

08/11/2024

- Consider the standard linear model  $y = X\beta + \epsilon$  and  $L\beta = z$

$$\tilde{\beta} = \beta^{LS} - (X^T X)^{-1} L^T (L (X^T X)^{-1} L^T)^{-1} (L \beta^{LS} - z)$$

Lemma:  $P(L) = P(L (X^T X)^{-1} L^T)$

Assumptions: we can take  $L_{m \times p}$  s.t.  $P(L) = m$ .

Also, suppose  $P(X) = p$ .

$$P(L (X^T X)^{-1} L^T) = m$$

$$\tilde{\beta} = \beta^{LS} - (X^T X)^{-1} L^T (L (X^T X)^{-1} L^T)^{-1} (L \beta^{LS} - z)$$

if  $m=1$

$$\tilde{\beta} = \beta^{LS} - \frac{(X^T X)^{-1} L^T (L \beta^{LS} - z)}{(L (X^T X)^{-1} L^T)}$$

- $y_i = \theta_i + \epsilon_i \quad \forall i=1(1)4$

$$\theta_1 + \theta_2 + \theta_3 + \theta_4 = 0$$

$$L = (1, 1, 1, 1)$$

$$L\theta = 0 \Rightarrow (1, 1, 1, 1) \begin{pmatrix} \theta_1 \\ \theta_2 \\ \theta_3 \\ \theta_4 \end{pmatrix} = 0$$

$$X = \begin{pmatrix} 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{pmatrix}$$

$$\theta^{LS} = (X^T X)^{-1} X^T y = y$$

$$\tilde{\theta} = y - \begin{pmatrix} \bar{y} \\ \bar{y} \\ \bar{y} \\ \bar{y} \end{pmatrix}$$

$$\cdot \quad y_{ij} = \mu + \theta_i + \epsilon_{ij} \quad \forall i = 1(1)k \quad j = 1, \dots, n_i$$

$$\text{the restriction is } \theta_k - \theta_l = 0 \quad \forall k \neq l$$

$$\min_{\beta} (y - X\beta)^T (y - X\beta) \quad \text{subject to } L\beta = z$$

$$(y - X\beta)^T (y - X\beta)$$

$$= \sum_{i=1}^k \sum_{j=1}^{n_i} (y_{ij} - \mu - \theta_i)^2$$

$$\mu + \theta_i \stackrel{LS}{=} \bar{y}_i = \frac{1}{n_i} \sum_{j=1}^{n_i} y_{ij}$$

$$\left\{ \begin{array}{l} \sum_{i=1}^n (y_i - \theta)^2 \\ \mu \text{ minimized} \\ \text{when } \theta = \bar{y} \end{array} \right.$$

$$\theta_i = \theta_j \quad \forall i \neq j$$

$$\sum_{i=1}^k \sum_{j=1}^{n_i} (y_{ij} - \mu - \theta)^2$$

$$\mu + \theta \stackrel{LS}{=} \bar{y} = \frac{1}{k} \sum_{i=1}^k \left( \frac{1}{n_i} \sum_{j=1}^{n_i} y_{ij} \right)$$

$$\cdot \quad p(x) = n, \quad p = n \quad \text{and} \quad p(x) = n$$

$$\beta^{LS} = (X^T X)^{-1} X^T y = X^{-1} y$$

$$RSS = (y - X\beta^{LS})^T (y - X\beta^{LS}) = 0.$$

• Fisher-Cochran Theorem (Matrix Theoretic Version)

Theorem: let  $A_1, \dots, A_k$  be  $n \times n$  s.t.  
 $\sum_{i=1}^k A_i = I_n$ . Then the following conditions  
 are equivalent.

(i)  $\sum_{i=1}^k \text{rank}(A_i) = n$

(ii)  $A_i^2 = A_i \quad \forall i=1(1)k$

(iii)  $A_i A_j = 0 \quad \forall i \neq j$

Proof: See any Linear Algebra Book.

$S + T = \mathbb{R}^n$   
 projection into  
 $S$  along  $T$   
 $P$  is a projection  
 matrix

Multivariate Normal distribution

• let  $u_1, \dots, u_n$  be standard Normal  
 variables which are independent

$u_i \sim N(0, 1)$  and  $u_i$ s are independent.

$$u = \begin{pmatrix} u_1 \\ \vdots \\ u_n \end{pmatrix} \sim N(0, I_n)$$

Consider  $X_{n \times n}$  and a vector  $M_{n \times 1}$

$$y = Xu + M$$

Then  $y$  follows a multivariate Normal distribution

$$E(y) = M$$

$$\text{var}(y) = X X^T = \Sigma$$

$$f_y(w) = \frac{1}{(\sqrt{2\pi})^n \sqrt{|\Sigma|}} e^{-\frac{1}{2}[(w-\mu)^T \Sigma^{-1} (w-\mu)]}$$

well exist iff  $\Sigma$  is of full row rank.

- For a random vector  $y$ , Characteristic function  
 $E(e^{it^T y}) \quad t \in \mathbb{R}^n.$