

Implicit differentiation

$$f(x, y) = 0 \quad \text{find } \frac{\partial y}{\partial x}.$$

$$\text{Let } u = f(x, y)$$

$$\frac{\partial u}{\partial x} = \frac{\partial f}{\partial x} \frac{\partial x}{\partial x} + \frac{\partial f}{\partial y} \frac{\partial y}{\partial x} = 0$$

$$\frac{\partial y}{\partial x} = -\frac{\partial f}{\partial y} / \frac{\partial f}{\partial x} = -\frac{f_x}{f_y}$$

Transformation partial differential equation

One-dimension wave equation

$$\frac{\partial^2 u}{\partial t^2} = c^2 \frac{\partial^2 u}{\partial x^2}$$

Surface

$$\text{Curve } r(t). \quad f(x, y, z) = f(x(t), y(t), z(t)) = 0 \quad \frac{\partial f}{\partial t} = 0$$

$$\frac{\partial f}{\partial t} = \frac{\partial f}{\partial x} \cdot \frac{\partial x}{\partial t} + \frac{\partial f}{\partial y} \cdot \frac{\partial y}{\partial t} + \frac{\partial f}{\partial z} \cdot \frac{\partial z}{\partial t} = 0$$

Scalar product.

$$(\frac{\partial f}{\partial x} i + \frac{\partial f}{\partial y} j + \frac{\partial f}{\partial z} k) \cdot (\frac{\partial x}{\partial t} i + \frac{\partial y}{\partial t} j + \frac{\partial z}{\partial t} k) = 0$$

Define a vector  $\nabla F$ .

$$\nabla F = (\frac{\partial F}{\partial x} i + \frac{\partial F}{\partial y} j + \frac{\partial F}{\partial z} k)$$

gradient.

$$|\nabla F| \parallel \left| \frac{\partial h}{\partial t} \right| = 0 \rightarrow \text{tangent vector}$$

$\nabla F$  is a normal vector to  $S$  at  $p$ .

Tangent plane : one plane through  $P$  perpendicular to  $\nabla F$ .

Summary

## 5

Surfaces in parametric coordinates

Surface  $r(u, v)$ .

$r_u \times r_v$  is a normal vector to  $S$  at  $P$

(cross product)

Taylor theorem  $f(x, y)$

$$f(a+h, b+k) = \sum_{n=0}^{\infty} \frac{1}{n!} (h \frac{\partial}{\partial x} + k \frac{\partial}{\partial y})^n f(x, y) \Big|_{a, b} + R_n$$

Remainder:  $R_n = \frac{1}{(n+1)!} (h \frac{\partial}{\partial x} + k \frac{\partial}{\partial y})^{n+1} f(x, y) \Big|_{a+th, b+tk}$   
 $(0 \leq t \leq 1)$

$$f(a+h, b+k) = f(a, b) + (hf_x + kf_y) + \frac{1}{2} (h^2 f_{xx} + 2hk f_{xy} + k^2 f_{yy})$$

Linear approximation

$$f(x, y) \approx f(a, b) + (x-a)f_x(a, b) + (y-b)f_y(a, b) = L(x, y)$$

$Z = L(x, y)$  is linear in  $x, y$  and is a plane

↳ tangent plane

Quadratic Approximation

$$f(x, y) \approx f + (x-a)f_x + (y-b)f_y + \frac{1}{2} [(x-a)^2 f_{xx} + 2(x-a)(y-b)f_{xy} + (y-b)^2 f_{yy}] = Q(x, y)$$

Stationary point

Definition:  $\boxed{f_x = f_y = 0}$ ,  $(a, b)$  is a stationary point of  $f(x, y)$

Summary

Local Maximum and Local minima

maximum if  $f(x,y) \leq f(a,b)$

minima if  $f(x,y) \geq f(a,b)$

Saddle point

A saddle point is  
a stationary point  
which is not a maximum  
nor a local minimum

$(a,b)$  is a saddle point of  $f(x,y)$  if there is a  
point  $(x_1, y_1)$  in  $D$  where  $f(x_1, y_1) > f(a,b)$  and a  
point  $(x_2, y_2)$  in  $D$  where  $f(x_2, y_2) < f(a,b)$

### Hessian

$$H = f_{xx}f_{yy} - f_{xy}^2$$

If  $H < 0$ ,  $f$  has a saddle point

If  $H > 0$ , if  $f_{xx} > 0$ ,  $f$  has a minimum at  $(a,b)$

If  $f_{xx} < 0$ ,  $f$  has a maximum at  $(a,b)$

$H = 0$ , no conclusion.

Lagrange multiplier

$$f(x,y, \lambda) = f(x,y) - \lambda g(x,y) \quad \nabla f = \lambda \nabla g(x,y)$$

$$\text{solve } F_x = F_y = F_\lambda = 0$$

Generalisation

$$F(x_1, \dots, x_n, \lambda_1, \dots, \lambda_m) = f - \sum_{k=1}^m \lambda_k g_k$$

Double integral

$$\iint_R f(x,y) dx dy$$

$$A = \iint_R 1 dx dy \quad \text{Area of } R$$

Summary

Average and centroids

$$\bar{f} = \frac{1}{n} \sum_{i=1}^n f_i$$

$$\frac{\sum_{i=1}^n f_i w_i}{\sum_{i=1}^n w_i}$$

$$\bar{f} = \frac{\iint_R f(x,y) dx dy}{\iint_R dx dy}$$

$$\bar{x} = \frac{\iint_R x dx dy}{\iint_R dx dy} \quad \bar{y} = \frac{\iint_R y dx dy}{\iint_R dx dy}$$

$\bar{x}, \bar{y}$  centroid.

Change of variable and Jacobian

$$x = x(u,v) \quad y = y(u,v)$$
$$\frac{\partial(x,y)}{\partial(u,v)} = \left( \frac{\partial x}{\partial u} \frac{\partial y}{\partial v} - \frac{\partial x}{\partial v} \frac{\partial y}{\partial u} \right)$$
$$= \begin{vmatrix} \frac{\partial x}{\partial u} & \frac{\partial x}{\partial v} \\ \frac{\partial y}{\partial u} & \frac{\partial y}{\partial v} \end{vmatrix}$$

$$J = \iint_S f(x(u,v), y(u,v)) \left| \frac{\partial(x,y)}{\partial(u,v)} \right| du dv$$

Useful Property.

$$\frac{\partial(u,v)}{\partial(x,y)} = 1 / \frac{\partial(x,y)}{\partial(u,v)}$$

Summary

5

Triple integral

$$\iiint_V f(x, y, z) dx dy dz$$

$$\text{Volume} = \iiint_D dV$$

$$J = \frac{\partial(x, y, z)}{\partial(u, v, w)} = \begin{vmatrix} \frac{\partial x}{\partial u} & \frac{\partial x}{\partial v} & \frac{\partial x}{\partial w} \\ \frac{\partial y}{\partial u} & \frac{\partial y}{\partial v} & \frac{\partial y}{\partial w} \\ \frac{\partial z}{\partial u} & \frac{\partial z}{\partial v} & \frac{\partial z}{\partial w} \end{vmatrix}$$

given  $x = x(u, v, w)$ ,  $y = y(u, v, w)$ ,  $z = z(u, v, w)$

### Integration Techniques

#### O2 Method:

$$\text{Theorem: } \int f(x) g(x) dx = \sum_{j=0}^{n-1} (-1)^j f^{(j)}(x) g^{(-j+1)}(x) + (-1)^n \int f^{(n)}(x) g(x) dx$$

Corollary: Suppose that  $f^{(n)}(x) \equiv 0$

$$\int f(x) g(x) dx = \sum_{j=0}^{n-1} (-1)^j f^{(j)}(x) g^{(-j+1)}(x) + C$$

$$\begin{array}{c}
 D \quad L \\
 f \quad + \quad g \quad + - + - + - \\
 f^{(1)} \quad \searrow \quad g^{(-1)} \\
 f^{(2)} \quad \searrow \quad g^{(-2)} \\
 \vdots \\
 D = f^{(n)} \quad g^{-n}
 \end{array}$$

Summary

5

$$\text{eg: } \int x^3 \cos x \, dx$$

$$\begin{array}{cc}
 D & I \\
 x^3 & + \cos x \\
 3x^2 & - \sin x \\
 6x & -\cos x \\
 6 & + \sin x \\
 0 & \cos x
 \end{array}$$

$$\int x^3 \cos x \, dx = x^3 \sin x - 3x^2(-\cos x) + 6x(-\sin x) - 6 \cos x + C.$$

$$\int_a^b f(x) \, dx = \int_a^c f(x) \, dx + \int_c^b f(x) \, dx$$

$$\int_1^\infty e^{-[x]} \, dx = \int_1^\infty e^{-1} \, dx + \int_2^\infty e^{-2} \, dx + \dots +$$

$$= \int_1^\infty e^{-1} \, dx + \int_2^\infty e^{-2} \, dx + \int_3^\infty e^{-3} \, dx + \dots$$

$$= e^{-1} \int_1^\infty 1 \, dx + e^{-2} \int_2^\infty 1 \, dx$$

$$= e^{-1} + e^{-2} + e^{-3} + \dots$$

$$\lim_{n \rightarrow \infty} \frac{e^{-1} + e^{-2} + \dots + e^{-n}}{1 - e^{-1}} = \frac{e^{-1}}{1 - e^{-1}} = \frac{1}{e-1}$$

Advanced substitution

$$I = \int \frac{2\sin x - \cos x}{\sin x + 2\cos x} \, dx$$

$$= A \int \frac{\sin x + 2\cos x}{\sin x + 2\cos x} \, dx + B \int \frac{\cos x - 2\sin x}{\sin x + 2\cos x} \, dx$$

$$2\sin x - \cos x = A(\sin x + 2\cos x) + B(\cos x - 2\sin x)$$

$$(A - 2B)\sin x + (2A + B)\cos x$$

$$\begin{cases} A - 2B = 2 \\ 2A + B = -1 \end{cases} \quad \begin{cases} A = 0 \\ B = -1 \end{cases}$$

Summary

Useful facts:

$$\lim_{x \rightarrow 0} \frac{\sin x}{x} = 1$$

$$\lim_{x \rightarrow 0} \frac{e^x - 1}{x} = 1$$

$$\lim_{x \rightarrow 0} \frac{1 - \cos x}{x^2} = \frac{1}{2}$$

$$\lim_{x \rightarrow a} \frac{x^n - a^n}{x - a} = na^{n-1}$$

$$\lim_{x \rightarrow 0} \frac{\ln(1+x)}{x} = 1$$

$$\lim_{x \rightarrow \infty} \left(1 + \frac{1}{x}\right)^x = e$$

Trapezium Rule with error

$$\int_a^b f(x) dx \approx \frac{h}{2} [f_0 + 2f_1 + \dots + f_{n-1}] + f_n]$$

$$h = (b-a)/n$$

$$|E_T| \leq \frac{M(b-a)^3}{12n^2} \quad M = \max_{a \leq x \leq b} |f''(x)|$$

$n^{\text{th}}$  derivative and stationary point

Theorem: lowest nonvanishing derivative of  $f$  at stationary point  $x_0$  is  $f^{(k)}(x_0)$ ,  $k \geq 2$ ,  $f^{(k+1)}(x_0)$  exist.

i)  $k$  is even  $f^{(k)}(x_0) < 0 \Rightarrow x_0$  is maximum

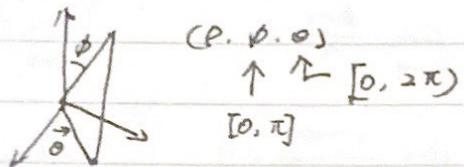
ii)  $k$  is even  $f^{(k)}(x_0) > 0 \Rightarrow x_0$  is minimum

iii)  $k$  is odd  $\Rightarrow x_0$  is a point of inflection.

Summary

注意顺序  
课堂上和试卷写的  
是法是  $(\rho, \theta, \phi)$

Spherical coordinates



Q:  $(0, 2\sqrt{3}, -2)$  in rectangular coordinate.

$$\left\{ \begin{array}{l} \rho = \sqrt{x^2 + y^2 + z^2} \\ \text{using } \cos\theta = x \\ \text{Polar } \sin\theta = y \\ \rho \cos\phi = z \end{array} \right. \quad \phi \in [0, \pi]$$

$$\left\{ \begin{array}{l} \rho = \sqrt{0^2 + (2\sqrt{3})^2 + (-2)^2} \\ \cos\theta = \frac{-2}{\rho} = -\frac{1}{2} \end{array} \right. \quad \theta \in [0, \pi)$$

$$\rho \cos\phi = -2$$

$$\rho = \sqrt{0^2 + (2\sqrt{3})^2 + (-2)^2} = 4$$

$$2\sqrt{3} \rightarrow \cos\theta = \frac{-2}{4} = -\frac{1}{2} \quad \theta = \frac{2\pi}{3}$$

$$2\sqrt{3} = 4 \times \frac{\sqrt{3}}{2} \times \sin\theta. \quad \sin\theta = 1 \quad \theta = \frac{\pi}{2}$$

$$(4, \frac{2\pi}{3}, \frac{\pi}{2})$$

注意代数式的导致  
几次求导之后会为0  
的情况

右边的可拆看

Leibniz's Theorem

$$f^{(n)} = \sum_{r=0}^n C_r n^r v^{n-r}$$

$$y'' - 2y' + y = e^x + e^{-x}$$

通解:  $y_{cp} = Ae^x + Be^{-x}$ .

特解:  $\left\{ \begin{array}{l} e^x \rightarrow \text{重根} \quad y_{p1} = Ax^2 e^x \\ e^{-x} \rightarrow \text{不是根} \quad y_{p2} = Be^{-x} \end{array} \right.$

$$\begin{aligned} y &= y_{cf} + y_{p1} + y_{p2} \\ &= Ae^x + Be^{-x} + \frac{1}{2}x^2 e^x + \frac{1}{4}Be^{-x} \end{aligned}$$

Summary

formal definition to get partial derivative

$$f(x,y) = \begin{cases} \frac{x^3+x^4-y^3}{x^2+y^2} & \text{if } (x,y) \neq (0,0) \\ 0 & \text{if } (x,y) = (0,0) \end{cases}$$

$$\frac{\partial f}{\partial x}(0,0) = \lim_{h \rightarrow 0} \frac{f(0+h,0) - f(0,0)}{h}$$

$$\frac{\partial f}{\partial y}(0,0) = \lim_{h \rightarrow 0} \frac{f(0,0+h) - f(0,0)}{h}$$

想计算  $\frac{\partial f}{\partial x}(0,0)$  及  $\frac{\partial f}{\partial y}(0,0)$  用公式

但想求  $\frac{\partial f}{\partial x}(0,0)$ , 必须用到  $f(0,0) = 0$

$$f(h,0) = \frac{h^3 + h^4 - 0^3}{h^2 + 0^2} = \frac{h^3 + h^4}{h^2} = h + h^2$$

$$\begin{aligned} \frac{\partial f}{\partial x}(0,0) &= \lim_{h \rightarrow 0} \frac{f(0+h,0) - f(0,0)}{h} = \lim_{h \rightarrow 0} \frac{f(h,0) - f(0,0)}{h} \\ &= \lim_{h \rightarrow 0} \frac{h + h^2 - 0}{h} \\ &= \lim_{h \rightarrow 0} (1+h) \\ &= 1 \end{aligned}$$

$$\text{Hessian} = f_{xx}f_{yy} - f_{xy}^2$$

$\boxed{H < 0}$ , saddle point

$H > 0$ ,  $f_{xx} < 0$  local maxim

$f_{xx} > 0$  local mini

$H = 0$  no conclusion.

Summary

5

Lagrange multipliers 共同  
代回原函数找最值.

二重积分互换  $dx dy$ . 积分方向也要换.

$$\text{so } \int_0^1 \int_{\sqrt{y^2}}^1 e^{y^2} dy dx$$

$$\int_0^1 \int_0^{y^2} e^{y^2} dx dy \text{ 最好画个图}$$

三重积分.  $x, y, z$  分别在 3D 图和 2D 中寻找积分域

- 最后一个积分元变量
- from inside to outside.

Jacobian 算进去的是绝对值  $|J|$ .

画 domain 时, 如果是圆, 指半径和圆心.

判断 方程形状,  $let z=0$ . 俯视图.  $x, y=0$ . 侧视图

$$eq. z = \sqrt{x^2+y^2}. \text{ cone. 圆锥.}$$

Taylor 的误差

$$R_n = \frac{1}{(n+1)!} (h \frac{\partial}{\partial x} + k \frac{\partial}{\partial y})^{n+1} f(x, y) \Big|_{x=t, y=s} \cdot \\ a+th, b+th \quad (0 \leq t \leq 1)$$

Lagrange 算距离记得最后开平方

Summary

积分常数不要忘记

换元，积分域也要变，记得代回原值

Euler homogeneous

$$y = vx \quad \frac{dy}{dx} = \frac{dv}{dx}x + v$$

记得注明取值范围；分子不为0 etc.

三角积分，阅历里 $\alpha$ ，原来是无 $\alpha$ 的。

最后将其换回去

$$\sqrt{a^2 - x^2} \quad x = a \sin \theta$$

$$x^2 + a^2 \quad x = a \tan \theta$$

$$\sqrt{x^2 - a^2} \quad x = a \sec \theta$$

$$\sqrt{\frac{a-x}{ax}} \quad x = a \cos \theta$$

$$\int \sec x dx = \int \sec x \cdot \left( \frac{\sec x + \tan x}{\sec x + \tan x} \right) dx \\ = \ln |\sec x + \tan x| + C$$

$$\int \csc x dx = \int \csc x \cdot \left( \frac{\csc x - \cot x}{\csc x - \cot x} \right) dx \\ = \ln |\csc x - \cot x| + C$$

$$\int \frac{1}{\sqrt{x^2 + a^2}} dx = \ln |x + \sqrt{x^2 + a^2}| + C$$

Summary

5

$$\frac{1}{(x+a)(x+b)} = \frac{A}{x+a} + \frac{B}{x+b}$$

$$\frac{1}{(x+a)^2(x+b)} = \frac{A}{x+a} + \frac{B}{(x+a)^2} + \frac{C}{(x+b)}$$

$$\frac{1}{(x^2+a)(x+b)} = \frac{Ax+B}{x^2+a} + \frac{C}{x+b}$$

form

$$\frac{1}{a+b\cos x + c\sin x} \quad \frac{1}{a+b\cos x} \quad \frac{1}{a+b\sin x}$$

$$\tan(\frac{x}{2}) = \pm \quad \sin x = \frac{2t}{1+t^2}$$

$$dx = \frac{2dt}{1+t^2} \quad \cos x = \frac{1-t^2}{1+t^2}$$

$$\tan x = \frac{2t}{1-t^2}$$

$$\int p(x) f(x) dx$$

1. 2. Method.

$$\downarrow \quad + \quad \downarrow \quad + - + - - .$$

$$\int \frac{1}{x^2+a^2} dx = \frac{1}{a} \tan^{-1}(\frac{x}{a}) + C$$

$$\int \frac{1}{\sqrt{a^2-x^2}} dx = \sin^{-1}(\frac{x}{a}) + C$$

$$\int \frac{1}{|x|\sqrt{x^2-a^2}} = \frac{1}{a} \sec^{-1}(\frac{|x|}{a}) + C$$

$$\int \frac{1}{\sqrt{x^2-a^2}} = \ln |x + \sqrt{x^2-a^2}| + C$$

Summary

5

$$\int \frac{1}{x^2-a} dx = \frac{1}{2a} \ln \left| \frac{x-a}{x+a} \right| + C$$

$$\int \frac{1}{a^2-x^2} dx = \frac{1}{2a} \ln \left| \frac{x+a}{x-a} \right| + C$$

$$e^x = \sum_{n=0}^{\infty} \frac{x^n}{n!} = 1 + x + \frac{x^2}{2!} + \frac{x^3}{3!} + \dots$$

$$\cos x = \sum_{n=0}^{\infty} \frac{(-1)^n x^{2n}}{(2n)!} = 1 - \frac{x^2}{2!} + \frac{x^4}{4!} - \frac{x^6}{6!} + \dots$$

$$\sin x = \sum_{n=0}^{\infty} \frac{(-1)^n x^{2n+1}}{(2n+1)!} = x - \frac{x^3}{3!} + \frac{x^5}{5!} - \frac{x^7}{7!} + \dots$$

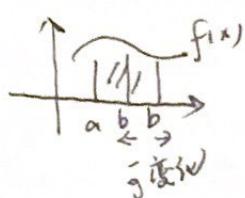
$$\frac{1}{1+x} = \sum_{n=0}^{\infty} (-1)^n x^n = 1 - x + x^2 - x^3 + x^4 - \dots$$

$$\ln(1+x) = \sum_{n=1}^{\infty} \frac{(-1)^{n+1} x^n}{n} = x - \frac{x^2}{2} + \frac{x^3}{3} - \frac{x^4}{4} + \dots$$

可整体代换

Summary

5



度量积分.

$$\int_a^b f(x) dx = F(b) - F(a)$$

而积是函数的函数

积分值与和分变量无关.

$$F(x) = \int_a^x f(t) dt$$

便于区分

$$F(x) = \int_a^x f(t) dt = F(x) - F(a)$$

$$(F(x))' = [F(x) - F(a)]' = f(x) = f(x)$$

推导

$$F(x) = \int_a^{x^2} f(t) dt = F(x^2) - F(a)$$

$$(F(x))' = [F(x^2)]' = 2x f(x^2)$$

- 例题:

$$H(x) = \int_{\varphi_1(x)}^{\varphi_2(x)} f(t) dt$$

求  $H'(x)$ 

$$H'(x) = (\varphi_2'(x)f(\varphi_2(x))) - (\varphi_1'(x)f(\varphi_1(x)))$$

$$\text{eg. } \frac{d}{dx} \left( \int_{3x}^{x^2} \sin t dt \right)$$

$$= 2x \sin x^2 - 3 \sin 3x$$

Summary

5

## Statistics Summary:

Central limit theorem:

If  $X_1, \dots, X_n$  are independent and identically distributed, with same mean  $\mu$  and final variance  $\sigma^2$ , then  $\bar{X}$  is approximately  $N(\mu, \frac{\sigma^2}{n})$  for large  $n$  no matter what the distribution of the  $X_i$ .

泊松的期望方差相等

Poisson distribution:  $p\{X=k\} = \frac{\lambda^k}{k!} e^{-\lambda}$ ,  $k=0, 1, 2, \dots$

$$E(X) = \lambda$$

$$\text{Var}(X) = \lambda$$

Exponential distribution:  $f_X(x) = \begin{cases} \lambda e^{-\lambda x} & x \geq 0 \\ 0 & x < 0 \end{cases}$

$$f_X(x) = \begin{cases} 1 - e^{-\lambda x} & x \geq 0 \\ 0 & x < 0 \end{cases}$$

$$E(X) = \frac{1}{\lambda}$$

$$\text{Var}(X) = \frac{1}{\lambda^2}$$

significance & p-value & power

(Type I)

$P$  (Obtaining a test statistic at least  $H_0$  is true as extreme as that observed)

$1 - \beta$

• significance level of a hypothesis  $\alpha = P(\text{reject } H_0 \mid H_0 \text{ true})$

• p-value is the probability of observing an outcome "more extreme" than what was actually observed given  $H_0$  is true

• power of a hypothesis test is  $P(\text{reject } H_0 \mid H_0 \text{ false})$ .

Summary

5

Assumptions of least square linear regression

$$\text{(i)} \quad \mathbb{E}(\epsilon_i) = 0 \quad \text{for all } i=1, \dots, n$$

the error has zero mean

$$\text{(ii)} \quad \text{var}(\epsilon_i) = \sigma^2 \quad \text{for all } i=1, \dots, n$$

errors have the same variances

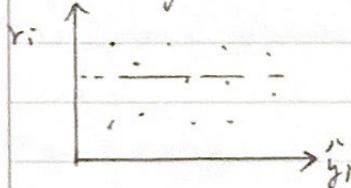
$$\text{(iii)} \quad \text{cov}(\epsilon_i, \epsilon_j) = 0 \quad \text{for all } i \neq j$$

the errors are uncorrelated

residual plot

$$r_i = y_i - \hat{y}_i$$

$r_i$  against fitted values.  $\hat{y}_i$



Contingency table

$H_0$ : independent vs  $H_1$ : not independent

Exp. 3. 列联表. 组和比例一样.

Normal Approximation to Binomial Distribution

$$n \geq 20, np \geq 5, n(1-p) \geq 5$$

$$X \approx N(np, np(1-p))$$

$$X = X_1 + \dots + X_n$$

Summary

5

Preparation for  $\chi^2$  test.

$$\chi^2_{\text{obs}} = \sum \frac{(O_i - E_i)^2}{E_i} = \sum \frac{(O_i)^2}{E_i} - n$$

(χ<sup>2</sup>检验数)

Explanation for 0.2:

If the experiment were to be repeated multiple times and a 90% confidence interval was calculated in this way for each experiment, then 90% of the confidence interval would contain the true value of the parameter.

lower the confidence  $\Rightarrow$  increasing  $\alpha$

$$\alpha \propto \frac{1}{\sqrt{n}}$$

Q-Q plot:

① straight line ② through the origin ③ gradient 1

$$\begin{array}{ll} \alpha & \downarrow \\ 1-\alpha & \uparrow \\ \downarrow \text{confidence} \end{array}$$

if  $k$  degrees of freedom is 计算整合之后的  $k$ .

形成上下的数据，计算向右看清楚的 two sample 还是 difference. 如果不看合乎假设的错误。

Normal Q-Q plot

Summary

known  $\mu, \sigma^2$  1.  $b_1, b_2, \dots, b_n$

①  $(0, \sigma)$  2. Standardize  $y_i: z_{ij} = \frac{y_{ij} - \mu}{\sigma}$  比较  $y_i$

② straight line 3. Find  $p(x_{ij}) = \frac{i - \frac{1}{2}}{n}$  (查表)

4. Plot  $z_{ij}$  against  $x_{ij}$

unknown  $\mu, \sigma^2$

1.  $b_1, b_2, \dots, b_n$

2. Find  $p(x_{ij}) = \frac{i - \frac{1}{2}}{n}$

3. Plot

centered  $\hat{y}_{ij}$  on top for interpretation to slope

George Wilkinson

$\hat{y}_{ij} \rightarrow \mu + \sigma$

5

看剖題問中有 difference in the variance  
between ... 想主) F-test.

Test for regression  $\neq$  原為 test model 的斜率

$$H_0: \beta_1 = \beta^* \text{ vs } H_1: \beta_1 \neq \beta^*$$

reject if

$$|t| = \left| \frac{\hat{\beta}_1 - \beta^*}{\sqrt{\frac{1}{n-2} s_{\hat{\beta}_1}^2}} \right| > t_{n-2, \alpha/2}$$

題中可能對  $\beta^*$  有個表述: increasing by 1 ...

$\chi^2$  ~ test 12式推導

$$\begin{aligned} \sum_i \frac{(o_i - e_i)^2}{e_i} &= \sum_i \left( \frac{o_i^2}{e_i} - 2o_i + e_i \right) = \sum_i \frac{o_i^2}{e_i} - 2 \sum_i o_i + \sum_i e_i \\ &= \sum_i \frac{o_i^2}{e_i} - 2n + n \\ &= \sum_i \frac{o_i^2}{e_i} - n \rightarrow \text{這步數理格子數} \end{aligned}$$

核心是  $\chi^2$  之  $o_i, e_i$  都是頻數. 總和為  $n$

p-value 定義:

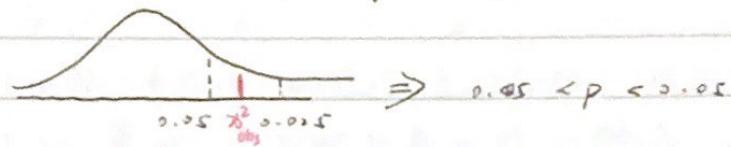
is the probability of observing a test statistic  
more extreme than observed assuming the null  
hypothesis is true.]

$p\text{-value} = p \left( \text{obtaining a test statistic more extreme than observed} \mid H_0 \text{ is true} \right)$

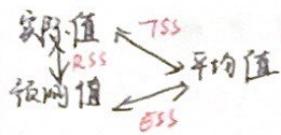
Small p-value, implies there is evidence against  
the null hypothesis being true.

Summary

因为 P 值判断是一个具体范围  
所以需要自己测试在多少  $\alpha\%$  significance level  
会或不会拒绝  $H_0$  可以得出 P 的范围.



⇒ 得出结论



$$\text{Total Sum of squares: } TSS = \sum (y_i - \bar{y})^2 = (n-1) s_y^2 \quad \text{总平方和}$$

$$\text{The residual sum of square: } RSS = \sum (y_i - \hat{y}_i)^2 = (n-2) \hat{\sigma}^2 \quad \text{残差平方和}$$

$$\text{The regression sum of square: } ESS = \sum (\hat{y}_i - \bar{y})^2 \quad \text{回归平方和}$$

(explained sum of square)

$$TSS = RSS + ESS$$

$$\begin{aligned} \sum r_i^2 &= (n-1) [s_y^2 - \hat{\sigma}^2] \quad \text{或 } \hat{\sigma}^2 = \frac{1}{n-2} \sum r_i^2 \\ &= (n-1) \cdot (n-2) \hat{\sigma}^2 / (n-1) \\ &= (n-2) \hat{\sigma}^2 \end{aligned}$$

$$\text{proportion of the variance explained is } R^2 = \frac{ESS}{TSS}$$

solve:  $R^2$  to use ←  
the proportion to  
determine the sample  
correlation between  
 $x$  and  $y$ .

## Summary

TSS: how much variance there is in the dependent variable

RSS: how much of the variation your model did not explain

ESS: how much of the variation your model did explain

5

$\chi^2$  goodness of fit / Contingency table 都是单边检验  
适用条件: valid for large sample.

Levene's test

做假设原因为  $\leftarrow$   
sample size 太小了.  
 $\neq \sigma^2$ .

two population 检验中. 给出了 SD 不对称方差公式,  
仍用假设  $\sigma^2$  相等  $\rightarrow S_p^2 = \frac{(n_1-1)s_1^2 + (n_2-1)s_2^2}{n_1+n_2-2}$



接着, 检验  $\sigma^2$  相等这个假设. 用 F-test  
这个题可谓完美过渡.

2019:

Standard deviation / appropriate no outlier, robust  
not appropriate outlier  
( $Q_1 - 1.5 \text{ IQR}, Q_3 + 1.5 \text{ IQR}$ )

机理:  
样本太小, 无法用  $\leftarrow$   
正态分布.

可以用计算机算,  
也可以用表  
但是要注意有没有  
等号

$$P(X \geq 5)$$

$$= 1 - P(X \leq 4) \quad \text{因为是} \\ = 1 - P(X \leq 4) \quad \text{离散的}$$

Summary

intercept:  $\mu$

slope:  $\sigma$ .

$y_{ij} = \mu + \sigma z_{ij}$

QQ plot 画正态分布 (未知方差和均值)

利用公式 + table

未知方差和均值，先将样本

$$T(x_i) = \frac{i - \frac{1}{2}}{n}$$

排序并计算  $T$  值以判断

$$E(x_i) = \frac{i - \frac{1}{2}}{n}$$

$\sigma$  (0.25) straight line  $\Rightarrow$  slope 1.

$x_i = E(x_i) \pm \frac{\sigma}{n}$  白色取值对称性,  $\sigma$  值是~~对称的~~

加不加负号的问题.

画出一个大致直线 估计即可.

## 检验假设检验序贯的假设管理

Test for means:  $\mu$  known (Z-test)

Assume:  $n$  is large ( $n \geq 30$ ) OR data are normally distributed

Test for means:  $\mu$  unknown (t-test)

Assume:  $x_1, \dots, x_n \sim i.i.d. N(\mu, \sigma^2)$

Two populations

这样方差  $\approx \sigma^2$

Case I: Assume  $\sigma_1^2, \sigma_2^2$  known [OR  $n_1, n_2$  large]

Case II: Assume  $\sigma_1^2, \sigma_2^2$  unknown but equal

Case III: Assume  $\sigma_1^2, \sigma_2^2$  unknown unequal.

①

If  $n_1, n_2$  large ( $n_1 \geq 30, n_2 \geq 30$ ) treat  $\sigma_1, \sigma_2$  as known

using  $s_1, s_2 \Rightarrow$  Case I

If  $\sigma_1, \sigma_2$  too different  $\Rightarrow$  Case II.

Summary

## 5

Pair samples

Assume:  $(Z_i \sim N(\mu, \sigma^2))$  independently,  
normally distributed and matched pairs

If there is a positive relationship between  
the scores in a paired situation, then  
matched pairs test has a higher power than  
the independent sample test. (i.e. has  
higher probability of correctly rejecting  $H_0$   
when  $H_0$  is true/false.)

Inference about Variance  $\sigma^2$

• Chi-square test

Assume:  $X_1, \dots, X_n \sim i.i.d. N(\mu, \sigma^2)$

• F-test

Assume:  $X_1, X_2, \dots, X_{n,x} \sim N(\mu_x, \sigma_x^2)$  } independent  
 $Y_1, Y_2, \dots, Y_{n,y} \sim N(\mu_y, \sigma_y^2)$  }

Correlation not means causality

i.e. do not imply that one has any direct or  
indirect effect on the other)

Summary

5

T-test for linear correlation

Assume:  $(x_1, y_1), \dots, (x_n, y_n)$  are normally distributed

Require  $x, y$  have a linear relationship

Non-parametric test for linear correlation

Assume: non-linear  $\Rightarrow$  no correlation  $H_0: \text{line corrl}$

Least squares linear regression

1.  $E(\epsilon_i) = 0$  i=1...n errors have zero mean

2.  $\text{cov}(\epsilon_i, \epsilon_j) = 0 \quad i \neq j$  errors are uncorrelated

3.  $\text{var}(\epsilon_i) = \sigma^2 \quad i=1...n$  have same variances

Assume: Data do not have to be normally distributed

Test for the existence of a regression  $H_0: \beta = 0$

Inference for individual / mean  $H_1: \beta \neq 0$

why individual + 1

An individual observation

will have greater

variability than a

mean value)

extra error should be

taken into account

Population two samples

Assume: two samples are independent and n<sub>1</sub>, n<sub>2</sub>

are large

Summary

5

Spearman rank order

 $H_0$ : no correlation  $H_1$ : some correlation

线性方程最小二乘法的公式推导

$$\epsilon_i = y_i - \beta_0 - \beta_1 x_i$$

$$S = \sum \epsilon_i^2 = \sum (y_i - (\beta_0 + \beta_1 x_i))^2$$

Minimize  $S$ 

$$\frac{\partial S}{\partial \beta_0} = \frac{\partial}{\partial \beta_0} \sum_{i=1}^n (y_i - \beta_0 - \beta_1 x_i)^2 = 0$$

$$\frac{\partial S}{\partial \beta_1} = \frac{\partial}{\partial \beta_1} \sum_{i=1}^n (y_i - \beta_0 - \beta_1 x_i)^2 = \sum_{i=1}^n -2x_i(y_i - (\beta_0 + \beta_1 x_i)) = 0$$

$$\begin{cases} \sum (y_i - \beta_0 - \beta_1 x_i) = 0 \Rightarrow \sum \hat{\epsilon}_i = 0 \\ \sum \epsilon_i (y_i - \beta_0 - \beta_1 x_i) = 0 \Rightarrow \sum \hat{\epsilon}_i x_i = 0 \end{cases}$$

$$\begin{cases} \sum y_i = n\beta_0 + \beta_1 \sum x_i = 0 \\ \sum x_i y_i - \beta_0 \sum x_i - \beta_1 \sum x_i^2 = 0 \end{cases}$$

$$\begin{aligned} \beta_0 &= \frac{1}{n} (\sum y_i - \beta_1 \sum x_i) \\ &= \bar{y} - \hat{\beta}_1 \bar{x} \end{aligned}$$

$$\sum x_i y_i - (\bar{y} - \hat{\beta}_1 \bar{x}) \sum x_i - \beta_1 \sum x_i^2 = 0$$

$$\beta_1 (\sum x_i^2 - \bar{x} \sum x_i) = \sum x_i y_i - \bar{y} \sum x_i$$

$$\beta_1 = \frac{\sum x_i y_i - \bar{y} \sum x_i}{\sum x_i^2 - \bar{x} \sum x_i}$$

$$= \frac{\sum x_i y_i - \frac{1}{n} \sum x_i \sum y_i}{\sum x_i^2 - \frac{1}{n} (\sum x_i)^2} = \frac{S_{xy}(n-1)}{S_x(n-1)} = \frac{S_{xy}}{S_x^2}$$

Summary