

Homework 3

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Question 4.2

Problem

(a)

Show that, under random sampling and the zero conditional mean assumption $\mathbb{E}(u|\mathbf{x}) = 0$, $\mathbb{E}(\hat{\beta}|\mathbf{X}) = \beta$ if $\mathbf{X}\mathbf{X}'$ is nonsingular. (Hint: use property CE.5 in the appendix of chapter 2)

(b)

In addition to the assumption from part a, assume that $\text{var}(u|\mathbf{x}) = \sigma^2$. Show that $\text{var}(\hat{\beta}|\mathbf{X}) = \sigma^2(\mathbf{X}\mathbf{X}')^{-1}$

Solutions

(a)

Proof. The OLS estimator is given by $\hat{\beta} = (X'X)^{-1}X'y$ and the linear model gives us that $y = X\beta + u$. If we substitute that into the formula for $\hat{\beta}$, we get $\hat{\beta} = (X'X)^{-1}X'X\beta + (X'X)^{-1}X'u$. Next, taking the conditional expectation of $\hat{\beta}$ given X and using the zero mean assumption, we can show that $\mathbb{E}(\hat{\beta}|X) = (X'X)^{-1}X\beta + (X'X)^{-1}X'0 \rightarrow \mathbb{E}(\hat{\beta}|X) = (X'X)^{-1}X'X\beta \rightarrow \mathbb{E}(\hat{\beta}|X) = \beta$ \square

(b)

Proof. The formula for the variance of the OLS estimator $\hat{\beta} = \mathbb{E}[(\hat{\beta} - \mathbb{E}(\hat{\beta}|X))(\hat{\beta} - \mathbb{E}(\hat{\beta}|X))'|X']$. Using the result from (a), we can say that $\text{var}(\hat{\beta}|X) = \mathbb{E}[(X'X)^{-1}X'u((X'X)^{-1}X'u)'|X] = (X'X)^{-1}X'\mathbb{E}(uu'|X)X(X'X)^{-1}$. As the problem tells us that $\text{var}(u|X) = \sigma^2$, we can state $\mathbb{E}(uu'|X) = \sigma^2 I$. Thus, $\text{var}(\hat{\beta}|X) = (X'X)^{-1}X'(\sigma^2 I)X(X'X)^{-1} = \sigma^2(X'X)^{-1}$ \square

Question 4.3

Problem

Suppose that in the linear model 4.5, $\mathbb{E}(\mathbf{x}'u) = \mathbf{0}$ (where \mathbf{x} contains unity), $\text{var}(u|\mathbf{x}) = \sigma^2$, but $\mathbb{E}(u|\mathbf{x}) \neq \mathbb{E}(u)$.

(a)

Is it true that $\mathbb{E}(u^2|\mathbf{x}) = \sigma^2$?

(b)

What relevance does part a have for OLS estimation?

Solutions

(a)

No, this is not true.

Proof. $\text{var}(u|x) = \mathbb{E}(u^2|x) - \mathbb{E}(u|x)^2$. From the question, we can say that $\sigma^2 = \mathbb{E}(u^2|x) - \mathbb{E}(u|x)^2$. Due to the fact that $\mathbb{E}(u|x) \neq \mathbb{E}(u)$, we know that u varies with x , and the same with $\mathbb{E}(u|x)^2$. So, rearranging the equation, $\mathbb{E}(u^2|x) = \sigma^2 + \mathbb{E}(u|x)^2$. Since, as stated, $\mathbb{E}(u|x)^2$ varies with x , so too must $\mathbb{E}(u^2|x)$. With that conclusion we can say that $\mathbb{E}(u^2|x) \neq \sigma^2$ \square

(b)

When heteroskedasticity is present, the usual OLS S.E. will be incorrect due to the assumption of homoskedasticity. Thus, robust S.E. will be needed to produce meaningful estimation with OLS.

Question 4.17

Problem

Consider the standard linear model $y = \mathbf{x}\boldsymbol{\beta} + \mathbf{u}$ under Assumptions OLS.1 and OLS.2. Define $h(\mathbf{x}) \equiv \mathbb{E}[u^2|\mathbf{x}]$. Let $\hat{\boldsymbol{\beta}}$ be the OLS estimator, and show that we can always write:

$$\text{Avar}\sqrt{N}(\hat{\boldsymbol{\beta}} - \boldsymbol{\beta}) = [\mathbb{E}[\mathbf{x}'\mathbf{x}]]^{-1} \mathbb{E}[h(\mathbf{x})\mathbf{x}'\mathbf{x}] [\mathbb{E}[\mathbf{x}'\mathbf{x}]]^{-1} \quad (1)$$

This expression is useful when $\mathbb{E}[u|\mathbf{x}] = \mathbf{0}$ for comparing the asymptotic variances of OLS and weighted least squares estimators; see, for example, Wooldridge (1994b).

Question 5.1

Problem

In this problem you are to establish the algebraic equivalence between 2SLS and OLS estimation of an equation containing an additional regressor. Although the result is completely general, for simplicity consider a model with a single (suspected) endogenous variable:

$$\begin{aligned} y_1 &= z_1\delta_1 + \alpha_1 y_2 + u_1, \\ y_2 &= z\pi_2 + v_2. \end{aligned}$$

For notational clarity, we use y_2 as the suspected endogenous variable and z as the vector of all exogenous variables. The second equation is the reduced form for y_2 . Assume that z has at least one more element than z_1 . We know that one estimator of (δ_1, α_1) is the 2SLS estimator using instruments x . Consider an alternative estimator of (δ_1, α_1) : (a) estimate the reduced form by OLS, and save the residuals \hat{v}_2 ; (b) estimate the following equation by OLS:

$$y_1 = z_1\delta_1 + \alpha_1 y_2 + \rho\hat{v}_2 + \text{error}. \quad (5.52)$$

Show that the OLS estimates of δ_1 and α_1 from this regression are identical to the 2SLS estimators. (Hint: Use the partitioned regression algebra of OLS. In particular, if $\hat{y} = x_1\beta_1 + x_2\beta_2$ is an OLS regression, β_1 can be obtained by first regressing x_1 on x_2 , getting the residuals, say \hat{x}_1 , and then regressing y on \hat{x}_1 ; see, for example, Davidson and MacKinnon (1993, Section 1.4). You must also use the fact that z_1 and \hat{v}_2 are orthogonal in the sample.)

Question 5.2

Problem

Consider a model for the health of an individual:

$$\begin{aligned} \text{health} = & \beta_0 + \beta_1 \text{age} + \beta_2 \text{weight} + \beta_3 \text{height} \\ & + \beta_4 \text{male} + \beta_5 \text{work} + \beta_6 \text{exercise} + u_1, \end{aligned} \quad (2)$$

where *health* is some quantitative measure of the person's health; *age*, *weight*, *height*, and *male* are self-explanatory; *work* is weekly hours worked; and *exercise* is the hours of exercise per week.

Parts

(a)

Why might you be concerned about *exercise* being correlated with the error term u_1 ?

(b)

Suppose you can collect data on two additional variables, *disthome* and *distwork*, the distances from home and from work to the nearest health club or gym. Discuss whether these are likely to be uncorrelated with u_1 .

(c)

Now assume that *disthome* and *distwork* are in fact uncorrelated with u_1 , as are all variables in equation (2) with the exception of *exercise*. Write down the reduced form for *exercise*, and state the conditions under which the parameters of equation (2) are identified.

(d)

How can the identification assumption in part c be tested?

Question 5.11

Problem

A model with a single endogenous explanatory variable can be written as

$$y_1 = z_1\delta_1 + \alpha_1 y_2 + u_1, \quad (3)$$

$$E(z'u_1) = 0, \quad (4)$$

where $z = (z_1, z_2)$. Consider the following two-step method, intended to mimic 2SLS:

(a)

Regress y_2 on z_2 , and obtain fitted values, \hat{y}_2 . (That is, z_1 is omitted from the first-stage regression.)

(b)

Regress y_1 on z_1, \hat{y}_2 to obtain $\hat{\delta}_1$ and $\hat{\alpha}_1$. Show that $\hat{\delta}_1$ and $\hat{\alpha}_1$ are generally inconsistent. When would $\hat{\delta}_1$ and $\hat{\alpha}_1$ be consistent? (Hint: Let y_2^0 be the population linear projection of y_2 on z_2 , and let a_2 be the projection error: $y_2^0 = z_2\lambda_2 + a_2$, $E(z'a_2) = 0$. For simplicity, pretend that λ_2 is known rather than estimated; that is, assume that \hat{y}_2 is actually y_2^0 . Then, write)

$$y_1 = z_1\delta_1 + \alpha_1 y_2^0 + \alpha_1 a_2 + u_1 \quad (5)$$

and check whether the composite error $\alpha_1 a_2 + u_1$ is uncorrelated with the explanatory variables.