

Modeling frequency-conditioned paradigm uniformity in Japanese voiced velar nasalization*

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Abstract

Recent quantitative work on the variable [g]-[ŋ] alternation in compounds of certain dialects of Japanese has revealed *token frequency* of the compound as a whole, and of the compound's second-member (N2) in its freestanding form, to be important predictors of the alternation. In this paper, we propose a formal phonological analysis that integrates usage-based factors like frequency with the action of the phonological grammar, extending mechanisms of lexicon-grammar interaction previously proposed in the context of Lexical Conservatism. We demonstrate that our model fits the experimental data better than—or at least comparably to—a theoretically-naïve statistical model proposed in the previous work. Based on the success of our modeling, we discuss the role of token frequency in phonological patterning more broadly, and how the mechanism that we propose might be extended to unify a range of contradictory frequency-dependent processes that have been observed in the literature.

1 Introduction

This paper is about how to integrate information about *usage frequency*—here, the token frequency of morphemes in the language experience of an individual speaker—into a constraint-based phonological grammar formalism that characterizes that speaker's generative linguistic knowledge.

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19 We take as our empirical case the frequency-conditioned variability in optional paradigm uni-
20 formity in voiced velar nasalization (henceforth “nasalization”) in phonologically-conservative
21 Japanese dialects, recently studied using corpus data by Breiss et al. (2021b) and experimentally
22 verified in Breiss, Katsuda and Kawahara (*to appear*) — henceforth BKK. These studies are the
23 latest in a long research tradition centered on the allophonic distribution of /g/ in conserva-
24 tive Japanese dialects, where a static phonotactic restriction enforces /g/ to be realized as [g]
25 prosodic-word-initially and [ŋ] elsewhere (e.g. Kindaichi 1942; Trubetskoy 1969; Labrune 2012).
26 This correspondence is disrupted in compounds with /g/-initial second member (N2) that can oc-
27 cur as a free morpheme: in compounds with N2s that do not occur as free-standing words, the
28 /g/ → [ŋ] alternation is exceptionless, but in compounds where N2 may additionally occur as a
29 free-standing word (that is, with initial [g]) the nasalization process is optional (Ito and Mester,
30 1996, 2003).

31 The contribution of recent work by BKK (reviewed in section 3) is to characterize this variation
32 in quantitative detail, and in particular to highlight how the token frequency of both the com-
33 pound and the free N2 impact the outcome of optional nasalization: higher frequency compounds
34 encourage more nasalization of medial /g/ to [ŋ], while higher frequency free N2s encourage more
35 retention of medial /g/ as [g], remaining uniform across the paradigm of their free-standing forms
36 and compound forms (Steriade, 2000; Benua, 2000).

37 The novel contribution of this paper is to provide a formally explicit model of the experimental
38 data. The model builds upon the Voting Bases model of lexicon-grammar interaction (Breiss,
39 2024), originally proposed to model Lexical Conservatism (Steriade, 1997). Lexical Conservatism
40 is a type of paradigm uniformity where the distribution of stem allomorphs (referred to as “bases”)
41 in a paradigm influences the way that paradigm accommodates new members. The canonical
42 example comes from Steriade (1997), who observed that the phonologically-similar forms *rémedy*
43 and *párody* differ in their behavior when affixed with *-able*, yielding *remédiable* with shifted stress,
44 but *párodiable*, with fixed stress. She argued that this difference stems not from the forms *remedy*
45 and *parody* themselves, but from the fact that *remedy* has a stem allomorph *remédi-* in *remédiál*
46 that satisfies the marked lapse arising from affixation.

47 Breiss (2024) examined the same Lexical Conservatism dependency using novel derived forms
48 (like *lábor* + *-able*, with related form *labórious*, and *pláster* + *-able* with no phonologically-
49 advantageous related form), and found that in experimental settings, speakers are sensitive not
50 only to the *presence* of the phonologically-beneficial stem allomorph (like *remédiál* and *labórious*),
51 but also to its salience in the lexicon as manipulated by priming. To account for these data, he
52 proposed a formal phonological model that integrates the influence of the contents of the lex-
53 icon along with their resting activation, enabling the phonological grammar to be sensitive to the
54 psycholinguistic properties of the morphemes which it manipulates. Breiss (2024) termed this

55 formal model of lexicon-grammar interaction the *Voting Bases* model.

56 In this paper, we demonstrate that the Voting Bases model extends, without modification, to
57 the separate case of lexicon-grammar interaction found in Japanese nasalization. The success
58 of the model suggests that the foundational principles of the Voting Bases model may be a good
59 candidate for a general theory of the way that the lexicon and grammar interact. This finding also
60 underscores the explanatory value to be gained for phonological phenomena by adopting a more
61 psycholinguistically-nuanced portrait of the lexicon as a dynamic substrate that can influence the
62 computations of the grammar on the items which it contains. In section 6.3 we take up a series of
63 questions which arise when adopting this boundary-blurring approach, in light of the traditional
64 dichotomy between generative and usage-based perspectives on linguistic data.

65 The layout of the paper is as follows: the first two sections of the paper review in some
66 depth basic facts about Japanese nasalization drawn from the literature (section 2), and then
67 specifically reviews in detail Experiment 1 of BKK (section 3). Though this may not constitute new
68 information, we hope the reader will find its inclusion helpful in contextualizing the theoretical
69 analysis. The following section, 4, focuses on the Voting Bases model, and how we apply it
70 to the context of optional paradigm uniformity. Section 5 then actually fits the model to the
71 experimental results, and discusses relative and absolute model fit in comparison to minimally-
72 different models that incorporate only some of the assumptions of the Voting Bases model. The
73 paper closes in section 6 with a discussion of broader issues, touching on how such a system
74 might come to be in the mind of the learner, on the merits of a joint model of psycholinguistic
75 and grammatical influence on word formation, and on what a unified theory of token frequency
76 effects on the phonological grammar might look like.

77 2 The traditional picture of Japanese nasalization

78 The data that we model in this paper comes from Experiment 1 of BKK, which investigated the
79 variation between [g] and [ŋ] induced by the phonotactics of phonologically-conservative di-
80 alects of Japanese. The pattern, which has been well-studied in both descriptive (Kindaichi, 1942;
81 Trubetskoy, 1969; Hibiya, 1995) and generative (Labrune, 2012; Ito and Mester, 1996, 2003) lit-
82 erature on Japanese linguistics, is exemplified in the complementary distribution of [g] and [ŋ]
83 shown in the monomorphemic data in example (1) below, where the voiced oral velar stop is only
84 permitted word-initially, and the velar nasal is only permitted word-medially.

85 (1) a. /kaŋami/ → [kaŋami]

86 “mirror”

87 b. /gake/ → [gake]

88 “cliff”

89 We assume throughout that non-alternating forms are stored surface-true as URs in the lexi-
90 con, in accordance with the phonological tradition of (Strong) Lexicon Optimization (Prince and
91 Smolensky, 1993; Sanders, 2003). This stance is supported by psycholinguistic research on the
92 contents of lexical representations, reviewed in section 4.1.

93 Japanese’s extensive use of compounding in word-formation gives the opportunity for the
94 phonotactic restriction to drive alternations, seen in examples (2)-(5) below. Here we see that
95 when a /g/-initial morpheme is word-initial (either as a prosodically-free word, in examples (2)–
96 (4), or as the first member (N1) of a compound, in example (5)¹), it is realized with an initial [g],
97 while when it occurs as the second member of a compound (N2) it is realized with initial [ŋ].
98 Critically for the current study, Ito and Mester (1996) observed that although in all cases the /g/-
99 initial N2 *may* be realized word-medially with initial [ŋ], nasalization is optional when the N2
100 can stand on its own as a prosodically-free form (cf. the “b” series in (2)–(4) vs. (5c))—a case of
101 optional paradigm uniformity.

- 102 (2) a. /hai + gan/ → [hai-ŋan] ~ [hai-gan]
lung cancer

“lung cancer”

- 104 b. /gan/ → [gan]

cancer

“cancer”

- 107 (3) a. /noo + geka/ → [noo-ŋeka] ~ [noo-geka]
brain surgery

“brain surgery”

- 109 b. /geka/ → [geka]

surgery

“surgery”

- 112 (4) a. /doku + ga/ → [doku-ŋa] ~ [doku-ga]
poison moth

“poison moth”

- 114 b. /ga/ → [ga], “moth” (a free-standing morpheme)

- 115 (5) a. /doku + ga/ → [doku-ŋa], *[doku-ga]
poison fang

“poison fang”

¹We temporarily adopt here for the traditional assumption that the [g]-initial form of a free N2 is underlying, for expository ease and continuity with the previous literature. Our own proposal is laid out in section 4.

- 117 b. /ga + zoo/ → [ga-zoo]
 fang castle
 “main castle”
- 118 c. /ga/ → *[ga] (a bound morpheme)
 fang
 “fang”
- 119
- 120

121 Breiss et al. (2021b) examined this variation in a corpus derived from a pronunciation dic-
 122 tionary (NHK, 1993) and found that among compounds with free N2s, the two most prominent
 123 predictors of whether an item would be nasalized was the frequency of the N2’s free [g]-initial
 124 form, and the frequency of the whole compound. These effects ran in opposite directions: higher
 125 frequency compounds were more likely to be nasalized (the left facet of Figure 1); on the other
 126 hand, the more frequent the free N2, the less likely the nasalization was (the right facet of Figure
 127 1).

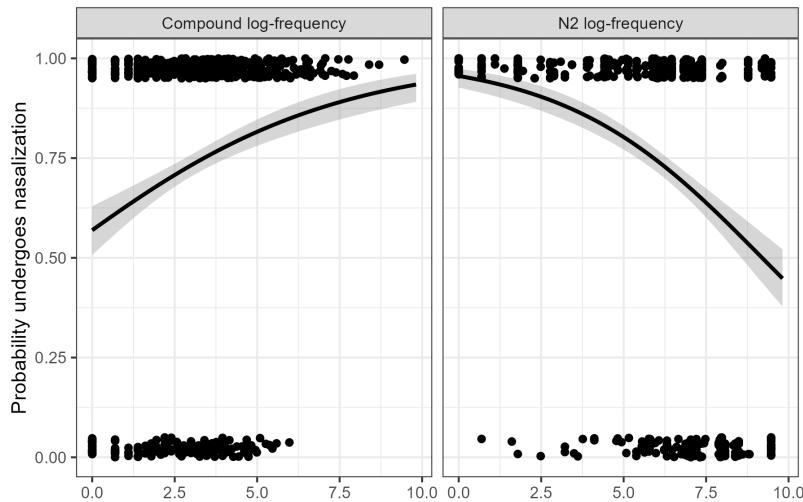


Figure 1: The effects of whole compound frequency (left facet) and N2 frequency (right facet) on the probability of nasalization (vertical axis), with binomial smooths in the corpus data. One dot represents one lexical item; vertical jitter has been added for readability. Figure and caption adapted with permission from Breiss et al. (*to appear*), data from Breiss et al. (2021b).

128 The corpus data was modeled as a case of probabilistic paradigm uniformity in Breiss et al.
 129 (2021a) using Output-Output Faithfulness constraints (Benua, 2000) indexed to items binned by
 130 the relative frequency of each compound and N2. The paper was limited, however, by the untested
 131 assumption of their model that the frequency-modulation of paradigm uniformity in their cor-
 132 pus data actually represents the synchronic knowledge of speakers. Additionally, their formal
 133 model was not explicitly informed by psycholinguistic considerations and thus its linking hy-
 134 pothesis between frequency (necessarily a lexical characteristic) and the phonological grammar

had a problem of simply being stipulative—in other words, there was nothing in their model that prevented the opposite relation between frequency and paradigm uniformity from holding.

In this paper, we offer two improvements on the state of affairs in Breiss et al. (2021a). First, we model experimental data from Breiss et al. (*to appear*) (BKK) where the frequency-conditioning of the variable paradigm uniformity is reproduced in existing compounds and extended to novel ones. Second, we do this by extending the Voting Bases model of Breiss (2024) which is compatible with consensus understanding of the way lexical frequency is connected to the lexical representation and activation, and which offers an explicit linking hypothesis relating the real-time dynamics of the lexicon to the representation and computations of the phonological grammar.

3 BKK’s Experiment 1

BKK carried out two experiments on Japanese nasalization, with the goal of seeing whether the corpus patterns were representative of speakers’ generalizable knowledge, both in the aggregate and as individuals. They found that both individually and in aggregate, speakers’ propensity to nasalize displayed sensitivity to the frequency of the free N2 and compound, in existing and novel compounds. In this paper, we focus our modeling efforts on the results of their Experiment 1, which we describe in some detail below.²

BKK’s stimuli were roughly balanced between existing Japanese compounds of varying frequencies, and novel (i.e., zero-frequency) semantically-compositional Sino-Japanese compounds. Both existing and novel stimuli had attested free N2s of a range of frequencies. Out of a desire to sample compounds with a wide range of frequencies that would likely be known to participants, existing compounds ranged from two to eight moras in length, while all novel compounds were four moras long. Complete details of the experimental materials are available from BKK’s OSF repository³.

BKK recruited speakers of the phonologically-conservative Tōhoku dialect of Japanese, and used a short dialect questionnaire to ensure that their speech exhibited the allophonic distribution of word-initial [g] and word-medial [ŋ]. For the purposes of the model which we develop, we will see that these monomorphemic words provide crucial evidence for the lower bound of the weight of the markedness constraint driving nasalization, since with data from compounds alone, it is not uniquely identified against the background of faithfulness constraints that the Voting Bases model uses (see section 5 for further details).

²They also sought to determine whether correlation between nasalization and the overall prosodic size of the compound, which is observed in the corpus (Breiss et al., 2021b) but is a typologically unusual pattern, was replicated in participants’ online productions (Experiment 2). They actually found that there was no evidence of a direct relationship between nasalization and global prosodic length (cf. Jiang 2023). We therefore do not address this experimental data here, as our point is made in the simpler case of data from Experiment 1.

³https://osf.io/avnpw/?view_only=cd2afdcc183f4de3ac1261b4af66f08d

165 The dialect questionnaire consisted of a production task where speakers were asked to read
166 aloud 10 monomorphemic words with word-initial [g] of varying frequencies, and 10 monomor-
167 phemic words with word-medial [ŋ]. The stimuli were written with *kanji* orthography, which
168 does not distinguish between [g] and [ŋ]—this is also true of the main production experiment
169 described below, so we follow BKK in assuming that the participants’ production was not influ-
170 enced by orthographic factors. The twenty words were shown to the participant in a random
171 order, and their productions were recorded; only the eight participants who exhibited the target
172 pattern of allophony in all monomorphemes were invited to participate in the main experiment.

173 After this knowledge check, participants saw each compound one at a time in a random order,
174 and produced the form aloud while their speech was recorded. Participants also produced and
175 indicated knowledge of all of the free N2s in the experiment, as well as all of the compounds. See
176 Breiss et al. (*to appear*) for complete details.

177 3.1 Results

178 BKK found that the participants reflected at an individual level the frequency-conditioned vari-
179 ability seen in the corpus study of Breiss et al. (2021b). In existing compounds (Figure 2), their
180 productions were influenced by both the frequency of the compound (the left facet), for which
181 higher values correlated with more nasalization, and by the frequency of the free N2 (the right
182 facet), where higher values correlated with less nasalization.

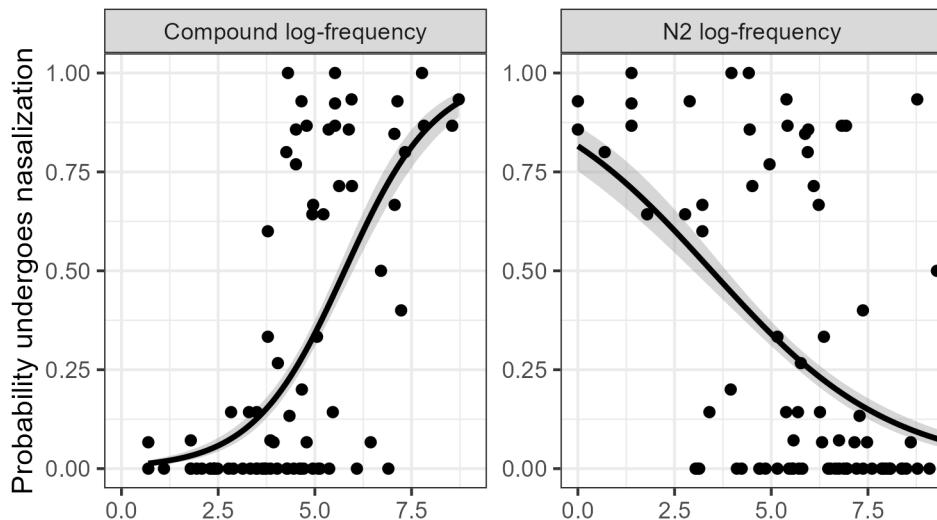


Figure 2: Probability of nasalization (the vertical axis) plotted against compound log-frequency (the left facet) and N2 log-frequency (the right facet) in existing words, with binomial smooths for readability, in the experiment by BKK. Plot and caption reproduced with permission from Breiss et al. (*to appear*).

183 Figure 3 plots the same effect of N2 frequency in novel compounds: forms with higher-
184 frequency N2s were less likely to undergo nasalization relative to those with lower-frequency
185 N2s.

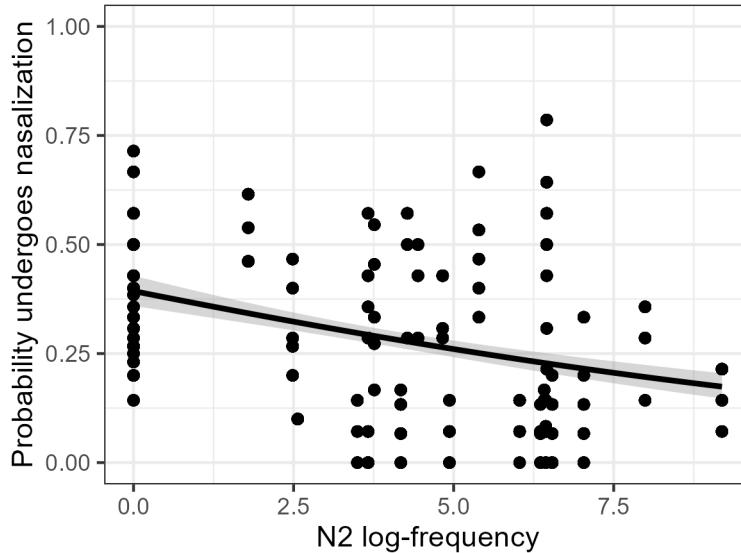


Figure 3: The probability of undergoing nasalization in novel compounds, plotted against N2 log-frequency, with a binomial smooth to aid readability. Plot and caption reproduced with permission from Breiss et al. (*to appear*).

186 Finally, BKK found that the frequency effect was stable at the level of the individual, across
187 existing and novel compounds, which is plotted in Figure 4. In this figure, the horizontal axis plots
188 the strength and direction of the effect of N2 log-frequency in novel compounds, and the vertical
189 axis plots the strength and direction of the effect of N2 log-frequency in the existing compounds;
190 see the caption of Figure 4 for further details. Although different participants were more or less
191 sensitive to the frequency of a given N2, lying higher or lower on each axis, there was uniformity
192 in this degree of sensitivity such that the two co-varied along a diagonal line through the center
193 of the plot. BKK interpreted this correlation as evidence that morpheme usage frequency and
194 phonological markedness have separable, distinct influences on speaker productions.

195 3.2 Summary and goals for modeling

196 To summarize, the findings of BKK that are relevant for the modeling task of this paper are
197 the following. Among those speakers for whom the phonotactic restriction enforcing [g]~[ŋ]
198 allophony was exceptionless in monomorphemic words:

- 199 1. Phonotactically-driven nasalization is variable in compounds with free N2s.

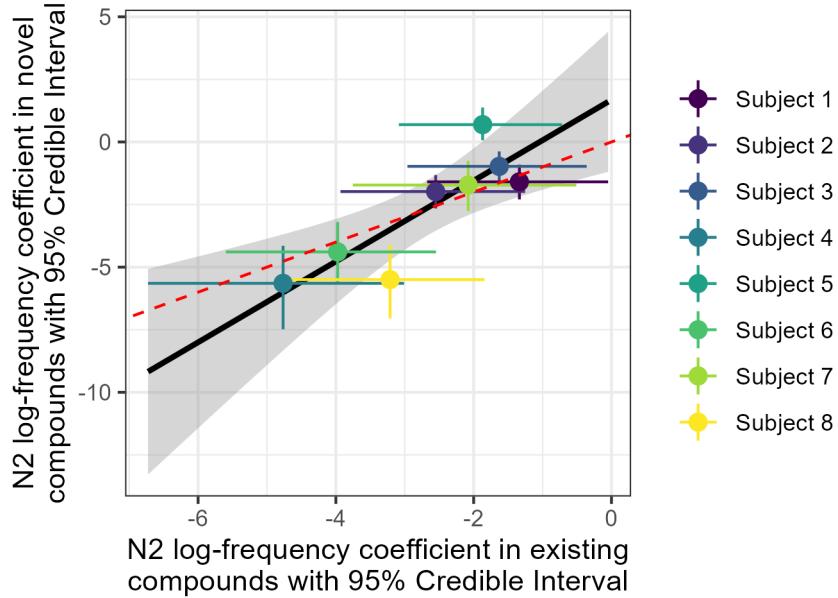


Figure 4: The coefficient of N2 log-frequency in novel compounds, derived from the model in Table 1 of Breiss et al. (*to appear*), is plotted on the horizontal axis, and the coefficient for N2 log-frequency in existing compounds, derived from the model summarised in Table 3 of Breiss et al. (*to appear*), is plotted on the vertical axis. Points represent median values of the posterior with ranges encompassing the 95% Bayesian Credible Intervals, colors represent speakers, and a linear smooth has been added for readability, with the line of slope 1 intersecting the origin in dotted red. Plot and caption adapted with permission from Breiss et al. (*to appear*).

- 200 2. In these compounds, the probability of nasalization is increased by higher compound frequency, and decreased by higher N2 frequency.
- 201
- 202 3. The frequency effect is uniform within individuals across existing and novel compounds.
- 203 Below, we propose a formal model of these facts, using the Voting Bases model to relate a lexicon
- 204 containing usage-frequency information to a phonological grammar couched in the Maximum
- 205 Entropy (MaxEnt) framework.⁴

206 4 Modeling token frequency in the phonological grammar

- 207 Based on the facts laid out above, we seek a model of the phonological grammar that allows
- 208 non-phonological properties of individual lexical items (here, frequency) to influence their par-

⁴We do not attempt to model the frequency of the first compound member, N1, on the probability of nasalization in compounds, since this was not manipulated by BKK. Future work might profitably pursue this question experimentally and formally, since corpus data in Breiss et al. (2021b) suggests that higher N1 frequency may also independently lower the probability of nasalization; see Rebrus and Törkenczy (2017) for a similar finding of N1 frequency on compound coherence in Hungarian vowel harmony.

209 ticipation in phonological processes (here, paradigm uniformity). Note that we specifically aim
210 to model phonological and non-phonological influences on the outputs of the phonological gram-
211 mar, rather than any possible morphological or paradigmatic effects on phonetic realization (see
212 Purse et al. 2022 for a review), about which the Voting Bases model as laid out in Breiss (2024)
213 makes no predictions.

214 4.1 The contents of a lexical entry

215 As prolegomena to the grammatical model, it will be important to establish some relevant context
216 regarding the contents of the lexicon, because it is these representations that are at stake in
217 discussions of token frequency. Psycholinguistic research has amassed a large body of evidence
218 that the lexicon is richly structured, with numerous types of linked representations of various
219 levels of detail grouped under the same lexical entry. We do not review this research in depth
220 here, but simply highlight the findings relevant to developing the type of integrated phonological
221 theory referenced above. For a thorough discussion and literature review on the (phonologically-
222 relevant) contents of a lexical entry, see Pierrehumbert (2016); for more on how this information
223 interacts with the Voting Bases theory in cases beyond those relevant for the nasalization, see
224 Breiss (2021, 2024).

225 Since nasalization concerns paradigm uniformity, we assume the lexical entry for an existing
226 word lists (among many other things) their allomorphs (cf. Strong Lexicon Optimization em-
227 braced by Sanders 2006, as well as arguments by Wang and Hayes 2025 on the sufficiency of
228 less-abstract URs): for a non-alternating monomorpheme like [kanjami] “mirror”, this would be
229 simply /kanjami/; for a monomorpheme that can appear as an N2 and undergo nasalization, such
230 as [ga]–[ŋa] “moth”, the lexical entry would list both /ga/ and /ŋa/. Finally, we assume that ex-
231 isting compounds are stored whole, with nasalization applied so as to respect the phonotactic in
232 the lexicon (Albright, 2008; Martin, 2007).⁵

233 With regard to non-phonological characteristics of the lexicon, we follow a large body of
234 evidence that lexical representations have differing degrees of salience or strength of encoding.
235 Following Breiss (2021, 2024), we refer to this quantity as *resting activation*, borrowing the term
236 (though not the theory) from Morton (1970), which corresponds to the strength of a memory
237 representation itself, not a number or rank stored in long-term memory as a characteristic of the
238 lexical item. Thus, characteristics (long-term or dynamic) of lexical items like their frequency,
239 and whether or not they were recently activated (for example, by priming), all contribute dynam-

5On the suggestion of a reviewer, we relaxed this assumption by fitting a comparable model but assuming stored allomorphs for both oral and nasal forms of the compound, with corresponding faithfulness constraints for each. Such a model returns weights and fits to the data identical to the one without the relevant faithfulness constraint, indicating that it is thus at best superfluous in explaining the data. This exercise shows that our assumption here is well-founded, or at least benign. Details of the model fit can be found in the supplementary materials.

240 ically to an item’s resting activation. Importantly, also following Breiss, we use the term “resting
241 activation” as a stand-in for any scalar summary statistic that can be derived from an imple-
242 mented model of lexical dynamics. We remain intentionally agnostic as to the specific model of
243 these dynamics, whether the specific model endorsed by Morton (1970) or not, simply stressing
244 that so long as such a model can be used to drive a measure of relative salience influenced by the
245 factors just mentioned, the Voting Bases model can make reference to it to scale faithfulness con-
246 straint violations (cf. e. g. Luce and Pisoni, 1998). We discuss how resting activation is modeled
247 as influencing the phonological grammar below in section 4.4.

248 4.2 The Voting Bases model

249 We now turn to a formal phonological model of the Japanese nasalization data. We use the Voting
250 model of Base competition proposed in Breiss (2021, 2024). The Voting model has been used to
251 model data in Lexical Conservatism in English and Spanish, and is broadly compatible with the
252 view of the lexicon laid out above. Here, we extend the scope of the model by analyzing the
253 probabilistic paradigm-uniformity found in Japanese nasalization.

254 The Voting Bases model has two parts: the first is that all listed stem allomorphs in the lexi-
255 con exert an analogical pull on derivatives (operationalized using allomorph-specific faithfulness
256 constraints), violations of which are scaled in proportion to the resting activation of the repre-
257 sentation to which faithfulness is being assessed. We note that the terminology of “bases” comes
258 from the original context for which the model was developed, but here the term can be read
259 as a synonym for “stored allomorph”.⁶ The second part is that markedness constraints evaluate
260 candidates in the standard way for any constraint-based phonological models.

261 The Voting Bases model assumes a probabilistic, weighted-constraint phonological grammar;
262 here, we use MaxEnt Harmonic Grammar (Smolensky, 1986a; Goldwater and Johnson, 2003), but
263 in principle we could also use another grammar formalism that has these characteristics, like
264 Stochastic (or Noisy) Harmonic Grammar (Boersma and Pater, 2016). We use MaxEnt since it has
265 various strengths; e.g. it directly relates Harmony to probability (Hayes, 2022), permits constraint
266 cumulativity by default (Jäger and Rosenbach, 2006; Breiss, 2020), has a learning algorithm to set
267 its weights, and is rooted in well-understood statistical techniques used widely outside linguistics
268 (Jurafsky and Martin, 2009, ch. 5). We stress, however, that our analyses can be recast in terms
269 of other stochastic constraint-based frameworks.

⁶The probabilistic paradigm-uniformity might, as a reviewer points out, be captured in terms of Output-Output faithfulness (Benua, 2000) instead of the Voting Bases model. This approach was pursued in Breiss et al. (2021a), but ultimately we abandon it here because it fails to explain the correlation between the degree of faithfulness to a non-local paradigm form and the relative frequency of the two forms in question. In the Voting Bases model, this relationship has a clear source by virtue of the explicit location of both URs in a psycholinguistically-dynamic lexicon; for a more extended comparison between these two approaches, see discussion in Breiss (2024, fn. 5).

270 **4.3 Constraints**

271 In the analysis developed in this paper, we adopt the general approach of Ito and Mester (1996,
272 2003), following loosely Breiss et al. (2021a). We only use three constraints: a single marked-
273 ness constraint to motivate nasalization (extending the spirit of the constraint *VgV from Ito and
274 Mester 2003 to be compatible with nasal-final N1s, which pattern identically to vowel-final N1s),
275 and a pair of faithfulness constraints which correspond to the second member’s free form and to
276 the analogical pull of the compound as a whole, if one exists. They are listed below.⁷

- 277 • ***INTERNAL-[g]**: Assign one violation for each word-internal [g].
- 278 • **ID-[nasal]-N2**: Assign one violation for each segment in the listed allomorph for the free-
279 standing N2 that does not match its corresponding segment in the feature [nasal].
- 280 • **ID-[nasal]-COMPOUND**: Assign one violation for each segment in the listed allomorph for
281 the full compound that does not match its corresponding segment in the feature [nasal].

282 Note that the constraint definitions do not make reference to scaling or the contents of the
283 lexicon; the proposal in the Voting Bases model is an architectural proposal about how psycholin-
284 guistic, “extra-grammatical” factors act within and beside the phonological grammar to influence
285 certain variable phenomena.

286 **4.4 Modeling resting activation**

287 The discussion in 4.1 above left open how a specific numerical value for resting activation might
288 be calculated on the basis of the psycholinguistic characteristics of item’s lexical entry. Here,
289 we model the data using the log-frequency of the allomorph, passed through a sigmoid function
290 $\frac{1}{1+e^{-\log freq}}$ that translates the linear predictor (i.e. $-\log freq$) into the bounded interval of {0,1},
291 which will be the scaling factor applied to faithfulness violations. This is illustrated in Figure 5.
292 The effect of this non-linear transformation will be to preserve the idea that it is less penalized to
293 be unfaithful to low-frequency lexical items compared to higher-frequency ones, while damping
294 down the difference between extreme values of the scale and rendering it bounded.

295 The final move we make here is rather than using *raw* log-frequencies, we use *scaled and*
296 *centered* log-frequencies, following the statistical analysis in BKK. This corresponds to the notion
297 that it is not so much the *absolute* frequency of each item that is important, but how frequent it
298 is relative to the other competitor items in the lexicon (here approximated by the population of

⁷The first faithfulness constraint plays the same role as faithfulness to the Remote Base in an analysis of Lexical Conservatism. The second faithfulness constraint parallels faithfulness to the Local Base in a Lexical Conservatism analysis (Breiss, 2021, 2024). We use more transparent names here for the sake of clarity, since nothing in the Voting Bases model structurally prioritizes Local Bases over Remote ones.

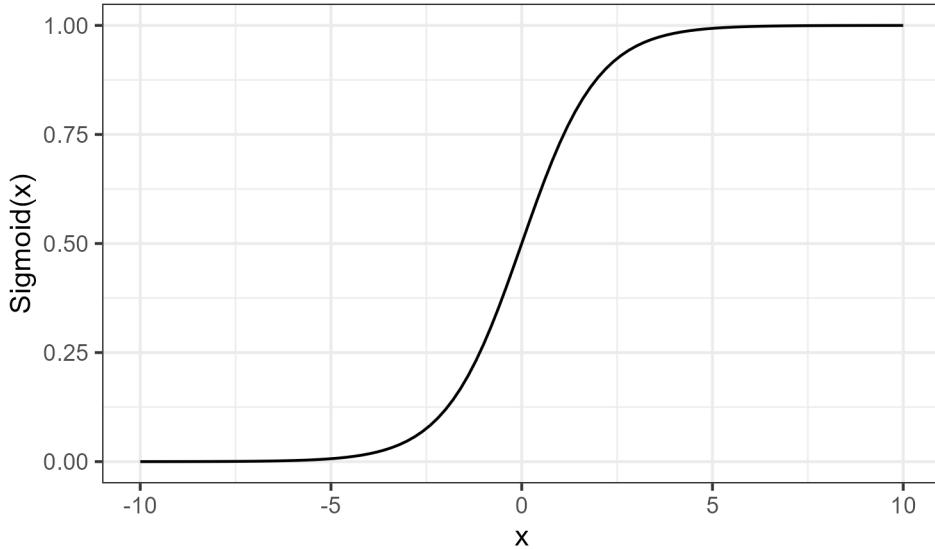


Figure 5: Sigmoid function that translates the (centered) frequencies into the scaling factors. See text for details.

299 items in the experiment), which is in line with previous work on morphological decomposition
 300 in stored forms (Hay, 2001). Finally, in the analysis that we develop below, we do not model the
 301 priming of N2, since BKK did not find substantial evidence that it affected their experimental
 302 data.⁸

303 4.5 Schematic illustrations

304 Before modeling the experimental data itself, it will be useful to work with some toy data to get
 305 a feel for how resting-activation-scaled faithfulness violations interact with the dynamics of a
 306 MaxEnt grammar. First, let us consider the case of novel compounds, since they are the simplest
 307 case to lay out the workings of the analysis. Recall the empirical pattern: here, although the
 308 frequency of the compound is zero, we nevertheless find that nasalization is modulated by the
 309 frequency of N2. Now, consider the case of two hypothetical novel compounds, one with a higher-
 310 frequency N2, and one with a lower-frequency N2, such that when the sigmoid transformation
 311 is applied to their frequencies the higher-frequency form scales its violations of Id-[nasal]-N2 by
 312 0.7, and the lower scales its own violations of the same constraint by 0.3 (these specific numbers

⁸The Voting Bases framework is easily extensible to multiple predictors of resting activation: to incorporate priming, for instance, one could simply treat the term passed into the sigmoid as itself a log-linear model, adding a coefficient (weight) for the effect of priming, in addition to a coefficient for the effect of lexical frequency. This is beyond the current scope of this paper, however, and so we simply assume a fixed coefficient for lexical frequency, since there being only one predictor in the log-linear model for resting activation would make the coefficient of frequency redundant with the weight of the faithfulness constraint being scaled. Similarly extensions of the Voting Bases model could also model by-participant variability in the priming effect using a hierarchical model structure.

313 are chosen purely for the sake of illustration). Using the constraints defined in section 4.3 above,
 314 we can define the tableaux below in Figure 6.

| $/.../N_1, /g.../High-freq.N_2$ | *INTERNAL-[g] | ID-[nas] _{N2} | H | p |
|---------------------------------|---------------|------------------------|----|-----|
| Weight: | 2 | 1 | | |
| a. [...g ...] | 1 | | 2 | .21 |
| b. [...ŋ ...] | | .7 | .7 | .79 |
| $/.../N_1, /g.../Low-freq.N_2$ | *INTERNAL-[g] | ID-[nas] _{N2} | H | p |
| Weight: | 2 | 1 | | |
| c. [...g ...] | 1 | | 2 | .15 |
| d. [...ŋ ...] | | .3 | .3 | .85 |

Figure 6: Schematic application of the Voting model of Base Competition to the formation of a novel compound in the *wug*-test.

315 We can see that the pull of faithfulness to the N2 with higher frequency is stronger than
 316 the one with lower frequency, though both are relatively marginal outcomes since the weight of
 317 *INTERNAL-[g] dominates the distribution of probabilities in this scenario.

318 Moving on to existing compounds, we now must add another item to the lexical entry we are
 319 considering in our left-hand input cell to our tableaux, shown in Figure 7. For the sake of minimal
 320 contrasts, we assume that the frequency of both N2s are equal and medial relative to the examples
 321 in Figure 6 above, allowing us to examine the effect of compound frequency holding N2 frequency
 322 constant. However, in our analysis of the actual data, both scaling factors are independently set
 323 on a per-item basis.

| $/.../N_1, /g.../N_2,$ $/...ŋ.../High-freq.compound$ | *INTERNAL-[g] | ID-[nas] _{N2} | ID-[nas] _{Compound} | H | p |
|---|---------------|------------------------|------------------------------|-----|-----|
| Weight: | 2 | 1 | 1 | | |
| e. [...g ...] | 1 | | .7 | 2.7 | .09 |
| f. [...ŋ ...] | | .5 | | .5 | .91 |
| $/.../N_1, /g.../N_2,$ $/...ŋ.../Low-freq.compound$ | *INTERNAL-[g] | ID-[nas] _{N2} | ID-[nas] _{Compound} | H | p |
| Weight: | 2 | 1 | 1 | | |
| g. [...g ...] | 1 | | .3 | 2.3 | .14 |
| h. [...ŋ ...] | | .5 | | .5 | .86 |

Figure 7: Schematic application of the Voting model of Base Competition to the formation of an existing compound in the *wug*-test.

324 Here we see that the scaling of the compound again depends on frequency, but because of the

assumption we made about the listed form of the compound—specifically, that phonologically well-formed words are preferentially the target of lexicalization (Albright, 2008; Martin, 2007)—we find that the faithfulness to the compound’s UR penalizes the candidate that does not exhibit nasalization and violates markedness.

Finally, we lay out the case where the competition between candidates is driven primarily by faithfulness. Above, where markedness had a high weight, the candidate that satisfied markedness had a higher probability than the one which violated it, and the effects of the faithfulness constraints were on the probability of the minority candidate. In the scenario where markedness is low and the weights of the faithfulness constraints are dominant, the majority candidate is the one that satisfies faithfulness to the whole compound, and the presence of the N2 is the main reason that the unfaithful (but markedness-satisfying) candidate gets appreciable probability; this is a type of “analogical” effect where markedness has little role, as in Figure 8, in which the markedness constraint is assigned a very low weight (here, arbitrarily set as 0.1). Below, we will see that this scenario is most similar to the state of the VVN alternation.

| $/.../N_1, /g.../N_2,$ $/...ŋ.../High-freq.\text{compound}$ | *INTERNAL-[g] | ID-[nas] _{N2} | ID-[nas] _{Compound} | H | p |
|--|---------------|------------------------|------------------------------|-----|-----|
| Weight: | 0.1 | 1 | 2 | | |
| i. [...g ...] | 1 | | .7 | 1.5 | .27 |
| j. [...ŋ ...] | | .5 | | 0.5 | .73 |
| $/.../N_1, /g.../N_2,$ $/...ŋ.../Low-freq.\text{compound}$ | *INTERNAL-[g] | ID-[nas] _{N2} | ID-[nas] _{Compound} | H | p |
| Weight: | 0.1 | 1 | 2 | | |
| k. [...g ...] | 1 | | .3 | 0.7 | .45 |
| l. [...ŋ ...] | | .5 | | 0.5 | .55 |

Figure 8: Schematic application of the Voting model of Base Competition to the formation of an existing compound in the *wug-test*, in a regime where faithfulness is strong and markedness weak.

5 The model in action

Moving on to the analysis itself, we fit a model to the data from existing compounds and monomorphemes, assessing its fit in that setting as well as its generalization to data from novel compounds. We fit the MaxEnt models using the *Solver()* function in Microsoft Excel (Fylstra et al., 1998), and used a weakly informative Gaussian prior of $\text{Normal}(0,10)$ on constraint weights, which has the effect of allowing weights to vary in response to values that best fit the data, while making extreme values (here, above twenty or so) less appealing. For more on priors on weights in MaxEnt

³⁴⁶ phonological models, see Wilson (2006) and White (2017). All models fit in this paper are provided
³⁴⁷ in the supplementary materials.

³⁴⁸ 5.1 Existing compounds

³⁴⁹ We first applied the analysis sketched in section 4.5 to data from existing compounds. Recall that
³⁵⁰ in these forms, compounds with higher-frequency N2s are more likely to resist nasalization than
³⁵¹ those with lower-frequency N2s, but that compound frequency itself also influences nasalization,
³⁵² with higher-frequency compounds favoring the surface-realization of their underlying [ŋ]. We
³⁵³ model the counts of compounds produced having undergone nasalization or not.

³⁵⁴ We also integrate the fact that the participants were included in the experiment on the ba-
³⁵⁵ sis of exhibiting complementary distribution of [g] and [ŋ] in monomorphemes. Therefore, the
³⁵⁶ model included the monomorphemes used in the dialect questionnaire to screen participants for
³⁵⁷ inclusion in the experiment, including frequency-based scaling of their faithfulness violations.
³⁵⁸ Since we assume lexicon optimization (i.e., non-alternating monomorphemes are restructured to
³⁵⁹ be /ŋ/-ful), and the monomorphemes we surveyed are only a small subset of the lexicon that ex-
³⁶⁰hibits the complementary distribution of [g] and [ŋ] and so do not allow us to train phonotactic
³⁶¹ learning models that rely on implicit negative evidence (Hayes and Wilson, 2008), we cannot ac-
³⁶²curately assess the weight of *INTERNAL-[g]. However, we can find a lower bound on its weight
³⁶³ by constraining the sets of weights we consider to those that maximize the likelihood of the com-
³⁶⁴pound data, while simultaneously preserving allophony in monomorphemes (operationalized as
³⁶⁵ having 95% or greater probability of faithful realization). The final model yielded weights listed
³⁶⁶ in Table 9, and predictions plotted in Figure 10.

³⁶⁷ The weights of the two faithfulness constraints were not significantly different from one an-
³⁶⁸other, as assessed via a likelihood ratio test: $\Delta\text{log-likelihood} = 1.3$, $p = 0.10$; a similar conclusion
³⁶⁹ was suggested by the near-zero difference in the sample-size corrected AIC of the two models:
³⁷⁰ $\Delta\text{AICc} = 1.8$. AICc differences greater than 10 are typically taken to indicate strong support for
³⁷¹ the model with the lower AICc value; for more on model-comparison in statistical models and
³⁷² phonological grammars, see Shih (2017) and Wilson and Obdeyn (2009). This result suggests that
³⁷³ the attractive influence of both bases is critical in driving the alternation in attested forms; the
³⁷⁴ zero weight of the markedness constraint *INTERNAL-[g] indicates that in existing compounds,
³⁷⁵ analogical faithfulness is doing all the work, despite the assumption in the literature that the
³⁷⁶ alternation is markedness-driven. We will revisit the role of markedness below in section 6.2.

³⁷⁷ We also compared the full model to one where the two faithfulness constraints were allowed
³⁷⁸ to take on different values but were not scaled by frequency. As one might expect, since low- and
³⁷⁹ high-frequency forms have the same violation profiles in the phonological grammar, a grammar
³⁸⁰ without access to frequency information can only predict one rate of nasalization across all forms;

| Constraint | Weight |
|---------------------|--------|
| *INTERNAL-[g] | 0.0 |
| ID-[nasal]-COMPOUND | 7.09 |
| ID-[nasal]-N2 | 7.39 |

Figure 9: Best-fitting weights for the experimental data, existing and novel compounds combined, that preserves the allophony in monomorphemes.

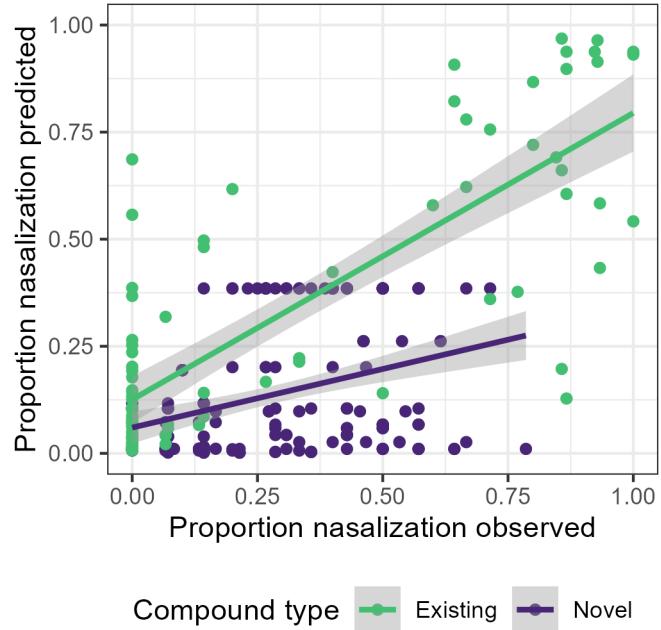


Figure 10: Predicted (vertical axis) vs. observed (horizontal axis) rates of nasalization for categories existing (green) and novel (purple) compounds under the combined model (weights in Table 9).

381 this model fits the data dramatically less well ($\Delta \log\text{-likelihood} = 264.57$, $p < .001$ with one degree
382 of freedom, $\Delta \text{AICc} = 526.50$).

383 Finally, we evaluate the absolute performance of the model by examining how well it fits
384 the data it was trained on: although the two models have different internal structures, we can
385 ask whether the theoretically-informed MaxEnt model here does as good a job in explaining the
386 data patterns as the theory-neutral mixed effects logistic regression model reported by BKK.⁹ We
387 do this using the measure of R^2 , which ranges from zero to one, and can be thought of as the
388 proportion of the variation in the dependent variable (here, whether nasalization applies or not)
389 explained by the collection of independent variables (the phonological and lexical characteristics
390 of interest).

391 We used the *r2_bayes()* function from the *performance* package (Lüdecke et al., 2021) to obtain
392 the marginal R^2 of the statistical model—that is, the amount of variance in the data explained
393 by the fixed effects—and compared it to the R^2 for the MaxEnt model.¹⁰ Since the statistical

⁹The model specification in BKK was as follows: Nasalization ~ 1 + LogN2Freq*N2Primed + LogCompoundFreq + NasalFinalN1 + (1 + LogN2Freq*N2Primed + LogCompoundFreq + NasalFinalN1 | Speaker) + (1 + N2Primed | Compound); see BKK section 3.3 and 3.4.1 for details.

¹⁰We used marginal R^2 , which makes reference to fixed effects only, since the conditional R^2 that takes into account the variance explained by both fixed and random effects has no direct comparison in the MaxEnt model

394 model is Bayesian, we obtain a median and 95% Credible Interval for our R^2 : 0.48 and [0.31, 0.56],
395 respectively. This is lower, though still relatively comparable, to the MaxEnt models R^2 of 0.63, for
396 which we have only a point estimate. Although the two are relatively close, the point value for
397 the marginal R^2 of the MaxEnt model is outside the 95% Credible Interval of the statistical model;
398 this comparison suggests that the theoretically-structured model out-performs the theory-blind
399 statistical one. While we find this result to be encouraging, this conclusion is tentative, however—
400 since the MaxEnt model does not capture variation at the level of the speaker, it may be that the
401 non-hierarchical structure of the model mismatches the structure in the data in a way that distorts
402 the results, attributing to the population grammar variance that should more conservatively be
403 attributed to speaker-level idiosyncrasies.

404 5.2 Novel compounds

405 We take advantage of the fact that we have data for both existing and novel forms to administer
406 a more severe test of the model. We do this by asking how well the grammar that was fit to the
407 existing items generalizes to the novel forms. We evaluate the probability of the two candidate
408 outcomes in the novel forms using the learned weights reported in Figure 9, with the relevant
409 frequency information for the N2s, and zero frequency for compounds (since they are novel). The
410 fit to the data is shown in Figure 10, alongside the fit to the existing compounds.

411 We found that the model of the existing forms generalizes to the novel forms quite poorly; the
412 obvious problem is that while the range of attested proportions of nasalization range between 0
413 and 0.79, the model predicts outcomes only in the range of 0.002-0.39. This indicates that either the
414 grammar that best fits the existing compounds is a poor estimate of the knowledge that speakers
415 used when generalizing to novel compounds (due to incompatible weights, constraints, or both),
416 or that the novel compound data is simply extremely variable. To check whether the mis-fit is due
417 to incompatible constraint weights, we fit a model with the same constraints to the novel forms
418 directly, without regard for data from monomorphemes or existing forms. This yields a shifted
419 range of predicted proportions (0.29-0.70), but a only marginally lower R^2 (0.11, compared to 0.12
420 based on the grammar fit to only the existing forms), which indicates that the data are still poorly
421 fit. Therefore, the unexpected finding about the faithfulness-driven nature of the alternation is
422 not to blame for the poor generalization performance of the model.

423 Next, we compare the fit of our theoretically-motivated MaxEnt model to the purely statistical
424 model fit by BKK¹¹, and find that the R^2 of our primary model when generalizing to the novel

we fit. For more on the relationship between mixed effects models and hierarchical structures in linguistic data, see Zymet (2019).

¹¹Model specification: Nasalization ~ 1 + LogN2Freq*N2Primed + (1 + LogN2Freq*N2Primed | Speaker) + (1 + N2Primed | Compound); see BKK section 3.3 and 3.4.2 for details.

425 compounds, 0.11, falls within the 95% Credible Interval of the median of that of the statistical
426 model, 0.06 [0.00, 0.18]. The model also predicts a range of 0.07-0.64 in terms of proportion
427 undergoing nasalization, which more closely matches the true data. While still low in absolute
428 terms, it lessens the possibility that our theoretical commitments are what is limiting us in being
429 able to account for the data well. Therefore, we suspect that the cause of the poor model fit may
430 be that there is simply less signal in the novel compound data.

431 6 Discussion

432 This paper has proposed a model of variable voiced velar nasalization in Japanese, drawing on ex-
433 perimental data published in Breiss et al. (*to appear*). The model integrates grammatical and func-
434 tional determinants of variation, drawing on the Voting Bases framework of lexicon-grammar
435 interaction, which was originally developed to model an entirely separate phonological phe-
436 nomenon, Lexical Conservatism in English and Spanish (Breiss, 2024). Here, we address several
437 major issues that the model raises, notably about whether the proposed system can be learned
438 from the actual Japanese lexicon (section 6.1), about the unexpected (lack of) role markedness
439 appears to play in driving the alternation (section 6.2), about the competence-performance dis-
440 tinction (section 6.3), and about how the Voting Bases model’s mechanism for integrating usage
441 frequency and formal grammar compares to other propositions in the literature (section 6.4). Fi-
442 nally, we close the paper with a more general discussion about how we might understand the
443 broader empirical landscape of frequency effects in phonological patterning in light of the pro-
444 posal in this paper.

445 6.1 Whence the weights? Evidence in the lexicon

446 Having observed that there is robust frequency-conditioning of nasalization in both existing and
447 novel compounds, we can ask what the source of this frequency-conditioning might be. A sen-
448 sible null hypothesis would be that the relationship between frequency and resting activation is
449 one that is automatic and not overtly learned. However, we find that the model performs signif-
450 icantly better when allowed to set the weights of faithfulness constraints referencing different
451 allomorphs to different weights. This result suggests that, setting aside the relationship between
452 frequency and activation, the speakers must be able to attribute different amounts of influence
453 to different faithfulness constraint violations depending on which base the violation is assessed
454 against. Put another way, the learner needs to be able to figure out how analogically-driven
455 her lexicon is. Here, we present a preliminary investigation of what kind of evidence might
456 exist in the Japanese lexicon that could allow speakers to assign different weights to ID-[nasal]-
457 COMPOUND and ID-[nasal]-N2.

458 We fit a grammar with the constraints in section 4.3 and frequency-driven scaling of faithfulness
459 violations to the set of compounds in the corpus analyzed by Breiss et al. (2021b) that had
460 a free N2. We found that the optimal weights of the grammar were zero for both *INTERNAL-[g]
461 and ID-[nasal]-COMPOUND, and 1.08 for ID-[nasal]-N2. We had anticipated there being little to
462 no weight assigned to the markedness constraint in this dataset for the same reasons discussed
463 above in section 5, but we also found that instead of a tension between faithfulness to the com-
464 pound itself and faithfulness to the N2, the grammar instead left it to the paradigm uniformity
465 effect alone to perturb the otherwise at-chance distribution of variation (at chance because the
466 weight of ID-[nasal]-COMPOUND was at zero, indicating, all else equal, that the alternating and
467 non-alternating candidates were equiprobable). This is qualitatively the same finding as for the
468 novel compounds.

469 We compared the model fit to the corpus data to one where the grammar was forced to assign
470 the same weight to ID-[nasal]-COMPOUND and ID-[nasal]-N2, and found that it was significantly
471 out-performed by the model that allowed the grammar to allot differing weights to different
472 faithfulness constraints to different bases ($\Delta\log\text{-likelihood} = 45.3$, $p < .001$ with one degree of
473 freedom). We take this as tentative evidence that there is an empirical basis in the lexicon for
474 assigning different degrees of faithfulness to different bases.

475 6.2 On the role of markedness

476 We began our discussion of voiced velar nasalization by reviewing various sources that have as-
477 sumed that the alternation observed in monomorphemes is a byproduct of a word-level marked-
478 ness constraint banning word-medial /g/. This is a typologically-common scenario, and is built
479 quite deeply into the foundations of constraint-based models (cf. Prince and Smolensky (1993),
480 and the more recent summary in Chong (2017)). Separating marked structures from their repair
481 makes it possible to derive both alternations and phonotactic restrictions from a common source.
482 This, in turn, helps resolve the “duplication problem” (Kenstowicz and Kisseeberth, 1977).

483 However, the weight of evidence drawn from BKK’s data to this point suggest that rather
484 than being driven by markedness, the VVN may instead be driven by competing faithfulness
485 pressures. Evidence comes from the zero weight assigned to the markedness constraint *INTER-
486 NAL-[g] in the model fit to the existing data in section 5, as well as the zero weight assigned to
487 the same constraint when fitting the data from the corpus, and also when trying to model the
488 novel N2 data directly. In both these scenarios, however, faithfulness constraints to both /g/-ful
489 and /ŋ/-ful allomorphs received nonzero weight, allowing the data to nevertheless be accounted
490 for. Only in a model that assumes no scaling of faithfulness constraints by resting activation
491 does markedness get weight, underscoring the importance of jointly modeling usage-based and

492 grammatical influences on probabilistic phonology (see 6.3 directly below).¹²

493 Further, though more indirect, evidence that the weight of the markedness constraint may
494 be in decline comes from the general pattern of change in many Japanese dialects, including
495 the spoken style of the Tokyo dialect, which has almost entirely lost the allophony in favor of
496 retaining /g/ as [g] in all contexts. This fact does not bear directly on the actual formal model we
497 propose, but it suggests that something in the learning data—be it phonetic, phonological, or
498 otherwise—is contributing to the loss of the allophony and the markedness constraint behind it,
499 that is common to many dialects which is precipitating the loss of the allophony and its driving
500 markedness constraint.

501 Although this type of faithfulness-driven alternation is unexpected based on the literature
502 reviewed in section 2, nevertheless the Voting Bases model predicts these outcomes should occur,
503 as shown in Figure 8.

504 **6.3 Competence, performance, and formal modeling**

505 This paper has proposed a model of Japanese nasalization that integrates token frequency into the
506 workings of the phonological grammar. Since the prospect of integrating a putatively performance-
507 related factor like token frequency into a formal phonological model is not an uncontroversial
508 one, below we directly address some possible criticisms of this approach. We certainly do not
509 think that these are the last words on the topic, but we do feel that by explicitly discussing what
510 we are doing and our motivations for doing it, we take a first step towards a clearer understanding
511 of the stakes and consequences of the choices made in modeling information about usage jointly
512 with the phonological grammar.

513 One initial objection to formally modeling the frequency-conditioned variation in nasalization
514 might be that there is nothing competence-related to model here at all—the variation is solely
515 driven by “performance” factors (Chomsky, 1965). We respond that this cannot be true of Japanese
516 nasalization: the fact that only compounds whose N2 is morphologically free exhibit frequency-
517 sensitive variation, despite the existence of bound morphemes with [g]- and [ŋ]-initial forms like
518 [ga]/[ŋa] “fang”, as shown by the examples in (5), requires an explanation that makes reference
519 to grammatical structures.

520 Further afield, cases like Lexical Conservatism much more strongly blur the line between the
521 contents of the lexicon and the phonological grammar and are well-modeled by a framework
522 like Voting Bases. The fact that this paper demonstrates both paradigm uniformity and Lexical
523 Conservatism emerge as special cases of the same theory speaks to the theoretical insight that
524 can be gained by jointly modeling “performance-related” and “competence-related” influences on

¹²In such a model, there is also weight given to a faithfulness constraint indexed to a /g/-ful UR for the compound, following the intuition of a reviewer.

525 the phonological grammar.

526 Another objection is that by incorporating both resting activation (a psycholinguistic construct) and phonological markedness (a grammatical one), the model blurs the line between competence and performance, raising the question of what exactly the model is modeling. If so, this
527 would be a legitimate concern. However, a virtue of the Voting Bases model is that lexical influence
528 on the grammar is clearly delimited: the model only allows the lexicon to scale the weights
529 of faithfulness constraints to corresponding lexical representations. Manipulating the resting
530 activation of a given UR has identifiable, localized influences on the computations of the phonological
531 grammar, and instantiates a linking hypothesis consistent with a consensus view of the
532 basic structure of the lexicon. This mechanism can be seen as one way of implementing the idea
533 of “grammar dominance” put forth, for example, by Coetzee (2016) and Coetzee and Kawahara
534 (2013). The “core” phonological grammar—weighted constraints which can assess violations of
535 novel candidates—can be recovered by simply ignoring the influence of the lexicon on constraint
536 violations, and can be studied in novel contexts like *wug*-tests, where there is no relevant lexical
537 representation to bear on the grammar.
538

539 A final objection that we consider is that the very act of jointly modeling usage frequency
540 and the phonological grammar risks leading the analyst to think of fundamentally performance-
541 related factors as in fact competence-related, thus undercutting the goal of researchers whose
542 focus is only understanding linguistic competence. We contend that this is simply false, and in
543 fact, the reverse is true: for a researcher who *only* cares about linguistic competence, modeling
544 usage factors jointly with theories of competence is vital. When confronting data derived from
545 language use (that is, modeling corpus data as in Breiss et al. (2021a), or experimental data where
546 stimuli are existing morphemes of the language as in Breiss et al. (*to appear*)), a joint model
547 will better expose the true influence of competence-related factors on the data under study, with
548 the performance-related parts of the model accounting for the otherwise-distorting influence of
549 these factors. Simply ignoring performance-related factors in a formal model makes the strong
550 claim that they have no effect, an assumption which is untenable in the cases examined here,
551 and, we suggest, is also false in many (if not all) types of linguistic data that speakers might have
552 prior usage-based experience with (Arnon and Snider, 2010; Smith and Moore-Cantwell, 2017;
553 Zymet, 2018; Morgan and Levy, 2016, 2023). Rather, an integrated approach that jointly models
554 grammar and usage is essential to disentangle and distill and understanding of competence from
555 its entanglement with performance factors, if this is the goal of the analysis.

556 The foregoing discussion, as well as comments from reviewers, raise the question of whether
557 the analysis proposed here still cleaves to the generative roots of the constraint-based model for-
558 malism that it adopts (though cf. Smolensky (1986a); Legendre et al. (1990); Smolensky (1986b) on
559 the shared roots of Optimality Theory, Harmonic Grammar, MaxEnt, and connectionism (Rumel-

561 hart et al., 1986)). This, in our opinion, is somewhat a matter of perspective, and is in any case
562 rather beside the point. Depending on how one defines “generative” or “functionalist”, our model
563 may be seen as aligned with either point of view — since it, too, aims to model grammar, its use,
564 and acquisition at a certain necessary level of abstraction. What we hope this exercise demon-
565 strates, rather, is that by reifying our theories about what the data-generating process is in a
566 computational model, we can confront complex data with many interlocking or moving parts,
567 and recover transportable analytical insights that we are confident are common desiderata shared
568 by many strands of linguistic analysis. We also note that we are far from the first to pursue this
569 approach — for very closely-related discussions of what it means for a linguistic theory to model
570 frequency, see Coetzee and Kawahara (2013); Coetzee (2016), among others.

571 **6.4 Comparison with other models**

572 The Voting Bases model is one of several approaches in the literature that propose to model the
573 interaction of usage frequency and phonological grammar. In particular, it is similar to the meth-
574 ods proposed in Coetzee (2016) and Coetzee and Kawahara (2013) which directly scale the weight
575 of faithfulness constraints by the frequency of the form they make reference to, and that of Baird
576 (2021) where a simulated perception-production loop comes to the same result via online learn-
577 ing. This family of approaches involves lowering the weight of faithfulness constraints to high-
578 frequency forms relative to lower-frequency forms which enables them to model data like coronal
579 stop deletion in English (Coetzee and Kawahara, 2013), where higher-frequency monomorphemes
580 (like *just*) tend to get produced more often with a deleted coronal stop than phonologically-similar
581 words (like *jest*). Common to these models is that they assume that the underlying form is /t/-ful,
582 and thus the task of their model must relate higher frequency to therefore have lower constraint
583 weights for it.

584 A weakness of these models is that, with the possible exception of Baird (2021), the direction-
585 ality between frequency and constraint weight is arbitrary—the primary goal set in these studies
586 was to fit the data, which is better than the alternative which does not model the effects of lex-
587 ical frequencies at all, but they suffered somewhat for the lack of clear functional grounding the
588 relation.

589 By contrast, the frequency-faithfulness relation that Voting Bases model adopts runs in the op-
590 posite direction—more frequent forms exact a greater penalty for unfaithful realizations relative
591 to less frequent forms; constraint violations are less severe for low-frequency vs. high-frequency
592 forms. This allows the model to fit a similar range of data, but with a linking hypothesis that
593 is explicitly rooted in resting activation, a construct that is externally justified by a large body
594 of work in psycholinguistics, as reviewed in Breiss (2021, 2024). Lexical items with higher rest-
595 ing activation are more insistent on faithfulness to themselves, corresponding to their increased

596 salience in the language processing system. The main contribution of the Voting Bases model
597 in modeling this phenomenon is that the influence of the lexicon on the grammar should be, in
598 principle, derivable without reference to any facts about the experiment in question; given some
599 independently-established computationally-implemented model of lexical dynamics that repre-
600 sents a scalar quantity of resting activation (or similar construct), the strong prediction of the
601 Voting Bases model is that that quantity should be able to be a fully adequate scaling factor for
602 faithfulness constraint violations. The specific mechanism that is used in this paper—scaling the
603 weights by the sigmoidal transformation of the resting activation—is used since it represents,
604 to us, a reasonable first stab, but the linking function may need to be revised in light of future
605 findings.

606 In summary, we suggest that the Voting Bases model, because of its functional grounding
607 of frequency effects in externally-motivated psycholinguistic phenomena, is on firmer footing
608 than theories that have alternative linking functions between frequency and grammar, which are
609 arguably arbitrary.

610 6.5 Towards a unified picture of token frequency in phonology

611 In this section, we broaden our view of token frequency effects in phonology, and discuss how
612 considering the varying functional roles of frequency can reconcile some seemingly-contradictory
613 bodies of evidence (cf. also Bybee, 2003).

614 First, there is evidence that higher token-frequency leads to more markedness-reducing al-
615 ternations. Coetze and Kawahara (2013) found that higher-frequency lexical items were more
616 likely to undergo phonological processes of simplification and (markedness-)reduction: high-
617 frequency English words like *jus(t)* underwent an optional process of coronal stop deletion at a
618 higher rate than low-frequency words like *jes(t)*, and high-frequency Japanese words like [baggu]
619 “bag” underwent geminate devoicing more often than low-frequency words like [budda] “Bud-
620 dha” (Kawahara and Sano, 2013). Zuraw (2007) examines frequency-conditioned application of
621 markedness-reducing phonological processes in a corpus of written Tagalog, and likewise finds
622 higher rates of repair within higher-frequency units (words, clitic groups, etc), subject to the
623 markedness principles of the language.

624 On the other hand, there is also evidence to show that higher-frequency forms are more
625 likely to be exceptional, and thus marked with regard to the overall properties of the grammar.
626 Smith and Moore-Cantwell (2017) found that higher-frequency comparative constructions are
627 more likely to flout grammar-wide trends driven by markedness. In a similar vein, Anttila (2006)
628 and Mayer (2021) found that higher-frequency morphologically-complex forms were more likely
629 to behave opaquely with respect to grammar-wide phonological processes.

630 We can compare these effects to the ones observed in Breiss et al. (2021b) and Breiss et al.

631 (to appear): higher-frequency N2s act as stronger attractors, yielding *more* faithfulness to their
632 preserved surface [g] resulting in lower rates of nasalization, whereas higher compound fre-
633 quency as a whole yielded higher rates of nasalization. Thus it seems that for compounds, higher
634 frequency is correlated with more phonological-process application and markedness-reduction;
635 this is broadly in line with the findings of Coetze and Kawahara (2013) where higher-frequency
636 words undergo more phonological alternations. However, we found that at the same time, in
637 compounds with free N2s, higher free N2 frequency is related to *less* process application, with
638 higher-frequency supporting the retention of a marked structure (word-medial [g]).

639 We suggest we can resolve this tension by distinguishing between the processes that token
640 frequency can impact: one is whether to set up an independent lexical representation for a surface
641 allomorph, and the other is influencing the strength of that representation in the lexicon of the
642 speaker.

643 If a form is exceptional and high-frequency, it may be more economical for a speaker to pay a
644 one-time “cost” of encoding the exception as a listed form that is not derived by the grammar, thus
645 relieving the phonology of the difficulty of having to generate the exceptional or idiosyncratic
646 form on each of the many frequent occasions of use (cf. Adaptor Grammars (Johnson et al., 2007,
647 *et seq.*) or Fragment Grammars (O’Donnell, 2015) which offer computationally-explicit imple-
648 mentations of this general idea). For lower-frequency exceptional forms, the likelihood of listing
649 is less since the price trades off less favorably with the amount of times it is used; thus lower-
650 frequency forms are more susceptible to change and regularization to the dominant grammatical
651 trends over time compared to higher-frequency forms.

652 Another aspect of this trade-off is the emergence of Lexicon Optimization (Prince and Smolen-
653 sky, 1993; Sanders, 2003, 2006); even if a form is not particularly exceptional, if a UR almost
654 always surfaces with a phonological process applied to it, with sufficient frequency it becomes
655 less costly to just store the form with phonological process applied—that is, create a separate
656 allomorph that is specific to the environment that would trigger the phonological rule. This, sim-
657 ilarly, relieves the grammar of the job of having to repair the form every time. Thus, we find
658 Lexicon Optimization targeting forms like *jus(t)* over forms like *jest*, making these forms restruc-
659 tured to automatically have the phonological alternation applied, thus giving the appearance of
660 having undergone a markedness-improving repair in the grammar, but actually the frequency of
661 the form has resulted in restructuring to the lexicon (see Breiss and Wilson (2020) for an initial
662 attempt at a computational model of the phonological grammar and lexicon that exhibits this
663 property).

664 As reviewed above, lexical frequency also influences the resting activation of a lexical item
665 once it is listed in the lexicon. In the Voting Bases model, higher resting activation leads to
666 the listed form exerting a stronger pull on the surface realization of a related form; where this

667 pressure goes against the broader principle of markedness in the grammar, as in cases of paradigm
668 uniformity, we find that marked structures with high-frequency output-bases are preserved; in
669 cases where the listed form coincides with the output of the markedness-reducing process, as
670 in many cases of Lexical Conservatism (Steriade, 1997; Steriade and Stanton, 2020; Breiss, 2021),
671 then the higher-frequency form promotes an unmarked surface form.

672 Recent work by Jarosz et al. (2024) has laid out a class of models which exhibit characteristics
673 that align favorably with the dynamics of frequency laid out here, suggesting that an integrated,
674 implemented model that can jointly account for the variety of frequency effects reviewed in this
675 section is perhaps quite close at hand. Future work may profitably explore how well these mod-
676 els can provide converging evidence from computational learning simulations to support the
677 psycholinguistic, experimental, and diachronic evidence for the contents of the lexicon that the
678 Voting Bases theory relies on. In sum, the broader landscape of token frequency in phonology
679 is compatible with the functional grounding given to frequency under the Voting Base model,
680 though much empirical and formal work remains to be done to further support the predictions of
681 the framework more broadly as a candidate for a general theory of the influence of the dynamic
682 lexicon on the probabilistic grammar.

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