* of X , denoted $\text{Var}(X)$, by

$$\text{Var}(X) = \mathbb{E} \left[(X - \mu_X)^2 \right] = \mathbb{E} \left[(X - E(X))^2 \right].$$

We also write σ_X^2 for $\text{Var}(X)$.

The *standard deviation* of X , denoted σ_X , is the square root of variance, that is,

$$\sigma_X = \sqrt{\text{Var}(X)}.$$

Properties of Variance

V1. $\text{Var}(X) \geq 0$. (Variance is non-negative.)

V2. $\text{Var}(X) = \mathbb{E}(X^2) - [\mathbb{E}(X)]^2$. (This is simpler for calculating variance.) (Compare with $\sum_{i=1}^n (x_i - \bar{x})^2 = \sum_{i=1}^n x_i^2 - n\bar{x}^2$.)

V3. $\text{Var}(aX + b) = a^2 \text{Var}(X)$. (Scaling.) (Compare with $u_i = ax_i + b \Rightarrow s_u = a^2 s_x^2$.)

Proof

V1.

$$\text{Var}(X) = \mathbb{E} \left[(X - \mu_X)^2 \right] = \sum_x \underbrace{(x - \mu_X)^2}_{\geq 0} \underbrace{p_X(x)}_{\geq 0} \geq 0,$$

as each term in the sum is non-negative (≥ 0).

Proof (ctd)

V2.

$$\begin{aligned}\text{Var}(X) &= \mathbb{E} \left[(X - \mu_X)^2 \right] \\&= \mathbb{E} [(X - \mu_X)(X - \mu_X)] \\&= \mathbb{E} [X(X - \mu_X) - \mu_X(X - \mu_X)] \\&= \mathbb{E} [X^2 - \mu_X X] - \mu_X \underbrace{\mathbb{E}(X - \mu_X)}_{=0} \\&= \mathbb{E} (X^2) - \mu_X \mathbb{E}(X) \\&= \mathbb{E} (X^2) - \mu_X^2 = \mathbb{E} (X^2) - (\mathbb{E}(X))^2.\end{aligned}$$

Proof (ctd)

V3. See slide 15. The proof is similar, with Σ replaced by \mathbb{E} .

$$\begin{aligned}\text{Var}(aX + b) &= \mathbb{E} \left[(aX + b - \mathbb{E}(aX + b))^2 \right] \\ &= \mathbb{E} \left[(aX + b - a\mathbb{E}(X) - b)^2 \right] \\ &= \mathbb{E} \left([a(X - \mu_X)]^2 \right) \\ &= \mathbb{E} \left[a^2 (X - \mu_X)^2 \right] \\ &= a^2 \mathbb{E} \left[(X - \mu_X)^2 \right] \\ &= a^2 \text{Var}(X).\end{aligned}$$

Note that the standard deviation of $aX + b$, denoted σ_{aX+b} , is $|a| \sigma_X$.

Standardised random variable

Let the random variable X have mean μ_X and standard deviation σ_X . Put

$$Z = \frac{X - \mu_X}{\sigma_X} = \underbrace{\frac{1}{\sigma_X}}_a X - \underbrace{\frac{\mu_X}{\sigma_X}}_b.$$

Then

$$\mathbb{E}(Z) = \mathbb{E}(aX - b) = a \mathbb{E}(X) - b = \frac{1}{\sigma_X} \mu_X - \frac{\mu_X}{\sigma_X} = 0,$$

$$\text{Var}(Z) = \text{Var}(aX - b) = a^2 \text{Var}(X) = \frac{1}{\sigma_X^2} \text{Var}(X) = \frac{1}{\sigma_X^2} \times \sigma_X^2 = 1.$$

We call Z a *standardised random variable*. This result is similar to that for standardising data (see §1.8):

$$z_i = \frac{x_i - \bar{x}}{s} \Rightarrow \bar{z} = 0 \text{ and } s_z^2 = 1.$$

Note

$$\begin{aligned} Y &= aX + b \\ \Rightarrow \text{Var}(Y) &= a^2 \text{Var}(X) \\ \text{and } \sigma_Y &= |a| \sigma_X, \end{aligned}$$

since standard deviation is always non-negative. This is similar to the result for linear transformation of data. Note that

$$|a| = \begin{cases} a & \text{if } a \geq 0, \\ -a & \text{if } a < 0. \end{cases}$$

Facts

$$\sum_{i=0}^{\infty} x^i = \frac{1}{1-x}$$
$$\sum_{i=1}^{\infty} ix^{i-1} = \frac{1}{(1-x)^2}.$$

3.3 Some standard distributions

Bernoulli distribution

A random variable $X \sim \text{Bern}(p)$ if

$$p_X(x) = p^x(1-p)^{1-x}, \quad x = 0, 1.$$

Then

$$\mathbb{E}(X) = p, \text{Var}(X) = p(1-p).$$

Geometric distribution

Consider a sequence of Bernoulli trials, and let the random variable X denote the trial on which the first success occurs. Then $X \sim \text{Geom}(p)$ where p is the probability of success, with probability mass function

$$p_X(x) = p(1 - p)^{x-1}, \quad \text{for } x = 1, 2, 3, \dots$$

Now

$$\mathbb{E}(X) = \frac{1}{p} \quad \text{and} \quad \text{Var}(X) = \frac{1 - p}{p^2}.$$

Proof

$$\begin{aligned}\mathbb{E}(X) &= \sum_{x=1}^{\infty} xp(1-p)^{x-1} \\ &= \end{aligned}$$

Proof (ctd)

Geometric distribution: variance

$$\begin{aligned}\mathbb{E}(X^2) &= \sum_{x=1}^{\infty} x^2 p(1-p)^{x-1} \\ &= \end{aligned}$$

Geometric distribution: variance (ctd)

Binomial distribution

Let the random variable X denote the number of successes in n iid Bernoulli trials. Then $X \sim \text{Bin}(n, p)$ where p is the probability of success, with probability mass function

$$p_X(k) = \binom{n}{k} p^k (1-p)^{n-k}, \quad k = 0, 1, 2, \dots, n.$$

Then

$$\mathbb{E}(X) = np, \text{Var}(X) = np(1-p).$$

Proof

Proof(ctd)

Proof(ctd)

Poisson distribution

Let the random variable X denote the number of successes in a fixed volume. Then $X \sim Poi(\lambda)$ where λ is the mean number of occurrences, with probability mass function

$$p_X(k) = \frac{\lambda^k e^{-\lambda}}{k!}, \quad k = 0, 1, 2, \dots$$

Now

$$\mathbb{E}(X) = \lambda \text{ and } \text{Var}(X) = \lambda.$$

Proof

Proof(ctd)

Proof(ctd)